# LARGE DEVIATIONS OF GOODNESS OF FIT STATISTICS AND LINEAR COMBINATIONS OF ORDER STATISTICS

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Asymptotic behavior of large deviations of empirical distribution functions (df's) is considered. Borovkov (1967) and Hoadley (1967) obtained results for functionals continuous in the sup norm topology on the set of df's. Groeneboom, Oosterhoff, and Ruymgaart (1979) extended this to functionals continuous in a stronger  $\tau$ -topology. This result is now extended to functionals that are  $\tau$ -continuous only on a particular useful subset of df's. Applications to the Anderson-Darling statistic and linear combinations of order statistics are considered. We begin by correcting the work of Abrahamson (1967); from this the role of the key weight function  $\psi(t) = -\log t(1-t)$  is discovered. It is then exploited to the end indicated above, and it is considered as a weight function in tests of fit

1. Introduction and summary. Let  $X_1, \dots, X_n$  be iid F with empirical df  $\hat{F}_n$ . We consider weight functions  $\psi$  that are finite, positive and continuous on (0, 1). The weighted Kolmogorov-Smirnov statistics with weight function  $\psi$  are defined by

(1.1) 
$$K^{+} \equiv K_{\psi,n}^{+} \equiv \sup_{x} n^{1/2} [\hat{F}_{n}(x) - F(x)] \psi(F(x))$$

$$K^{-} \equiv K_{\psi,n}^{-} \equiv \sup_{x} n^{1/2} [F(x) - \hat{F}_{n}(x)] \psi(F(x))$$

$$K \equiv K_{\psi,n} \equiv K_{\psi,n}^{+} \vee K_{\psi,n}^{-}.$$

and the weighted Cramér-von Mises statistic is

(1.2) 
$$C \equiv C_{\psi,n} \equiv \int_{-\infty}^{\infty} n[\hat{F}_n(x) - F(x)]^2 \psi^2(F(x)) \ dF(x).$$

These statistics are often used to test the null hypothesis that F is the true df of the  $X_i$ 's. If  $\xi$  is uniform (0, 1), then  $X \equiv F^{-1}(\xi)$  (where  $F^{-1}(t) \equiv \inf\{x : F(x) \ge t\}$ ) has df F. By this inverse transformation we may suppose that

$$(1.1') K^+ = \sup_x U_n(F(x))\psi(F(x))$$

(1.2') 
$$C = \int_{-\infty}^{\infty} U_n^2(F(x)) \psi^2(F(x)) \ dF(x)$$

where  $\hat{\Gamma}_n$  is the empirical df of n independent uniform (0, 1) rv's  $\xi_1, \dots, \xi_n$  and

$$U_n(t) = n^{1/2} [\hat{\Gamma}_n(t) - t]$$
 for  $0 \le t \le 1$ 

is the associated empirical process. We note that

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(1.1") 
$$K^{+} = \sup_{0 < t < 1} U_n(t) \psi(t) \quad \text{if } F \text{ is continuous}$$

(1.2") 
$$C = \int_0^1 U_n^2(t) \psi^2(t) dt \qquad \text{if } F \text{ is continuous.}$$

We now assume that F is continuous.

The evaluation of Bahadur efficiency of tests based on such statistics requires a large deviation result for (1.1") and (1.2"). It is closely linked to Chernoff's theorem for independent Bernoulli trials; that is,

(1.3) 
$$n^{-1}\log P(\hat{\Gamma}_n(t) \ge t + a) \to -f(a, t)$$
 from below as  $n \to \infty$ ,

for  $0 \le t \le 1$ ,  $a \ge 0$  where (in the notation of Bahadur (1971))

$$(1.4) f(a,t) = \begin{cases} (a+t)\log\frac{a+t}{t} + (1-a-t)\log\frac{1-a-t}{1-t} & \text{if } 0 \le a \le 1-t \\ \infty & \text{if } a > 1-t. \end{cases}$$

This function is given in Table 1. We now define

(1.5) 
$$g_{\psi}(\alpha) \equiv \inf_{0 < t < 1} f(\alpha/\psi(t), t).$$

We consider  $g_1$  and  $g_2$  in Table 2, where

(1.6) 
$$g_i \equiv g_{\psi_i}$$
 with  $\psi_1(t) \equiv 1$  and  $\psi_2(t) = -\log[t(1-t)]$ .

THEOREM 1. Let  $X_1, X_2, \cdots$  be iid with continuous df F. Let  $\psi$  denote a finite, positive and continuous function on (0, 1) that is symmetric about  $t = \frac{1}{2}$  and for which  $\lim_{t\to 0} \psi(t)$  exists in  $[0, \infty]$ . Letting  $K^*$  denote any of  $K^+$ ,  $K^-$  or K, we have

(1.7) 
$$\lim n^{-1} \log P(K_{\psi,n}^{\#} \ge n^{1/2}a) = -g_{\psi}(a)$$

for each  $a \ge 0$ . Moreover, (when  $\psi$  puts too much weight in the tails)

(1.8) 
$$\lim \inf_{t\downarrow 0} (\log 1/t)/\psi(t) = 0 \quad implies \quad g_{\psi}(a) = 0 \quad for \ all \quad a \geq 0.$$

REMARK 1. The Theorem (1.7) still holds if the assumption of symmetry is dropped provided we replace  $g_{\psi}$  by  $g_{\psi}^{\#}$  where  $g_{\psi}^{\#}$  is any of

(1.9) 
$$g_{\psi}^{+} = g_{\psi}$$
 of (1.5),  $g_{\psi}^{-} = \inf_{0 < t < 1} f(\alpha/\psi(t), 1 - t), \quad g_{\psi}^{-} = g_{\psi}^{+} \wedge g_{\psi}^{-}.$ 

REMARK 2. The function  $\psi_2(t) \equiv -\log(t(1-t))$  of (1.6) satisfies Theorem 1. Moreover

$$(1.10) f(a/\psi_2(t), t) \to a as t \downarrow 0 or as t \uparrow 1,$$

and the positive minimum of this function will be attained in (0, 1). In fact

(1.11) 
$$g_{\psi_0}(a) \sim e^2 a^2 / 8$$
 as  $a \downarrow 0$ ,

while the point of minimization  $t_a$  converges to a solution of  $t(1-t)=\exp(-2)$  as  $a\downarrow 0$ . Moreover,  $g_{\psi_2}(a)\to \infty$  as  $a\to \infty$  (see Section 4). The function  $\psi_1(t)\equiv 1$  of (1.6) also satisfies Theorem 1. It is known from Lemma 5.1 of Bahadur (1971) that

$$(1.12) g_{\psi_1}(a) \sim 2a^2 \text{ as } a \downarrow 0$$

and that this function approaches  $\infty$  as  $a \uparrow 1$ .

Remark 3. The functions  $\psi_{\delta}(t) \equiv [t(1-t)]^{-\delta}$  with  $\delta > 0$ , satisfy Theorem 1 and (1.8); thus their g functions are identically zero. Intuitively, these functions are too severe because they put too much weight on the extreme order statistics; we now demonstrate

Table 1.  $f(a,t) = (a+t)\log\frac{a+t}{t} + (1-a-t)\log\frac{1-a-t}{1-t}.$ 

.60 .50	.0053	.0216 .0201 .0204	.0915 .0823 .0811	.226 .193 .184	.511 .368 .335	.693 .551	.916					
.40 .30 .20	.0051 .0058 .0074	.0204	.0872	.192	.339 .382	.534 .583	.794 .832	1.20 1.15	1.61			
.10 .05	.0122	.0444	.1537 .2251	.311 .434	.511 .688	.751 .983	1.033 1.318	1.36 1.70	1.76 2.13	$2.30 \\ 2.65$	3.00	
.01	.0588	.1690	.4611	.815	1.217	1.661	2.144	2.67	3.25	3.89	4.25	<b>4.</b> 61
t/a	.05	.10	.20	.30	.40	.50	.60	.70	.80	.90	.95	.99

Table 2.  $\hat{g}_1(a) = 2a^2; \quad \hat{g}_2(a) = e^2a^2/8 = .9236a^2$  Interpolate linearly in columns. See (1.6) for definition of  $g_1$  and  $g_2$ .

