

ON THE INCREMENTS OF MULTIDIMENSIONAL RANDOM FIELDS

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For a nondifferentiable random field $\{X_t: t \in \mathbb{R}^N\}$ with values in \mathbb{R}^d , it is often easy to check that with probability 1 $\liminf_{s \rightarrow t} \|X_s - X_t\|/\sigma(s, t) = 0$ and $\limsup_{s \rightarrow t} \|X_s - X_t\|/\sigma(s, t) = \infty$ for a.e. t , where $\sigma^2(s, t) = E\|X_s - X_t\|^2$. In this note we discuss the "proportion" of s 's near t for which $\|X_s - X_t\|/\sigma(s, t)$ is small or large.

Let $\{X_t, t \in \mathbb{R}^N\}$ be an M -valued stochastic process, \mathbb{R}^N being N -dimensional Euclidean space and $(M, \|\cdot\|)$ a separable Banach space; (Ω, \mathcal{F}, P) denotes the probability space. We assume that $(t, \omega) \rightarrow X_t(\omega)$ is measurable $\mathcal{B} \otimes \mathcal{F} \rightarrow \mathcal{M}$ (\mathcal{B}, \mathcal{M} the Borel σ -fields in \mathbb{R}^N, M respectively), that $E\|X_t\|^2 < \infty$ for all t , and that $\sigma^2(s, t) = E\|X_s - X_t\|^2$ is jointly continuous.

In this paper we will consider the approximate local behavior of the normalized increments $\|X_s - X_t\|/\sigma(s, t)$. When $M = \mathbb{R}^d, d \leq N$, and X_t is nondifferentiable, it is usually easy to check that with probability one,

$$(1) \quad \liminf_{s \rightarrow t} \frac{\|X_s - X_t\|}{\sigma(s, t)} = 0 \quad \text{and} \quad \limsup_{s \rightarrow t} \frac{\|X_s - X_t\|}{\sigma(s, t)} = \infty$$

for every t or λ -a.e. t , where λ is Lebesgue measure on \mathcal{B} . However, behavior such as (1) provides no information on the "proportion" of s 's near t for which $\|X_s - X_t\|/\sigma(s, t)$ is small or large. Indeed, (1) can be arranged by altering $X_t(\omega)$ on a countable set.

Let μ be a positive, σ -finite Borel measure on \mathcal{B} , positive on open sets, and let $B(t, \varepsilon)$ be the open ball centered at $t \in \mathbb{R}^N$ of radius ε . Let $f(s)$ be a real, \mathcal{B} -measurable function. The *approximate lower limit* $\text{ap} \liminf_{s \rightarrow t} f(s)$ (relative to μ and the Euclidean topology) is the supremum of those v such that $\{s: f(s) < v\}$ has density 0 at t —i.e.,

$$\lim_{\varepsilon \rightarrow 0} \frac{\mu\{B(t, \varepsilon) \cap \{s: f(s) < v\}\}}{\mu\{B(t, \varepsilon)\}} = 0.$$

The *approximate upper limit* is defined analogously. (See [2] and [7] for the role of approximate limits, derivatives, etc., in classical and modern analysis.)

Received May 17, 1976; revised June 1, 1977.

¹ This work partially supported by NSF grant MPS 72-04813-A03.

² This work partially supported by NSF grant MPS 75-07605.

AMS 1970 subject classifications. Primary 60G10, 60G15, 60G17.

Key words and phrases. Random field, approximate limit, Gaussian process, stationary increments.



For a large class of processes we will show that for each $t \in \mathbb{R}^M$, with probability one:

$$(2) \quad \text{ap lim inf}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\sigma(s, t)} = 0 \quad \text{and} \quad \text{ap lim sup}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\sigma(s, t)} = \infty .$$

Obviously, (2) implies (1).

Before stating the main theorems, we give two examples; the first shows how these theorems complement certain results in [3], [4] and [6].

