

## AN ASYMPTOTIC 0–1 BEHAVIOR OF GAUSSIAN PROCESSES

BY CLIFFORD QUALLS<sup>1</sup> AND HISAO WATANABE<sup>2</sup>

University of North Carolina, Chapel Hill  
University of New Mexico and Kyushu University

Let  $\{X(t), -\infty < t < \infty\}$  be a stationary Gaussian process with covariance function satisfying: (1)  $r(t) = 1 - C|t|^\alpha + o(|t|^\alpha)$  as  $t \rightarrow 0$ :  $C > 0$ ,  $0 < \alpha \leq 2$ ; and (2)  $r(t) = O(t^{-\gamma})$  as  $t \rightarrow \infty$ :  $\gamma > 0$ . Then for all positive increasing functions  $\phi(t)$  on  $[a, \infty)$ ,  $P[X(t) > \phi(t)$  infinitely often] = 0 or 1 as  $\int_a^\infty \phi(t)^{2/\alpha-1} \exp\{-\phi^2(t)/2\} dt < \infty$  or  $= \infty$ .

This result generalizes the paper of Watanabe [*Trans. Amer. Math. Soc.* **148** 233–248] by replacing his condition that  $r(t) = o(1/t)$  as  $t \rightarrow \infty$  by condition (2). Our result is extended also to the nonstationary process treated by Watanabe. Our proof treats the problem as a crossing problem using a recent result of Pickands [*Trans. Amer. Math. Soc.* **145** 51–73] and a modification of the Borel lemmas.

**0. Introduction.** Let  $\{X(t), -\infty < t < \infty\}$  be a real separable Gaussian process defined on a probability space  $(\Omega, \mathcal{A}, P)$ . We assume  $EX(t) \equiv 0$  and  $v^2(t) = EX^2(t) > 0$ . Denote the correlation function by  $\rho(s, t) = EX(s)X(t)/(v(s)v(t))$ . This paper is concerned with the probability of the event

$$E_\phi = [\exists t_0(\omega): X(t) \leq v(t)\phi(t) \quad \text{for all } t \geq t_0(\omega)].$$

One of the authors in [3] gives conditions on the correlation function so that  $PE_\phi = 1$  or 0 as  $I_\phi < \infty$  or  $= \infty$ , where the quantity  $I_\phi = \int_a^\infty \phi(t)^{2/\alpha-1} \exp\{-\phi^2(t)/2\} dt$ , and  $\alpha$  is given below. He considers the problem as a type of the so-called law of the iterated logarithm which appeared in the study of sums of independent random variables. In this paper, we treat the problem from a different point of view as a type of crossing problem. The resulting simpler proof shows the above 0–1 behavior holds for a larger class of processes, and also makes the intuitive content of the result clearer. The Gaussian processes (or rather the corresponding correlation functions) now included satisfy:

I There are positive constants  $\Delta$ ,  $C_1$ ,  $C_2$ ,  $T$ , and  $\alpha$  with  $0 < \alpha \leq 2$ , such that  $1 - C_1 h^\alpha \leq \rho(t, t+h) \leq 1 - C_2 h^\alpha$  for  $0 \leq h < \Delta$  and all  $t \geq T$ ; and

II  $\rho(t, t+s) = O(s^{-\gamma})$  uniformly in  $t$  as  $s \rightarrow \infty$  for some  $\gamma > 0$ .  
Condition II replaces the condition that  $\rho(t, t+s) = o(1/s)$  in Watanabe (1970).

---

Received February 2, 1971.

<sup>1</sup> Research supported in part by the Office of Naval Research under Contract N00014-67-A-0321-0002.

<sup>2</sup> Research supported in part by the National Science Foundation under Grant GU-2059.

*Key words and phrases:* Gaussian process, stationary process, crossing problem, a 0–1 law, a law of the iterated logarithm.

There is also a slight improvement in condition I.

Section 1 gives the proof for stationary processes making use of a recent result due to Pickands [1].

A well-known theorem of Slepian [2] is used in Section 2 to extend the results of Section 1 to nonstationary processes described by conditions I and II above. It has been pointed out in [3] that the result of Section 2 can be made to yield (by a time transformation) the analogous 0–1 behavior for a class of Gaussian processes containing Brownian motion.

