

**PROOF OF CONJECTURES ABOUT THE EXPECTED VALUES
OF THE ELEMENTARY SYMMETRIC FUNCTIONS
OF A NONCENTRAL WISHART MATRIX**

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This paper provides proof of conjectures about the expected values of the j th elementary symmetric function of the roots (i.e. $E(\text{tr}_j A)$) of a non-central Wishart matrix. Utilizing this result we have obtained computationally convenient formula for the expected value of the j th elementary symmetric function of the roots of central positive definite quadratic form in matrix argument.

1. Introduction. Let $X: p \times n$ be a matrix whose column vectors are independently and identically distributed in multivariate normal distributions having $E(X) = M$ and variance covariance matrix Σ . The matrix $XX' = A$ is distributed as a noncentral Wishart distribution $W_p(n, \Sigma, \Omega; A)$ with n degrees of freedom and noncentrality parameter $\Omega = \Sigma^{-1}MM'$.

In the next section we provide the mathematical proof of conjectures made by de Waal [3] about (i) $E(\text{tr}_p A)$ and (ii) $E(\text{tr}_j A)$, $j = 1, 2, \dots, p - 1$. In the last section, we have extended the above results to the case of central positive definite quadratic form in matrix argument.

2. Central and noncentral Wishart case.

LEMMA 1. For a central Wishart matrix $W_p(n, \Sigma, 0; S)$

(2.1) (a) $E(\text{tr}_p S) = E(|S|) = n^{(p)}|\Sigma|$,

(2.2) (b) $E(\text{tr}_j S) = n^{(j)} \text{tr}_j \Sigma$, $j = 1, 2, \dots, p - 1$,

where $n^{(p)} = n(n - 1) \dots (n - p + 1)$.

PROOF. (a) This is a well-known result and proof is given in many standard text books.

(b) By definition, $\text{tr}_j S$ is the sum of all principal minors of order j of matrix S . Thus,

$$\begin{aligned} E(\text{tr}_j S) &= E(\text{sum of all principal minors of order } j \text{ of matrix } S) \\ &= n^{(j)}(\text{sum of all principal minors of order } j \text{ of matrix } \Sigma) \\ &= n^{(j)} \text{tr}_j \Sigma. \end{aligned}$$

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This establishes Lemma 1. We may note here that de Waal [3] has established the lemma by providing zonal polynomial argument.

LEMMA 2. For a noncentral Wishart matrix $W_p(n, \Sigma, \Omega; A)$ with $\Omega = \Sigma^{-1}W$, $W = MM'$,

$$(2.3) \quad (a) \quad E(\text{tr}_p A) = E(|A|) = n^{(j)}|\Sigma| + \sum_{i=1}^p (n - i)^{(p-i)}|\Sigma| \text{tr}_i \Omega$$

$$(2.4) \quad (b) \quad E(\text{tr}_j A) = n^{(j)} \text{tr}_j \Sigma + \sum_{k=1}^j (n - k)^{(j-k)} \sum_{i(j)} |\Sigma(i(j))| \text{tr}_k \{[\Sigma(i(j))]^{-1}W(i(j))\}, \quad j = 1, 2, \dots, p - 1,$$

where $\Sigma(i(j))$ and $W(i(j))$ are submatrices obtained by considering i_1, i_2, \dots, i_j rows and i_1, i_2, \dots, i_j columns of matrices Σ and W respectively and $\sum_{i(j)} = \sum_{i_1=1; i_1 > i_2 > \dots > i_j} \sum_{i_j=1}^p$.

