

## Note on Random Riemann Sum.

By Makiko NISIO

(Received July 6, 1957)

§ 1. Let  $f \in L_p(a, b)$  and consider the Riemann sum  $S_n \equiv S_n(f; t_1, \dots, t_n)$  for random division points  $t_i$ . It was the idea of P. Lévy [1] to define 'generalized integrals' as 'limits' of such Riemann sums. S. Takahashi [2] proved that  $S_n$  converges to  $\int_0^1 f(t) dt$  with probability 1, for  $p > 1$  and in probability for  $p = 1$ , if the  $t_i$ ,  $1 \leq i < \infty$ , are independent with the uniform distribution on  $(0, 1)$ .

We shall now consider the case  $p = 1$ ,  $a = 0$ ,  $b = \infty$ , taking, for each  $n \geq 1$ , division-points  $t_i^n$  with probability  $n dt$  in  $(t, t + dt)$  and mutually independently in non-overlapping time-intervals. Rigorously speaking, we adopt the jumping times of the Poisson process with parameter  $n$  as the division-points. We then let  $n$  tend to  $\infty$ .

In § 2 we shall prove that if  $f \in L_1(0, \infty)$ ,  $S_n$  converges to the Lebesgue integral of  $f$  over  $(0, \infty)$  in probability as  $n$  tends to infinity and in § 3 that if  $f \in L_1(0, \infty) \cap L_2(0, \infty)$ , the subsequence  $S_{2^n}$  converges to the same value with probability one.

It should be noticed that our way of picking the division-points by means of Poisson process has made much easier the analytical treatment compared with the case treated by Takahashi [2].

The author wishes to express her sincere thanks to Professor K. Ito for his valuable suggestions.

§ 2. Let  $(\Omega, B, P)$  be the probability space on which we define Poisson processes  $\{P_n(t, \omega), t \in (0, \infty)\}$ ,  $n$  being the mean value of  $P_n(1, \omega)$ ; it makes no difference here whether these processes are independent or dependent. Let  $t_i^n(\omega)$  be the  $i$ -th jumping point of the Poisson process  $P_n(t, \omega)$ .

In this note we shall often use the following well-known fact.

LEMMA.  $\{t_{i+1}^n(\omega) - t_i^n(\omega)\}$ ,  $i = 0, 1, 2, \dots$  ( $t_0^n(\omega) \equiv 0$ ) are independent random variables with the following probability law:

$$(2.1) \quad P\{t_{i+1}^n(\omega) - t_i^n(\omega) < t\} = 1 - e^{-nt} \quad (t \geq 0),$$

and so we have

$$(2.2) \quad P\{t_{i+j}^n(\omega) - t_j^n(\omega) \in (t, t + dt)\} = \frac{(nt)^{i-1} e^{-nt}}{(i-1)!} n dt$$

Now we shall prove the following

**THEOREM 1.** *If  $f \in L_1(0, \infty)$ , then  $S_n(\omega) = \sum_{i=1}^{\infty} [f(t_i^n(\omega)) \cdot (t_{i+1}^n(\omega) - t_i^n(\omega))]$  converges to  $\int_0^{\infty} f(t) dt$  in probability as  $n$  tends to infinity.*

**PROOF.** Keeping the above Lemma in mind, we have

$$E|S_n(\omega) - \int_0^{\infty} f(t) dt| \leq E\left|\sum_{i=1}^{\infty} \int_{t_i^n(\omega)}^{t_{i+1}^n(\omega)} (f(t_i^n(\omega)) - f(t)) dt\right| + E\left|\int_0^{t_1^n(\omega)} f(t) dt\right|$$

$$\equiv A_n + B_n.$$

$$A_n \leq \sum_{i=1}^{\infty} E\left\{\int_{t_i^n(\omega)}^{t_{i+1}^n(\omega)} |f(t_i^n(\omega)) - f(t)| dt\right\}$$