а	$g_1(a)/\hat{g}_1(a)$	$g_2(a)/\hat{g}_2(a)$		
.00	1.0000	1.0000		
.01	1.0000	.9917		
.02	1.0001	.9835		
.03	1.0002	.9755		
.05	1.0006	.9596		
.07	1.0011	.9442		
.10	1.0022	.9219		
.12	1.0032	.9076		
.15	1.0051	.8868		
.20	1.0091	.8541		
.25	1.0144	.8237		
.30	1.0210	.7953		
.40	1.0390	.7440		
.50	1.0646	.6989		
.60	1.1009	.6591		
.70	1.1540	.6236		
.80	1.2394	.5918		
.90	1.4211	.5632		
.97	1.9342			
1.00		.5373		
1.50		.4368		
2.00		.3679		
3.00		.2790		

this. Let  $\xi_{n:1} \leq \cdots \leq \xi_{n:n}$  denote the order statistics of uniform (0, 1) rv's  $\xi_1, \dots, \xi_n$ . Then

$$\begin{split} [K_{\psi_{\delta},n}^+ \geq n^{1/2}a] \supset [1/n \geq a(\xi_{n:1}(1-\xi_{n:1}))^{\delta} + \xi_{n:1}] \\ \supset [1/n \geq a\xi_{n:1}^{\delta} + \xi_{n:1} \quad \text{and} \quad \xi_{n:1} \leq 1/(2n)] \\ \supset [\xi_{n:1} \leq (2n)^{-1} \wedge (2na)^{-1/\delta}] \supset [\xi_1 \leq (2n)^{-1} \wedge (2na)^{-1/\delta}], \end{split}$$

and even the probability of the smallest event does not go to zero exponentially fast since  $\lim_{n \to \infty} n^{-1} \log((2n)^{-1} \wedge (2na)^{-1/\delta}) = 0$ .

Asymptotic distributions of truncated or restandardized forms of  $K_{\psi_{1/2},n}$  are computed in Jaeschke (1979) and Eicker (1979). Again, truncation or standardization is needed to avoid too heavy weights on the extreme order statistics.

COROLLARY 1. If  $X_1, X_2, \cdots$  are iid with continuous df G and if

(1.13) 
$$K_{\psi,n}^{\#}/n^{1/2} \to \phi(G) \text{ a.s.} \quad as \ n \to \infty$$

for some  $\phi(G) \ge 0$ , then the exact Bahadur slope of the  $K_{\psi,n}^*$ -test for testing against the alternative G is

$$(1.14) 2g_{\nu}(\phi(G)).$$

REMARK 4. Abrahamson (1967) states a similar theorem. However, her proof fails (the part when i=1 is incorrect) in case  $\psi(t)\to\infty$ , as  $t\downarrow 0$  or as  $t\uparrow 1$ . Moreover, her conclusion is incorrect, because (as in paragraph two of her proof) she claims that  $f(a/\psi(t),t)\to\infty$  as  $t\downarrow 0$  or  $t\uparrow 1$ ; whereas in the case  $\psi(t)=(t(1-t))^{-1/2}$  she considered,  $f(a/\psi(t),t)\to 0$  as  $t\downarrow 0$  or  $t\uparrow 1$ . Thus, her basic expansion of her  $\log \rho_{\psi}^*(\epsilon)$  following her (3.16) is incorrect, and also her Table 1 is incorrect.

We will now relate Theorem 1 to some papers in the literature that consider the Sanov problem, and we will present our second theorem along these lines. To this end, we introduce the following notation. Let D denote the set of all df's on the real line  $\mathbb{R}$ . For F,  $G \in D$  the Kullback-Leibler number K(G, F) of G with respect to F is defined by

(1.15) 
$$K(G, F) = \begin{cases} \int_{R} \log\left(\frac{dG}{dF}\right) dG & \text{if } G \ll F \\ \infty & \text{otherwise} \end{cases}$$

For a set  $\Omega$  of df's we define

$$(1.16) K(\Omega, F) = \inf\{K(G, F): G \in \Omega\}.$$

REMARK 5. For  $\psi$  as in Theorem 1 or Remark 1, define  $T_F:D\to\mathbb{R}$  by

(1.17) 
$$T_F(G) = \sup_{x \in R} |G(x) - F(x)| \psi(F(x)).$$

For  $a \ge 0$  let

$$\Omega_a = \{G: T_F(G) \ge a\}.$$

In Section 4, it is shown that if  $\psi$  is bounded, then

$$(1.19) \quad \lim n^{-1} \log P(K_{\psi,n}^{\#} \ge n^{1/2}a) = \lim n^{-1} \log P(T_F(\hat{F}_n) \ge a) = -K(\Omega_a, F) = -g_{\psi}(a)$$

follows from Theorem 3.1 of Groeneboom, Oosterhoff and Ruymgaart (1979) (henceforth referred to as GOR(1979)). This also follows from Hoadley's (1967) Theorem 1. That the conclusion (1.19) holds more broadly follows from (1.7) and (1.25) below.

REMARK 6. As a matter of curiosity, we note that Theorem 1 implies that Sievers' (1976) conditions are satisfied. In fact, the proof of Theorem 1 shows that the event  $\{\hat{F}_n \in \Omega_a\}$  is contained in the union of events  $\bigcup_{i=1}^{k_n} \{\hat{F}_n \in U_{i,n}\}$ , where  $k_n = o(n)$ ,  $U_{i,n}$  is a convex set of df's and  $\lim_{n\to\infty} K(\bigcup_{i,n}, F) \geq K(\Omega_a, F)$ . Using this fact, one can show that Sievers' (1976) Condition II is satisfied. However, it seems difficult to prove this directly (i.e. without essentially repeating the proof of Theorem 1).

Hoadley (1967) established the Sanov result (1.18) for certain functionals  $T:D\to\mathbb{R}$  uniformly continuous in the sup norm topology on D. GOR (1979) extended this to certain functionals continuous in a stronger  $\tau$ -topology. This result is now extended to certain functionals that are  $\tau$ -continuous only on particular subsets of D. (The  $\tau$ -topology is described in detail in Section 4 just after the proof of Corollary 1.)

THEOREM 2. Let F be a continuous df on  $\mathbb{R}$  and let  $T:D \to \mathbb{R}$  be  $\tau$ -continuous on

(1.20) 
$$D_m \equiv \{G \in D: d_F(G, F) \le m\}, \text{ where } d_F(G, F) \equiv \sup_{x \in \mathcal{R}} |G(x) - F(x)| \psi_2(F(x))$$
 for each  $m > 0$ . Let

$$\Omega_a \equiv \{G \in D: T(G) \ge a\},\$$

and suppose  $K(\Omega_a, F) < \infty$ . Then

(1.22) 
$$\lim \sup_{n\to\infty} n^{-1} \log P(T(\hat{F}_n) \ge a) \le -K(\Omega_a, F).$$

Moreover, if the function  $t \to K(\Omega_t, F)$  is right continuous at t = a and  $u_n \to 0$  then

$$\lim_{n\to\infty} n^{-1}\log P(T(\hat{F}_n) \ge \alpha + u_n) = -K(\Omega_a, F).$$

There are several important facts that prove useful in the applications and throughout the proofs. To avoid the impression of circular arguments, we prove these in this section.

PROPOSITION 1. Let  $D_m$  and  $d_F$  be defined as in (1.20).

