

CORRECTION NOTE

CORRECTIONS TO

“STRONG CONSISTENCY OF CERTAIN SEQUENTIAL ESTIMATORS”

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In the above paper (*Ann. Math. Statist.* **40** 1492-1495), the main consistency result (Theorem 3.4) uses Theorem 2.7 to justify the assertion that $\mathcal{C}^* = \downarrow \lim_i \mathcal{C}_{t_i} \equiv \{\emptyset, \Omega\}$. However, Theorem 2.7 as stated is incorrect: $\mathcal{C}^* = \downarrow \lim \mathcal{C}_{N_i}$ can properly contain \mathcal{C}_{N_∞} . Proposition 1 below illustrates an instance of this. To get around this difficulty, Proposition 2 below gives a sufficient condition that $\mathcal{C}^* \equiv \{\emptyset, \Omega\}$. The condition is seen to hold for a large variety of examples, including those considered in the paper.

PROPOSITION 1. *Let N be a random index and let $N_n = \max \{N, n\}$, $n = 1, 2, \dots$. Then for any decreasing sequence $\{\mathcal{C}_i\}$ with $\mathcal{C}_\infty = \downarrow \lim_i \mathcal{C}_i$, $\{N_n: 1 \leq n \leq \infty\}$ is C -ordered and $\mathcal{C}_{N_n} \downarrow \mathcal{B}(\mathcal{C}_\infty, (N < \infty))$, the σ -field generated by \mathcal{C}_∞ and the set $(N < \infty)$.*

PROOF. If $C \in \mathcal{C}_{N_n}$,

$$(1) \quad C = \sum_{k=1}^{\infty} C_{kn}(N_n = k) \cup C_{\infty n}(N_n = \infty) \\
 = C_n(N \leq n) \cup \sum_{n+1}^{\infty} C_{kn}(N = k) \cup C_\infty(N = \infty),$$

where $C_{kn} \in \mathcal{C}_k$, $1 \leq k \leq \infty$ and we write $C_{nn} = C_n$ and $C_{\infty n} = C_\infty$ (note that the latter set does not depend on n). Clearly any set in $\mathcal{C}_{N_{n+1}}$ is of this form (with $C_n = C_{n,n+1} = C_{n+1}$), so $\mathcal{C}_{N_{n+1}} \subset \mathcal{C}_{N_n}$. Thus $\{N_n: 1 \leq n \leq \infty\}$ is C -ordered. Let $C^* = \downarrow \lim_n \mathcal{C}_{N_n}$. We note that for all n , $\mathcal{C}_\infty \subset \mathcal{C}_{N_n}$ and $(N < \infty) = (N_n < \infty) \in \mathcal{C}_{N_n}$. Hence $\mathcal{C}^* \supset \mathcal{B}(\mathcal{C}_\infty, (N < \infty))$. This already contradicts Theorem 2.7, which asserts in this case that $\mathcal{C}^* = \mathcal{C}_\infty$.

To establish the reverse inclusion for \mathcal{C}^* , choose $C \in \mathcal{C}^*$. Then for all n , C has a representation as in (1). Fix m . For $n > m$, it follows from (1) that $C(N \leq m) = C_m(N \leq m) = C_n(N \leq m)$. Thus $\lim 1_{C_n} = 1$ on $(N \leq m)$. Let $C_{\infty*} = \limsup C_n \in \mathcal{C}_\infty$. Then $C(N \leq m) = C_{\infty*}(N \leq m)$. Letting $m \rightarrow \infty$ then shows that $C = C_{\infty*}(N < \infty) \cup C_\infty(N = \infty)$. Thus $\mathcal{C}^* \subset \mathcal{B}(\mathcal{C}_\infty, (N < \infty))$. \square

We note that if N is a stopping time in Proposition 1, then so are the N_n . Thus Theorem 2.7 is not even true in general for C -ordered stopping times. If one adds the hypothesis $N_\infty < \infty$ with probability one, Theorem 2.7 is true and the proof given is valid. (Whether the theorem remains true under the weaker hypothesis: for all i , $N_i < \infty$ with probability one, is not known. Note that in Proposition 1, $N_n < \infty$ with probability one if and only if $N < \infty$ with probability one and then $\mathcal{C}^* \equiv \mathcal{C}_\infty$.) Of course the case of primary interest in the paper is $N_\infty \equiv \infty$, so some suitable alternative to Theorem 2.7 seems necessary.

In the sequel we assume the structure of Section 3. Moreover, all random indices are assumed to be \mathcal{A}_∞ measurable. Let Σ_n denote the permutation group on the first n positive integers and let $\Sigma = \bigcup_n \Sigma_n$ be all finite permutations of the positive integers. An element σ in Σ acts on (x_1, x_2, \dots) by sending it into $(x_{\sigma_1}, x_{\sigma_2}, \dots)$. For a random index N , we let $\sigma N = N^\circ \sigma^{-1}$. Note that $\sigma(N = k) = (\sigma N = k)$ and, since the $\{x_i\}$ are i.i.d., N and σN are equidistributed.

