

**SOME IMPROVEMENTS IN SETTING LIMITS FOR THE EXPECTED  
NUMBER OF OBSERVATIONS REQUIRED BY A SEQUENTIAL  
PROBABILITY RATIO TEST**

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**Summary.** Upper and lower limits for the expected number  $n$  of observations required by a sequential probability ratio test have been derived in a previous publication [1]. The limits given there, however, are far apart and of little practical value when the expected value of a single term  $z$  in the cumulative sum computed at each stage of the sequential test is near zero. In this paper upper and lower limits for the expected value of  $n$  are derived which will, in general, be close to each other when the expected value of  $z$  is in the neighborhood of zero. These limits are expressed in terms of limits for the expected values of certain functions of the cumulative sum  $Z_n$  at the termination of the sequential test.

In section 7 a general method is given for determining limits for the expected value of any function of  $Z_n$ .

**1. Introduction.** Let  $x$  be a random variable and let  $f(x, \theta)$  be the elementary probability law of  $x$  involving an unknown parameter  $\theta$ . Let  $H_0$  denote the hypothesis that  $\theta = \theta_0$ , and  $H_1$  the hypothesis that  $\theta = \theta_1$ , where  $\theta_0$  and  $\theta_1$  are given specified values. The sequential probability ratio test for testing  $H_0$  against  $H_1$ , as defined in [1], is given as follows: Put

$$(1.1) \quad z_i = \log \frac{f(x_i, \theta_1)}{f(x_i, \theta_0)}$$

where  $x_i$  denotes the  $i$ -th observation on  $x$ . Two constants,  $a$  and  $b$  are chosen where  $a > 0$  and  $b < 0$ . At each stage of the experiment, at the  $m$ -th trial for each positive integral value  $m$ , the cumulative sum

$$(1.2) \quad Z_m = z_1 + \cdots + z_m$$

is computed. Experimentation is continued as long as  $b < Z_m < a$ . The first time that  $Z_m$  does not lie between  $b$  and  $a$ , experimentation is terminated. The hypothesis  $H_1$  is accepted if  $Z_m \geq a$ , and  $H_0$  is accepted if  $Z_m \leq b$ .

Let  $n$  denote the smallest value of  $m$  for which  $Z_m$  does not lie between  $b$  and  $a$ . Then  $n$  is the number of observations required by the sequential test. The expected value of  $n$  is a function of the true parameter value  $\theta$  and is denoted by  $E_\theta(n)$ .

Upper and lower limits for  $E_\theta(n)$  have been derived in section 4 of [1]. These limits, however, are of little practical value when the expected value of

$$(1.3) \quad z = \log \frac{f(x, \theta_1)}{f(x, \theta_0)}$$

is in the neighborhood of zero, for they converge to  $+\infty$  and  $-\infty$ , respectively, as the expected value of  $z$  approaches zero. It can be shown that the expected value of  $z$  is negative when  $\theta = \theta_0$ , and positive when  $\theta = \theta_1$ .<sup>1</sup> Thus, if the expected value of  $z$  is a continuous function of  $\theta$ , there will be a value  $\theta'$  between  $\theta_0$  and  $\theta_1$  such that the expected value of  $z$  is zero when  $\theta = \theta'$ . Hence, the limits for  $E_\theta(n)$ , as given in [1], are of no practical value when  $\theta$  is near  $\theta'$ .

The purpose of this paper is to derive upper and lower limits for  $E_\theta(n)$  which will be, in general, close to each other when  $\theta$  is in the neighborhood of  $\theta'$ . Thus, it will generally be possible to obtain close limits for  $E_\theta(n)$  over the whole range of  $\theta$ , if the limits given here are used for values in a certain small interval containing  $\theta'$ , and the limits given in [1] are used when  $\theta$  is outside this interval.

**2. Notation.** We shall use the following notations throughout the paper. For any random variable  $u$ , the symbol  $E_\theta(u)$  will denote the expected value of  $u$  when  $\theta$  is the true value of the parameter. The conditional expected value of  $u$ , under the restriction that some relationship  $R$  is fulfilled will be denoted by  $E_\theta(u | R)$ . The symbol  $P(R | \theta)$  will denote the probability that the relationship  $R$  holds when  $\theta$  is true.

