

CROSSINGS AND OCCUPATION MEASURES
 FOR A CLASS OF SEMIMARTINGALES¹

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We show that

$$\frac{1}{\sqrt{\varepsilon}} \left\{ \int_{-\infty}^{\infty} f(u) k_{\varepsilon} N_{\tau}^{X_{\varepsilon}}(u) du - \int_0^{\tau} f(X_t) a_t dt \right\}$$

converges in law (as a continuous process) to $c_{\psi} \int_0^{\tau} f(X_t) a_t dB_t$, where $X_t = \int_0^t a_s dW_s + \int_0^t b_s ds$, with W a standard Brownian motion, a and b regular and adapted processes, $X_{\varepsilon}(t) = \int_{-\infty}^{\infty} (1/\varepsilon) \psi((t-u)/\varepsilon) X_u du$, ψ a smooth kernel, $N_{\tau}^g(u)$ the number of roots of the equation $g(s) = u$, $s \in (0, t]$, $k_{\varepsilon} = \sqrt{\pi\varepsilon/2}/\|\psi\|_2$, f a smooth function, B a standard Brownian motion independent of W and c_{ψ} a constant depending only on ψ .

1. Introduction. Let $X = \{X_t: t \geq 0\}$ be a real-valued continuous semimartingale of the form

$$(1) \quad X_t = \int_0^t a_s dW_s + \int_0^t b_s ds,$$

where $W = \{W_t: t \geq 0\}$ is a standard Brownian motion (BM for short) adapted to a filtration $F = \{F_t: t \geq 0\}$, where $F_t \perp \sigma\{W_r - W_s: t \leq s \leq r\} \forall t \geq 0$. Here $a = \{a_t: t \geq 0\}$ and $b = \{b_t: t \geq 0\}$ are F -adapted processes verifying a certain number of regularity and boundedness conditions to be precised later on. We shall also assume that $a_t > 0$.

The purpose of this paper is to compute the speed of convergence of the normalized number of crossings of regularizations of X to the local time of X .

More precisely, let ψ be a C^{∞} kernel, $\psi: \mathbb{R} \rightarrow \mathbb{R}^+$, with compact support (say $\text{supp } \psi \subset [-1, 1]$), and $\int_{-1}^1 \psi(u) du = 1$.

Define

$$(2) \quad X_{\varepsilon}(t) = \int_{-\infty}^{\infty} \frac{1}{\varepsilon} \psi\left(\frac{t-u}{\varepsilon}\right) X_u du = \int_{-1}^1 \psi(-u) X_{t+\varepsilon u} du,$$

where we have extended X by means of $X_t = 0$ if $t < 0$.

A comment on notation: the symbol " $\xrightarrow{\varepsilon \rightarrow 0^+}$ " denotes convergence of real numbers in the ordinary sense, and " $\xrightarrow[\varepsilon \rightarrow 0^+]{w}$ " indicates weak convergence of processes or measures.

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In Azais and Wschebor (1995) it is proved that if M is a real-valued and continuous local martingale with bracket $A = \{A_t: t \geq 0\}$ then, almost surely, for any bounded interval I contained in $[0, \infty)$ and any continuous function $f: \mathbb{R} \rightarrow \mathbb{R}$, we have

$$(3) \quad k_\varepsilon \int_{-\infty}^{\infty} f(u) N_I^{M_\varepsilon}(u) du \xrightarrow{\varepsilon \rightarrow 0^+} \int_I f(M_t) (\dot{A}_t)^{1/2} dt;$$

$N_I^g(u)$ is the number of roots of the equation $g(t) = u$, $t \in I$; $\dot{A} = \{\dot{A}_t: t \geq 0\}$ is the (almost everywhere) derivative of A , and

$$k = \frac{1}{\|\psi\|_2} \sqrt{\frac{\pi}{2}}, \quad k_\varepsilon = k\sqrt{\varepsilon}, \quad \|\psi\|_2 = \left(\int_{-1}^1 \psi(u)^2 du \right)^{1/2}.$$

We will also denote $N_I^g(u) = N_{(0, t]}^g(u)$.

Theorem 1 gives a speed of convergence in (3) for semimartingales of the form (1). Note that in the statement of Theorem 1 neither the centering nor the limit distribution depend on the drift term in (1). The constant c_ψ depends only on the regularizing kernel ψ and not on the process.

Theorem 1 can be used to make inference on the martingale part of X . It also allows measuring the local time of X from the observation of the number of crossings of X_ε . In fact, introduce the modified local time:

$$\hat{L}_I^X(u) = \int_I \frac{1}{a_t} L_{dt}^X(u),$$

where $L_J^X(u)$ is the value at $u \in \mathbb{R}$ of the canonical bicontinuous local time of the continuous martingale M on the interval J [see Revuz and Yor (1991), page 209, (1.6)].

Taking into account that in this case $\dot{A}_t^{1/2} = a_t$, we can rewrite the right-hand term of (3) as

$$\int_I f(X_t) a_t dt = \int_{-\infty}^{\infty} f(u) \hat{L}_I^X(u) du.$$

To see this, argue as follows:

$$\int_I f(X_t) a_t dt = \int_I f(X_t) a_t^2 \frac{1}{a_t} dt.$$

If f is nonnegative, consider the measure defined by

$$\nu(J) = \int_{-\infty}^{\infty} f(u) L_J^X(u) du = \int_J f(X_t) a_t^2 dt;$$

then

$$\begin{aligned} \int_I f(X_t) a_t dt &= \int_I \frac{1}{a_t} \nu(dt) = \int_I \frac{1}{a_t} \int_{-\infty}^{\infty} f(u) L_{dt}^X(u) du \\ &= \int_{-\infty}^{\infty} f(u) \left(\int_I \frac{1}{a_t} L_{dt}^X(u) \right) du = \int_{-\infty}^{\infty} f(u) \hat{L}_I^X(u) du. \end{aligned}$$

Then (3) means that, almost surely,

$$(4) \quad k_\varepsilon N_I^{X_\varepsilon}(u) du \xrightarrow{\varepsilon \rightarrow 0^+} \hat{L}_I^X(u) du, \text{ where convergence takes place as weak convergence of measures on } \mathbb{R}.$$

Theorem 1 refers to the speed of convergence in (4) enabling measuring the discrepancy between the approximation and its limit.