$$= \sum_{i=1}^{\infty} \int_{s=0}^{\infty} e^{-ns} \frac{(ns)^{i-1}}{(i-1)!} n ds \int_{h=0}^{\infty} e^{-nh} n dh \int_{t=s}^{s+h} |f(s) - f(t)| dt$$

$$= \int_{s=0}^{\infty} n ds \int_{h=0}^{\infty} e^{-nh} n dh \int_{t=0}^h |f(t+s) - f(s)| dt$$

$$= \int_{s=0}^{\infty} dt \int_{t=0}^{\infty} |f(t+s) - f(s)| dt \int_{h=t}^{\infty} n^2 e^{-nh} dh$$

$$= \int_{s=0}^{\infty} ds \int_{t=0}^{\infty} |f(t+s) - f(s)| n e^{-nt} dt$$

$$= \int_0^{\infty} \|f_t - f\| n e^{-nt} dt = \int_0^{\infty} \|f_{t/n} - f\| e^{-t} dt \rightarrow 0 \quad (n \rightarrow \infty)$$

where  $f_t(s) \equiv f(t+s)$  and  $\| \cdot \|$  means the norm in  $L_1(0, \infty)$ .

$$B_n \leq \int_0^{\infty} e^{-nh} n dh \int_0^h |f(t)| dt = \int_0^{\infty} e^{-h} dh \int_0^{h/n} |f(t)| dt \rightarrow 0 \quad (n \rightarrow \infty)$$

Hence we have  $p\text{-}\lim_{n \rightarrow \infty} S_n(\omega) = \int_0^{\infty} f(t) dt$ .

**§ 3. THEOREM 2.** *For  $f \in L_1(0, \infty) \cap L_2(0, \infty)$ ,  $S_{2^n}(\omega)$  converges to  $\int_0^{\infty} f(t) dt$  with probability one as  $n$  tends to infinity.*

**PROOF.** Let us put  $k = \int_0^{\infty} f^2(t) dt$ ,  $k' = \int_0^{\infty} |f(t)| dt$  and  $I = \int_0^{\infty} f(t) dt$ . Making use of the above Lemma, we shall estimate  $ES_{2^n}^2(\omega)$ .

$$\begin{aligned}
ES_n^2(\omega) &= E\left(\sum_{i=1}^{\infty} (t_{i+1}^n(\omega) - t_i^n(\omega))^2 f^2(t_i^n(\omega))\right) \\
&\quad + 2E\left(\sum_{i=1}^{\infty} \sum_{j=i+2}^{\infty} (t_{j+1}^n(\omega) - t_j^n(\omega))(t_{i+1}^n(\omega) - t_i^n(\omega)) f(t_j^n(\omega)) f(t_i^n(\omega))\right) \\
&\quad + 2E\left(\sum_{i=1}^{\infty} (t_{i+2}^n(\omega) - t_{i+1}^n(\omega))(t_{i+1}^n(\omega) - t_i^n(\omega)) f(t_{i+1}^n(\omega)) f(t_i^n(\omega))\right) \\
&\equiv A_n + B_n + C_n. \\
A_n &= \sum_{i=1}^{\infty} E(t_{i+1}^n(\omega) - t_i^n(\omega))^2 f^2(t_i^n(\omega)) = \sum_{i=1}^{\infty} \left[ \int_0^{\infty} f^2(t) e^{-nt} \frac{(nt)^{i-1}}{(i-1)!} n dt \cdot \right. \\
&\quad \left. \int_0^{\infty} h^2 e^{-nh} n dh \right] = 2k/n. \\
B_n &= 2 \sum_{i=1}^{\infty} \sum_{j=i+2}^{\infty} \left[ \int_0^{\infty} f(t) e^{-nt} \frac{(nt)^{i-1}}{(i-1)!} n dt \int_0^{\infty} h e^{-nh} n dh \cdot \right. \\
&\quad \left. \int_{t+h}^{\infty} f(s) e^{-n(s-t-h)} \frac{\{n(s-t-h)\}^{j-i-2}}{(j-i-2)!} n ds \int_0^{\infty} u e^{-nu} n du \right].
\end{aligned}$$