EXAMPLE 1. Let $\{X_t = (X_1(t), \dots, X_d(t)), t \in \mathbb{R}^M\}$ be the d -dimensional Gaussian random field with i.i.d. components, zero means, and covariance $EX_j(t)X_j(s) = \|t\|^\alpha + \|s\|^\alpha - \|t - s\|^\alpha$, $0 < \alpha < 2$. It follows from Theorem 1 of [6] that, if $N - \alpha d/2 < 0$,

$$(3) \quad \lim_{s \rightarrow t} \frac{\|X_s - X_t\|}{\|s - t\|^r} = \infty \quad \text{for } \lambda\text{-a.e. } t, \text{ a.s.}$$

for any $r > \alpha/2$. On the other hand, if $N - \alpha d/2 = \varepsilon > 0$, the results of [3] yield

$$(4) \quad \text{ap lim}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\|s - t\|^r} = \infty \quad \text{for } \lambda\text{-a.e. } t, \text{ a.s.}$$

for $r = \alpha/2 + \varepsilon/d$. Now for standard one-dimensional Brownian motion, i.e., $\alpha, N, d = 1$, it is not hard to show, using the zero-one law, that

$$(5) \quad \text{ap lim inf}_{s \rightarrow t} \frac{|X_s - X_t|}{|s - t|^{1/2}} = \infty \quad \text{for } \lambda\text{-a.e. } t, \text{ a.s.}$$

Finally, by Theorem 3 of this paper we find that for any N and d (in particular when $N = \alpha d/2$),

$$(6) \quad \begin{aligned} \text{ap lim inf}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\|s - t\|^{\alpha/2}} &= 0, \\ \text{ap lim inf}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\|s - t\|^{\alpha/2}} &= \infty \quad \lambda\text{-a.e.}, \text{ a.s.} \end{aligned}$$

It is clear that in some sense (6) represents the boundary case between (3) and (4).

EXAMPLE 2. Let $W(s, t)$, $s, t \in [0, 1]$, be the Yeh-Wiener process, i.e., the mean 0, real Gaussian field on $[0, 1]^2$ with $W(0, t) = 0$ a.s., $W(s, 0) = 0$ a.s., and

$$EW(s_1, t_1)W(s_2, t_2) = \min(s_1, s_2) \cdot \min(t_1, t_2).$$

One can also view W as a one-parameter process with values in $M = C[0, 1]$: $W_s: \Omega \rightarrow C[0, 1]$, $W_s(\omega)(t) = W(s, t, \omega)$.

From Theorem 1, then,

$$\begin{aligned} \text{ap lim sup}_{(s,t) \rightarrow (a,b)} \frac{|W(s, t) - W(a, b)|}{\sigma_1((s, t), (a, b))} &= \infty, \\ \text{ap lim inf}_{(s,t) \rightarrow (a,b)} \frac{|W(s, t) - W(a, b)|}{\sigma_1((s, t), (a, b))} &= 0 \quad \text{a.s.} \end{aligned}$$

for each $0 \leq a, b \leq 1$, where $\sigma_1^2((s, t), (a, b)) = st + ab - 2(s \wedge a)(t \wedge b)$. Also,

$$\begin{aligned} \text{ap lim sup}_{s \rightarrow a} \frac{\sup_{0 \leq t \leq 1} |W(s, t) - W(a, t)|}{|s - a|^{\frac{1}{2}}} &= \infty, \\ \text{ap lim inf}_{s \rightarrow a} \frac{\sup_{0 \leq t \leq 1} |W(s, t) - W(a, b)|}{|s - a|^{\frac{1}{2}}} &= 0 \quad \text{a.s.} \end{aligned}$$

Returning to the general case, our basic assumptions about X_t will be

(A1) There exist measurable functions $\phi_j: \mathbb{R}^N \rightarrow M$ and independent random variables $\xi_j, j = 1, 2, \dots$ such that for each t ,

- (a) $X_t(\omega) = \lim_{n \rightarrow \infty} \sum_{j=1}^n \xi_j(\omega) \phi_j(t)$ a.s. and
- (b) $\lim_{s \rightarrow t} \|\phi_j(s) - \phi_j(t)\|/\sigma(s, t) = 0$;

(A2) For each $t \in \mathbb{R}^N$ and $Q > 0$ there are numbers $\varepsilon, \delta > 0$ such that

$$\delta \leq P(\|X_s - X_t\| \leq Q\sigma(s, t)) \leq 1 - \delta$$

for all $s \in B(t, \varepsilon)$.