Extensions of Theorem 1 to \mathbb{R}^d -valued semimartingales will be considered elsewhere. In fact, the general setting consists of the study of the asymptotic behavior (as ε goes to zero) of functionals defined on the smoothed paths $X_\varepsilon(\cdot)$ having the form

$$\int_I F(X_\varepsilon(t), \sqrt{\varepsilon} \dot{X}_\varepsilon(t)) dt$$

for suitable choices of the function F . Theorem 1 corresponds to $d = 1$ and $F(x, y) = f(x)|y|$.

In the case X is BM, a proof of Theorem 1, based on convergence of moments, has been given in Berzin and León (1994).

2. Main results and examples. In what follows, a and b will be \mathbb{R} -valued adapted and continuous processes, with $a > 0$, and such that the following hold:

- (A) For every $T, p > 0, \sup_{t \in [0, T]} E\{|b_t|^p\} < \infty$.
- (B) $\forall \varepsilon > 0, (a_{s+\varepsilon} - a_s)/\sqrt{\varepsilon} = \alpha_s^* Z_{s, \varepsilon} + r_{s, \varepsilon}$, where we have the following:
 - (i) α^* adapted, $Z_{\bullet, \varepsilon}$ and $r_{\bullet, \varepsilon}$ $F_{\bullet+\varepsilon}$ -predictable, $Z_{s, \varepsilon} \perp F_s$.
 - (ii) For almost every pair $s, t \geq 0, t \neq s$, we have the following weak convergence (in $\mathbb{R}^2 \times \mathbb{C} ([0, \infty))^2$):

$$(Z_{t, \varepsilon}, Z_{s, \varepsilon}, W_{\bullet}^{\varepsilon, t}, W_{\bullet}^{\varepsilon, s}) \xrightarrow[\varepsilon \rightarrow 0]{w} (Z_t, Z_s, W_{\bullet}^t, W_{\bullet}^s),$$

where

$$W_{\gamma}^{\varepsilon, t} = \frac{W_{t+\varepsilon\gamma} - W_t}{\sqrt{\varepsilon}};$$

$\{W_{\bullet}^t: t \geq 0\}$ is a collection of independent Brownian motions, $V_{\bullet}(t, s) = (Z_t, Z_s, W_{\bullet}^t, W_{\bullet}^s) \perp F_{\infty}$; $V_{\bullet}(t, s)$ has a symmetric distribution [i.e., $V_{\bullet}(t, s)$ and $-V_{\bullet}(t, s)$ have the same distribution] and if $\{s, t\}$ and $\{s', t'\}$ are disjoint, $V_{\bullet}(s, t) \perp V_{\bullet}(s', t')$.

- (iii) For every $p > 0, T > 0$, and some $\delta > 0$,

$$\begin{aligned} & \sup_{s \in [0, T], \varepsilon \in [0, \delta]} E(|Z_{s, \varepsilon}|^p) < \infty, \\ & \sup_{s \in [0, T]} E(|r_{s, \varepsilon}|^p) \xrightarrow{\varepsilon \rightarrow 0^+} 0, \\ & \sup_{s \in [0, T]} E(|\alpha_s^*|^p) < \infty. \end{aligned}$$

We will consider Brownian semimartingales defined by (1) with a and b as before. In addition, we set

$$(5) \quad \Delta_\varepsilon(t) = X_\varepsilon(t) - X_t,$$

$$(6) \quad X_u^{\varepsilon, t} = \frac{X_{t+\varepsilon u} - X_t}{\sqrt{\varepsilon}}.$$

Observe that if $\text{supp } \psi \subset [-1, 0]$, $X_u^{\varepsilon, t} = \int_0^u a_{t+\varepsilon v} dW_v^{\varepsilon, t} + \sqrt{\varepsilon} \int_0^u b_{t+\varepsilon v} dv$. Hence, $X_\bullet^{\varepsilon, t}$ is the solution of the SDE:

$$(7) \quad d_u X_u^{\varepsilon, t} = a_{t+\varepsilon u} d_u W_u^{\varepsilon, t} + \sqrt{\varepsilon} b_{t+\varepsilon u} du, \quad u \geq 0, \quad X_0^{\varepsilon, t} = 0.$$

Let us denote by C_b^2 the set of real-valued functions with bounded continuous second derivative and set

$$E_\varepsilon(\tau) := \frac{1}{\sqrt{\varepsilon}} \left\{ \int_{-\infty}^{\infty} f(u) k_\varepsilon N_\tau^{X_\varepsilon}(u) du - \int_0^\tau f(X_t) a_t dt \right\}$$

Our main result is the following theorem.

THEOREM 1. *If X is as in (1), $f \in C_b^2$ then*

$$(W_\tau, E_\varepsilon(\tau)) \xrightarrow[\varepsilon \rightarrow 0]{w} (W_\tau, c_\psi \int_0^\tau f(X_t) a_t dB_t) \quad \text{in } \mathbb{C}([0, \infty))^2,$$

where B is a BM, $B \perp W$, and c_ψ is the constant

$$c_\psi^2 = \int_{-1}^1 \int_{-1}^1 E \left\{ \prod_{i=1}^{i=2} H(R_{\gamma_i}, \beta^2(\gamma_i)) \psi(-\gamma_i) \right\} d\gamma_2 d\gamma_1,$$

where $H(x, \theta) := k[2\Phi(x/\theta) - 1]$, Φ is the standard normal distribution, $R_\gamma := \int_0^\gamma \psi(-u) dW_u$ and $\beta^2(\gamma) := \int_\gamma^1 \psi^2(-u) du$.