The cumulative distribution function of  $z$  will be denoted by  $F(z, \theta)$  when  $\theta$  is the true value of the parameter. The moment generating function of  $z$ , when  $\theta$  is true, will be denoted by  $\varphi(t, \theta)$ , i.e.

$$(2.1) \quad \varphi(t, \theta) = \int_{-\infty}^{\infty} e^{tz} dF(z, \theta).$$

**3. Assumptions concerning the family of distribution functions  $F(z, \theta)$ .** In this section we shall formulate two assumptions concerning  $F(z, \theta)$  which will then be used to prove various lemmas and theorems. Since we are interested in values of  $\theta$  near  $\theta'$ , we shall restrict the domain of  $\theta$  to a finite closed interval  $I$  containing  $\theta'$  in its interior. It will be understood throughout the paper that any statements concerning  $\theta$  refer to the domain  $I$ , even if this is not explicitly stated.

**ASSUMPTION 1.** *The moment generating function  $\varphi(t, \theta)$  exists for any point  $t$  in the complex plane and any value  $\theta$ , and is a continuous function of  $\theta$ .*

**ASSUMPTION 2.** *There exists a positive  $\delta$  such that  $P(e^z > 1 + \delta | \theta)$  and  $P(e^z < 1 - \delta | \theta)$  have positive lower bounds with respect to  $\theta$ .*

**4. Proof that  $\varphi(t, \theta)$  is continuous in  $t$  and  $\theta$  jointly and that all moments of  $z$  are continuous functions of  $\theta$ .**<sup>2</sup> In this section we shall prove the following theorem:

<sup>1</sup> This follows easily from Lemma 1 in [1], p. 156.

<sup>2</sup> The original proof of the author was somewhat lengthy. The present proof was suggested by T. E. Harris.

**THEOREM 4.1.** *It follows from Assumption 1 that  $\varphi(t, \theta)$  is continuous in  $t$  and  $\theta$  jointly and all moments of  $z$  are continuous functions of  $\theta$ .*

**PROOF:** First we show that  $\varphi(t, \theta)$  is a bounded function of  $t$  and  $\theta$  in the domain  $|t| \leq t_0$ , for any finite positive value  $t_0$ . Clearly,

$$(4.1) \quad 0 \leq |\varphi(t, \theta)| \leq 2[\varphi(t_0, \theta) + \varphi(-t_0, \theta)]$$

for all values  $t$  for which  $|t| \leq t_0$ . The boundedness of  $\varphi(t_0, \theta)$  and  $\varphi(-t_0, \theta)$  follows from Assumption 1. Hence  $\varphi(t, \theta)$  is a bounded function of  $\theta$  and  $t$  over any bounded  $t$ -domain.

Let  $\{t_m, \theta_m\}$  ( $m = 1, 2, \dots$ , ad inf.) be a sequence of pairs converging to the pair  $(t', \theta')$ . We have

$$(4.2) \quad \varphi(t_m, \theta_m) - \varphi(t', \theta') = [\varphi(t_m, \theta_m) - \varphi(t', \theta_m)] + [\varphi(t', \theta_m) - \varphi(t', \theta')].$$

The second expression in brackets converges to zero by continuity in  $\theta$ . Thus the first part of Theorem 4.1 is proved if we show that

$$(4.3) \quad \lim_{m \rightarrow \infty} [\varphi(t_m, \theta_m) - \varphi(t', \theta_m)] = 0.$$

It follows from Assumption 1 that for any given  $\theta$ ,  $\varphi(t, \theta)$  is an analytic function with no singularities in any finite  $t$ -domain. Hence we can expand  $\varphi(t_m, \theta_m)$  in a Taylor series around  $t = t'$ , i.e.

$$(4.4) \quad \varphi(t_m, \theta_m) - \varphi(t', \theta_m) = \sum_{k=1}^{\infty} \frac{1}{k!} \left( \frac{\partial^k \varphi(t, \theta_m)}{\partial t^k} \Big|_{t=t'} \right) (t_m - t')^k.$$

