

## PROJECTION WITH THE WRONG INNER PRODUCT AND ITS APPLICATION TO REGRESSION WITH CORRELATED ERRORS AND LINEAR FILTERING OF TIME SERIES

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**1. Introduction.** In many places in statistics one wants to calculate the orthogonal projection  $Px$  of some vector  $x$  on a subspace  $\mathcal{P}$ . Oftentimes the inner product function is specified by the unknown covariances  $C$  of a set of random variables. The usual procedure is to estimate  $C$  by  $C^*$  and approximate  $Px$  by  $P^*x$ , the orthogonal projection with respect to  $C^*$ ; that is,  $x$  is projected on  $\mathcal{P}$  using a wrong inner product. There is, therefore, interest in knowing when  $P^*$  will be a good approximation of  $P$ .

In Section 2, the question of calculating orthogonal projections with the wrong inner product in a general Hilbert space is investigated. The results are then applied to the problem of regression with correlated errors in Section 3 and to linear filtering operations on multi-channel, wide-sense stationary, stochastic processes in Section 4.

**2. Projection with the wrong inner product in a general Hilbert space.** Let  $\mathcal{H}$  be a Hilbert space with inner product  $(\cdot, \cdot)$  and norm  $\|\cdot\| = (\cdot, \cdot)^{\frac{1}{2}}$ . Let  $[\cdot, \cdot]$ , which will be thought of as the wrong inner product, be a bilinear functional with the following properties:

- (1)  $[\cdot, \cdot]$  is defined on  $\mathcal{D} \times \mathcal{D}$  where  $\mathcal{D}$  is a linear subset of  $\mathcal{H}$  whose closure is  $\mathcal{H}$ ,  $[x, y] = \overline{[y, x]}$ , and for fixed  $x$  the linear functional  $[x, \cdot]$  on  $\mathcal{D}$  is bounded.
- (2) If  $z_n$  is a sequence of vectors in  $\mathcal{H}$  such that  $z_n \rightarrow z$  and  $z_n$  is a  $[\cdot, \cdot]$  Cauchy sequence, that is  $[z_n - z_m, z_n - z_m] \rightarrow 0$  as  $n, m \rightarrow \infty$ , then  $z \in \mathcal{D}$ .

In (1), it has not been assumed that  $[\cdot, \cdot]$  is a bounded bilinear functional. Assumption (2) has been made to ensure that  $[\cdot, \cdot]$  is defined everywhere it is possible to do so and maintain the properties in (1).

From (1), for fixed  $x \in \mathcal{D}$ , the definition of the linear functional  $[x, \cdot]$  may be extended boundedly to all of  $\mathcal{H}$ . From the Riesz Representation Theorem (Halmos, 1957, page 31) there exists  $y \in \mathcal{H}$  such that  $[x, \cdot] = (y, \cdot)$ . Let  $B$  be the mapping defined by  $Bx = y$ .  $B$  is a linear and self-adjoint, and  $B$  is bounded if and only if  $[\cdot, \cdot]$  is a bounded bilinear functional.

Let  $\mathcal{P}$  be a subspace of  $\mathcal{H}$  and  $P$  the orthogonal projection operator onto  $\mathcal{P}$ . Let  $\mathcal{P}^* = \mathcal{P} \cap \mathcal{D}$  and let  $\mathcal{Q}^*$  be the set of all  $x \in \mathcal{D}$  such that  $[x, y] = 0$  for all

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$y \in \mathcal{P}^*$ . Let  $\mathcal{D}^*$  be the set of all  $x \in \mathcal{D}$  such that  $x = p + q$  with  $p \in \mathcal{P}^*$  and  $q \in \mathcal{Q}^*$ . This decomposition of  $x$  will be unique if and only if  $[\cdot, \cdot]$  is positive definite on  $\mathcal{P}^* \times \mathcal{P}^*$ . For, suppose  $[\cdot, \cdot]$  is not positive definite, that is there is a  $z \in \mathcal{P}^*$  with  $z \neq 0$  such that  $[z, z] = 0$ . Since the Cauchy–Schwartz inequality holds for  $[\cdot, \cdot]$  (Helmberg, 1969, page 10),  $[z, y] = 0$  for all  $y \in \mathcal{P}^*$ . Thus  $z \in \mathcal{Q}^*$  and  $x = (p + z) + (q - z)$  so that the decomposition is not unique. Suppose  $[\cdot, \cdot]$  is positive definite and  $x = p' + q'$  with  $p' \in \mathcal{P}^*$  and  $q' \in \mathcal{Q}^*$ . Then

$$0 = [p' + q' - p - q, p' + q' - p - q] = [p' - p, p' - p] + [q' - q, q' - q]$$

which implies  $p' = p$  and  $q' = q$ .

