Electron. Commun. Probab. **25** (2020), no. 22, 1–14. https://doi.org/10.1214/20-ECP302 ISSN: 1083-589X

ELECTRONIC COMMUNICATIONS in PROBABILITY

Coexistence in chase-escape

Rick Durrett^{*} Matthew Junge[†] Si Tang[‡]

Abstract

We study a competitive stochastic growth model called chase-escape in which red particles spread to adjacent uncolored sites and blue particles only to adjacent red sites. Red particles are killed when blue occupies the same site. If blue has rate-1 passage times and red rate- λ , a phase transition occurs for the probability red escapes to infinity on \mathbb{Z}^d , *d*-ary trees, and the ladder graph $\mathbb{Z} \times \{0, 1\}$. The result on the tree was known, but we provide a new, simpler calculation of the critical value, and observe that it is a lower bound for a variety of graphs. We conclude by showing that red can be stochastically slower than blue, but still escape with positive probability for large enough *d* on oriented \mathbb{Z}^d with passage times that resemble Bernoulli bond percolation.

Keywords: growth model; particle system; phase transition. **AMS MSC 2010:** 60K35. Submitted to ECP on June 5, 2019, final version accepted on February 25, 2020.

1 Introduction

First-passage percolation is a stochastic growth model that can be used to describe the spread of species. The classic first passage percolation model on a connected graph G = (V, E) assumes that each edge e is assigned a random variable t_e independently drawn from a common nonnegative distribution F, and these random variables represent the passage time needed to traverse the edges. Specifically, for any edge $e = (x, y) \in E$, if at time t_0 , a particle occupies site x, then after time $t_{x,y}$, the particle at x places a copy of itself at y, provided y is empty at time $t_0 + t_{x,y}^-$.

For any two vertices $x, y \in V$, define $T(x, y) = \inf_{\gamma:x \to y} \sum_{e \in \gamma} t_e$ to be the shortest time needed for a particle at x to reach y. When $G = \mathbb{Z}^d$, the ball $D_t := \{\mathbf{x} \in \mathbb{R}^d : T(\mathbf{0}, \lfloor \mathbf{x} \rfloor) \leq t\}$ has a limit shape [9], where $\lfloor \mathbf{x} \rfloor \in \mathbb{Z}^d$ is the nearest lower-left lattice point for $\mathbf{x} \in \mathbb{R}$. Specifically, under a mild moment condition on F, there exists a convex, axis-symmetric set $\mathcal{D}_F \subseteq \mathbb{R}^d$ such that for any $\epsilon > 0$,

$$P\left((1-\epsilon)\mathcal{D}_F \subseteq \frac{D_t}{t} \subseteq (1+\epsilon)\mathcal{D}_F, \text{ for all } t \text{ sufficiently large}\right) = 1.$$
(1.1)

Understanding \mathcal{D}_F and the boundary fluctuations of D_t remains a challenging area of research [5]. The Markovian setting—in which the passage times are unit exponential

^{*}Duke University. E-mail: rtd@math.duke.edu

[†]Bard College. E-mail: mjunge@bard.edu

[‡]Lehigh University. E-mail: sit218@lehigh.edu

random variables ($F(x) = 1 - e^{-x}$) known as the *Richardson growth model* [24]—is understood about as well as the setting with general edge weights.

There are various extensions of the first-passage percolation model to allow two types of particles (say, red and blue) to spread and interact [14, 15, 21, 16, 18]. In these extensions, each edge e is assigned two passage times (t_e^R, t_e^B) , one for red and one for blue. A rule is imposed to define the interaction of the two types. For example, we may declare that the first color to reach a site occupies it for all time and blocks the other type from expanding through the site. In this case, if one color is surrounded by the other, then it cannot grow any further. Coexistence in this setting means that red and blue particles can both reach infinity. Häggström and Pemantle [16] demonstrated that this occurs with positive probability on \mathbb{Z}^2 for the special case where t_e^R and t_e^B are both exponential random variables with equal rates. Hoffman [19] extended this to the four-color setting. Coexistence in the competitive Richardson growth process is especially relevant, because it implies the existence of at least as many geodesics in ordinary first passage percolation as coexisting types. This is discussed in detail by Ahlberg in [2]. It is known for all but countable many spreading rates that coexistence cannot occur if the red and blue rates are different [17].

We consider a two-type spatial growth process called *chase-escape*. In this process, red particles only spread over the empty sites using the passage times t_e^R , and blue particles take over the red sites using the passage time t_e^B . More precisely, for any edge $(x, y) \in E$, (i) if there is some $t_0 \ge 0$ so that site x is red and site y is empty during time period $[t_0, t_0 + t_{x,y}^R)$, then y turns red at time $t_0 + t_{x,y}^R$; (ii) if there is some $t_0 \ge 0$ so that site x is blue and site y is red during time period $[t_0, t_0 + t_{x,y}^R]$, then y turns blue at time $t_0 + t_{x,y}^R$.

We learned of this process from Lalley. It is inspired by the spread of two species—host and parasite—through an environment. For example, brush may spread to neighboring empty patches of soil while an infection transmits among the roots. It is interesting to ask how the environment (graph) and spreading rates affect the coexistence of both species. A closely related variant of chase-escape was studied by Kordzakhia on trees [20]. Later, Kortchemski considered the process on trees and the complete graph [23, 22]. Chaseescape can also be viewed as scotching a rumor. This interpretation was studied by Bordenave in [7]. The continuous limit of rumor scotching is known as the assassination process and was considered many years earlier by Aldous and Krebs [4].

Unless stated otherwise, we always assume that t_e^R is sampled from a rate- λ exponential distribution (denoted $\text{Exp}(\lambda)$), with t_e^B from a rate-1 exponential distribution. We let $P_{\lambda}(\cdot)$ be the probability measure on chase-escape with these passage times. Let A = A(G) be the event that red and blue particles coexist at all times on the graph G. For simplicity, we always assume that G is rooted at a vertex ρ and consider the initial configuration, where a red particle is at ρ and a blue particle is at an auxiliary vertex $\mathfrak{b} \notin G$ attached to ρ by an edge \mathfrak{e} . We say that *coexistence* occurs if $P_{\lambda}(A) > 0$. Figure 1 shows three realizations of chase-escape on \mathbb{Z}^2 . On a given graph, it is natural to ask how λ affects coexistence. Define

$$\lambda_c(G) := \sup\{\lambda \colon P_\lambda(A) = 0\}$$

to be the fastest red expansion rate for which coexistence does not occur. Note that it is not obvious how to prove that $P_{\lambda}(A)$ is monotonic in λ , thus it is not clear if $\lambda_c(G)$ is also equal to $\inf\{\lambda: P_{\lambda}(A) > 0\}$. This is discussed in more detail at the end of this section.

