

Distribution of the particles of a critical branching Wiener process

PÁL RÉVÉSZ

Technische Universität Wien, Institut für Statistik, Wiedner Hauptstr. 8–10, A-1040 Vienna, Austria

Consider a critical branching Wiener process in \mathbb{R}^d . Let $\{F_T(x), T=1, 2, \dots, x \in \mathbb{R}^d\}$ be the distribution of the particles living at time T . The main result of this paper tells us that any given absolutely continuous function $F(x)$ will be well approximated by $F_T(x)$ with positive probability if T is big enough and the process does not die out up to T .

Keywords: critical branching Wiener process, empirical distribution of particles, limit theorems, measure-valued process

1. Introduction

Consider the following model:

- (i) A particle starts from position $\mathbf{0} \in \mathbb{R}^d$ and executes a Wiener process $W(t) \in \mathbb{R}^d$.
- (ii) Arriving at time $t = 1$ at the new location $W(1)$, it dies.
- (iii) At death it is replaced by Y off-springs where

$$\mathbf{P}\{Y = \ell\} = p_\ell \quad (\ell = 0, 1, 2, \dots)$$

and

$$p_\ell \geq 0, \quad \sum_{\ell=0}^{\infty} p_\ell = 1.$$

(iv) Each offspring, starting from where its ancestor dies, executes a Wiener process (from its starting point) and repeats the above given steps and so on. All Wiener processes and offspring numbers are assumed independent of one another.

A more formal definition is given in Révész (1994, p. 91).

Let $A \subset \mathbb{R}^d$ be a Borel set and let $\lambda(A, t)$ ($t = 0, 1, 2, \dots$) be the number of particles located in A at time t . Then

$$B(t) = \lambda(\mathbb{R}^d, t)$$

is the number of particles living at t and $\{B(t), t = 0, 1, 2, \dots\}$ is a branching process. From now on we assume that

$$1 \leq m = \sum_{k=0}^{\infty} k p_k < \infty$$

and

$$0 < \sigma^2 = \sum_{k=0}^{\infty} (k-m)^2 p_k < \infty.$$

It is well known (cf. Athreya and Ney 1972) that

$$\lim_{t \rightarrow \infty} \frac{B(t)}{m^t} = B \quad \text{a.s.}$$

where

$$P\{B = 0\} = \begin{cases} 1 & \text{if } m = 1, \\ q & \text{if } m > 1 \end{cases}$$

and $q < 1$ depends on the distribution $\{p_k\}$.

On the limit properties of $\lambda(A, t)$, ($t \rightarrow \infty$) in the case $m > 1$ we have the following theorem:

Theorem A. (Révész 1994, Theorem 6.4, p. 107). For any $x \in \mathbb{R}^d$ and $\epsilon > 0$

$$\lim_{T \rightarrow \infty} T^{1/2-\epsilon} \left| \frac{\lambda(\{y : y \leq xT^{1/2}\}, T)}{m^T} - B\Phi(x) \right| = 0 \quad \text{a.s.} \quad (1.1)$$

where

$$\Phi(x) = \Phi(x_1, x_2, \dots, x_d) = \frac{1}{(2\pi)^{d/2}} \int_{-\infty}^{x_d} \dots \int_{-\infty}^{x_1} \exp\left(-\frac{y_1^2 + y_2^2 + \dots + y_d^2}{2}\right) dy_1 dy_2 \dots dy_d.$$

Further, for any fixed $x \in \mathbb{R}^d$ and $0 < \epsilon < 1$ we have

$$\lim_{T \rightarrow \infty} T^{1-\epsilon} \left| (2\pi T)^{d/2} \frac{\lambda(C(x), T)}{m^T} - B \right| = 0 \quad \text{a.s.} \quad (1.2)$$

where

$$C(x) = \{y : \|y - x\| \leq r_d\}$$

and

$$r_d = \begin{cases} 2^{-1} & \text{if } d = 1, \\ \pi^{-1/2} & \text{if } d = 2, \\ \pi^{-1/2} (\Gamma(d/2 + 1))^{1/d} & \text{if } d \geq 3, \end{cases}$$

i.e. $C(x)$ is the ball in \mathbb{R}^d around x of volume 1.

Theorem A tells us that in the case $m > 1$ the particles at time T (if T is big enough) are distributed according to the normal law.

In the present paper we investigate the case $m = 1$. Hence from now on we assume that $m = 1$. Since if $m = 1$ then $B = 0$ a.s., we study the limit behaviour of $\lambda(A, T)$ as $T \rightarrow \infty$,

under the condition $\{B(T) > 0\}$. In this case it turns out that the distribution of the particles may agree to any given smooth enough distribution with positive probability. In fact we prove the following theorem:

Theorem 1. *Let $F(x)$ ($x \in \mathbb{R}^d$) be an arbitrary, given, absolutely continuous distribution function with bounded density. Then for any $\epsilon > 0$ there exist a $\delta = \delta(\epsilon) > 0$ and a $T_0 = T_0(\epsilon) > 0$ such that*

$$P\left\{\sup_{x \in \mathbb{R}^d} |F_T(x) - F(x)| \leq \epsilon | B(T) > 0\right\} \geq \delta$$

if $T \geq T_0$, where

$$F_T(x) = \frac{\lambda(\{y : y \leq xT^{1/2}\}, T)}{B(T)}.$$

We are also interested in studying the expectation vector of the empirical distribution $F_T(x) = F_T(x_1, x_2, \dots, x_d)$. Let

$$F_T^{(i)}(x_i) = F_T(\infty, \dots, \infty, x_i, \infty, \dots, \infty) = \frac{\lambda(\{y = (y_1, \dots, y_d) : y_i \leq x_i T^{1/2}\}, T)}{B(T)},$$

$$m_i = m_i(T) = \int_{-\infty}^{+\infty} x_i dF_T^{(i)}(x_i), \quad (i = 1, 2, \dots, d)$$

and

$$m = m(T) = (m_1, m_2, \dots, m_d).$$

Then we have the following theorem:

Theorem 2.