The operator  $P^*$ , orthogonal projection with respect to  $[\cdot, \cdot]$ , will be defined in the following manner. Let  $\mathcal{N}$  be the set of vectors  $y \in \mathcal{D}$  such that  $[y, y] = 0$ . Clearly  $\mathcal{N}$  is a linear space. Let  $z_n$  be a Cauchy sequence in  $\mathcal{N}$  and let  $z$  be the limit of  $z_n$ . From (2),  $z \in \mathcal{D}$  and

$$[z, z] = (Bz, z) = \lim_{n \rightarrow \infty} (Bz, z_n) = \lim_{n \rightarrow \infty} (z, Bz_n) = 0.$$

Thus  $\mathcal{N}$  is a subspace. Let  $N$  be the orthogonal projection operator onto  $\mathcal{N}$ . Let  $x \in \mathcal{D}^*$  have the decomposition  $x = p + q$ . Then define  $P^*x$  to be  $p - Np$ . We could have taken  $P^*x$  to be any vector in  $p + \mathcal{N}$  or even all of them. However, in the special cases of Sections 3 and 4, it is seen that  $p - Np$ , the vector in  $p + \mathcal{N}$  with the smallest norm, is a natural assignment.

(3) THEOREM. Suppose  $\overline{\mathcal{D}^*} = \mathcal{D}$ . Then  $P = P^*$  on  $\mathcal{D}$  if and only if  $\overline{B\mathcal{P}^*} = \mathcal{P}$ .

PROOF. Suppose  $P = P^*$  on  $\mathcal{D}$ . For all  $x \in \mathcal{D}$  and  $y \in \mathcal{P}^*$

$$(x, By) = (Bx, y) = [x, y] = [Px, y] = (BPx, y) = (Px, By).$$

Since  $P$  and  $(\cdot, By)$  are continuous and  $\mathcal{D}$  is dense in  $\mathcal{H}$ ,  $(x, By) = (Px, By)$  for all  $x \in \mathcal{H}$ . In particular it holds for all vectors in the orthogonal complement of  $\mathcal{P}$  in  $\mathcal{H}$ , which implies  $By \in \mathcal{P}$ ; hence  $B\mathcal{P}^* \subset \mathcal{P}$ . To show equality, we first prove  $\overline{\mathcal{P}^*} = \mathcal{P}$  and  $B$  is one-to-one on  $\mathcal{P}^*$ . Let  $y \in \mathcal{P}$ . Since  $\mathcal{D}$  is dense in  $\mathcal{H}$ , there is a sequence  $y_n \in \mathcal{D}$  with  $y_n \rightarrow y$ ; thus  $Py_n \rightarrow Py = y$ . But  $Py_n = P^*y_n \in \mathcal{P}^*$  so that  $\overline{\mathcal{P}^*} = \mathcal{P}$ . Suppose that for some  $x \in \mathcal{P}^*$ ,  $Bx = 0$ . Then  $[x, y] = 0$  for all  $y \in \mathcal{P}^*$  which implies  $P^*x = 0$ . But  $P^*x = Px = x$ . Thus  $B$  is one-to-one on  $\mathcal{P}^*$ . The facts that  $\overline{\mathcal{P}^*} = \mathcal{P}$  and  $B$  is a one-to-one self-adjoint transformation on  $\mathcal{P}^*$  with  $B\mathcal{P}^* \subset \mathcal{P}$  together with (Helmberg, 1969, Theorem 6, page 121), imply  $\overline{B\mathcal{P}^*} = \mathcal{P}$ .

Suppose  $\overline{B\mathcal{P}^*} = \mathcal{P}$ . Let  $x \in \mathcal{D}$  then for all  $y \in \mathcal{P}^*$ ,  $(x, By) = (Bx, y) = [x, y] = [P^*x, y] = (BP^*x, y) = (P^*x, By)$ . Thus  $(x, z) = (P^*x, z)$  for all  $z \in B\mathcal{P}^*$ . From the hypothesis, this last equality may be extended to all  $z \in \mathcal{P}$ , which implies  $P^*x = Px$ .