Let  $r$  be a given positive value. Because of the boundedness of  $\varphi(t, \theta)$  in any finite  $t$ -domain, there exists a constant  $M$  such that  $|\varphi(t, \theta)| < M$  for all  $\theta$  and for all  $t$  in the domain  $|t - t'| \leq r$ . From the Cauchy integral formula for an analytic function it follows that

$$(4.5) \quad \frac{1}{k!} \left| \frac{\partial^k \varphi(t, \theta_m)}{\partial t^k} \Big|_{t=t'} \right| \leq \frac{M}{r^k}.$$

From (4.4) and (4.5) we obtain

$$(4.6) \quad |\varphi(t_m, \theta_m) - \varphi(t', \theta_m)| \leq M \sum_{k=1}^{\infty} \frac{|t_m - t'|^k}{r^k}.$$

Equation (4.3) is an immediate consequence of (4.6). This proves the first half of Theorem 4.1.

Let  $C$  be a circle in the complex  $t$ -plane with finite radius and center at the origin. According to the Cauchy integral formula we have

$$(4.7) \quad \frac{1}{2\pi i} \int_C \frac{\varphi(t, \theta)}{t^{k+1}} dt = \frac{1}{k!} \frac{\partial^k \varphi(t, \theta)}{\partial t^k} \Big|_{t=0} = \frac{1}{k!} E_{\theta}(z^k).$$

Since  $\varphi(t, \theta)$  is continuous in  $t$  and  $\theta$  jointly, the integral on the left hand side of (4.7) is a continuous function of  $\theta$ . This proves the second half of Theorem 4.1.

**5. Some lemmas.** In this section we shall prove several lemmas which will then be used to derive the results contained in sections 6 and 8.

LEMMA 5.1. *It follows from assumptions 1 and 2 that for any given  $\theta$  the equation in  $t$*

$$(5.1) \quad \varphi(t, \theta) = 1$$

*has exactly two real roots, one of which is zero. The other real root is different from zero if  $E_\theta(z) \neq 0$ . If  $E_\theta(z) = 0$ , both roots are equal to zero, i.e., zero is a double root of (5.1).*

This lemma is essentially the same as Lemma 2 in [2] and the proof is therefore omitted.<sup>3</sup>

Let  $h(\theta)$  denote the non-zero root of (5.1), if  $E_\theta(z) \neq 0$ . If  $E_\theta(z) = 0$ , we put  $h(\theta) = 0$ .

In what follows the variable  $t$  will be restricted to real values, unless the contrary is explicitly stated.

LEMMA 5.2. *It follows from assumptions 1 and 2 that  $h(\theta)$  is a continuous function of  $\theta$ .*

PROOF: It follows from assumption 2 that

$$(5.2) \quad \lim_{t \rightarrow \pm\infty} \varphi(t, \theta) = +\infty$$

uniformly in  $\theta$ . Hence, since by definition

$$\varphi[h(\theta), \theta] = 1$$

identically in  $\theta$ ,  $h(\theta)$  must be a bounded function of  $\theta$ .

Let  $\{\theta_m\}$  be a sequence of parameter values which converges to  $\theta^*$ . From Theorem 4.1 it follows that

$$(5.3) \quad \lim_{m \rightarrow \infty} [\varphi(t, \theta_m) - \varphi(t, \theta^*)] = 0$$

uniformly in  $t$  over any finite interval. Since  $h(\theta)$  is bounded, we obtain from (5.3)

$$(5.4) \quad \lim_{m \rightarrow \infty} \{\varphi[h(\theta_m), \theta_m] - \varphi[h(\theta_m), \theta^*]\} = 0.$$

Since  $\varphi[h(\theta_m), \theta_m] = 1$ , it follows from (5.4) that

$$\lim_{m \rightarrow \infty} \varphi[h(\theta_m), \theta^*] = 1.$$

It follows from assumption 1 that for any limit point  $h$  of the bounded sequence  $\{h(\theta_m)\}$  ( $m = 1, 2, \dots$ , ad inf.) we have

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<sup>3</sup> Condition IV of Lemma 2 in [2] is not postulated here, since the validity of this condition is implied by assumption 1. Condition IV could have been omitted also in [2], since it follows from condition III.

$$(5.5) \quad \varphi(h, \theta^*) = 1$$

If  $h(\theta^*) = 0$ , then equation  $\varphi(t, \theta^*) = 1$  has the only root  $t = 0$ . Consequently, all limit points of  $\{h(\theta_m)\}$  must be equal to zero, that is

$$(5.6) \quad \lim_{m \rightarrow \infty} h(\theta_m) = 0 \quad \text{if} \quad h(\theta^*) = 0.$$

