



Vol. 14 (2009), Paper no. 50, pages 1495–1512.

Journal URL

<http://www.math.washington.edu/~ejpecp/>

A functional approach for random walks in random sceneries

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Abstract

A functional approach for the study of the random walks in random sceneries (RWRS) is proposed. Under fairly general assumptions on the random walk and on the random scenery, functional limit theorems are proved. The method allows to study separately the convergence of the walk and of the scenery: on the one hand, a general criterion for the convergence of the local time of the walk is provided, on the other hand, the convergence of the random measures associated with the scenery is studied. This functional approach is robust enough to recover many of the known results on RWRS as well as new ones, including the case of many walkers evolving in the same scenery.

Key words: Weak convergence, Random walk, Random scenery, Local time.

AMS 2000 Subject Classification: Primary 60F05, 60F25, 60G50.

Submitted to EJP on April 5, 2009, final version accepted May 24, 2009.

1 Introduction

Let $S = (S_n)_{n \in \mathbb{N}}$ be a *random walk* in \mathbb{Z}^d starting from zero. The term "random walk" has to be understood in the broad sense of a discrete process, not necessarily with independent and identically distributed (i.i.d.) increments. Let $\xi = \{\xi_x, x \in \mathbb{Z}^d\}$ be a collection of random variables called the *random scenery*. The *random walk in random scenery* (RWRS) is the cumulative process

$$Z_n = \sum_{k=0}^n \xi_{S_k}.$$

Motivated by the construction of a new class of self-similar stationary increments processes, Kesten and Spitzer [17] and Borodin [3; 4] introduced RWRS in dimension one and proved functional limit theorems. More precisely, they considered the case when the random walk has i.i.d. increments belonging to the domain of attraction of a stable distribution with index $\alpha \in (1, 2]$, the random scenery is i.i.d. and belongs to the domain of attraction of a stable distribution with index $\beta \in (0, 2]$; the scenery and the walk are independent. Under these assumptions, the renormalised process $(n^{-\delta} Z_{[nt]})_{t \geq 0}$ is shown to converge to a δ -self-similar stationary increments process where $\delta = 1 - \alpha^{-1} + (\alpha\beta)^{-1}$.

Since then, many authors considered RWRS models and related functional limit theorems. Bolthausen [2] studied the case of a recurrent \mathbb{Z}^2 -random walk. Maejima [21] considered the case of a random walk in dimension one and of a multivariate random scenery in link with operator self-similar random processes. All these results require independence properties on the increments of the walk and on the scenery. This constraint can be relaxed by assuming some reasonable dependence structure: in Wang [26], the increments of the random walk are assumed to be strongly correlated, Gaussian and in the domain of attraction of a fractional Brownian motion; it evolves in a i.i.d. square integrable random scenery. In Lang & Xahn [18] a functional limit theorem is proved for an i.i.d. increments \mathbb{Z} -random walk in a strongly correlated scenery satisfying a non central limit theorem of Dobrushin & Major [10]. Weakly dependent sceneries have also recently been studied by Guillotin-Plantard & Prieur [15; 16] under a θ -mixing condition. Let us also mention the work of Le Borgne [19] in which the random scenery verifies some strong decorrelation condition.

One dimensional RWRS arise in the study of random walks evolving on oriented versions of \mathbb{Z}^2 (Guillotini-Plantard & Le Ny [13; 14]; Pène [24]) and in the context of charged polymers (Chen & Khoshnevisan [6]). Motivated by models for traffic in a network, Cohen & Samorodnitsky [9] introduced a random reward schema consisting of sums of independent copies of RWRS and studied its convergence. Extensions of this work have been considered in Guillotin-Plantard & Dombry [12] and Cohen & Dombry [8].

In this paper, a functional approach for RWRS is developed and the convergence of the finite-dimensional distributions of RWRS under very general assumptions that cover many of the above results are proved. A functional approach using stochastic integration was suggested by Borodin [4] and Cadre [5] in the case when the random scenery is a martingale. Our approach is rather based on the convergence of random measures to random noises (see [11]) and is more adapted when the scenery is not a martingale (including weakly and strongly correlated RS). Under the assumption of independence of the RW and of the RS (assumed in all the above cited works), it allows to study separately the convergence of the local time associated with the random walk on the one hand, and the convergence of the random measures associated with the random scenery on the other hand.

The technique is also robust enough to deal with a variant of RWRS: the case when many walkers evolve in a single scenery, which might appear as a more realistic model of the random reward scheme considered in [9] since all the users share the same network. To our knowledge, these are the first results in that direction.

Our paper is organized as follows. In Section 2, after introducing a general criterion for the L^p -convergence of the local time of the random walk, we verify that it applies in the classical cases above-mentioned. In Section 3 a sequence of random measures is associated with the random scenery and conditions to ensure its convergence to a stable, Gaussian or fractional Gaussian random noise are given. In Section 4 we explain how these results allow to recover limit theorems for RWRS and also consider the case of many walkers evolving in one single random scenery.

2 Convergence in L^p of the local times of the random walk

2.1 A general criterion

Let $S = (S_n)_{n \in \mathbb{N}}$ be a \mathbb{Z} -valued random sequence. We define the discrete local time at time $n \in \mathbb{N}$ and point $x \in \mathbb{Z}$ by

$$N(n, x) = \text{card}\{k \in [0, n]; S_k = x\}.$$

Proposition 2.1. *Suppose that the following assumption (RW) is satisfied:*

- **(RW1)** *There is some sequence a_n verifying $a_n \rightarrow \infty$ and $n^{-1}a_n \rightarrow 0$ such that the renormalized process $(a_n^{-1}S_{[nt]})_{t \geq 0}$ converges in $\mathcal{D}([0, \infty))$ to some process $(Y_t)_{t \geq 0}$ admitting a local time $(L(t, x))_{t \geq 0, x \in \mathbb{R}}$.*
- **(RW2)** *There is some $p \geq 1$ such that for all $M > 0$:*

$$\lim_{\delta \rightarrow 0} \limsup_{n \rightarrow \infty} \int_{[-M, M]} \mathbb{E}|L_n(t, x) - L_n(t, [x]_\delta)|^p dx = 0$$

where $L_n(t, x)$ is the rescaled discrete local time

$$L_n(t, x) = n^{-1}a_n N([nt], [a_n x]) \tag{1}$$

and $[x]_\delta = \delta[\delta^{-1}x]$.

