

EFFICIENCY-ROBUST ESTIMATION OF LOCATION¹

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1. Introduction. Let U_1, \dots, U_m and V_1, \dots, V_n be two independent samples from distributions with distribution functions F and G , where $F(x) = G(x - \Delta)$. The usual procedures for finding estimates of Δ which are "efficient" require knowledge of the function F . For many F , estimates are known whose asymptotic variance coincides with the Cramér-Rao lower bound; such estimates are sometimes called asymptotically efficient for F . It is usually the case that, if the populations from which the samples are taken are not in the translation family of F , the estimate computed on the assumption that they are, is not asymptotically efficient in the above sense.

Here, estimates are proposed which are asymptotically efficient, uniformly for all F in a large class; that is, each such estimate is a sequence of functions which can be constructed without knowledge of F (functions of the observations only) and whose asymptotic variance is the Cramér-Rao lower bound for F , no matter which F in a class \mathcal{F} is the underlying population. That such estimates exist was indicated by Stein [6]. Bhattacharya [1] proposed estimates of Δ that are "universally almost efficient" for all F in a class \mathcal{F}' . By reducing one sample to a frequency distribution over a fixed set of intervals, he obtains a sequence of estimates (functions of the observations only) whose asymptotic variance equals the Cramér-Rao lower bound for the distribution of the grouped data, no matter which $F \in \mathcal{F}'$ is the underlying distribution. The classes \mathcal{F} and \mathcal{F}' are somewhat different. \mathcal{F}' contains the Cauchy distribution, which is not contained in \mathcal{F} and \mathcal{F} contains the double exponential distribution which is not contained in \mathcal{F}' .

The estimates proposed here are based on Hájek's [2] uniformly asymptotically efficient test of the hypothesis $\Delta = 0$. A modification of this test, satisfying the conditions of Hodges and Lehmann [4], will be shown to lead to a uniformly asymptotically efficient estimate of Δ .

In Section 2 the definition of the estimates will be given, as well as their asymptotic distribution. The proofs of the results of Section 2 will be given in Section 3; Section 4 contains the analogous results for the one-sample problem.

2. The estimates and their asymptotic distribution. Let \mathcal{F} be the set of all distribution functions F with the following properties

1. F has an absolutely continuous density f ,

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- (2.1) 2. $\int_0^1 \varphi^2(u, f) du < \infty$, where $\varphi(u, f) = -\frac{f'(F^{-1}(u))}{f(F^{-1}(u))}$
 $(F^{-1}(u) = \inf \{x \mid F(x) \geq u\})$ and f' is the derivative of f ,
 3. $\varphi(u, f)$ is nondecreasing in $u (0 < u < 1)$.

Let, for $v = 1, 2, \dots$, $\{m_v, n_v\}$ be a sequence of pairs of integers with $\min(m_v, n_v) \rightarrow \infty$ and let $\{(U_v), (V_v)\} \equiv \{(U_{v,1}, \dots, U_{v,m_v}), (V_{v,1}, \dots, V_{v,n_v})\}$ be a sequence of pairs of independent samples from distributions with distribution functions F and G , where $F(x) = G(x - \Delta)$ and $F \in \mathcal{F}$. Let $\{K_v\}$ be a sequence of integers with

$$(2.2) \quad K_v \rightarrow \infty, \quad K_v / \min(m_v, n_v) \rightarrow 0$$

and let $(W_v) \equiv (U_{v,1}, \dots, U_{v,K_v}), (Z_v) \equiv (V_{v,1}, \dots, V_{v,K_v}), (\dot{X}_v) \equiv (X_{v,1}, \dots, X_{v,m_v - K_v}) \equiv (U_{v,K_v+1}, \dots, U_{v,m_v})$ and $(Y_v) \equiv (Y_{v,1}, \dots, Y_{v,n_v - K_v}) \equiv (V_{v,K_v+1}, \dots, V_{v,n_v})$. Hájek [2] proposed an estimate of $\varphi(u, f)$ based on the observations (W_v, Z_v) and used this estimate and the observations (X_v) and (Y_v) to construct a sequence of tests of the hypothesis $\Delta = 0$ that is uniformly asymptotically efficient for a sequence of alternatives Δ_v with $\Delta_v \rightarrow 0$. In order to obtain a sequence of estimates of Δ that is uniformly asymptotically efficient for each fixed Δ , an estimate of $\varphi(u, f)$ will be used that is an average of two estimates, one based on (W_v) and one based on (Z_v) . The estimate of Hájek will be further modified by constructing it in such a way that it is nondecreasing in u for each (W_v, Z_v) . This estimate of $\varphi(u, f)$ and the observations (X_v) and (Y_v) will be used to construct a sequence of estimates of Δ that is uniformly asymptotically efficient.