As mentioned earlier, a similar model called *escape* was studied first by Kordzakhia on homogeneous trees [20]. In the escape model, blue can also spread to empty sites. Note that when the underlying graph is a tree, the survival of red is equivalent in either escape or chase-escape. Kordzakhia gave an explicit formula for $\lambda_c(\mathbb{T}_d)$, but did not work

Coexistence in chase-escape



Figure 1: Chase-escape on \mathbb{Z}^2 with $\lambda = 1$ (left), $\lambda = 0.75$ (middle) and $\lambda = 0.5$ (right). Simulations suggest that red can survive despite being slower.

out whether survival occurs when $\lambda = \lambda_c(\mathbb{T}_d)$. This was answered later by Bordenave [8] in the more general setting of Galton-Watson trees. He proved that red does not survive on any tree with branching number $d \in (0, \infty)$ at criticality. Kortchemski also studied the process on trees and made some progress at describing the number of surviving particles at each level [23]. A recent paper by Beckman, Cooke, Eikmeier, Hernandez-Torres, and Junge [6] studies chase-escape in which red particles die on regular trees. They find the expected number of sites occupied by blue is different at criticality than for chase-escape without death. In this work, we provide a much simpler calculation of $\lambda_c(\mathbb{T}_d)$ in Theorem 1.1 that includes the behavior at criticality. Our main interest, however, is in the process on graphs with cycles, which are a more realistic environment for these dynamics and for which much less is known about chase-escape.

The shape theorem at (1.1) implies that for $G = \mathbb{Z}^d$ and $\rho = \mathbf{0} := (0, \dots, 0)$

$$\lambda_c(\mathbb{Z}^d) \leq 1$$
, for all $d \geq 1$.

To see this, suppose $\lambda > 1$ and let \mathcal{D}_1 be the limit shape of the one-type Richardson growth model [24]. If red and blue were to grow over the empty sites on two \mathbb{Z}^d lattices separately, starting from 0 and at rates λ and 1, respectively, then, for $\epsilon = (\lambda - 1)/3$, there exists some T such that

$$B(t) \subseteq (1+\epsilon)t\mathcal{D}_1 \subsetneq (1+2\epsilon)t\mathcal{D}_1 \subseteq R(t), \text{ for all } t > T,$$
(1.2)

where R(t) (resp. B(t)) denote the sites that can be reached by red (resp. blue) particles by time t. Recall that in the chase-escape model, we always assume that at time t = 0, a blue particle is attached to the red particle at 0 by a special edge c. Thus, red survives whenever $t_{\mathfrak{e}}^B > T$, where T is the special time such that (1.2) holds, and the event $\{t_{\mathfrak{e}}^B > T\}$ occurs with positive probability.

By comparing to a nearest-neighbor random walk, it is straightforward to prove that $\lambda_c(\mathbb{Z}) = 1$ (see Lemma 2.1). The discussion at the end of [22] credits James Martin with the conjecture that $\lambda_c(\mathbb{Z}^2) < 1$. Simulations in [25] suggest that red can survive for $\lambda < 1$ and that $\lambda_c(\mathbb{Z}^2) = 1/2$ (see Figures 1 and 2). This is surprising, because, if true, red can still survive even when it spreads significantly slower than blue. One "advantage" red has is that if blue reaches the red boundary it gets slowed down. Another is that the optimal routes for blue are unlikely to coincide with those for red. So, blue might have to travel at a suboptimal rate. Using these advantages to form a rigorous proof seems closely related to understanding the boundary fluctuation of the Richardson growth model, and geodesics in first passage percolation. Despite some recent progress from Damron and Hanson as well as Ahlberg and Hoffman [11, 3], both of these objects are notoriously difficult to describe [5].

ECP 25 (2020), paper 22.

Ì



Figure 2: Chase-escape on \mathbb{Z}^2 at the conjectured critical value $\lambda = 0.5$ (left). The right image shows that there are still surviving red buds at the boundary of the fractal-like blue region.

Proving coexistence in \mathbb{Z}^d for some pair (d, λ) with $d \ge 2$ and $\lambda \le 1$ would be very interesting. We further conjecture that $\lambda_c(\mathbb{Z}^d) \downarrow 0$ as $d \uparrow \infty$. This may be hard to prove since it is unclear that the process is monotonic in λ or d. Intuitively, $P_{\lambda}(A)$ ought to increase with λ and decrease with d, and this is indeed the case on trees. However, there is no obvious coupling that establishes this on graphs with cycles. The issue is that when red spreads faster, then blue also speeds up. It is not even obvious that there is a single phase transition.

1.1 Results

Establishing a coexistence phase where red is slower than blue is tractable on the infinite *d*-ary tree \mathbb{T}_d (a rooted tree where each vertex has *d*-children). Our first result is a new proof that computes $\lambda_c(\mathbb{T}_d)$ and determines the behavior at criticality. Note, as discussed earlier, this result was already proved in [20, 8]. Our proof is different though and much shorter.

Theorem 1.1. On the *d*-ary tree with $d \ge 2$,

$$\lambda_c(\mathbb{T}_d) = 2d - 1 - 2\sqrt{d^2 - d} \sim \frac{1}{4d},$$

and $P_{\lambda_c}(A) = 0$.

The proof that $P_{\lambda}(A) = 0$ for $\lambda \leq \lambda_c$ uses the fact that, restricted to a ray on \mathbb{T}_d , the distance between red and blue is a nearest neighbor random walk on the nonnegative integers. We can use the setup from Proposition 2.1 and the asymptotic behavior for the return time to 0 of a random walk to show that the expected number of sites at distance n to be colored red is summable. This implies coexistence does not occur. To prove $P_{\lambda}(A) > 0$ for $\lambda > \lambda_c$ we embed a Galton-Watson process that describes the number of red particles that reach generation kN, $k \in \mathbb{N}$. When N is large enough, the Galton-Watson process is supercritical, and thus $P_{\lambda}(A) > 0$.