Then, for every $m \geq 1$, for every $\theta_1 \in \mathbb{R}, \dots, \theta_m \in \mathbb{R}$, for every (t_1, \dots, t_m) such that $0 < t_1 < \dots < t_m$, the following convergence holds in $L^p, p \geq 1$:

$$\sum_{i=1}^m \theta_i L_n(t_i, \cdot) \xrightarrow{\mathcal{L}} \sum_{i=1}^m \theta_i L(t_i, \cdot) \quad \text{as } n \rightarrow \infty.$$

Remark: A direct application of Lebesgue's dominated convergence Theorem shows that Assumption (RW2) is a consequence of the following two conditions:

(RW2.a) There is some $A > 0$ such that for all $n \geq 1$,

$$\sup_{x \in \mathbb{R}} \mathbb{E}|L_n(t, x)|^p \leq A,$$

(RW2.b) For every $t > 0$, for almost every $x \in \mathbb{R}$,

$$\lim_{y \rightarrow 0} \limsup_{n \rightarrow \infty} \mathbb{E} |L_n(t, x) - L_n(t, x + y)|^p = 0.$$

We notice that if these conditions are verified for some $p > 1$ then they hold for any $p' \in [1, p]$.

Proof:

We prove the theorem for $m = 1$ and $t_1 = t > 0$. Generalizations to any $m \geq 1$ can be easily obtained.

We denote by q the conjugate of p .

According to Theorem 2.4 of De Acosta (1970) we need to check two things:

- first that there is a w^* -dense set $D \in L^q(\mathbb{R})$ such that for each $f \in D$,

$$\int_{\mathbb{R}} f(x) L_n(t, x) dx \rightarrow \int_{\mathbb{R}} f(x) L(t, x) dx,$$

- second that the sequence $(L_n(t, \cdot))_{n \geq 1}$ is flatly concentrated, that is, for every $\varepsilon > 0$, there is a finite dimensional subspace F of $L^p(\mathbb{R})$ such that

$$\liminf_{n \rightarrow \infty} \mathbb{P} [L_n(t, \cdot) \in F^\varepsilon] \geq 1 - \varepsilon$$

where F^ε is the ε -neighborhood of F .

The first assertion follows from the invariance principle (RW1) for $a_n^{-1}S_{[n\cdot]}$. We take D as the set of continuous functions with bounded support and observe that the definition of the discrete occupation time L_n and the uniform continuity of f imply

$$\begin{aligned} \int_{\mathbb{R}} f(x) L_n(t, x) dx &= n^{-1} \int_{\mathbb{R}} f(a_n^{-1}x) N([nt], [x]) dx \\ &= n^{-1} \int_{\mathbb{R}} f(a_n^{-1}[x]) N([nt], [x]) dx + o(1) \\ &= n^{-1} \sum_{k=1}^{[nt]} f(a_n^{-1}S_k) + o(1) \\ &= \int_0^t f(a_n^{-1}S_{[nu]}) du + o(1) \end{aligned}$$

From the continuity of the application

$$\mathcal{D}([0, t]) \rightarrow \mathbb{R}, \phi \mapsto \int_0^t f \circ \phi,$$

from the invariance principle (RW1) and the existence of the local time L for Y , we deduce

$$\int_{\mathbb{R}} f(x) L_n(t, x) dx \rightarrow \int_0^t f(Y_u) du = \int_{\mathbb{R}} f(x) L(t, x) dx.$$

We turn to the second assertion and consider the partition $x_j = j\delta$, $j \in \mathbb{Z}$ for a given $\delta > 0$. Define the finite dimensional space F of $L^p(\mathbb{R})$ as the space of functions that are constant on any interval (x_j, x_{j+1}) , $j \in \mathbb{Z}$ and that vanish outside $[-M, M]$ (for some $M > 0$ to be specified later). Define \tilde{L}_n by $\tilde{L}_n(x) = L_n(t, x_j)$ if $x_j \leq x < x_{j+1}$, i.e. $\tilde{L}_n(x) = L_n(t, [x]_\delta)$. Notice that $\tilde{L}_n \in F$ on the event

$$\{L_n(t, \cdot) \equiv 0 \text{ on } \mathbb{R} \setminus [-M, M]\} = \left\{ \max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| \leq M \right\}.$$

Therefore,

$$\begin{aligned} & \mathbb{P} [L_n(t, \cdot) \notin F^\varepsilon] \\ & \leq \mathbb{P} \left[\max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| > M \right] + \mathbb{P} \left[\max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| \leq M \text{ and } \|L_n(t, \cdot) - \tilde{L}_n\|_{L^p} \geq \varepsilon \right] \end{aligned}$$

By the invariance principle (RW1),

$$\lim_{n \rightarrow \infty} \mathbb{P} \left[\max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| > M \right] = \mathbb{P} \left[\max_{0 \leq s \leq t} |Y_s| > M \right]$$

is small for large M . Hence it remains only to show that

$$\lim_{\delta \rightarrow 0} \limsup_{n \rightarrow \infty} \mathbb{P} \left[\max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| \leq M \text{ and } \|L_n(t, \cdot) - \tilde{L}_n\|_{L^p} \geq \varepsilon \right] = 0.$$