We note that the set of pairs (t, ω) for which $\sum_1^n \xi_j(\omega) \phi_j(t)$ converges to $X_t(\omega)$ is jointly measurable; hence, for almost every ω we have $X_t(\omega) = \sum_1^\infty \xi_j(\omega) \phi_j(t)$ for each $t \in A_\omega, \mu(A_\omega^c) = 0$.

THEOREM 1. *Suppose $\{X_t\}$ satisfies (A1) and (A2). Then for each $t \in \mathbb{R}^N$, (2) holds with probability one.*

NOTE. For t, ω , and $Q > 0$ fixed, the expression

$$(7) \quad g(\varepsilon) \equiv \frac{\mu\{s \in B(t, \varepsilon) : \|X_s(\omega) - X_t(\omega)\| \leq Q\sigma(s, t)\}}{\mu\{B(t, \varepsilon)\}}$$

need not have a limit as $\varepsilon \rightarrow 0$. (See Remark (e).) To prove the theorem, we must show that for each t fixed, $\limsup_{\varepsilon \downarrow 0} g(\varepsilon) > 0$ and $\liminf_{\varepsilon \downarrow 0} g(\varepsilon) < 1$ for all $Q > 0$, with probability one.

PROOF. Fix $t \in \mathbb{R}^N$ and define (for a.e. ω):

$$X^m(t, \omega) = \sum_{j=m+1}^\infty \xi_j(\omega) \phi_j(t), \quad X_m(t, \omega) = \sum_{j=1}^m \xi_j(\omega) \phi_j(t), \quad m \geq 0,$$

and

$$\begin{aligned} K^m(Q, n, \omega) &= [\mu\{B(t, 1/n)\}]^{-1} \mu\{s \in B(t, 1/n) \cap A_\omega : \\ &\quad \|X^m(s, \omega) - X^m(t, \omega)\| \leq Q\sigma(s, t)\}, \quad n \geq 1, Q > 0. \end{aligned}$$

(Since $\mu(A_\omega^c) = 0$ a.s., we will ignore A_ω^c in what follows.) Further, define

$$\bar{G}^m(Q, \omega) = \limsup_{n \rightarrow \infty} K^m(Q, n, \omega), \quad \underline{G}^m(Q, \omega) = \liminf_{n \rightarrow \infty} K^m(Q, n, \omega).$$

We will show that, with probability one, $\bar{G}^0(Q, \omega) > 0$ for all Q and $\underline{G}^0(Q, \omega) < 1$ for all Q .

Now fix $m \geq 0, Q > \eta > 0$, and $\omega \in \Omega$. By (A1)(b), there exists an integer n_0 such that $\|s - t\| \leq 1/n_0$ implies $\|X_m(s, \omega) - X_m(t, \omega)\| \leq \eta\sigma(s, t)$. Hence,

for $n \geq n_0$ and a.e. ω

$$\begin{aligned} & \mu \left\{ s \in B \left(t, \frac{1}{n} \right) : \|X_s - X_t\| \leq (Q - \eta)\sigma(s, t) \right\} \\ & \leq \mu \left\{ s \in B \left(t, \frac{1}{n} \right) : \|X^m(s) - X^m(t)\| - \|X_m(s) - X_m(t)\| \leq (Q - \eta)\sigma(s, t) \right\} \\ & \leq \mu \left\{ s \in B \left(t, \frac{1}{n} \right) : \|X^m(s) - X^m(t)\| \leq Q\sigma(s, t) \right\} \\ & \leq \mu \left\{ s \in B \left(t, \frac{1}{n} \right) : \|X^m(s) - X^m(t)\| + \|X_m(s) - X_m(t)\| \leq (Q + \eta)\sigma(s, t) \right\} \\ & \leq \mu \left\{ s \in B \left(t, \frac{1}{n} \right) : \|X_s - X_t\| \leq (Q + \eta)\sigma(s, t) \right\}. \end{aligned}$$