Our main interest is studying chase-escape on graphs with cycles. The argument that proves $P_{\lambda}(A) = 0$ for $\lambda \leq \lambda_c(\mathbb{T}_d)$ can be easily extended to any graph on which the number of length-n paths grows no faster than exponential.

Corollary 1.2. Fix an integer $d \ge 2$ and an infinite, connected graph G. Let Γ_n be the number of self-avoiding length-n paths starting from the root. If $|\Gamma_n| \le Cd^n$ for some C > 0 and all $n \ge 1$, then $\lambda_c(\mathbb{T}_d) \le \lambda_c(G)$.

We initially thought it possible that $\lambda_c(\mathbb{Z} \times \{0,1\})$ was strictly less than 1. This was because blue is slowed down when it reaches the not-entirely-filled red boundary. However, this effect is negligible, because the mixed region of red and uncolored sites is exponentially unlikely to be large.

Theorem 1.3. For $G = \mathbb{Z} \times \{0, 1\}$ with the subgraph structure induced by \mathbb{Z}^2 and $\rho = 0$, $\lambda_c(G) = 1$

To prove $\lambda_c(\mathbb{Z} \times \{0,1\}) \leq 1$, we show that for any $\lambda < 1$, there is $K < \infty$ such that blue is within a distance K of the red boundary infinity often. Whenever this occurs, there is a positive probability that blue overtakes all of the surviving red sites. Thus, red is caught eventually. While we did not work out the details, this result suggests that red is caught on any strip of a fixed height. Additionally it ought to hold that $P_1(A) = 0$ on the ladder, however, our proof relies heavily on the fact that blue moves strictly faster than red and some new ideas are required.

Next, we show that it is possible on the lattice for red to spread slower than blue, but still have a chance to survive. We consider the chase-escape model on the oriented lattice \mathbb{Z}^d , in which particles can only occupy their neighboring sites along the oriented edges $\{\vec{e}_{x,x+b_i}\}_{x\in\mathbb{Z}^d,i=1,\ldots,d}$, with $\{b_1,\ldots,b_d\}$ being the standard, positive basis vectors for \mathbb{Z}^d . Initially, $\mathbf{0} = (0,\ldots,0)$ is red and $(-1,0,\ldots,0)$ is blue. Red and blue passage times for each directed edge \vec{e} are sampled independently from the two distributions

$$t_{\vec{e}}^R \sim p\delta_1 + (1-p)\delta_{\infty}$$

$$t_{\vec{e}}^B \sim p\delta_0 + (1-p)\delta_m.$$
 (1.3)

With atomic passage times there could be moments that blue arrives at a red site at the exact time that red would spread to the next site. To make the process well-defined, we invoke the rule that blue catches red and prevents it from spreading if blue arrives at or before the time red would move. Similar competitive growth models that mimic Bernoulli percolation are studied by Garet and Marchand in [14, 15]. The competition dynamics they consider are not chase-escape, but rather more like the dynamics in [16].

Since edges with ∞ passage times are never used, the collection of all edges for which red passage time equals one is equivalent to the open clusters in *bond percolation*, the random subgraph obtained when each edge in \mathbb{Z}^d is independently retained (open) with probability p and removed (closed) with probability 1 - p. Many aspects of bond percolation are well understood. Letting C be the connected component containing the origin, it is known that there is a critical value $\vec{p}_c(d)$ and if $p > \vec{p}_c(d)$ we have $P_p(|\mathcal{C}| = \infty) > 0$ [12].

Red particles never use edges with ∞ passage times to expand. However, blue can still "jump" across such an edge if the other endpoint was colored red along some other path with no infinite red passage time. Uninhibited, blue would spread much faster than red. However, the presence of sufficiently many "dead ends"—places where red has ∞ passage times in all directions—cuts blue off from spreading and makes it possible for red to escape for small p and large d and m. Let A = A(d, p, m) be the event that there are always surviving red particles in the percolation-like chase-escape model on \mathbb{Z}^d . We show that if d and p are chosen appropriately, and m is made large then red can survive.

Theorem 1.4. For $G = \mathbb{Z}^d$ and $\rho = \mathbf{0}$, there is a choice of d and $p > \vec{p}_c(d)$ such that

$$\lim_{m \to \infty} P_{p,m}(A) = (1-p)P_p(|\mathcal{C}| = \infty).$$

In particular, $P_{p,m}(A) > 0$ for some choice of d, p, and m.

ECP 25 (2020), paper 22.

It is slightly unsatisfying to require that d be large. The reasoning in the proof of Corollary 1.2 can also be used to show that for $p > \vec{p}_c(2)$, and small values of m, that $P_{p,m}(A(\vec{\mathbb{Z}}^2)) = 0$. Thus, if

$$P_{p_0,m_0}(A(\vec{\mathbb{Z}}^2)) > 0$$
 for some p_0 and m_0 ,

then there is a phase transition in $P_{p_0,m}(A(\vec{\mathbb{Z}}^2))$ as we increase m. As before there is no obvious monotonicity in m, so we cannot rule out the possibility of multiple phase transitions.

2 The path, tree, and ladder

2.1 The path

Proposition 2.1. $\lambda_c(\mathbb{Z}) = 1$ and $P_1(A) = 0$.

Proof. Since the process evolves independently in the positive and negative directions, it suffices to prove that red survives on $\{-1, 0, 1, 2, ...\}$ with a blue particle at -1 and a red particle at 0 initially. Let R_t be the number of sites ever occupied by the red particles up to time t and let B_t be the corresponding quantity for blue particles. Define $\tau_n = \inf\{t: R_t + B_t = n + 2\}$, and $\tau_n = \infty$ if there is no such t. Let D_t be the distance between the rightmost red and blue sites at time t and $S_n = D_{\tau_n}$. Notice that (S_n) is a nearest-neighbor random walk on the integers, starting at $S_0 = 1$, where 0 is an absorbing state. Due to the independence of red and blue passage times, we have

$$p := P_{\lambda}(S_{n+1} = S_n + 1) = P(\operatorname{Exp}(\lambda) \le \operatorname{Exp}(1)) = \frac{\lambda}{\lambda + 1}.$$
(2.1)

When $\lambda \leq 1$, the extinction of red is equivalent to the above *p*-biased random walk visiting zero. This is well known to be a.s. finite. On the other hand, when $\lambda > 1$, there is a positive probability that S_n never visits 0 and thus $P_{\lambda}(A) > 0$ for such λ .