Using Markov's inequality, we get

$$\begin{aligned} & \mathbb{P} \left[\max_{0 \leq k \leq [nt]} |a_n^{-1} S_k| \leq M \text{ and } \|L_n(t, \cdot) - \tilde{L}_n\|_{L^p} \geq \varepsilon \right] \\ & \leq \varepsilon^{-p} \mathbb{E} \left[\int_{[-M, M]} |L_n(t, x) - \tilde{L}_n(t, [x]_\delta)|^p dx \right] \end{aligned}$$

and the result follows from Assumption (RW2). \square

2.2 Local time of a random walk in the domain of attraction of a stable Lévy motion

Let $S = (S_n)_{n \in \mathbb{N}}$ be a \mathbb{Z} -random walk such that

$$\begin{cases} S_0 = 0, \\ S_n = \sum_{k=1}^n X_k, \quad n \geq 1, \end{cases}$$

where the X_i are i.i.d. integer-valued random variables belonging to the normal domain of attraction of a strictly stable distribution $\mathcal{S}_\alpha(\sigma, \nu, 0)$ with parameters

$$\alpha \in (1, 2], \sigma > 0 \text{ and } \nu \in [-1, 1].$$

This means that the following weak convergence holds:

$$n^{-\frac{1}{\alpha}} S_n \xrightarrow[n \rightarrow \infty]{\mathcal{L}} Z_\alpha, \tag{2}$$

where the characteristic function of the random variable Z_α is given by

$$\forall u \in \mathbb{R}, \quad \mathbb{E} \exp(iuZ_\alpha) = \exp\left(-\sigma^\alpha |u|^\alpha \left(1 - iv \tan\left(\frac{\pi\alpha}{2}\right) \operatorname{sgn}(u)\right)\right). \quad (3)$$

Assumption **(RW1)** holds with $a_n = n^{1/\alpha}$ and $(Y(t))_{t \geq 0}$ an α -stable Lévy process such that $Y(0) = 0$ and $Y(1)$ is distributed according to Z_α .

The local time $(N(n, x))_{n \in \mathbb{N}; x \in \mathbb{Z}}$ of the random walk S is defined as in the previous section. Its rescaled local time $(L_n(t, x))_{t \geq 0, x \in \mathbb{R}}$ is defined as

$$L_n(t, x) = n^{-(1-1/\alpha)} N([nt], [n^{1/\alpha}x]).$$

We denote by $(L(t, x))_{t \geq 0, x \in \mathbb{R}}$ the jointly continuous version of the local time of the process Y . We are interested in the convergence in $L^p, p \geq 1$ of the rescaled local time L_n to L . By applying the general criterion of Section 2.1, we get the following result.

Proposition 2.2. *For every $m \geq 1$, for every $\theta_1 \in \mathbb{R}, \dots, \theta_m \in \mathbb{R}$, for every (t_1, \dots, t_m) such that $0 < t_1 < \dots < t_m$, the following convergence holds in $L^p, p \in [1, 2]$:*

$$\sum_{i=1}^m \theta_i L_n(t_i, \cdot) \xrightarrow{\mathcal{L}} \sum_{i=1}^m \theta_i L(t_i, \cdot) \quad \text{as } n \rightarrow \infty.$$

Proof:

From the remark following Proposition 2.1, it is enough to prove that conditions **(RW2.a – b)** are satisfied in our setting with $p = 2$.

Condition (RW2.a): Kesten and Spitzer proved in [17] ((2.10) in Lemma 1) that if τ_x denotes the hitting time of the point x by the random walk S , then the following inequality holds for any $x \in \mathbb{Z}$, $r \geq 0$ and $n \geq 1$,

$$\mathbb{P}(N(n, x) \geq r) \leq \mathbb{P}(N(n+1, 0) \geq r) \mathbb{P}(\tau_x \leq n+1).$$

It implies that the moment of order 2 of the random variable $N(n, x)$ is uniformly bounded by the moment of order 2 of $N(n+1, 0)$ which is equivalent to $Cn^{2(1-1/\alpha)}$ for some $C > 0$ (see (2.12) in Lemma 1 of [17]).

Condition (RW2.b): We know (see Lemma 2 and 3 in [17]) that there exists some constant $C > 0$ independent of $x, y \in \mathbb{Z}$ and $n \geq 1$ s.t.

$$\mathbb{E}(|N(n, [n^{1/\alpha}x]) - N(n, [n^{1/\alpha}(x+y)])|^2) \leq C \left| [n^{1/\alpha}x] - [n^{1/\alpha}(x+y)] \right|^{\alpha-1} n^{1-1/\alpha}.$$

Then, we have

$$\begin{aligned} & \lim_{y \rightarrow 0} \limsup_{n \rightarrow \infty} \mathbb{E}(|L_n(t, x) - L_n(t, x+y)|^2) \\ & \leq C \lim_{y \rightarrow 0} \limsup_{n \rightarrow \infty} \left| \frac{[n^{1/\alpha}x]}{n^{1/\alpha}} - \frac{[n^{1/\alpha}(x+y)]}{n^{1/\alpha}} \right|^{\alpha-1} \\ & = C \lim_{y \rightarrow 0} |y|^{\alpha-1} = 0 \quad (\alpha > 1 \text{ by hypothesis}). \quad \square \end{aligned}$$

2.3 Local time of a strongly correlated walk in the domain of attraction of a fractional Brownian motion

Let $S = (S_n)_{n \in \mathbb{N}}$ be a \mathbb{Z} -random walk such that

$$\begin{cases} S_0 = 0, \\ S_n = \left[\sum_{k=1}^n X_k \right], \quad n \geq 1, \end{cases}$$

where $(X_k)_{k \geq 1}$ is a stationary Gaussian sequence with mean 0 and correlations $r(i-j) = \mathbb{E}(X_i X_j)$ satisfying as $n \rightarrow \infty$,

$$\sum_{i=1}^n \sum_{j=1}^n r(i-j) \sim n^{2H} l(n)$$

with $0 < H < 1$ and l is a slowly varying function at ∞ . Assumption **(RW1)** is satisfied with $a_n = n^H l(n)^{1/2}$ and Y the standard fractional Brownian motion $(B_H(t))_{t \geq 0}$ of index $0 < H < 1$, i.e. the centered Gaussian process such that $B_H(0) = 0$ and $\mathbb{E}(|B_H(t_2) - B_H(t_1)|^2) = |t_2 - t_1|^{2H}$. The local time $N_H(n, x)$ with $n \in \mathbb{N}$ and $x \in \mathbb{Z}$ is defined as the number of visits of $(S_n)_{n \in \mathbb{N}}$ at point x up to time n . Its rescaled local time $(L_{n,H}(t, x))_{t \geq 0, x \in \mathbb{R}}$ is defined as

$$L_{n,H}(t, x) = n^{-(1-H)} l(n)^{1/2} N([nt], [n^H l(n)^{1/2} x]).$$

We denote by $(L_H(t, x))_{t \geq 0, x \in \mathbb{R}}$ the jointly continuous version of the local time of the process B_H . We are interested in the convergence in L^p , $p \geq 1$ of the rescaled local time $L_{n,H}$ to L_H . By applying the general criterion of Section 2.1, we get the following result.

