

EQUIVALENT COMPARISONS OF EXPERIMENTS¹

BY DAVID BLACKWELL

Stanford University and Howard University

1. Summary. Sherman [8] and Stein [9] have shown that a method given by the author [1] for comparing two experiments is equivalent, for experiments with a finite number of outcomes, to the original method introduced by Bohnenblust, Shapley, and Sherman [4]. A new proof of this result is given, and the restriction to experiments with a finite number of outcomes is removed. A class of weaker comparisons—comparison in k -decision problems—is introduced, in three equivalent forms. For dichotomies, all methods are equivalent, and can be described in terms of errors of the first and second kinds.

2. Introduction. An ordered collection $\alpha = (m_1, \dots, m_n)$ of probability measures on a Borel field \mathfrak{B} of subsets of a space X will be called an *experiment*. Any pair (α, A) , where A is a closed bounded convex subset of n -space corresponds to a decision problem as follows. A point $x \in X$ is selected according to one of the distributions m_i ; the statistician observes x and then chooses an action d from a given set D , incurring a loss $L(i, d)$. If we associate with d the vector $w(d) = (L(1, d), \dots, L(n, d))$, the range of $w(d)$ as d varies over D is the set A associated with the problem. Thus we may replace D by A , and suppose that the statistician chooses a point $a = (a_1, \dots, a_n) \in A$, incurring loss a_i when m_i is the distribution of x . By using randomized decision procedures we increase A to its convex hull, and for simplicity we suppose A closed and bounded as well as convex.

A decision function in the problem (α, A) is a \mathfrak{B} -measurable function f from X into A , specifying for each x the action $a = f(x)$ to be taken when x is observed. When m_i is the distribution of x , the expected loss from f is $v_i(f) = \int a_i(x) dm_i(x)$; the vector $v(f) = (v_1(f), \dots, v_n(f))$ is called the loss vector of f , and the range of $v(f)$ as f varies over all decision functions in the problem (α, A) will be denoted by $B(\alpha, A)$. The set $B(\alpha, A)$ will be a closed, bounded, convex subset of n -space [2].

For two experiments α, β with the same n , following Bohnenblust, Shapley, and Sherman, we say that α is *more informative than* β , written $\alpha \supset \beta$, if for every A we have $B(\alpha, A) \supset B(\beta, A)$, that is if every loss vector attainable in problem (β, A) is also attainable in (α, A) . For any experiment $\alpha = (m_1, \dots, m_n)$, let $p_i(x)$ be the density of m_i with respect to $\sum_{i=1}^n m_i$, let $p(x) = [p_1(x), \dots, p_n(x)]$, and let m_α denote the distribution of $p(x)$ when x has distribution $\sum_{i=1}^n m_i/n$. Then m_α is a probability measure defined on the set P of all vectors

Received 11/17/52.

¹ This paper was prepared under an Office of Naval Research Contract.

$p = (p_1, \dots, p_n)$ with $p_i \geq 0$ and $\sum_1^n p_i = 1$, and

$$(1) \quad \int p_i dm_\alpha = 1/n;$$

the center of gravity of m_α is the point $(1/n, \dots, 1/n)$. The measure m_α is called by Bohnenblust, Shapley, and Sherman the standard measure associated with the experiment α . Their basic results connecting m_α and \supset are summarized as Theorem 1 below (for a proof see [1]).

THEOREM 1. *Every probability measure on P with property (1) is the standard measure of some experiment; two experiments α and β have the same standard measure if and only if $B(\alpha, A) = B(\beta, A)$ for all A ; $\alpha \supset \beta$ if and only if for every continuous convex function $\phi(p)$ on P , $\int \phi dm_\alpha \geq \int \phi dm_\beta$.*