- (i) For  $m \equiv m_{F,\Omega}$  sufficiently large
- (1.24)  $K(\Omega \cap D_m, F) = K(\Omega, F)$  for any F and  $\Omega$  having  $K(\Omega, F) < \infty$ .
- (ii) For continuous F any  $\psi$  positive on (0, 1),
- (1.25)  $K(\Omega_a, F) = g_{\psi}(a) \text{ where } \Omega_a \text{ is defined as in (1.18)}.$
- (iii) Suppose F is continuous and Gr are arbitrary df's; then
- $(1.26) \quad \lim \sup_{r \downarrow 0} r^{-2} K(G_r, F) \le B \text{ implies } \lim \sup_{r \downarrow 0} r^{-1} d_F(G_r, F) \le (8B/e^2)^{1/2}.$

PROOF. (i) Since  $t\psi_2(t)$  is bounded for  $t \le \frac{1}{2}$  and likewise  $(1-t)\psi_2(t)$  is bounded for  $t \ge \frac{1}{2}$ , we have for all large m that

$$D_m^c \subset \{G \in D: \sup_{F(x) \le 1/2} G(x) \psi_2(F(x)) \ge m/2\}$$

$$\bigcup \{G \in D: \sup_{F(x) > 1/2} (1 - G(x)) \psi_2(F(x)) \ge m/2\}.$$

Also for m sufficiently large, we have

$$\sup_{F(x) \le 1/2} G(x) \ \psi_2(F(x)) \ge m/2 \quad \text{implies} \quad (K(G, F)) \ge m/4$$

(and a similar implication with 1 - G(x)) since

$$K(G, F) \ge G(x)\log(G(x)/F(x)) + (1 - G(x))\log(1 - G(x))/(1 - F(x)).$$

Thus for some large  $m_1$  we have that

$$K(G, F) \ge m_1/4 \ge K(\Omega, F) + 1$$
 whenever  $d_F(G, F) \ge m_1$ .

(ii) Since  $F \circ F^{-1}$  is the identity for continuous df's F, we can define the df  $G_t$  on  $\mathbb{R}$  so that the df  $\overline{G}_t \equiv G_t \circ F^{-1}$  on the unit interval has the uniform on intervals density

(a) 
$$\bar{g}_t(u) = \begin{cases} [t + a/\psi(t)]/t & \text{if } 0 < u < t \\ [(1-t) - a/\psi(t)]/(1-t) & \text{if } t < u < 1. \end{cases}$$

Note that  $G_t \in \Omega_a$  since  $T_F(G_t) = a$ , and note that

(b) 
$$K(G_t, F) = f(\alpha/\psi(t), t).$$

Thus

(c) 
$$K(\Omega_a, F) \leq \inf_t K(G_t, F) = g_{\psi}(a).$$

To obtain the reverse inequality, we note that for  $\epsilon > 0$  there exists a df  $G_a^* \in \Omega_a$  for which

(d) 
$$\dot{K}(G_a^*, F) < K(\Omega_a, F) + \epsilon.$$

Now  $G_a^* \in \Omega_a$  implies  $|\bar{G}_a^*(t_a) - t_a| \ge a/\psi(t_a)$  for some  $t_a$  in (0, 1). But the df  $\bar{G}_a^\circ$  that is linear from (0, 0) to  $(t_a, \bar{G}_a^*(t_a))$  to (1, 1) has the smallest Kullback-Leibler number, when compared to the Uniform (0, 1) df, among all df's that pass through the point  $(t_a, \bar{G}_a^*(t_a))$ ; and its information number is  $\ge K(G_{t_a}, F)$  (note that  $\bar{G}_a^\circ = \bar{G}_{t_a}$  when  $\bar{G}_a^*(t_a) - t_a = a/\psi(t_a)$ ). Thus for  $\epsilon > 0$ 

(e) 
$$K(\Omega_a, F) + \epsilon \ge K(G_a^*, F) \ge K(G_{t_a}, F) = f(a/\psi(t_a), t_a) \ge g_{\psi}(a).$$

(iii) Assume there exist df's  $G_r$  contradicting (1.26), so that  $c_r \equiv r^{-1}d_F(G_r, F)$  satisfies  $\limsup c_r > (8B/e^2)^{1/2}$ . Then with  $\Omega_r$  as in (1.18) we have

(1.27) 
$$B \ge \limsup r^{-2}K(G_r, F) \ge \limsup r^{-2}K(\Omega_{rc_r}, F)$$

$$= \limsup r^{-2}g_{\psi}(rc_r) \quad \text{by (1.25) for any such } \psi$$

$$= \limsup r^{-2}e^2r^2c_r^2/8 \quad \text{for } g_2 \text{ by (1.11)}$$

$$> B$$

providing the necessary contradiction.

Results analogous to (1.11) for other  $\psi$ 's produce, in conjunction with (1.27), results analogous to (1.26) for other  $\psi$ 's; equation (1.12) allows this for  $\psi_1$ . Thus, if F is continuous and  $G_r$  are arbitrary df's, then

(1.28) 
$$\limsup_{r\downarrow 0} r^{-2}K(G_r, F) \leq B$$
 implies  $\limsup_{r\downarrow 0} r^{-1} \sup_{x} |G(x) - F(x)| \leq (B/2)^{1/2}$ .

2. Applications to tests of fit. Three statistics will be considered in this section: the two-sided Kolmogorov-Smirnov statistics, call them  $K_1$  and  $K_2$ , of (1.1) with weight functions  $\psi_1(t) \equiv 1$  and  $\psi_2(t) \equiv -\log(t(1-t))$ , and the Anderson-Darling statistic A that corresponds to (1.2) with weight function  $[t(1-t)]^{-1/2}$ .

The key results for  $K_1$  and  $K_2$  are contained in (1.7), (1.11), and (1.12). The corollary below presents the corresponding results for A.

EXAMPLE 1. The Anderson-Darling statistic  $A \equiv A_n = T(\hat{\Gamma}_n)$ , where  $T: D \to \mathbb{R} \cup \{\infty\}$  is defined via (1.2'') by

(2.1) 
$$T(G) = \int_0^1 [G(t) - t]^2 / [t(1-t)] dt \quad \text{for } G \in D.$$

We will now show that this functional is finite and  $\tau$ -continuous on each subset

$$(2.2) D_m = \{G \in D : \sup_{t \in (0,1)} |G(t) - t| \psi_2(t) \le m\}$$

with m>0. To see this, notice that the integrals over  $(0,\eta]$  and  $[1-\eta,1)$  in (2.1) tend to zero as  $\eta\downarrow 0$ , uniformly in  $G\in D_m$  for each fixed m>0. Also, for fixed  $\eta\in (0,\frac{1}{2})$ , the functional  $G\to \int_{\eta}^{1-\eta} [G(t)-t]^2/[t(1-t)] dt$  is obviously  $\tau$ -continuous (even continuous with respect to the topology of weak convergence) on D. From these two statements the  $\tau$ -continuity of T on  $D_m$  follows. (However, T is not continuous with respect to the  $\tau$ -topology on the whole set of df's.) From Theorem 2, we obtain (2.3) below.

COROLLARY 2. Let F denote the uniform (0, 1) distribution. Let  $\Omega_a = \{G \in D : T(G) \geq a\}$  for each  $a \geq 0$ , with T as in (2.1). Then:

(2.3) 
$$\lim_{n\to\infty} n^{-1}\log P(A_n \ge a) = -K(\Omega_a, F) \quad \text{for each} \quad a \ge 0.$$

(2.4) The function 
$$a \to K(\Omega_a, F)$$
,  $a \ge 0$ , is continuous.

(2.5) 
$$K(\Omega_a, F) = a + o(a) \quad as \quad a \downarrow 0.$$

REMARK 7. Result (2.5) of Corollary 2 appears in Nikitin (1976), where it is stated that it can be derived using the theory of branching of solutions of non-linear equations in Banach spaces in Vainberg and Trenogin (1974). Some of the statements in Nikitin (1976) are proved along these lines in Nikitin (1979) and a proof of (2.5) is announced as a part of a forthcoming paper, Nikitin (1980). A more straightforward proof of (2.5) is given in Section 4 below. Result (2.3) of Corollary 2 is used implicitly in Nikitin's paper, he only remarks that asymptotic probabilities of large deviations of functionals of empirical distribution functions can be computed by finding the Kullback-Leibler numbers with respect to the corresponding set of distribution functions. It is clear, however, that the functionals, and also the Kullback-Leibler numbers of the sets of distribution functions, have to satisfy certain regularity conditions (for examples where the relation between Kullback-Leibler numbers and large deviation limits does not hold, see e.g. Bahadur and Zabell (1979) and GOR (1979)).

