

## THE RATE OF CONVERGENCE OF A RANDOM WALK TO BROWNIAN MOTION

BY DAVID F. FRASER

*Worcester Polytechnic Institute*

This paper establishes a rate of convergence of a random walk to Brownian motion which is nearly best possible. The Skorokhod representation is employed in the proof.

**1. Introduction and summary.** Let  $x_i$  be a sequence of independent random variables with mean 0 and variance 1. Let  $s_n(t) = \sum_{i=1}^{\lfloor nt \rfloor} x_i/n^{\frac{1}{2}}$ , and  $w(t)$  be the standard Brownian motion.

**THEOREM.** *Suppose  $E(\exp |x_i|^c) \leq M < \infty$ . Then*

$$(1.1) \quad P(\max_{0 \leq t \leq 1} |s_n(t) - w(t)| > \alpha(\log n)^\beta n^{-\frac{1}{2}}) = O(n^{-q})$$

for all  $q$ , where  $\alpha = \alpha(c, q)$  and  $\beta = c^{-1} + \frac{3}{2}$ .

Rosencrantz (1967) proves that

$$|F_n(\lambda) - F(\lambda)| < A(\log n)^{\frac{1}{2}} n^{-(p-2)/2(p+1)}$$

where  $F_n(\lambda) = P(\max_{1 \leq k \leq n} |\sum_{i=1}^k x_i/n^{\frac{1}{2}}| < \lambda)$ , and  $F(\lambda) = P(\max_{0 < t < 1} |w(t)| < \lambda)$ , if  $E(|x_i|^p) < \infty$ ,  $2 < p < 4$ , and gets a Lévy rate-of-convergence theorem. Heyde (1969) obtains a rate of convergence  $A(\log n)^{\frac{1}{2}} n^{-p/4(p+1)}$  for  $p \geq 4$ , but his estimates are not sufficient for (1.1). By Theorem 2 of Sawyer (1972),

$$P(|s_n(1) - w(1)| \leq \lambda c/n^{\frac{1}{2}}) = G(\lambda),$$

it is clear that the rate  $O(n^{-\frac{1}{2}})$  cannot be improved, so a result like (1.1) is of value, since often the variables with which one is involved satisfy the hypothesis. From (1.1) one can draw the usual conclusions, namely

(i) If  $\Phi(x)$  is any functional on  $C[0, 1]$  such that  $P(\Phi(w) \leq \lambda)$  has a bounded density and  $|\Phi(x) - \Phi(y)| \leq C||x - y||$  then

$$\sup_{\lambda} |P(\Phi(s_n(1)) \leq \lambda) - P(\Phi(w(1)) \leq \lambda)| = O((\log n)^\beta n^{-\frac{1}{2}}).$$

(ii) By Lemma 1.2 of Prokhorov (1956), if  $P_n = P(s_n(\cdot) \in A)$ ,  $L(P_n, W) = O((\log n)^\beta n^{-\frac{1}{2}})$ , where  $L(\cdot, \cdot)$  is the Prokhorov metric.

**2. Establishing the result.** By means of the Skorokhod (1965) representation, we can find independent random variables  $\tau_i$  such that  $w(\sum_{i=1}^k \tau_i)$  and  $\sum_{i=1}^k x_i/n^{\frac{1}{2}}$

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have the same joint distribution. Then

$$\begin{aligned} P(\max_{0 \leq t \leq 1} |s_n(t) - w(t)| > 2\varepsilon) &\leq P(\max_{1 \leq l \leq n} |w(\sum_{i=1}^l \tau_i) - w(l/n)| > \varepsilon) \\ &\quad + P(\max_{0 \leq t < n} \max_{0 \leq t \leq 1/n} |w(l/n) - w(l/n + t)| > \varepsilon) \\ &= A + B. \end{aligned}$$

Now

$$\begin{aligned} A &\leq P(\max_{1 \leq l \leq n} \max_{|s| \leq \delta, l/n + s \geq 0} |w(l/n + s) - w(l/n)| > \varepsilon) \\ &\quad + P(\max_{1 \leq l \leq n} |\sum_{i=1}^l \tau_i - l/n| > \delta) \\ &= C + D, \end{aligned}$$

where

$$C \leq 4nP(\sup_{0 \leq s \leq \delta} w(s) > \varepsilon) \leq 4n \exp(-\varepsilon^2/2\delta)$$

and by a submartingale inequality

$$\begin{aligned} D &= P(\max_{1 \leq l \leq n} |\sum_{i=1}^l (n\tau_i - 1)| > n\delta) \\ &\leq (n\delta)^{-2p} E((\sum_{i=1}^n (n\tau_i - 1))^{2p}) \quad \text{for all } p \geq 1. \end{aligned}$$

Let  $y_i = n\tau_i - 1$ . The  $y_i$  are independent with mean 0.

$$E((\sum_{i=1}^n y_i)^{2p}) = \sum_{|\alpha|=2p} \frac{(2p)!}{\alpha_1! \alpha_2! \dots \alpha_n!} E(y_1^{\alpha_1} y_2^{\alpha_2} \dots y_n^{\alpha_n}).$$

If any  $\alpha_i = 1$ ,  $E(y^\alpha) = 0$ , and so

$$E((\sum_{i=1}^n y_i)^{2p}) = \sum_{|\alpha|=2p, n\alpha_i=1} \binom{2p}{\alpha} E(y^\alpha),$$

where

$$\begin{aligned} E(y^\alpha) &= E(y_1^{\alpha_1} \dots y_n^{\alpha_n}) \\ &\leq (E(y_1^{2p}))^{\alpha_1/2p} \dots (E(y_n^{2p}))^{\alpha_n/2p}. \end{aligned}$$

Using the estimate of Sawyer (1967, (2.6)) we get

$$\begin{aligned} E(y_i^{2p}) &\leq 2^{2p-1} (2(2p)! E(x_i^{4p}) + 1) \\ &\leq 2^{2p} ((2p)! 4pMc^{-1}\Gamma(4p/c) + 1). \end{aligned}$$

Let  $P(j, k)$  be the number of ways of putting  $j$  envelopes into  $k$  slots such that each slot gets at least two envelopes. From the trivial estimate  $P(2p, k) < k^{2p}$  we estimate

$$\sum_{|\alpha|=2p, n\alpha_i=1} \binom{2p}{\alpha} \leq pn^p p^{2p}/p!$$

Using the above estimate, we obtain

$$D \leq D' = \delta^{-2p} n^{-p} M(2p)^{4p} (4p/c\varepsilon)^{4p/c}.$$

We now set  $D' = Mn^{-q}$  and solve for  $\delta$ , with  $q > 1$ .

$$\delta = 4p^2 (4p/c\varepsilon)^{2/c} n^{-\frac{1}{2} + q/2p}.$$

Then set  $C' = 4ne^{-\varepsilon^2/8\delta} = 4n^{1-q}$  and, with  $p = \log n$ , solve for  $\varepsilon$

$$\begin{aligned}\varepsilon &= (2q\delta \log n)^{\frac{1}{2}} \\ &= 2^{\frac{1}{2}}q^{\frac{1}{2}}e^{q/4}(4/ce)^{1/c}n^{-\frac{1}{4}}(\log n)^{c^{-1}+\frac{1}{2}}.\end{aligned}$$

Term  $C$  dominates term  $B$  since  $\delta > 1/n$ . The proof of the result is now concluded. (For an alternate approach to this see Dudley (1972).)

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DEPARTMENT OF MATHEMATICS  
WORCESTER POLYTECHNIC INSTITUTE  
WORCESTER MASSACHUSETTS 01609