An alternative method of comparing two experiments α, β , introduced by the author [1], can best be described in terms of the concept of *stochastic transformation*. If $\mathfrak{B}, \mathfrak{C}$ are Borel fields of subsets of X, Y respectively, a *stochastic transformation* T is a function $Q(x, E)$ defined for all $x \in X$ and $E \in \mathfrak{C}$ which for fixed E is a \mathfrak{B} -measurable function of x and for fixed x is a probability measure on \mathfrak{C} . For any probability measure m on \mathfrak{B} , the function $M(E) = \int Q(x, E) dm(x)$ is a probability measure on \mathfrak{C} , denoted by Tm . If X, Y are Borel sets in n -space and $\mathfrak{B}, \mathfrak{C}$ are the Borel subsets of X, Y , T is called *mean-preserving* if $\int y dQ(x, y) = x$ for all x .

If $\alpha = (m_1, \dots, m_n)$ and $\beta = (M_1, \dots, M_n)$ are two experiments, with m_i, M_i defined on Borel fields $\mathfrak{B}, \mathfrak{C}$ of X, Y respectively, we shall say that α is *sufficient* for β , written $\alpha > \beta$, if there is a stochastic transformation T with $Tm_i = M_i$ for $i = 1, \dots, n$. Thus $\alpha > \beta$ means that the statistician, observing the result x of α , can, by selecting y according to $Q(x, E)$, obtain a result equivalent to the result of observing β .

The concept $>$ also has a description in terms of standard measures, summarized in

THEOREM 2. [1]. $\alpha > \beta$ if and only if there is a mean-preserving stochastic transformation T with $Tm_\beta = m_\alpha$.

If $\alpha > \beta$ and ϕ is any continuous convex function on P ,

$$\begin{aligned} \int \phi dm_\alpha &= \int \left(\int \phi(p) dQ(q, p) \right) dm_\beta(q) \\ &\geq \int \phi \left(\int p dQ(q, p) \right) dm_\beta(q) \\ &= \int \phi dm_\beta, \end{aligned}$$

so that, from Theorem 1 we obtain

THEOREM 3. $\alpha > \beta$ implies $\alpha \supset \beta$.

The converse of Theorem 3 has been proved, for experiments with a finite number of outcomes, by Sherman [8] and Stein [9]. In Section 3 we give a new proof of the Sherman-Stein theorem, and in Section 4 extend the theorem to arbitrary experiments.

3. The Sherman-Stein Theorem. If the space X of outcomes of the experiment α is finite, consisting say of x_1, \dots, x_N , then α is characterized by the $n \times N$ Markov matrix $P = \| p_{ij} \|$, where $p_{ij} = m_i(x_j)$, and conversely every Markov matrix can be interpreted as an experiment. For two Markov matrices P, Q with the same n , we write $P \supset Q, P > Q$ if the corresponding experiments are related by $\supset, >$ respectively.

THEOREM 4. If P, Q are $n \times N_1, n \times N_2$ Markov matrices with $P \supset Q$, then for every $N_2 \times n$ matrix D there is an $N_1 \times N_2$ Markov matrix M with

$$\text{Trace}(PMD) \leq \text{Trace}(QD).$$

PROOF. Let A be the convex hull of the rows of D . The decision function f in problem (Q, A) selecting the j th row of D when j is observed has $v_i(f) = \sum_j q_{ij} d_{ji}$, the i th diagonal element of QD .

Since $P \supset Q$, there is a decision function g in problem (P, A) , selecting say $a_j \in A$ when j is observed, with $v_i(g) = \sum_j p_{ij} a_{ji} = v_i(f)$ for all i . Since $a_j \in A$, there are nonnegative numbers m_{jk} with $\sum_k m_{jk} = 1$ such that $a_{ji} = \sum_k m_{jk} d_{ki}$ for all i . Thus $v_i(g) = \sum_{j,k} p_{ij} m_{jk} d_{ki}$, which is the i th diagonal element of PMD . It follows that M has not only the property asserted in the theorem but the stronger property that PMD and QD have identical diagonal elements.

THEOREM 5. $P > Q$ if and only if there is a Markov matrix M with $PM = Q$.