- (i) m_1, m_2, \dots, m_d are i.i.d. random variables.
- (ii) $Em_i = 0$ ($i = 1, 2, \dots, d$).
- (iii) m_1, m_2, \dots, m_d are normally distributed.
- (iv) $\lim_{T \rightarrow \infty} Em_i^2 = \sigma^2/2$.

Theorem 1 is an analogue of (1.1). We are also interested in finding an analogue of (1.2), i.e. we intend to study the limit properties of $\lambda(C, T)$ where $C = C(0)$.

First we mention the following trivial facts:

$$E(\lambda(C, T) | B(T)) \sim B(T)(2\pi T)^{-d/2}$$

and

$$E(\lambda(C, T) | B(T) > 0) \sim \frac{\sigma^2 T}{2} (2\pi T)^{-d/2}.$$

Now we formulate our next theorem.

Theorem 3. (i) *In the case $d = 1$*

$$\lim_{T \rightarrow \infty} T^{-1/2} E(\lambda(C, T) | \lambda(C, T) > 0) = \mathcal{L}_1$$

where

$$\mathcal{L}_1 = \frac{\sigma^2}{2(2\pi)^{1/2}} \sum_{k=1}^{\infty} \int_0^1 \cdots \int_0^1 \left(\frac{x_1 \cdots x_k}{2 - x_1 \cdots x_k} \right)^{1/2} dx_1 \cdots dx_k = \frac{\sigma^2}{4} \left(\frac{\pi}{2} \right)^{1/2}.$$

(ii) In the case $d = 2$, for any T big enough,

$$\ell_2 \log T \leq \mathbb{E}(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) \leq L_2 \log T$$

where

$$\ell_2 = \frac{\sigma^2}{8\pi}, \quad L_2 = \frac{\sigma^2}{4\pi}.$$

(iii) In the case $d = 3$, for any T big enough,

$$\ell_3 \leq \mathbb{E}(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) \leq L_3$$

where

$$\ell_3 = \frac{\sigma^2}{\pi^{3/2} 2^5}, \quad L_3 = \frac{\sigma^2}{\pi^{3/2} 2^{7/2}}.$$

(iv) In the case $d \geq 4$, for any T big enough,

$$\ell_d \leq \mathbb{E}(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) \leq L_d$$

where the values of L_d and ℓ_d are not given exactly.

Since

$$P\{\lambda(\mathcal{C}, T) > 0 | B(T) > 0\} = \frac{\mathbb{E}(\lambda(\mathcal{C}, T) | B(T) > 0)}{\mathbb{E}(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0)},$$

Theorem 3 clearly implies the following result:

Theorem 4. In the case $d = 1$

$$\lim_{T \rightarrow \infty} P\{\lambda(\mathcal{C}, T) > 0 | B(T) > 0\} = \frac{\sigma^2}{2\mathcal{L}_1(2\pi)^{1/2}} = \frac{2}{\pi}.$$

In the case $d \geq 2$, for any T big enough,

$$P\{\lambda(\mathcal{C}, T) > 0 | B(T) > 0\} \leq \begin{cases} \frac{\sigma^2}{4\pi\ell_2 \log T} = \frac{2}{\log T} & \text{if } d = 2, \\ \frac{\sigma^2 T}{2\ell_3} (2\pi T)^{-3/2} = 2^{5/2} T^{-1/2} & \text{if } d = 3, \\ \frac{\sigma^2 T}{2\ell_d} (2\pi T)^{-d/2} & \text{if } d \geq 4 \end{cases}$$

and

$$P\{(C, T) > 0 | B(T) > 0\} \geq \begin{cases} \frac{\sigma^2}{4\pi L_2 \log T} \approx \frac{1}{\log T} & \text{if } d = 2, \\ \frac{\sigma^2 T}{2L_3} (2\pi T)^{-3/2} = 2T^{-1/2} & \text{if } d = 3, \\ \frac{\sigma^2 T}{2L_d} (2\pi T)^{-d/2} & \text{if } d \geq 4. \end{cases}$$

It is also easy to see by Theorem 3 that the following result holds:

Theorem 5. *Uniformly in T*

$$\begin{aligned} \lim_{K \rightarrow \infty} P\{\lambda(C, T) > KT^{1/2} | \lambda(C, T) > 0\} &= 0 & \text{if } d = 1, \\ \lim_{K \rightarrow \infty} P\{\lambda(C, T) > K \log T | \lambda(C, T) > 0\} &= 0 & \text{if } d = 2, \\ \lim_{K \rightarrow \infty} P\{\lambda(C, T) > K | \lambda(C, T) > 0\} &= 0 & \text{if } d \geq 3. \end{aligned}$$

Remark. Theorems 3–5, in the case $d = 1$ ($d = 2$) are closely related to Theorem 3.11 (2.11) of Fleischman (1978).

2. Lemmas on the critical branching process

Lemma A. *For any $t \geq 0$*

$$EB(t) = 1, \quad (2.1)$$

and as $t \rightarrow \infty$

$$P\{B(t) > 0\} \sim \frac{2}{\sigma^2 t}, \quad (2.2)$$

$$E(B(t) | B(t) > 0) = \frac{1}{P\{B(t) > 0\}} \sim \frac{\sigma^2 t}{2}, \quad (2.3)$$

$$\lim_{t \rightarrow \infty} P\left\{\frac{B(t)}{t} > z | B(t) > 0\right\} = \exp\left(-\frac{2z}{\sigma^2}\right) \quad (z \geq 0). \quad (2.4)$$

For any $0 \leq s < t < \infty$, let $Q(s, t)$ be the number of those particles which are living at time s and which have at least one offspring living at time t . Clearly

$$\begin{aligned} B(s) &\geq Q(s, t), & B(t) &\geq Q(s, t), \\ \{Q(s, t) = 0\} &= \{B(t) = 0\} & (0 \leq s < t) \end{aligned}$$

and as a function of s , $Q(s, t)$ is non-decreasing. Hence on the set $\{B(t) > 0\}$ one can define a sequence $\nu_2 = \nu_2(t) \leq \nu_3 = \nu_3(t) \leq \dots \leq \nu_\mu = \nu_\mu(t) < t$ and a random variable $\mu = \mu(t)$

as follows:

$$\nu_k = \inf \{s : 0 \leq s \leq t, Q(s, t) \geq k\}$$

and μ is the largest integer for which $\nu_\mu < t$.

