

ASYMPTOTIC EXPECTED INFERIOR SAMPLE SIZE OF A SEQUENTIAL TEST INVOLVING TWO POPULATIONS

BY H. H. PETER CHENG

National Central University

Let X_1, \dots be i.i.d. $\sim N(\mu_1, 1)$ and Y_1, \dots be i.i.d. $\sim N(\mu_2, 1)$. A symmetric sequential procedure for $H_0: \mu_1 > \mu_2$ vs. $H_1: \mu_1 < \mu_2$ is proposed in this paper. The expected number of observations taken from the inferior population is given in an asymptotic form, which is optimum in Farrell's sense.

1. Introduction and Summary. Let X_1, \dots be i.i.d. $\sim N(\mu_1, 1)$ and Y_1, \dots be i.i.d. $\sim N(\mu_2, 1)$. A sequential procedure for testing $H_0: \mu_1 > \mu_2$ vs. $H_1: \mu_1 < \mu_2$ consists of a terminal decision rule for choosing H_0 or H_1 , a stopping rule for deciding when to stop, and a sampling rule for deciding (before stopping) whether the next observation should be an X or a Y . Let us assume that the X population is inferior if H_1 is true and the Y population is inferior if H_0 is true. It is desirable to have a procedure not only with controlled error probabilities but with minimized expected number of observations taken from the inferior population. This kind of problem has been considered by Robbins and Siegmund [5].

A sequential procedure is proposed in Section 2 and its symmetric nature is discussed in Section 4. Section 5 gives other properties of the procedure. In Section 6 the expected number of observations taken from the inferior population is estimated asymptotically as $|\mu_2 - \mu_1| \rightarrow 0$. The optimum property of this procedure in the sense of Farrell [1] is discussed in Section 7.

2. The sampling rule, the stopping rule, and the terminal decision rule. Let $f(y) = (|y| \log_2 |y|^{-1} (\log_2 |y|^{-1})^{\frac{3}{2}})^{-1}$ for $0 < |y| < e^{-e}$ and $f(y) = 0$ elsewhere. We write $\log_2 x = \log \log x$, $\log_3 x = \log \log_2 x$, etc. Define $H(x, t) = \int_{-\infty}^{\infty} \exp(xy - (t/2)y^2) f(y) dy$. Given $0 < \alpha < \frac{1}{2}$, set

$$(1) \quad b = (2\alpha)^{-1}$$

$$(2) \quad A(t) = A(t, b) = \inf\{x > 0 : H(x, t) \geq b\}.$$

From (2) clearly

$$(3) \quad H(x, t) \geq b \quad \text{if and only if} \quad |x| \geq A(t).$$

By the dominated convergence theorem, $b = \int_{-\infty}^{\infty} \exp(A(t)y - (t/2)y^2) f(y) dy = \int_0^{\infty} \exp(A(t)y - (t/2)y^2) f(y) dy + o(1)$ as $t \rightarrow \infty$. Arguing as Section 4 of [3], we have

$$(4) \quad A(t) = (2t \log_2 t + 4t \log_3 t + \log b + \log(\pi)^{\frac{1}{2}} + o(1))^{\frac{1}{2}} \text{ as } t \rightarrow \infty.$$

Received February 1978; revised December 1978.

AMS 1970 subject classifications. Primary 62L10; secondary 60G40.

Key words and phrases. Sequential tests, expected inferior sample size.

LEMMA 1. Let $W(t), t \geq 0$ be the standard Wiener process; then $P(H(W(t), t) \geq b, \text{ for some } t > 0) = b^{-1}$

PROOF. Denote $z(t) = H(W(t), t)$. It can be verified that $\{z(t), \beta(W(s), s \leq t), t \geq 0\}$ is a martingale. (cf. [3], Section 3). For any $\delta > 0, P(z(t) > \delta) = P(|W(t)| > A(t, \delta)) \rightarrow 0$ as $t \rightarrow \infty$ by (3) and (4). Since $z(0) = 1$ and $b > 1$, this lemma is proven by Lemma 1 in [3].