Let $\{p_v\}$ be a sequence of integers satisfying

$$(2.3) \quad K_v^{4/5} < p_v \leq K_v^{4/5} + 1$$

and let $\{0 = h_{v,0} < h_{v,1} < \dots < h_{v,q_v-1} < h_{v,q_v} = K_v\}$ be a sequence of $(q_v + 1)$ -tuples of integers, satisfying

$$(2.4) \quad \lim_{v \rightarrow \infty} \max_{0 \leq j < q_v} |h_{v,j+1} - h_{v,j}| / K_v^{5/6} = \lim_{v \rightarrow \infty} \min_{0 \leq j < q_v} |h_{v,j+1} - h_{v,j}| / K_v^{5/6} = 1.$$

Let $U_v^{(1)} < \dots < U_v^{(K_v)}$ be the order statistics of $U_{v,1}, \dots, U_{v,K_v}$ and let $N_v = m_v + n_v$. Then Hájek's estimate $\tilde{\varphi}_v(u, W_v)$ of $\varphi(u, f)$ based on (W_v) , is given by

$$(2.5) \quad \tilde{\varphi}_v \left(\frac{i}{N_v - 2K_v + 1}, W_v \right) = \frac{1}{2} K_v^{-1/30} \left\{ \frac{1}{U_v^{(h_{v,j} + p_v)} - U_v^{(h_{v,j} - p_v)}} - \frac{1}{U_v^{(h_{v,j+1} + p_v)} - U_v^{(h_{v,j+1} - p_v)}} \right\}$$

for $\frac{h_{v,j}}{K_v} < \frac{i}{N_v - 2K_v + 1} \leq \frac{h_{v,j+1}}{K_v} \quad j = 1, \dots, q_v - 1, i = 1, \dots, N_v - 2K_v$

= 0 otherwise

and the definition is completed by taking $\tilde{\varphi}_v(u, W_v)$ constant on the intervals $[(i-1)/(N_v-2K_v), i/(N_v-2K_v)]$, $i = 1, \dots, N_v-2K_v$.

For the same sequences $\{p_v\}$, $\{q_v\}$ and $\{h_{v,0}, \dots, h_{v,q_v}\}$, let $\tilde{\varphi}_v(u, Z_v)$ be Hájek's estimate based on (Z_v) and let

$$(2.6) \quad \tilde{\varphi}_v(u) = \tilde{\varphi}_v(u, W_v, Z_v) = \frac{1}{2}\{\tilde{\varphi}_v(u, W_v) + \tilde{\varphi}_v(u, Z_v)\} \quad 0 \leq u \leq 1.$$

This estimate $\tilde{\varphi}_v(u)$ is a function that is, for $0 \leq u \leq 1$, constant on each of a finite number of intervals. Call these interval $I_{v,i}$ ($i = 1, \dots, Q_v$), in such a way that, for each $i = 1, \dots, Q_v-1$, $u_i < u_{i+1}$ if $u_i \in I_{v,i}$ and $u_{i+1} \in I_{v,i+1}$. Let, for $i = 1, \dots, Q_v$, $\tilde{\varphi}_{v,i}$ be the value of $\tilde{\varphi}_v(u)$ for $u \in I_{v,i}$ and let $l_{v,i}$ be the length of $I_{v,i}$. Then define

$$(2.7) \quad \hat{\varphi}_v(u) = \hat{\varphi}_v(u, W_v, Z_v) \\ = \max_{1 \leq j \leq i} \min_{i \leq k \leq Q_v} \frac{l_{v,j} \tilde{\varphi}_{v,j} + \dots + l_{v,k} \tilde{\varphi}_{v,k}}{l_{v,j} + \dots + l_{v,k}} \quad \text{for } u \in I_{v,i}, \\ i = 1, \dots, Q_v.$$

The estimate of $\varphi(u, f)$, used in the construction of the sequence of estimates of Δ , is then given by

$$(2.8) \quad \hat{\varphi}_v^*(u) = \hat{\varphi}_v^*(u, Z_v, W_v) = \hat{\varphi}_v(u) - (N_v - 2K_v)^{-1} \\ \times \sum_{i=1}^{N_v-2K_v} \hat{\varphi}_v(i(N_v - 2K_v + 1)^{-1}) \\ = \hat{\varphi}_v(u) - \int_0^1 \hat{\varphi}_v(u) du.$$

Now define

$$(2.9) \quad \hat{h}_v^*(X_v, Y_v) = \sum_{i=1}^{m_v-K_v} \hat{\varphi}_v^*(R_{v,i}(N_v - 2K_v + 1)^{-1}),$$

where, for $i = 1, \dots, m_v - K_v$, $R_{v,i}$ is the rank of $X_{v,i}$ in $(X_{v,1}, \dots, X_{v,m_v-K_v}, Y_{v,1}, \dots, Y_{v,m_v-K_v})$. The statistic $\hat{h}_v^*(X_v, Y_v)$ satisfies the condition (see Section 3) that, for each (U_v, V_v) , $\hat{h}^*(X_v - b, Y_v)$ is a nonincreasing function of b . Further (see Section 3) $\hat{h}^*(X_v - b, Y_v)$ satisfies, for each (U_v, V_v) , one of the following two conditions.