**1. Stationary case.**

**THEOREM 1.1.** *Let  $\{X(t), -\infty < t < \infty\}$  be a real separable stationary Gaussian process with  $EX(t) \equiv 0$  and covariance function  $r$  satisfying*

(1)  $r(t) = 1 - C|t|^\alpha + o(|t|^\alpha)$  as  $t \rightarrow 0$  for some  $C > 0$  and some  $\alpha$  with  $0 < \alpha \leq 2$ ; and

(2)  $r(t) = O(t^{-\gamma})$  as  $t \rightarrow \infty$  for some  $\gamma > 0$ .

Then for all functions  $\phi(t)$  that are positive and nondecreasing on some interval  $[a, \infty)$ , it follows that

$$PE_\phi \equiv P[\exists t_0(\omega) > a: X(t) < \phi(t) \text{ for all } t \geq t_0(\omega)] = 1 \text{ or } 0$$

as the integral

$$I_\phi \equiv \int_a^\infty \phi(t)^{2/\alpha-1} \exp\{-\phi^2(t)/2\} dt \text{ is finite or infinite.}$$

Note that the condition (1) implies the process  $X(t)$  has continuous sample functions. The following lemma will be needed for the first half of the proof of this theorem.

**LEMMA 1.2.** *If condition (1) of Theorem 1.1 holds and  $A(t) = \inf\{(1-r(s))/|s|^\alpha: 0 < s \leq t\} > 0$ , then*

$$\lim_{x \rightarrow \infty} \frac{P[\max_{0 \leq s \leq t} X(s) > x]}{tx^{2/\alpha}\Psi(x)} = C^{1/\alpha}H_\alpha, \quad \text{where } \Psi(x) = (2\pi)^{-1/2}x^{-1}e^{-x^2/2} \text{ and}$$

$$0 < H_\alpha \equiv \lim_{T \rightarrow \infty} T^{-1} \int_0^\infty e^s P[\sup_{0 \leq t \leq T} Y(t) > s] ds < \infty,$$

and  $Y(t)$  is a nonstationary Gaussian process with mean  $EY(t) = -|t|^\alpha$  and covariance function  $r(s, t) = -|s-t|^\alpha + |s|^\alpha + |t|^\alpha$ .

**PROOF.** This is Lemma 2.9 of Pickands [1]. In addition to condition (1), Pickands required that  $A_1(t) = \inf\{(1-r^2(s))/|s|^\alpha: 0 < s \leq t\} > 0$ . However, checking Pickands' proofs, we note that this requirement can be replaced by only the requirement that  $A(t) > 0$ .

In other words, these proofs permit  $r(s) = -1$  but not  $r(s) = 1$  in the interval  $0 < s \leq t$ .

REMARK. If it should happen that  $A(t) = 0$ , then there is a smallest  $s_0 > 0$  such that  $r(s_0) = 1$  and then both  $r(t)$  and  $X(t)$  are periodic with period  $s_0$ . Since  $\max_{[0,\tau]} X(s) = \max_{[0,t]} X(s)$  where  $\tau = \min(t, s_0)$ , the requirement that  $A(t) > 0$  can be eliminated from Lemma 1.2 by replacing the denominator  $tx^{2/\alpha}\Psi(x)$  by  $\tau x^{2/\alpha}\Psi(x)$ .

PROOF OF THEOREM 1.1 WHEN  $I_\phi < \infty$ . Only condition (1) of Theorem 1.1 will be required for this half of the proof. Considering the sequence of intervals  $[n, n + 1]$  with integer end points, we obtain  $\infty > I_\phi \geq \sum_{n=N}^\infty \phi(n+1)^{2/\alpha}\Psi(\phi(n+1))(2\pi)^{\frac{1}{2}}$  where the lower limit  $a = N$  of the integral  $I_\phi$  is chosen large enough that the integrand of  $I_\phi$  is a decreasing function in the argument  $\phi$ . Define  $F_n = [\max_{n \leq s \leq n+1} X(s) \leq \phi(n)]$ .

Since by Lemma 1.2, there is a positive constant  $K$  such that

$$PF_n^c \sim K\phi(n)^{2/\alpha}\Psi(\phi(n))$$

as  $\phi(n) \rightarrow \infty$ , we obtain  $\sum_{n=N}^\infty PF_n^c < \infty$ . The Borel-Cantelli lemma completes this half of the proof.

