### SOME RESULTS ON RESTRICTED OCCUPANCY THEORY<sup>1</sup>

By John E. Freund and Arthur N. Pozner

# Virginia Polytechnic Institute

1. Introduction. If k judges rate a product on a discrete scale, say  $0, 1, 2, \dots$ , or m, it is not only important to know the average rating assigned to the product, but it is also important to know the consistency of the ratings. Whereas the average rating is r/k, where r is the total number of points assigned to the product by the k judges, it seems reasonable to use the variance of the ratings, i.e.,

(1) 
$$s^{2} = \sum_{i=0}^{m} \left(i - \frac{r}{k}\right)^{2} \left(\frac{v_{i}}{k}\right),$$

as a measure of consistency. Here  $v_i$  stands for the number of judges who gave the product the rating i.

In order to test the consistency of such ratings, it will be necessary to find a suitable mathematical model which will assign low probabilities to cases in which the ratings are either very inconsistent or overly consistent. It is felt that the appropriate model is provided by that of restricted occupancy theory, in which we consider as equally likely all possible distributions of r indistinguishable objects among k cells with at most m objects per cell. With this model we shall then test whether it is reasonable to suppose that the r points given by the k judges are randomly distributed among the k judges.

Related problems dealing with the probability that x cells contain more than q objects when r objects are distributed among k cells, and the probability that x cells contain m objects when r objects are distributed among k cells with at most m objects per cell were investigated by Baticle, [1], [2], and [3], with reference to applications to casualty insurance, merchandizing, and transportation.

2. A restricted occupancy distribution. Restricted occupancy theory deals with the distribution of r objects among k cells if a maximum of m objects is permitted per cell, the cells are distinguishable, and empty cells are permitted. In this paper we shall investigate the distribution of the variables  $v_i$  ( $i = 0, 1, 2, \dots, m$ ), standing for the number of cells occupied by i objects, respectively, if r indistinguishable objects are distributed among k cells with a maximum of m objects per cell.

Since r and k are assumed to be fixed, it should be noted that the variables  $v_i$  are subject to the two linear restrictions

(2) 
$$\sum_{i=0}^{m} v_i = k \quad \text{and} \quad \sum_{i=0}^{m} i \cdot v_i = r,$$

and any m-1 of the  $v_i$  will thus determine the remaining two.

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If we now write as N(r, k, m) the total number of ways in which r indistinguishable objects can be distributed among k cells with at most m objects per cell, we find that this restricted occupancy coefficient is the coefficient of  $x^r$  in

(3) 
$$f(x; k, m) = (1 + x + x^2 + \cdots + x^m)^k.$$

Since

$$f(x; k, m) = (1 - x^{m+1})^k (1 - x)^{-k}$$

$$= (1 - x^{m+1})^k \cdot \sum_{n=0}^{\infty} {\binom{-k}{n}} (-1)^n x^n$$

$$= (1 - x^{m+1})^k \cdot \sum_{n=0}^{\infty} {\binom{k+n-1}{k-1}} x^n,$$

we can write explicitly

(4) 
$$N(r,k,m) = \sum_{j=0}^{k} (-1)^{j} {k \choose j} {k+r-j(m+1)-1 \choose k-1}.$$

If we now let  $i_1$ ,  $i_2$ ,  $\cdots$ ,  $i_{m-1}$  stand for an arbitrary permutation of any m-1 of the m+1 numbers 0, 1, 2,  $\cdots$ , m, the joint distribution of the  $v_{i_j}$ , for  $j=1, 2, \cdots, m-1$ , may be written as

(5) 
$$f(v_{i_1}, \dots, v_{i_{m-1}}) = \frac{\binom{k}{v_{i_1}, v_{i_2}, \dots, v_{i_{m-1}}} \cdot \binom{k - \sum_{j=1}^{m-1} v_{i_j}}{q}}{N(r, k, m)},$$

where

$$q = \frac{k \cdot i_{m+1} - r - \sum_{j=1}^{m-1} (i_{m+1} - i_j)v_{i_j}}{i_{m+1} - i_m},$$

and

$$\binom{k}{v_{i_1}, v_{i_2}, \cdots, v_{i_{m-1}}}$$

is a multinomial coefficient. Also,  $i_m$  and  $i_{m+1}$  are the two numbers which were omitted when choosing the m-1 numbers  $i_j$ .