2.2 The tree

To study the process on the tree we need an asymptotically precise estimate for the probability that red can reach a site at distance n in the $G = \mathbb{Z}$ case. We prove it here.

Lemma 2.2. Let $A_n = A_n(\lambda)$ be the event that site n is ever colored red in the chase-escape model on \mathbb{Z} , and set $a_n = \frac{[4p(1-p)]^n}{n^{3/2}}$, where $p = \lambda/(\lambda + 1)$.

- (i) For some c > 0 and all $n \ge 1$, $P_{\lambda}(A_n) \ge c a_n$
- (ii) If $\lambda < 1$, then for some $C_{\lambda} > 0$ and all $n \ge 1$, $P_{\lambda}(A_n) \le C_{\lambda} a_n$;

Proof. Let (S_k) be the nearest-neighbor *p*-biased random walk as in (2.1). Note that the event A_n is equivalent to the event that S_k remains strictly positive for the first 2n steps, i.e.,

$$P_{\lambda}(A_n) = P_{\lambda}(S_k \ge 1, k \le 2n | S_0 = 1) = \sum_{a=0}^n P_{\lambda}(S_k \ge 1, k \le 2n; S_{2n} = 2a + 1 | S_0 = 1).$$

Since S_{2n} must have the same parity as S_0 , we only considered the cases where S_{2n} is odd. For any random walk path of length 2n from 1 to 2a + 1 that does not hit zero, there must be (n + a) steps to the right and (n - a) steps to the left. Using the reflection principle (see [13, Theorem 4.3.2]), the total number of such paths is

$$\binom{2n}{n+a} - \binom{2n}{n+a+1}.$$

ECP 25 (2020), paper 22.

http://www.imstat.org/ecp/

We then have

$$P_{\lambda}(A_n) = p^{2n} + \sum_{a=0}^{n-1} \left[\binom{2n}{n+a} - \binom{2n}{n+a+1} \right] p^{n+a} (1-p)^{n-a}$$

$$= p^{2n} + p^n (1-p)^n \sum_{a=0}^{n-1} \binom{2n}{n+a} \frac{2a+1}{n+a+1} \left(\frac{p}{1-p} \right)^a$$

$$\leq p^{2n} + \frac{p^n (1-p)^n}{n+1} \binom{2n}{n} \sum_{a=0}^{\infty} (2a+1) \left(\frac{p}{1-p} \right)^a.$$
 (2.2)

When $\lambda < 1$, we have $p < \frac{1}{2} < 1 - p$, the summation above is finite. Moreover, $\frac{1}{n+1} {\binom{2n}{n}}$ is the *n*-th Catalan number, which is known to be of order $\frac{4^n}{n^{3/2}}$ for large *n*. Putting $C_{\lambda} = \sum_{a=0}^{\infty} (2a+1) \left(\frac{p}{1-p}\right)^a$, we have the desired upper bound. The lower bound is obtained by looking at the a = 0 term in (2.2) and using the asymptotic behavior of Catalan numbers.

Proof of Theorem 1.1. We consider the initial configuration in which the root is red and a special vertex b attached to the root is blue. First note that if $\lambda > 1$, then the distance between red and blue along an arbitrary path to ∞ is equivalent to chase escape on \mathbb{Z} . By Lemma 2.1 we have $P_{\lambda}(A) > 0$ in this case.

Now, suppose that $\lambda \leq 1$. Let R_n be the number of sites at distance n that are ever colored red and $R = \sum_{n=1}^{\infty} R_n$ be the total number of sites of colored red. Notice that red survives a.s. if and only if it occupies infinitely many sites. Thus, $P_{\lambda}(A) = P_{\lambda}(R = \infty)$. We show that $P_{\lambda}(R = \infty) = 0$ for λ small enough by proving $ER < \infty$.

For any vertex $v \in \mathbb{T}_d$, let |v| denote its graph distance from the root. Let A(v) be the event that v is ever colored red. Since the tree has no cycles, we have $P_{\lambda}(A(v)) = P_{\lambda}(A_n)$ for any $v \in \mathbb{T}_d$ with |v| = n, with A_n the event that red reaches a distance n on a fixed path as in Lemma 2.2. Linearity of expectation and the bound from Lemma 2.2 gives

$$ER_n = E \sum_{|v|=n} \mathbf{1}_{A(v)} = \sum_{|v|=n} P_{\lambda}(A(v)) \le C_{\lambda} \frac{d^n (4p(1-p))^n}{n^{3/2}}.$$

Observe that $\lambda_c(d)$ is the smallest solution of

$$4p(1-p)d = \frac{4d\lambda}{(1+\lambda)^2} = 1.$$

It is straightforward to verify that $4dp(1-p) \leq 1$ for $\lambda \leq \lambda_c(d)$, and in this case ER_n is summable, and thus $ER < \infty$.

To prove that $P_{\lambda}(A) > 0$ for $\lambda > \lambda_c(d)$, observe that the lower bound in Lemma 2.2 ensures that for some fixed, large N we have $d^N P_{\lambda}(A_N) > 1$. Thus, the expected number of sites at distance N that are ever colored red is strictly greater than 1. When first occupied by red, the distance from each of these sites to the nearest blue particle is at least one. Since the tree has no cycles, the survival probability of chase-escape is monotonic on a tree; both respect to λ and the initial distance blue starts from red. This means that moving the chasing blue particles to distance 1 from each of red site at distance N will result in fewer surviving red particles at distance 2N. Thus, the number of sites colored red at distances $N, 2N, \ldots$ dominates a Galton-Watson process with mean $d^N P_{\lambda}(A_N) > 1$. This expression for the mean comes from linearity of expectation applied to the d^N sites at distance N from the root. This is supercritical, and thus $P_{\lambda}(R = \infty) > 0$.

