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THE RATE OF CONVERGENCE OF THE MEAN LENGTH OF THE LONGEST COMMON SUBSEQUENCE[1]

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Given two i.i.d. sequences of n letters from a finite alphabet, one can consider the length L_n of the longest sequence which is a subsequence of both the given sequences. It is known that EL_n grows like γn for some $\gamma \in [0, 1]$. Here it is shown that $\gamma n \geq EL_n \geq \gamma n - C(n \log n)^{1/2}$ for an explicit numerical constant C which does not depend on the distribution of the letters. In simulations with $n = 100{,}000$, EL_n/n can be determined from k such trials with 95% confidence to within $0.0055/\sqrt{k}$, and the results here show that γ can then be determined with 95% confidence to within $0.0225 + 0.0055/\sqrt{k}$, for an arbitrary letter distribution.

1. Introduction. Given a finite alphabet A and two sequences $x_1, \ldots, x_n$ and $y_1, \ldots, y_n$ in A, there is said to be a common subsequence of length k if for some $1 \leq i_1 < \cdots < i_k \leq n$ and $1 \leq j_1 < \cdots < j_k \leq n$, we have $x_{i_m} = y_{j_m}$ for all $1 \leq m \leq k$. We wish to consider the length L_n of the longest common subsequence (LCS) of two A-valued i.i.d. sequences $X_1, \ldots, X_n$ and $Y_1, \ldots, Y_n$ with a common law μ. This problem and its variants have been much studied in probability theory [6, 7, 19], computer science [1, 3, 14] and mathematical biology [11, 15, 16, 18]; see also the volume [17] for several articles. In mathematical biology, the alphabet $A = \{a, c, t, g\}$ of greatest interest consists of the four DNA bases, and one may want to test whether an observed common subsequence between two base sequences could be due to chance. The quantity $2(n - L_n)$ is the minimal number of insertions and deletions needed to change either sequence to the other one; in computer science this "edit distance" is used as a metric on strings.

It is easy to see that $\{EL_n,\ n \geq 1\}$ is a superadditive sequence, that is,

$$EL_{n+m} \geq EL_n + EL_m \quad \text{for all } n, m \geq 1. \tag{1.1}$$

It therefore follows from standard methods that EL_n/n has a limit and the convergence is from below, that is, there exists $\gamma = \gamma(\mu) \in [0, 1]$ such that

$$\lim_n EL_n/n = \sup_n EL_n/n = \gamma. \tag{1.2}$$

Kingman's subadditive ergodic theorem [13] further implies that $L_n/n \to \gamma$ a.s. For fair coin tossing, where $A = \{H, T\}$ and $\mu(H) = \mu(T) = 1/2$, simula-

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tions and heuristics (see [19] and Section 3 below) suggest that γ is between 0.81 and 0.82.

What interests us here is the rate at which the convergence in (1.2) occurs. The following is our main result.

THEOREM 1.1. *There exists a constant C such that for every alphabet A, law μ and $n \geq 1$,*

$$\gamma n \geq EL_n \geq \gamma n - C(n \log n)^{1/2}. \tag{1.3}$$

For a given n_0 our calculations will give an explicit value of C valid for $n \geq n_0$. This C will be smaller for larger n_0 as lower-order terms become negligible. In fact, we will show in Section 2 that for any $C > 3.42$, (1.3) is valid for all sufficiently large n.

The bound in (1.3) is useful in conjunction with simulations, which can really only estimate EL_n, in estimating γ. Simulations with $n = 100{,}000$ will be discussed in Section 3, together with simulations which suggest that the $(n \log n)^{1/2}$ rate in Theorem 1.1 is nearly the best obtainable by the methods of this paper.

Our method is modeled after that of [2], where a rate of convergence problem for first-passage percolation was considered. The analog of L_n is the passage time from the origin to a point n units out on an axis. The applicability of the method is not surprising in view of the fact that the LCS problem can be reformulated as a dependent first-passage percolation problem, as noted in [1, 4, 14, 20].