Proposition 2.3. *For every $m \geq 1$, for every $\theta_1 \in \mathbb{R}, \dots, \theta_m \in \mathbb{R}$, for every (t_1, \dots, t_m) such that $0 < t_1 < \dots < t_m$, the following convergence holds in L^p , $p \in [1, 2]$:*

$$\sum_{i=1}^m \theta_i L_{n,H}(t_i, \cdot) \xrightarrow{\mathcal{L}} \sum_{i=1}^m \theta_i L_H(t_i, \cdot) \quad \text{as } n \rightarrow \infty.$$

Proof:

Conditions **(RW2.a – b)** directly follow from Lemma 4.4 and Lemma 4.6 of Wang [26]. \square

2.4 The local time of many independent random walks

We consider independent and identically distributed \mathbb{Z} -random walks

$$S^{(i)} = (S_n^{(i)})_{n \in \mathbb{N}}, \quad i \geq 1,$$

verifying assumption **(RW)**. The local time of the i -th random walk is denoted by $(N^{(i)}(n, x))_{n \in \mathbb{N}; x \in \mathbb{Z}}$ and the rescaled one by $(L_n^{(i)}(t, x))_{t \geq 0; x \in \mathbb{R}}$. The independent limit processes of the correctly renormalized random walks $S^{(i)}$ have local time denoted by $(L^{(i)}(t, x))_{t \geq 0; x \in \mathbb{R}}$. We prove that

Proposition 2.4. Let $c_n \rightarrow \infty$ as $n \rightarrow \infty$. Suppose that the random walks $S^{(i)}, i \geq 1$ are i.i.d. and satisfy assumption **(RW)** for some $p \geq 1$. Suppose furthermore that for any fixed $t \geq 0$,

$$\mathbb{E} \left[\left(\int_{\mathbb{R}} L^{(1)}(t, x)^p dx \right)^{1/p} \right] < \infty. \quad (4)$$

Then, for every $m \geq 1$, for every $\theta_1 \in \mathbb{R}, \dots, \theta_m \in \mathbb{R}$, for every (t_1, \dots, t_m) such that $0 < t_1 < \dots < t_m$, the following convergence holds in L^p :

$$\frac{1}{c_n} \sum_{j=1}^m \theta_j \sum_{i=1}^{c_n} L_n^{(i)}(t_j, \cdot) \xrightarrow{\mathcal{L}} \sum_{j=1}^m \theta_j \mathbb{E}(L^{(1)}(t_j, \cdot)) \quad \text{as } n \rightarrow \infty. \quad (5)$$

Proof:

From Proposition 2.1 and Skorohod's representation theorem, there exists a probability space on which are defined independent copies of the processes $L_n^{(i)}$ and $L^{(i)}$ denoted by $\tilde{L}_n^{(i)}$ and $\tilde{L}^{(i)}$ such that for every $i \geq 1$, the sequence $\sum_{j=1}^m \theta_j \tilde{L}_n^{(i)}(t_j, \cdot)$ converges almost surely in L^p to $\sum_{j=1}^m \theta_j \tilde{L}^{(i)}(t_j, \cdot)$ as n tends to infinity. So it is enough to prove the result for these later sequences.

Step 1:

First, from triangular inequality, we have

$$\begin{aligned} & \left\| \frac{1}{c_n} \sum_{j=1}^m \theta_j \sum_{i=1}^{c_n} [\tilde{L}_n^{(i)}(t_j, \cdot) - \tilde{L}^{(i)}(t_j, \cdot)] \right\|_{L^p(\mathbb{R})} \\ & \leq \frac{1}{c_n} \sum_{i=1}^{c_n} \left\| \sum_{j=1}^m \theta_j (\tilde{L}_n^{(i)}(t_j, \cdot) - \tilde{L}^{(i)}(t_j, \cdot)) \right\|_{L^p(\mathbb{R})}. \end{aligned}$$

and we prove that this quantity converges to zero as $n \rightarrow \infty$. To see this, let

$$X_n^{(i)} = \left\| \sum_{j=1}^m \theta_j (\tilde{L}_n^{(i)}(t_j, \cdot) - \tilde{L}^{(i)}(t_j, \cdot)) \right\|_{L^p(\mathbb{R})}.$$

The triangular array of random variables $\{X_n^{(i)}, n \geq 1, i \in \{1, \dots, c_n\}\}$ is such that:

- on each row (fixed n), the random variables $X_n^{(i)}, i \in \{1, \dots, c_n\}$ are i.i.d..
- on each column (fixed i), $(X_n^{(i)})_{n \geq 1}$ weakly converges to zero as $n \rightarrow \infty$.

Under these general assumptions, the following weak law of large numbers holds

$$\frac{1}{c_n} \sum_{i=1}^{c_n} X_n^{(i)} \xrightarrow{\mathcal{L}} 0, \quad (6)$$

implying the required result for the local time.

We finally give some indications for the proof of the weak law of large numbers. Denote by A_n the random variable in the left hand side of (6). Using the independence of the $X_n^{(i)}$'s, its characteristic function is given by

$$\mathbb{E} [\exp(iuA_n)] = \left(\mathbb{E} [\exp(iuc_n^{-1}X_n^{(1)})] \right)^{c_n}$$

and hence

$$|\mathbb{E} [\exp(iuA_n)] - 1| \leq \mathbb{E} [\min (c_n |\exp(iuc_n^{-1}X_n^{(1)}) - 1|, 2)] \leq \mathbb{E} [\min (|u|X_n^{(1)}, 2)].$$

The function $x \rightarrow \min(|u|x, 2)$ is continuous and bounded on \mathbb{R}^+ . Then, since the sequence $(X_n^{(1)})_{n \geq 1}$ weakly converges to 0, we deduce that $(\mathbb{E} [\exp(iuA_n)])_{n \geq 1}$ converges to 1 as $n \rightarrow \infty$ and this proves the weak law of large numbers.