The following inequality is due to Kantorovich (1948, page 142).

(4) LEMMA. Let  $0 < a_1 \leq a_2 \cdots \leq a_n < \infty$  and  $c_i \geq 0$  for  $i = 1, \dots, n$ . Then

$$\sum_{i=1}^n c_i a_i \sum_{i=1}^n c_i a_i^{-1} \leq \frac{(a_1 + a_n)^2}{4a_1 a_n} \sum_{i=1}^n c_i,$$

and equality holds for the proper choice of  $c_i$ .

The inequality of the following theorem is an addition to the class of relatives of the Kantorovich Inequality in a Hilbert space (cf. (Mond, 1966)).

(5) THEOREM. Let  $x \in \mathcal{D}^*$ . Let

$$\alpha_1 = \inf \left\{ \frac{[z, z]}{(z, z)} : z \in \mathcal{D} \right\} \quad \text{and} \quad \alpha_2 = \sup \left\{ \frac{[z, z]}{(z, z)} : z \in \mathcal{D} \right\}.$$

Then

$$\frac{\|x - P^*x\|^2}{\|x - Px\|^2} \leq \frac{(\alpha_1 + \alpha_2)^2}{4\alpha_1\alpha_2},$$

where the right side of the inequality is assigned the value  $\infty$  if either  $\alpha_1 = 0$  or  $\alpha_2 = \infty$ ; and the left side of the inequality is assigned the value 1 if both numerator and denominator are 0, and the value  $\infty$  if the denominator is 0 and the numerator is not. The inequality is best in the sense that the left side can be made arbitrarily close to the right by the proper choice of  $\mathcal{P}$  and  $x$ .

PROOF. If  $\alpha_1 = 0$  or  $\alpha_2 = \infty$  the inequality is clearly true. Thus assume  $\alpha_1 > 0$  and  $\alpha_2 < \infty$ . In this case  $\mathcal{D}^* = \mathcal{D} = \mathcal{H}$  since  $[\cdot, \cdot]$  is defined on all of  $\mathcal{H} \times \mathcal{H}$  in view of (2), and  $[\cdot, \cdot]$  together with the elements  $\mathcal{H}$  is a Hilbert space. If  $P^*x = Px$  the inequality is clearly true since the right side is  $\geq 1$ . Thus assume  $P^*x \neq Px$ . Let  $z_1 = x - Px$  and  $z_2 = P^*z_1$ . Note that neither  $z_1$  nor  $z_2$  is 0 since  $P^*x \neq Px$ . Let  $R$  be the orthogonal projection operator onto the space spanned by  $z_2$ , and  $R^*$  the orthogonal projection operator with respect to  $[\cdot, \cdot]$ . Since  $P^*z_1 = z_2$  and  $P^*Px = Px$  we have

$$(6) \quad z_1 - R^*z_1 = x - P^*x.$$

Since  $(z_1, z_2) = 0$ , we have  $Rz_1 = 0$  and thus

$$(7) \quad z_1 - Rz_1 = x - Px.$$

From (6) and (7) there is no loss of generality in assuming  $\mathcal{P}$  is one-dimensional,  $x$  is orthogonal to  $\mathcal{P}$ , and  $\|x\| = 1$ . Let  $y$  be a basis for  $\mathcal{P}$ , where  $\|y\| = 1$ . Let  $\mathcal{S}$  be the space spanned by  $x$  and  $y$ . Then  $Px = 0$  and  $P^*x = ([x, y]/[y, y])y$  and

$$\alpha = \frac{\|x - P^*x\|^2}{\|x - Px\|^2} = 1 + \frac{[x, y]^2}{[y, y]^2}.$$