ECP 25 (2020), paper 22.

We can quickly deduce that the critical speed for any graph with less than d^n selfavoiding paths of length n is no smaller that of a d-ary tree.

Proof of Corollary 1.2. Let G_n be the event that there is a length-n self-avoiding path started at the root that red reaches the terminal vertex of. It is easy to see that red survives only if G_n occurs for infinitely many n. A union bound and our hypothesis $|\Gamma_n| \leq Cd^n$ gives $P(G_n) \leq Cd^n P(A_n)$ with A_n as in Lemma 2.2. As in the proof of Theorem 1.1, we have $\lambda \leq \lambda_c(\mathbb{T}_d)$ implies that $\sum_{n\geq 1} P(G_n) < \infty$. The Borel-Cantelli lemma implies that almost surely finitely many G_n occur, and thus $P_{\lambda}(A) = 0$ for such λ .

2.3 The ladder

Proof of Theorem 1.3. As with the path, it suffices to consider the one-sided ladder. Note that $\lambda_c(\mathbb{Z}_+ \times \{0,1\}) \leq 1$, following a similar shape argument as in the introduction and the shape theorem for first passage percolation on one-dimensional graphs [1, Theorem 1.1]. For the lower bound, we start by considering ordinary first-passage percolation with only one-type, that is, red particles spread at rate λ and blue particles are absent. We will say that empty sites have state 0 and red sites have state 1. Let $\xi_t(m, n)$ denote the state of site (m, n) at time t:

$$\xi_t(m,n) = \begin{cases} 1, & \text{if } (m,n) \text{ is red at time } t \\ 0, & \text{if } (m,n) \text{ is empty at time } t \end{cases}$$

Let $\eta_t(m) = \xi_t(m, 0) + \xi_t(m, 1)$. If $\eta_t(m) = 1$ then we say that there is a *hole* at distance m at time t. Define $H_t := |\{m : \eta_t(m) = 1\}|$ to be the number of holes at time t. The process $(H_t)_{t\geq 0}$ is dominated by a time-homogeneous Markov chain with jump rates

$$q(k-1,k) = 2\lambda$$
$$q(k,k-1) = k\lambda$$

for all $k \ge 1$. This is because holes can be filled in by any neighbor and new holes can only be created at the front. A comparison with an asymmetric random walk, in which $k \to k - 1$ happens at rate 3λ for all $k \ge 3$ and all other transition rates are held the same, shows that the first time to return to 0 (denoted by T_0) satisfies that for all n, $P_1(T_0 \ge n) \le ce^{-\gamma n}$ for some $c, \gamma > 0$.

For any $x, y \in \mathbb{Z}_+ \times \{0, 1\}$, let T(x, y) denote the passage time to travel to site y starting from x. Define

$$u := \lim_{n \to \infty} \frac{T(\mathbf{0}, \mathcal{H}_n)}{n},$$

where $\mathcal{H}_n = \{(n,0), (n,1)\}$ and $T(\mathbf{0}, \mathcal{H}_n)$ represents the minimal passage time to the set \mathcal{H}_n . Notice that the ladder graph is a one-dimensional graph as defined in [1]. Using the same regenerative structure employed in [1] (especially, Section 2.1 in the arXiv preprint for the special case of the ladder with exponential passage times), we see that $T(0, \mathcal{H}_n)$ can be approximated by a sum of i.i.d random variables, each dominated by a sum of Geo(1/2) independent exponential random variables and thus has an exponential tail. It follows from standard large deviation estimates that, for each $\epsilon > 0$, there are constants A, B > 0 (depending on ϵ) such that for all $n \geq 1$,

$$P_{\lambda}(|T(\mathbf{0},\mathcal{H}_n) - n\nu| \ge n\epsilon) < Ae^{-Bn}.$$
(2.3)

Let $s_0 = 0$ and for $k \ge 0$, define

$$\tau_k := \inf\{t \ge s_k \colon H_t = 0\},\\ s_{k+1} := \inf\{t > \tau_k \colon H_t > 0\}.$$

ECP 25 (2020), paper 22.

http://www.imstat.org/ecp/

At each τ_k , the red region contains no holes, we call this state a *solid red block*, and we call each interval $[\tau_k, \tau_{k+1})$ a *red cycle*. Note that when blue particles are absent, $\{\tau_{k+1} - \tau_k\}_{k>0}$ is stationary. Denote $\sigma := E(\tau_{k+1} - \tau_k)$.

Consider now the chase-escape model, where each site can be in one of the three states 0 (empty), 1 (red), and 2 (blue). At each time $t \ge 0$, let \tilde{R}_t denote the first coordinate of the rightmost red particle and \tilde{B}_t that of the rightmost blue particle. The definition of a "hole" is the same as above: a hole is present at distance m at time t only if the two sites at (m, 0) and (m, 1) consist of one red site and one empty site at time t. Note that this configuration can only appear in the region where blue has not reached but red has. Specifically, if $\{(m, 0), (m, 1)\}$ consists of one blue site and one empty site, we do not count location m as a hole. The sequences, $\{\tilde{s}_k\}_{k\geq 0}$ and $\{\tilde{\tau}_k\}_{k\geq 0}$, and red cycles $\{[\tilde{\tau}_k, \tilde{\tau}_{k+1})\}_{k\geq 0}$ are defined accordingly in the chase-escape setting.

Red spreads at rate λ with $\lambda < 1$ and the blue at rate 1. Suppose the process is in some configuration at t = 0. It suffices to show that, independent of this configuration, there are constants K > 0 and $\delta > 0$ such that with probability at least δ at some later time $T < \infty$, we have $\tilde{R}_T - \tilde{B}_T \leq K$ and that $[\tilde{B}_T + 1, \tilde{R}_T] \times \{0, 1\}$ is a solid red block. It follows that such desired configuration will occur infinitely many times with probability one, and if this configuration appears, blue has a positive probability to take over the entire solid red block before it expands further.