Step 2:

From Assumption (4) and the strong law of large numbers in the Banach space $L^p(\mathbb{R})$ (see [20] for instance), the sequence

$$\frac{1}{c_n} \sum_{j=1}^m \theta_j \sum_{i=1}^{c_n} \tilde{L}^{(i)}(t_j, \cdot)$$

converges almost surely in L^p to the function

$$\sum_{j=1}^m \theta_j \mathbb{E}(\tilde{L}^{(1)}(t_j, \cdot)).$$

Conclusion:

Combining steps 1 and 2 gives (5). \square

Examples: The above proposition directly applies in the case of random walks with i.i.d. increments in the domain of attraction of a stable Lévy motion (see Section 2.2) or of a strongly correlated Gaussian random walk in the domain of attraction of a fractional Brownian motion (see Section 2.3). In both cases, equation (4) is satisfied for every $p \in [1, 2]$: Lemma 2.1 of [12] states that the local time of a stable Lévy motion is L^p -integrable and Theorem 3.1 of [9] gives a similar result for the fractional Brownian motion.

3 Convergence of the random measures associated with the random scenery

Let $\xi = \{\xi_x, x \in \mathbb{Z}\}$ be a family of real random variables. In the sequel, we consider $\mu_h = \mu_h(\xi)$ the random signed measure on \mathbb{R} absolutely continuous with respect to Lebesgue measure with random density

$$\frac{d\mu_h}{dx}(x) = \gamma_h h^{-1} \sum_{k \in \mathbb{Z}} \xi_k \mathbf{1}_{[hk, h(k+1))}(x), \tag{7}$$

where $\gamma_h > 0$ is a normalisation constant.

For a locally integrable function $f \in L^1_{loc}$, we want to define

$$\mu_h[f] = \gamma_h \sum_{k \in \mathbb{Z}} \xi_k h^{-1} \int_{hk}^{h(k+1)} f(x) dx. \tag{8}$$

Denote by \mathcal{F}_{μ_h} the set of functions $f \in L^1_{loc}$ such that this series is convergent in a sense that will be precised later according to the considered cases: either almost-sure semi-convergence or convergence in $L^2(\Omega)$. Clearly, any integrable function with bounded support belongs to the set

\mathcal{F}_{μ_h} .

The following scaling relation is worth noting:

$$\mu_h[f(c\cdot)] = \frac{\gamma_h}{c\gamma_{ch}} \mu_{ch}[f(\cdot)], \quad (9)$$

whenever these quantities are well-defined.

We introduce the cumulative scenery $(w_x)_{x \in \mathbb{Z}}$

$$w_x = \begin{cases} \sum_{i=0}^{x-1} \xi_i & \text{if } x > 0 \\ \sum_{i=x}^{-1} \xi_k & \text{if } x < 0 \\ 0 & \text{if } x = 0 \end{cases}, \quad x \in \mathbb{Z},$$

and we extend it to a continuous process by the linear interpolation

$$w_x = w_{[x]} + (x - [x])(w_{[x]+1} - w_{[x]}).$$

Note that

$$\mu_h[1_{[x_1, x_2]}] = \gamma_h(w_{h^{-1}x_2} - w_{h^{-1}x_1}) = W_h(x_2) - W_h(x_1), \quad x_1, x_2 \in \mathbb{R}, x_1 \leq x_2. \quad (10)$$

with W_h the rescaled cumulative scenery $W_h(x) = \gamma_h w_{h^{-1}x}$.

3.1 Independent and identically distributed scenery

Let $1 < \beta \leq 2$. We consider the case when the $\xi_x, x \in \mathbb{Z}$ satisfy the following assumption:

(RS1) the ξ_x 's are i.i.d. random variables in the normal domain of attraction of the stable law $\mathcal{S}_\beta(\sigma, \nu, 0)$ with characteristic function given in (3).

In this case, with the scaling $\gamma_h = \sigma^{-1}h^{\frac{1}{\beta}}$, the rescaled cumulative scenery W_h converges to a bilateral β -stable Lévy process W such that $W(1)$ is distributed according to the stable distribution $\mathcal{S}_\beta(1, \nu, 0)$. In [11], the convergence of the random measures μ_h to a stable white noise was proved. Here a function $f \in L^1_{loc}$ belongs to \mathcal{F}_{μ_h} if the random series (8) defining $\mu_h[f]$ converges almost surely.

Proposition 3.1. *Suppose that ξ satisfies assumption (RS1) for some $1 < \beta \leq 2$. Then for any $h > 0$, $L^\beta(\mathbb{R}) \subset \mathcal{F}_{\mu_h}$ and if $f_n \rightarrow f$ in L^β and $h_n \rightarrow 0$ as $n \rightarrow \infty$,*

$$\mu_{h_n}[f_n] \xrightarrow{\mathcal{L}} \int_{\mathbb{R}} f(x)W(dx)$$

where W denotes the independently scattered β -stable random noise on \mathbb{R} with Lebesgue intensity and constant skewness ν .

3.2 A general criterion when the scenery is square integrable

We consider the case of a centered scenery with finite variance in the domain of attraction of some square integrable random process W . Our assumption **(RS2)** is:

- **(RS2.a)** The scenery ξ is centered and square integrable.
- **(RS2.b)** The finite dimensional distributions of the rescaled cumulative scenery W_h converge as $h \rightarrow 0$:

$$(W_h(x))_{x \in \mathbb{R}} \xrightarrow{\mathcal{L}} (W(x))_{x \in \mathbb{R}}$$

where W is a non-degenerate centered square integrable process.

- **(RS2.c)** There is some $C_1 > 0$ such that for any $h > 0$, $K \geq 1$, $x_1 \leq \dots \leq x_{K+1}$, $\theta_1, \dots, \theta_K \in \mathbb{R}$,

$$\left\| \sum_{i=1}^K \theta_i (W_h(x_{i+1}) - W_h(x_i)) \right\|_{L^2(\Omega)} \leq C_1 \left\| \sum_{i=1}^K \theta_i (W(x_{i+1}) - W(x_i)) \right\|_{L^2(\Omega)}.$$

Note that $W(0) = 0$ and that W non degenerate means that for any pairwise distinct nonzero $x_1, \dots, x_p \in \mathbb{R}$, the random variables $W(x_1), \dots, W(x_p)$ are linearly independent in $L^2(\Omega)$.