$$(2.10) \quad \begin{aligned} &1. \hat{h}_v^*(X_v - b, Y_v) = 0 \quad \text{for all } b \quad \text{or} \\ &2. \hat{h}_v^*(X_v - b, Y_v) > 0 \quad \text{for } b < X_v^{(1)} - Y_v^{(n_v-K_v)} \\ &\quad < 0 \quad \text{for } b > X_v^{(m_v-K_v)} - Y_v^{(1)}. \end{aligned}$$

Let S_v be the set of points (U_v, V_v) where (2.10.2) is satisfied, then (see Section 3), for every Δ and any $F \in \mathcal{F}$, $P_{v,\Delta}((U_v, V_v) \in S_v) \rightarrow 1$. For $(U_v, V_v) \in S_v$ the estimate $\hat{\Delta}_v(U_v, V_v)$ of Δ is defined as follows.

Let

$$(2.11) \quad \begin{aligned} \Delta_v^*(U_v, V_v) &= \sup \{b \mid \hat{h}_v^*(X_v - b, Y_v) > 0\} \\ \Delta_v^{**}(U_v, V_v) &= \inf \{b \mid \hat{h}_v^*(X_v - b, Y_v) < 0\} \end{aligned}$$

and let α be a fixed number with $0 \leq \alpha \leq 1$; then

$$(2.12) \quad \hat{\Delta}_v(U_v, V_v) = \alpha \Delta_v^*(U_v, V_v) + (1 - \alpha) \Delta_v^{**}(U_v, V_v) \quad \text{for } (U_v, V_v) \in S_v.$$

Because, for every Δ and any $F \in \mathcal{F}$, $P_{v,\Delta}((U_v, V_v) \notin S_v) \rightarrow 0$, the asymptotic distribution of the estimate $\hat{\Delta}_v(U_v, V_v)$ does not depend on the definition of $\hat{\Delta}_v(U_v, V_v)$ for $(U_v, V_v) \notin S_v$.

The following theorem will be proved in Section 3.

THEOREM 2.1. *For every fixed Δ and any $F \in \mathcal{F}$*

$$(2.13) \quad \lim_{v \rightarrow \infty} P_{v,\Delta}((m_v n_v)^{\frac{1}{2}}(N_v)^{-\frac{1}{2}}(\hat{\Delta}_v(U_v, V_v) - \Delta) \leq u) = \sigma^{-1}(2\pi)^{-\frac{1}{2}} \int_{-\infty}^u \exp(-\frac{1}{2}\sigma^{-2}x^2) dx,$$

where

$$(2.14) \quad \sigma^2 = [\int_0^1 \varphi^2(u, f) du]^{-1} = [\int_{-\infty}^{+\infty} (f'(x)/f(x))^2 f(x) dx]^{-1}$$

3. Proofs of the results of Section 2.

LEMMA 3.1. *For each (U_v, V_v) , $\hat{h}_v^*(X_v - b, Y_v)$ is a nonincreasing function of b .*

PROOF. From the definition of $\hat{\varphi}_v^*(u)$ it follows that $\hat{\varphi}_v^*(u)$ is a nondecreasing function of u . Further $\hat{\varphi}_v^*(u)$ is independent of (X_v) and (Y_v) so that $\hat{h}_v^*(X_v - b, Y_v)$ depends on b only through the ranks $R_{v,i}(b)$ of $X_{v,i} - b$ among $(X_{v,1} - b, \dots, X_{v,m_v - K_v} - b, Y_{v,1}, \dots, Y_{v,n_v - K_v})$ and not through $\hat{\varphi}_v^*$ itself. It is easily seen that $R_{v,i}(b)$ is a nonincreasing function of b , so that $\hat{h}_v^*(X_v - b, Y_v)$ is a sum of $m_v - K_v$ nonincreasing functions, $\hat{\varphi}_v^*(R_{v,i}(b)(N_v - 2K_v + 1)^{-1})$, of b .