Consequently, $\bar{G}^0(Q - \eta, \omega) \leq \bar{G}^m(Q, \omega) \leq \bar{G}^0(Q + \eta, \omega)$ and $\underline{G}^0(Q - \eta, \omega) \leq \underline{G}^m(Q, \omega) \leq \underline{G}^0(Q + \eta, \omega)$. Since $\bar{G}^0(\cdot, \omega)$ and $\underline{G}^0(\cdot, \omega)$ are nondecreasing in Q , each has at most countably many discontinuities. It follows (upon letting $\eta \downarrow 0$) that for a.e. ω , $\bar{G}^0(Q, \omega) = \bar{G}^m(Q, \omega)$ and $\underline{G}^0(Q, \omega) = \underline{G}^m(Q, \omega)$ for all m , for all but at most countably many Q 's. Now \bar{G}^m and \underline{G}^m being jointly measurable, we then find, for λ -a.e. $Q > 0$, that $\bar{G}^0(Q, \omega) = \lim_{m \rightarrow \infty} \bar{G}^m(Q, \omega)$, $\underline{G}^0(Q, \omega) = \lim_{m \rightarrow \infty} \underline{G}^m(Q, \omega)$ a.s.

Therefore, for such Q , \bar{G}^0 and \underline{G}^0 are measurable with respect to the tail σ -field for the sequence $\{\xi_j(\omega)\}$.

Finally,

$$\begin{aligned} E \underline{G}^0(Q, \omega) &= E \liminf_{n \rightarrow \infty} K^0(Q, n, \omega) \leq \liminf_{n \rightarrow \infty} EK^0(Q, n, \omega) \\ &= \liminf_{n \rightarrow \infty} [\mu\{B(t, 1/n)\}]^{-1} \int_{B(t, 1/n)} P(\|X_s - X_t\| \\ &\leq Q\sigma(s, t)) \mu(ds) \leq 1 - \delta_Q \end{aligned}$$

for some $0 < \delta_Q < 1$, where the last inequality uses (A2). As a result, for λ -a.e. $Q > 0$, $\underline{G}^0(Q, \omega) < 1$ a.s., from which it follows—by Fubini's theorem and the monotonicity of $\underline{G}^0(\cdot, \omega)$ —that $\underline{G}^0(Q, \omega) < 1$ for all Q a.s. As for \bar{G}^0 , $E \bar{G}^0 = E \limsup_{n \rightarrow \infty} K^0(Q, n, \omega) \geq \limsup_{n \rightarrow \infty} EK^0(Q, n, \omega) \geq \delta_Q > 0$, which leads to $\bar{G}^0(Q, \omega) > 0$ for all Q a.s. This completes the proof.

We intend to apply Theorem 1 to multidimensional Gaussian random fields $X(t) = \{X_i(t)\}_{i=1}^d$, $t \in \mathbb{R}^N$ (i.e., $M = \mathbb{R}^d$). Roughly speaking, we will show that (A1) and (A2) hold whenever $\sigma(s, t)/\|s - t\| \rightarrow \infty$ and the components $X_i(t, \omega)$ are independent and each covariance has a suitable "spectral representation." The following lemma extends Theorem 1 of Klein [5]. We omit the proof, it being an easy modification of Klein's.

LEMMA 2. *Let $\{X_t, t \in \mathbb{R}^N\}$ be a real-valued, second-order process with mean 0 and let $R(t, s) = EX_t X_s$. Suppose that there exists a locally compact Hausdorff space L and a σ -finite, regular Borel measure π on L , finite on compacts, such that: (a) there exists a family of real functions on $L\{g(t, \cdot), t \in \mathbb{R}^N\}$ such that the finite linear*

combinations of the $g(t, \cdot)$ are dense in $L^2(\pi)$ and (b) $(\partial/\partial s_k)g(s, \lambda)$ exists and is jointly continuous for all $(s, \lambda) \in \mathbb{R}^N \times L$ and $1 \leq k \leq n$, and $R(t, s) = \int_L g(t, \lambda)g(s, \lambda)\pi(d\lambda)$.

Then there exist real functions $\phi_j(t)$, $t \in \mathbb{R}^N$, and orthogonal random variables $\xi_j(\omega)$, $j = 1, 2, \dots$, such that:

- (i) $\phi_j(t) = E\xi_j X_t$ for all t ;
- (ii) $\limsup_{s \rightarrow t} |\phi_j(s) - \phi_j(t)|/|s - t| < \infty$ for all t ;
- (iii) $\{\phi_j\}$ forms a complete orthonormal system in the reproducing kernel Hilbert space $H(R)$; and
- (iv) $E|X_t - \sum_{j=1}^m \xi_j \phi_j(t)|^2 \rightarrow 0$ as $m \rightarrow \infty$ for all t .