2. Proof of the main result. In place of L_n it is more convenient to work with

$$U_n := 2(n - L_n).$$

If we think of the corresponding letters of the two maximal identical subsequences as being matched, then U_n represents the number of letters unused in this matching. More generally, for $1 \leq i \leq j + 1$ and $1 \leq m \leq n + 1$, we define $U([i, j], [m, n])$ to be the number of letters unused after matching the corresponding letters of a longest common subsequence of $X_i, \ldots, X_j$ and $Y_m, \ldots, Y_n$. When $j + 1 = i$ and/or $n + 1 = m$, we interpret the corresponding sequence here as being empty, so that, for example, $U([i, i - 1], [m, n]) = n - m + 1$ if $m \leq n + 1$. When $j + 1 < i$ and/or $n + 1 < m$, we use the convention that $U([i, j], [m, n]) = \infty$. We will abbreviate $U([1, j], [1, n])$ to $U(j, n)$. Define

$$V_n := \min_{-n \leq k \leq n} U(n + k, n - k).$$

These quantities appear naturally in the first-passage reformulation of LCS, so we will briefly describe that reformulation now. Consider the integer lattice in $[0, 2n] \times [0, 2n]$, with horizontal and vertical bonds between nearest-neighbor sites of the lattice (that is, pairs x, y with $|x - y| = 1$) and a

diagonal bond from each $(i-1, j-1)$ to (i, j), $1 \le i \le 2n$ and $1 \le j \le 2n$. The passage time of each horizontal and vertical bond is defined to be 1, and the passage time of the diagonal bond from $(i-1, j-1)$ to (i, j) is 0 if $X_i = Y_j$, and ∞ otherwise. Then $U([i, j], [r, s])$ represents the minimal total passage time among all paths from $(i-1, r-1)$ to (j, s) for which each coordinate in nondecreasing. We will call such a path a nondecreasing path. Let l_n denote the diagonal from $(0, 2n)$ to $(2n, 0)$. Then V_n represents the minimal total passage time among all nondecreasing paths starting at $(0, 0)$ and ending on l_n.

For $n \ge 1$ and $\beta > 0$, define the generating functions

$$g_n(\beta) := -\log\Big(\sum_{-n \le k \le n} E \exp(-\beta(U(n+k, n-k) - 2)) \Big).$$

Heuristically one expects the sum in the definition of $g_n(\beta)$ to behave like its largest term, so that $g_n(\beta)$ behaves like $-\log E \exp(-\beta V_n)$. In fact, by Jensen's inequality we have

$$g_n(\beta) \le -\log(Ee^{-\beta(V_n - 2)}) \le \beta(EV_n - 2). \tag{2.1}$$

The key property of $g_n(\beta)$ is given in the following result.

PROPOSITION 2.1. *For each $\beta > 0$, the sequence $\{g_n(\beta): n \ge 1\}$ is superadditive, that is,*

$$g_{n+m}(\beta) \ge g_n(\beta) + g_m(\beta) \quad \text{for all } m, n \ge 0. \tag{2.2}$$

Consequently, for some constants $\nu_\beta \le 2(1-\gamma)$,

$$\lim_n g_n(\beta)/n = \sup_n g_n(\beta)/n = \beta\nu_\beta \quad \text{for each } \beta > 0. \tag{2.3}$$

Before proving Proposition 2.1 we note that together with (1.2) it tells us that, for each fixed n and β, we have

$$g_n(\beta)/\beta \le \nu_\beta n \le 2(1-\gamma)n \le EU_n. \tag{2.4}$$

Thus because EU_n is subadditive and $g_n(\beta)$ is superadditive, $g_n(\beta)/\beta$ and EU_n are on opposite sides of the limiting approximation $2(1-\gamma)n$. It follows that

$$\begin{aligned} EU_{2n} - 4(1-\gamma)n &\le EU_{2n} - 2g_n(\beta)/\beta \\ &= (EU_{2n} - 2EV_n) + 2(EV_n - g_n(\beta)/\beta). \end{aligned} \tag{2.5}$$

Note that all of this is valid even if β is chosen depending on n.

To prove Proposition 2.1 we will need the following result.

LEMMA 2.2. *For each $n, m \ge 0$ and $0 \le k \le n+m$,*

$$\begin{aligned} &U(n+m+k, n+m-k) + 2 \\ &\quad \ge \min\{U([1, n+j], [1, n-j]) + U([n+j+1, n+m+k], \\ &\qquad [n-j+1, n+m-k]): -n \le j \le n, k-m \le j \le k+m\}. \end{aligned}$$

PROOF. Let Γ be a nondecreasing path of minimal total passage time from $(0,0)$ to $(n+m+k, n+m-k)$. Then Γ intersects l_n in a unique point $(n+x, n-x)$, and for some integer j, $-n \le j < n$, either $x = j$ or $x = j + 1/2$. Since the path is nondecreasing, we have $-n \le x \le n$ and $k - m \le x \le k + m$. If $x = j$, then breaking Γ into two pieces at $(n+j, n-j)$ shows that

$$\begin{aligned} &U(n+m+k, n+m-k) \\ &\quad = U([1, n+j], [1, n-j]) \\ &\qquad + U([n+j+1, n+m+k], [n-j+1, n+m+k]). \end{aligned}$$