Now let us assume that  $h(\theta^*) \neq 0$ . Since the second derivative of  $\varphi(t, \theta)$  with respect to  $t$  is positive, it can be seen that  $\varphi(t, \theta) < 1$  for values  $t$  in the open interval  $(0, h(\theta))$ , and  $\varphi(t, \theta) > 1$  for any  $t$  outside the closed interval  $[0, h(\theta)]$ . Hence,  $\varphi(t, \theta) < 1$  implies that  $|h(\theta)| > |t|$  and  $h(\theta)$  and  $t$  have the same sign. Now let  $t_0$  be a value in the open interval  $(0, h(\theta^*))$ . Then we have

$$(5.7) \quad \varphi(t_0, \theta^*) < 1$$

It follows from assumption 1 that

$$(5.8) \quad \varphi(t_0, \theta_m) < 1$$

for sufficiently large  $m$ . Hence  $h(\theta_m)$  and  $t_0$  have the same sign and

$$(5.9) \quad |h(\theta_m)| > |t_0|$$

Inequality (5.9) implies that zero cannot be a limit point of the sequence  $\{h(\theta_m)\}$ . Since  $\varphi(t, \theta^*) = 1$  has only the roots  $t = 0$  and  $t = h(\theta^*)$ , it follows from (5.5) that the sequence  $\{h(\theta_m)\}$  cannot have a limit point different from  $h(\theta^*)$ . Thus,

$$(5.10) \quad \lim_{m \rightarrow \infty} h(\theta_m) = h(\theta^*)$$

and Lemma 5.2 is proved.

**LEMMA 5.3.** *It follows from assumption 1 that for any given  $t$ ,  $E_\theta(e^{tz})$  is a bounded function of  $\theta$ .*

**PROOF:** We have

$$(5.11) \quad E_\theta(e^{tz}) \leq E_\theta(e^{tz} + e^{-tz}) = \varphi(t, \theta) + \varphi(-t, \theta)$$

It follows from assumption 1 that  $\varphi(t, \theta)$  and  $\varphi(-t, \theta)$  are bounded functions of  $\theta$ . Hence Lemma 5.3 is proved.

**LEMMA 5.4.** *Let  $\theta'$  be a value of  $\theta$  such that  $E_{\theta'}(z) = 0$ , but  $E_\theta(z) \neq 0$  for all  $\theta \neq \theta'$  in an open interval containing  $\theta'$ . It follows from assumptions 1 and 2 that*

$$(5.12) \quad \lim_{\theta \rightarrow \theta'} \left( -\frac{2E_\theta(z)}{h(\theta)} \right) = E_{\theta'}(z^2).$$

**PROOF:** We have

$$(5.13) \quad e^{h(\theta)z} = 1 + h(\theta)z + \frac{[h(\theta)]^2}{2} z^2 + \frac{[h(\theta)]^3}{6} z^3 e^{uh(\theta)z}$$

where  $0 \leq u \leq 1$ . Hence

$$(5.14) \quad E_\theta(e^{h(\theta)z}) = 1 + h(\theta)E_\theta(z) + \frac{[h(\theta)]^2}{2} E_\theta(z^2) + \frac{[h(\theta)]^3}{6} E_\theta(z^3 e^{uh(\theta)z}).$$

Since  $E_\theta(e^{h(\theta)z}) = 1$ , we obtain from (5.14)

$$(5.15) \quad h(\theta)E_\theta(z) + \frac{[h(\theta)]^2}{2} E_\theta(z^2) + \frac{[h(\theta)]^3}{6} E_\theta(z^3 e^{uh(\theta)z}) = 0.$$

We shall consider only values  $\theta$  for which  $h(\theta) \neq 0$ . For such values of  $\theta$ , also  $E_\theta(z) \neq 0$ . Dividing (5.15) by  $h(\theta)E_\theta(z)$ , we obtain

$$(5.16) \quad 1 + \frac{h(\theta)}{2E_\theta(z)} \left[ E_\theta(z^2) + \frac{h(\theta)}{3} E_\theta(z^3 e^{uh(\theta)z}) \right] = 0.$$

Let  $t_0$  be an upper bound of  $|h(\theta)|$  with respect to  $\theta$ . Then for a suitably chosen constant  $C$  we have

$$(5.17) \quad |z^3 e^{uh(\theta)z}| < C e^{|t_0 z|}$$

From this and Lemma 5.3 it follows that  $E_\theta(z^3 e^{uh(\theta)z})$  is a bounded function of  $\theta$ .