If, in addition, $\{X_t\}$ is Gaussian, then ξ_1, ξ_2, \dots may be taken independent and standard normal.

As we shall see in the proof of Theorem 3, the conclusions of Lemma 2 hold for any real, mean 0, Gaussian random field $\{X_t, t \in \mathbb{R}^N\}$ with stationary increments.

THEOREM 3. Let $\{X_t, t \in \mathbb{R}^N\} = \{(X_1(t), X_2(t), \dots, X_d(t)), t \in \mathbb{R}^N\}$ be a mean 0, d -dimensional, Gaussian random field with independent components. Assume further that $\sigma(s, t)/|s - t| \rightarrow \infty$ as $s \rightarrow t$ for all t , and that each component $\{X_j(t), t \in \mathbb{R}^N\}$ has stationary increments. Then for each $t \in \mathbb{R}^N$, (2) holds with probability 1.

PROOF. First, we show that (i)–(iv) of Lemma 2 hold for $\{X_j(t) - X_j(0), t \in \mathbb{R}^N\}$ for each $j = 1, 2, \dots, d$. For simplicity, let $j = 1$.

It is well known (see e.g., [1], Theorem 3.1) that there exists a unique measure π on $\mathbb{R}^N \setminus \{0\}$

$$\begin{aligned} E(X_1(t) - X_1(0))(X_1(s) - X_1(0)) \\ = \int_{\mathbb{R}^N \setminus \{0\}} (e^{it \cdot \lambda} - 1)(e^{is \cdot \lambda} - 1)\pi(d\lambda) \quad \text{for all } s, t \end{aligned}$$

where π is a Lévy measure, i.e., with $D = \{\lambda \in \mathbb{R}^N : \|\lambda\| \geq 1\}$ and $C = \{\lambda \in \mathbb{R}^N : 0 < \|\lambda\| \leq 1\}$,

$$\pi(D) < \infty \quad \text{and} \quad \int_C \|\lambda\|^2 \pi(d\lambda) < \infty.$$

To apply Lemma 2, we choose $L = \mathbb{R}^N \setminus \{0\}$; now π is finite on compacts since these are bounded away from ∞ and 0. We need to show that the functions of the form $\sum_{j=1}^m c_j(e^{it \cdot j \cdot \lambda} - 1)$ are dense in $L^2(\pi)$. To this end, let $f(\lambda) \in L^2(\pi)$ be orthogonal to each of the functions $\lambda \rightarrow (e^{it \cdot \lambda} - 1)$, $t \in \mathbb{R}^N$, or what is the same,

$$(8) \quad \int_L f^+(\lambda)(e^{it \cdot \lambda} - 1)\pi(d\lambda) = \int_L f^-(\lambda)(e^{it \cdot \lambda} - 1)\pi(d\lambda) \quad \text{for all } t$$

where f^+, f^- are the positive and negative parts of f . Since f^+ and f^- are in $L^2(\pi)$:

$$\begin{aligned} \int_C \|\lambda\|^2 f^\pm(\lambda)\pi(d\lambda) &\leq \int_C \|\lambda\| f^\pm(\lambda)\pi(d\lambda) \\ &\leq (\int_C \|\lambda\|^2 \pi(d\lambda))^{1/2} (\int_C (f^\pm(\lambda))^2 \pi(d\lambda))^{1/2} < \infty \end{aligned}$$

and

$$\int_D f^\pm(\lambda)\pi(d\lambda) \leq (\int_D (f^\pm(\lambda))^2 \pi(d\lambda))^{1/2} (\pi(D))^{1/2} < \infty.$$

Hence $f^+(\lambda)\pi(d\lambda)$ and $f^-(\lambda)\pi(d\lambda)$ are Lévy measures. By (8) and the uniqueness of Lévy measures, $f^+ = f^-$ π -a.e.