If $x = j + 1/2$, then replacing the bond from $(n+j, n-j-1)$ to $(n+j+1, n-j)$ in Γ with the bond from $(n+j, n-j-1)$ to $(n+j, n-j)$ and the bond from $(n+j, n-j)$ to $(n+j+1, n-j)$ adds 2 to the passage time. Breaking the altered Γ at $(n+j, n-j)$ then shows that

$$\begin{aligned} &U(n+m+k, n+m-k) + 2 \\ &\quad \ge U([1, n+j], [1, n-j]) \\ &\qquad + U([n+j+1, n+m+k], [n-j+1, n+m-k]). \end{aligned}$$

In both cases, the desired result follows. □

PROOF OF PROPOSITION 2.1. From Lemma 2.2, independence and translation invariance,

$$\begin{aligned} &\sum_{-(n+m) \le k \le n+m} E \exp(-\beta [U(n+m+k, n+m-k) - 2]) \\ &\quad \le \sum_{-(n+m) \le k \le n+m} E \exp(-\beta \min\{U([1, n+j], [1, n-j]) \\ &\qquad + U([n+j+1, n+m+k], [n-j+1, n+m-k]) - 4: \\ &\qquad\qquad -n \le j \le n, k - m \le j \le k + m\}) \\ &\quad \le \sum_{-(n+m) \le k \le n+m} \sum_{\substack{j: -n \le j \le n \\ k-m \le j \le k+m}} E \exp(-\beta\{U([1, n+j], [1, n-j]) \\ &\qquad + U([n+j+1, n+m+k], [n-j+1, n+m-k]) - 4\}) \\ &\quad = \sum_{-(n+m) \le k \le n+m} \sum_{\substack{j: -n \le j \le n \\ k-m \le j \le k+m}} E \exp(-\beta [U(n+j, n-j) - 2]) \\ &\qquad\qquad \times E \exp(-\beta [U(m+k-j, m-k+j) - 2]) \\ &\quad = \Big(\sum_{j: -n \le j \le n} E \exp(-\beta [U(n+j, n-j) - 2]) \Big) \\ &\qquad \times \Big(\sum_{r: -m \le r \le m} E \exp(\{\beta [U(m+r, m-r) - 2]) \Big). \end{aligned}$$

Taking logarithms then yields (2.2). Standard subadditivity arguments then give (2.3). The fact that $\nu_\beta \le 2(1-\gamma)$ follows from (2.1) and the fact that $V_n \le U_n$. $\square$

The following lemma gives a special case of Azuma's inequality [5] and is essentially a martingale version of Theorem 2 of Hoeffding [10].

LEMMA 2.3. *Suppose $f(x_1,\dots,x_n,y_1,\dots,y_n)$ is a function on A^{2n} with the property that changing any one argument of f while holding the others fixed changes the value of f by at most 2. Then for $Z := f(X_1,\dots,X_n,Y_1,\dots,Y_n)$ and $u \ge 0$,*

$$P[Z - EZ \ge u] \le \exp(-u^2/4n).$$

In particular,

$$P[L_{2n} - EL_{2n} \ge u/2] = P[U_{2n} - EU_{2n} \le -u] \le \exp(-u^2/8n), \tag{2.6}$$

$$P[L_{2n} - EL_{2n} \le -u/2] = P[U_{2n} - EU_{2n} \ge u] \le \exp(-u^2/8n), \tag{2.7}$$

$$P[V_n - EV_n \le -u] \le \exp(-u^2/8n) \tag{2.8}$$

and

$$P[V_n - EV_n \ge u] \le \exp(-u^2/8n). \tag{2.9}$$

Theorem 1.1 is an immediate consequence of the next proposition, since monotonicity of EU_n handles odd indices n. Any fixed values of $\lambda > 1$ and $\theta > \sqrt{2}$ suffice for proving Theorem 1.1, so if $C > (2+\sqrt{2})$, then (1.3) is valid for all sufficiently large n. For the explicit confidence intervals of Section 3, the more detailed requirements in (2.10) and (2.11) are important.