Let M_ν and N_ν denote the number of x 's and the number of y 's observed at stage ν of the experiment. Hence $M_\nu + N_\nu = \nu$. Set $\bar{X}_{M_\nu} = \sum_1^{M_\nu} X_i / M_\nu, \bar{Y}_{N_\nu} = \sum_1^{N_\nu} Y_i / N_\nu, t_\nu = M_\nu N_\nu / \nu$, and

$$(5) \quad Z_\nu = Z_{M_\nu, N_\nu} = t_\nu (\bar{Y}_{N_\nu} - \bar{X}_{M_\nu}).$$

The stopping rule is defined:

$$(6) \quad T = \text{First } \nu \geq 2 \text{ such that } |Z_\nu| \geq A(t_\nu); T = \infty \text{ if no such } \nu \text{ occurs.}$$

Set $\eta(\nu) = (1 + (\log_3 \nu - 1)^+)^{-1}$ for $\nu \geq 2$. Define the sampling rule: observe X_1, Y_1 first. When $T > \nu$ at stage ν , observe a Y if

$$(7) \quad (N_\nu - M_\nu) / \nu \leq Z_\nu [A(t_\nu)(1 + \eta(\nu))]^{-1};$$

otherwise observe an X .

The terminal decision rule is defined on $(T < \infty)$: assert H_0 if $Z_T < 0$; assert H_1 if $Z_T > 0$.

$$(8) \quad \begin{aligned} \text{On } (T < \infty) \text{ set } M &= M_T \text{ and } N = N_T; \\ \text{on } (T = \infty) \text{ set } M &= \lim_{\nu \rightarrow \infty} M_\nu \text{ and } N = \lim_{\nu \rightarrow \infty} N_\nu. \end{aligned}$$

Set $\lambda(\nu) = \eta(\nu) / 2(1 + \eta(\nu))$ for $\nu \geq 2$. Clearly $\lambda(\nu) \downarrow 0$ and $\nu\lambda(\nu) \uparrow \infty$ as $\nu \rightarrow \infty$.

LEMMA 2. $(T > \nu, M_\nu \leq \nu\lambda(\nu)) \subset (T > \nu, M_{\nu+1} > M_\nu) \subset (T > \nu, N_\nu > \nu\lambda(\nu))$.

PROOF. Cf. (5.4) and (5.13) in [3].

From Lemma 2 it is easily seen

$$(9) \quad M_\nu \geq \frac{\nu}{2}\lambda(\nu) \text{ and } N_\nu \geq \frac{\nu}{2}\lambda(\nu) \text{ on } (T \geq \nu) \text{ for all } \nu \geq 2.$$

Hence $M_\nu N_\nu / \nu \geq (\nu/4)\lambda(\nu) \rightarrow \infty$ as $\nu \rightarrow \infty$ and $(T = \infty) \subset (\lim_{\nu \rightarrow \infty} (M_\nu N_\nu) / \nu = \infty)$. Since $\nu \geq M_\nu N_\nu / \nu, (\lim_{\nu \rightarrow \infty} M_\nu N_\nu / \nu = \infty) \subset (T = \infty)$. Denoting $\lim_{\nu \rightarrow \infty} (M_\nu N_\nu) / \nu$ by $MN / (M + N)$ on $(T = \infty)$, we have

LEMMA 3. $(T = \infty) = (MN / (M + N) = \infty)$.

3. Wiener process with drift d . Let $d = \mu_2 - \mu_1$. The definition of a Wiener process with drift d is in [5] and the following Lemma 4 is the Lemma 1 of [5].

LEMMA 4. Let $X_1, \dots; Y_1, \dots$ be as in Section 1, $Z_{m,n} = (m \sum_1^n Y_i - n \sum_1^m X_j) / (m + n), d = \mu_2 - \mu_1$, and $\{W_d(t), t \geq 0\}$ be a Wiener process with drift d .

For any sequence of pairs (m, n) of positive integers which is nondecreasing in each coordinate, the sequences $\{Z_{m, n}\}$ and $\{W_d(mn/(m + n))\}$ have the same distribution.

Lemma 4 suggests that the sequential procedure proposed in Section 2 can be considered in terms of Wiener process with drift $d = \mu_2 - \mu_1$. To be precise, the following are defined: $\bar{M}_2 = \bar{N}_2 = 1$. For $\nu \geq 2$, define $\bar{N}_{\nu+1} = \bar{N}_\nu + 1$ if

$$(10) \quad (\bar{N}_\nu - \bar{M}_\nu)/\nu < \frac{W_d(\bar{t}_\nu)}{A(\bar{t}_\nu)(1 + \eta(\nu))};$$

otherwise $\bar{N}_{\nu+1} = \bar{N}_\nu$. Set $\bar{M}_\nu = \nu - \bar{N}_\nu$, $\bar{t}_\nu = \bar{M}_\nu \bar{N}_\nu / \nu$.