By Lemma 2, for each $j = 1, 2, \dots, d$ there are independent standard normal rv's $\{\xi_{jk}\}_{k=1}^\infty$ and real functions $\{\tilde{\phi}_{jk}\}_{k=1}^\infty$ on \mathbb{R}^N such that, for each t , $\sum_{k=1}^n \xi_{jk}(\omega)\tilde{\phi}_{jk}(t)$ converges to $X_j(t, \omega) - X_j(0, \omega)$ in mean square, and hence pointwise. Since the components $\{X_j\}_{j=1}^d$ are independent, we can and do assume *all* the ξ_{jk} 's independent. To see that (A1) holds, define $\phi_{jk}: \mathbb{R}^N \rightarrow \mathbb{R}^d$ by $\phi_{jk}(t) = (0, \dots, 0, \tilde{\phi}_{jk}(t), 0, \dots, 0)$ where $\tilde{\phi}_{jk}(t)$ occupies the j th coordinate. Then for each t ,

$$X_t(\omega) - X_0(\omega) = \sum_{j=1}^d \sum_{k=1}^\infty \xi_{jk}(\omega)\phi_{jk}(t) \quad \text{a.s.}$$

so that (A1) holds for $\{X_t - X_0, t \in \mathbb{R}^N\}$, which is clearly enough for Theorem 1. (A1)(b) follows from conclusion (ii) of Lemma 2 together with the assumption that $\sigma(s, t)/\|s - t\| \rightarrow \infty$ as $s \rightarrow t$.

By Theorem 1, to finish the proof, we need only check that (A2) holds. Let $\sigma_j^2(s, t) = E(X_j(t) - X_j(s))^2$, $1 \leq j \leq d$, and let $\psi_j(s) = \sigma_j^2(s, t)/\sigma^2(s, t)$ for $\|s - t\|$ small; here t is fixed. Then

$$\frac{\|X_s - X_t\|^2}{\sigma^2(s, t)} = \sum_{j=1}^d Y_j^2(\omega)\psi_j(s)$$

where Y_1, Y_2, \dots, Y_d are independent and standard normal, and of course $\sum \psi_j(s) = 1$. Hence, for any $Q > 0$,

$$P\{\|X_s - X_t\| \leq Q\sigma(s, t)\} = P\{\sum_{j=1}^d Y_j^2\psi_j(s) \leq Q^2\}.$$

But,

$$0 < P\{\max_{1 \leq j \leq d} |Y_j| \leq Q\} \leq P\{\sum_{j=1}^d Y_j^2\psi_j(s) \leq Q^2\} \\ \leq P\{\min_{1 \leq j \leq d} |Y_j| \leq Q\} < 1,$$

which gives (A2).

REMARKS. (a) The zero-one law for subgroups implies that, for Gaussian processes, the event $\text{ap lim sup}_{s \rightarrow t} \|X_s - X_t\|/\sigma(s, t) < \infty$ has probability 0 or 1.

(b) Theorem 2 remains valid if the components $\{X_j(t), t \in \mathbb{R}^N\}$ have "mth order stationary increments" (see [7]).

(c) If, in Theorem 3, the components are identically distributed, but not necessarily independent, one still obtains $\text{ap lim sup}_{s \rightarrow t} \|X_s - X_t\|/\sigma(s, t) = \infty$ a.s. for all t , as follows:

$$\mu\{s \in B(t, \varepsilon) : \sum_{j=1}^d |X_j(s) - X_j(t)|^2 \leq Q^2 d \sigma_1^2(s, t)\} \\ \leq \mu\{s \in B(t, \varepsilon) : |X_1(s) - X_1(t)| \leq Q(d)^{1/2} \sigma_1(s, t)\};$$

hence, since $\sigma_1^2(s, t)/\|s - t\| \rightarrow \infty$,

$$\text{ap lim sup}_{s \rightarrow t} \frac{\|X_s - X_t\|}{\sigma(s, t)} \geq \text{ap lim sup}_{s \rightarrow t} \frac{|X_1(s) - X_1(t)|}{\sigma_1(s, t)} = \infty \quad \text{a.s.}$$

(d) Let $X_t = (X_1(t), \dots, X_d(t))$, $t \in \mathbb{R}^N$, where X_1, \dots, X_d are independent, symmetric stable processes of index $\alpha_1, \dots, \alpha_d$, each with independent increments.

