

# ABSOLUTE CONTINUITY OF INFINITELY DIVISIBLE DISTRIBUTIONS

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**1. Introduction and summary.** A probability distribution function  $F$  is said to be infinitely divisible if and only if for every integer  $n$  there is a distribution function  $F_n$  whose  $n$ -fold convolution is  $F$ . If  $F$  is infinitely divisible, its characteristic function  $f$  is necessarily of the form

$$(1) \quad f(u) = \exp \left\{ iu\gamma + \int_{-\infty}^{\infty} \left( e^{iux} - 1 - \frac{iux}{1+x^2} \right) \frac{1+x^2}{x^2} dG(x) \right\},$$

where  $u \in (-\infty, \infty)$ ,  $\gamma$  is some constant, and  $G$  is a bounded, non-decreasing function. J. R. Blum and M. Rosenblatt [1] have found necessary and sufficient conditions that  $F$  be continuous and necessary and sufficient conditions that  $F$  be discrete. The purpose of this note is to add to the results of Blum and Rosenblatt by giving sufficient conditions under which an infinitely divisible probability distribution  $F$  is absolutely continuous. These conditions are that  $G$  be discontinuous at 0 or that  $\int_{-\infty}^{\infty} (1/x^2) dG_{ac}(x) = \infty$ , where  $G_{ac}$  is the absolutely continuous component of  $G$ . In § 2 some lemmas will be proved, and in § 3 the proof of the sufficiency of these conditions will be given. All notation used here is standard and may be found, for example, in Loève [2].

**2. Some lemmas.** In this section three lemmas are proved which will be used in the following section.

**LEMMA 1.** *If  $F$  and  $H$  are probability distribution functions, and if  $F$  is absolutely continuous, then the convolution of  $F$  and  $H$ ,  $F * H$ , is absolutely continuous.*

This lemma is well known, and the proof is omitted.

**LEMMA 2.** *If  $\{F_n\}$  is a sequence of absolutely continuous distribution functions, and if  $p_n \geq 1$  and  $\sum_{n=1}^{\infty} p_n = 1$ , then  $\sum_{n=1}^{\infty} p_n F_n$  is an absolutely continuous distribution function.*

*Proof.* By using the Lebesgue monotone convergence theorem it is easy to verify that  $\sum_{n=1}^{\infty} p_n f_n$  is the density of  $\sum_{n=1}^{\infty} p_n F_n$ , where  $f_n$  is the density of  $F_n$ .

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**LEMMA 3.** *Let  $\{Y, X_1, X_2, \dots\}$  be independent random variables. Assume that the  $X_i$ 's have the same absolutely continuous distribution  $F$ , and assume that the distribution of  $Y$  is Poisson with expectation  $\lambda$ . Then  $Z = X_1 + \dots + X_Y$  has a distribution function which has a saltus  $e^{-\lambda}$  at 0 and is absolutely continuous elsewhere, and has as characteristic function*

$$f_Z(u) = \exp \lambda \int_{-\infty}^{\infty} (e^{iux} - 1) dF(x) .$$

*Proof.* Let  $E(x)$  be the distribution function degenerate at 0, and let  $F^{*n}(x)$  denote the convolution of  $F$  with itself  $n$  times. Then it is easy to see that the distribution function of  $Z, F_Z(z)$ , may be written as  $F_Z(z) = e^{-\lambda}E(z) + \sum_{n=1}^{\infty} e^{-\lambda}(\lambda^n/n!)F^{*n}(z)$ . By lemma 1, each  $F^{*n}$  is absolutely continuous and has a density  $f^{*n}$ . We need only show that  $F_Z(z) - e^{-\lambda}E(z)$  is absolutely continuous. If we write

$$F_Z(z) - e^{-\lambda}E(z) = \sum_{n=1}^{\infty} e^{-\lambda}(\lambda^n/n!) \int_{-\infty}^z f^{*n}(t) dt$$

and apply the Lebesgue monotone convergence theorem we obtain this conclusion.