$$(11) \quad \bar{T} = \text{First } \nu \geq 2 \text{ such that } |W_d(\bar{t}_\nu)| \geq A(\bar{t}_\nu); \quad \bar{T} = \infty \text{ otherwise.}$$

\bar{M} and \bar{N} are similarly defined as in (8). Lemma 3 becomes

$$(12) \quad (\bar{T} = \infty) = (\bar{t}_{\bar{T}} = \infty).$$

The $M_\nu, N_\nu, \bar{M}_\nu, \bar{N}_\nu$ are random variables; the m, n in Lemma 4 are constants. By sampling rule (7), $M_{\nu+1}$ and $N_{\nu+1}$ are $\beta(Z_i, i \leq \nu)$ -measurable; by (10) $\bar{M}_{\nu+1}$ and $\bar{N}_{\nu+1}$ are $\beta(W_d(\bar{t}_i), i \leq \nu)$ -measurable. Using Lemma 4, it is straightforward to justify the following

LEMMA 5. $\{Z_\nu, \nu \geq 2\}$ and $\{W_d(\bar{t}_\nu), \nu \geq 2\}$ have the same distribution.

4. The symmetry of the procedure. Let $1 = m_2 \leq m_3 \dots$, and $1 = n_2 \leq n_3 \dots$ be any sequences of positive integers such that $m_\nu + n_\nu = \nu$. Let $d = \mu_2 - \mu_1$ and $\delta > 0$ be any positive number. By Lemma 4 $\{Z_{m_\nu, n_\nu}\}$ in the case $d = \delta$ has the same distribution as $\{-Z_{n_\nu, m_\nu}\}$ in the case $d = -\delta$. By (7) it is direct to justify $P_\delta(M_i = m_i, N_i = n_i, i \leq \nu; |Z_{m_i, n_i}| < A(s_i), i \leq \nu - 1; Z_{m_\nu, n_\nu} \geq A(s_\nu)) = P_{-\delta}(M_i = n_i, N_i = m_i, i \leq \nu; |Z_{n_i, m_i}| < A(s_i), i \leq \nu - 1; Z_{n_\nu, m_\nu} \leq -A(s_\nu))$, where $s_i = m_i n_i / i$. From the above equality the next lemma follows easily.

LEMMA 6. $P_\delta(M = k) = P_{-\delta}(N = k), E_\delta(M) = E_{-\delta}(N);$
 $P_\delta(T = \nu) = P_{-\delta}(T = \nu), P_\delta(T < \infty) = P_{-\delta}(T < \infty);$
 $P_\delta(\text{reject } H_1) = P_{-\delta}(\text{reject } H_0).$

In the following sections, only $d = \mu_2 - \mu_1 = \delta > 0$ and $d = 0$ are considered.

5. Other properties of the procedure. By the strong law of large numbers $P(\lim_{t \rightarrow \infty} W_\delta(t)/t = \delta) = 1$. By (4) $\lim_{t \rightarrow \infty} A(t)/t = 0$. With Lemma 5 and Lemma 6 we have

LEMMA 7. $P_\delta(T < \infty) = 1$.

Note $P_\delta(\bar{T} < \infty) = 1$ by Lemma 5. These kinds of paralleled results will not be stated explicitly in the following discussion.

LEMMA 8. $P_0(T = \infty) > 0; P_0(\bar{t}_{\bar{T}} = \infty) > 0$.

PROOF. By (12), $P_0(T = \infty) = P_0(\bar{T} = \infty) = P_0(\bar{t}_{\bar{T}} = \infty) = P(|W_0(t_\nu)| < A(t_\nu), \text{ for all } \nu \geq 2) \geq P(H(W_0(t), t) < b, \text{ for all } t \geq 0) = 1 - 1/b = 1 - 2\alpha > 0$, by (3) and Lemma 1.

LEMMA 9. $P_\delta(\text{reject } H_1) \leq \alpha$.

PROOF. This follows from Lemmas 1 and 5 and symmetry.

For all $\gamma > 0$, $(\bar{t}_{\bar{T}} = \gamma) = \cup_{i \geq 2} (T = i, \bar{t}_i = \gamma) \in \beta(W_\delta(s), s \leq \gamma)$. Hence $\bar{t}_{\bar{T}}$ is a stopping time.