Before proceeding to the second half of the proof, we will need the following lemmas.

LEMMA 1.3. *If condition I of Theorem 1.1 holds and  $A(t) > 0$ , then*

$$\lim_{x \rightarrow \infty} \frac{P[\max_{0 \leq k \leq m} X(kax^{-2/\alpha}) > x]}{tx^{2/\alpha}\Psi(x)} = C^{1/\alpha} \frac{H_\alpha(a)}{a},$$

where  $a > 0$ ,  $m = [tax^{-2/\alpha}]$  and  $[ ]$  denotes the greatest integer function, and  $H_\alpha(a)$  is a certain positive constant.

PROOF. This is Lemma 2.5 of Pickands [1]. All the remarks given for Lemma 1.2 also apply here.

LEMMA 1.4. *If Theorem 1.1 for the case  $I_\phi = \infty$  is true under the additional restriction that for large  $t$ ,  $2 \log t \leq \phi^2(t) \leq 3 \log t$ , then it is true without this restriction.*

PROOF. A similar statement could have been made for the case  $I_\phi < \infty$ , but it was not needed in the proof. The restriction that  $\phi^2(t) \leq 3 \log t$  is treated in Lemma 4.1 of [3], and the restriction that  $\phi^2(t) \geq 2 \log t$  is only a slight modification of Lemma 4.1 [3] when  $\alpha < 2$ . Suppose  $\alpha = 2$ . Since  $X(t) > \hat{\phi}(t)$  occurs infinitely often (as  $t \rightarrow \infty$ ) implies  $X(t) > \phi(t)$  occurs i.o. for any  $\hat{\phi} \geq \phi$ , we need only show that  $I_\phi = \infty$  implies  $I_{\hat{\phi}} = \infty$  for  $\hat{\phi} = \max(\phi, u)$  and  $u(t) = (2 \log t)^{\frac{1}{2}}$ .

Letting  $A = \{t > a: \phi(t) \leq u(t)\}$  and  $B = \{t > a: \phi(t) > u(t)\}$  we write  $I_\phi = A_\phi + B_\phi = \infty$ ,  $I_u = A_u + B_u = \infty$ , and  $I_{\hat{\phi}} = A_{\hat{\phi}} + B_{\hat{\phi}} = A_u + B_\phi$ , where for example  $B_\phi$  means  $\int_B \exp\{-\phi^2(t)/2\} dt$ . We may suppose  $B_\phi < \infty$  for otherwise the lemma follows immediately. Note that if  $t_0 \in B$ , then there is a (largest) nonempty interval in  $B$  containing  $t_0$ . Consequently,  $B$  is a union of such (disjoint) intervals  $I_n$  whose lengths and left end points will be denoted by  $\Delta_n$  and  $t_n$ . (Unfortunately, it may not be possible to index the  $t_n$ 's according to their order on the line.) However, we

assume  $\phi(t)$  crosses  $u(t)$  infinitely often as  $t \rightarrow \infty$ , and therefore that the  $\Delta_n$ 's are finite numbers and the sequence  $\{t_n\}$  is infinite. If  $\phi$  does not cross  $u(t)$  i.o., then either  $\phi \leq u$  and  $I\hat{\phi} = I_u = \infty$ , or  $\phi > u$  and  $I\hat{\phi} = I_\phi = \infty$  for some  $a$ . Note that  $\phi(t_n + \Delta_n) = u(t_n + \Delta_n)$  since the jumps of  $\phi$  are never downward. Now

$$\infty > B_\phi = \sum_n \int_{I_n} \exp\{-\phi^2/2\} dt \geq \sum_n \Delta_n \exp\{-\phi^2(t_n + \Delta_n)/2\} = \sum_n \frac{\Delta_n}{t_n + \Delta_n},$$

and

$$B_u \leq \sum_n \Delta_n \exp\{-u^2(t_n)/2\} = \sum_n \Delta_n/t_n = \sum_n \frac{\Delta_n}{t_n + \Delta_n} \left\{ \frac{1}{1 - \Delta_n/(t_n + \Delta_n)} \right\}.$$

Since  $\Delta_n/(t_n + \Delta_n) \rightarrow 0$ , we have  $\sum_n \Delta_n/t_n < \infty$ . Finally  $B_u < \infty$  implies  $A_u = \infty$  which in turn implies  $I\hat{\phi} = \infty$ .