Since the quadratic forms  $Q_1(v) = (v, v)$  and  $Q_2(v) = [v, v]$  for  $v \in \mathcal{S}$  are non-singular with respect to each other, we may coordinatize  $\mathcal{S}$  so that if  $v$  corresponds to the two-tuple  $v_1, v_2$  then

$$Q_1(v) = |v_1|^2 + |v_2|^2 \quad \text{and} \quad Q_2(v) = \beta_1|v_1|^2 + \beta_2|v_2|^2,$$

where  $\beta_1 = \min \{Q_1(v)/Q_2(v) : v \in \mathcal{S}\}$  and  $\beta_2$  is defined similarly with min replaced by max. Let  $x$  correspond to  $x_1, x_2$  and  $y$  to  $y_1, y_2$  then

$$\alpha = 1 + \left[ \frac{\beta_1 x_1 \bar{y}_1 + \beta_2 x_2 \bar{y}_2}{\beta_1 |y_1|^2 + \beta_2 |y_2|^2} \right]^2.$$

Since  $\|x\| = \|y\| = 1$  and  $(x, y) = 0$ , we have  $|x_1| = |y_2|$  which implies

$$\alpha = \frac{\beta_1^2 |y_1|^2 + \beta_2^2 |y_2|^2}{(\beta_1 |y_1|^2 + \beta_2 |y_2|^2)^2}.$$

Letting  $a_j = \beta_j$  and  $c_j = \beta_j |y_j|^2 / (\beta_1 |y_1|^2 + \beta_2 |y_2|^2)$  we may apply (4) with the result that

$$(8) \quad \alpha \leq \frac{(\beta_1 + \beta_2)^2}{4\beta_1\beta_2}.$$

The inequality of the theorem now follows from this and the fact that  $\beta_1 \geq \alpha_1$  and  $\beta_2 \leq \alpha_2$ .

It will now be shown that the inequality cannot be improved. (We are still assuming  $\alpha_1 > 0$  and  $\alpha_2 < \infty$ .) For  $\alpha_1 = \alpha_2$  it is obvious. Thus assume  $\alpha_1 < \alpha_2$ . For  $\varepsilon > 0$ , let  $u, v \in \mathcal{H}$  be such that

$$\frac{[u, u]}{(u, u)} < \alpha_1 + \varepsilon \quad \text{and} \quad \frac{[v, v]}{(v, v)} > \alpha_2 - \varepsilon.$$

Clearly  $u, v$  are linearly independent for  $\varepsilon$  small enough. Now let  $\mathcal{S}$  be the space spanned by  $u$  and  $v$ . Then  $\beta_1 < \alpha_1 + \varepsilon$  and  $\beta_2 > \alpha_2 - \varepsilon$ . But from (4) equality holds in (8) for the proper choice of  $y_1, y_2$ , which shows the inequality cannot be improved for the case  $\alpha_1 > 0$  and  $\alpha_2 < \infty$ .

If  $[\cdot, \cdot]$  is positive definite on  $\mathcal{D} \times \mathcal{D}$  and either  $\alpha_1 = 0$  or  $\alpha_2 = \infty$  then by a proof analogous to that in the previous paragraph it may be shown that the inequality cannot be improved.

Suppose  $x \in \mathcal{D}$  is such that  $[x, x] = 0$  and  $x \neq 0$ . Let  $\mathcal{P}$  be the space spanned by  $x$  then  $Px = x$  and  $P^*x = 0$ . Thus the inequality cannot be improved in this case.

(9) COROLLARY.

$$\|Px - P^*x\|^2 \leq \frac{(\alpha_1 - \alpha_2)^2}{4\alpha_1\alpha_2} \|x - Px\|^2 \leq \frac{(\alpha_1 - \alpha_2)^2}{4\alpha_1\alpha_2} \|x - P^*x\|^2.$$

PROOF. The first inequality follows easily from (5) and the equality  $\|x - P^*x\|^2 = \|x - Px\|^2 + \|Px - P^*x\|^2$ . The second follows from the fact that  $\|x - Px\|^2 \leq \|x - P^*x\|^2$ .

**3. Regression with correlated errors.** Suppose  $x = x_1, \dots, x_n$  has a multivariate normal distribution with mean  $m$  and nonsingular covariance matrix  $C$ . Suppose  $m$  lies in  $\mathcal{P}$ , a subspace of Euclidean  $n$ -space  $\mathcal{E}$ . The norm  $(y, y)^\dagger = \|y\| = (yC^{-1}y)^\dagger$ , where  $y \in \mathcal{E}$ , is oftentimes referred to as the Mahalanobis distance. It provides a measure of the distance of two normal populations with the same covariance matrix  $C$ , in the sense that the closer the means of the two populations under this norm, the more difficult it is to discriminate them (Rao, 1965, Section 8e).