LEMMA 10. $\int_{(\bar{t}_{\bar{T}} \leq c)} (W_\delta(\bar{t}_{\bar{T}}) - A(\bar{t}_{\bar{T}})) dP \leq \delta + 1$.

PROOF. Cf. Lemma 2 of [6]. Note $\bar{t}_{i+1} - \bar{t}_i \leq 1$ for all $i \geq 2$.

LEMMA 11. $E_\delta(\bar{t}_{\bar{T}}) < \infty$.

PROOF. The main idea appears in several places, e.g., [4]; the detailed proof is omitted.

LEMMA 12. $E_\delta(\bar{t}_{\bar{T}}) \sim 2\delta^{-2} (\log_2 \delta^{-2}) P_0(\bar{t}_{\bar{T}} = \infty)$ as $\delta \downarrow 0$.

PROOF. $\bar{t}_{\bar{T}}$ is written as T to simplify notation. $E_\delta(T) \geq t P_\delta(T > t) = t \int_{(T > t)} \exp(\delta W_\delta(t) - t\delta^2/2) dP_0$ (cf. [4]). By Fatou's lemma,

$$(13) \quad \liminf_{\delta \downarrow 0} E_\delta(T) \geq t \liminf_{\delta \downarrow 0} P_\delta(T > t) \geq t P_0(T > t).$$

Letting $t \rightarrow \infty$, we have $\lim_{\delta \downarrow 0} E_\delta T = \infty$ by Lemma 8. Since $E_\delta(T) < \infty$, by Wald's lemma, $P_\delta(T < \infty) = 1$ and Lemma 10,

$$\begin{aligned} \delta E_\delta(T) &= E W_\delta(T) = E_\delta(W_\delta(T) - A(T)) + E_\delta(A(T)) \\ &\leq \delta + 1 + E_\delta(A(T)). \end{aligned}$$

Arguing exact as in Section 3 of [4] gives that for all $\epsilon > 0$

$$(14) \quad E_\delta(T) \leq 2(1 + \epsilon) \delta^{-2} (\log_2 \delta^{-1}) P_0(T = \infty) \text{ as } \delta > 0.$$

On the other hand, note that $(t_1 \leq T \leq t_2) \subset (|W_\delta(t)| \geq A(t) \text{ for some } t_1 \leq t \leq t_2)$ and $A(t) \sim (2t \log_2 t)^{1/2}$. Arguing as in [6] (cf. Lemma 9, Lemma 12, and (40) in [6]) gives that for all $\epsilon > 0$

$$(15) \quad \liminf_{\delta \rightarrow \infty} P_\delta(T \geq 2(1 - \epsilon) \delta^{-2} \log_2 \delta^{-1}) \geq P_0(T = \infty).$$

By (14) and (15), this lemma is proven.

THEOREM 1. $E_\delta(MN/(M + N)) \sim 2\delta^{-2} (\log_2 \delta^{-1}) P_0(T = \infty)$ as $\delta \downarrow 0$.

PROOF. By Lemma 5, Lemma 8, Lemma 11 and Lemma 12.

6. The asymptotic behavior of $E_\delta(M)$ as $\delta \downarrow 0$.

LEMMA 13. Let $X_n \geq 0, n \geq 1$ be a sequence of random variables. If $\lim_{n \rightarrow \infty} EX_n = a > 0$ and for all $\gamma > 0, \liminf_{n \rightarrow \infty} P(X_n > 1 - \gamma) \geq a$, then for all $\epsilon > 0, \lim_{n \rightarrow \infty} \int_{(X_n \geq 1 + \epsilon)} X_n dP = 0$.

PROOF. Given $\varepsilon > 0$, let $q > 0$ be such that $1 > \varepsilon q$. Set $I_n = \int_{(X_n \geq 1 + \varepsilon)} X_n dP$.

$$\begin{aligned} EX_n &\geq I_n + \int_{(1 + \varepsilon > X_n > 1 - \varepsilon q)} X_n dP \\ &\geq I_n + (1 - \varepsilon q)(P(X_n > 1 - \varepsilon q) - P(X_n \geq 1 + \varepsilon)) \\ &\geq I_n + (1 - \varepsilon q)P(X_n > 1 - \varepsilon q) - \frac{1 - \varepsilon q}{1 + \varepsilon} I_n. \end{aligned}$$

As $n \rightarrow \infty$, $a \geq \varepsilon(1 + q)/(1 + \varepsilon) \limsup I_n + (1 - \varepsilon q)a$. Hence $qa \geq ((1 + q)/(1 + \varepsilon)) \limsup I_n$. As $q \downarrow 0$, $\lim_{n \rightarrow \infty} I_n = 0$. To simplify notation, denote $f(\delta) = 2\delta^{-2} \log \log \delta^{-1}$ from now on.