LEMMA 1.5. *Let  $X(t)$  be a Gaussian process with zero mean function and covariance function  $r(s, t)$  with  $r(t, t) \equiv 1$ . Let  $E_n = [X(t_{n,v}) \leq x_{n,v} : v = 0, \dots, m_n]$  with all  $t_{n,v}$  distinct. Then*

$$\left| P\left(\bigcap_1^n E_k\right) - \prod_1^n P E_k \right| \leq \sum \sum_{1 \leq i < j \leq n} \sum_{\mu=0}^{m_j} \sum_{\nu=0}^{m_i} |r| \int_0^1 \phi(x_{i,\nu}, x_{j,\mu}; \lambda r) d\lambda,$$

where  $\phi(x, y; \lambda r)$  is the standard bivariate normal density with correlation coefficient  $\lambda r = \lambda r(t_{i,\nu}, t_{j,\mu})$ .

PROOF. This type of lemma now appears in many proofs of asymptotic independence for crossing problems. We include the "standard" proof with the necessary differences.

The event  $\bigcap E_k$  involves  $N = \prod_1^n (m_k + 1)$  random variables  $X(t_{n,v})$ , and the corresponding covariance matrix will be denoted by  $\sum_1 = (r_{kl})$ , where the doubly indexed random variables  $X(t_{n,v})$  have been renumbered by a single index  $k$  (the  $k$  in  $r_{kl}$ ).

Partition  $\sum_1 = [\sum_{ij}]$  so that each submatrix  $\sum_{ij}$  is the covariances of the random variables of  $E_i$  with those of  $E_j$ . Now the events  $E_k$  would be independent if and only if the corresponding covariance matrix were  $\sum_0 = [\sum_{ij}^0]$  with  $\sum_{ii}^0 = \sum_{ii}$  but  $\sum_{ij}^0 = 0$  matrix for  $i \neq j$ .

Let  $\sum_\lambda = \lambda \sum_1 + (1 - \lambda) \sum_0 = (r_{\lambda ij})$  be the covariance matrix for the standardized multivariate normal density  $\phi_\lambda(\mathbf{y})$ , and

$$F(\lambda) = \int_{-\infty}^{x_{1,0}} \dots \int_{-\infty}^{x_{1,m_1}} \dots \int_{-\infty}^{x_{n,0}} \dots \int_{-\infty}^{x_{n,m_n}} \phi_\lambda(\mathbf{y}) d\mathbf{y},$$

where  $d\mathbf{y} = dy_1 dy_2 \dots dy_N$ . We now have

$$\left| P\left(\bigcap_1^n E_k\right) - \prod_1^n P E_k \right| = |F(1) - F(0)| = \left| \int_0^1 F'(\lambda) d\lambda \right| \leq \int_0^1 |F'(\lambda)| d\lambda.$$

Since  $F'(\lambda) = \int \partial \phi_\lambda / \partial \lambda d\mathbf{y}$  and  $dr_{\lambda kl} / d\lambda = 0$  or  $r_{kl}$  according to whether  $(k, l)$  refers to a diagonal  $\sum_{kk}$  or not, we obtain by the chain rule for  $\partial \phi_\lambda / \partial \lambda$

$$|F'(\lambda)| = \left| \sum_{i < j} \sum^* r_{kl} \int \frac{\partial^2 \phi_\lambda}{\partial y_k \partial y_l} d\mathbf{y} \right| \leq \sum_{i < j} \sum^* |r_{kl}| \phi(x_k, x_l; \lambda r_{kl}).$$

The double sum  $\sum^*$  extends over all  $(k, l)$  that refer to covariances  $r_{kl}$  of  $\sum_{ij}$ . Integrating this inequality with respect to  $\lambda$  finishes the proof.

PROOF OF THEOREM 1.1 WHEN  $I_\phi = \infty$ . Note that condition (2) of Theorem 1.1 eliminates the periodic case discussed immediately following Lemma 1.2. Define a sequence of intervals by  $I_n = [n\Delta, n\Delta + \beta]$  for  $\Delta > 0$  and  $0 < \beta < \Delta$ . Let  $G_k = \{t_{k,v} = k\Delta + (v/n_k); v = 0, \dots, [ \beta n_k ]\}$  be a set of points in  $I_k$  where  $n_k$  shall be chosen later.