EXAMPLE 1 (Diffusions). Consider the diffusion process

$$(8) \quad X_t = \int_0^t \sigma(X_s) dW_s + \int_0^t b(X_s) ds,$$

where $\sigma > 0$ and assume

$$(9) \quad \sigma(x)^2 + b(x)^2 \leq K(1 + x^2) \quad \forall x \in \mathbb{R},$$

$$(10) \quad |b(x) - b(y)| + |\sigma(x) - \sigma(y)| \leq L_N |x - y| \\ \forall |x|, |y| \leq N, \quad \forall N \in \mathbb{N},$$

in which case existence and uniqueness of a strong solution of (8) are guaranteed, with all its moments uniformly bounded over compact intervals. Furthermore, assume that σ belongs to C_b^2 .

Denote $a_s = \sigma(X_s)$, $b_s = b(X_s)$. We have ($0 < \theta < 1$)

$$\begin{aligned} \frac{a_{s+\varepsilon} - a_s}{\sqrt{\varepsilon}} &= \sigma'(X_s) \frac{(X_{s+\varepsilon} - X_s)}{\sqrt{\varepsilon}} + \sigma''(X_{s+\theta\varepsilon}) \frac{(X_{s+\varepsilon} - X_s)^2}{2\sqrt{\varepsilon}} \\ &= \sigma'(X_s)\sigma(X_s)W_1^{\varepsilon,s} + \sigma'(X_s) \int_0^1 [\sigma(X_{s+\varepsilon v}) - \sigma(X_s)] d_v W_v^{\varepsilon,s} \\ &\quad + \sigma'(X_s)\sqrt{\varepsilon} \int_0^1 b(X_{s+\varepsilon v}) d_v W_v^{\varepsilon,s} + \sigma''(X_{s+\theta\varepsilon}) \frac{(X_{s+\varepsilon} - X_s)^2}{2\sqrt{\varepsilon}}. \end{aligned}$$

Thus, we have the representation (B) for a , with

$$\begin{aligned} a_s^* &= \sigma'(X_s)\sigma(X_s), \\ Z_{s,\varepsilon} &= W_1^{\varepsilon,s}, \\ r_{s,\varepsilon} &= \sigma'(X_s) \int_0^1 [\sigma(X_{s+\varepsilon v}) - \sigma(X_s)] d_v W_v^{\varepsilon,s} \\ &\quad + \sigma'(X_s)\sqrt{\varepsilon} \int_0^1 b(X_{s+\varepsilon v}) d_v W_v^{\varepsilon,s} \\ &\quad + \sigma''(X_{s+\theta\varepsilon}) \frac{(X_{s+\varepsilon} - X_s)^2}{2\sqrt{\varepsilon}}, \end{aligned}$$

which clearly satisfy all the required conditions.

The statement of Theorem 1 can then be rewritten as

$$\left(W_\tau, \frac{1}{\sqrt{\varepsilon}} \int_{-\infty}^{\infty} f(u) \left[k_\varepsilon N_\tau^{X_\varepsilon}(u) - \frac{L_\tau^X(u)}{\sigma(u)} \right] du \right) \xrightarrow[\varepsilon \rightarrow 0]{w} \left(W_\tau, c_\psi \int_0^\tau f(X_t) \sigma(X_t) dB_t \right)$$

[in $\mathbb{C}([0, \infty)^2)$].

EXAMPLE 2 (Non-Markovian martingales). Consider $X_t = \int_0^t f(W_s) dW_s$, where $f: \mathbb{R} \rightarrow \mathbb{R}$ is a C^3 function such that $\underline{f} = \inf\{f(x): x \in \mathbb{R}\} > 0$, f'' and $f^{(3)}$ are bounded, $f''(0) \neq 0$ and $\|f''\|_\infty < 2\underline{f}$ [e.g., $f(x) = 1 + \mathbb{C} \exp(-x^2)$, with $0 < \mathbb{C} < 1$].

Then X verifies the hypothesis of Theorem 1 with $a_s = f(W_s)$, $b_s \equiv 0$, $a_s^* = f'(W_s)$, $Z_{\varepsilon,s} = W_1^{\varepsilon,s}$. However, X is non-Markovian; hence it is not a diffusion [cf. Nualart and Wschebor (1991), page 106].

EXAMPLE 3 (Smoother integrands). Suppose a satisfies a Hölder condition of the form

$$\sup_{0 \leq t \leq T-\varepsilon} |a_{t+\varepsilon}(\omega) - a_t(\omega)| \leq C_T(\omega) \varepsilon^\alpha \quad (\alpha > \frac{1}{2})$$

for each $T > 0$ and $C_T \in L^p$ for all $p > 0$. Then the process X is included in our framework with $a_t^* = 0$, $Z_{s,\varepsilon} = 0$.

3. Proof of the main result. With no loss of generality, we will restrict the parameter to vary in $[0, 1]$. We also may suppose $\text{supp } \psi \subset [-1, 0]$ (see Proof of Step 1) and a localization argument implies that we can assume a and b uniformly bounded by a (nonrandom) constant and a bounded away from zero (see Lemma 1).

Throughout the proof, \mathbb{C} will denote a generic positive constant that may change from line to line. We divide the proof into several steps, and include further a series of auxiliary lemmas.

STEP 1. Denote $Z_\varepsilon(\tau) = (1/\sqrt{\varepsilon}) \int_0^\tau f(X_t) g^t(\sqrt{\varepsilon} \dot{X}_\varepsilon(t)) dt$, where $g^t(x) := k|x| - a_t$; then we have the following:

- (i) $E_\varepsilon(\tau) - Z_\varepsilon(\tau) \xrightarrow{\varepsilon \rightarrow 0^+} 0$ (in L^2);
- (ii) $\{E_\varepsilon - Z_\varepsilon; \varepsilon > 0\}$ is $\mathbb{C}([0, 1])$ -tight.

