

ACCURACY OF CONVERGENCE OF SUMS OF DEPENDENT RANDOM VARIABLES WITH VARIANCES NOT NECESSARILY FINITE¹

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Let $S_n = \sum_{k=1}^{k_n} X_{nk}$ and X be random variables with distribution functions $F_n(x)$ and $F(x)$. No assumptions are made that the (X_{nk}) have finite means or variances. Also, no independence conditions are assumed. A bound is found for

$$M_n = \sup_{-\infty < x < \infty} |F_n(x) - F(x)|.$$

This bound involves various truncated moments and conditional probabilities and expectations. A typical quantity involved is $\sum_{k=1}^{k_n} E|E(X_{nk} | \sum_{j=1}^{k-1} X_{nj}) - E(X_{nk})|$. Using this bound, particular conditions are found so that S_n converges in distribution to X .

1. Introduction and summary. Let (X_{nk}) , $k = 1, 2, \dots, k_n$, $n = 1, 2, \dots$ be a system of random variables with distribution functions $F_{nk}(x)$. Let $S_n = \sum_{k=1}^{k_n} X_{nk}$ have distribution function $F_n(x)$ and X be a random variable with distribution function $F(x)$. If $\mathcal{L}(S_n) \rightarrow \mathcal{L}(X)$ (i.e., S_n converges in distribution to X), it has been of interest to investigate bounds on

$$M_n = \sup_{-\infty < x < \infty} |F_n(x) - F(x)|.$$

In this paper we obtain a bound on M_n and use the bound to give conditions for the convergence of S_n to X where neither finite variances of the X_{nk} nor any independence conditions are assumed. The convergence theorem is of a type considered by Loève in [5] and by the author in [1] assuming finite variances of the X_{nk} . Theorem 1, whose proof is given in Section 3, gives the bound on M_n . The main component of this bound, $g^a(n, m, r)$, is given at the end of Section 2.

THEOREM 1. Let (X_{nk}) be a system of random variables and X be an infinitely divisible random variable with distribution function $F(x)$. Assume $F'(x) = dF(x)/d(x)$ exists for each x and is bounded by B . Then for each r and a such that $0 < r \leq 1$, $a > 1$

$$M_n = \sup_{-\infty < x < \infty} |F_n(x) - F(x)| \leq \sum_{k=1}^{k_n} (F_{nk}(-a) + 1 - F_{nk}(a)) + h(B)g^a(n, m, r)$$

if $\sigma_{nk}^2(a) \leq 1$ for all n, k where $h(b)$ depends only on B .

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² This is a very mild assumption as seen by the remark of the next to the last paragraph of [1].

Theorems of this type were discussed by Shapiro [6] and Boonyasombut and Shapiro [2] and were applied by them to specific limit theorems where $X_{n1}, X_{n2}, \dots, X_{nk_n}$ were assumed to be independent for each n .

2. Notation. For $a > 0$, let

$$X_{nk}^a = X_{nk} \quad \text{if } -a < X_{nk} \leq a$$

$$= 0 \quad \text{otherwise}$$

with $S_n^a = \sum_{k=1}^{k_n} X_{nk}^a$, $F_{nk}^a(x)$ and $F_n^a(x)$ the corresponding distribution functions, means $\mu_{nk}(a)$ and $\mu_n(a)$, $\sigma_{nk}^2(a)$ the variance of X_{nk}^a and $\sigma_n^2(a) = \sum_{k=1}^{k_n} \sigma_{nk}^2(a)$. Also let

$$F_{nk}^{a'}(x) = P(X_{nk}^a \leq x \mid \sum_{j=1}^{k-1} X_{nj}^a)$$

$$E'(X_{nk}^a) = E(X_{nk}^a \mid \sum_{j=1}^{k-1} X_{nj}^a)$$

$$K_{nk}^a(x) = \int_{-\infty}^x u^2 dF_{nk}^a(u + \mu_{nk}(a)), K_n^a(x) = \sum_{k=1}^{k_n} K_{nk}^a(x)$$

$$K_{nk}^{a'}(x) = \int_{-\infty}^x u^2 dF_{nk}^{a'}(u + \mu_{nk}(a)).$$

We say (X_{nk}^*) is the independent version of (X_{nk}) if for all n, k , X_{nk}^* and X_{nk} have the same distribution and for each n , $X_{n1}^*, X_{n2}^*, \dots, X_{nk_n}^*$ are independent.