This is simply a restatement of the definition of $>$ for the special case $X = (1, \dots, N_1), Y = (1, \dots, N_2)$, since a stochastic transformation becomes simply an $N_1 \times N_2$ Markov matrix.

THEOREM 6. (Sherman-Stein theorem). $P \supset Q$ implies $P > Q$.

PROOF. Consider the function $h(D, M) = \text{Trace}(Q - PM)D$, as M varies over all $N_1 \times N_2$ Markov matrices and D varies over all $N_2 \times n$ matrices with $0 \leq d_{ki} \leq 1$ for all k, i . Since h is bilinear and the ranges of D, M are closed, bounded, and convex, h has a saddle point [3], that is there exist D_0, M_0 with $h(D_0, M) \geq h(D_0, M_0) \geq h(D, M_0)$ for all D, M . From Theorem 4, there is an M with $h(D_0, M) \leq 0$, so that $h(D, M_0) \leq 0$ for all D . Writing $U = Q - PM_0$, we have

$$\text{Trace}(UD) \leq 0 \text{ for all } D,$$

so that $u_{ik} \leq 0$ for all i, k . Since U is the difference of two Markov matrices, $\sum_{i,k} u_{ik} = 0$, so that $u_{ik} = 0$ for all i, k and $PM_0 = Q$. Thus by Theorem 5, $P > Q$.

An alternative form of the Sherman-Stein theorem is

THEOREM 7. If m_1 and m_2 are any two probability measures on a finite subset X of n -space such that for every continuous convex ϕ defined on the convex hull of X ,

$\int \phi dm_1 \geq \int \phi dm_2$, then there is a mean-preserving stochastic transformation T with $Tm_2 = m_1$.

From Theorems 1 and 2, Theorem 7 implies Theorem 6. Theorem 7 was proved for $n = 1$ by Hardy, Littlewood, and Polya [6], for $n = 2$ without the restriction that X be finite by the author, and in the form given here by Sherman [8] and Stein [9].

PROOF OF THEOREM 7. From Theorems 1 and 2, Theorem 6 implies Theorem 7 if $X \subset P$ and the common center of gravity of m_1, m_2 is $(1/n, \dots, 1/n)$, since in this case m_1, m_2 are the standard measures of experiments. Imbedding X in $n + 1$ space and performing an appropriate linear transformation reduces the general case in n -space to that of standard measures in $n + 1$ space and completes the proof.

A direct proof of Theorem 7, using the methods of Theorem 6 and not appealing to Theorems 1 and 2 can be given.

4. Equivalence of \supset and $>$. In this section we extend Theorem 7, replacing the requirement that X be finite by the weaker requirement that X be bounded. For any two probability measures m, M on a bounded subset X of n -space, we write $M \supset m$ if for every continuous convex ϕ on the convex hull of X $\int \phi dM \geq \int \phi dm$ and $M > m$ if there is a mean-preserving stochastic transformation (abbreviated m.p.s.t.) T with $Tm = M$. We shall prove

THEOREM 8. *If $M \supset m$, then $M > m$.*

The method of proof consists of approximating m, M by measures concentrated on finite sets and using Doob's martingale convergence theorems. We first prove

A. *There exist sequences of measures m_n, M_n each concentrated on a finite set, with $m_N < m_{N+1} \subset m \subset M < M_{N+1} < M_N$ for all N , and for every open set O*

$$m_N(O) \rightarrow m(O), \quad M_N(O) \rightarrow M(O) \quad \text{as } N \rightarrow \infty.$$

PROOF. For any n -vector $a = (a_1, \dots, a_n)$ with integral coordinates, let $C(N, a)$ denote the cube consisting of all $t = (t_1, \dots, t_n)$ with $2^{-N}a_i \leq t_i < 2^{-N}(a_i + 1)$, let $Z(N, a)$ be the center of gravity of m on $C(N, a)$ and let m_N assign to $Z(N, a)$ measure $m(C(N, a))$. It is easily verified that m_N has the required properties.