Define $\phi_j(s, t) \geq 0$ by $E(e^{iu[X(s)-X(t)]}) = \exp\{-\phi_j^{\alpha_j}(s, t)|u|^{\alpha_j}\}$ and let $\phi(s, t) = \sum_{j=1}^d \phi_j(s, t)$. Then $\text{ap lim inf}_{s \rightarrow t} \|X_s - X_t\|/\phi(s, t) = 0$ and $\text{ap lim sup}_{s \rightarrow t} \|X_s - X_t\|/\phi(s, t) = \infty$. To see this note that the zero-one law still applies by the independence of the increments and (A2) holds by the choice of ϕ .

(e) We will show that for one-dimensional Brownian motion $\{W_t : 0 \leq t \leq 1\}$ the limit in (7) does not exist for $t = 0$ (and hence for any $t \geq 0$). $A =_D B$ will mean that A and B have the same distribution, and m will denote Lebesgue measure.

Fix $Q > 0$ and let

$$\begin{aligned} Z_\varepsilon^Q &= \frac{1}{\varepsilon} m\{0 \leq s \leq \varepsilon : |W(s)| \leq Qs^\frac{1}{2}\} \\ &= \frac{1}{\varepsilon} \int_0^\varepsilon 1_{[0, Q]}(|W(s)|/s^\frac{1}{2}) ds = \int_0^1 1_{[0, Q]}(|W(\varepsilon s)|/(\varepsilon s)^\frac{1}{2}) ds \\ &=_D \int_0^1 1_{[0, Q]}(|W(s)|/s^\frac{1}{2}) ds, \quad \text{since } \{W(\varepsilon s)/\varepsilon^\frac{1}{2} : 0 \leq s \leq 1\} \\ &=_D \{W(s) : 0 \leq s \leq 1\}. \end{aligned}$$

Hence the rv's $Z_\varepsilon^Q, 0 < \varepsilon \leq 1$, are identically distributed. If Z_ε^Q converges a.s. (as $\varepsilon \downarrow 0$), say to Z^Q , we have $Z^Q =_D Z_\varepsilon^Q = m\{0 \leq s \leq 1 : |W(s)| \leq Qs^\frac{1}{2}\}$. Also, by the independence of the increments of $\{W_s\}$, the zero-one law implies that Z^Q is constant a.s., and hence a.s.:

$$\begin{aligned} Z^Q &= EZ^Q = \int_0^1 P(|W(s)| \leq Qs^\frac{1}{2}) ds \\ &= 2\Phi(Q) - 1, \end{aligned}$$

where $\Phi(Q)$ is the standard normal distribution function. As a result, then, $Z_\varepsilon^Q = 2\Phi(Q) - 1$ a.s. for all ε , and consequently

$$\int_0^\varepsilon 1_{[0, Q]}(|W(s)|/s^\frac{1}{2}) ds = \varepsilon(2\Phi(Q) - 1) \quad \text{for all } \varepsilon, \quad \text{a.s.}$$

Differentiating both sides above with respect to ε , we arrive at a contradiction. Hence, for each $Q > 0$,

$$P\{\lim_{\varepsilon \downarrow 0} Z_\varepsilon^Q \text{ exists}\} = 0.$$

REFERENCES

[1] DUDLEY, R. M. (1973). Sample functions of the Gaussian process. *Ann. Probability* **1** 66-103.
 [2] FEDERER, H. (1969). *Geometric Measure Theory*. Springer-Verlag, Berlin.
 [3] GEMAN, D. (1977). On the approximate local growth of multidimensional random fields. *Z. Wahrscheinlichkeitstheorie und Verw. Gebiete* **38** 237-251.
 [4] GEMAN, D. and HOROWITZ, J. (1976). Local times for real and random functions. *Duke Math. J.* **43** 809-828.
 [5] KLEIN, R. (1976). A representation theorem on stationary Gaussian processes and some local properties. *Ann. Probability* **4** 844-849.
 [6] KONO, N. (1975). Sur la minoration asymptotique et la caractere transitoire des trajectoires des fonctions aleatoires gaussiennes a valeurs dans \mathbb{R}^d . *Z. Wahrscheinlichkeitstheorie und Verw. Gebiete* **33** 95-112.
 [7] SAKS, S. (1964). *Theory of the Integral*. Dover, New York.

- [8] YAGLOM, A. M. (1962). *An Introduction to the Theory of Stationary Random Functions*.
Prentice-Hall, Englewood Cliffs, N.J.

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