Because of the continuity of  $h(\theta)$  we have

$$(5.18) \quad \lim_{\theta \rightarrow \theta'} h(\theta) = 0.$$

Lemma 5.4 follows from (5.16), (5.18), the boundedness of  $E_\theta(z^3 e^{uh(\theta)z})$  and the fact that  $E_\theta(z^2)$  is a continuous function of  $\theta$  and  $E_\theta(z^2) > 0$ .

LEMMA 5.5. *From assumptions 1 and 2 it follows that for any given  $t$ ,  $E_\theta(e^{t z_n})$  exists and is a bounded function of  $\theta$ .*

PROOF: It is sufficient to show that  $E_\theta(e^{t z_n})$  is a bounded function of  $\theta$  for any  $t$ , since

$$(5.19) \quad e^{|t z_n|} \leq e^{t z_n} + e^{-t z_n}$$

Clearly,  $e^{t z_n}$  lies between  $e^{b t + z_n t}$  and  $e^{a t + z_n t}$ . Hence Lemma 5.5 is proved if we show that  $E_\theta(e^{z_n t})$  is a bounded function of  $\theta$ .

It follows from Assumption 2 that there exists a positive integer  $k$  and a positive constant  $g$  such that

$$(5.20) \quad P(|z_1 + \dots + z_k| \geq a - b | \theta) \geq g$$

for all  $\theta$ . For any positive integer  $m$  and for any real values  $\lambda_1 < \lambda_2$  we have

$$(5.21) \quad \frac{P[(m-1)k < n \leq mk | \theta]}{P[(m-1)k < n | \theta]} \geq g \quad (m = 1, 2, \dots, \text{ad inf.})$$

and

$$(5.22) \quad \frac{P[(m-1)k < n \leq mk \ \& \ \lambda_1 \leq z_n < \lambda_2 | \theta]}{P[(m-1)k < n | \theta]} \leq 1 - [1 - P(\lambda_1 \leq z < \lambda_2 | \theta)]^k.$$

Hence

$$(5.23) \quad \frac{P[(m-1)k < n \leq mk \text{ \& } \lambda_1 \leq z_n < \lambda_2 | \theta]}{P[(m-1)k < n \leq mk | \theta]} \leq \frac{1 - [1 - P(\lambda_1 \leq z < \lambda_2 | \theta)]^k}{g}.$$

Multiplying (5.23) by  $P[(m-1)k < n \leq mk | \theta]$  and summing with respect to  $m$  we obtain

$$(5.24) \quad P(\lambda_1 \leq z_n < \lambda_2 | \theta) \leq \frac{1 - [1 - P(\lambda_1 \leq z < \lambda_2 | \theta)]^k}{g}.$$

From (5.24) it follows readily that

$$(5.25) \quad \frac{P(\lambda_1 \leq z_n < \lambda_2 | \theta)}{P(\lambda_1 \leq z < \lambda_2 | \theta)}$$

is a bounded function of  $\lambda_1$ ,  $\lambda_2$  and  $\theta$ . Let  $A$  be an upper bound of the ratio (5.25). Then

$$(5.26) \quad E_\theta(e^{tz_n}) \leq AE_\theta(e^{tz}) = A\varphi(t, \theta).$$

Because of Assumption 1,  $\varphi(t, \theta)$  is a bounded function of  $\theta$ . Hence also  $E_\theta(e^{tz_n})$  is bounded and Lemma 5.5 is proved.