Local Bahadur efficiencies of the Anderson-Darling test with respect to other well-known goodness of fit tests are computed in Nikitin (1976), indicating that the Anderson-Darling test is to be preferred for normal and logistic location alternatives and that Watson's test is best for Cauchy location alternatives.

EXAMPLE 2. (Logistic scale alternatives). Let the df  $F(x) = 1/(1 + \exp(-x))$  have density f; note that F(1 - F) = f. For  $\theta > 0$  let  $F_{\theta} = F(\cdot/\theta)$ . Then two times

$$\int_{R} [F(x/\theta) - F(x)]^{2} / [F(x)(1-F)] f(x) \ dx$$

behaves as

(2.6) 
$$2(1-\theta)^2 \int_R x^2 f^2(x) \ dx = .430(1-\theta)^2 \quad \text{as} \quad \theta \to 1.$$

Thus, by (2.5), the exact slope of the Anderson-Darling statistic for testing  $\theta = 1$  against  $\theta \neq 1$  also behaves as (2.6). The analogous result for the  $K_2$ -test comes from

(2.7) 
$$\sup_{x \in \mathcal{R}} |F(x/\theta) - F(x)| \psi_2(F(x)) \\ \sim |(1 - \theta) \sup_{x \in \mathcal{R}} x f(x) \log(1/f(x))| = .476 |1 - \theta| \quad \text{as} \quad \theta \to 1$$

and (1.11); so that the exact slope of the  $K_2$ -statistic behaves as

(2.8) 
$$(2e^2/8)(.476(1-\theta))^2 = .419(1-\theta)^2 \quad \text{as} \quad \theta \to 1.$$

(The maximum of .476 is attained at approximately  $x \approx 2.168$ ). Analogously, the  $K_1$ -statistic has exact slope

(2.9) 
$$4(.224(1-\theta))^2 = .201(1-\theta)^2$$
 as  $\theta \to 1$ 

(since xf(x) attains a maximum of  $\approx .224$  at  $x \approx 1.543$ ). Thus the Bahadur efficiencies of the A and  $K_2$  tests with respect to the  $K_1$  test converge to 2.14 and 2.08 respectively as  $\theta \to 1$ ; see Table 3 for the rate of convergence to 2.08.

Exact Bahadur efficiencies of  $K_1$  and  $K_2$ -tests for any fixed alternative can be computed from (1.13) via Table 2. This was done for logistic and normal location and scale alternatives; the results are given in Table 3. We note that the weight function  $\psi_2$  outperforms  $\psi_1$  for scale alternatives with a reversal for location alternatives.

3. Applications to linear combinations of order statistics. Theorem 2 can also be used to obtain large deviation results (and hence Bahadur efficiencies) for linear combinations of order statistics. We consider linear combinations of order statistics of the form

(3.1) 
$$T(\hat{F}_n) = \int_0^1 J(u) \hat{F}_n^{-1}(u) du,$$

	L	$\mathbf{s}$	$\mathbf{a_1}$	$\mathbf{a_2}$	$\mathbf{e}_{2,1}$
Logistic scale	0.00	2.00	.150	.259	1.13
	"	1.50	.0897	.167	1.38
	"	1.10	.0213	.0439	1.88
	"	1.01	.00223	.00472	2.07
Normal Scale	0.00	2.00	.161	.282	1.13
	"	1.50	.0968	.185	1.46
	"	1.10	.0231	.0498	2.07
	″	1.01	.00241	.00540	2.31
Logistic location	1.00	1.00	.245	.378	.820
	.50	"	.124	.178	.821
	.10	"	.0250	.0347	.863
	.01	"	.00250	.00347	.885
Normal location	1.00	1.00	.3829	.6983	.926

TABLE 3.  $a_i \equiv \sup_x [F(x) - F((x-L)/S)] \psi_i(F(x))$  for i = 1, 2.  $e_{2,1}$  is the Bahadur efficiency of the  $K_{\psi_2}^-$ -test with respect to the  $K_{\psi_1}^-$ -test; here we are testing F while F((.-L)/S) is actually true.

where J is an  $L_1$ -function (i.e.  $\int_0^1 |J(u)| du < \infty$ ) and for a df  $G \in D$ , the inverse  $G^{-1}$  is defined by  $G^{-1}(u) = \inf\{x \in \mathbb{R} : G(x) \ge u\}$ .

.1974

.03988

.00399

.3048

.05582

.00553

.865

.865

884

In GOR (1979), large deviation results for statistics of the form (3.1) with score functions having support contained in  $[\alpha, \beta] \subset (0, 1)$  are given. By Theorem 2, we can obtain an extension of these results to statistics with score functions J, whose support is not necessarily restricted to a closed subinterval of (0, 1). This is shown in the next corollary.

COROLLARY 3. Suppose that F is a continuous df and that J is an  $L_1$ -function satisfying

(i) 
$$\int_{1/2}^1 |J(u)F^{-1}(1-\exp(-c/(1-u)))| \ du < \infty \quad and$$
 
$$\int_0^{1/2} |J(u)F^{-1}(\exp(-c/u))| \ du < \infty \quad for \ each \quad c>0.$$

(ii) 
$$J \ge 0$$
 on an interval  $(\gamma, \delta) \subset (0, 1)$  and  $\int_0^{\delta} J(u) du > 0$ .

Let  $\Omega_r = \{G \in D : r \leq \int_0^1 J(u) G^{-1}(u) du < \infty\}$ . Then, if  $\hat{F}_n$  is the empirical df of independent random variables  $X_1, \dots, X_n$  with df F,

(3.2) 
$$\lim_{n\to\infty} n^{-1}\log P(T(\hat{F}_n) \ge r) = -K(\Omega_r, F).$$

.50

.10

.01

(3.3) The function  $t \to K(\Omega_t, F)$  is a continuous mapping from  $\mathbb{R}$  into  $[0, \infty]$ , where  $[0, \infty]$  is endowed with the topology of the extended real line.

REMARK 8. It is shown in Oosterhoff (1978) that if J is monotonically non-decreasing, a condition like (i) is not needed to ensure (3.2).

EXAMPLE 3. Let J(u) = u(1-u),  $u \in (0, 1)$ , and let F be the logistic df  $F(x) = (1 + \exp(-x))^{-1}$ . It can be shown that Corollary 3 applies, and that

(3.4) 
$$K(\Omega_r, F) \sim 6r^2 \quad \text{as} \quad r \downarrow 0.$$

Technical details can be found in Groeneboom and Shorack (1980).

Consider location alternatives  $F_{\theta}(x) = F(x - \theta)$ ,  $\theta > 0$ . Then  $\int_0^1 J(u) F_{\theta}^{-1} du = \theta \int_0^1 J(u) du = (\frac{1}{2}\theta) \theta$  and hence by (3.4) the Bahadur slope of the sequence  $\{\int_0^1 J(u) \hat{F}_n^{-1}(u) du\}$  at  $\theta$  is asymptotically equivalent to  $12(\int_0^1 J(u) F^{-1}(u) du)^2 = (\frac{1}{2}\theta)^2$ , as  $\theta \downarrow 0$ . The Bahadur slope of the most powerful test for testing  $\mathcal{H}_0: \theta = 0$  against an alternative  $\theta > 0$  is given by

$$c(\theta) = 2\{\log 2 - \log(1 + e^{2\theta}) + 2e^{\theta}(e^{2\theta} - 1)^{-1}(\theta e^{\theta} - \arctan(\frac{1}{2}(e^{\theta} - e^{-\theta})))\}$$

(see Groeneboom and Oosterhoff (1977), page 21). By Taylor series expansion it is seen that  $c(\theta) \sim (\frac{1}{2})\theta^2$ , as  $\theta \downarrow 0$ . Hence the Bahadur slope of a test based on the sequence of test statistics  $\{\int_0^1 J(u)\hat{F}_n^{-1}(u)\ du\}$  coincides locally with the Bahadur slope of the most powerful test, as  $\theta \downarrow 0$ . It follows from Corollary 2 that the Anderson-Darling test also has this property.