**3. The theorem.** If  $G$  is a bounded nondecreasing function used in (1), then we may write  $G(x) = G_s(x) + G_{ac}(x)$ , where  $G_s$  is a singular nondecreasing function and  $G_{ac}(x)$  is an absolutely continuous nondecreasing function.

**THEOREM.** *Let  $F$  be an infinitely divisible distribution function with characteristic function (1). Then  $F$  is absolutely continuous if at least one of the following two conditions is satisfied:*

- (i)  $G$  is not continuous at 0, or
- (ii)  $\int_{-\infty}^{\infty} (1/x^2) dG_{ac}(x) = \infty$ .

*Proof.* If condition (i) is satisfied, then by Lemma 1 it easily follows that  $F$  is absolutely continuous, since in that case  $F$  is a convolution of a normal distribution with another infinitely divisible distribution. We now prove that condition (ii) is sufficient. By Lemma 1 it is sufficient to prove that the distribution function  $F_0$  whose characteristic function is

$$(2) \quad \exp \int_{-\infty}^{\infty} \left( e^{iux} - 1 - \frac{iux}{1+x^2} \right) \frac{1+x^2}{x^2} dG_{ac}(x)$$

is absolutely continuous. Let  $\epsilon_n > \epsilon_{n+1} > 0$  for each  $n$  be such that  $\epsilon_n \rightarrow 0$  as  $n \rightarrow \infty$  and such that

$$\lambda_n = \int_{S_n} ((1 + x^2)/x^2) dG_{ac}(x) > 0 ,$$

where

$$S_n = (-\varepsilon_{n-1}, -\varepsilon_n] \cup [\varepsilon_n, \varepsilon_{n-1}) , \quad n = 1, 2, \dots ,$$

and where  $\varepsilon_0 = \infty$ . Let  $U_n$  be a random variable with characteristic function

$$(3) \quad f_{U_n}(u) = \exp \int_{S_n} \left( e^{iux} - 1 - \frac{iux}{1+x^2} \right) \frac{1+x^2}{x^2} dG_{ac}(x) ,$$

and let

$$H_n(x) = (1/\lambda_n) \int_{(-\infty, x] \cap S_n} ((1 + x^2)/x^2) dG_{ac}(x) .$$

One easily sees that  $\lambda_n < \infty$  and that  $H_n(x)$  is an absolutely continuous distribution function of a bounded random variable. For each positive integer  $n$  we may write, by Lemma 3, that

$$U_n = X_{n,1} + X_{n,2} + \dots + X_{n,Z_n} - \int_{S_n} (1/x) dG_{ac}(x)$$

where  $Z_n$  is a random variable with Poisson distribution with expectation  $\lambda_n$ , where  $\{X_{n,1}, X_{n,2}, \dots\}$  have the common absolutely continuous distribution function  $H_n(x)$ , and where  $\{Z_n, X_{n,1}, X_{n,2}, \dots\}$  are independent. If we assume that

$$\{\{Z_n, X_{n,1}, X_{n,2}, \dots\}, n = 1, 2, \dots\}$$

are all independent, then the distribution function of

$$U_0 = \sum_{n=1}^{\infty} U_n = \sum_{n=1}^{\infty} \left( \sum_{j=1}^{Z_n} X_{n,j} - \int_{S_n} (1/x) dG_{ac}(x) \right)$$

is equal to  $F_0$ . Now let us define a sequence of events  $\{C_n\}$  by

$$C_1 = [Z_1 \neq 0] , \quad C_2 = [Z_1 = 0][Z_2 \neq 0] ,$$

and, in general,

$$C_n = [Z_n \neq 0] \bigcap_{i=1}^{n-1} [Z_i = 0] .$$

These events are easily seen to be disjoint. If we define

$$(4) \quad C_0 = \left( \bigcup_{n=1}^{\infty} C_n \right)^c = \bigcap_{n=1}^{\infty} [Z_n = 0] ,$$

then  $\Omega = \bigcup_{n=1}^{\infty} C_n$ , where  $\Omega$  is the sure event. The distribution function of  $U_0$  is

$$F_{U_0}(u) = \sum_{n=1}^{\infty} P([U_0 \leq u] | C_n) P(C_n) + P([U_0 \leq u] | C_0).$$