Lemma B (Révész, 1994).

$$E(Q(s, t) | B(t) > 0) = \frac{1}{P\{B(t) > 0\}} \frac{1}{E(B(t-s) | B(t-s) > 0)} = \frac{t}{t-s} (1 + o(t-s)). \quad (2.5)$$

$$E(Q^2(s, t) | B(t) > 0) = \frac{t(t+s)}{(t-s)^2} (1 + o(t-s)) \quad (2.6)$$

and for any fixed $k = 2, 3, \dots$,

$$E((t - \nu_k) | B(t) \geq k) \sim \frac{t}{k}, \quad (2.7)$$

$$E((t - \nu_k)^2 | B(t) \geq k) \sim \frac{2t^2}{k(k+1)}, \quad (2.8)$$

$$\lim_{t \rightarrow \infty} P\left\{\frac{\nu_k}{t} < x | B(t) \geq k\right\} = x^{k-1} \quad (0 \leq x \leq 1). \quad (2.9)$$

Consider a fixed $[0, T]$ -branch $\{P_0, P_1, \dots, P_T\}$ of the underlying branching process, i.e. P_0 is the particle at 0, P_1 is an offspring of P_0, \dots, P_i is an offspring of P_{i-1} ($i = 1, 2, \dots, T$). Clearly such a branch exists if and only if $B(T) > 0$. Let $\xi_1 = \xi_1(T)$ be the first time-point where a new $[\xi_1, T]$ -branch starts which lives up to time T , i.e. ξ_1 is the smallest i for which P_i has an offspring $Q_{i+1} \neq P_{i+1}$ having at least one offspring living at time T . Let $\xi_2 = \xi_2(T)$ be the second element of the fixed $[0, T]$ -branch where a new $[\xi_2, T]$ -branch starts which lives up to time T , i.e. ξ_2 is the smallest j for which $j > \xi_1$ and P_j has an offspring $Q_{j+1} \neq P_{j+1}$ which has an offspring living at time T . Continuing this procedure, we get a random sequence $1 \leq \xi_1 < \xi_2 < \dots < \xi_\nu < T$ where $\nu = \nu(T)$ is the largest integer for which $\xi_\nu < T$.

Now consider a partition of the $B(T) - 1$ particles living at time T , not considering the terminal point P_T of the fixed $[0, T]$ -branch. The first class $C_1 = C_1(T)$ consists of the terminal points of those $[\xi_1, T]$ -branches which branch from the fixed $[0, T]$ -branch at ξ_1 . $C_2 = C_2(T)$ consists of the terminal points of those $[\xi_2, T]$ -branches which branch from the fixed $[0, T]$ -branch at ξ_2 , etc.

Let U_1, U_2, \dots be a sequence of independent random variables uniformly distributed on $[0, 1]$ and introduce the following notation:

$$V_1 = 1,$$

$$V_k = \prod_{j=1}^{k-1} (1 - U_j) \quad (k = 2, 3, \dots),$$

$$L_k = \sum_{j=1}^k U_j V_j = 1 - \prod_{j=1}^k (1 - U_j) \stackrel{\text{a}}{=} 1 - \prod_{j=1}^k U_j \quad (k = 1, 2, \dots).$$

Then we have the following lemma:

Lemma 1. For any k fixed and $\epsilon > 0$

$$\lim_{T \rightarrow \infty} P \left\{ \frac{\xi_1}{T} < x_1, \frac{\xi_2}{T} < x_2, \dots, \frac{\xi_k}{T} < x_k \mid B(T) > 0 \right\} = P\{L_1 < x_1, L_2 < x_2, \dots, L_k < x_k\}, \quad (2.10)$$

$$\lim_{T \rightarrow \infty} (T - \xi_k)^{-1} \mathbf{E}(|C_k| \mid \xi_k, B(T) > 0) = \frac{\sigma^2}{2}, \quad (2.11)$$

$$\lim_{T \rightarrow \infty} P \left\{ \left| \frac{\nu(T)}{\log T} - 1 \right| \geq \epsilon \mid B(T) > 0 \right\} = 0, \quad (2.12)$$

$$\lim_{T \rightarrow \infty} \mathbf{E} \left(\frac{\nu(T)}{\log T} \mid B(T) > 0 \right) = 1 \quad (2.13)$$

and

$$\lim_{T \rightarrow \infty} T^{-1/2} \mathbf{E} \left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T} \right)^{-1/2} \right) = 2, \quad (2.14)$$

where $|C_k|$ is the cardinality of C_k .

In order to prove Lemma 1, we first prove the two following lemmas.