PROOF. Part (a): The noncentral moments of the generalized variance $|A|$ have been given by Anderson [1] for the case of Ω of rank ≤ 2 , and in general by Herz [4] who quoted by Constantine [2] as:

$$(2.5) \quad E(\text{tr}_p A) = |2\Sigma|\Gamma_p(\frac{1}{2}n + 1)\{\Gamma_p(\frac{1}{2}n)\}^{-1}F_1(-1; \frac{1}{2}n; -\frac{1}{2}\Omega),$$

where ${}_pF_q(a, b; S) = \sum_{\kappa=0}^{\infty} \sum_{\kappa} ((a)_{\kappa}/(b)_{\kappa})(C_{\kappa}(S)/k!)$, \sum_{κ} is taken over all the partitions $\kappa = (k_1, k_2, \dots, k_p)$, $k_1 \geq k_2 \geq \dots \geq k_p \geq 0$ so that $k_1 + k_2 + \dots + k_p = k$, $(a)_{\kappa} = \prod_{i=1}^p \{a - (i - 1)/2\}_{k_i}$, $(x)_{\kappa} = x(x + 1) \dots (x + k - 1)$, $C_{\kappa}(s)$ is a polynomial of degree k in the latent roots of the matrix S and $\Gamma_p(x) = \pi^{p(p-1)/4} \prod_{j=1}^p \Gamma\{x - (j - 1)/2\}$. The difficulty, as quoted by de Waal ([3], page 346), is to simplify expression (2.5). From definitions

$$(2.6) \quad |2\Sigma|\Gamma_p(\frac{1}{2}n + 1)\{\Gamma_p(\frac{1}{2}n)\}^{-1} = n^{(p)}|\Sigma|, \quad \text{and}$$

$$(2.7) \quad {}_1F_1(-1; \frac{1}{2}n; -\frac{1}{2}\Omega) = \sum_{\kappa=0}^{\infty} \sum_{\kappa} \frac{(-1)_{\kappa}}{(\frac{1}{2}n)_{\kappa}} \frac{C_{\kappa}(-\Omega/2)}{k!}.$$

Now $(-1)_{\kappa} = 0$ if $k_1 > 1$. Hence $(-1)_{\kappa} = 0$ if $k > p$. Further, $(-1)_{\kappa} = \prod_{i=1}^k \{-1 - (i - 1)/2\}$, if $k_1 = k_2 = \dots = k_k = 1$ and other $k_j = 0$ for $k \leq p$. Hence $(-1)_{\kappa} = (-1)^k(k + 1)! (2^{-k})$, for $\kappa = \{1^k, 0\}$. Therefore,

$$(2.8) \quad {}_1F_1(-1; \frac{1}{2}n; -\frac{1}{2}\Omega) = \sum_{k=0}^p (k + 1)\{n^{(k)}2^k\}^{-1}C_{\{1^k, 0\}}(\Omega).$$

For every partition $\kappa = (k_1, k_2, \dots, k_p)$, $k_1 \geq \dots \geq k_p \geq 0$, the zonal polynomials $C_{\kappa}(\Omega)$ can be written as (see [5], equation 18, page 478)

$$(2.9) \quad C_{\kappa}(\Omega) = \{\chi_{[2\kappa]}(1)2^k k!\}Z_{\kappa}(\Omega)/(2k).$$

Equation (4.6) is to be evaluated at $\kappa = \{1^k, 0\}$. Using equation 19, page 479 of James [5], we have

$$(2.10) \quad \chi_{[2\{1^k, 0\}]}(1) = (2k)! \{k! (k + 1)!\}^{-1},$$

and noting (120) and (121) of James [5] on page 492, we have

$$(2.11) \quad Z_{\{1^k, 0\}}(\Omega) = k! \text{tr}_k \Omega.$$

Substituting the values of equations (2.10) and (2.11) in (2.9), we obtain

$$(2.12) \quad C_{(1^k, 0)}(\Omega) = 2^k(\text{tr}_k \Omega)(k + 1)^{-1}.$$

Thus,

$$(2.13) \quad {}_1F_1(-1; \frac{1}{2}n; -\frac{1}{2}\Omega) = \sum_{k=0}^p \text{tr}_k \Omega / n^{(k)}.$$

Combining expressions (2.12) and (2.13) establishes a conjectured result.

