GAMES ASSOCIATED WITH A RENEWAL PROCESS

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- 1. Introduction. Consider a sequence of occurrences of a recurrent event \mathcal{E} for which the intervals, X_1 , X_2 , \cdots , are independent identically distributed non-negative random variables (a renewal process) with common cdf (cumulative distribution function) F(x). Robbins [3] considered games when X_1 is an integer-valued random variable. It seems of interest to extend his results to games when X_1 is not necessarily integer-valued. Thus, for example, $\{X_i\}$ may denote the lifetimes of similar articles, or the times between accidents of automobiles insured by a certain company. We will also consider games associated with a general renewal process where $\{X_i\}$ is preceded by another random variable X_0 which is independent of $\{X_i\}$ and may have a different distribution. Since the discrete case has been fully dealt with by Feller [1] and Robbins [3], the emphasis in this paper will be on the continuous case. However, the results will be presented in a general form which will include all such F(x) which do not have a jump at zero.
- 2. Fixed-time games. Let us consider the following game called $\mathfrak G$. The game starts at t=0 when an event has just occurred. At the kth occurrence of the event $\mathfrak E$, player A receives an amount $c(X_k)$ and pays an amount a_k . Here c(t) is a given function which vanishes for t<0, and a_i is a sequence of constants. For example, A may be a buyer of certain articles. The benefit which A derives from the kth article is a function $c(X_k)$ of its lifetime, X_k , and he pays the price a_k for the purchase. An insurance company A pays an amount a_k at the occurrence of kth death (life insurance) or kth accident (automobile or other accident insurance) and receives premiums and interests which are a function of time between such occurrences.

Let

T(t) = total amount received by A in (0, t];

(1)
$$T_i(t) = \text{total amount received by } A \text{ in } (X_i, X_i + t];$$

U(t) = total amount paid by A in (0, t].

Obviously the $T_i(t)$ have the same distribution as T(t). If ET(t) = EU(t) for all $0 \le t \le t_0$, we shall say that \mathfrak{g} is fair for $[0, t_0]$; if $t_0 = \infty$, we shall say that \mathfrak{g} is fair.

To avoid triviality we shall assume that c(t) is not a null function. c(t) will be called *realistic* if it is non-negative and finite for finite t. The game g will be called realizable if, given a realistic c(t), a sequence $\{a_i\}$ of non-negative constants exists for which g is fair.

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If f(t) is a function of the real variable t defined on the interval $0 \le t < \infty$, and of bounded variation in the interval $a \le t \le b$ for every positive a and every positive b, write

(2)
$$f^*(s) = \int_{0+}^{\infty} e^{-st} df(t) = \lim_{\substack{a \to 0 \\ b \to a}} \int_{a}^{b} e^{-st} df(t)$$

when this limit exists. $f^*(s)$ will be called the Laplace-Stieltjes transform (LST) [5, ch: II] of f(t).

Define

$$G(t) = \int_0^t c(x) dF(x), \quad \text{if } t \ge 0; \quad 0, \quad \text{if } t < 0;$$

and write $\bar{T}(t) = ET(t)$, and so on. Assuming that $G^*(s)$ exists for some $s = s_0$, where $s_0 = \sigma_0 + i\tau_0$, by arguments similar to those of Robbins [3, p. 190], we find

(3)
$$\bar{T}^*(s) = G^*(s)/[1 - F^*(s)].$$

 $\bar{T}^*(s)$ is analytic in the half plane $\sigma > c = \max(0, \sigma_0)$. Similarly, formally introducing the series

$$A(y) = \sum_{n=1}^{\infty} a_n y^n,$$

we obtain

(5)
$$\bar{U}^*(s) = A(F^*(s)).$$

Hence the following theorem.

THEOREM 1. 9 is fair if $G^*(s_0)$ exists, A(y) converges for $|y| < F^*(c)$, and

(6)
$$A(F^*(s)) = G^*(s)/[1 - F^*(s)].$$

Further, if g is fair and $G^*(s_0)$ exists, then A(y) converges for $|y| < F^*(c)$ and (6) holds.