Thus if  $\hat{m}$  and  $m^*$  are estimates of  $m$ , a measure of their proximity is  $\|\hat{m} - m^*\|$ . In fact, if  $\hat{m}$  denotes the posterior mean of  $m$  under the assumption that the prior on  $m$  is uniform, then  $\hat{m}$  is the vector in  $\mathcal{P}$  which minimizes  $\|x - p\|$  as  $p$  ranges over  $\mathcal{P}$ . ( $\hat{m}$  is also the Gauss–Markov estimate even if normality is not assumed.) That is,  $\hat{m} = Px$  where  $P$  is the orthogonal projection operator onto  $\mathcal{P}$ .

Suppose, however, that  $C$  is not known but  $C^*$  is available where  $C^*$  is a nonsingular approximation or estimate of  $C$  or perhaps just a convenient choice, as in least squares estimation. Now suppose  $m$  is estimated as above with  $C$  replaced by  $C^*$ . That is,  $[y, y] = y(C^*)^{-1}y'$  and  $m$  is estimated by  $m^* = P^*x$ , where  $P^*$  is the orthogonal projection operator on  $\mathcal{P}$  with respect to  $[\cdot, \cdot]$ .

In the notation of Section 2, let  $\mathcal{H}$  be the subspace of  $\mathcal{E}$  spanned by  $x$  and  $\mathcal{P}$ . Let  $\alpha_1$  and  $\alpha_2$  be defined as in (5). Then (5) provides a bound for  $\|x - m^*\|/\|x - \hat{m}\|$  and (9) gives a bound for  $\|\hat{m} - m^*\|$ , the Mahalanobis distance between  $\hat{m}$  and  $m^*$ . (Compare this with (Watson, 1967, page 1685).)

Instead of choosing  $\mathcal{H}$  as above we can take  $\mathcal{H} = \mathcal{E}$ . This, in general, results in less sharp bounds for a particular  $x$  and  $\mathcal{P}$ , but bounds which now hold whatever the choice of  $x$  and  $\mathcal{P}$ . From (5), these bounds are in fact best among all bounds which do not depend on  $x$  and  $\mathcal{P}$ . The operator  $B$  of Section 2 is  $C(C^*)^{-1}$  and  $\alpha_1$  and  $\alpha_2$  are the minimum and maximum eigenvalues of this operator. Clearly in this example  $\mathcal{D}^* = \mathcal{D}$  and  $C(C^*)^{-1}\mathcal{P} = \overline{C(C^*)^{-1}\mathcal{P}}$  so that from (3),  $m^* = \hat{m}$  if and only if  $C(C^*)^{-1}\mathcal{P} = \mathcal{P}$ , a result first proved by Kruskal (1968).

**4. Linear filtering operations.** The calculation of projections, in particular linear predictions, interpolations, and signal extractions, in the space spanned by a multi-channel, wide-sense stationary time series  $X_t$  has received great attention since the initial works of Wiener (1949) and Kolmogorov (1941). The theory is based on the assumption that the spectral distribution matrix of  $X_t$  is known. But it is not really known in practice, and the traditional remedy is to estimate the matrix and then calculate the projections as though the estimate were the true matrix.

Suppose  $X_t$  is an  $r$ -channel process. Let  $F(\lambda) = [F_{jk}(\lambda)]$  be the  $r \times r$  spectral distribution matrix of  $X_t$ , where  $-\infty < \lambda < \infty$  if  $X_t$  is a continuous parameter process and  $0 \leq \lambda \leq 1$  if it is discrete parameter. Let  $F^*(\lambda) = [F^*_{jk}(\lambda)]$  be an estimate of  $F(\lambda)$ . Assume  $\sum_{j=1}^r F^*_{jj}$  is absolutely continuous with respect to  $\sum_{j=1}^r F_{jj}$ .