Example 4. If  $\Phi$  is the standard normal df, then

$$\Phi^{-1}(\exp(-c/u)) \sim -\sqrt{2c/u}, \quad \text{as} \quad u \downarrow 0 \quad \text{and}$$

$$\Phi^{-1}(1 - \exp(-c/(1 - u))) \sim \sqrt{\frac{2c}{1 - u}}, \quad \text{as} \quad u \uparrow 1.$$

Hence, in this case, Corollary 3 allows us even to take score functions J which tend to infinity as  $u \downarrow 0$  and  $u \uparrow 1$ , for example score functions of type  $J(u) = \{u(1 - u)\}^{-(1/2) + \epsilon}$  with  $u \in (0, 1)$  and  $\epsilon > 0$ .

#### 4. Proofs.

PROOF OF THEOREM 1. We first consider  $K^+$ . Fix  $a \ge 0$ . For any fixed t we have from (1.3) that

(a) 
$$n^{-1}\log P(K^+ \ge n^{1/2}a) \ge n^{-1}\log P(\hat{\Gamma}_n(t) \ge t + a/\psi(t)) \to -f(a/\psi(t), t).$$

Hence

(b) 
$$\lim\inf n^{-1}\log P(K^{+}\geq n^{1/2}a)\geq -\inf_{t\in(0,1)}f(a/\psi(t),t)=-g_{\psi}(a).$$

We now turn to the other inequality. Note first that

(c) 
$$\lim_{t \downarrow 0} \psi(t) = \infty$$
 implies  $f(a/\psi(t), t) = (a/\psi(t))(\log 1/t) + o(1)$ 

where  $|o(1)| \le \epsilon$  for all  $0 < t \le \text{some } \delta_{\epsilon}$ .

If the condition of (1.8) holds, it follows immediately from (c) that  $g_{\psi}(a) = 0$ . Thus (1.7) follows trivially from (b) in this case. The conclusion (1.7) also holds trivially when a = 0. Thus we now assume that a > 0 and

(d) 
$$\lim \inf_{t \downarrow 0} (\log 1/t)/\psi(t) > 0.$$

For each n we choose numbers  $t_{n,1} < \cdots < t_{n,m(n)}$  such that

(e) 
$$1 - 2/n < t_{n,i}/t_{n,i+1} < 1 - 1/n$$
 for  $1 \le i < m(n)$  and

(f) 
$$t_{n,1} = 1 - t_{n,m(n)} = \exp(-n^2).$$

For  $1 \le i < m(n)$ , we let  $\psi_{n,i} \equiv \max\{\psi(t) : t_{n,i} \le t \le t_{n,i+1}\}$ . Then,

$$P(\sup_{t \in [t_n, t_{n,i+1}]} [\widehat{\Gamma}_n(t) - t] \psi(t) \ge a) \le P(\widehat{\Gamma}_n(t_{n,i+1}) - t_{n,i} \ge a/\psi_{n,i})$$

(g) 
$$\leq \exp(-nf(a/\psi_{n,i} + t_{n,i} - t_{n,i+1}, t_{n,i+1})) \text{ by (1.3)}$$
 
$$\leq \exp(-n(g_{\psi}(a) - \epsilon)) \text{ for all } i \text{ and large } n$$

by continuity and an argument analogous to (c). Note that

(i)

$$m(n) = o(n^3)$$

(solve  $(1 - 1/n)^m = \exp(-n^2)$  for m). Also

$$P(\sup_{t \le t_{n,1}} | \hat{\Gamma}_n(t) - t | \psi(t) \ge a) = P(\sup_{t \ge t_{n,m(n)}} | \hat{\Gamma}_n(t) - t | \psi(t) \ge a)$$

$$= P(\sup_{t \le t_{n,1}} [\hat{\Gamma}_n(t) - t] \psi(t) \ge a) \quad \text{for all large } n$$

$$\le P(\hat{\Gamma}_n(t_{n,1}) > 0) \le nP(\xi_1 \le t_{n,1}) = n \exp(-n^2)$$

Hence (g), (h) and (i) give, for arbitrary  $\epsilon > 0$ ,

 $\lim\sup n^{-1}\log P(K^+ \ge n^{1/2}a)$ 

(j) 
$$\leq \limsup n^{-1} \log\{m(n) \exp(-n(g_{\psi}(a) - \epsilon)) + 2 \exp(-n^2/2)\} = -g_{\psi}(a) + \epsilon.$$

 $\leq \exp(-n^2/2)$ 

for all large n

Thus (b) and (j) complete the proof for  $K^+$ .

Finally,  $K^-$  is completely analogous to  $K^+$ ; while

$$P(K^{+} \ge n^{1/2}a) \le P(K \ge n^{1/2}a)$$

$$\le P(K^{+} \ge n^{1/2}a) + P(K^{-} \ge n^{1/2}a) \le 2P(K^{+} \ge n^{1/2}a). \quad \Box$$

PROOF OF (1.11). Let  $\phi \equiv \phi_a(t) \equiv t + a/\psi_2(t) \equiv t + a/\psi$  for  $t \in (0, 1)$ . Then for a, t such that  $\phi < 1$  we have

(a) 
$$f \equiv f(\alpha/\psi, t) = \phi \log(\phi/t) + (1 - \phi)\log((1 - \phi)/(1 - t)).$$

Letting

(b) 
$$\phi' = (d/dt)\phi_a(t) = 1 + \frac{a(1-2t)}{t(1-t)\psi^2},$$

we note that

$$f' \equiv (d/dt)f(a/\psi, t) = \phi'[\log(\phi/t) - \log((1-\phi)/(1-t))]$$
(c) 
$$-(\phi/t) + ((1-\phi)/(1-t))$$

$$= \phi'[\log(1 + a/(t\psi)) - \log(1 - a/((1-t)\psi))] - a/(t(1-t)\psi).$$

We first fix  $0 < \eta < 1$ , and then fix  $\epsilon > 0$  so that  $(1 - 2\epsilon)x - \log(1 + x) > 0$  for  $x \ge \eta$ . Let  $T \equiv T_a(t) \equiv (1/\psi)\log(1 + a/(t\psi))$ . We will now choose  $\delta$  so small that (for small a) the function  $f(a/\psi, t)$  does not reach a minimum on  $[0, \delta]$ . We will do this by showing that when a is sufficiently small and  $t \le \text{some } \delta$  we have

- (i) f' < 0 if  $a/(t\psi) \le \eta$ ,
- (ii) f' < 0 if  $a/(t\psi) > \eta$  and  $T \le \epsilon$ , and
- (iii)  $f > \inf_r f(\alpha/\psi(r), r)$  if  $T > \epsilon$ .

It follows by analogy that  $f(a/\psi, t)$  does not reach a minimum on  $[1 - \delta, 1]$  when a is small. It is trivial that

(d) 
$$f = a^2/[2t(1-t)\psi^2] + o(a^3) \text{ uniformly for } t \in [\delta, 1-\delta];$$

while differentiating the main term in (d) yields a minimum of  $e^2\alpha^2/8$  whenever t solves  $t(1-t)=e^{-2}$ . Thus the result holds.

It remains only to prove (i), (ii) and (iii):

(i) In equation (c) we expand  $\log(1 + a/(t\psi))$  to three terms, expand  $\log(1 - a/((1 - t)\psi))$  in an infinite series, and plug in (b) to obtain

(e) 
$$f' \le \frac{a^2}{[t(1-t)\psi]^2} \left\{ \frac{1-2t}{\psi} + \frac{t^2 - (1-t)^2}{2} + \frac{\phi'}{3} \left\{ \left(\frac{a}{t\psi}\right)^3 + \sum_{j=3}^{\infty} \frac{3}{j} \left(\frac{a}{(1-t)\psi}\right)^j \right\} \right\}$$

whenever  $a/[t(1-t)\psi] < \eta$ . For  $t < \text{some } \delta_1$  and  $a/(t\psi) \le \eta$  the rhs of (e) is dominated by the term involving  $-(1-t)^2/2$ .