Hence, E_ε has the same asymptotic distribution as Z_ε .

REMARK 1. It follows from the proof that

$$E \left[\sup_{0 \leq \tau \leq 1} |E_\varepsilon(\tau) - Z_\varepsilon(\tau)| \right] \xrightarrow{\varepsilon \rightarrow 0^+} 0.$$

STEP 2. We can decompose

$$g^t(\sqrt{\varepsilon} \dot{X}_\varepsilon(t)) = a_t \int_0^1 \Phi_{\varepsilon, t}(v) d_v W_v^{\varepsilon, t} + R_\varepsilon(t),$$

where

$$\mathbb{R}_\varepsilon(\tau) := \frac{1}{\sqrt{\varepsilon}} \int_0^\tau f(X_t) R_\varepsilon(t) dt \xrightarrow[\varepsilon \rightarrow 0]{w} 0 \quad \text{in } \mathbb{C}([0, 1]);$$

$$\Phi_{\varepsilon, t}(v) := k \left[2\Phi \left(\frac{Y_v^{\varepsilon, t}}{\beta(v)a_t} \right) - 1 \right];$$

$$Y_v^{\varepsilon, t} := \int_0^v \psi(-u) d_u X_u^{\varepsilon, t}.$$

REMARK 2. As in Step 1, we obtain $E[\sup_{0 \leq \tau \leq 1} |\mathbb{R}_\varepsilon(\tau)|] \xrightarrow{\varepsilon \rightarrow 0^+} 0$.

STEP 3. We can decompose: $Z_\varepsilon(\tau) = \int_0^\tau f(X_t) K_\varepsilon(t) a_t dW_t + \alpha_\varepsilon(\tau)$, where

$$K_\varepsilon(t) := \int_{\max(0, t-\varepsilon)}^t \frac{\Phi_{\varepsilon, v}((t-v)/\varepsilon)}{\varepsilon} dv; \quad \alpha_\varepsilon \xrightarrow[\varepsilon \rightarrow 0]{w} 0 \quad \text{in } \mathbb{C}([0, 1]).$$

STEP 4. Let $V = \{V_t; t \geq 0\}$ be an adapted process such that

$$\sup_{s \in [0, 1]} E(|V_s|^p) < \infty \quad \forall p > 0.$$

Then if $V_\varepsilon^*(\tau) := \int_0^\tau V_t K_\varepsilon^2(t) dt$, $\hat{V}_\varepsilon(\tau) := \int_0^\tau V_t K_\varepsilon(t) dt$, we have

$$V_\varepsilon^* \xrightarrow[\varepsilon \rightarrow 0]{w} c_\psi \int_0^\tau V_t dt \quad \text{in } \mathbb{C}([0, 1]); \quad \hat{V}_\varepsilon \xrightarrow[\varepsilon \rightarrow 0]{w} 0 \quad \text{in } \mathbb{C}([0, 1]).$$

STEP 5. If $S_\varepsilon(\tau) := \int_0^\tau K_\varepsilon(t) dW_t$, then $(W, S_\varepsilon) \xrightarrow[\varepsilon \rightarrow 0]{w} (W, c_\psi B)$ [in $\mathbb{C}([0, 1])^2$], where B is a BM, $B \perp W$.

STEP 6. If $\hat{Z}_\varepsilon(\tau) := \int_0^\tau f(X_t) a_t dS_\varepsilon(t)$, then $(W_\bullet, \hat{Z}_\varepsilon(\cdot)) \xrightarrow[\varepsilon \rightarrow 0]{w} (W_\bullet, c_\psi B_{\theta(\cdot)})$ [in $\mathbb{C}([0, 1])^2$], with $\theta(\tau) := \int_0^\tau f(X_t)^2 a_t^2 dt$.

Hence, from Step 3, $(W_\bullet, Z_\varepsilon(\cdot)) \xrightarrow[\varepsilon \rightarrow 0]{w} (W_\bullet, c_\psi B_{\theta(\cdot)})$ [in $\mathbb{C}([0, 1])^2$], with $\theta(\tau) = \int_0^\tau f(X_t)^2 a_t^2 dt$.

The theorem follows from Steps 1 and 6, which we will prove, with the help of some auxiliary lemmas presented in Section 4.

PROOF OF STEP 1. The formula $\int_{-\infty}^\infty u(x) N_I^v(x) dx = \int_I u(v(t)) |v'(t)| dt$ can be easily checked for $u: \mathbb{R} \rightarrow \mathbb{R}$ continuous and $v: I \rightarrow \mathbb{R}$ of class C^1 , and I a bounded interval in the line [see, e.g., Nualart and Wschebor (1991), page 88, (2.4)].

Hence,

$$\begin{aligned} E_\varepsilon(\tau) &= \frac{1}{\sqrt{\varepsilon}} \left[\int_0^\tau f(X_\varepsilon(t)) \sqrt{\frac{\pi}{2}} \frac{|\sqrt{\varepsilon} \dot{X}_\varepsilon(t)|}{\|\psi\|_2} dt - \int_0^\tau f(X_t) a_t dt \right] \\ &= Z_\varepsilon(\tau) + k \int_0^\tau [f(X_\varepsilon(t)) - f(X_t)] |\dot{X}_\varepsilon(t)| dt. \end{aligned}$$

Applying Lemma 3c, we deduce that (i) and (ii) hold, which concludes the proof of this step.

PROOF OF STEP 2. First we will prove tightness. Set

$$G^t(x, \theta) = E\{g^t(x + \sqrt{\theta} a_t N) / F_t\} = k \int_{-\infty}^\infty |x + \sqrt{\theta} a_t s| \phi(s) ds - a_t,$$

where N is a standard normal random variable, $N \perp F_\infty$, $x \in \mathbb{R}$, $\theta > 0$ and ϕ stands for the standard normal density.