To define M_N , let $Q_N(t, E)$ for $t \in C(N, a)$ concentrate on the 2^n vertices of $C(N, a)$ assigning to vertex $2^{-N}(a_1 + \epsilon_1, \dots, a_n + \epsilon_n)$, where $\epsilon_i = 0$ or 1 , measure $b_1 b_2 \dots b_n$, where $b_i = 2^{-N}A_i + 1 - t_i$ if $\epsilon_i = 0$ and $b_i = t_i - 2^{-N}a_i$ if $\epsilon_i = 1$. The function $Q_N(t, E)$ is a m.p.s.t. U_N , and if we define $M_N = U_N M$, we have also $M_N = U_N M_{N+1}$, so that M_N has the required properties.

B. *There exist sequences T_N, V_N, W_N of m.p.s.t. each from a finite set of n -space to a finite set of n -space with*

$$(a) \quad m_{N+1} = T_N m_N, \quad (b) \quad M_{N-1} = V_N M_N, \quad (c) \quad M_N = W_N m_N,$$

and

$$(d) \quad W_N = V_{N+1}W_{N+1}T_N.$$

PROOF. From A there exist sequences T_N and V_N with properties (a) and (b). Also from A , $m_N \subset M_N$, so that, from Theorem 7, there is a m.p.s.t. Y_N from a finite set to a finite set with $M_N = Y_N m_N$. For $D > N$, write

$$Y_{ND} = V_{N+1} \cdots V_D Y_D T_{D-1} \cdots T_N,$$

so that

$$Y_{ND} = V_{N+1} Y_{N+1,D} T_N \quad \text{for } D > N + 1,$$

and

$$M_N = Y_{ND} m_N \quad \text{for } D > N.$$

Let $D \rightarrow \infty$ through a subsequence for which Y_{ND} converges for all N , say to W_N . Then W_N satisfies (c) and (d).

PROOF OF THEOREM 8. We specify the joint distribution of two sequences $x_1, x_2, \dots, y_1, y_2, \dots$, of n -dimensional chance variables by

C. For any N , the variables $x_1, \dots, x_N, y_N, \dots, y_1$ form a Markov chain in the order written. The distribution of x_1 is m_1 and the conditional distributions of x_{i+1} given x_i, y_N given x_N , and y_{i-1} given y_i , are specified by T_i, W_N , and V_i respectively.

Part (d) of B guarantees that the requirements C are consistent, and Kolmogorov's extension theorem [7] then asserts the existence of $x_1, x_2, \dots, y_1, y_2, \dots$, with property C. Parts (a), (b), (c) of B imply that x_N, y_N have distributions m_N, M_N respectively. Also the sequence

$$x_1, x_2, \dots, \dots, y_2, y_1$$

forms a martingale [5] in the order written; by Doob's martingale theorem [5], $x_N \rightarrow x^*, y_N \rightarrow y^*$ as $N \rightarrow \infty$, and $E(y^* | x^*) = x^*$. From A, x^* and y^* have distributions m, M respectively, so that $Q(x, E) = \text{Prob} \{y^* \in E | x^* = x\}$ is a m.p.s.t. T with $Tm = M$. This completes the proof.

5. k -decision problems. In this section we introduce a comparison somewhat weaker than $>$. The following lemma will be useful.

LEMMA. For any experiment α and any closed, bounded set C with convex hull A , $B(\alpha, A) = \text{convex hull of } B(\alpha, C)$.