We now define integration of deterministic functions with respect to W . This construction is classical in the case when W is a Gaussian process (e.g. Brownian or fractional Brownian motion), and the reader can refer to [22], but we simply assume here that W is centered square integrable and non degenerate.

Let \mathcal{E} be the class of step functions on \mathbb{R} , i.e. $f \in \mathcal{E}$ if and only if $f = \sum_{i=1}^K \theta_i 1_{]x_i, x_{i+1}]}$ for some $\theta_1, \dots, \theta_K$ and $x_1 < \dots < x_{K+1}$. For $f \in \mathcal{E}$, we define the random variable $W[f] \in L^2(\Omega)$ by

$$W[f] = \sum_{i=1}^K \theta_i (W(x_{i+1}) - W(x_i))$$

and use the notation $W[f] = \int_{\mathbb{R}} f(x) W(dx)$ to emphasize the analogy with integration (even if the path W need not to have bounded variations on compact sets).

Define the scalar product $\langle \cdot, \cdot \rangle_W$ on \mathcal{E} by

$$\langle f_1, f_2 \rangle_W = \text{Cov}(W[f_1], W[f_2]) \quad , f_1, f_2 \in \mathcal{E}.$$

This is indeed a scalar product since the process W is non degenerate. With these notations, $W : \mathcal{E} \rightarrow L^2(\Omega), f \mapsto W[f]$ is a linear isometry. Define the Hilbert space L_W^2 as the closure of \mathcal{E} with respect to the scalar product $\langle \cdot, \cdot \rangle_W$. Since $L^2(\Omega)$ is complete and \mathcal{E} is dense in L_W^2 (by construction), we can extend the isometry into $W : L_W^2 \rightarrow L^2(\Omega)$.

Examples: If W is the Brownian motion, the scalar product is given by

$$\langle f_1, f_2 \rangle_W = \int_{\mathbb{R}} f_1(u) f_2(u) du, \quad f_1, f_2 \in \mathcal{E}$$

and then $L_W^2 = L^2(\mathbb{R})$. Then $W[f]$ is the Wiener integral of f with respect to Gaussian white noise. If W is the fractional Brownian motion of index $H \in (1/2, 1)$, then

$$\langle f_1, f_2 \rangle_W = H(2H - 1) \int_{\mathbb{R} \times \mathbb{R}} |u - v|^{2H-2} f_1(u) f_2(v) dudv, \quad f_1, f_2 \in \mathcal{E}$$

and elements of L_W^2 may not be functions but rather distributions of negative order (see [23] or [25]). To our purpose, we will not have to consider this distributions space, but it will be enough to remark that $L^1 \cap L^2 \subset L_W^2$ and that $\|f\|_{L_W^2} \leq C_2 \|f\|_{L^1 \cap L^2}$ for some constant C_2 , where $L^1 \cap L^2$ is equipped with the norm $\|f\|_{L^1 \cap L^2} = \max(\|f\|_{L^1}, \|f\|_{L^2})$.

In order to compare integration with respect to μ_h and to W , we have to suppose that L_W^2 contains some suitable subspace of locally integrable functions since $\mathcal{F}_{\mu_h} \subseteq L_{loc}^1$ is the set of functions $f \in L_{loc}^1$ such that the random series (8) defining $\mu_h[f]$ converges in $L^2(\Omega)$. In view of the above examples, we make the following further assumption:

- **(RS2.d)** $L^1 \cap L^2 \subseteq L_W^2$ with continuous injection, i.e. there is some $C_2 > 0$ such that for any $f \in L^1 \cap L^2$,

$$\|f\|_{L_W^2} \leq C_2 \|f\|_{L^1 \cap L^2}.$$

We then prove the convergence of the random measures μ_h defined by (7) to the random noise W on the space $L^1 \cap L^2 \subseteq L_W^2$.

Proposition 3.2. *Suppose that ξ satisfies assumptions **(RS2)**.*

Then for any $h > 0$, $L^1 \cap L^2 \subset \mathcal{F}_{\mu_h}$ and if $f_n \rightarrow f$ in $L^1 \cap L^2$ and $h_n \rightarrow 0$ as $n \rightarrow \infty$,

$$\mu_{h_n}[f_n] \xrightarrow{\mathcal{L}} W[f] = \int_{\mathbb{R}} f(x)W(dx).$$

Remark: It is worth noting that in the important case when W is a Brownian motion, assumption **(RS2.d)** is trivial. Proposition 3.2 applies under assumptions **(RS2.b – c)** and the conclusion can be strengthened: the convergence holds on the whole space $L^2 = L_W^2 = \mathcal{F}_{\mu_h}$ rather than only on the subspace $L^1 \cap L^2$.

Proof:

For the sake of clarity, we divide the proof into four steps.

Step 1: First note that in the simple case when $f_n \equiv f \in \mathcal{E}$, Proposition 3.2 directly follows from assumption **(RS2.b)**. Let $f = \sum_{i=1}^K \theta_i 1_{]x_i, x_{i+1}]}$. Using equation (10),

$$\mu_{h_n}[f] = \sum_{i=1}^K \theta_i (W_{h_n}(x_{i+1}) - W_{h_n}(x_i)).$$

Then, assumption **(RS2.b)** states that as $n \rightarrow +\infty$,

$$\mu_{h_n}[f] \xrightarrow{\mathcal{L}} \sum_{i=1}^K \theta_i (W(x_{i+1}) - W(x_i)) = W[f].$$

Step 2: We now prove that for every fixed integer n , $L^1 \cap L^2 \subset \mathcal{F}_{\mu_{h_n}}$ and that μ_{h_n} induces a linear functional $L^1 \cap L^2 \rightarrow L^2(\Omega)$ with norm operator bounded by $C_1 C_2$.