Since, by assumption, F(x) does not have a jump at zero, for real s, $F^*(s)$ is a monotone decreasing function on $[0, \infty]$ to [0, 1]. It therefore has a unique inverse F^{*-1} on [0, 1] to $[0, \infty]$ which is a monotone decreasing function. Let

(7)
$$K(y) = G^*(F^{*-1}(y)), \qquad 0 \le y \le F^*(c).$$

The condition (6) then implies that

(8)
$$A(y) = K(y)(1-y)^{-1}.$$

If the Maclaurin's series expansion of K(y) is given by

(9)
$$K(y) = \sum_{n=0}^{\infty} b_n y^n,$$

We must have $b_0 = 0$, as $K(0) = G^*(\infty) = 0$, and the series converges for $0 \le y < F^*(c)$, as $G^*(s)$ is analytic for $\sigma > c$. Thus from (1)-(9) we have the following corollary.

COROLLARY 1. For any reward function c(t) for which the LST of G(t) exist, there is a unique sequence of fees $\{a_i\}$ which makes G(t) fair. The $\{a_i\}$ are given by

$$a_n = \sum_{j=1}^n b_j, \qquad n = 1, 2, \cdots.$$

Furthermore, if c(t) is realistic, and $\sum_{j=1}^{n} b_j \ge 0$, for $n = 1, 2, \dots$, the game is realizable.

REMARK 1. If the LST of G(t) does not exist for any finite s, we can still find a sequence $\{a_i\}$ which will make G fair for $[0, t_0]$, where t_0 is an arbitrary positive number. We simply set dT(t) = 0 if $t \ge t_0$, and the theorem applies with $G^*(s)$ replaced by

$$G_0^*(s) = \int_{0^+}^{t_0} e^{-st} c(t) dF(t).$$

The analogue of Robbins' Corollary 2 [3, p. 192] is COROLLARY 3. G is fair if and only if

$$\int_0^t c(x) dF(x) = \sum_{n=1}^\infty a_n H_n(t), \qquad 0 \le t \le \infty,$$

where $H_n(t) = F_n(t) - F_{n+1}(t)$ is the probability of exactly n occurrences in the interval (0, t].

3. A continuous version of the Petersburg game. A continuous analogue of the Petersburg game [1, ch. 10] may be described as follows. We consider a Poisson process where the events are occurring at the average rate of λ , $0 < \lambda < \infty$, per unit time, and at the kth occurrence A receives an amount exp (νX_k) . We have

(10)
$$dF(t) = \lambda e^{-\lambda t} dt, \qquad c(t) = e^{rt}, \qquad \text{if } t \ge 0,$$

and zero if t < 0. We note that $Ec(X_k) = \infty$ if $\nu \ge \lambda$. A straightforward calculation gives $A(y) = y(1 - \nu y/\lambda)^{-1}(1 - y)^{-1}$, i.e.,

(11)
$$a_n = \sum_{j=0}^{n-1} (\nu/\lambda)^j = \begin{cases} n, & \text{if } \nu = \lambda, \\ \frac{(\nu/\lambda)^n - 1}{(\nu/\lambda) - 1}, & \text{if } \nu \neq \lambda. \end{cases}$$

The game is realizable if $\nu \geq -\lambda$. Also,

(12)
$$\bar{T}(t) = \bar{U}(t) = \lambda t + \frac{1}{2}\lambda^2 t^2, \qquad \text{if } \nu = \lambda,$$

$$= \lambda \nu (\nu - \lambda)^{-2} [e^{(\nu - \lambda)t} - 1] - \lambda^2 (\nu - \lambda)^{-1} t, \qquad \text{if } \nu \neq \lambda.$$