Any projection in the Hilbert space spanned by the process can be described as a

projection in  $\mathcal{L}_2(F)$ , the Hilbert space of vector functions  $v(\lambda) = v_1(\lambda), \dots, v_r(\lambda)$  such that

$$\|v\|^2 = \sum_{j,k=1}^r \int \overline{v_j(\lambda)} v_k(\lambda) dF_{jk}(\lambda) < \infty,$$

where the range of integration is  $-\infty$  to  $\infty$  if  $X_i$  is continuous parameter and 0 to 1 if it is discrete (Rosanov, 1967, page 28). Let  $m$  be any measure such that  $\sum_{j=1}^r F_{jj}$  is absolutely continuous with respect to  $m$ . Let  $f(\lambda) = (dF/dm)(\lambda)$  and  $f^*(\lambda) = (dF^*/dm)(\lambda)$ . Then

$$\|v\|^2 = \int \overline{v(\lambda)} f(\lambda) v'(\lambda) dm(\lambda),$$

where  $v'(\lambda)$  denotes the transpose of  $v(\lambda)$ .

To apply the results of Section 2, we let  $\mathcal{H} = \mathcal{L}_2(F)$  and define  $[\cdot, \cdot]$  by

$$[u, v] = \int \overline{u(\lambda)} f^*(\lambda) v'(\lambda) dm(\lambda),$$

so that  $\mathcal{D}$  is the set of  $v \in \mathcal{L}_2(F)$  with  $[v, v] < \infty$ . Using  $F^*$  in place of  $F$  to calculate a projection is equivalent to using  $[\cdot, \cdot]$  in place of  $(\cdot, \cdot)$ . It is easily seen that assumptions (1) and (2) hold, and the operator  $B$  maps  $v$  to  $u$ , where  $u$  at the point  $\lambda$  is

$$u(\lambda) = \left[ \frac{\overline{v(\lambda)} f^*(\lambda) v'(\lambda)}{\overline{v(\lambda)} f(\lambda) v'(\lambda)} \right] v(\lambda).$$

With these definitions, the results of Section 2 may now be applied.

The values  $\alpha_1$  and  $\alpha_2$  in (5) can be written in terms of  $f$  and  $f^*$  as shown by the following theorem.

(10) THEOREM. *Let*

$$\Delta_1(\lambda) = \min \left\{ \frac{\overline{ef^*(\lambda)}c'}{\overline{ef(\lambda)}c'} : c = c_1, \dots, c_r, \text{ a complex } r\text{-tuple} \right\}$$

for all  $\lambda$ . Define  $\Delta_2(\lambda)$  similarly with  $\min$  replaced by  $\max$ . Then  $\Delta_1$  and  $\Delta_2$  are  $m$ -measurable. Let  $\zeta_1 = \text{ess}_\lambda \inf \Delta_1(\lambda)$  and  $\zeta_2 = \text{ess}_\lambda \sup \Delta_2(\lambda)$ , where  $\text{ess inf}$  and  $\text{sup}$  are with respect to  $m$ . Then  $\alpha_1 = \zeta_1$  and  $\alpha_2 = \zeta_2$ . If  $f(\lambda)$  is nonsingular,  $\Delta_1(\lambda)$  is the minimum eigenvalue of  $f^*(\lambda)f^{-1}(\lambda)$  and  $\Delta_2(\lambda)$  the maximum.

PROOF. Since the complex  $r$ -tuples with rational real and imaginary parts are dense in the space of  $r$ -tuples with norm  $(\overline{ef(\lambda)}c')^{\frac{1}{2}}$ , the minimum in the definition of  $\Delta_1(\lambda)$  may be taken over such rational  $c$ . Thus  $\Delta_1$  is  $m$ -measurable since it is the minimum of a countable number of  $m$ -measurable functions. Similarly,  $\Delta_2$  is  $m$ -measurable.

Since for each  $\lambda$ ,  $\Delta_2(\lambda)$  is the maximum eigenvalue of  $f^*(\lambda)$  with respect to  $f(\lambda)$ , there exists an eigenvector  $e(\lambda) = e_1(\lambda), \dots, e_r(\lambda)$  such that

$$e(\lambda) f^*(\lambda) = \Delta_2(\lambda) e(\lambda) f(\lambda).$$

It is easy to specify a routine for choosing  $e(\lambda)$  to ensure that  $e$  is  $m$ -measurable. Also  $e(\lambda)$  may be chosen so that  $|e_j(\lambda)| < 1$ , since a constant times an eigenvector is also an eigenvector.