Lemma 2. For any $0 < \epsilon < 1$, let

$$A_n = A_n(\epsilon) = \{(x_1, x_2, \dots, x_n) : 0 < x_i < 1, x_1 x_2 \dots x_n > \epsilon\}$$

and

$$I_n = I_n(\epsilon) = \int_{A_n} (x_1 x_2 \dots x_n)^{-1/2} dx_1 dx_2 \dots dx_n.$$

Then we have

$$\lim_{\epsilon \searrow 0} \epsilon^{1/2} \sum_{n=1}^{\infty} I_n = 2.$$

Proof. Let

$$J_n = J_n(\epsilon) = \int_{A_n} (x_1 x_2 \dots x_n)^{-1} dx_1 dx_2 \dots dx_n.$$

Then

$$\begin{aligned} I_n &= \int_{A_{n-1}} (x_1 x_2 \dots x_{n-1})^{-1/2} \left(\int_{\epsilon(x_1 x_2 \dots x_{n-1})^{-1}}^1 x_n^{-1/2} dx_n \right) dx_1 dx_2 \dots dx_{n-1} \\ &= 2 \int_{A_{n-1}} (x_1 x_2 \dots x_{n-1})^{-1/2} (1 - \epsilon^{1/2} (x_1 x_2 \dots x_{n-1})^{-1/2}) dx_1 dx_2 \dots dx_{n-1} \\ &= 2I_{n-1} - 2\epsilon^{1/2} J_{n-1}. \end{aligned}$$

By a simple calculation we get

$$I_1 = 2(1 - \epsilon^{1/2}),$$

$$J_n = \frac{\left(\log \frac{1}{\epsilon}\right)^n}{n!}.$$

Hence, by induction,

$$I_n = 2^n \left(1 - \epsilon^{1/2} \sum_{k=0}^{n-1} \frac{\left(\frac{1}{2} \log \frac{1}{\epsilon}\right)^k}{k!} \right) = 2^n \epsilon^{1/2} \sum_{k=n}^{\infty} \frac{\left(\frac{1}{2} \log \frac{1}{\epsilon}\right)^k}{k!}.$$

Consequently

$$\sum_{n=1}^{\infty} I_n = 2\epsilon^{-1/2} - 2$$

□

and Lemma 2 is proved.

Lemma 3.

$$\mathbf{E}L_k = 1 - 2^{-k}, \quad (2.15)$$

$$\mathbf{E} \log(eU_k) = 0, \quad (2.16)$$

$$\mathbf{E}(\log(eU_k))^2 = 1, \quad (2.17)$$

$$-\infty < \mathbf{E}(\log(eU_k))^r < \infty \quad (r = 1, 2, \dots), \quad (2.18)$$

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n \log(eU_k) = 0 \quad \text{a.s.}, \quad (2.19)$$

$$\lim_{n \rightarrow \infty} P \left\{ n^{-1/2} \sum_{k=1}^n \log(eU_k) < x \right\} = \Phi(x), \quad (2.20)$$

$$\lim_{\epsilon \searrow 0} \frac{\mu(\epsilon)}{\log(1/\epsilon)} = 1 \quad \text{a.s.}, \quad (2.21)$$

$$\lim_{\epsilon \searrow 0} \mathbf{E} \left(\frac{\mu(\epsilon)}{\log(1/\epsilon)} \right) = 1, \quad (2.22)$$

$$\lim_{\epsilon \searrow 0} \epsilon^{1/2} \mathbf{E} \left(\sum_{k=1}^{\mu} (U_1 U_2 \dots U_k)^{-1/2} \right) = 2 \quad (2.23)$$

where

$$\mu(\epsilon) = \min\{k : L_k \geq 1 - \epsilon\}.$$

Proof. Expressions (2.15)–(2.18) can be obtained by trivial calculations. Equations (2.19) and (2.20) are, respectively, consequences of (2.16) and (2.17). In order to prove (2.22), observe that

$$\begin{aligned} P\{\mu > k\} &= P\{U_1 U_2 \dots U_k > \epsilon\} \\ &= P\left\{\frac{\log(eU_1) + \log(eU_2) + \dots + \log(eU_k)}{k} > 1 - \frac{\log(1/\epsilon)}{k}\right\}, \end{aligned}$$

and for any $0 < \delta < 1$ we have

$$\begin{aligned} E\mu &= \sum_{k=1}^{\infty} P\{\mu \geq k\} \\ &= \sum_{k=1}^{(1-\delta)\log(1/\epsilon)} P\{\mu \geq k\} + \sum_{k=(1-\delta)\log(1/\epsilon)}^{(1+\delta)\log(1/\epsilon)} P\{\mu \geq k\} + \sum_{k=(1+\delta)\log(1/\epsilon)}^{\infty} P\{\mu \geq k\}. \end{aligned}$$

Clearly

$$\begin{aligned} \lim_{\epsilon \searrow 0} P\{\mu \geq k\} &= 1 \text{ uniformly in } k \leq (1-\delta)\log(1/\epsilon), \\ \sum_{k=(1-\delta)\log(1/\epsilon)}^{(1+\delta)\log(1/\epsilon)} P\{\mu \geq k\} &\leq 2\delta \log(1/\epsilon), \\ \lim_{\epsilon \searrow 0} \sum_{k=(1+\delta)\log(1/\epsilon)}^{\infty} P\{\mu \geq k\} \\ &\leq \lim_{\epsilon \searrow 0} \sum_{k=(1+\delta)\log(1/\epsilon)}^{\infty} P\left\{\frac{\log(eU_1) + \log(eU_2) + \dots + \log(eU_k)}{k} > \frac{\delta}{1+\delta}\right\} = 0. \end{aligned}$$

The last three relations clearly imply (2.22).

In order to prove (2.21), we note that by (2.19) we have

$$\begin{aligned} U_1 U_2 \dots U_n &= \exp(-n(1 + o(1))) \quad \text{a.s.}, \\ U_1 U_2 \dots U_\mu &= \exp(-\mu(1 + o(1))) \leq \epsilon, \end{aligned} \tag{2.24}$$

$$U_1 U_2 \dots U_{\mu-1} = \exp(-(\mu-1)(1 + o(1))) > \epsilon; \tag{2.25}$$

in the last two equations $o(1)$ is a function of μ converging to 0 a.s. as $\mu \rightarrow \infty$, i.e. as $\epsilon \rightarrow 0$. Clearly (2.24) implies (2.25).