PROOF. Since both $B(\alpha, A)$ and $B(\alpha, C)$ are closed and $B(\alpha, A)$ is convex [2], it suffices to show that every $v(f) \in B(\alpha, A)$ can be approximated by points in the convex hull of $B(\alpha, C)$. We may suppose that f assumes only a finite number of values a_1, \dots, a_N , since every f can be approximated by f 's of this kind. Say

$$S_j = \{f(x) = a_j\}, \quad a_j = \sum_{i=1}^r \lambda_{ji} c_i, \quad \lambda_{ji} \geq 0, \quad \sum_i \lambda_{ji} = 1.$$

For any $h = (h_1, \dots, h_N), 1 \leq h_i \leq r$, define

$$f(h) = c_{h_j} \text{ for } x \in S_j, \lambda(h) = \prod_{j=1}^r \lambda_{jh_i}.$$

Then $v(f(h)) \in B(\alpha, C)$, and $\sum_n \lambda(h) v[f(h)]$ has for its s th coordinate

$$\begin{aligned} \sum_h \lambda(h) \sum_j \int_{S_j} c_{h_j} dm_s &= \sum_{h,j} m_s(S_j(c_{h_j}, \lambda(h))) \\ &= \sum_j m_s(S_j) \left(\sum_i c_{is} \left(\sum_{h: h_j=i} \lambda(h) \right) \right) \\ &= \sum_j m_s(S_j) \left(\sum_i \lambda_{ij} c_{is} \right) = \sum_j a_{js} m_s(S_j) \\ &= s^{\text{th}} \text{ coordinate of } v(f). \end{aligned}$$

This completes the proof.

APPLICATION 1. Let α be any experiment, let $S = (S_1, \dots, S_k)$ be any partition of X into k disjoint \mathcal{G} -measurable sets, let $P(S)$ be the $n \times k$ Markov matrix with $p_{ij} = m_i(S_j)$, let $\mathcal{P}_{\alpha k}^*$ be the range of $P(S)$, and let $\mathcal{P}_{\alpha k}$ be the set of all $n \times k$ Markov matrices P which have the property $\alpha > P$. Then $\mathcal{P}_{\alpha k}$ is the convex hull of $\mathcal{P}_{\alpha k}^*$.

This is the special case of the lemma applied to the experiment α' consisting of nk measures M_{ij} with $M_{ij} = m_i$ for $j = 1, \dots, k$ and C consisting of the $kn \times k$ Markov matrices P_1, \dots, P_k , where P_j has the j th column identically 1 and the remaining columns identically zero.

APPLICATION 2. For any experiment α and any closed bounded convex set A which is the convex hull of the set of k points d_1, \dots, d_k , $B(\alpha, A)$ is the range of $\text{diag } PD$ as P varies over $\mathcal{P}_{\alpha k}$, where $\text{diag } U$ for any $n \times n$ matrix $U = \|u_{ij}\|$ denotes the n -vector $(u_{11}, u_{22}, \dots, u_{nn})$ and D is the $k \times n$ matrix whose rows are d_1, \dots, d_k .

If C consists of d_1, \dots, d_k , and f is any decision function in (α, C) , say $S_j = \{f = d_j\}$. Then the s th coordinate of $v(f)$ is

$$\sum_j m_s(S_j) d_{js},$$

so that

$$v(f) = \text{diag } P(S)D.$$

Thus $B(\alpha, C) = \text{range of } \text{diag } PD \text{ as } P \text{ varies over } \mathcal{P}_{\alpha k}^*$. From the lemma, the convex hull of $B(\alpha, C)$ is $B(\alpha, A)$, and from Application (1) the convex hull of the range of $\text{diag } PD$ as P varies over $\mathcal{P}_{\alpha k}^*$ is the range of PD as P varies over $\mathcal{P}_{\alpha k}$.

THEOREM 9. Let α, β be two experiments with the same n . The following conditions are equivalent:

- (1) $\mathcal{P}_{\alpha k} \supset \mathcal{P}_{\beta k}$
- (2) For every A which is the convex hull of a set of k points, $B(\alpha, A) \supset B(\beta, A)$.
- (3) For every convex function ϕ on n -space which is the maximum of k linear functions, $\int \phi dm_\alpha \geq \int \phi dm_\beta$.

PROOF. Suppose (1) and let $v \in B(\beta, A)$, where A is the convex hull of d_1, \dots, d_k . Then $v = \text{diag } PD$ for some $P \in \mathcal{P}_{\beta k}$.