Part (b): By definition of trace, it is easy to write

$$(2.14) \quad \text{tr}_j A = \sum_{i(j)} |A(i(j))|$$

where $A(i(j))$ is a submatrix obtained by considering i_1, i_2, \dots, i_j rows and i_1, i_2, \dots, i_j columns of matrix A . We may further note that the submatrix $A(i(j))$ is distributed as noncentral Wishart $W_j(n, \Sigma(i(j)), [\Sigma(i(j))]^{-1}W(i(j)); A(i(j)))$. Applying the result of Lemma 2 part (a) to the submatrix $A(i(j))$, we get

$$(2.15) \quad E(\text{tr}_j A) = \sum_{i(j)} E\{|A(i(j))|\} \cdot \sum_{i(j)} \{n^{(j)} + \sum_{k=1}^j (n - k)^{(j-k)} \text{tr}_k [\Sigma(i(j))]^{-1}W(i(j))\} \Sigma(i(j))\}$$

which establishes part (b).

COROLLARY. When $\Sigma = \sigma^2 I_p$

$$(2.16) \quad E(\text{tr}_j A) = \sigma^{2j} \sum_{k=0}^j (n - k)^{(j-k)} \binom{p-k}{j-k} \text{tr}_k \Omega, \quad j = 1, 2, \dots, p.$$

PROOF. Since $\Sigma = \sigma^2 I_p$, we have

$$\sum_{i(j)} \text{tr}_k [\Sigma(i(j))]^{-1}W(i(j)) = \binom{p-k}{j-k} \text{tr}_k \Omega,$$

which establishes the corollary.

3. Quadratic forms in matrix argument. Let $Q = XLX'$, where $X: p \times n$ is a matrix whose column vectors are independently and identically distributed in multivariate normal distribution having $E(X) = 0$ and variance covariance matrix Σ , and $L: n \times n$ is a positive definite matrix. Previously Khatri [6] gave $E(|Q|)$ involving hypergeometric series representation in matrix argument. This expression is very difficult to evaluate numerically. Lemma 3 gives a simpler and computationally convenient expression.

LEMMA 3. For a central positive definite quadratic form

$$(3.1) \quad (a) \quad E(|Q|) = n^{(p)} |\Sigma| q^p \sum_{i=0}^p (-1)^i p^{(i)} \text{tr}_i T / n^{(i)} i!,$$

$$(3.2) \quad (b) \quad E(\text{tr}_j Q) = n^{(j)} q^j (\text{tr}_j \Sigma) \sum_{i=0}^j (-1)^i j^{(i)} \text{tr}_i T / n^{(i)} i!,$$

where $q > 0$ and $T = I_n - q^{-1}L$.

PROOF. (a) Equation (46) of [6] can be written as,

$$(3.3) \quad \begin{aligned} E(|Q|) &= n^{(p)} |\Sigma| q^{p(\frac{1}{2}n+1)} |L|^{-p/2} {}_1F_0^{(n)}(\frac{1}{2}n + 1; I_n - qL^{-1}, I_p) \\ &= n^{(p)} |\Sigma| |qL^{-1}|^{\frac{1}{2}p} q^p {}_2F_1(\frac{1}{2}n + 1, p; \frac{1}{2}n; I_n - qL^{-1}) \\ &= n^{(p)} |\Sigma| q^p {}_2F_1(-1, \frac{1}{2}p; \frac{1}{2}n; T), \end{aligned}$$

since ${}_2F_1(a; b; c; S) = |I - S|^{-b} {}_2F_1(c - a, b; c; -S(I - S)^{-1})$. Equation (3.3) can then be simplified in the manner described in Lemma 2, giving a final result as mentioned in (3.1). We may note that when $L = I$, (3.1) reduces to (2.1).

(b) By applying the argument of Lemma 1, we can easily establish Lemma 3(b). This result generalizes Lemma 1.

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