G^t is the $C^\infty(\mathbb{R} \times (0, \infty))$ solution of

$$(11) \quad D_\theta G^t = \frac{a_t^2}{2} D_{xx}^2 G^t; \quad G^t(x, 0^+) = g^t(x).$$

Denoting by Φ , the standard normal distribution, we have

$$(12) \quad D_x G^t(x, \theta) = k \left[2\Phi\left(\frac{x}{\sqrt{\theta} a_t}\right) - 1 \right],$$

$$(13) \quad D_{xx}^2 G^t(x, \theta) = \frac{2k}{\sqrt{\theta}} \phi\left(\frac{x}{\sqrt{\theta} a_t}\right).$$

Note that $D_x G^t(x, \theta)$ and $\sqrt{\theta} D_{xx}^2 G^t(x, \theta)$ are continuous and bounded. In addition, $H(x, \theta) := D_x^t(a_t x, \theta)$ and $J(x, \theta) := a_t D_{xx}^2 G^t(a_t x, \theta)$ do not depend on t , $H(\cdot, \theta)$ is odd and $J(\cdot, \theta)$ is even.

Define $Y_\gamma^{\varepsilon, t} = \int_0^\gamma \psi(-u) d_u X_u^{\varepsilon, t}$. Applying Itô's formula to

$$\eta_\gamma^{\varepsilon, t} = G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)),$$

and noting that $\sqrt{\varepsilon} \dot{X}_\varepsilon(t) = Y_1^{\varepsilon, t}$, we get

$$\begin{aligned} g^t(\sqrt{\varepsilon} \dot{X}_\varepsilon(t)) &= \eta_1^{\varepsilon, t} - \eta_0^{\varepsilon, t} \\ &= \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) d_\gamma Y_\gamma^{\varepsilon, t} \\ (14) \quad &+ \int_0^1 D_\theta G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \left(\frac{d\beta^2(\gamma)}{d\gamma} \right) d\gamma \\ &+ \frac{1}{2} \int_0^1 D_{xx}^2 G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) d_\gamma \langle Y^{\varepsilon, t}, Y^{\varepsilon, t} \rangle_\gamma. \end{aligned}$$

Using (7), (11) and (14) we obtain

$$\begin{aligned} g^t(\sqrt{\varepsilon} \dot{X}_\varepsilon(t)) &= \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) a_t d_\gamma W_\gamma^{\varepsilon, t} \\ &+ \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) [a_{t+\varepsilon\gamma} - a_t] d_\gamma W_\gamma^{\varepsilon, t} \\ &+ \frac{1}{2} \int_0^1 D_{xx}^2 G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi^2(-\gamma) [a_{t+\varepsilon\gamma} - a_t] [a_{t+\varepsilon\gamma} + a_t] d\gamma \\ &+ \sqrt{\varepsilon} \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) b_{t+\varepsilon\gamma} d\gamma \\ &= a_t \int_0^1 \Phi_{\varepsilon, t}(\gamma) d_\gamma W_\gamma^{\varepsilon, t} + R_\varepsilon(t), \end{aligned}$$

with

$$(15) \quad \Phi_{\varepsilon, t}(\gamma) = D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma);$$

$$\begin{aligned} R_\varepsilon(t) &= \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) [a_{t+\varepsilon\gamma} - a_t] d_\gamma W_\gamma^{\varepsilon, t} \\ (16) \quad &+ \frac{1}{2} \int_0^1 D_{xx}^2 G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi^2(-\gamma) [a_{t+\varepsilon\gamma} - a_t] [a_{t+\varepsilon\gamma} + a_t] d\gamma \\ &+ \sqrt{\varepsilon} \int_0^1 D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) b_{t+\varepsilon\gamma} d\gamma. \end{aligned}$$

We have

$$(17) \quad \mathbb{R}_\varepsilon(\tau) = A_\varepsilon^1(\tau) + A_\varepsilon^2(\tau) + A_\varepsilon^3(\tau)$$

where

$$(18) \quad A_\varepsilon^1(\tau) = \int_0^\tau \int_0^1 f(X_t) D_x G^t(Y_\gamma^{\varepsilon, t}, \beta^2(\gamma)) \psi(-\gamma) \frac{[a_{t+\varepsilon\gamma} - a_t]}{\sqrt{\varepsilon}} d_\gamma W_\gamma^{\varepsilon, t} dt;$$

$$(19) \quad A_\varepsilon^2(\tau) = \frac{1}{2} \int_0^\tau \int_0^1 f(X_t) D_{xx}^2 G^t(Y_\gamma^{\varepsilon,t}, \beta^2(\gamma)) \psi^2(-\gamma) \\ \times \frac{[a_{t+\varepsilon\gamma} - a_t]}{\sqrt{\varepsilon}} [a_{t+\varepsilon\gamma} + a_t] d\gamma dt;$$

$$(20) \quad A_\varepsilon^3(\tau) = \int_0^\tau \int_0^1 f(X_t) D_x G^t(Y_\gamma^{\varepsilon,t}, \beta^2(\gamma)) \psi(-\gamma) b_{t+\varepsilon\gamma} d\gamma dt.$$

We will prove in what follows that $\{A_\varepsilon^i: \varepsilon > 0\}$ is $\mathbb{C}([0, 1])$ -tight for $i = 1, 2, 3$.

Using Hölder's inequality and Lemma 2, we see that the integrand in (18) is bounded in L^p for each $p > 0$; applying Jensen and the Burkholder–Davis–Gundy inequality [cf. Revuz and Yor (1991), page 152] we obtain

$$(21) \quad E\{|A_\varepsilon^1(\tau') - A_\varepsilon^1(\tau)|^4\} \leq \mathbb{C} |\tau' - \tau|^3,$$

which proves tightness for $i = 1$. The case $i = 3$ is even easier. For the case $i = 2$, it suffices to remark that $D_{xx}^2(x, \beta^2(\gamma)) \psi^2(-\gamma)$ is bounded by $\mathbb{C}(\psi^2(-\gamma)/\beta(\gamma))$ and that

$$\int_0^1 \frac{\psi^2(-\gamma)}{\beta(\gamma)} d\gamma = \int_0^{\|\psi\|_2^2} \frac{1}{\sqrt{u}} du < \infty$$

and tightness follows.