We will make use of a regenerative structure and we will explain the first step in full detail. Note that whenever blue takes over a red site, it never creates new holes, and by the same argument as in the red-only case, one can see that $P_{\lambda}(\tilde{\tau}_0 < \infty) = 1$, starting from any initial configuration. If $\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} \leq K$, then we are done. Now suppose at $\tilde{\tau}_0$, we have $\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} = M > K$. We run the chase-escape model for another

$$N_1 := \left\lfloor \frac{M(1-2\epsilon)\lambda\nu}{\sigma} \right\rfloor$$

red cycles. Let R_t and τ_k be defined as above for the one-type case at rate λ , and let B_t denote the frontier at time t of another independent one-type process of rate-1. These two one-type models are coupled with the two-type chase-escape model $(\tilde{R}_t, \tilde{B}_t)_{t\geq 0}$ so that (i) passage times for $(R_t)_{t\geq 0}$ (resp., $(B_t)_{t\geq 0}$) are the same as those used by the red (resp., blue) particles in $(\tilde{R}_t, \tilde{B}_t)_{t\geq 0}$; (ii) $(B_t)_{t\geq 0}$ starts from the same configuration as $\tilde{B}_{\tilde{\tau}_0}$ and (iii) $(R_t)_{t\geq 0}$ starts from a solid red block that spans from \mathcal{H}_0 to $\mathcal{H}_{\tilde{R}_{\tilde{\tau}_0}}$ (and hence $\tau_0 = 0$).

We observe that, due to the memoryless property of exponential random variables,

$$P_{\lambda}(\tilde{R}_{\tilde{\tau}_{N_{1}}} - \tilde{B}_{\tilde{\tau}_{N_{1}}} \le M\alpha \mid \tilde{R}_{\tilde{\tau}_{0}} - \tilde{B}_{\tilde{\tau}_{0}} = M)$$

$$\ge P_{\lambda}(R_{\tau_{N_{1}}-\tau_{0}} - B_{\tau_{N_{1}}-\tau_{0}} \le M\alpha - M, \ B_{\tau_{N_{1}}-\tau_{0}} < M).$$
(2.4)

This is because in the right side of the inequality, $\{B_{\tau_{N_1}-\tau_0} < M\}$ corresponds to the event that in the two independent one-type models, $(R_t)_{t\geq 0}$ finishes N_1 red cycles before the blue frontier B_t advances for a distance M. Then, choose any $\epsilon < \frac{1-\lambda}{4}$ and $\alpha := 1 - 4\lambda\epsilon^2 \in (0, 1)$ fixed, we have (2.4) is at least

$$P_{\lambda} \left(R_{\tau_{N_1} - \tau_0} < \lambda M (1 - 4\epsilon^2), \ M (1 - 4\epsilon) \le B_{\tau_{N_1} - \tau_0} \le M \right).$$

Since $\tau_k - \tau_{k-1}$ has an exponential tail, the total time $(\tau_{N_1} - \tau_0)$ satisfies

$$P_{\lambda}(|\tau_{N_1}-\tau_0-N_1\sigma|\leq\epsilon N_1\sigma)\geq 1-e^{-CN_1},\quad\text{for some }C>0.$$

On the event $\{|\tau_{N_1} - \tau_0 - N_1\sigma| < \epsilon N_1\sigma\}$, the blue displacement during a time period $N_1(1 \pm \epsilon)\sigma$ is within the interval

$$\left[\frac{N_1(1-2\epsilon)\sigma}{\lambda\nu}, \frac{N_1(1+2\epsilon)\sigma}{\lambda\nu}\right] \subseteq [M(1-4\epsilon), M]$$

ECP 25 (2020), paper 22.

http://www.imstat.org/ecp/

with probability $\geq 1 - C_2 e^{-C_3 N_1}$, due to (2.3). We thus have

$$P_{\lambda}\left(R_{\tau_{N_{1}}-\tau_{0}} < \lambda M(1-4\epsilon^{2}), \ M(1-4\epsilon) \le B_{\tau_{N}-\tau_{0}} \le M\right)$$
$$\ge P_{\lambda}\left(R_{N_{1}\sigma(1+\epsilon)} < \lambda M(1-4\epsilon^{2}), \ |\tau_{N_{1}}-\tau_{0}-N_{1}\sigma| < \epsilon N_{1}\sigma\right)(1-C_{2}e^{-C_{3}N})$$

According to (2.3), the red frontier extends by a distance at most $N_1\sigma(1+2\epsilon)/\nu$ with a probability $\geq 1 - Ae^{-BN_1}$. Combining all these estimates above, we have for some $C_0 = C_0(\epsilon), \theta = \theta(\epsilon)$, with probability at least

$$(1 - C_2 e^{-C_3 N_1})(1 - A e^{-BN_1} - e^{-CN_1}) = 1 - C_0 e^{-\theta M},$$

the distance between the red and blue frontiers after N_1 red cycles is less than $M\alpha$, and moreover, at $\tilde{\tau}_{N_1}$, the rightmost blue particle is following a solid red block. We iterate this procedure for $s_M := \min\{j: M\alpha^j \leq K\} = \lceil \log_{1/\alpha}(M/K) \rceil$ times. For any M > K, after s_M iterations, the distance between the red and blue frontiers is no more than K with probability at least

$$\delta_0 := \inf_{M > K} (1 - C_0 e^{-\theta M}) (1 - C_0 e^{-\theta M \alpha}) \cdots (1 - C_0 e^{-\theta M \alpha^{s_M - 1}}).$$

To see $\delta_0 > 0$, we note that $K \leq M \alpha^{s_M - 1} \leq K / \alpha$, and the right side is at least

$$(1 - C_0 e^{-\theta K(1/\alpha)^{s_M - 1}}) \cdots (1 - C_0 e^{-\theta K}) \ge \prod_{n=0}^{\infty} (1 - C_0 e^{-\theta K(1/\alpha)^n}) := \prod_{n=0}^{\infty} (1 - u_n),$$

where $u_n = C_0 e^{-\theta K(1/\alpha)^n} \in (0,1)$ if we pick K large enough. The infinite product $\prod_{n=0}^{\infty} (1-u_n)$ converges to a positive number if and only if $\sum_{n=0}^{\infty} u_n$ converges, which can be easily verified using a ratio test:

$$\frac{u_{n+1}}{u_n} = e^{-\theta K(1/\alpha)^n (1/\alpha - 1)} \le e^{-\theta K(1/\alpha - 1)} < 1.$$

Therefore, we have

$$\begin{split} P(\text{at some } T < \infty, \hat{R}_T - \hat{B}_T \leq K) \\ \geq \sum_{M=K+1}^{\infty} P(\text{at some } T < \infty, \tilde{R}_T - \tilde{B}_T \leq K | \tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} = M) P(\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} = M) \\ + P(\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} \leq K) \\ \geq \delta_0 [1 - P(\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} \leq K)] + P(\tilde{R}_{\tilde{\tau}_0} - \tilde{B}_{\tilde{\tau}_0} \leq K) > \delta > 0, \end{split}$$

and the proof is complete.