It is interesting to compare this game with the classical type of game associated with a Poisson process. Here, F(t) and c(t) are given by (10) and the

game is played until the event occurs n times. The interesting case, of course, is when $Ec(X_k) = \infty$, i.e., when $\nu \ge \lambda$. The problem is to find the total fee, e_n , to be paid by A, so that

(13)
$$\operatorname{plim}_{n\to\infty} \{ [c(X_1) + \cdots + c(X_n)] / e_n \} = 1.$$

Writing $a = \lambda/\nu$, and assuming $0 < a \le 1$, we have the common distribution of $c(X_i)$

(14)
$$dP(t) = at^{-a-1} dt, \quad \text{if } 1 \le t \le \infty; 0, \text{ if } t < 1.$$

Let S_n denote the numerator in (13). Applying Theorem 3 of Gnedenko and Kolmogorov [2, p. 57], we have (13), if and only if,

$$\lim_{n\to\infty} |\phi_n(t)| = 1$$

uniformly in every finite interval of t. Here $\phi_n(t)$ is the characteristic function of $S_n/e_n - 1$ and is given by $\phi_n(t) = \{P^*(-it/e_n)\}^n e^{-it}$.

A series expansion of $\phi_n(t)$ shows that no solution exists for 0 < a < 1; and for a = 1

$$e_n = n \log_e n$$

satisfies the condition (15). This compares with Feller's [1, p. 237] solution of the Petersburg paradox, where he obtains the solution $e_n = n \log_2 n$, when X_1 is integer valued, $\Pr(X_1 = n) = 2^{-n}$, $n = 1, 2, \dots$, and $c(n) = 2^n$.

4. Fixed-time games with a general renewal process. We consider a general renewal process [4, p. 245] $\{X_0, X_1, X_2, \cdots\}$, where $\{X_i\}_{i=1}^{\infty}$ is a renewal process which is preceded by another non-negative random variable X_0 which is independent of $\{X_i\}$, and has cdf B(x). For example, the game may be started at some arbitrary time origin which does not necessarily coincide with the occurrence of an initial event.

Let $F_n(x)$ denote the cdf of $X_1 + \cdots + X_n$ and $K_n(x)$ that of

$$X_0 + X_1 + \cdots + X_{n-1}$$
.

Then

(16)
$$K_n(x) = B * F_{n-1}(x)$$

where * denotes the convolution of two cdfs.

In addition to the random variables introduced in (1), introduce $T_0(t)$ obtained from $T_i(t)$ by setting i = 0. Then $T_0(t)$, $T_1(t)$, \cdots , have identical distributions, but T(t) may have a different one.

In this case, we obtain

(17)
$$d\bar{T}(t) = \int_0^t d\bar{T}_0(t-u) dB(u) + c(t) dB(t),$$
$$d\bar{T}_0(t) = \int_0^t d\bar{T}_0(t-u) dF(u) + c(t) dF(t).$$

Thus writing

$$dG_1(t) = c(t) dB(t),$$

and assuming that $G_1^*(s)$ and $G_1^*(s)$ both exist for some $s = s_0$, we find

(19)
$$\tilde{T}^*(s) = G^*(s)B^*(s)/[1 - F^*(s)] + G_1^*(s).$$

Also

(20)
$$\bar{U}^*(s) = B^*(s)A(F^*(s))/F^*(s).$$

Thus, if G is associated with this process, we have the following theorem. Theorem 2. G is fair if $G^*(s_0)$ and $G_1^*(s_0)$ exist, A(y) converges for

$$|y| < F^*(c),$$

where $c = \max(0, \sigma_0)$, and if

$$(21) A(F^*(s)) = G^*(s)F^*(s)/[1 - F^*(s)] + G_1^*(s)F^*(s)/B^*(s).$$

Further, if g is fair and $G^*(s_0)$ and $G_1^*(s_0)$ exist, then A(y) converges for

$$|y| < F^*(c)$$

and (21) holds.

If we set $B^* = F^*$ so that $G_1^* = G^*$, we have Theorem 1. Hence, Theorem 1 may be considered a corollary of Theorem 2.

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