(ii) In equation (c) we have, for  $t < \text{some } \delta_2$ , that

$$\begin{split} f' &\leq [1 + (a/(t\psi))/\psi] [\log(1 + a/(t\psi)) - \log(1 - a/((1-t)\psi))] - a/[t(1-t)\psi] \\ &\leq \log(1 + a/(t\psi)) + (a/(t\psi))\epsilon - [1 + a/(t\psi^2)] [\log(1 - a/((1-t)\psi)) - a/(t\psi) \\ &\leq (1 - 2\epsilon)(a/(t\psi)) + (a/(t\psi))\epsilon - a/(t\psi) - [1 + a/(t\psi^2)] [\log(1 - a/((1-t)\psi)) \\ &< -\epsilon a/(t\psi) + [1 + a/(t\psi^2)] 2at/(\psi t(1-t)) \\ &= -[\epsilon - 2(1 + a/(t\psi^2))(t/(1-t))] a/(t\psi) < 0 \end{split}$$

using  $T \le \epsilon$  in the second inequality, the definition of  $\epsilon$  in the third inequality and  $a \le$  some  $a_0$ .

(iii) From (a) we have for  $t < \text{some } \delta_3$ 

$$f \ge (a/\psi) \log(1 + a/(t\psi)) + (1 - \phi) \log(1 - a/((1 - t)\psi))$$
  
 
$$\ge a\epsilon = (a\epsilon/2) + a\epsilon/2$$

using  $T > \epsilon$ . But from (d) we have  $f \le \text{(some } K)$   $\alpha^2$  on  $[\delta, 1 - \delta]$  with  $\delta = \delta_1 \wedge \delta_2 \wedge \delta_3$ . Thus for small  $\alpha, f(t) > \inf_r f(\alpha/\psi(r), r)$  if  $T > \epsilon$  and  $t \le \delta$ .

Note that (1.10), (1.4), and a slight modification of the above argument yield  $g_{\psi}(a) \rightarrow \infty$  as  $a \rightarrow \infty$ .

PROOF OF COROLLARY 1. Use Theorem 1 and Theorem 7.2 of Bahadur (1971).

We now define the  $\tau$ -topology. Let  $\mathscr{P} = \{B_1, \cdots, B_m\}$  denote a partition of  $\mathbb{R}$  into Borel measurable sets. Let

$$(4.1) K_{\mathscr{P}}(G, F) \equiv \sum_{i=1}^{m} P_G(B_i) \log P_G(B_i) / P_F(B_i)$$

where  $P_F$  (or  $P_G$ ) is the probability distribution corresponding to F (or G). For a set of df's  $\Omega$  let

$$(4.2) K_{\mathscr{P}}(\Omega, F) = \inf\{K_{\mathscr{P}}(G, F): G \in \Omega\}.$$

(We use the conventions  $0 \log \infty = 0$  and  $a \log 0 = -\infty$  for a > 0.) Consider also the pseudometric  $d_{\mathscr{P}}$  on D given by

$$(4.3) d_{\mathscr{P}}(G, F) \equiv \max_{1 \le i \le m} |P_F(B_i) - P_G(B_i)|.$$

The topology on D generated by all such  $d_{\mathscr{P}}$  will be denoted by  $\tau$  (cf. GOR (1979)); thus  $\tau$  is the smallest topology such that the sets  $\{G \in D: d_{\mathscr{P}}(G, F) < \epsilon\}$  are open for each  $\epsilon > 0$ , each  $F \in D$  and each finite partition  $\mathscr{P}$ . In fact, these sets form a subbase for the topology.

PROOF OF (1.19). For bounded  $\psi$ , the function  $T_F$  of (1.17) is clearly  $\tau$ -continuous. Thus, the second equality of (1.19) follows from Theorem 3.1 of GOR (1979). The third equality of (1.19) follows from (1.25).

REMARK 9. If  $\psi$  is unbounded, then the proof of (1.19) fails since  $T_F$  is not  $\tau$ -continuous (hence a fortiori not continuous with respect to the topology on D induced by the supremum metric, cf. GOR (1979) Lemma 2.1). Even the weaker condition

$$K(\Omega_a, F) = \sup\{K_{\mathscr{P}}(\Omega_a, F) : \mathscr{P} \text{ is a finite partition of } \mathbb{R}\}\$$

fails, if  $K(\Omega_a, F) > 0$ . This condition is used to obtain upper bounds for the probabilities  $P(\hat{F}_n \in \Omega_a)$  in Stone (1974) and also in GOR (1979). To show that the condition is not satisfied we suppose for simplicity that F is the uniform df on (0, 1). If  $\limsup_{t\downarrow 0} \psi(t) = \infty$ , there exists a decreasing sequence of points  $t_n \in (0, 1)$  such that  $\lim_{n\to\infty} t_n = 0$  and  $\lim_{n\to\infty} \psi(t_n) = \infty$ . Define the df  $G_n \in \Omega_a$  by the density

$$g_n(u) = \begin{cases} 1 + a/(t_n \psi(t_n)), & 0 < u \le t_n \\ 1 - a/\{(1 - t_n) \psi(t_n)\}, & t_n < u < 1 \\ 0, & u \notin (0, 1). \end{cases}$$

For each Borel measurable set B we have

$$\lim_{n\to\infty}\int_{R}dG_{n}(x)=\int_{R}1_{(0,1)}\ du,$$

and hence  $K_{\mathscr{P}}(\Omega_a, F) = 0$  for each partition  $\mathscr{P} = \{B_1, \dots, B_m\}$ . Note, however, that although the usual proofs fail, we still have (by Theorem 1)

$$\lim_{n\to\infty} n^{-1}\log P(\hat{F}_n\in\Omega_a) = -K(\Omega_a, F),$$

since  $K(\Omega_a, F) = g_{\psi}(a)$ .

PROOF OF THEOREM 2. We will show below that, for all  $m \ge \text{some } m_1$ , we have both

(a) 
$$K(\Omega_a, F) = K(\Omega_a \cap D_m, F)$$
 and

(b) 
$$K(\Omega_a \cap D_m, F) = K(c1_\tau(\Omega_a \cap D_m), F).$$

Then, using Lemmas 2.4 and 3.1 of GOR (1979), and (b) at inequality (d) we have

$$\limsup_{n\to\infty} n^{-1}\log P(T(\hat{F}_n) \geq a) = \limsup_{n\to\infty} n^{-1}\log P(\hat{F}_n \in \Omega_a)$$

(c) = 
$$\limsup_{n\to\infty} n^{-1}\log P(\hat{F}_n \in \Omega_a \cap D_m)$$
 for sufficiently large m

(d) 
$$\leq -K(\Omega_a \cap D_m, F)$$
  
=  $-K(\Omega_a, F)$  by (a):

equation (c) holds because Theorem 1 implies

(e) 
$$\limsup_{n\to\infty} n^{-1}\log P(\hat{F}_n \in D_m^c) \to -\infty$$
 as  $m\to\infty$ 

since  $g_{\psi_2}$  (a)  $\to \infty$  as  $a \to \infty$  (see Remark 2).

Since (a) is merely (1.24), it remains only to prove (b). To establish (b), it suffices to show that

(f) 
$$G \in c1_{\sigma}(\Omega_a \cap D_m)$$
 with  $K(G, F) \leq K(\Omega_a, F) + 1$ 

implies  $G \in \Omega_a \cap D_m$ . We suppose (f). We will first establish that  $G \in D_m$ . Now for some constants  $C_i$  we have

$$\begin{split} d_F(G,F) & \leq \sup_{F(x) \leq 1/2} G \log 2/F + \sup_{F(x) \geq 1/2} (1-G) \log 2/(1-F) \\ & + \sup_x 2F(1-F) \log 1/(F(1-F)) \\ & \leq \sup_x \left[ G \log 1/F + (1-G) \log 1/(1-F) \right] + C_1 \\ & \leq \sup_x \left[ G \log G/F + (1-G) \log (1-G)/(1-F) \right] + C_2 \\ & \leq K(G,F) + C_2 \\ & \leq K(\Omega_a,F) + C_3 \quad \text{by} \quad (f) \\ & \leq m \quad \text{for $m$ sufficiently large.} \end{split}$$

Thus  $G \in D_m$ . From the first half of (f) we thus have  $G \in c1_{\tau}(\Omega_a \cap D_m) \cap D_m$ . Since T is  $\tau$ -continuous on  $D_m$ ,  $\Omega_a \cap D_m$  is closed in the relative  $\tau$ -topology on  $D_m$ ; thus  $\Omega_a \cap D_m = c1_{\tau}(\Omega_a \cap D_m) \cap D_m$ . Thus  $G \in \Omega_a \cap D_m$  as we set out to show. Thus (b) holds. Using the  $\tau$ -continuity of T on  $D_m$ , one can prove (1.23) along the same lines as Theorem 3.2 in GOR (1979).