Let

$$B_k = \{U_1 U_2 \dots U_k > \epsilon\}.$$

Then

$$E\left(\sum_{k=1}^{\mu} (U_1 U_2 \dots U_k)^{-1/2}\right) = \sum_{k=1}^{\infty} \int_{B_k} (U_1 U_2 \dots U_k)^{-1/2} dP = \sum_{k=1}^{\infty} I_k.$$

Hence we have (2.23) by Lemma 2 and Lemma 3 is proved.

Proof of Lemma 1. Since $\xi_1 = \nu_2$, we have (2.10) for $k = 1$ by (2.9). When $k = 2$, observe that

$$\frac{\xi_2}{T} = \frac{\xi_2 - \xi_1}{T - \xi_1} \left(1 - \frac{\xi_1}{T}\right) + \frac{\xi_1}{T},$$

where

$$\lim_{T \rightarrow \infty} P \left\{ \frac{\xi_2 - \xi_1}{T - \xi_1} < x_1, \frac{\xi_1}{T} < x_2 \mid B(T) \geq 2 \right\} = x_1 x_2 \quad (0 \leq x_1, x_2 \leq 1),$$

and we have (2.10) for $k = 2$. Continuing this procedure, we complete the proof of (2.10) by induction.

(2.11) is a simple consequence of (2.3). Equations (2.12), (2.13) and (2.14) follow from Lemma 3. Lemma 1 is proved. \square

3. Proofs of Theorems 1 and 2

We introduce the following notation:

(i) $\mathcal{P}(s, t) = \{P_1^{(s)}, P_2^{(s)}, \dots, P_{Q(s,t)}^{(s)}\} \subset \mathbb{R}^d$ is the set of locations of those particles at time s which have an offspring living at time t .

(ii) $\mathcal{P}_i(s, u, t) = \{P_{1i}^{(s,u)}, P_{2i}^{(s,u)}, \dots\} \subseteq \mathbb{R}^d$ is the set of the locations of those offspring of $P_i^{(s)}$ at time u ($0 \leq su \leq t$) which have an offspring living at time t .

(iii) $\mathcal{P}(s, u, t) = \bigcup_{i=1}^{Q(s,t)} \mathcal{P}_i(s, u, t) = \mathcal{P}(u, t)$.

(iv) $X_i^{s,u} = \max_j \|P_i^{(s)} - P_{i,j}^{(s,u)}\|$.

(v) $D(s, u) = \max_{1 \leq i \leq Q(s,t)} X_i^{(s,u)}$.

Clearly $D(s, u)$ tells us how far from their ancestors the particles go during the time interval (s, u) . Note that the number of elements of $\mathcal{P}(s, u, t)$ is $Q(u, t)$.

Lemma 4. Let $0 \leq s < u \leq t < \infty$. Then

$$P\{D(s, u) \geq K(u-s)^{1/2} \mid Q(u, t)\} \leq (1 + o(u-s))Q(u, t) \exp\left(-\frac{K^2}{2}\right).$$

Proof. The probability that a particle moves at least $K(u-s)^{1/2}$ during a time interval (s, u) is

$$P\{W(u-s) \geq K(u-s)^{1/2}\} \leq (1 + o(u-s)) \exp\left(-\frac{K^2}{2}\right).$$

Since we have to take into consideration $Q(u, t)$ particles, we get Lemma 4. \square

Lemma 5. For any $\epsilon > 0$ there exists a $0 < \delta = \delta(\epsilon) < 1$ such that

$$P\left\{D\left(\frac{t}{2}, t\right) \geq \epsilon t^{1/2}\right\} \leq \delta \quad (3.1)$$

for any $t > 0$.

Proof. Observe that

$$D\left(\frac{t}{2}, t\right) \leq \sum_{k=1}^{\infty} D\left(t - \frac{t}{2^k}, t - \frac{t}{2^{k+1}}\right),$$

and applying Lemma 4 with

$$s = s_k = t - \frac{t}{2^k}, \quad u = u_k = t - \frac{t}{2^{k+1}}, \quad K = \frac{\epsilon}{5} 2^{(k+1)/4},$$

we get

$$\begin{aligned} & P\left\{D\left(t - \frac{t}{2^k}, t - \frac{t}{2^{k+1}}\right) \geq K\left(\frac{t}{2^{k+1}}\right)^{1/2} \mid Q\left(t - \frac{t}{2^{k+1}}, t\right)\right\} \\ & \leq (1 + o(2^{-k}t)) Q\left(t - \frac{t}{2^{k+1}}, t\right) \exp\left(-\frac{K^2}{2}\right). \end{aligned}$$

Hence

$$\begin{aligned} & P\left\{D\left(t - \frac{t}{2^k}, t - \frac{t}{2^{k+1}}\right) \geq \frac{\epsilon}{5} 2^{-(k+1)/4} t^{1/2} \mid Q\left(t - \frac{t}{2^{k+1}}, t\right), k = 1, 2, \dots\right\} \\ & \geq (1 + o(2^{-k}t)) Q\left(t - \frac{t}{2^{k+1}}, t\right) \exp\left(-\frac{\epsilon^2}{50} 2^{(k+1)/2}\right). \end{aligned}$$

Since

$$\frac{\epsilon}{5} t^{1/2} \sum_{k=1}^{\infty} 2^{-(k+1)/4} \leq \epsilon t^{1/2},$$

and by (2.5)

$$\sum_{k=1}^{\infty} Q\left(t - \frac{t}{2^{k+1}}, t\right) \exp\left(-\frac{\epsilon^2}{50} 2^{(k+1)/2}\right) < \frac{1}{2}$$

with positive probability, we have (3.1). \square

For any $x \in \mathbb{R}^d$ let

$$D_x\left(\frac{t}{2}\right) = \max_{1 \leq i \leq Q(t/2, t)} \|P_i^{(t/2)} - xt^{1/2}\|.$$

Clearly $D_x(t/2)$ tells us how close to $xt^{1/2}$ at time $t/2$ were located those particles which have an offspring living at time t .