Since $\mathcal{O}_{\beta k} \subset \mathcal{O}_{\alpha k}$, $v = \text{diag } PD$ for some $P \in \mathcal{O}_{\alpha k}$ and $v \in B(\alpha, A)$. Thus (1) implies (2).

Now suppose (2) and let $P \in \mathcal{O}_{\beta k}$. Then for any closed bounded convex set R , let $v \in B(P, R)$, say $v = v(f)$, where $f(j) = r_j \in R, j = 1, \dots, k$. Then $v \in B(P, R^*)$, where R^* is the convex hull of r_1, \dots, r_k . Since $B(P, R^*) \subset B(\beta, R^*) \subset B(\alpha, R^*)$, $v \in B(\alpha, R^*)$ and consequently $v \in B(\alpha, R)$. Thus $\alpha \supset P$ for any $P \in \mathcal{O}_{\beta k}$ and, by Theorem 8, $\alpha \succ P$. Since $\mathcal{O}_{\alpha k}$ contains all $n \times k$ Markov matrices P with $\alpha \succ P$, $P \in \mathcal{O}_{\alpha k}$ and $\mathcal{O}_{\beta k} \subset \mathcal{O}_{\alpha k}$. Thus (2) implies (1).

In considering (3), we use the fact that the standard measure m_P of an $n \times k$ Markov matrix P is concentrated on k points, which follows immediately from the definition. Suppose (3), let ϕ be the maximum of any finite set \mathcal{L} of linear functions, and let $P \in \mathcal{O}_{\beta k}$. There is a ψ , the maximum of k functions in \mathcal{L} , which agrees with ψ on the k points on which m_P is concentrated. Then $\int \phi dm_\alpha \geq \int \psi dm_\alpha \geq \int \psi dm_\beta \geq \int \psi dm_P = \int \phi dm_P$, so that from Theorems 1 and 8, $\alpha \succ P$. Thus $P \in \mathcal{O}_{\alpha k}$, $\mathcal{O}_{\beta k} \subset \mathcal{O}_{\alpha k}$ and (3) implies (1).

Finally, suppose (1) and let $\phi = \max(L_1, \dots, L_k)$; say

$$U_j = \{L_j(p) = \phi(p), L_i(p) < \phi(p) \text{ for } i < j\}.$$

If $S_j = \{p(x) \in U_j\}$, $\mathcal{S} = (S_1, \dots, S_k)$ is a partition of X and the experiment $P = P(\mathcal{S})$ associated with β and \mathcal{S} (see Application 1) has a standard measure m_P with

$$m_P(U_j) = m_\beta(U_j),$$

so that

$$\int \phi dm_\beta = \int \phi dm_P.$$

Since $P \in \mathcal{O}_{\beta k}$, (1) implies $P \in \mathcal{O}_{\alpha k}$, so that $\int \phi dm_\alpha \geq \int \phi dm_P = \int \phi dm_\beta$. This completes the proof.

If two experiments α, β with the same n satisfy any of the three equivalent conditions of Theorem 9, we shall say that α is more *informative than* β for k -decision problems, written $\alpha \succ_k \beta$. Condition (2) is the direct analogue of \supset , and condition (1) is analogous to \succ , since it requires that every experiment with k outcomes producible from β is also producible from α . Clearly \succ_{k+1} implies \succ_k , and if $\alpha \succ_k \beta$ for all k , then $\alpha \succ \beta$, since $\alpha \succ_k \beta$ for all k implies $\int \phi dm_\alpha \geq \int \phi dm_\beta$ for every ϕ which is the maximum of a finite number of linear functions and hence, by approximation, for every continuous convex ϕ . An alternative statement is: if every experiment with a finite number of outcomes which is producible from β is also producible from α , then β is itself producible from α .

Stein (unpublished paper) has shown that in general \succ_{k+1} is actually stronger than \succ_k . For $n = 2$, however, all \succ_k for $k \geq 2$ are equivalent.