PROOF OF COROLLARY 2. It is shown in Example 2.1 that the function T defined by (2.1) is  $\tau$ -continuous on each subset  $D_m$ , with m > 0. Hence, according to Theorem 2, result (2.3) holds if the function  $a \to K(\Omega_a, F)$  is continuous from the right.

First of all,  $K(\Omega_a, F)$  is finite for all  $a \in (0, \infty)$  (and also trivially for  $a \le 0$ ). To see this, consider the df's  $G_x$ ,  $x \in (0, 1)$ , defined by

$$G_x(t) = x^{-1}t1_{[0,x)}(t) + 1_{[x,\infty)}(t).$$

Then

$$T(G_x) = \int_0^1 \frac{(G_x - t)^2}{t(1 - t)} dt = \left(\frac{1}{x} - 1\right)^2 \int_0^x \left(t/(1 - t)\right) dt + \int_x^1 \left((1 - t)/t\right) dt.$$

Hence  $T(G_x) \to \infty$ , as  $x \downarrow 0$ , but  $K(G_x, F) = x(1/x)\log(1/x) < \infty$ , all  $x \in (0, 1)$ . Fix a > 0. By (1.24), there exists an m > 0 such that

$$K(\Omega_a, F) = K(\Omega_a \cap D_m, F).$$

Since  $\Omega_a \cap D_m$  is  $\tau$ -closed, there exists by Lemma 3.2 in GOR (1979) a df  $G \in \Omega_a \cap D_m$  such that  $K(G, F) = K(\Omega_a, F)$ . Because  $K(G, F) < \infty$ , we have  $G \ll F$ . Hence G has a density g. Since  $G \in \Omega_a$  and a > 0, there exists  $x \in (0, 1)$  such that  $G(x) \neq x$ , say G(x) > x. Without loss of generality we may assume G(x) < 1. Let  $y \leq \inf\{t \geq x : G(t) \leq t\}$  have x < y < 1. Define the df  $G_\delta$  by the density

$$g_{\delta}(t) = \begin{cases} g(t) + \delta, t \in [x, x + \delta] \\ g(t) - (y - x - \delta)^{-1} \delta^{2}, t \in (x + \delta, y] \\ g(t), & \text{elsewhere.} \end{cases}$$

where  $\delta > 0$  is a suitably chosen small number. Then  $G_{\delta}(t) > G(t)$  for  $t \in (x, y)$  and  $G_{\delta}(t) = G(t)$ , elsewhere.

Hence  $T(G_{\delta}) \downarrow T(G)$ . It is clear that the Kullback-Liebler number  $K(G_{\delta}, F)$  tends to K(G, F) as  $\delta \downarrow 0$ . This implies that the function  $a \to K(\Omega_a, F)$  is continuous from the right. Thus, part (2.3) of the corollary follows.

By the continuity of T on  $D_m$  and Lemma 3.3 in GOR (1979), the function  $a \to K(\Omega_a, F)$  is also continuous from the left, implying (2.4).

Finally, to prove (2.5), we first observe that the df's  $\bar{G}_a$  defined by  $\bar{G}_a(t) = t + t(1 - t)\sqrt{6a}$ ,  $t \in [0, 1]$ ,  $6a \le 1$  have  $K(\bar{G}_a, F) \sim a$ , as  $a \downarrow 0$ . Hence

(a) 
$$K(\Omega_a, F) \le a + o(a)$$
, as  $a \downarrow 0$ 

Fix  $\epsilon > 0$ . We shall prove

(b) 
$$\lim \inf_{a \downarrow 0} a^{-1} K(\Omega_a, F) \ge 1 - \epsilon.$$

Since  $\epsilon$  is arbitrary, (2.5) will follow from (a) and (b).

Suppose that the collection of df's  $\{G_a\}$  satisfies  $K(G_a, F) = K(\Omega_a, F)$  and  $T(G_a) = a$  (by the first part of the proof such a collection exists). By (a) we have  $K(G_a, F) \le a + o(a)$ , as  $a \downarrow 0$ . Hence, by (1.26),

(c) 
$$d_F(G_a, F) \le k\sqrt{a},$$

uniformly in a, where k > 0 is some fixed number.

Let  $\eta \in (0, 1/2)$ , then  $\int_0^{\eta} (G_a(t) - t)^2 (t(1-t))^{-1} dt < k^2 a \int_0^{\eta} \{t(1-t)\log^2(t(1-t))\}^{-1} dt$ . Thus by choosing  $\eta > 0$  sufficiently small, we obtain  $\int_0^{\eta} (G_a(t) - t)^2 dt \le \epsilon a/2$ , uniformly in a. By the same argument, we may choose  $\eta > 0$  such that  $\int_{1-\eta}^1 (G_a(t) - t)^2 (t(1-t))^{-1} dt$  $\leq \epsilon a/2$ .

Define the df  $H_a$  by

$$H_a(t) = egin{cases} \eta^{-1}tG_a(\eta), & t \in [0,\,\eta) \ & G_a(t), & t \in [\eta,\,1-\eta] \ & G_a(1-\eta) + \eta^{-1}(t-1+\eta)(1-G_a(1-\eta)), \, t \in (1-\eta,\,1]. \end{cases}$$

Then  $K(H_a, F) \leq K(G_a, F)$  and  $H_a \in \Omega_{(1-\epsilon)a}$ . The df's  $H_a$  have the property

(d) 
$$\sup_{t\in(0,1)}\frac{|H_a(t)-t|}{t(1-t)}\leq M\sqrt{a}$$

where M is some fixed number independent of a, since  $|H_a(t) - t| = |G_a(t) - t| \le$  $K\sqrt{a/\psi(\eta)}$  by (c) if  $t \in [\eta, 1-\eta]$  and

$$|H_a(t) - t| < \frac{t(1-t)}{\eta(1-\eta)} \max\{|G_a(\eta) - \eta|, |G_a(1-\eta) - (1-\eta)|\}, \quad \text{if} \quad t \in [\eta, 1-\eta].$$

Using (d), we now prove that  $K(H_a, F) \ge (1 - \epsilon)a + o(a)$ , as  $a \downarrow 0$ . Define the function  $\phi(t, y, z)$  by

$$\phi(t, y, z) = (1+z)\log(1+z) - y^2/(t/(1-t))$$

where  $t \in (0, 1)$ ,  $y \in [-1, 1]$  and  $z \ge -1$ . We shall prove that

$$\int_0^1 \phi(t, H_a(t) - t, h_a(t) - 1) dt \ge o(a) \quad \text{as} \quad a \downarrow 0,$$

where  $h_a$  is the density of the df  $H_a$ . Let p(t, y) = y(1 - 2t)/(t(1 - t)),  $t \in (0, 1)$  and  $y \in$ [-1, 1], and let

(e) 
$$I(H_a) = \int_0^1 \{ \phi(t, y_a, p) - p \phi_z(t, y_a, p) + z_a p_z(t, y_a, p) \} dt$$

where

$$\phi_z(t, y, z) = \frac{\partial}{\partial z} \phi(t, y, z), y_a = y_a(t) = H_a(t) - t, z_a = z_a(t) = H_a(t) - 1$$

and

$$p=p(t,y_a(t)).$$

(It is true, but not needed in this proof, that the functions p(t, y) correspond to an "approximate" field of extremals for the function  $\phi(t, y, z)$ , since they correspond to approximate solutions  $f(t) = \alpha t(1-t)$  of the differential equation

$$y'' + \frac{2}{t(1-t)} \{y + yy'\} = 0,$$

which is the Euler-Lagrange equation for the minimization of the functional  $\int_0^1 \phi(t, y(t))$ y'(t)) dt.)