Since we can exchange the integral and summation by virtue of  $f \in L_1(0, \infty)$ , we have

$$\begin{aligned}
B_n &= 2 \int_0^{\infty} f(t) n dt \int_0^{\infty} h e^{-nh} n dh \int_{t+h}^{\infty} f(s) ds \\
&= 2 \int_0^{\infty} f(t) dt \int_t^{\infty} f(s) ds - 2 \int_0^{\infty} f(t) n dt \int_0^{t+h} h e^{-nh} n dh \int_t^{t+h} f(s) ds = B'_n - B''_n. \\
B'_n &= \left( \int_0^{\infty} f(t) dt \right)^2 = I^2.
\end{aligned}$$

By Schwarz' inequality we have  $|\int_t^{t+h} f(s) ds| \leq (kh)^{1/2}$ , so that we get

$$|B''_n| = 2 \left| \int_0^{\infty} f(t) n dt \int_0^{\infty} h e^{-nh} n dh \int_t^{t+h} f(s) ds \right| \leq 2k^{1/2} k' \int_0^{\infty} h^{3/2} n^2 e^{-nh} dh = O(n^{-1/2}).$$

Therefore we have

$$\begin{aligned}
B_n &= I^2 + O(n^{-1/2}). \\
C_n &= 2 \sum_{i=1}^{\infty} \int_0^{\infty} h' e^{-nh'} n dh' \int_0^{\infty} f(t) e^{-nt} \frac{(nt)^{i-1}}{(i-1)!} n dt \int_0^{\infty} h f(t+h) e^{-nh} n dh \\
&= \frac{2}{n} \int_0^{\infty} f(t) n dt \int_0^{\infty} h f(t+h) e^{-nh} n dh
\end{aligned}$$

$$\begin{aligned}
 &= \frac{2}{n} \int_0^\infty f(t) dt \int_0^\infty h f(t+h/n) e^{-h} dh \\
 &= \frac{2}{n} \int_0^\infty \int_0^\infty f(t) f(t+h/n) dt h e^{-h} dh
 \end{aligned}$$

Using Schwarz' inequality, we have

$$|C_n| \leq \frac{2}{n} \int_0^\infty \left( \int_0^\infty f^2(t) dt \right)^{1/2} \left( \int_0^\infty f^2(t+h/n) dt \right)^{1/2} h e^{-h} dh \leq 2k/n.$$

Summing up the above estimation, we obtain

$$ES_n^2(\omega) = I^2 + O(n^{-1/2}).$$

On the other hand

$$\begin{aligned}
 ES_n(\omega) &= \sum_{i=1}^\infty \int_0^\infty f(t) e^{-nt} \frac{(nt)^{i-1}}{(i-1)!} n dt \int_0^\infty h e^{-nh} n dh \\
 &= n \int_0^\infty f(t) dt \int_0^\infty h e^{-nh} n dh \\
 &= \int_0^\infty f(t) dt = I.
 \end{aligned}$$

Thus we have

$$E(S_n(\omega) - I)^2 = ES_n^2(\omega) - I^2 = O(n^{-1/2}),$$

so that

$$E \left\{ \sum_{n=1}^\infty (S_{2^n}(\omega) - I)^2 \right\} < \infty,$$

which implies

$$\sum_{n=1}^\infty (S_{2^n}(\omega) - I)^2 < \infty$$

and so  $S_{2^n}(\omega) \rightarrow I$  with probability one.

Mathematics Department  
Kobe University.

### References

- [ 1 ] P. Lévy, Integrales de Stieltjes généralisées, Ann. Soc. Polon. Math., 25 (1952), 17-26.
- [ 2 ] S. Takahasi, On some Random Riemann-Sums, Tôhoku. Math. Jour., 8 (1956), 245-257.