PROPOSITION 2.4. *Suppose $n \ge 8$ and $\lambda, \theta > 0$ satisfy*

$$\lambda^2 \ge 1 + 1/(2n\log 2n) + 2\lambda/(2n\log 2n)^{1/2} + (\log 5.1\lambda)/\log 2n + (\log\log 2n)/(2\log 2n) \tag{2.10}$$

and

$$\theta^2 \ge 2 + (\log 4)/\log 2n. \tag{2.11}$$

Then

$$EU_{2n}/(2n) \le 2(1-\gamma) + 2(2\lambda+\theta)((\log 2n)/(2n))^{1/2} + ((8\log 2)/n)^{1/2} \tag{2.12}$$

and

$$EL_{2n}/(2n) \ge \gamma - (2\lambda+\theta)((\log 2n)/(2n))^{1/2} - ((2\log 2)/n)^{1/2}. \tag{2.13}$$

PROOF. We will use (2.5). Let

$$\beta_n := \lambda((\log 2n)/(2n))^{1/2}.$$

Let us first bound $EV_n - g_n(\beta_n)/\beta_n$. From integration by parts and Lemma 2.3,

$$\begin{aligned} &E\exp(-\beta_n V_n) \\ &\quad = \int_0^\infty \beta_n \exp(-\beta_n x) P[V_n \le x]\, dx \\ &\quad \le \exp(-\beta_n EV_n) + \int_0^{EV_n} \beta_n \exp(-\beta_n x)\exp\bigl(-(EV_n - x)^2/(8n)\bigr)\, dx \\ &\quad \le \exp(-\beta_n EV_n) + \beta_n (8\pi n)^{1/2} \exp\bigl(-\beta_n EV_n + 2n\beta_n^2\bigr). \end{aligned}$$

Therefore,

$$\begin{aligned} \exp(-g_n(\beta_n)) &= \sum_{-n\le k\le n} E\exp(-\beta_n(U(n+k, n-k) - 2)) \\ &\le (2n+1) E\exp(-\beta_n(V_n - 2)) \\ &\le (2n+1)\{\exp(-\beta_n(EV_n - 2))\}\bigl[1 + \beta_n(8\pi n)^{1/2}\exp\bigl(2n\beta_n^2\bigr)\bigr]. \end{aligned} \tag{2.14}$$

Observe that

$$\begin{aligned} &\log\bigl(1.011(8\pi n\beta_n^2)^{1/2}\exp(2n\beta_n^2)\bigr) \\ &\quad \le \log(2.860\pi^{1/2}\lambda) + \log\bigl((n\beta_n^2)^{1/2}/\lambda\bigr) + 2n\beta_n^2. \end{aligned} \tag{2.15}$$

Taking logs in (2.14), rearranging and using $n\beta_n^2 \ge 2\log 2$, (2.15) and (2.10), we obtain

$$\begin{aligned} &EV_n - g_n(\beta_n)/\beta_n \\ &\quad \le 2 + \beta_n^{-1}\log(2n+1) + \beta_n^{-1}\log\bigl(1 + (8\pi n\beta_n^2)^{1/2}\exp(2n\beta_n^2)\bigr) \\ &\quad \le 2 + \beta_n^{-1}(\log 2n + 1/(2n)) + \beta_n^{-1}\log\bigl(1.011(8\pi n\beta_n^2)^{1/2}\exp(2n\beta_n^2)\bigr) \\ &\quad \le \lambda^{-1}(2n\log 2n)^{1/2} \\ &\qquad \times\bigl[1 + 1/(2n\log 2n) + 2\lambda/(2n\log 2n)^{1/2} \\ &\qquad\quad + (\log 5.1\lambda)/\log 2n + (\log\log 2n)/2\log 2n + \lambda^2\bigr] \\ &\quad \le 2\lambda(2n\log 2n)^{1/2}. \end{aligned} \tag{2.16}$$

Let us next bound $EU_{2n} - 2EV_n$. In terms of the first-passage formulation, we use what is essentially a reflection argument across the diagonal l_n. From (2.6) of Lemma 2.3, we have

$$\begin{aligned} 1/2 &\le P\bigl[V_n \le EV_n + (8n\log 2)^{1/2}\bigr] \\ &\le \sum_{-n<j<n} P\bigl[U(n+j, n-j) \le EV_n + (8n\log 2)^{1/2}\bigr], \end{aligned} \tag{2.17}$$

where $j = n$ and $-n$ need not be included in the sum because the minimum of $U(n+j, n-j)$ always occurs with $-n < j < n$. Therefore, there exists an

index j with

$$P\big[U(n+j,n-j) \le EV_n + (8n\log 2)^{1/2}\big] \ge 1/(4n-2).$$

Since $U([n+j+1,2n],[n-j+1,2n])$ has the same distribution as $U(n+j,n-j) = U([1,n+j],[1,n-j])$ and is independent of it, we have

$$\begin{aligned} 1/(4n-2)^2 \le P\big[&U([1,n+j],[1,n-j]) \le EV_n + (8n\log 2)^{1/2}, \\ &U([n+j+1,2n],[n-j+1,2n]) \\ &\le EV_n + (8n\log 2)^{1/2}\big] \\ \le P\big[&U_{2n} \le 2\big(EV_n + (8n\log 2)^{1/2}\big)\big]. \end{aligned} \tag{2.18}$$