DEFINITION. A random index N is called tail-symmetric if for every σ in Σ there is an integer p so that N and σN coincide on $(N > p, \sigma N > p)$. That is, for all $k > p$, $(N = k)(N > p, \sigma N > p) = (\sigma N = k)(N > p, \sigma N > p)$, or

$$(2) \quad \forall k > p, (N = k, \sigma N > p) = (\sigma N = k, N > p).$$

Of course such a p , if it exists, is not unique. Then we denote by $p(\sigma, N)$ the least positive integer p for which (2) holds. A collection $\{N_i\}$ of random indices is called homogeneously tail-symmetric if each is tail-symmetric and for every σ in Σ , $\sup_i p(\sigma, N_i) < \infty$.

PROPOSITION 2. Suppose $N_1 \leq N_2 \leq \dots$ are C -ordered and $\lim N_i = +\infty$ with probability one. If, in addition, the $\{N_i\}$ are homogeneously tail-symmetric, then $\mathcal{C}_{N_i} \downarrow \mathcal{C}^* \equiv \{\emptyset, \Omega\}$.

PROOF. The C -ordering implies that \mathcal{C}_{N_i} decreases, to \mathcal{C}^* , say. Choose $C \in \mathcal{C}^*$. Since $C \in \mathcal{C}_{N_i}$, $C = \sum C_k(N_i = k)$, where $C_k \in \mathcal{C}_k$, $1 \leq k \leq \infty$. Choose an integer m , an element σ of Σ_m and let $n = \max\{m, \sup_i p(\sigma, N_i)\}$. Then $C(N_i > n) = \sum_{k>n} C_k(N_i = k)$, where $C_k \in \mathcal{C}_k$, for all $k > n$. Thus $\sigma\{C(N_i > n)\} = \sigma C(\sigma N_i > n) = \sum_{k>n} C_k(\sigma N_i = k)$, since for $k > n \geq m$ the sets in \mathcal{C}_k are Σ_m -invariant. Thus

$$(3) \quad \begin{aligned} \sigma C(\sigma N_i > n)(N_i > n) &= \sum_{k>n} C_k(\sigma N_i = k)(N_i > n) \\ &= \sum_{k>n} C_k(N_i = k)(\sigma N_i > n) = C(N_i > n)(\sigma N_i > n), \end{aligned}$$

where the second equality in (3) follows from homogeneous tail-symmetry. Since N_i and σN_i are equidistributed, it follows that $\sigma N_i \uparrow +\infty$ with probability one. Letting $i \rightarrow \infty$ in (3) then shows that $\sigma C = C$ with probability one. Since $\sigma \in \Sigma$ is arbitrary, that $\mathcal{C}^* \equiv \{\emptyset, \Omega\}$ follows from the Hewitt–Savage 0–1 law. \square

REMARK. Since the \mathcal{C}_{N_i} decrease, it is enough for the conclusion of Proposition 2 to hold that some infinite subset of $\{N_i\}$ be homogeneously tail-symmetric.

Theorem 3.4 is then correct if one adds the requirement that the stopping times $\{t_i\}$ be homogeneously tail-symmetric. We show next that the structure assumed in Theorem 3.5 assures this, in addition to the C -ordering. Specifically, we isolate the following sufficient condition that a random index N be tail-symmetric.

PROPOSITION 3. Suppose there are sets $\{D_n\}$, $D_n \in \mathcal{B}(z_n, x_{n+1}, \dots)$, so that for all n , $(N > n) = \bigcap_1^n D_k$. Then N is tail-symmetric and if $\sigma \in \Sigma_n$, $p(\sigma, N) \leq n$.

PROOF. Choose $n, \sigma \in \Sigma_n$ and $k > n$. We note that $(N = k) = D_k^c \bigcap_1^{k-1} D_i$ and that for $i \geq n$, D_i is Σ_n -invariant. It follows directly that

$$(N = k, \sigma N > n) = (\sigma N = k, N > n) = D_k^c \bigcap_n^{k-1} D_i \bigcap_1^{n-1} \{D_i \cap \sigma D_i\}. \quad \square$$

REMARK. In Theorem 3.5, the condition of Proposition 3 is satisfied for t_i with $D_n = (t_n \notin V_{ni})$. Thus such $\{t_i\}$ are homogeneously tail-symmetric.

Regarding the examples, it is easily seen that Proposition 3 applies to (i), (iii) and (iv). In example (ii), it is easily checked that $\{t_i\}$ is homogeneously tail-symmetric. In fact, if $\sigma \in \Sigma_n$, then for $i \geq n$, $\sigma t_i = t_i$.