Now if  $v \in \mathcal{D}$ ,

$$\begin{aligned} [v, v] &= \int \overline{v(\lambda)} f^*(\lambda) v'(\lambda) dm(\lambda) \\ &= \int \frac{\overline{v(\lambda)} f^*(\lambda) v'(\lambda)}{v(\lambda) f(\lambda) v'(\lambda)} \overline{v(\lambda)} f(\lambda) v'(\lambda) dm(\lambda) \\ &\leq \zeta_2 \|v\|^2. \end{aligned}$$

That  $\zeta_2 = \alpha_2$  will be proved by exhibiting a sequence  $v_n \in \mathcal{L}_2(F)$  such that  $\|v_n\| = 1$  and  $[v_n, v_n] \rightarrow \zeta_2$ . Define the set  $S_n$  as follows: if  $\zeta_2 = \infty$  then

$$S_n = \{\lambda: \Delta_2(\lambda) > n\};$$

if  $\zeta_2 < \infty$  then

$$S_n = \left\{ \lambda: \Delta_2(\lambda) > \zeta_2 - \frac{1}{n} \right\}.$$

There exists an  $m$ -measurable subset  $T_n$  of  $S_n$  with positive  $m$  measure such that  $T_n$  is contained in a bounded interval. Let  $T_n$  also denote the indicator function of the set  $T_n$  and define  $v_n(\lambda) = \|T_n e\|^{-1} T_n(\lambda) e(\lambda)$ .  $v_n \in \mathcal{L}_2(F)$  since  $e$  and  $T_n$  are  $m$ -measurable,  $|e_j(\lambda)| < 1$ , and  $T_n$  lies in a bounded interval. Now

$$[v_n, v_n] = \int \Delta_2(\lambda) \overline{v_n(\lambda)} f(\lambda) v_n'(\lambda) dm(\lambda)$$

so that

$$\text{ess inf}_{\lambda \in T_n} T_n(\lambda) \Delta_2(\lambda) \leq [v_n, v_n] \leq \text{ess sup}_{\lambda \in T_n} T_n(\lambda) \Delta_2(\lambda) \leq \zeta_2.$$

Since the left side tends to  $\zeta_2$ , so does  $[v_n, v_n]$ .

A similar proof shows  $\alpha_1 = \zeta_1$ . The final statement of the theorem is a well-known fact (Rao, 1965, page 15).

A special case of  $F$  and  $F^*$  deserves comment because of its frequent occurrence in practice. Suppose that  $X_t$  is a discrete parameter process,  $m$  is Lebesgue measure, and

$$p_1 I \leq f(\lambda) \leq p_2 I$$

for all  $\lambda$  where  $I$  is the  $r \times r$  identity matrix and  $p_1, p_2$  are positive constants. A common way of obtaining  $f^*$  is to assume  $X_t$  is a  $k$ th order autoregression

$$A_0 X_t + A_1 X_{t-1} + \cdots + A_k X_{t-k} = W_t$$

where the  $k \times k$  matrices  $A_j$  are such that the roots of the polynomial in  $z$

$$\det [A_0 + A_1 z + \cdots + A_k z^k]$$

lie outside the unit circle, and  $W_t$  is  $r$ -channel white noise; that is,  $EW_{t_1}W'_{t_2} = I$  for  $t_1 = t_2$  and 0 for  $t_1 \neq t_2$ . For such a model

$$f(\lambda) = (A_0 + A_1 e^{2\pi i \lambda} + \dots + A_k e^{2\pi i k \lambda})^{-1} (A_0' + A_1' e^{-2\pi i \lambda} + \dots + A_k' e^{-2\pi i k \lambda})^{-1}$$

(cf. Whittle, 1953, page 125). The parameters  $A_0, \dots, A_k$  are estimated by  $A_0^*, \dots, A_k^*$  and  $f^*(\lambda)$  is formed by replacing  $A_j$  by  $A_j^*$  in the above expression for  $f(\lambda)$ .