Lemma 6. For any $\epsilon > 0, K > 0$ there exists a $\delta = \delta(\epsilon, K) > 0$ such that

$$P\left\{D_{Kx}\left(\frac{t}{2}\right) \leq \epsilon t^{1/2}\right\} \geq \delta,$$

provided that $\|x\| \leq 1$.

Proof. Essentially the same as that of Lemma 4. □

Lemma 7. For any $d = 1, 2, \dots, \delta > 0, K > 0$ there exists an $\epsilon = \epsilon(\delta, K) > 0$ such that

$$P\{\lambda(\tilde{C}(x, \delta t^{1/2}), t) = 0 \mid B(t) > 0\} \geq \epsilon,$$

where

$$C(a, r) = \{x : x \in \mathbb{R}^d, \|x - a\| \leq r\},$$

$$\|x\| \leq Kt^{1/2},$$

and

$$\tilde{C}(\cdot, \cdot) = \mathbb{R}^d - C(\cdot, \cdot).$$

Lemma 7 tells us that for any $x \in \mathbb{R}^d$ with $\|x\| \leq Kt^{1/2}$, all particles living at time t will be located in a ball around x of radius $\delta t^{1/2}$ with positive probability.

Proof. The lemma is a trivial consequence of Lemmas 5 and 6. □

Proof of Theorem 1. Clearly for an $\epsilon > 0$ there exist a $1 < \ell = \ell(\epsilon) < \infty$, a $C = C(\epsilon) > 0$ and a partition A_1, A_2, \dots, A_ℓ of \mathbb{R}^d such that

$$\int_{A_i} dF(x) \leq \epsilon \quad (i = 1, 2, \dots, \ell)$$

and an A_i contains a ball of radius $C\epsilon$.

Let

$$\nu_\ell = \inf\{s : 0 \leq s \leq t, Q(s, t) = \ell = \ell(\epsilon)\}$$

and let the locations of these ℓ particles at time ν_ℓ be

$$Q_1, Q_2, \dots, Q_\ell.$$

Observe that

$$P\{\max_{1 \leq i \leq \ell} \|Q_i\| \leq t^{1/2}\} \geq \delta$$

and that by (2.9)

$$P\{t - \nu_\ell \geq \epsilon t\} \geq \delta$$

for some $\delta = \delta(\epsilon) > 0$ if ℓ is big enough, say $\ell > 1/\epsilon$.

Let $B_i(t - \nu_\ell)$ be the number of offspring of the particle located in Q_i at time ν_ℓ . Then

$$P\{\max_{1 \leq i \leq j \leq \ell} |B_i(t - \nu_\ell) - B_j(t - \nu_\ell)| \leq \epsilon^2 t\} \geq \delta.$$

Then by Lemma 7 all $B_i(t - \nu_i)$ offspring of the particle located in Q_i at time ν_i will be located in A_i with positive probability. Hence we have Theorem 1. \square

In order to prove Theorem 2 we first prove the following:

Lemma 8. *Let $d = 1$ and $P_1, P_2, \dots, P_{B(T)}$ be the locations of the particles at time T . Then we have*

$$\lim_{T \rightarrow \infty} T^{-3} \mathbb{E}((P_1 + P_2 + \dots + P_{B(T)})^2 | B(T) > 0) = \frac{\sigma^2}{2}.$$

Proof. Let P_1 be fixed and consider a point $P_j \in C_k$ (cf. Lemma 1). Then

$$\mathbb{E}(P_1 P_j | \xi_k) = \xi_k.$$

Hence, by (2.11),

$$\mathbb{E}\left(\sum_{j \in C_k} P_1 P_j \mid \xi_k\right) \sim \frac{\sigma^2}{2} \xi_k (T - \xi_k)$$

and

$$\begin{aligned} \mathbb{E}\left(\sum_{j \in C_k} P_1 P_j \mid B(T) > 0\right) &\sim \frac{\sigma^2}{2} T^2 \mathbb{E}\left(\frac{\xi_k}{T} \left(1 - \frac{\xi_k}{T}\right)\right) \\ &\sim \frac{\sigma^2}{2} T^2 \mathbb{E}((1 - U_1 U_2 \dots U_k) U_1 U_2 \dots U_k) \\ &= \frac{\sigma^2}{2} T^2 \left(\frac{1}{2^k} - \frac{1}{3^k}\right). \end{aligned}$$

Consequently

$$\mathbb{E}\left(\sum_{j=2}^{B(T)} P_1 P_j \mid B(T) > 0\right) \sim \frac{\sigma^2}{4} T^2$$

which implies Lemma 8. \square

Proof of Theorem 2. (i)–(iii) are trivial. Hence it is enough to prove (iv) in the case $d = 1$, which is a straight consequence of Lemma 8. \square

4. Proofs of Theorems 3–5

Let

$$b(t) = (b_1(t), b_2(t), \dots, b_d(t)) \quad (0 \leq t \leq 1)$$

where $b_1(t), b_2(t), \dots, b_d(t)$ are independent Brownian bridges. Let $\{W(t) \in \mathbb{R}^d, t \geq 0\}$ be a Wiener process. Assume that $b(\cdot)$ and $W(\cdot)$ are independent.