For the convergence to zero in L^2 , note that

$$A_\varepsilon^1(\tau) = \int_0^\tau \int_t^{t+\varepsilon} f(X_t) D_x G^t\left(Y_{((v-t)/\varepsilon)}^{\varepsilon,t}, \beta^2\left(\left(\frac{v-t}{\varepsilon}\right)\right)\right) \\ \times \psi\left(-\left(\frac{v-t}{\varepsilon}\right)\right) \frac{[a_v - a_t]}{\varepsilon} dW_v dt \\ = \int_0^\tau I_\varepsilon^1(t) dt.$$

Because of the martingale property of the stochastic integral it is clear that $I_\varepsilon^1(t)$, $I_\varepsilon^1(s)$ are uncorrelated for $|t - s| > \varepsilon$ and it follows that $E\{|A_\varepsilon^1(\tau)|^2\} = O(\varepsilon)$.

Now

$$(22) \quad Y_\gamma^{\varepsilon,t} = a_t R_\gamma^{\varepsilon,t} + \sqrt{\varepsilon} O_{L^p}(1) \quad \forall p > 0,$$

where $T(\varepsilon, t) = O_{L^p}$ means $\sup_{t \in [0,1], \varepsilon > 0} E\{|T(\varepsilon, t)|^p\} < \infty$, and $R_\gamma^{\varepsilon,t} = \int_0^\gamma \psi(-u) d_u W_u^{\varepsilon,t}$.

(23) For almost every pair $s, t > 0$, $s \neq t$, $(R_\gamma^{s,s}, Z_{s,\varepsilon}, R_\gamma^{\varepsilon,t}, Z_{t,\varepsilon}) \xrightarrow[\varepsilon \rightarrow 0]{w} (R_\gamma^s, Z_s, R_\gamma^t, Z_t)$, where $\{R_\bullet^s: s \in [0, 1]\}$ are independent copies of R_\bullet (defined in the statement of Theorem 1), $(R_\bullet^s, Z_s, R_\bullet^t, Z_t) \perp W$, $(R_\bullet^s, Z_s) \perp (R_\bullet^t, Z_t)$, $(R_\bullet^s, Z_s) \perp F_{s^+}$, (R_\bullet^t, Z_t) have symmetric distributions.

Equation (22) follows from (8), the definition of $Y_\gamma^{\varepsilon,t}$, $R_\gamma^{\varepsilon,t}$, and condition (B)(iii), (23) follows from condition (B)(ii).

Set $c_t = \alpha_t a_t^*$. Since the integrands in (19), (20) are $O_{L^p}(1) \forall p > 0$, applying dominated convergence and (22), our problem reduces to show that, for all $\gamma, \gamma' > 0$, and almost every pair $s, t > 0$, $s \neq t$,

$$(24) \quad \lim_{\varepsilon \rightarrow 0} E\{f(X_t)f(X_s)J(R_\gamma^{\varepsilon,t}, \beta^2(\gamma))J(R_{\gamma'}^{\varepsilon,s}, \beta^2(\gamma'))c_t c_s Z_{t,\varepsilon} Z_{s,\varepsilon}\} = 0,$$

$$(25) \quad \lim_{\varepsilon \rightarrow 0} E\{f(X_t)f(X_s)H(R_\gamma^{\varepsilon,t}, \beta^2(\gamma))H(R_{\gamma'}^{\varepsilon,s}, \beta^2(\gamma'))b_t b_s\} = 0.$$

Assume that $s > t$ are such that (23) holds and take ε so that $0 < \varepsilon < s - t$. Conditioning on F_s and using that $(R_{\gamma'}^{\varepsilon,s}, Z_{s,\varepsilon}) \perp F_s$, we reduce the problem to show that

$$\lim_{\varepsilon \rightarrow 0} E\{J(R_{\gamma'}^{\varepsilon,s}, \beta^2(\gamma'))Z_{s,\varepsilon}\} = 0 = \lim_{\varepsilon \rightarrow 0} E\{H(R_{\gamma'}^{\varepsilon,s}, \beta^2(\gamma'))\}.$$

From (23) and uniform integrability it suffices to prove that

$$E\{J(R_\gamma^t, \beta^2(\gamma))Z_t\} = E\{H(R_\gamma^t, \beta^2(\gamma))\} = 0,$$

which is obvious by the symmetry of the distribution of (R_γ^t, Z_t) and the fact that $J(\cdot, \theta)$ (resp. $H(\cdot, \theta)$) is even (resp. odd).

PROOF OF STEP 3. Replacing $g^t(\sqrt{\varepsilon}\dot{X}_\varepsilon(t))$ by the formula in Step 2 we obtain

$$Z_\varepsilon^\tau = \int_0^\tau \int_0^1 f(X_t) a_t \frac{\Phi_{\varepsilon,t}(v)}{\sqrt{\varepsilon}} d_v W_v^{\varepsilon,t} dt + \frac{1}{\sqrt{\varepsilon}} \int_0^\tau f(X_t) R_\varepsilon(t) dt.$$

For $\varepsilon > 0$ fixed and every $p > 0$, it is obvious that the integrand in the first term of the right-hand member is $O_{L^p}(1)$; hence the Fubini-type Lemma 4 shows that

$$\begin{aligned} & \int_0^\tau \int_0^1 f(X_t) a_t \frac{\Phi_{\varepsilon,t}(v)}{\sqrt{\varepsilon}} d_v W_v^{\varepsilon,t} dt \\ &= \int_0^\tau \int_t^{t+\varepsilon u} f(X_t) a_t \frac{\Phi_{\varepsilon,t}((u-t)/\varepsilon)}{\varepsilon} dW_u dt \\ &= \int_0^{\tau+\varepsilon} \int_{\max(0, u-\varepsilon)}^{\min(u, \tau)} f(X_t) a_t \frac{\Phi_{\varepsilon,t}((u-t)/\varepsilon)}{\varepsilon} dt dW_u. \end{aligned}$$