However, from (2.6) of Lemma 2.3 and (2.11),

$$P\big[U_{2n} \le EU_{2n} - 2\theta(2n\log 2n)^{1/2}\big] \le \exp(-\theta^2\log 2n) \le 1/(4n)^2, \tag{2.19}$$

which with (2.18) shows

$$EU_{2n} \le 2EV_n + 2(8n\log 2)^{1/2} + 2\theta(2n\log 2n)^{1/2}.$$

With (2.5) and (2.16) this proves (2.12), which implies (2.13). □

Our method should be applicable to other problems which have a first-passage formulation. The main ingredients for which one must have analogs are, first, that there is enough independence that something like Lemma 2.3 holds, and second, that for a path Γ as in the proof of Lemma 2.2 which meets l_n at a particular point, the two segments into which Γ is split by that point are independent or at least are appropriately comparable to independent segments as in [2].

3. Confidence bounds and simulations. For $2n \ge 100{,}000$, (2.10) and (2.11) are satisfied with $\lambda = 1.123$ and $\theta = 1.457$. Therefore, from Proposition 2.4,

$$EL_{2n}/2n \le \gamma \le EL_{2n}/(2n) + 0.0450. \tag{3.1}$$

Given k independent observations of L_{2n}, let $\overline{L}_{2n}$ denote the sample mean of these observations. For optimal confidence bounds, we place our estimate of γ in the center of the interval suggested by (3.1), defining

$$\hat{\gamma}_{2n} := \overline{L}_{2n}/(2n) + 0.0225. \tag{3.2}$$

From Lemma 2.3, for $2n \ge 100{,}000$, $x \ge 0.0055/\sqrt{k}$ and $t := \gamma - EL_{2n}/2n$,

$$\begin{aligned} &P\big[|\hat{\gamma}_{2n} - \gamma| > x + 0.0225\big] \\ &\quad\le P\big[\overline{L}_{2n}/(2n) - EL_{2n}/(2n) \ge x + t\big] \\ &\qquad + P\big[\overline{L}_{2n}/(2n) - EL_{2n}/(2n) \le -(x + 0.0450 - t)\big] \\ &\quad\le \exp\big(-2kn(x+t)^2\big)2 + \exp\big(-2kn(x+0.0450-t)^2\big) \\ &\quad\le \exp(-2knx^2) + \exp\big(-2kn(x+0.0450)^2\big) \\ &\quad\le 0.05. \end{aligned}$$

In particular, for $2n \geq 100{,}000$ and $k = 2$,

$$P\big[|\hat{\gamma}_{2n} - \gamma| \leq 0.0264\big] \geq 0.95. \tag{3.3}$$

Eggert and Waterman [9] simulated two trials of L_{2n} for fair coin tossing with $2n = 100{,}000$. The observed values were 81,223 and 81,146, yielding the estimate $\hat{\gamma}_{2n} = 0.8343$. Therefore, from (3.3),

$$0.8079 \leq \gamma \leq 0.8607. \tag{3.4}$$

with 95% confidence. By contrast, the best bounds known with certainty for fair coin tossing are $0.7615 \leq \gamma \leq 0.8376$ [7, 8]. By using this upper bound we can improve on (3.4) as follows. From Lemma 2.3 we have $P[\bar{L}_{2n}/2n - \delta \leq \gamma] \geq 1 - \exp(-2kn\delta^2)$. In particular, with $k = 2$ and $\delta = 0.0039$ this yields $0.8079 \leq \gamma \leq 0.8376$ with 95% confidence.

Additional simulations in [21] suggest that the variance of L_n is approximately proportional to n. If this is true, then Lemma 2.3 cannot be valid with any smaller power of n in the denominator of the exponent. This means that our method cannot yield a better power of n than the $1/2$ which appears in (1.3). Of course, the actual difference $\gamma n - EL_n$ may well be $o(n^{1/2})$, but quite different methods would apparently be needed to obtain such an improved result.

NOTE ADDED IN PROOF. An additional reference by P. Jaillet [(1992) *Math Oper. Res.* **17** 964–980), has a proof of Lemma 2.3 (ii) and gives two-sided bounds with rates as in (1.4) for several functionals including TSP and MST.

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