The moments of the  $v_{ij}$  may be obtained directly from (5), but their derivation is simplified if we refer to the generating function given in (3). Writing

$$f(x; k, m) = [x^{i} + (1 + x + x^{2} + \cdots + x^{m} - x^{i})]^{k}$$

$$= \sum_{j=0}^{k} {k \choose j} x^{ij} (1 + x + x^{2} + \cdots + x^{m} - x^{i})^{k-j},$$

we find that for a fixed value of j, the coefficient of  $x^r$  represents the number of ways in which r indistinguishable objects can be distributed among k cells with

at most m objects per cell under the condition that exactly j cells have exactly i occupants. Writing this coefficient as  $N(r, k, m|v_i = j)$ , the marginal distribution of  $v_i$  becomes

(6) 
$$\Pr \{v_i = j\} = \frac{N(r, k, m \mid v_i = j)}{N(r, k, m)}.$$

If we let  $E[v_i^{(t)}]$  stand for the tth factorial moment of  $v_i$ , we have

(7) 
$$N(r, k, m)E[v_i^{(t)}] = \sum_{j=0}^k j^{(t)}N(r, k, m) \mid v_i = j),$$

and the right-hand member of (7) is the coefficient of  $x^r$  in

$$\sum_{j=0}^{k} {k \choose j} j^{(i)} x^{ij} (1 + x + x^2 + \dots + x^m - x^i)^{k-j}$$

$$= k^{(i)} (1 + x + x^2 + \dots + x)^{k-i} x^{ii},$$

where  $j^{(t)}$  and  $k^{(t)}$  are the tth factorial powers of j and k. We thus find that

(8) 
$$E[v_i^{(t)}] = \frac{k^{(t)}N(r-ti,k-t,m)}{N(r,k,m)}.$$

Similarly, if we let  $N(r, k, m|v_{i_1} = j_1, v_{i_2} = j_2)$  stand for the number of ways in which r indistinguishable objects can be distributed among k cells with at most m objects per cell under the condition that exactly  $j_1$  cells have exactly  $i_1$  occupants and exactly  $j_2$  cells have exactly  $i_2$  occupants (with  $i_1$  not equal to  $i_2$ ), this quantity is given by the coefficient of  $x^r$  in the corresponding term of

$$f(x; k, m) = [x^{i_1} + x^{i_2} + (1 + x + \dots + x^m - x^{i_1} - x^{i_2})]^k$$

$$= \sum_{i_1=0}^k \sum_{j_2=0}^k {k \choose j_1, j_2} x^{i_1 j_1 + i_2 j_2} (1 + x + \dots + x^m - x^{i_1} - x^{i_2})^{k-j_1 - j_2}.$$

Using the same steps as above, it may then be shown that

(9) 
$$E[v_{i_1}^{(t_1)}v_{i_2}^{(t_2)}] = \frac{k^{(t_1+t_2)}N(r-i_1t_1-i_2t_2,k-t_1-t_2,m)}{N(r,k,m)},$$

and higher factorial product moments can be obtained in the same way.

The calculation of the restricted-occupancy coefficients needed in the evaluation of the moments of the  $v_{ij}$  and subsequently of the mean and variance of  $s^2$  [as defined in (1)] is greatly simplied by the use of the recursion relations

(10a) 
$$N(r, k, m) = \sum_{i=0}^{m} N(r - i, k - 1, m),$$

(10b) 
$$N(r, k-1, m) = \sum_{i=1}^{m} \frac{ki-r}{r} \cdot N(r-i, k-1, m),$$

or

(10c) 
$$N(r, k, m) = N(r - 1, k, m) + N(r, k - 1, m) - N(r - m - 1, k - 1, m)$$

These recursion formulas may be obtained directly by equating the coefficients of  $x^r$  in

(11a) 
$$f(x; k, m) = f(x; 1, m)f(x; k - 1, m),$$

(11b) 
$$f(x; 1, m)f'(x; k, m) = k \cdot f(x; k, m)f'(x; 1, m),$$

or

(11c) 
$$(1-x)f(x; k, m) = (1-x^{m+1})f(x; k-1, m).$$

Equation (10c) may also be obtained as an immediate consequence of (10a). Suitable tables for an exact test of significance for m=2 are in preparation. Tables of the restricted-occupancy coefficients N(r, k, 4) for values of k from 1 to 20 and of r from 1 to 40 are given in [4]. It was also found experimentally for m=2 and m=4 that the chi-square criterion applied to the observed  $v_i$  and their expectations provides a good approximate test of significance.

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## A VECTOR FORM OF THE WALD-WOLFOWITZ-HOEFFDING THEOREM<sup>1</sup>

### By D. A. S. Fraser

University of Toronto and Princeton University

- 1. Summary. Hotelling and Pabst [1] showed that the rank correlation coefficient had a limiting normal distribution under the equally likely permutations of the hypothesis of independence. Wald and Wolfowitz [2] developed a general theorem of this type, and Noether [3] and Hoeffding [4] have relaxed the conditions used therein. In this paper a vector form of the theorem is proved along the lines used in an example by Wald and Wolfowitz [1] but taking account of the singular cases in which the correlations approach one.
- 2. The theorem. For each positive integer n let  $||C_{nl}(i,j)||, \dots, ||C_{nk}(i,j)||$  be  $n \times n$  matrices of real numbers. Also let  $(R_1, \dots, R_n)$  be a random variable which takes each permutation of  $(1, \dots, n)$  with the same probability,  $(n!)^{-1}$ .

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