**6. The limiting value of  $E_\theta(n)$  when  $\theta$  approaches a value  $\theta'$  for which  $E_{\theta'}(z) = 0$ .** In this section we shall prove the following theorem:

**THEOREM 6.1.** *Let  $\theta'$  be a value of  $\theta$  such that  $E_{\theta'}(z) = 0$ , but  $E_\theta(z) \neq 0$  for all  $\theta \neq \theta'$  in an open interval containing  $\theta'$ . If assumptions 1 and 2 hold, we have*

$$(6.1) \quad \lim_{\theta \rightarrow \theta'} \left[ E_\theta(n) - \frac{E_\theta(Zn^2)}{E_{\theta'}(z^2)} \right] = 0.$$

**PROOF:** Consider the Taylor expansion

$$(6.2) \quad e^{\lambda(\theta)z_n} = 1 + h(\theta)Z_n + \frac{[h(\theta)]^2}{2} Z_n^2 + \frac{[h(\theta)]^3}{6} Z_n^3 e^{\lambda h(\theta)z_n}$$

where  $0 \leq \lambda \leq 1$ . It was shown in [2] (p. 286) that

$$(6.3) \quad E_\theta e^{\lambda(\theta)z_n} = 1.$$

Hence, taking expected values on both sides of (6.2), we obtain

$$(6.4) \quad h(\theta)E_\theta(Z_n) + \frac{[h(\theta)]^2}{2} E_\theta(Z_n^2) + \frac{[h(\theta)]^3}{6} E_\theta(Z_n^3 e^{\lambda h(\theta)z_n}) = 0.$$

We consider only values of  $\theta$  for which  $E_\theta(z) \neq 0$ . For such values, also  $h(\theta) \neq 0$ . Thus, we can divide both sides of (6.4) by  $h(\theta)E_\theta(z)$ . We then obtain

$$(6.5) \quad \frac{E_\theta(Z_n)}{E_\theta(z)} + \frac{h(\theta)}{2E_\theta(z)} \left[ E_\theta Z_n^2 + \frac{h(\theta)}{3} E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n}) \right] = 0.$$

It was shown in [1] (p. 142) that

$$(6.6) \quad E_\theta(n) = \frac{E_\theta(Z_n)}{E_\theta(z)}.$$

Hence

$$(6.7) \quad E_\theta(n) + \frac{h(\theta)}{2E_\theta(z)} \left[ E_\theta(Z_n^2) + \frac{h(\theta)}{3} E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n}) \right] = 0.$$

Let  $t_0$  be an upper bound of  $|h(\theta)|$ . Then for a properly chosen constant  $C$  we have

$$(6.8) \quad |Z_n^3 e^{\lambda h(\theta) Z_n}| \leq C e^{t_0 Z_n}$$

From this and Lemma 5.5 it follows that  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$  is a bounded function of  $\theta$ . Since  $\lim_{\theta \rightarrow \theta'} h(\theta) = 0$  and  $E_\theta(Z_n^2)$  has a positive lower bound, Theorem

6.1 follows from 6.7, Lemma 5.4 and Theorem 4.1.

If  $\lim_{\theta \rightarrow \theta'} E_\theta Z_n^2 = E_{\theta'} Z_n^2$ , Theorem 6.1 gives<sup>4</sup>

$$(6.9) \quad E_{\theta'}(n) = \frac{E_{\theta'}(Z_n^2)}{E_{\theta'}(z^2)}.$$

Limits for  $E_\theta(n)$  can be obtained by computing limits for  $E_\theta(Z_n^2)$ . In the next section we shall give a general method for obtaining limits for  $E_\theta[\psi(Z_n)]$ , where  $\psi(Z_n)$  is any function of  $Z_n$ .

**7. Determination of lower and upper limits for the expected value of any function of  $Z_n$ .** Let  $\psi(Z_n)$  be a function of  $Z_n$ . Limits for  $E_\theta[\psi(Z_n)]$  may be determined as follows: First we determine limits for  $E_\theta[\psi(Z_n) | Z_n \geq a]$ . Let  $r$  be a positive variable. Clearly, for any given value  $r$  we have

$$(7.1) \quad E_{-\theta}(\psi Z_n | Z_{n-1} = a - r \text{ and } Z_n \geq a) = E_\theta[\psi(a - r + z) | z \geq r]$$

From (7.1) we obtain the limits

$$(7.2) \quad \begin{aligned} \text{g.l.b.}_{0 < r < a-b} E_\theta[\psi(a - r + z) | z \geq r] &\leq E_\theta[\psi(Z_n) | Z_n \geq a] \\ &\leq \text{l.u.b.}_{0 < r < a-b} E_\theta[\psi(a - r + z) | z \geq r]. \end{aligned}$$