Let X be an infinitely divisible random variable with Lévy-Khintchine representation (see (2.1) of [2]) determined by the function $G(u)$ and the constant γ . Let

$$G^a(u) = 0 \quad \text{if } u \leq -a,$$

$$= G(u) - G(-a) \quad \text{if } -a < u \leq a, \quad \gamma^a = \gamma - \int_{|u|>a} u^{-1} dG(a)$$

$$= G(a) - G(-a) \quad \text{if } u > a,$$

which by the above representation determines a unique infinitely divisible random variable X^a with distribution function $F^a(x)$, mean $\mu(a)$, variance $\sigma^2(a)$ (which can be shown to be finite), and Kolmogorov representation (see page 85 of [4]) given by a bounded non-decreasing function $K^a(x)$ and the constant $\mu(a)$.

For any $A > 0$ such that $\pm A$ are continuity points of $G(u)$, let $0 < \delta \leq 2A$ and define $m = m(A, \delta) = [2A/\delta] + 1$, $-A = x_0 < x_1 < \dots < x_m = A$ with x_i continuity points of $G(u)$ and such that $\max(x_i - x_{i-1}) < \delta$. Then for $r > 0$ let

$$g^a(n, m, r) = \left[\frac{5}{16} \sigma_n^2(a) \max_{1 \leq k \leq k_n} \sigma_{nk}^2(a) \right]^{\frac{1}{2}} + \left[\frac{5}{6} \delta (3\sigma_n^2(a) + \sigma^2(a)) \right]^{\frac{1}{2}}$$

$$+ \left[\frac{1}{2} \sum_{i=0}^m \sum_{k=1}^{k_n} E |K_{nk}^{a'}(x_i) - K_{nk}^a(x_i)| + \frac{1}{2} \sum_{i=0}^m |K_n^a(x_i) - K^a(x_i)| \right]^{\frac{1}{2}}$$

$$+ \left[2 \sum_{k=1}^{k_n} E |E'(X_{nk}^a) - \mu_{nk}(a)| + 2 |\mu_n(a) - \mu(a)| \right]$$

$$+ (4/A) (2\sigma_n^2(a) + K_n^a(\infty) - K_n^a(A) + K^a(\infty) - K^a(A) + K_n^a(-A)$$

$$+ K^a(-A))]^{\frac{1}{2}}$$

$$+ \left[\frac{8 \int_{|u|>a} |u|^r dG(u)}{r} \right]^{1/1+r}$$

3. Proof of Theorem 1. The following proof combines the techniques of Theorem 1 of [2] and Theorem 2 of [1]. The crucial approximation of $|\phi_n^a(t) - \phi^a(t)|$ derives from Lemma 2 of [1].

PROOF. To obtain the bound for M_n observe that

$$\begin{aligned} |F_n(x) - F(x)| &\leq |F_n(x) - F_n^a(x)| + |F_n^a(x) - F(x)| \\ &\leq \sum_{k=1}^{k_n} (F_{nk}(-a) + 1 - F_{nk}(a)) + |F_n^a(x) - F(x)| \end{aligned}$$

where the second inequality follows from Lemma 1 of [2]. A bound is now found for $|F_n^a(x) - F(x)|$. Let $\phi_n^a(t)$, $\phi^a(t)$ and $\phi(t)$ be the characteristic functions of S_n^a , X^a , and X respectively. Then

$$|\phi_n^a(t) - \phi(t)| \leq |\phi_n^a(t) - \phi^a(t)| + |\phi^a(t) - \phi(t)|.$$