3 Percolation-like chase-escape

Before getting to the proof of Theorem 1.4, we discuss the chase-escape dynamics on the infinite path $[-1,\infty) \cap \mathbb{Z}$ with -1 blue and 0 red initially. Suppose that all red passage times are 1 and that blue chases according to (1.3). Let t_i^B be the passage time for blue along the edge (i-1,i). Define

$$T_n = \sum_{i=0}^n t_i^B.$$

We will suppress the dependence on p and m during the proof and denote the probability measure on this process by $P(\cdot) = P_{p,m}(\cdot)$.

Lemma 3.1. Let $B = \{t_0 = m\}$. For any $\epsilon > 0$, there exists $M = M(\epsilon)$ such that for all n, m > M,

$$P(T_n \le n+1 \mid B) \le (p+\epsilon)^n.$$

Proof. If there are more than (n + 1 - m)/m edges with passage time m then $T_n > n + 1$ conditional on B. Let $a_{n,m} = \lceil (n + 1 - m)/m \rceil$ and $X \stackrel{d}{=} Bin(n, 1 - p)$. It follows from the previous observation that

$$P(T_n \le n+1 \mid B) \le P(X \le a_{n,m}).$$

Taking $m \ge 2$, it follows from a union bound and that P(X = k) is an increasing function for $k \in [0, n/2]$ that

$$P(X \le (n+1)/m) \le a_{n,m} P(X = a_{n,m}) = a_{n,m} {n \choose a_{n,m}} p^{n-a_{n,m}}.$$

Stirling's formula implies that given $\epsilon' > 0$, there exists a $\delta > 0$ such that for all $n > N(\delta)$ it holds that

$$(n+1)\binom{n}{\delta n} \le (1+\epsilon')^n.$$

Let ϵ' and δ be small enough so that $(1 + \epsilon')p^{1-\delta} < (p + \epsilon)$. Take $M(\delta)$ large enough so that, for $m > M(\delta)$, we have $a_{n,m} \leq \delta n$ for all $n \geq N(\delta)$. Setting $M = \max\{N(\delta), M(\delta)\}$ gives for n, m > M

$$P(X \le a_{n,m}) \le [(1+\epsilon')p^{(1-\delta)}]^n \le (p+\epsilon)^n.$$

Since this is also a bound for $P(T_n \le n+1 \mid B)$, the proof is complete.

Proof of Theorem 1.4. The idea is to set up and then apply a union bound that uses the one-dimensional bound in Lemma 3.1 and also an observation that bounds the probability of a continuous path in the cluster of sites accessible by red via 1-passage times.

First we must choose p and d appropriately. Let $p = d^{-1} + d^{-2}$ and d > 3 be large enough so that $p > p_c(d)$ and also

$$1 - (1 - p)^d < 1 - 3d^{-1}. ag{3.1}$$

Such a *d* exists by the main result of [10], which proved $p_c(d) = d^{-1} + d^{-3} + O(d^{-4})$. Furthermore, for this fixed choice of *d*, let $\epsilon = d^{-2}$ so that for $\delta = d^{-1} + 6d^{-2}$ we have

$$d(d^{-1} + d^{-2} + \epsilon)(1 - 3d^{-1}) = 1 - \delta.$$
(3.2)

We describe two events whose occurrence is necessary for red to survive. We must condition on the event $B = \{t_{e}^{B} = m\}$ i.e., that the first blue edge has passage time m. Otherwise, red is instantly caught. Let C be the cluster of vertices connected to 0 via red edges with passage time 1 and $C = \{|C| = \infty\}$. Our choice of p is larger than $p_{c}(d)$, thus P(C) > 0. Since red can only survive if it reaches infinity in the absence of being chased and B and C are independent, we have

$$P(A) = P(A \mid B, C)P(B)P(C).$$

Recall that A is the event that infinitely sites are at some point colored red. As P(B) = 1 - p, our result will follow from proving that $P(A \mid B, C) \to 1$ as $m \to \infty$.

Let γ_n be an oriented path started at 0 with *n* vertices and edges. Define the event $R(\gamma)$ that every vertex of γ belongs to C. Let $T(\gamma_n)$ be the passage times for blue along

ECP 25 (2020), paper 22.

the edges of γ_n . If blue never reaches the red boundary, then red spreads as if it is not being chased and survives on the event C. Since red spreads at unit speed along C, we can express this idea formally as

$$P(A \mid B, C) \ge 1 - P(\bigcup_{n \ge 1, \gamma_n} \{T(\gamma_n) \le n+1\} \cap R(\gamma_n) \mid B, C).$$

Using a union bound and independence of red and blue passage times gives

$$P(A \mid B, C) \ge 1 - \sum_{n \ge 1} \sum_{\gamma_n} P(T(\gamma_n) \le n + 1 \mid B) P(R(\gamma_n) \mid C).$$

Noting that $P(T(\gamma_n) \le n+1 \mid B) = 0$ for n < m, we then have

$$P(A \mid B, C) \ge 1 - \sum_{n \ge m} \sum_{\gamma_n} P(T(\gamma_n) \le n + 1 \mid B) P(R(\gamma_n) \mid C).$$

The $T(\gamma_n)$ are identically distributed with the same distribution as T_n from Lemma 3.1, and there are d^n oriented paths of length n. So, the above line reduces to

$$P(A \mid B, C) \ge 1 - \sum_{n \ge m} d^n P(T_n \le n + 1 \mid B) P(R(\gamma_n) \mid C).$$