Let  $E_k = [\max_{s \in G_k} X(s) \leq \phi(k\Delta + \beta)]$ . Now for  $a = N\Delta$  sufficiently large  $\infty = I_\phi \leq \sum_{k=N}^\infty \Delta \phi(k\Delta)^{2/\alpha} \Psi(\phi(k\Delta)) (2\pi)^{\frac{1}{2}}$  implies that  $\sum \beta \phi(k\Delta + \beta)^{2/\alpha} \Psi(\phi(k\Delta + \beta)) = \infty$ . If we choose  $n_k = [\phi(k\Delta + \beta)^{2/\alpha}]$ , then Lemma 1.3 implies there is a positive constant  $K$  such that  $PE_k^c \sim K\beta \phi(k\Delta + \beta)^{2/\alpha} \Psi(\phi(k\Delta + \beta))$  as  $\phi(k\Delta + \beta) \rightarrow \infty$ . So we have  $\sum PE_k^c = \infty$ .

As in the Borel lemma, the main step is

$$1 - P[E_k^c \text{ i.o.}] = \lim_{m \rightarrow \infty} [\prod_m^\infty PE_k + \lim_{m \rightarrow \infty} \{P(\bigcap_m^\infty E_k) - \prod_m^\infty PE_k\}].$$

The first limit is zero because  $\sum PE_k^c = \infty$ , and the second limit will be zero because of the asymptotic independence of the events  $E_k$ . Note the separation between  $I_k$ 's is  $\Delta - \beta$ . By Lemma 1.5, we have

$$A_{m,n} = |P(\bigcap_m^n E_k) - \prod_m^n PE_k| \leq \sum \sum_{m \leq i < j \leq n} \sum_{\mu=0}^{[\beta n_j]} \sum_{v=0}^{[\beta n_i]} |r| \int_0^1 g(\phi(i\Delta + \beta), \phi(j\Delta + \beta); \lambda r) d\lambda,$$

where  $r = r(t_{j,v} - t_{i,\mu})$  and  $g(x, y; \lambda p)$  denotes the bivariate normal density. Because  $t_{j,v} - t_{i,\mu} \geq j\Delta - i\Delta - \beta \geq \Delta - \beta$ , and because of condition (2) of Theorem 1.1,  $\Delta$  can be chosen large enough that  $|r(t_{j,v} - t_{i,\mu})| \leq M_0(j\Delta - i\Delta - \beta)^{-\gamma}$  for all  $j > i \geq m$  and some positive constant  $M_0$  and such that  $|r| < \delta \equiv \min(\frac{1}{3}, \gamma/6)$ . In fact, for  $M = M_0(1 - \beta/\Delta)^{-\gamma}$ , we have  $|r| \leq M(j\Delta - i\Delta)^{-\gamma}$ . Now by Lemma 1.4, we can choose  $m$  large enough that  $\phi^2(k\Delta + \beta) \geq u^2(k\Delta + \beta) = 2 \log(k\Delta + \beta)$  and  $\phi^2(k\Delta + \beta) \leq w^2(k\Delta + \beta) = 3 \log(k\Delta + \beta)$  for all  $k \geq m$ . We obtain

$$\begin{aligned} &g(\phi(i\Delta + \beta), \phi(j\Delta + \beta); \lambda r) \\ &\leq (2\pi)^{-1} (1 - \delta^2)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} [\phi^2(i\Delta + \beta) - 2|r|\phi(i\Delta + \beta)\phi(j\Delta + \beta) + \phi^2(j\Delta + \beta)] \right\} \\ &\leq (2\pi)^{-1} (1 - \delta^2)^{-\frac{1}{2}} \exp \left\{ -\frac{1}{2} [u^2(i\Delta + \beta) - 2|r|w(i\Delta + \beta)w(j\Delta + \beta) + u^2(j\Delta + \beta)] \right\} \\ &\leq (2\pi)^{-1} (1 - \delta^2)^{-\frac{1}{2}} \left( \frac{1}{i\Delta + \beta} \right) \left( \frac{1}{j\Delta + \beta} \right)^{1 - 3|r|}. \end{aligned}$$