LEMMA 3.2. *For each (U_v, V_v) , $\hat{h}_v^*(X_v - b, Y_v)$ satisfies one of the following two conditions*

$$(3.1) \quad \begin{aligned} &1. \hat{h}_v^*(X_v - b, Y_v) = 0 \quad \text{for all } b \quad \text{or} \\ &2. \hat{h}_v^*(X_v - b, Y_v) > 0 \quad \text{for } b < X_v^{(1)} - Y_v^{(m_v - K_v)} \\ &\quad < 0 \quad \text{for } b > X_v^{(m_v - K_v)} - Y_v^{(1)}. \end{aligned}$$

PROOF. $\hat{h}_v^*(X_v - b, Y_v) = 0$ for all b if $\hat{\varphi}_v^*(u) = 0$ for all u . If $\hat{h}_v^*(X_v - b, Y_v) \neq 0$ for some b , then $\hat{\varphi}_v^*(u) \neq 0$ for some u . Because $\hat{\varphi}_v^*(u)$ is nondecreasing and constant on each of the intervals $[(i-1)/(N_v - 2K_v), i/(N_v - 2K_v)]$ ($i = 1, \dots, N_v - 2K_v$) and because $\sum_{i=1}^{N_v - 2K_v} \hat{\varphi}_v^*(i(N_v - 2K_v + 1)^{-1}) = 0$, there exists then a value of i with $\hat{\varphi}_v^*(i(N_v - 2K_v + 1)^{-1}) < \hat{\varphi}_v^*((i+1)(N_v - 2K_v + 1)^{-1})$. Thus

$$(3.2) \quad \sum_{i=1}^k \hat{\varphi}_v^*(i(N_v - 2K_v + 1)^{-1}) < 0 < \sum_{i=k+1}^{N_v - 2K_v} \hat{\varphi}_v^*(i(N_v - 2K_v + 1)^{-1})$$

for all $k = 1, \dots, N_v - 2K_v - 1$

Further, if $b > X_v^{(m_v - K_v)} - Y_v^{(1)}$, $R_{v,i}(b) = i$ for $i = 1, \dots, m_v - K_v$, thus (see (3.2))

$$(3.3) \quad \hat{h}_v^*(X_v - b, Y_v) = \sum_{i=1}^{m_v - K_v} \hat{\varphi}_v^*(i(N_v - 2K_v + 1)^{-1}) < 0;$$

if $b < X_v^{(1)} - Y_v^{(n_v - K_v)}$, $R_{v,i}(b) = n_v - K_v + i$ for $i = 1, \dots, m_v - K_v$ and (see (3.2))

$$(3.4) \quad \begin{aligned} \hat{h}_v^*(X_v - b, Y_v) &= \sum_{i=1}^{m_v - K_v} \hat{\phi}_v^*(i)(n_v - K_v + i)(N_v - 2K_v + 1)^{-1} \\ &= \sum_{i=n_v - K_v + 1}^{N_v - 2K_v} \hat{\phi}_v^*(i)(N_v - 2K_v + 1)^{-1} > 0. \end{aligned}$$

LEMMA 3.3. For any $F \in \mathcal{F}$ and every fixed Δ , $\hat{\phi}_v^*(u)$ is a consistent estimate of $\phi(u, f)$ in the sense that

$$(3.5) \quad \lim_{v \rightarrow \infty} P_{v,\Delta} \{ \int_0^1 (\hat{\phi}_v^*(u) - \phi(u, f))^2 du > \varepsilon \} = 0.$$

PROOF. Let Φ be the set of all functions $\varphi(u)$ ($0 \leq u \leq 1$) satisfying

- (3.6) 1. $\int_0^1 \varphi^2(u) du < \infty$
2. $\varphi(u)$ is nondecreasing in u ,

then it will first be shown that $\hat{\phi}_v(u)$ is the function $\varphi(u)$ which minimizes $\int_0^1 (\varphi(u) - \tilde{\varphi}_v(u))^2 du$ for $\varphi(u) \in \Phi$. This can be seen as follows.

$$(3.7) \quad \int_0^1 (\varphi(u) - \tilde{\varphi}_v(u))^2 du = \sum_{i=1}^{Q_v} \int_{I_{v,i}} (\varphi(u) - \tilde{\varphi}_{v,i})^2 du$$

Further, for each $i = 1, \dots, Q_v$

$$(3.8) \quad \int_{I_{v,i}} (\varphi(u) - \tilde{\varphi}_{v,i})^2 du \geq \int_{I_{v,i}} (\bar{\varphi}_{v,i} - \tilde{\varphi}_{v,i})^2 du, \quad \text{where}$$

$\bar{\varphi}_{v,i} = l_{v,i}^{-1} \int_{I_{v,i}} \varphi(u) du$. Thus a function $\varphi(u)$ which minimizes $\int_0^1 (\varphi(u) - \tilde{\varphi}_v(u))^2 du$ for $\varphi(u) \in \Phi$ is a function which is constant on each of the intervals $I_{v,i}$. Thus the problem of minimizing $\int_0^1 (\varphi(u) - \tilde{\varphi}_v(u))^2 du$ for $\varphi(u) \in \Phi$ is reduced to the problem of finding Q_v numbers $\hat{\varphi}_{v,1}, \dots, \hat{\varphi}_{v,Q_v}$, satisfying $\hat{\varphi}_{v,1} \leq \dots \leq \hat{\varphi}_{v,Q_v}$ such that $\sum_{i=1}^{Q_v} l_{v,i} (\hat{\varphi}_{v,i} - \tilde{\varphi}_{v,i})^2$ is a minimum. The problem of finding the $\hat{\varphi}_{v,i}$ such that $\sum_{i=1}^{Q_v} l_{v,i} (\hat{\varphi}_{v,i} - \tilde{\varphi}_{v,i})^2$ is a minimum, subject to the conditions $\hat{\varphi}_{v,1} \leq \dots \leq \hat{\varphi}_{v,Q_v}$, is a special case of a problem solved by van Eeden ([7] and [8]). It is proved there that this problem has a unique solution given by (2.7).