It then follows from the bound in Lemma 3.1 that for any $\epsilon > 0$ and m sufficiently large we have

$$P(A \mid B, C) \ge 1 - \sum_{n \ge m} d^n (p + \epsilon)^n P(R(\gamma_n) \mid C).$$
(3.3)

A simple way to bound $P(R(\gamma_n) \mid C)$ is to observe that $R(\gamma_n)$ can only occur if each vertex of γ_n has at least one incident edge with red passage time 1. This occurs with probability $1 - (1 - p)^d$ at each of the *n* vertices of γ_n . Let c = 1/P(C). Using the trivial bound $P(R(\gamma_n) \cap C) \leq P(R(\gamma_n))$ and the formula for conditional probability, it follows that for $||x||_1 = n$

$$P(R(\gamma_n) \mid C) \le \frac{P(R(\gamma_n))}{P(C)} \le c(1 - (1 - p)^d)^n.$$

Our choice of d at (3.1) guarantees that $1 - (1 - p)^d < 1 - 3d^{-1}$. Thus,

$$P(R(\gamma_n) \mid C) \le c(1 - 3d^{-1})^n.$$
(3.4)

Applying (3.4) to (3.3) and recalling that $p = d^{-1} + d^{-3}$ gives

$$P(A \mid B, C) \ge 1 - c \sum_{n \ge m} [d(d^{-1} + d^{-3} + \epsilon)(1 - 3d^{-1})]^n.$$

Our choice of ϵ at (3.2) yields

$$P(A \mid B, C) \ge 1 - c \sum_{n \ge m} (1 - \delta)^n$$

This lower bound converge to 1 as $m \to \infty$, which completes the proof.

Acknowledgments. The second author was partially supported by NSF DMS grant 1855516.

ECP 25 (2020), paper 22.

 \square

References

- Daniel Ahlberg, Asymptotics of first-passage percolation on one-dimensional graphs, Advances in Applied Probability 47 (2015), no. 1, 182–209. MR-3327321
- [2] Daniel Ahlberg, Existence and coexistence in first-passage percolation, 2020.
- [3] Daniel Ahlberg and Christopher Hoffman, Random coalescing geodesics in first-passage percolation, 2016.
- [4] David Aldous and William B. Krebs, The birth-and-assassination process, Statistics & Probability Letters 10 (1990), 427–430. MR-1078244
- [5] Antonio Auffinger, Michael Damron, and Jack Hanson, 50 years of first-passage percolation, vol. 68, American Mathematical Soc., 2017. MR-3729447
- [6] Erin Beckman, Keisha Cook, Nicole Eikmeier, Sarai Hernandez-Torres, and Matthew Junge, Chase-escape with death on trees, 2019.
- [7] Charles Bordenave, On the birth-and-assassination process, with an application to scotching a rumor in a network, Electron. J. Probab. **13** (2008), 2014–2030. MR-2453554
- [8] Charles Bordenave, Extinction probability and total progeny of predator-prey dynamics on infinite trees, Electron. J. Probab 19 (2014), no. 20, 1–33. MR-3167884
- [9] J. Theodore Cox and Richard Durrett, Some limit theorems for percolation processes with necessary and sufficient conditions, Ann. Probab. **9** (1981), no. 4, 583–603. MR-0624685
- [10] J Theodore Cox and Richard Durrett, Oriented percolation in dimensions $d \ge 4$: bounds and asymptotic formulas, Mathematical Proceedings of the Cambridge Philosophical Society, vol. 93, Cambridge University Press, 1983, pp. 151–162. MR-0684285
- [11] Michael Damron and Jack Hanson, Busemann functions and infinite geodesics in twodimensional first-passage percolation, Communications in Mathematical Physics 325 (2014), no. 3, 917–963. MR-3152744
- [12] Richard Durrett, Oriented percolation in two dimensions, The Annals of Probability (1984), 999–1040. MR-0757768
- [13] Rick Durrett, Probability: theory and examples, vol. 49, Cambridge university press, 2019. MR-3930614
- [14] Olivier Garet and Régine Marchand, Coexistence in two-type first-passage percolation models, Ann. Appl. Probab. 15 (2005), no. 1A, 298–330. MR-2115045
- [15] Olivier Garet and Régine Marchand, Competition between growths governed by Bernoulli Percolation, Markov Processes and Related Fields (2006), 695–734. MR-2284399
- [16] Olle Häggström and Robin Pemantle, First passage percolation and a model for competing spatial growth, Journal of Applied Probability 35 (1998), no. 3, 683–692. MR-1659548
- [17] Olle Häggström and Robin Pemantle, Absence of mutual unbounded growth for almost all parameter values in the two-type richardson model, Stochastic Processes and their Applications 90 (2000), no. 2, 207–222. MR-1794536
- [18] Christopher Hoffman, Coexistence for richardson type competing spatial growth models, Ann. Appl. Probab. 15 (2005), no. 1B, 739–747. MR-2114988
- [19] Christopher Hoffman, Geodesics in first passage percolation, The Annals of Applied Probability (2008), 1944–1969. MR-2462555
- [20] George Kordzakhia, The escape model on a homogeneous tree, Electron. Commun. Probab.
 10 (2005), 113–124. MR-2150700
- [21] George Kordzakhia and Steven Lalley, A two-species competition model on \mathbb{Z}^d , Stochastic Processes and their Applications **115** (2006), 781–796. MR-2132598
- [22] Igor Kortchemski, A predator-prey SIR type dynamics on large complete graphs with three phase transitions, Stochastic Processes and their Applications 125 (2015), no. 3, 886 – 917. MR-3303961
- [23] Igor Kortchemski, Predator-prey dynamics on infinite trees: A branching random walk approach, Journal of Theoretical Probability 29 (2016), no. 3, 1027–1046. MR-3540488

Coexistence in chase-escape

- [24] Daniel Richardson, Random growth in a tessellation, Mathematical Proceedings of the Cambridge Philosophical Society, vol. 74, Cambridge University Press, 1973, pp. 515–528. MR-0329079
- [25] S. Tang, G. Kordzakhia, and S. P. Lalley, Phase Transition for the Chase-Escape Model on 2D Lattices, ArXiv e-prints: arXiv:1807.08387 (2018). MR-3824972