By Lemma 1.4,  $n_i = [\phi(i\Delta + \beta)^{2/\alpha}] \leq (3 \log(i\Delta + \beta))^{1/\alpha}$ , and since  $w(i\Delta + \beta) \leq w(j\Delta + \beta)$  and  $|r| < \delta$ , we have

$$A_{m,\infty} \leq K \sum_{m \leq i < j < \infty} \frac{\log(j\Delta + \beta)^{2/\alpha}}{(j\Delta - i\Delta)^\gamma} \left( \frac{1}{i\Delta + \beta} \right) \left( \frac{1}{j\Delta + \beta} \right)^{1 - 3\delta}$$

This double series can be seen to be convergent by letting  $k = j - i$  and obtaining

$$\begin{aligned} A_{m,\infty} &\leq K' \sum_{i=m}^{\infty} \sum_{k=1}^{\infty} \frac{(\log(k\Delta + i\Delta + \beta))^{2/\alpha}}{k^{6\delta}} \left(\frac{1}{i}\right) \left(\frac{1}{k+i}\right)^{1-3\delta} \\ &\leq K' \sum_{i=m}^{\infty} \sum_{k=1}^{\infty} \frac{(\log(k\Delta + i\Delta + \beta))^{2/\alpha}}{(k+i)^\delta} \left(\frac{1}{i}\right)^{1+\delta} \left(\frac{1}{k}\right)^{1+\delta} < \infty. \end{aligned}$$

Since this series is convergent, we have  $\lim_{m \rightarrow \infty} A_{m,\infty} = 0$  where  $A_{m,\infty} = |P(\bigcap_m^\infty E_k) - \prod_m^\infty PE_k|$ . This completes the proof.

**2. Non-stationary case.** We now extend the result for the stationary case to the nonstationary case treated in [3]. Here, let  $\{X(t), -\infty < t < \infty\}$  be a real separable Gaussian process with zero mean function.

We assume  $v^2(t) = EX^2(t) > 0$  and denote the covariance function of  $X(t)/v(t)$  by  $\rho(s, t) = EX(s)X(t)/(v(s)v(t))$ . We shall assume also  $\rho(s, t)$  is continuous.

**THEOREM 2.1.** *Suppose the above process  $X(t)$  satisfies*

(3) *there are positive constants  $\delta_1, C_1, T_1$  and  $\alpha$  with  $0 < \alpha \leq 2$  such that  $\rho(t, t+h) \geq 1 - C_1 h^\alpha$  for  $0 < h < \delta_1$  and all  $t > T_1$ . Then for all functions  $\phi(t)$  that are positive and nondecreasing on some interval  $[a, \infty)$  such that*

$$I_\phi \equiv \int_a^\infty \phi(t)^{2/\alpha-1} \exp\{-\phi^2(t)/2\} dt < \infty,$$

*we have*

$$P[\exists t_0(\omega) > a: X(t) \leq v(t)\phi(t) \text{ for all } t \geq t_0(\omega)] = 1.$$

This theorem is Theorem 1 of [3], but we include the following different proof.

**PROOF.** Let  $Y(t)$  be a separable stationary Gaussian process having a covariance function  $q(h)$  satisfying  $q(h) = 1 - C_1^* \cdot |h|^\alpha + o(|h|^\alpha)$  as  $h \rightarrow 0$  and  $q(h) \leq 1 - C_1 h^\alpha \leq \rho(t, t+h)$  for  $0 < h < \delta_1^*$  and  $t \geq T_1$ . This second requirement follows for  $C_1^* > C_1$ . By the nonstationary version of a well-known result due to Slepian (1962) (see Theorem 1), the fact that  $q(h) \leq \rho(t, t+h)$  for all  $t \geq T_1$  implies

$$P[\sup_{(n\Delta, n\Delta + \Delta)} Y(s) \leq u] \leq P[\sup_{(n\Delta, n\Delta + \Delta)} X(s) \leq u]$$

for  $\Delta < \delta_1^*$  and  $n\Delta > T_1$ . Now following the ‘‘proof of Theorem 1.1 when  $I_\phi < \infty$ ,’’ we have  $\sum PG_n^c \leq \sum PF_n^c < \infty$ , where  $F_n = [\sup_{(n\Delta, n\Delta + \Delta)} Y(s) \leq \phi(n\Delta)]$  and  $G_n = [\sup_{(n\Delta, n\Delta + \Delta)} X(s) \leq \phi(n\Delta)]$ . The Borel-Cantelli lemma applied to the  $G_n$ 's completes the proof.