Expansion of  $log(1 + p(t, y_a(t)))$  and integration by parts twice yields

$$\int_{0}^{1} z_{a}(t)\phi_{z}(t, y_{a}, p) dt = \int_{0}^{1} z_{a}(t)\log(1 + p(t, y_{a}(t))) dt$$

$$= \int_0^1 z_a(t) y_a(t) \frac{1 - 2t}{t(1 - t)} dt - \frac{1}{2} \int_0^1 z_a(t) y_a^2(t) \left(\frac{1 - 2t}{t(1 - t)}\right)^2 dt + O(a^{3/2})$$

$$= \frac{1}{2} \int_0^1 p^2(t, y_a(t)) dt + \int_0^1 \frac{y_a^2}{t(1 - t)} dt + O(a^{3/2}), \quad \text{as} \quad a \downarrow 0.$$

Here we used (d). Moreover, again by expanding log(1 + p) it is seen that

$$\int_0^1 \left\{ \phi(t, y_a, p) - p \phi_z(t, y_a, p) \right\} dt = \int_0^1 \left[ \log(1 + p) - p - y_a^2 / (t(1 - t)) \right] dt$$

$$= -\frac{1}{2} \int_0^1 p^2(t, y_a(t)) dt - \int_0^1 \frac{y_a^2(t)}{t(1 - t)} dt + O(a^{3/2}).$$

Hence  $I(H_a) = O(a^{3/2})$ . By the mean value theorem we have

$$\phi(t, y_a, z_a) - \phi(t, y_a, p) - (z_a - p)\phi_z(t, y_a, p) = \frac{1}{2} (z_a - p)^2 \phi_{zz}(t, y_a, \theta),$$

where  $\theta$  is a point between  $z_a = z_a(t)$  and  $p = p(t, y_a(t))$ , if  $z_a \neq p$ , and may be taken arbitrarily (> -1) otherwise (as usual we denote  $(\partial/\partial z^2) \phi$  by  $\phi_{zz}$ ). Thus  $\int_0^1 \phi(t, y_a(t), z_a(t)) dt - I(H_a) \ge 0$ , implying  $\int_0^1 \phi(t, y_a(t), z_a(t)) dt \ge o(a)$ , as  $a \downarrow 0$ . Since

$$\int_{0}^{1} \frac{y_a^2(t)}{t(1-t)} dt = \int_{0}^{1} \frac{(H_a(t)-t)^2}{t(1-t)} dt \ge (1-\epsilon)a,$$

this implies  $K(H_a, F) \ge (1 - \epsilon)a + o(a)$ ,  $a \downarrow 0$ . Hence we may conclude  $\liminf_{a \downarrow 0} a^{-1}K(G_a, F) \ge \liminf_{a \downarrow 0} a^{-1}K(H_a, F) > 1 - \epsilon$ , which we set out to prove.

REMARK 10. In the last part of the proof of Corollary 2, it was actually shown that Weierstrass' sufficiency conditions for a strong minimum of the functional  $\int_0^1 \phi(t, G(t) - t, g(t) - 1) dt$ , g the density of G, are "approximately" satisfied for the df's  $\bar{G}_a(t) = t + t(1 - t)\sqrt{6a}$ . Note that we only have to consider values of the function  $\phi(t, y, z)$  with  $z \ge -1$ , since  $g(t) - 1 \ge -1$  for each density g. For Weierstrass' sufficiency conditions, see e.g. Gelfand and Fomin (1963), page 148.

PROOF OF COROLLARY 3. Let the sets  $D_m$  be defined as in Theorem 2. We first show that the integrals

$$\int_0^1 J(u)G^{-1}(u) \ du \quad \text{are well-defined for} \quad G \in D_m, m > 0.$$

Fix m > 0,  $G \in D_m$  and let  $\psi = \psi_2$  (see (1.6)). Then

$$\sup_{t \in (0,1)} (t - G(F^{-1}(t))) \psi(t) < m.$$

There exists  $\delta > 0$  such that  $t - m/\psi(t) \ge 1 - 2m/\psi(t)$ , if  $t \in [1 - \delta, 1)$ . Hence, for these values of t

$$G^{-1}(1-2m/\psi(t)) \le G^{-1}\left(t-\frac{m}{\psi(t)}\right) \le F^{-1}(t).$$

Putting  $u = u(t) = 1 - 2m/\psi(t)$  (so  $1 - t \ge t(1 - t) = \exp(-2m/(1 - u))$ , we obtain for all u(t) such that  $t \in [1 - \delta, 1)$  (this is, for all u in some interval  $[1 - \delta', 1)$ )

(a) 
$$G^{-1}(u) \le F^{-1}(1 - \exp(-2m/(1 - u))).$$

Likewise there exists  $\delta'' > 0$  such that

(b) 
$$G^{-1}(u) \ge F^{-1}(\exp(-2m/u))$$

for all  $u \in (0, \delta'']$ .

Hence, by (a) and (b) and conditions (i) and (ii):

$$\int_0^1 |J(u)G^{-1}(u)| \ du < \infty.$$

Moreover, because of a common dominating function, the above argument shows that the integrals of  $JG^{-1}$  over  $(0, \eta]$  and  $[1 - \eta, 1)$  tend to zero as  $\eta \downarrow 0$ , uniformly in  $G \in D_m$ .

For fixed  $\eta \in (0, 1/2)$  the function  $T_{\eta}: D \to \mathbb{R}$  defined by

$$T_{\eta}(G)=\int_{-\pi}^{1-\eta}J(u)G^{-1}(u)\ du,\qquad G\in D,$$

is  $\tau$ -continuous on D (cf. Section 6 of GOR (1979)). This fact, together with the uniform integrability of  $J(u)G^{-1}(u)$  for  $G \in D_m$ , implies that the function  $T:D \to \mathbb{R}$  defined by

$$T(G) = \left\{egin{array}{ll} \int_0^1 J(u) \, G^{-1}(u) \, \, du, & ext{if} \quad G \in D_m & ext{for some} \quad m \geq 0 \ \ 0, \, ext{otherwise} \end{array}
ight.$$

is  $\tau$ -continuous on each set  $D_m$ .

Thus (1.22) of Theorem 2 holds, if  $K(\Omega_r, F) < \infty$ . It remains to show in this case only that  $t \to K(\Omega_t, F)$  is a continuous mapping. If  $K(\Omega_r, F) < \infty$ , then the continuity from the right at t = r can be proved by the argument of "displacement of mass" used in GOR (1979), page 579. Since, if  $K(\Omega_r, F) < \infty$ , we have by (1.24) that  $K(\Omega_r, F) = K(\Omega_r \cap D_m, F)$  for m sufficiently large, and moreover T is  $\tau$ -continuous on  $D_m$ , the left continuity of the mapping  $t \to K(\Omega_t, F)$  in t = r follows from Lemma 3.3 in GOR (1979). If  $K(\Omega_r, F) = \infty$ , then (3.2) must still be established. In this case this mapping is certainly continuous from the right at t = r (by monotonicity). It is continuous from the left since trivially  $K(\Omega_r \cap D_m, F) = K(\Omega_r, F)$  for each m > 0, and then apply Lemma 3.3 of GOR (1979) to get left continuity of  $K(\Omega_r \cap D_m, F)$ .

Let  $r_0 = \inf\{t: K(\Omega_t, F) = \infty\}$  (possibly  $r_0 = \infty$ ). Then  $\lim_{t \uparrow r_0} K(\Omega_t, F) = \infty$  and hence, by the last paragraph, we have

$$\begin{split} \lim\sup_{n\to\infty} n^{-1}\!\log P(T(\hat{F}_n) \geq r) & \leq \lim_{t\uparrow r_0} \lim\sup_{n\to\infty} n^{-1}\!\log P(T(\hat{F}_n) \geq t) \\ & \leq \lim_{t\uparrow r_0} \left( -K(\Omega_t, F) \right) = -\infty \end{split}$$

using the case already proved. This completes the proof.

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