The linearity of μ_{h_n} on \mathcal{E} is straightforward. Using equation (10), assumption **(RS2.c)** can be rewritten as

$$\forall f \in \mathcal{E}, \quad \|\mu_{h_n}[f]\|_{L^2(\Omega)} \leq C_1 \|W[f]\|_{L^2(\Omega)} = C_1 \|f\|_{L_W^2},$$

and then, in view of assumption **(RS2.d)**, for every $f \in \mathcal{E}$,

$$\|\mu_{h_n}[f]\|_{L^2(\Omega)} \leq C_1 C_2 \|f\|_{L^1 \cap L^2}.$$

Since \mathcal{E} is dense in $L^1 \cap L^2$ and $L^2(\Omega)$ is complete, the linear functional μ_{h_n} can be extended to $L^1 \cap L^2$ and its norm operator is bounded by $C_1 C_2$.

Step 3: We consider now the case when $f_n \equiv f$ belongs to $L^1 \cap L^2$.

We approximate f by a sequence of simple functions f_m such that $f_m \rightarrow f$ in $L^1 \cap L^2$ as $m \rightarrow \infty$. Let $Y_n = \mu_{h_n}[f]$, $X_{m,n} = \mu_{h_n}[f_m]$, $X_m = W[f_m]$, $X = W[f]$. We apply Theorem 4.2 in [1] by checking that :

- for fixed m , $X_{m,n}$ weakly converges to X_m (use step 1 with the simple function f_m),
- X_m weakly converges to X because using assumption **(RS2.d)**,

$$\|X_m - X\|_{L^2(\Omega)} = \|f_m - f\|_{L^2_W} \leq \|f_m - f\|_{L^1 \cap L^2} \rightarrow 0,$$

- $\lim_{m \rightarrow \infty} \limsup_{n \rightarrow \infty} \mathbb{P}(|X_{m,n} - Y_n| \geq \varepsilon) = 0$ for any $\varepsilon > 0$.

The last condition is satisfied since

$$\begin{aligned} \mathbb{P}(|X_{m,n} - Y_n| \geq \varepsilon) &= \mathbb{P}(|\mu_{h_n}[f_m - f]| \geq \varepsilon) \\ &\leq \varepsilon^{-2} \|\mu_{h_n}[f_m - f]\|_{L^2(\Omega)}^2 \leq (C_1 C_2)^2 \varepsilon^{-2} \|f_m - f\|_{L^1 \cap L^2}^2. \end{aligned}$$

Using Theorem 4.2 in [1], we deduce that Y_n weakly converges to X as n tends to infinity.

Step 4: We consider the general case $f_n \rightarrow f$ in $L^1 \cap L^2$, and we write

$$\mu_{h_n}[f_n] = \mu_{h_n}[f] + \mu_{h_n}[f_n - f].$$

From step 3, $\mu_{h_n}[f]$ weakly converges to $W[f]$. From step 2,

$$\|\mu_{h_n}[f_n - f]\|_{L^2(\Omega)} \leq C_1 C_2 \|f_n - f\|_{L^1 \cap L^2}$$

and hence $\mu_{h_n}[f_n - f]$ converges to 0 in $L^2(\Omega)$. As a consequence, $\mu_{h_n}[f_n]$ weakly converges to $W[f]$. \square

Note that the above proposition can be applied in the following cases.

Example: Weakly dependent sceneries.

Under the assumption of a L^2 stationary scenery satisfying some θ -mixing condition, a central limit theorem for triangular arrays implying assumption **(RS2.b)** was proved in [15] (Theorem 3.1) as well as a suitable control of the covariance implying assumption **(RS2.c)** (see Lemma 7.1 in [15]). By applying the techniques developed in [19], a Kesten-Spitzer result for a simple \mathbb{Z} -random walk in a random scenery coming from a partially hyperbolic dynamical system can be established. It holds when the scenery satisfies some strong decorrelation properties given in Theorem 1 of [24]. We just mention that conditions **(RS2.a – b)** should hold under these last assumptions.

Example: Moving averages sceneries.

Stationary sceneries obtained from moving averages of i.i.d. square integrable random variables are considered in [11], i.e.

$$\xi_k = \sum_{i \in \mathbb{Z}} c_{k-i} \eta_i$$

where $(\eta_i)_{i \in \mathbb{Z}}$ is an i.i.d. square integrable sequence. In the case when c is summable, then the convergence of the rescaled cumulative random scenery to a Brownian motion is shown and assumptions **(RS2.b – c)** hold.

In the case when c slowly decays, strong correlations persist in the limit. More precisely, if

$$\lim_{k \rightarrow +\infty} |k|^{-\gamma} c_k = p_1 \quad \text{and} \quad \lim_{k \rightarrow -\infty} |k|^{-\gamma} c_k = p_2$$

with $\gamma = (1/2, 1)$ and $p_1 p_2 \neq 0$, then convergence of the rescaled scenery to a fractional Brownian motion of index $H = 3/2 - \gamma \in (1/2, 1)$ is shown and assumptions **(RS2.b – c)** are satisfied. Note that in [11], the dimension d of the scenery needs not to be equal to one (spatial sceneries) and the innovations η_i are more generally assumed to belong to the domain of attraction of a β -stable distribution with $1 < \beta \leq 2$. In this framework, the convergence of the random measures μ_h to a fractional β -stable random noise on the space L^β is shown. The method of convergence of random measures to a random noise is thus very robust. We will show how this can be used to obtain new results on RWRS.

4 Limit theorems for RWRS

The above results can be applied to prove the convergence of RWRS and to obtain limit theorems under fairly general assumptions. The main idea is that the RWRS can be rewritten as a function of the rescaled local time $L_n(\cdot, \cdot)$ of the random walk S and of the random measures μ_h associated with the random scenery ξ . Using the scaling property (9) and introducing the suitable scalings for the local time (equation (1)) and for the random measures (equation (7)), we get

$$\begin{aligned} Z_{[nt]} &= \sum_{k=0}^{[nt]} \xi_{S_k} \\ &= \sum_{x \in \mathbb{Z}} N([nt], x) \xi_x \\ &= \gamma_1^{-1} \int_{\mathbb{R}} N([nt], [x]) \mu_1(dx) \\ &= \frac{n}{\gamma_{a_n^{-1}}} \mu_{a_n^{-1}}[L_n(t, \cdot)] \end{aligned}$$

Or equivalently,

$$n^{-1} \gamma_{a_n^{-1}} Z_{[nt]} = \mu_{a_n^{-1}}[L_n(t, \cdot)] \tag{11}$$

with $a_n^{-1} \rightarrow 0$. We then use the convergence of the local time and of the random measure to prove the convergence of the RWRS. Two cases are considered and stated in the same following theorem.