Define

$$(26) \quad Q_\varepsilon(u, \tau) = \int_{\max(0, u-\varepsilon)}^{\min(u, \tau)} f(X_t) a_t \frac{\Phi_{\varepsilon,t}((u-t)/\varepsilon)}{\varepsilon} dt,$$

$$(27) \quad K_\varepsilon(u) = \int_{\max(0, u-\varepsilon)}^u \frac{\Phi_{\varepsilon,t}((u-t)/\varepsilon)}{\varepsilon} dt.$$

By (15) and Lemma 2, it follows that

$$(28) \quad \sup_{u \in [0, 1], \tau \in [0, 1]} E\{|Q_\varepsilon(u, \tau)|^p\} < \infty, \quad \sup_{u \in [0, 1]} E\{|K_\varepsilon(u)|^p\} < \infty.$$

From this, the continuity of X, a , the Burkholder–Davies–Gundy inequality and Lemma 2, we obtain

$$(29) \quad \int_{\max(0, u-\varepsilon)}^{\min(u, \tau)} [f(X_u)a_u K_\varepsilon(u) - Q_\varepsilon(u, \tau)] dW_u = o_{L^p}(1)$$

[indeed, it is an $O_{L^p}(\sqrt{\varepsilon}) \forall p > 0$];

$$(30) \quad \int_\tau^{\tau+\varepsilon} Q_\varepsilon(u, \tau) dW_u = O_{L^p}(\varepsilon) \quad \forall p > 0.$$

After Step 2, (29) and (30), Step 3 is proved.

PROOF OF STEP 4. As a consequence of (28) and Jensen's inequality, both V^* and \hat{V} are tight. Equations (22),(23) also imply

$$(31) \quad E\{\hat{V}_\varepsilon(\tau)\} \xrightarrow{\varepsilon \rightarrow 0^+} c_\psi^2 \int_0^\tau E\{V_t\} dt,$$

$$(32) \quad E\{\hat{V}_\varepsilon^2(\tau)\} \xrightarrow{\varepsilon \rightarrow 0^+} c_\psi^4 E\left\{\left[\int_0^\tau V_t dt\right]^2\right\},$$

$$(33) \quad E\left\{\hat{V}_\varepsilon(\tau) \int_0^\tau V_t dt\right\} \xrightarrow{\varepsilon \rightarrow 0^+} c_\psi^2 E\left\{\left[\int_0^\tau V_t dt\right]^2\right\},$$

$$(34) \quad E\{V_\varepsilon^*(\tau)^2\} \xrightarrow{\varepsilon \rightarrow 0^+} 0;$$

and Step 4 follows.

PROOF OF STEP 5. Apply Step 4 with $V = 1$ and Rebolledo's theorem for convergence of martingales [cf. Revuz and Yor (1991), page 478].

PROOF OF STEP 6. Consider $P_t := f(X_t)a_t$ and

$$P_t^N := \sum_{i=0}^{i=N-1} f(X_{i\tau/N})a_{i\tau/N} \mathbb{1}_{[i\tau/N, (i+1)\tau/N)}(t).$$

It follows from Step 5 that

$$(35) \quad \int_0^\tau P_t^N dS_\varepsilon(t) \xrightarrow[\varepsilon \rightarrow 0]{w} c_\psi \int_0^\tau P_t dB_t \quad \text{in } \mathbb{C}([0, 1]).$$

Step 4 applied to $(P_t - P_t^N)^2$ shows that

$$(36) \quad \lim_{N \rightarrow \infty} \lim_{\varepsilon \rightarrow 0} \int_0^\tau (P - P^N) dS_\varepsilon(t) = 0 \quad \text{in } \mathbb{C}([0, 1]).$$

Since

$$\int_0^\tau P_t^N dB_t \xrightarrow[N \rightarrow \infty]{w} \int_0^\tau P_t dB_t$$

and by (35), (36), Step 6 follows and the theorem is proved. \square

4. Auxiliary lemmas.

LEMMA 1. *If α satisfies (A), (B) and $\varphi: \mathbb{R} \rightarrow \mathbb{R}$ is a bounded C^∞ function, then $\varphi \circ \alpha$ satisfies (A), (B).*

For the proof, use Taylor's expansion.

LEMMA 2. *Let X be as in (1), with α and b uniformly bounded by a (non-random) constant. Then, $\forall p \geq 2$ we have the following:*

- (a) $E\{\sup_{t \in [0, 1]} |X_t|^p\} < \infty$.
- (b) $E\{\sup_{t \in [0, 1]} |X_\varepsilon(t)|^p\} < \infty$.
- (c) $E\{\sup_{t \in [0, 1]} |\dot{X}_\varepsilon(t)|^p\} = O(\varepsilon^{-p/2})$.
- (d) $E\{\sup_{t \in [0, 1]} |\Delta_\varepsilon(t)|^p\} = O(\varepsilon^{p/2})$.

For the proof, use the Burkholder–Davis–Gundy inequality.