The consistency of $\hat{\phi}_v(u)$ can now be proved as follows. Hájek [2] proved that, for any F satisfying (2.1.1) and (2.1.2), $\tilde{\varphi}_v(u, W_v)$ and $\tilde{\varphi}_v(u, Z_v)$ are consistent estimates of $\varphi(u, f)$. From the definition (2.6) it is then clear that for any such F and every fixed Δ , $\tilde{\varphi}_v(u)$ is a consistent estimate of $\varphi(u, f)$. Now suppose that $\int_0^1 (\varphi(u, f) - \tilde{\varphi}_v(u))^2 du < \varepsilon$, then, because $\varphi(u, f) \in \Phi$ and because $\hat{\phi}_v(u)$ minimizes $\int_0^1 (\varphi(u) - \tilde{\varphi}_v(u))^2 du$ for $\varphi(u) \in \Phi$, we have

$$(3.9) \quad \int_0^1 (\varphi(u, f) - \tilde{\varphi}_v(u))^2 du < \varepsilon \rightarrow \int_0^1 (\hat{\phi}_v(u) - \tilde{\varphi}_v(u))^2 du < \varepsilon$$

and thus

$$(3.10) \quad \int_0^1 (\varphi(u, f) - \tilde{\varphi}_v(u))^2 du < \varepsilon \rightarrow \int_0^1 (\varphi(u, f) - \hat{\phi}_v(u))^2 du < 4\varepsilon$$

and the consistency of $\hat{\phi}_v(u)$ then follows from the consistency of $\tilde{\varphi}_v(u)$. Now let $\bar{\varphi}_v = \int_0^1 \hat{\phi}_v(u) du$, then

$$(3.11) \quad \int_0^1 (\varphi(u, f) - \hat{\phi}_v^*(u))^2 du = \int_0^1 (\varphi(u, f) - \hat{\phi}_v(u) + \bar{\varphi}_v)^2 du, \quad \text{where}$$

$$(3.12) \quad \bar{\varphi}_v^2 = (\int_0^1 \hat{\phi}_v(u) du)^2 = (\int_0^1 (\hat{\phi}_v(u) - \varphi(u, f)) du)^2 \leq \int_0^1 (\hat{\phi}_v(u) - \varphi(u, f))^2 du$$

so that

$$(3.13) \quad \int_0^1 (\varphi(u, f) - \hat{\varphi}_v(u))^2 du < \varepsilon \rightarrow \int_0^1 (\varphi(u, f) - \hat{\varphi}_v^*(u))^2 du < 4\varepsilon$$

and the consistency of $\hat{\varphi}_v^*(u)$ follows from that of $\hat{\varphi}_v(u)$.

LEMMA 3.4. For every fixed Δ and any $F \in \mathcal{F}$

$$(3.14) \quad \lim_{v \rightarrow \infty} P_{v, \Delta}(\hat{h}_v^*(X_v - b, Y_v) = 0 \text{ for all } b) = 0.$$

PROOF. $\hat{h}_v^*(X_v - b, Y_v) = 0$ if and only if $\hat{\varphi}_v^*(u) = 0$ for all u . (See Lemma 3.2.) If $\hat{\varphi}_v^*(u)$ is identically 0 then

$$(3.15) \quad \int_0^1 (\hat{\varphi}_v^*(u) - \varphi(u, f))^2 du = \int_0^1 \varphi^2(u, f) du > 0.$$

Thus

$$(3.16) \quad P_{v, \Delta} \{ \hat{h}^*(X_v - b, Y_v) = 0 \text{ for all } b \} \\ \leq P_{v, \Delta} \{ \int_0^1 (\hat{\varphi}_v^*(u) - \varphi(u, f))^2 du = \int_0^1 \varphi^2(u, f) du \}$$

and this last probability tends to zero as $v \rightarrow \infty$, because of the consistency of $\hat{\varphi}_v^*(u)$.