Limits for  $E_\theta[\psi(Z_n) | Z_n \leq b]$  can be obtained in a similar way. Again, let  $r$  be a positive variable. For any value of  $r$  we have

$$(7.3) \quad E_\theta[\psi(Z_n) | Z_n \leq b \text{ and } Z_{n-1} = b + r] = E_\theta[\psi(b + r + z) | z \leq -r]$$

Hence we obtain the limits

<sup>4</sup> The validity of (6.9) was shown by the author [3] using an entirely different method.



$$(7.4) \quad \begin{aligned} \text{g.l.b.}_{0 < r < a-b} E_\theta[\psi(b+r+z) | z \leq -r] &\leq E_\theta[\psi(Z_n) | Z_n \leq b] \\ &\leq \text{l.u.b.}_{0 < r < a-b} E_\theta[\psi(b+r+z) | z \leq -r]. \end{aligned}$$

Since

$$(7.5) \quad E_\theta[\psi(Z_n)] = P(Z_n \geq a)E_\theta[\psi(Z_n) | Z_n \geq a] + P(Z_n \leq b)E_\theta[\psi(Z_n) | Z_n \leq b],$$

a lower (upper) limit for  $E_\theta[\psi(Z_n)]$  can be obtained, by replacing the conditional expected values on the right hand side of (7.5) by their lower (upper) limits given in (7.2) and (7.4).

**8. Limits for  $E_\theta(n)$  when  $h(\theta)$  is near but unequal to zero.** Let  $\theta'$  be a value of  $\theta$  for which  $h(\theta') = 0$ . In this section we shall derive limits for  $E_\theta(n)$  which will generally be close to each other for values  $\theta$  in a small neighborhood of  $\theta'$ .

From equation (6.7) we obtain

$$(8.1) \quad E_\theta(n) = -\frac{h(\theta)}{2E_\theta(z)} \left[ E_\theta Z_n^2 + \frac{h(\theta)}{3} E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n}) \right]$$

where  $0 \leq \lambda \leq 1$ . Thus, limits for  $E_\theta(n)$  can be obtained by deriving limits for  $E_\theta Z_n^2$  and  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$ . Limits for  $E_\theta Z_n^2$  can be obtained by using the method described in section 7.

If  $\theta$  is near  $\theta'$ , any crude limits for  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$  will serve the purpose, since, as has been shown in section 6,  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$  is bounded and  $\lim_{\theta \rightarrow \theta'} h(\theta) = 0$ .

Limits for  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$  can be obtained as follows: For simplicity, let us assume that  $h(\theta) > 0$ . Then

$$(8.2) \quad Z_n^3 \leq Z_n^3 e^{\lambda h(\theta) Z_n} \leq Z_n^3 e^{h(\theta) Z_n} \quad (h(\theta) > 0)$$

Thus, to determine limits for  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$ , it is sufficient to determine a lower limit for  $E_\theta(Z_n^3)$  and an upper limit for  $E_\theta(Z_n^3 e^{h(\theta) Z_n})$ . The latter limits may be derived by using the method given in section 7.

If  $h(\theta) < 0$ , we have

$$(8.3) \quad Z_n^3 \geq Z_n^3 e^{\lambda h(\theta) Z_n} \geq Z_n^3 e^{h(\theta) Z_n}$$

and a similar procedure will yield the desired limits for  $E_\theta(Z_n^3 e^{\lambda h(\theta) Z_n})$ .

It should be emphasized that the limits of  $E_\theta(n)$ , as given in this section, can be expected to be close only if  $h(\theta)$  is near zero. For values of  $\theta$  for which  $h(\theta)$  is not near zero, the limits of  $E_\theta(n)$  given in [1] can be used.

REFERENCES

[1] A. WALD, "Sequential tests of statistical hypotheses," *Annals of Math. Stat.*, Vol. 16, (1945), pp. 117-186.  
 [2] A. WALD, "On cumulative sums of random variables," *Annals of Math. Stat.*, Vol. 15, (1944), pp. 283-296.  
 [3] A. WALD, "Differentiation under the expected sign in the fundamental identity of sequential analysis," *Annals of Math. Stat.*, Vol. 17 (1946), pp. 493-497.