LEMMA 14. For all $\varepsilon > 0$,

$$\int_{(MN > (1 + \varepsilon)f(\delta)(M + N))} \frac{MN}{(M + N)f(\delta)} dP_\delta \rightarrow 0 \quad \text{as } \delta \rightarrow 0.$$

PROOF. By Theorem 1, (15) and Lemma 13.

LEMMA 15. If $\lim_{\delta \downarrow 0} P_\delta(A) = 0$, then $\int_A MN / ((M + N)f(\delta)) dP_\delta \rightarrow 0$ as $\delta \rightarrow 0$.

PROOF. Note $A \subset (MN > 2(M + N)f(\delta)) \cup (A(MN \leq 2(M + N)f(\delta)))$ and Lemma 14.

LEMMA 16. For all $\varepsilon > 0 \lim_{\delta \downarrow 0} P_\delta(M > (1 + \varepsilon)f(\delta)) = 0$.

PROOF. Cf. (5.10) to (5.14) in [5] and (7) of [3].

LEMMA 17. If $\lim_{\delta \downarrow 0} P_\delta(A) = 0$, then $\int_A M / f(\delta) dP_\delta \rightarrow 0$ as $\delta \downarrow 0$, and $E_\delta(M) < \infty$.

PROOF. Set $M_1 = N_1 = 0$.

$$\begin{aligned} \int_A \frac{MN}{M + N} dP_\delta &= \sum_{\nu=2}^\infty \int_{A(T > \nu)} \{M_\nu N_\nu / \nu - M_{\nu-1} N_{\nu-1} / (\nu - 1)\} dP_\delta \\ &\geq \sum_{\nu=2}^\infty \int_{A(T > \nu, M_\nu > M_{\nu-1})} N_\nu^2 / \nu(\nu - 1) dP_\delta \quad (\text{Cf. (5.16) of [5]}) \\ &\geq \frac{1}{18} \sum_{\nu=2}^\infty P_\delta(A, T \geq \nu, M_\nu > M_{\nu-1}, N_{\nu-1} > (\nu - 1)/3) \\ &= \frac{1}{18} \{ \sum_{\nu=2}^\infty P_\delta(A, T \geq \nu, M_\nu > M_{\nu-1}) - \sum_{\nu=1}^\infty P_\delta(A, T > \nu, M_{\nu+1} > M_\nu, N_\nu \leq \nu/3) \} \\ &= \frac{1}{18} \{ \int_A M dP_\delta - \sum_\nu P_\nu \}, \text{ say.} \end{aligned}$$

Set $\nu_k = [kf(\delta)]$.

$$\begin{aligned} \sum_\nu P_\nu &= \sum_{k=0}^\infty \sum_{\nu=\nu_k+1}^{\nu_{k+1}} P_\nu \\ &\leq \nu_1 P_\delta(A) + \sum_{k \geq 1} \sum_\nu P_\delta \left(M_{\nu+1} > M_\nu, M_\nu \geq \frac{2\nu}{3}, \frac{\nu}{3} \geq N_\nu \geq \nu \lambda(\nu) \right), \\ &\hspace{15em} \text{by Lemma 2} \\ &\leq o(f(\delta)) + \sum_{k \geq 1} \sum_\nu P_\delta \left(\bar{Y}_{N_\nu} - \bar{X}_{M_\nu} \leq 0, M_\nu N_\nu / \nu \geq \frac{2}{3} \nu \lambda(\nu) \right), \quad \text{by (7)} \\ &\leq o(f(\delta)) + \sum_{k \geq 1} f(\delta) P_\delta \left(W_\delta(t_\nu) / t_\nu \leq 0, t_\nu \geq \frac{2}{3} \nu_k \lambda(\nu_k) \right) \\ &\leq o(f(\delta)) + f(\delta) \sum_{k \geq 1} 2 \left(1 - \Phi \left(\delta \left(\frac{2}{3} \nu_k \lambda(\nu_k) \right)^{\frac{1}{2}} \right) \right), \quad \text{by (7) of [3].} \end{aligned}$$