For any $0 \leq s < 1$ and $T > 0$, define the process

$$\Gamma(t, s, T) = \Gamma(t) = \begin{cases} T^{1/2}b(tT^{-1}) & \text{if } 0 \leq t \leq sT, \\ W(t - sT) + T^{1/2}b(s) & \text{if } sT \leq t \leq T. \end{cases}$$

Lemma 9. *The density function $\gamma(x)$ of $\Gamma(T)$ is*

$$\gamma(x) = (2\pi(1 - s^2)T)^{-d/2} \exp\left(-\frac{\|x\|^2}{2(1 - s^2)T}\right) \quad (x \in \mathbb{R}^d).$$

Proof. Consider the case $d = 1$. Then the density function of $T^{1/2}b(s)$ is

$$(2\pi s(1 - s)T)^{-1/2} \exp\left(-\frac{x^2}{2s(1 - s)T}\right)$$

and the conditional density function of $\Gamma(T)$ given $T^{1/2}b(s) = m$ is

$$(2\pi(1 - s)T)^{-1/2} \exp\left(-\frac{(x - m)^2}{2(1 - s)T}\right).$$

Hence

$$\begin{aligned} \gamma(x) &= \psi \frac{s^{-1/2}}{\pi} \int_{-\infty}^{+\infty} \exp\left(-\psi\left(\frac{m^2}{s} + (x - m)^2\right)\right) dm \\ &= \psi \frac{s^{-1/2}}{\pi} \left(\frac{s(1 - s)T}{1 + s}\right)^{1/2} \exp\left(-\frac{x^2}{2(1 - s^2)T}\right) \int_{-\infty}^{+\infty} \exp\left(-\frac{y^2}{2}\right) dy \\ &= (2\pi(1 - s^2)T)^{-1/2} \exp\left(-\frac{x^2}{2(1 - s^2)T}\right), \end{aligned}$$

where

$$\psi = (2(1 - s)T)^{-1}.$$

This proves the lemma for the case $d = 1$. The general case follows immediately from the case $d = 1$. \square

Lemma 10. *Assume that $0 \leq s = s(T) < 1$ and*

$$\lim_{T \rightarrow \infty} (1 - s^2)T = \infty.$$

Then

$$P\{\Gamma(T) \in C\} \sim (2\pi(1 - s^2)T)^{-d/2}$$

as $T \rightarrow \infty$.

Proof. This is a trivial consequence of Lemma 9. \square

Proof of Theorem 3. Assume that $\lambda(\mathcal{C}, T) > 0$. Then there exists a $[0, T]$ -branch of the underlying branching Wiener process having a terminal point in \mathcal{C} at time T . Fix this branch. Consider the time-points $0 < \xi_1 < \xi_2 < \dots < \xi_\nu < T$ and the sets C_1, C_2, \dots, C_ν of terminal points with respect to the fixed $[0, T]$ -branch. Then, by Lemma 10,

$$P\{\text{an element of } C_k \text{ belongs to } \mathcal{C}|\xi_k\} \sim \left(2\pi \left(1 - \left(\frac{\xi_k}{T}\right)^2\right) T\right)^{-d/2}$$

and, by (2.11),

$$\begin{aligned} & E(\#\text{ of those elements of } C_k \text{ which belong to } \mathcal{C}|\xi_k) \\ & \sim \frac{\sigma^2(T - \xi_k)}{2 \left(2\pi \left(1 - \left(\frac{\xi_k}{T}\right)^2\right) T\right)^{d/2}} = \frac{\sigma^2}{2(2\pi)^{d/2}} \frac{T^{1-d/2}}{\left(1 - \frac{\xi_k}{T}\right)^{d/2-1} \left(1 + \frac{\xi_k}{T}\right)^{d/2}}. \end{aligned}$$

Hence

$$E(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) \sim \frac{\sigma^2 T^{1-d/2}}{2(2\pi)^{d/2}} E\left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T}\right)^{1-d/2} \left(1 + \frac{\xi_k}{T}\right)^{-d/2}\right). \quad (4.1)$$

Since $0 < \xi_k/T < 1$, we have

$$\begin{aligned} & \frac{\sigma^2 T^{1-d/2}}{\pi^{d/2} 2^{d+1}} E\left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T}\right)^{1-d/2}\right) \\ & \leq E(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) \leq \frac{\sigma^2 T^{1-d/2}}{2(2\pi)^{d/2}} E\left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T}\right)^{1-d/2}\right). \end{aligned} \quad (4.2)$$

In the case $d = 1$, by (4.1) and Lemma 1,

$$\begin{aligned} E(\lambda(\mathcal{C}, T) | \lambda(\mathcal{C}, T) > 0) & \sim \frac{\sigma^2 T^{1/2}}{2(2\pi)^{1/2}} E\left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T}\right)^{1/2} \left(1 + \frac{\xi_k}{T}\right)^{-1/2}\right) \\ & \sim \frac{\sigma^2 T^{1/2}}{2(2\pi)^{1/2}} \sum_{k=1}^{\nu} \int_0^1 \dots \int_0^1 \left(\frac{x_1 x_2 \dots x_k}{2 - x_1 x_2 \dots x_k}\right)^{1/2} dx_1 dx_2 \dots dx_k. \end{aligned}$$

Now we prove the following identity.

$$\sum_{k=1}^{\infty} \int_0^1 \dots \int_0^1 \left(\frac{x_1 x_2 \dots x_k}{2 - x_1 x_2 \dots x_k}\right)^{1/2} dx_1 dx_2 \dots dx_k = \frac{\pi}{2}. \quad (4.3)$$

Since

$$(1 - x)^{-1/2} = \sum_{i=0}^{\infty} a_i x^i \quad (0 < x < 1)$$

where

$$a_i = \binom{-1/2}{i},$$

we have

$$\left(\frac{x}{2-x}\right)^{1/2} = \frac{1}{\sqrt{2}} \sum_{i=0}^{\infty} \frac{a_i}{2^i} x^{i+1/2}.$$