LEMMA 3.5. If, for $v = 1, 2, \dots$, $\Delta_v = b[N_v(m_v, n_v)^{-1}]^{\frac{1}{2}}$ is a sequence of values of Δ , then for any $F \in \mathcal{F}$

$$(3.17) \quad \lim_{v \rightarrow \infty} P_{v, \Delta_v}(\sigma_v^{-1}(\hat{h}_v^*(X_v, Y_v) - \mu_v) \leq u) = (2\pi)^{-\frac{1}{2}} \int_{-\infty}^u e^{-\frac{1}{2}x^2} dx,$$

$$(3.18) \quad \mu_v = b[(m_v - K_v)(n_v - K_v)(N_v - 2K_v)^{-1}]^{\frac{1}{2}} \int_0^1 \varphi^2(u, f) du \\ \sigma_v^2 = (m_v - K_v)(n_v - K_v)(N_v - 2K_v)^{-1} \int_0^1 \varphi^2(u, f) du$$

PROOF. Let, for $i = 1, \dots, N_v - 2K_v$

$$(3.19) \quad a_v(i, f) = E\varphi(T_v^{(i)}, f)$$

where $T_v^{(1)} < \dots < T_v^{(N_v - 2K_v)}$ are the order statistics of a sample of size $N_v - 2K_v$ from a uniform distribution between 0 and 1. Further let

$$(3.20) \quad g_v(X_v, Y_v) = \sum_{i=1}^{m_v - K_v} a_v(R_{v,i}, f)$$

then it follows from Theorem VI.2.3 and Theorem V.1.4b of Hájek and Sidák [3] and the fact that $K_v/\min(m_v, n_v) \rightarrow 0$, that

$$(3.21) \quad \lim_{v \rightarrow \infty} P_{v, \Delta_v}(\sigma_v^{-1}(g_v(X_v, Y_v) - \mu_v) \leq u) = (2\pi)^{-\frac{1}{2}} \int_{-\infty}^u e^{-\frac{1}{2}x^2} dx,$$

where μ_v and σ_v are given by (3.15). That

$$\sigma_v^{-1}(g_v(X_v, Y_v) - \mu_v) \quad \text{and} \quad \sigma_v^{-1}(\hat{h}_v^*(X_v, Y_v) - \mu_v)$$

have, for the sequence $\Delta_v = b[N_v(m_v, n_v)^{-1}]^{\frac{1}{2}}$, the same asymptotic distribution follows from Hájek and Sidák ([3] pages 264–265). This proof applies in our case, since $(R_{v,1}, \dots, R_{v, m_v - K_v})$ and (W_v, Z_v) are independent, $\hat{\varphi}_v^*(u)$ is a function of (W_v, Z_v) only and since $\hat{\varphi}_v^*(u)$ is a consistent estimate of $\varphi(u, f)$.

PROOF OF THEOREM 2.1. The proof is based on the results of Hodges and Lehmann [4]. They consider two sample statistics $h_v(X_v, Y_v)$ satisfying the conditions

A. for each (X_v, Y_v) , $h_v(X_v - b, Y_v)$ is a nonincreasing function of b (and is not identically zero) and

B. if $\Delta = 0$, the distribution of $h_v(X_v, Y_v)$ is, for every continuous F , symmetric around 0.

Their estimates derived from these statistics are estimates of the form (2.12) with $\alpha = \frac{1}{2}$ and one of their results is the following inequality

$$(3.22) \quad P_{v,\Delta}(h_v(X_v - b, Y_v) < 0) \leq P_{v,\Delta}(\hat{\Delta}_v(X_v, Y_v) \leq b) \leq P_{v,\Delta}(h_v(X_v - b, Y_v) \leq 0)$$

for any continuous F .

From the proofs by Hodges and Lehmann it can be seen (see also Hoyland [5]) that condition B is not necessary for (3.22). Further it can easily be seen from their proofs that the estimates satisfy (3.22) for any $\alpha \in [0,1]$. From Lemma 3.1 and Lemma 3.2 it then follows that, for any $F \in \mathcal{F}$ and for every v and every Δ ,

$$(3.23) \quad \begin{aligned} & P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) < 0 \mid (U_v, V_v) \in S_v) \\ & \leq P_{v,\Delta}(\hat{\Delta}_v(U_v, V_v) \leq b \mid (U_v, V_v) \in S_v) \\ & \leq P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) \leq 0 \mid (U_v, V_v) \in S_v). \end{aligned}$$

Thus, independent of the definition of $\hat{\Delta}_v(U_v, V_v)$ for $(U_v, V_v) \notin S_v$,

$$(3.24) \quad \begin{aligned} & P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) < 0, (U_v, V_v) \in S_v) + P_{v,\Delta}(\hat{\Delta}_v(U_v, V_v) \leq b, (U_v, V_v) \notin S_v) \\ & \leq P_{v,\Delta}(\hat{\Delta}_v(U_v, V_v) \leq b) \\ & \leq P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) \leq 0, (U_v, V_v) \in S_v) \\ & \quad + P_{v,\Delta}(\hat{\Delta}_v(U_v, V_v) \leq b, (U_v, V_v) \notin S_v). \end{aligned}$$