The adequacy of the autoregressive model is then checked by calculating the fitted residuals

$$W_t^* = A_0^* X_t + A_1^* X_{t-1} + \dots + A_k^* X_{t-k}$$

and checking them for the white noise assumption. If  $f^*$  is to be used for calculating projections then from (5) and (10) the model is adequate if

$$\alpha = \frac{(\alpha_1 + \alpha_2)^2}{4\alpha_1\alpha_2} = \frac{(\zeta_1 + \zeta_2)^2}{4\zeta_1\zeta_2}$$

is nearly 1. Let  $h(\lambda)$  be the derivative of the spectral distribution matrix of  $W_t^*$  with respect to Lebesgue measure. It will be shown below that if  $\Phi_1(\lambda)$  is the minimum eigenvalue of  $h(\lambda)$  and  $\Phi_2(\lambda)$  the maximum, then  $\Phi_1(\lambda) = \Delta_2^{-1}(\lambda)$  and  $\Phi_2(\lambda) = \Delta_1^{-1}(\lambda)$ . Thus letting  $\phi_1 = \text{ess}_\lambda \inf \Phi_1(\lambda)$  and  $\phi_2 = \text{ess}_\lambda \sup \Phi_2(\lambda)$  we have  $\phi_1 = \zeta_2^{-1}$  and  $\phi_2 = \zeta_1^{-1}$  so that

$$\alpha = \frac{(\phi_1 + \phi_2)^2}{4\phi_1\phi_2}.$$

Thus  $\alpha$  can be estimated by calculating the fitted residuals and estimating successively  $h(\lambda)$ ,  $\Phi_1(\lambda)$  and  $\Phi_2(\lambda)$ ,  $\phi_1$  and  $\phi_2$  and finally  $\alpha$ .

To see  $\Phi_1(\lambda) = \Delta_2^{-1}(\lambda)$ , let

$$A^*(\lambda) = A_0^* + A_1^* e^{2\pi i \lambda} + \dots + A_k^* e^{2\pi i k \lambda}.$$

It is easily seen that

$$h(\lambda) = A^*(\lambda) f(\lambda) A^{*'}(-\lambda).$$

Now  $\Phi_1^{-1}(\lambda)$  is the maximum eigenvalue of

$$h^{-1}(\lambda) = (A^{*'}(-\lambda))^{-1} f^{-1}(\lambda) (A^*(\lambda))^{-1},$$

which has the same maximum eigenvalue as

$$(A^*(\lambda))^{-1} (A^{*'}(-\lambda))^{-1} f^{-1}(\lambda) = f^*(\lambda) f^{-1}(\lambda),$$

which from (10) is  $\Delta_2(\lambda)$ . A similar argument shows  $\Phi_2(\lambda) = \Delta_1^{-1}(\lambda)$ .

As in Section 3, we can define  $\mathcal{H}$  to be the space spanned by  $x$  and  $\mathcal{P}$  rather than all of  $\mathcal{L}_2(F)$ . The new  $\alpha_1$  and  $\alpha_2$  will give better bounds, but bounds which now depend on  $x$  and  $\mathcal{P}$  and which might not be easily expressed in terms of  $f$  and  $f^*$ .



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## REFERENCES

- [1] HALMOS, P. R. (1957). *Hilbert Space*. Chelsea, New York.
- [2] HELMBERG, GILBERT (1969). *Spectral Theory in Hilbert Space*. Wiley, New York.
- [3] KANTOROVICH, L. W. (1948). Functional analysis and applied mathematics (in Russian). *Uspehi Mat. Nauk.* **3(6)** 89–189. (English translation, National Bureau of Standards, Los Angeles, 1952.)
- [4] KOLMOGOROV, A. N. (1941). Interpolation and extrapolation of stationary random sequences (Russian). *Izv. Akad. Nauk. SSSR. Ser. Mat.* **9** 3–14.
- [5] KRUSKAL, W. (1968). When are Gauss–Markov and least squares estimators identical? A coordinate-free approach. *Ann. Math. Statist.* **39** 70–75.
- [6] MOND, B. (1966). An inequality for operators in a Hilbert space. *Pacific J. Math.* **18** 161–163.
- [7] RAO, C. R. (1965). *Linear Statistical Inference*. Wiley, New York.
- [8] WATSON, G. S. (1967). Linear least squares regression. *Ann. Math. Statist.* **38** 1679–1699.
- [9] WHITTLE, P. (1953). The analysis of multiple stationary time series. *J. Roy. Statist. Soc. Ser. B* **15** 125–139.
- [10] WEINER, N. (1949). *Extrapolation, Interpolation, and Smoothing of Stationary Time Series*. M.I.T. Press, Cambridge.