THEOREM 10. *If α and β are two experiments with $n = 2$, then $\alpha \succ_2 \beta$ implies $\alpha \succ \beta$.*

PROOF. For $n = 2$, the standard measures m_α and m_β are defined on the line segment $p_i \geq 0, p_1 + p_2 = 1$. On this line segment, every function ϕ which is the

maximum of a finite number of linear functions is representable as $\sum a_i \phi_i$, where $a_i > 0$ and each ϕ_i is a maximum of two linear functions. Consequently $\alpha >_2 \beta$ implies $\alpha >_k \beta$ for all k and hence $\alpha > \beta$.

COROLLARY. Let A be the line segment joining $(0, 1)$ and $(1, 0)$. If $B(\alpha, A) \supset B(\beta, A)$, then $\alpha > \beta$.

PROOF. For any line segment A' in the plane, there is a transformation

$$\begin{aligned} L: \quad x' &= ax + b \\ y' &= cx + d \end{aligned}$$

with $LA = A'$. Since $LB(\alpha, A) = B(\alpha, LA)$ and similarly for β , we have $B(\alpha, A') \supset B(\beta, A')$, so that $\alpha >_2 \beta$ and consequently $\alpha > \beta$.

For the A of the corollary, the boundary of the set $B(\alpha, A)$ consists of two curves, joining $(0, 1)$ and $(1, 0)$, one of which is the reflection of the other about $(1/2, 1/2)$. Denote by $f_\alpha(t)$ the minimum of u for which $(t, u) \in B(\alpha, A)$. Then $\alpha > \beta$ if and only if $f_\alpha(t) \leq f_\beta(t)$ for all t , $0 \leq t \leq 1$. The function $f_\alpha(t)$ is a nonincreasing convex function of t , representing the minimum attainable error of the second kind when the error of the first kind is fixed at t . Thus an alternative statement of the corollary is:

α is more informative than β if and only if at every level t the error of the second kind with α is less than or equal to the corresponding error with β .

Since if $\alpha > \beta$, then an experiment with n independent observations with α is more informative than the corresponding experiment with β [1] we obtain

THEOREM 11. If for a sample of size 1 at every level t the probability of an error of the second kind with α does not exceed the corresponding probability for β , then the same is true for every sample size.

REFERENCES

- [1] DAVID BLACKWELL, "Comparison of experiments," *Proceedings of the Second Berkeley Symposium on Mathematical Statistics and Probability*, University of California Press, 1951, pp. 93-102.
- [2] DAVID BLACKWELL, "The range of certain vector integrals," *Proc. Amer. Math. Soc.*, Vol. 2 (1951), pp. 390-395.
- [3] H. F. BOHNENBLUST, S. KARLIN, AND L. S. SHAPLEY, "Games with continuous convex payoff," *Contributions to the Theory of Games*, Princeton University Press, 1950, pp. 181-192.
- [4] H. F. BOHNENBLUST, L. S. SHAPLEY, AND S. SHERMAN, Unpublished paper.
- [5] J. L. DOOB, "Continuous parameter martingales," *Proceedings of the Secondary Berkeley Symposium on Mathematical Statistics and Probability*, University of California Press, 1951, pp. 269-277.
- [6] G. H. HARDY, J. E. LITTLEWOOD, AND G. POLYA, *Inequalities*, Cambridge University Press, 1934.
- [7] A. N. KOLMOGOROV, "Grundbegriffe der wahrscheinlichkeitsrechnung," *Ergebnisse der Mathematik*, No. 3, 1933.
- [8] S. SHERMAN, "On a theorem of Hardy, Littlewood, Polya, and Blackwell," *Proc. Nat. Acad. Sci.*, U.S.A., Vol. 37 (1951), pp. 826-831.
- [9] CHARLES STEIN, "Notes on the Comparison of Experiments," (mimeographed), University of Chicago, 1951.