**THEOREM 2.2.** *Let the above process  $X(t)$  have a correlation function satisfying:*

(3') *There are positive constants  $\delta_2, C_2, T_2$  and  $\alpha'$  with  $0 < \alpha' \leq 2$  such that  $\rho(t, t+h) \leq 1 - C_2 h^{\alpha'}$  for  $0 < h < \delta_2$  and all  $t > T_2$ ; and*

(4)  *$\rho(t, t+s) = O(s^{-\gamma})$  uniformly in  $t$  as  $s \rightarrow \infty$  for some  $\gamma > 0$ .*

Then for all functions  $\phi$  as in Theorem 2.1 with  $I_\phi = \infty$ , we have

$$P[X(t) > v(t)\phi(t) \text{ i.o. in } t] = 1.$$

PROOF. Let  $Y(t)$  be a separable stationary Gaussian process having a covariance function  $q(h)$  satisfying  $q(h) = 1 - C_2^* |h|^{\alpha'} + o(|h|^{\alpha'})$  as  $h \rightarrow 0$  and  $\rho(t, t+h) \leq 1 - C_2 h^{\alpha'} \leq q(h)$  for  $0 < h < \delta_2^*$  ( $C_2^* < C_2$ ). Applying Slepian's result (see Theorem 1) and the "proof of Theorem 1.1 when  $I_\phi = \infty$ ," we obtain  $\infty = \sum P E_n^c \leq \sum P H_n^c$  for  $\beta < \delta_2^*$ , where  $E_k = [\max_{s \in G_k} Y(s) \leq \phi(k\Delta + \beta)]$  and  $H_k = [\max_{s \in G_k} X(s) \leq \phi(k\Delta + \beta)]$ . Consequently, it only remains to show that the events  $H_n$  are asymptotically independent (i.e. in the nonstationary case). Lemma 1.5 yields as in the "proof of Theorem 1.1 when  $I_\phi = \infty$ "

$$A_{m,n} = \sum \sum_{m \leq i < j \leq n} \sum_{\mu=0}^{[\beta n_j]} \sum_{\nu=0}^{[\beta n_i]} |\rho| \int_0^1 g(\phi(i\Delta + \beta), \phi(j\Delta + \beta); \lambda \rho) d\lambda,$$

where  $\rho = \rho(t_{i,\nu}, t_{j,\mu})$  and  $g(x, y; \lambda \rho)$  denotes the bivariate normal density. Again we chose  $\Delta$  large enough that  $|\rho| = |\rho(t_{i,\nu}, t_{j,\mu})| \leq M(j-i)^{-\gamma}$  and  $|\rho| < \delta \equiv \min(1/3, \gamma/6)$  for all  $j > i$  and some positive constant  $M$ . Consequently we may estimate  $A_{m,\infty}$  and complete the proof of Theorem 2.2 exactly as we did for the stationary case, i.e., Theorem 1.1 when  $I_\phi = \infty$ .  $\square$

REMARK. If conditions (3) and (3') hold simultaneously, then  $\alpha \leq \alpha'$ . The nonstationary case theorem exactly analogous to Theorem 1.1 holds if conditions (3) and (3') with  $\alpha = \alpha'$  and condition (4) hold.

Of course, Theorem 6 of [3], which is used by Watanabe as a proof of the asymptotic 0-1 behavior of Brownian motion, can be improved by using a hypothesis analogous to condition (4) of Theorem 2.2.

REFERENCES

[1] PICKANDS, J. (1969). Upcrossing probabilities for stationary Gaussian processes. *Trans. Amer. Math. Soc.* **145** 51-73.  
 [2] SLEPIAN, D. (1962). The one-sided barrier problem for Gaussian noise. *Bell System Tech. J.* **41** 463-501.  
 [3] WATANABE, H. (1970). An asymptotic property of Gaussian processes, I. *Trans. Amer. Math. Soc.* **148** 233-248.