Theorem 4.1. *Suppose that the random walk S and the random scenery ξ are independent and verify either:*

1. (First case) ξ satisfies **(RS1)** for some $\beta \in (1, 2]$ and S satisfies **(RW)** with $p = \beta$.
2. (Second case) ξ satisfies **(RS2)** and S satisfies **(RW)** with $p = 2$.

Then, the RWRS satisfy the following convergence in the sense of the finite dimensional distributions:

$$(n^{-1}\gamma_{a_n^{-1}}Z_{[nt]})_{t \geq 0} \xrightarrow{\mathcal{L}} \left(\int_{\mathbb{R}} L(t, x)W(dx) \right)_{t \geq 0}$$

where $(W(x))_{x \in \mathbb{R}}$ and $(L(t, x))_{t \geq 0; x \in \mathbb{R}}$ are defined in **(RS)** and **(RW)** respectively and independent of each other.

Proof:

We only prove the theorem in the first case. Thanks to Proposition 2.2 and Skorohod's representation theorem, there exists one probability space on which are defined copies of the processes L_n and L denoted by \tilde{L}_n and \tilde{L} such that the sequence $(\tilde{L}_n)_n$ converges almost surely in L^p to \tilde{L} as n tends to infinity. Then, by combining equality (11) and Proposition 3.1, we get that conditionally on these local times, for every $m \geq 1$, for every $\theta_1 \in \mathbb{R}, \dots, \theta_m \in \mathbb{R}$, for every $0 \leq t_1 < \dots < t_m$,

$$n^{-1}\gamma_{a_n^{-1}} \sum_{i=1}^m \theta_i Z_{[nt_i]} = \sum_{i=1}^m \theta_i \mu_{a_n^{-1}}[\tilde{L}_n(t_i, \cdot)] \tag{12}$$

converges in distribution to

$$\sum_{i=1}^m \theta_i \int_{\mathbb{R}} \tilde{L}(t_i, x)W(dx).$$

The result then easily follows using dominated convergence theorem. \square

We finally state a result for many walkers in random scenery. We consider c_n independent and identically distributed random walks $S^{(i)}, i = 1, \dots, c_n$ defined either as in Section 2.2 or as in Section 2.3. They are assumed to evolve in the same random scenery ξ . The i -th random walk induces a RWRS denoted by $Z^{(i)}$, but the processes $Z^{(1)}, \dots, Z^{(c_n)}$ are not independent since all random walkers evolve in the same scenery (see [9; 8; 12]). The total random reward at time n is then given by

$$Z_{n, c_n} = \sum_{i=1}^{c_n} Z_n^{(i)}.$$

Using equation (11), we get

$$c_n^{-1}n^{-1}\gamma_{a_n^{-1}}Z_{[nt], c_n} = \mu_{a_n^{-1}}[c_n^{-1} \sum_{i=1}^{c_n} L_n^{(i)}(t, \cdot)]. \tag{13}$$

For a finite number $c_n \equiv c$ of independent walkers in a random scenery the functional approach easily gives us the following convergence

$$(n^{-1}\gamma_{a_n^{-1}} \sum_{i=1}^c Z_{[nt]}^{(i)})_{t \geq 0} \xrightarrow{\mathcal{L}} \left(\int_{\mathbb{R}} \left(\sum_{i=1}^c L^{(i)}(t, x) \right) W(dx) \right)_{t \geq 0}.$$

When $c_n \rightarrow \infty$ as $n \rightarrow \infty$, thanks to Propositions 2.4 and 3.1, we get the following result, whose proof is very similar to the proof of Theorem 4.1 and is omitted.

Theorem 4.2. *Suppose that the random walk $S^{(i)}$ and the random scenery ξ are independent and verify either:*

1. (First case) ξ satisfies **(RS1)** for some $\beta \in (1, 2]$ and the $S^{(i)}$ verify **(RW)** as well as the integrability condition (4) with $p = \beta$.
2. (Second case) ξ satisfies **(RS2)** and the $S^{(i)}$ verify **(RW)** as well as the integrability condition (4) with $p = 2$.

Then, the total random reward satisfies the following convergence in the sense of the finite dimensional distributions:

$$(c_n^{-1} n^{-1} \gamma_{a_n^{-1}} Z_{[nt], c_n})_{t \geq 0} \xrightarrow{\mathcal{L}} \left(\int_{\mathbb{R}} \mathbb{E}(L(t, x)) W(dx) \right)_{t \geq 0}.$$

where $(W(x))_{x \in \mathbb{R}}$ and $(L(t, x))_{t \geq 0; x \in \mathbb{R}}$ are defined in **(RS)** and **(RW)** respectively.

5 Conclusion and further comments

In this paper a functional approach for the analysis of RWRS was proposed. The method is quite powerful and allows to recover many of the previous results about RWRS as well as new ones. However, our results are limited to the convergence of the finite-dimensional distributions and the question of the functional convergence (for example in the Skohorod topology) naturally arises. The classical techniques used for proving tightness such as Kolmogorov's criterion can be adapted. The reader should refer to [8], [12], [17] or [26] where tightness results were proved. In a few words, in the case of a random scenery satisfying **(RS2)**, the second order moment of the increments of the RWRS is given for any $0 \leq t_1 < t_2$ by

$$\begin{aligned} \mathbb{E} \left[|n^{-1} \gamma_{a_n^{-1}} Z_{[nt_2]} - n^{-1} \gamma_{a_n^{-1}} Z_{[nt_1]}|^2 \right] &= \mathbb{E} \left[|\mu_{a_n^{-1}} [L_n(t_2, \cdot) - L_n(t_1, \cdot)]|^2 \right] \\ &\leq (C_1 C_2)^2 \mathbb{E} \left[\int_{\mathbb{R}} |L_n(t_2, x) - L_n(t_1, x)|^2 dx \right]. \end{aligned}$$

The proof of the tightness then relies on the analysis of the self-intersection local time of the random walk (see [17] for RW with i.i.d. increments or [26] for strongly correlated Gaussian RW).

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