From Lemma 3.4 it follows that, for every $F \in \mathcal{F}$ and every fixed Δ ,

$$(3.25) \quad \lim_{v \rightarrow \infty} P_{v,\Delta}(\hat{\Delta}_v(U_v, V_v) \leq b, (U_v, V_v) \notin S_v) = 0$$

and

$$(3.26) \quad \begin{aligned} & \lim_{v \rightarrow \infty} [P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) < 0, (U_v, V_v) \in S_v) \\ & \quad - P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) < 0)] = 0 \\ & \lim_{v \rightarrow \infty} [P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) \leq 0, (U_v, V_v) \in S_v) \\ & \quad - P_{v,\Delta}(\hat{h}_v^*(X_v - b, Y_v) \leq 0)] = 0. \end{aligned}$$

From Lemma 3.5 it then follows that, for any $F \in \mathcal{F}$ and every Δ , (see also Hodges and Lehmann [4])

$$\begin{aligned}
 & \lim_{v \rightarrow \infty} P_{v,\Delta}([N_v^{-1}(m_v n_v)]^{\frac{1}{2}}(\hat{\Delta}_v(U_v, V_v) - \Delta) \leq b) \\
 &= \lim_{v \rightarrow \infty} P_{v,0}([N_v^{-1}(m_v n_v)]^{\frac{1}{2}}\hat{\Delta}_v(U_v, V_v) \leq b) \\
 &= \lim_{v \rightarrow \infty} P_{v,0}(\hat{\Delta}_v(U_v, V_v) \leq b[N_v(m_v n_v)^{-1}]^{\frac{1}{2}}) \\
 (3.27) \quad &= \lim_{v \rightarrow \infty} P_{v,0}(\hat{h}_v^*(X_v - b[N_v(m_v n_v)^{-1}]^{\frac{1}{2}}, Y_v) \leq 0) \\
 &= \lim_{v \rightarrow \infty} P_{v,-\Delta_v}(\hat{h}_v^*(X_v, Y_v) \leq 0) \\
 &= \lim_{v \rightarrow \infty} P_{v,-\Delta_v}(\sigma_v^{-1}(\hat{h}_v^*(X_v, Y_v) - \mu_v) \leq -\sigma_v^{-1}\mu_v) \\
 &= (2\pi)^{-\frac{1}{2}} \int_{-\infty}^{b[N_v(m_v n_v)^{-1}]^{\frac{1}{2}}} \int_0^1 \phi^2(u, f) du e^{-\frac{1}{2}x^2} dx = \sigma^{-1}(2\pi)^{-\frac{1}{2}} \int_{-\infty}^b \exp(-\frac{1}{2}\sigma^{-2}x^2) dx,
 \end{aligned}$$

where $\sigma^2 = (\int_0^1 \phi^2(u, f) du)^{-1}$.

4. The one-sample location problem. In this section the results for the one-sample location problem will be given. The proofs are analogous to those for the two-sample case and will not be given.

Let \mathcal{F}_1 be defined by

$$(4.1) \quad \mathcal{F}_1 = \{F \in \mathcal{F} \mid F \text{ is symmetric around } 0\}.$$

Let, for $v = 1, 2, \dots$, $\{N_v\}$ be a sequence of integers with $N_v \rightarrow \infty$ and let $\{U_v\} = \{U_{v,1}, \dots, U_{v,N_v}\}$ be a sequence of samples from a distribution with distribution function $F(x - \theta)$, where $F \in \mathcal{F}_1$. An estimate $\hat{\theta}_v(U_v)$ of θ that is uniformly asymptotically efficient for $F \in \mathcal{F}_1$ can be obtained as follows.

For every $F \in \mathcal{F}_1$, $\varphi(u, f)$ is nondecreasing in u and

$$(4.2) \quad \varphi(u, f) = -\varphi(1 - u, f) \quad 0 < u < 1.$$

An estimate $\hat{\varphi}_v^*(u)$ of $\varphi(u, f)$ satisfying these two conditions is obtained as follows. Let $\{K_v\}$ be a sequence of integers satisfying