By the proof of Theorem 2 of [1] if $T_n = 1/g^a(n, m, r)$, then for $|t| \leq T_n$

$$\begin{aligned} |\phi_n^a(t) - \phi^a(t)| &\leq |t|^4 \left[\frac{5}{8} \sigma_n^2(a) \max_{1 \leq k \leq k_n} \sigma_{nk}^2(a) \right] + |t|^3 \left[\frac{5}{4} \delta(3\sigma_n^2(a) + \sigma^2(a)) \right] \\ &\quad + |t|^2 \left[\frac{1}{2} \sum_{i=0}^m \sum_{k=1}^{k_n} E |K_{nk}^{a'}(x_i) - K_{nk}^a(x_i)| + \frac{1}{2} \sum_{i=0}^m |K_n^a(x_i) - K^a(x_i)| \right] \\ &\quad + |t| \left[\sum_{k=1}^{k_n} E |E'(X_{nk}^a) - \mu_{nk}(a)| + |\mu_n(a) - \mu(a)| + (2/A) \{ 2\sigma_n^2(a) \right. \\ &\quad \left. + K_n^a(\infty) - K_n^a(A) + K^a(\infty) - K^a(-A) + K^a(-A) \} \right]. \end{aligned}$$

From Lemma 2 of [2]

$$|\phi^a(t) - \phi(t)| \leq 4|t|^r \int_{|u|>a} |u|^r dG(u).$$

Now applying a result of Esseen [3], for any $p > 1$

$$\sup_{-\infty < x < \infty} |F_n^a(x) - F(x)| \leq \frac{p}{2\pi} \int_{-T_n}^{T_n} \left| \frac{\phi_n^a(t) - \phi(t)}{t} \right| dt + c(p) \cdot \frac{B}{T_n}$$

where $c(p)$ is a constant depending only on p . It is easy to show that

$$\int_{-T_n}^{T_n} \left| \frac{\phi_n^a(t) - \phi(t)}{t} \right| dt \leq g^a(n, m, r).$$

Thus

$$\sup_{-\infty < x < \infty} |F_n^a(x) - F(x)| \leq \left(\frac{p}{2\pi} + c(p) \cdot B \right) g^a(n, m, r)$$

and so by fixing p and letting $h(B) = (p/2\pi + c(p) \cdot B)$ the theorem follows.

4. A convergence theorem. In [1], [2], and [6] theorems similar to Theorem 1 were proven for specific limit theorems. Then in each case the bounds obtained on M_n were shown to converge to zero under conditions of the particular limit theorems under discussion. In this paper, the bound of the previous section, obtained without reference to a particular limit theorem, yields immediate conditions for a limit theorem. At the same time these conditions are such that the bound on M_n converges to zero. A corollary to Theorem 1 is thus obtained with conditions similar to those in [5] for random variables whose variances are not necessarily finite.

THEOREM 2. Let (X_{nk}) be a system of infinitesimal random variables with X as in Theorem 1. Also let $\mathcal{L}(S_n^*) \rightarrow \mathcal{L}(X)$ where $S_n^* = \sum_{k=1}^{k_n} X_{nk}^*$ and (X_{nk}^*) is the independent version of (X_{nk}) .³ Then $\mathcal{L}(S_n) \rightarrow \mathcal{L}(X)$ if

- (i) there is an $r > 0$ such that $\int_{-\infty}^{\infty} |u|^r dG(u) < \infty$,⁴
- (ii) $\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} E|E'(X_{nk}^a) - \mu_{nk}(a)| = 0$,
- (iii) $\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} E|K_{nk}^{a'}(x) - K_{nk}^a(x)| = 0$

for all x and a which are continuity points of $G(u)$, the Lévy–Khintchine function associated with X . Furthermore, the bound on M_n converges to zero as n goes to ∞ .