LEMMA 3. *Let $V = \{V_t: t \geq 0\}$ be a real-valued adapted process such that $\sup_{t \in [0, 1]} E\{|V_t|^p\} < \infty \forall p > 0$.*

If X is as in (1), with α and b uniformly bounded by a (nonrandom) constant, we have, for $0 < \varepsilon < 1$, $0 < h < 1$,

$$(a) \quad \sup_{0 \leq t \leq 1-h} E \left[\left\{ \int_t^{t+h} \Delta_\varepsilon^2(s) |\dot{X}_\varepsilon(s)| V_s ds \right\}^2 \right] \leq Ch^2 \varepsilon,$$

$$(b) \quad \sup_{0 \leq t \leq 1-h} E \left[\left\{ \int_t^{t+h} \Delta_\varepsilon(s) |\dot{X}_\varepsilon(s)| V_s ds \right\}^2 \right] \leq C \sqrt{\varepsilon} h^{3/2},$$

(c) *if $f \in C_b^2(\mathbb{R})$, then*

$$\sup_{0 \leq t \leq 1-h} E \left[\left\{ \int_t^{t+h} f(X_\varepsilon(s)) - f(X_s) |\dot{X}_\varepsilon(s)| ds \right\}^2 \right] \leq C \sqrt{\varepsilon} h^{3/2}.$$

PROOF. (a) Apply Lemma 2.

(b) Observe that

$$(37) \quad \begin{aligned} \sqrt{\varepsilon} \dot{X}_\varepsilon(s) &= \int_0^1 \psi(-u) a_{s+\varepsilon u} d_u W_u^{\varepsilon, s} + \sqrt{\varepsilon} \int \psi(-u) b_{s+\varepsilon u} du \\ &= a_s \int_0^1 \psi(-u) d_u W_u^{\varepsilon, s} + O_{L^p}(\varepsilon) \quad \forall p > 0; \end{aligned}$$

$$(38) \quad \begin{aligned} \frac{\Delta_\varepsilon(s)}{\sqrt{\varepsilon}} &= \int_0^1 \psi(-u) \int_0^u a_{s+\varepsilon v} d_v W_v^{\varepsilon, s} du + \sqrt{\varepsilon} \int_0^1 \psi(-u) \int_0^u b_{s+\varepsilon v} d_v du \\ &= a_s \int_0^1 \psi(-u) W_u^{\varepsilon, s} du + O_{L^p}(\varepsilon) \quad \forall p > 0. \end{aligned}$$

Set $H_\varepsilon(s, r) = E\{\Delta_\varepsilon(s)|\dot{X}_\varepsilon(s)|V_s\Delta_\varepsilon(r)|\dot{X}_\varepsilon(r)|V_r\}$.

Compute the second moment as follows:

$$\begin{aligned} &E\left\{\left[\int_t^{t+h} \Delta_\varepsilon(s)|\dot{X}_\varepsilon(s)|V_s ds\right]^2\right\} \\ &= \int_t^{t+h} \int_t^{t+h} E\{\Delta_\varepsilon(s)|\dot{X}_\varepsilon(s)|V_s\Delta_\varepsilon(r)|\dot{X}_\varepsilon(r)|V_r\} dr ds \\ &= \int_{\{t \leq r, s \leq t+h, |s-r| < \varepsilon\}} H_\varepsilon(s, r) dr ds + 2 \int_t^{t+h} \int_{s+\varepsilon}^{t+h} H_\varepsilon(s, r) dr ds \\ &= (I) + 2(II). \end{aligned}$$

Taking into account that the integrand H_ε is bounded, it is trivial to observe that

$$(39) \quad (I) \leq Ch \min\{\varepsilon, h\} \leq C\sqrt{\varepsilon}h^{3/2}.$$

For the second term, $A_{s,r} = a_s^2 V_s a_r^2 V_r$, and using (37), we deduce

$$(40) \quad \begin{aligned} H_\varepsilon(s, r) &= E\left\{A_{s,r} \left| \int_0^1 \psi(-u) d_u W_u^{\varepsilon, s} \right| \int_0^1 \psi(-u) W_u^{\varepsilon, s} du \right. \\ &\quad \left. \times \left| \int_0^1 \psi(-u) d_u W_u^{\varepsilon, r} \right| \int_0^1 \psi(-u) W_u^{\varepsilon, r} du \right\} + O(\sqrt{\varepsilon}). \end{aligned}$$

Since $s+\varepsilon \leq r$, conditioning to F_r and using the independence of the Brownian increments, we get

$$H_\varepsilon(s, r) = P(s, r) + O(\sqrt{\varepsilon})$$

with

$$(41) \quad \begin{aligned} P(s, r) &= E\left\{A_{s,r} \left| \int_0^1 \psi(-u) d_u W_u^{\varepsilon, s} \right| \int_0^1 \psi(-u) W_u^{\varepsilon, s} du \right\} \\ &\quad \times E\left\{\left| \int_0^1 \psi(-u) d_u W_u^{\varepsilon, r} \right| \int_0^1 \psi(-u) W_u^{\varepsilon, r} du\right\}. \end{aligned}$$

Since $(\int_0^1 \psi(-u) d_u W_u^{\varepsilon, r}, \int_0^1 \psi(-u) W_u^{\varepsilon, r} du)$ is a centered Gaussian vector, it follows by symmetry that

$$(42) \quad E\left\{\left| \int_0^1 \psi(-u) d_u W_u^{\varepsilon, r} \right| \int_0^1 \psi(-u) W_u^{\varepsilon, r} du\right\} = 0.$$

By (41) and (42) we deduce

$$(43) \quad (II) \leq C\sqrt{\varepsilon}h^2 \leq C\sqrt{\varepsilon}h^{3/2}.$$

This concludes the proof of part (b).

(c) Use Taylor's formula, apply (b) to the linear term and (a) to the quadratic one. \square

LEMMA 4. *Let $\{K(t, s): t, s \in [0, 1]\}$ be a real-valued random process such that*

$$(a) \quad \sup_{t, s \in [0, 1]} E\{|K(t, s)|^p\} < \infty \quad \forall p > 0,$$

$$(b) \quad \int_0^1 K(t, s) ds \text{ is predictable,}$$

$$(c) \quad \int_0^1 K(t, s) dW_t \text{ is measurable.}$$

Then

$$\int_0^\tau \int_0^1 K(t, s) ds dW_t = \int_0^1 \int_0^\tau K(t, s) dW_t ds.$$

The proof is an analogue to Lemma 1.4.1 of Ikeda and Watanabe (1981).

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