$$(4.3) \quad K_v \rightarrow \infty, \quad K_v/N_v \rightarrow 0,$$

and let $\{p_v\}, \{q_v\}$ and $\{h_{v,0}, \dots, h_{v,q_v}\}$ be sequences satisfying (2.3) and (2.4). Further let $(W_v) \equiv (U_{v,1}, \dots, U_{v,K_v})$, $(X_v) \equiv (X_{v,1}, \dots, X_{v,N_v - K_v}) \equiv (U_{v,K_v+1}, \dots, U_{v,N_v})$ and let $\tilde{\varphi}_v(u, W_v)$ be Hájek's estimate of $\varphi(u, f)$ based on (W_v) (see (2.5)). This function $\tilde{\varphi}_v(u, W_v)$ is a function that is constant on each of a finite number of intervals $I_{v,1}, \dots, I_{v,Q_v}$, where, for each $i = 1, \dots, Q_v - 1$, $u_i < u_{i+1}$ if $u_i \in I_{v,i}$ and $u_{i+1} \in I_{v,i+1}$. Let, for $i = 1, \dots, Q_v$, $\tilde{\varphi}_{v,i}$ be the value of $\tilde{\varphi}_v(u, W_v)$ for $u \in I_{v,i}$ and let $l_{v,i}$ be the length of $I_{v,i}$. Then define

$$\begin{aligned}
 (4.4) \quad \hat{\varphi}_v(u) &= \hat{\varphi}_v(u, W_v) \\
 &= \max_{1 \leq j \leq i} \min_{i \leq k \leq Q_v} \frac{l_{v,j} \tilde{\varphi}_{v,j} + \dots + l_{v,k} \tilde{\varphi}_{v,k}}{l_{v,j} + \dots + l_{v,k}}, \quad u \in I_{v,i}, \\
 & \hspace{20em} i = 1, \dots, Q_v
 \end{aligned}$$

and

$$(4.5) \quad \hat{\varphi}_v^*(u) = \hat{\varphi}_v^*(u, W_v) = \frac{1}{2}(\hat{\varphi}_v(u) - \hat{\varphi}_v(1 - u)) \quad 0 \leq u \leq 1.$$

Then $\hat{\phi}_v^*(u)$ is, for every W_v , nondecreasing in u and

$$(4.6) \quad \hat{\phi}_v^*(u) = -\hat{\phi}_v^*(1-u) \quad 0 \leq u \leq 1.$$

Let, for $i = 1, \dots, N_v$, $R_{v,i}^+$ be the rank of $|X_{v,i}|$ among $|X_{v,1}|, \dots, |X_{v,N_v-K_v}|$ and let

$$(4.7) \quad \hat{h}_v^*(X_v) = \sum_{i=1}^{N_v-K_v} \hat{\phi}_v^* \left(\frac{1}{2} [R_{v,i}^+(N_v-K_v+1)^{-1} + 1] \right) |X_{v,i}|^{-1},$$

then $\hat{h}_v^*(X_v - b)$ is, for every U_v , nonincreasing in b and satisfies one of the following two conditions

$$(4.8) \quad \begin{aligned} 1. & \hat{h}_v^*(X_v - b) = 0 \quad \text{for all } b \quad \text{or} \\ 2. & \hat{h}_v^*(X_v - b) > 0 \quad \text{for } b < X_v^{(1)} \\ & < 0 \quad \text{for } b > X_v^{(N_v-K_v)} \end{aligned}$$

with, for each fixed θ and any $F \in \mathcal{F}_1$,

$$(4.9) \quad \lim_{v \rightarrow \infty} P_{v,\theta}(\hat{h}_v^*(X_v - b) = 0 \quad \text{for all } b) = 0.$$

Let S_v be the set of points (U_v) such that (4.8.2) is satisfied, then the estimate $\hat{\theta}_v(U_v)$ for $(U_v) \in S_v$ is defined as follows. Let

$$(4.10) \quad \begin{aligned} \theta_v^*(U_v) &= \sup \{b \mid \hat{h}_v^*(X_v - b) > 0\} \\ \theta_v^{**}(U_v) &= \inf \{b \mid \hat{h}_v^*(X_v - b) < 0\} \end{aligned}$$

and let α be a fixed number with $0 \leq \alpha \leq 1$. Then

$$(4.11) \quad \hat{\theta}_v(U_v) = \alpha \theta_v^*(U_v) + (1-\alpha) \theta_v^{**}(U_v) \quad \text{for } (U_v) \in S_v.$$

Further, because of (4.9), the asymptotic distribution of the estimate $\hat{\theta}_v(U_v)$ does not depend on the definition of the estimate for $(U_v) \notin S_v$. In the same way as for the two-sample problem, the following theorem can be proved

THEOREM 4.1. For every fixed θ and any $F \in \mathcal{F}_1$

$$(4.12) \quad \lim_{v \rightarrow \infty} P_{v,\theta}(N_v^{\frac{1}{2}}(\hat{\theta}_v(U_v) - \theta) \leq u) = \sigma^{-1} (2\pi)^{-\frac{1}{2}} \int_{-\infty}^u \exp(-\frac{1}{2}\sigma^{-2}x^2) dx,$$

where

$$(4.13) \quad \sigma^2 = \left[\int_0^1 \phi^2(\frac{1}{2}(u+1), f) du \right]^{-1} = \left[\int_{-\infty}^{+\infty} (f'(x)/f(x))^2 f(x) dx \right]^{-1}.$$

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