PROOF. The proof of the last statement establishes the theorem. We first observe that if (X_{nk}) is infinitesimal so is (X_{nk}^a) . It then follows that $\lim_{n \rightarrow \infty} \max_{1 \leq k \leq k_n} \sigma_{nk}^2(a) = 0$. From the fact that $\mathcal{L}(S_n^*) \rightarrow \mathcal{L}(X)$ we have by Theorem 3 of [8] that for each a which is a continuity point of $G(u)$ that $\mathcal{L}(S_n^{*a}) \rightarrow \mathcal{L}(X^a)$ where $S_n^{*a} = \sum_{k=1}^{k_n} X_{nk}^{*a}$. Furthermore, by Theorem 6 of the same paper we have that $\lim_{n \rightarrow \infty} \sigma_n^2(a) = \sigma^2(a)$ and $\lim_{n \rightarrow \infty} \mu_n(a) = \mu(a)$. Since $\mathcal{L}(S_n^{*a}) \rightarrow \mathcal{L}(X^a)$ it follows from the proof of Theorem 2, page 100 of [4] and the remark following it that $K_n^a(x) \rightarrow K^a(x)$ and $K_n^a(+\infty) \rightarrow K^a(+\infty)$. We then have that as $n \rightarrow \infty$

$$K_n^a(\infty) - K_n^a(A) + K^a(\infty) - K^a(A) + K_n^a(-A) + K^a(-A) \rightarrow 2(K^a(\infty) - K^a(A) + K^a(-A)).$$

We now let $A = 1/\delta^{\frac{1}{2}}$ and realize that in $g^a(n, m(A, \delta), r)$, δ is a function of n so we write δ_n and $m(A, \delta) = m(\delta_n)$. Then we can find a sequence $\delta_n = \delta_n(a)$ so that $\pm \delta_n^{-\frac{1}{2}}$ are continuity points of $G(u)$ and such that

$$\left[\frac{1}{2} \sum_{i=1}^{m(\delta_n)} \sum_{k=1}^{k_n} E|K_{nk}^{a'}(x_i) - K_{nk}^a(x_i)| + \frac{1}{2} \sum_{i=0}^{m(\delta_n)} |K_n^a(x_i) - K^a(x_i)| \right] \rightarrow 0 \text{ as } n \rightarrow \infty.$$

Thus it is clear that for each a which is a continuity point of $G(u)$

$$\lim_{n \rightarrow \infty} \left\{ g^a(n, m(\delta_n(a)), r) - \left[\frac{8 \int_{|u|>a} |u|^r dG(u)}{r} \right]^{1/1+r} \right\} = 0$$

and so by Lemma 5 of [2] we can find $a_n \leq a_{n+1}$ so that $\lim_{n \rightarrow \infty} a_n = \infty$ and $\lim_{n \rightarrow \infty} g^{a_n}(n, m(\delta_n(a_n)), r) = 0$. Since by Lemma 4 of [2] $\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} (F_{nk}(-a_n) + 1 - F_{nk}(a_n)) = 0$, the result follows.

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³ This condition is equivalent to assuming certain conditions on $F_{nk}(x)$ such as those given in Theorem 1, Section 25 of [4].

⁴ This condition is satisfied by all stable laws, for example.

REFERENCES

- [1] BLOCK, H. W. (1970). Error estimation for a limit theorem for dependent random variables. *Ann. Math. Statist.* **41** 1334–1338.
- [2] BOONYASOMBUT, V. and SHAPIRO, J. M. (1970). The accuracy of infinitely divisible approximations to sums of independent variables with applications to stable laws. *Ann. Math. Statist.* **41** 237–250.
- [3] ESSEEN, C. G. (1945). Fourier analysis of distribution functions. *Acta Math.* **77** 1–125.
- [4] GNEDENKO, B. V. and KOLMOGOROV, A. N. (1954). *Limit Distributions for Sums of Independent Random Variables*, translated by K. L. Chung. Addison-Wesley, Reading.
- [5] LOÈVE, M. (1950). On sets of probability laws and their limit elements. *Univ. Calif. Publ. Statist.* **1** 53–87.
- [6] SHAPIRO, J. M. (1955). Error estimates for certain probability theorems. *Ann. Math. Statist.* **26** 617–630.
- [7] SHAPIRO, J. M. (1956). A condition for existence of moments of infinitely divisible distribution functions. *Canad. J. Math.* **8** 69–71.
- [8] SHAPIRO, J. M. (1957). Sums of independent truncated random variables. *Ann. Math. Statist.* **28** 754–761.