Note $\nu_k \rightarrow \infty$ as $\delta \downarrow 0$ and $\nu\lambda(\nu) = \nu/(2 \log_3 \nu)$ for large ν and $\delta(\frac{2}{3}\nu_k\lambda(\nu_k))^{\frac{1}{2}} \geq k^{\frac{1}{4}}((2 \log \delta^{-1})/(3 \log_3 f(\delta)))^{\frac{1}{2}} = k^{\frac{1}{4}}g(\delta)$, where $g(\delta) \rightarrow \infty$ as $\delta \downarrow 0$. $\sum_\nu P_\nu \leq o(f(\delta)) + f(\delta)\int_0^\infty (1 - \Phi(x^{\frac{1}{4}}g(\delta))) dx = o(f(\delta))$. Hence $\int_A M/f(\delta) dP_\delta \rightarrow 0$ as $\delta \rightarrow 0$ by Lemma 15. Let A be the whole space; it is easily seen $E_\delta(M) < \infty$.

LEMMA 18. For all $\epsilon > 0 \limsup_{\delta \downarrow 0} E_\delta(M/f(\delta)) \leq (1 + \epsilon)P_0(T = \infty)$.

PROOF. $E_\delta(M/f) = \int_{(M < \gamma f)} M/f dP_\delta + \int_{(\gamma f < M < (1+\epsilon)f)} M/f dP_\delta + \int_{((1+\epsilon)f < M)} M/f dP_\delta$, where $f = f(\delta)$ and $\gamma > 0$ arbitrarily small.

$$\begin{aligned} E_\delta(M/f) &\leq \gamma + (1 + \epsilon)P_\delta(M > \gamma f) + o(1), \text{ by Lemma 16 and Lemma 17} \\ &\leq \gamma + (1 + \epsilon)P_\delta(MN/(M + N) \geq \gamma f/4), \text{ by (9)} \\ &\leq \gamma + (1 + \epsilon)P_\delta(\bar{t}_T > t_0), \text{ where } t_0 < \frac{\gamma}{4}f \end{aligned}$$

$$\limsup_{\delta \downarrow 0} E_\delta(M/f) \leq \gamma + (1 + \epsilon)P_0(\bar{t}_T > t_0) \text{ (cf. (13) or (34) in [4]).}$$

Letting $t_0 \rightarrow \infty, \gamma \downarrow 0$, we have proved the lemma.

THEOREM 2. $E_\delta(M) \sim 2\delta^{-2} (\log \log \delta^{-1})P_0(T = \infty)$ as $\delta \downarrow 0$.

PROOF. Note $M \geq MN/(M + N)$ and Theorem 1 and Lemma 18.

7. **The optimum property of the procedure.** For any sequential procedure with $P_0(M'N'/(M' + N') = \infty) > 0$ (M', N' similarly defined as in (8)), we have

$$\limsup_{\delta \rightarrow 0} \{E_\delta M'/2\delta^{-2} \log_2 \delta^{-1}\} \geq P_0(M'N'/(M' + N') = \infty) \text{ (cf. [1]).}$$

The procedure in Section 2 is optimum in the sense that

$$\lim_{\delta \rightarrow 0} \{E_\delta M/2\delta^{-2} \log_2 \delta^{-1}\} = P_0(MN/(M + N) = \infty).$$

Acknowledgment. The author expresses his deep appreciation to Professor D. O. Siegmund for his very helpful advice in this paper.

REFERENCES

[1] FARRELL, R. H. (1964). Asymptotic behavior of expected sample size in certain one-sided tests. *Ann. Math. Statist.* **35** 36–72.
 [2] ROBBINS, H. (1970). Statistical method related to the law of the iterated logarithm. *Ann. Math. Statist.* **41** 1397–1409.
 [3] ROBBINS, H. and SIEGMUND, D. (1970). Boundary crossing probability for the Wiener process and sample sums. *Ann. Math. Statist.* **41** 1410–1429.
 [4] ROBBINS, H. and SIEGMUND, D. (1973). Statistical tests of power one and the integral representation of solutions of certain partial differential equations. *Bull. Math. Acad. Sinica.* **1** 94–120.
 [5] ROBBINS, H. and SIEGMUND, D. (1974a). Sequential tests involving two populations. *J. Amer. Statist. Assoc.* **69** 132–139.
 [6] ROBBINS, H. and SIEGMUND, D. (1974b). The expected sample size of some tests of power one. *Ann. Statist.* **2** 415–436.

26-1 MING TSU ST.
 PEITOU, TAIPEI, TAIWAN
 REPUBLIC OF CHINA