By (4) and by hypothesis, we obtain

$$P([U_0 \leq u] | C_0) \leq P(C_0) = \lim_{n \rightarrow \infty} \exp \left\{ - \int_{-\infty}^{-\varepsilon_n} + \int_{\varepsilon_n}^{\infty} (1/x^2) dG_{ac}(x) \right\} = 0.$$

Also,  $P([U_0 \leq u] | C_n)$  is the distribution function of  $X_{n,1} + W_n$ , where  $X_{n,1}$  and  $W_n$  are independent random variables. Since the distribution function of  $X_{n,1}$  is absolutely continuous, it follows by Lemma 1 that  $P([U_0 \leq u] | C_n)$  is absolutely continuous for each  $n$ . Lemma 2 then implies that  $F_{U_0}(u)$  is absolutely continuous, which concludes the proof of the theorem.

The condition given in this theorem is not necessary, as is shown in the following example. Let  $\gamma = 0$  in (1), and let  $\alpha, \beta$  be real numbers which satisfy  $\beta > 1, 1 > \alpha > \beta/2$ . For  $j = 1, 2, \dots$ , let us denote

$$x_j = j^{-\alpha} \quad \text{and} \quad \rho_j = j^{-\beta}.$$

Let  $G$  be a pure jump function with jumps at  $x_j$  and  $-x_j$  of size  $\rho_j$  for every  $j$ . (The total variation of  $G$  is  $2 \sum \rho_j < \infty$ .) In this case we obtain

$$f(u) = \exp 2 \sum_{n=1}^{\infty} \left( \cos \frac{u}{n^\alpha} - 1 \right) \frac{n^{2\alpha} - 1}{n^\beta}.$$

We shall show that there is a constant  $K$  such that for all  $|u| \geq \pi$ , the inequality

$$(5) \quad 0 < f(u) < \exp(-K|u|^{2-\beta/\alpha})$$

is true. This is equivalent to showing that

$$(6) \quad \sum_{n=1}^{\infty} \frac{n^{2\alpha} + 1}{n^\beta} \sin^2 \frac{|u|}{2n^\alpha} > K|u|^{2-\beta/\alpha}.$$

Let us consider, for each fixed  $|u| \geq \pi$  the integer  $N$  defined by

$$N = \left[ \frac{1}{2} \left( \frac{2|u|}{\pi} \right)^{1/\alpha} + 1 \right],$$

where the square brackets have their usual meaning. It is easy to verify that  $0 < |u|/2N^\alpha < \pi/2$ , and thus we may write

$$\begin{aligned} \frac{N^{2\alpha} + 1}{N^\beta} \sin^2 \frac{|u|}{2N^\alpha} &> N^{2\alpha-\beta} \sin^2 \frac{|u|}{2 \left[ \left( \frac{2|u|}{\pi} \right)^{1/\alpha} \right]^\alpha} \\ &> K|u|^{2-\beta/\alpha}, \end{aligned}$$

where  $K$  does not depend on  $u$ . This inequality implies that inequality (6) is true, thus implying (5). Inequality (5) implies that  $f(u) \in L_1(-\infty, +\infty)$ , which in turn implies that  $f(u)$  is the characteristic function of an absolutely continuous distribution. (See Theorem 3.2.2 on page 40 in [3].)

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#### REFERENCES

1. J. R. Blum and M. Rosenblatt, *On the structure of infinitely divisible distributions*, Pacific J. Math., **9** (1959), 1-7.
2. M. Loève, *Probability Theory*, D. Van Nostrand, Princeton, 1960 (Second Edition).
3. Eugene Lukacs, *Characteristic Functions*, Hafner, New York, 1960.

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