Let U_1, U_2, \dots be i.i.d. random variables uniformly distributed on $[0, 1]$. Then

$$\begin{aligned} \int_0^1 \cdots \int_0^1 \left(\frac{x_1 \cdots x_k}{2-x_1 \cdots x_k}\right)^{1/2} dx_1 \cdots dx_k &= \mathbb{E} \left(\frac{U_1 \cdots U_k}{2-U_1 \cdots U_k} \right)^{1/2} \\ &= \frac{1}{\sqrt{2}} \sum_{i=0}^{\infty} \frac{a_i}{2^i} \mathbb{E}(U_1 \cdots U_k)^{i+1/2} = \frac{1}{\sqrt{2}} \sum_{i=0}^{\infty} \frac{a_i}{2^i} \left(\frac{1}{i+3/2}\right)^k \end{aligned}$$

and

$$\begin{aligned} \sum_{k=1}^{\infty} \mathbb{E} \left(\frac{U_1 \cdots U_k}{2-U_1 \cdots U_k} \right)^{1/2} &= \frac{1}{\sqrt{2}} \sum_{i=0}^{\infty} \frac{a_i}{2^i} \sum_{k=1}^{\infty} \left(\frac{1}{i+3/2}\right)^k = \frac{1}{\sqrt{2}} \sum_{i=0}^{\infty} \frac{a_i}{2^i} \frac{2}{2i+1} \\ &= \sqrt{2} \int_0^{\infty} e^{-x} \sum_{i=0}^{\infty} a_i \left(\frac{e^{-2x}}{2}\right)^i dx = \sqrt{2} \int_0^{\infty} e^{-x} \left(1 - \frac{e^{-2x}}{2}\right)^{-1/2} dx \\ &= \sqrt{2} \int_0^1 \left(1 - \frac{y^2}{2}\right)^{-1/2} dy = 2 \int_0^1 \frac{1}{\sqrt{2-y^2}} dy \\ &= 2 \arcsin \left(\frac{y}{\sqrt{2}}\right) \Big|_0^1 = \frac{\pi}{2}. \end{aligned}$$

Hence we have (4.3) as well as Theorem 3 for $d = 1$.

In the case $d = 2$

$$\mathbb{E} \left(\sum_{k=1}^{\nu} \left(1 - \frac{\xi_k}{T}\right)^{1-d/2} \right) = \mathbb{E} \nu \sim \log T.$$

Hence by (4.2) we have

$$\frac{\sigma^2}{8\pi} \log T \leq \mathbb{E}(\lambda(C, T) | \lambda(C, T) > 0) \leq \frac{\sigma^2}{4\pi} \log T$$

as stated in part (ii).

In the case $d = 3$ by (2.14) we have our statement (iii).

The general case can be treated similarly and, in turn, we have Theorem 3. \square

As we mentioned in the Introduction, Theorems 4 and 5 easily follow from Theorem 3.

5. Questions

Question 1. $F_T(x)$ ($x \in \mathbb{R}^d$, $T = 1, 2, \dots$) of Theorem 1 for any T is a random, empirical distribution function. Hence we have a probability measure P_T on the space of the distributions on \mathbb{R}^d . Very likely P_T , as $T \rightarrow \infty$, converges weakly to a limit measure. Prove the existence of this limit measure and characterize it. This question was proposed by O. Barndorff-Nielsen. $m(T)$ of Theorem 2 is the random expectation of $F_T(x)$ according to the law of P_T . Theorem 2 came about as I tried to answer this question.

Question 2. Investigate the limit properties of

$$\max_{x \in \mathbb{R}^d} \lambda(\mathcal{C}(x), T) \quad (d = 1, 2, \dots)$$

as $T \rightarrow \infty$ on the set $\{B(T) > 0\}$.

Question 3. Let $d = 1$ and let

$$\lambda^-(T) = \max\{x : x < 0, \lambda(\mathcal{C}(x), T) = 0\},$$

$$\lambda^+(T) = \min\{x : x > 0, \lambda(\mathcal{C}(x), T) = 0\}.$$

Investigate the limit properties of

$$\lambda^+(T) - \lambda^-(T) \quad (T \rightarrow \infty)$$

on the set $\{\lambda(\mathcal{C}, T) > 0\}$.

6. A secret

Let $\{W_{ij}(t), t \geq 0\}$ ($i, j = 1, 2, \dots$) be an array of independent \mathbb{R}^d -valued Wiener processes. Consider a system of non-independent Wiener processes.

$$w(t) = w(i(1), i(2), \dots, i(\lfloor \lg T \rfloor), T, t)$$

$$= \begin{cases} W_{11}(t) & \text{if } 0 \leq t \leq \alpha_1 T, \\ w(\alpha_1 T) + W_{2, i(2)}(t - \alpha_1 T) & \text{if } \alpha_1 T \leq t \leq \alpha_2 T, \\ \dots & \dots \\ w(\alpha_k T) + W_{k+1, i(k+1)}(t - \alpha_k T) & \text{if } \alpha_k T \leq t \leq \alpha_{k+1} T, \end{cases}$$

where

$$\alpha_k = 1 - 2^{-k}, \quad k = 1, 2, \dots, \lfloor \lg T \rfloor - 1,$$

$$i(1) = 1, \quad i(2) = 1, 2 \quad \dots \quad i(k) = 1, 2, \dots, 2^{k-1},$$

$$0 \leq t \leq (1 - 2^{-\lfloor \lg T \rfloor})T, \quad T \geq 2.$$

It is easy to see that the process $w(\cdot)$ is very similar to a critical branching Wiener process. Hence my secret, which I share with you, is the following: I try to solve any question regarding a critical branching Wiener process by replacing it by the above model. If I succeed in doing so, I try to prove the same for the branching model by Lemma 1 which essentially claims that the two models are close to each other. In fact I followed this method in proving the above theorems. I have to confess that I could not answer the three questions in Section 5 even in this simple situation.

Acknowledgement

The author is indebted to the referee who proved (4.3).

References

- Athreya, K.B. and Ney, P.E. (1972) *Branching Processes*. New York: Springer-Verlag.
Fleischman, J. (1978) Limiting distributions for branching random fields. *Trans. Amer. Math. Soc.* **239**, 353–389.
Révész, P. (1994) *Random Walks of Infinitely Many Particles*. Singapore: World Scientific.

Received October 1994 and revised November 1995