

A CONVERSE TO THE DOMINATED CONVERGENCE THEOREM

BY

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1. Introduction and summary

On a probability space $(\Omega, \mathfrak{B}, P)$, let $\{f_n, n = 1, 2, \dots\}$ be a sequence of nonnegative random variables in L_1 such that $f_n \rightarrow f \in L_1$ with probability 1, and define $g = \sup_n f_n$. If $g \in L_1$, the Lebesgue dominated convergence theorem asserts that $E(f_n) \rightarrow E(f)$. More generally, as noted by Doob [1, p. 23], if $g \in L_1$, then for any Borel field \mathfrak{B}_0 contained in \mathfrak{B} ,

$$(1) \quad E(f_n | \mathfrak{B}_0) \rightarrow E(f | \mathfrak{B}_0) \quad \text{a.e.}$$

If one extends this result in a minor manner, Lebesgue's condition $g \in L_1$ is not only sufficient but necessary, as the following converse to the dominated convergence theorem asserts.

THEOREM 1. *If $f_n \geq 0, f_n \rightarrow f$ a.e., $f_n \in L_1, f \in L_1$, and $g = \sup_n f_n \notin L_1$, there are, on a suitable probability space, random variables $\{f_n^*, n = 1, 2, \dots\}, f^*$, and a Borel field \mathfrak{C} such that f^*, f_1^*, f_2^*, \dots have the same joint distribution as f, f_1, f_2, \dots , and*

$$(2) \quad P\{E(f_n^* | \mathfrak{C}) \rightarrow E(f^* | \mathfrak{C})\} = 0.$$

In view of this result, it is of interest to find conditions which will ensure that $g \in L_1$. As a special case of interest, let h be a nonnegative random variable in L_1 , let \mathfrak{B}_n be a monotone sequence of Borel fields contained in \mathfrak{B} , and let $f_n = E(h | \mathfrak{B}_n)$. Doob [1, p. 317] has shown that if $h \log h \in L_1$, then also $g = \sup_n f_n \in L_1$. It turns out that the condition $h \log h \in L_1$ is necessary, as well as sufficient, in the following sense:

THEOREM 2. *If $h \geq 0, h \in L_1, h \log h \notin L_1$, there are, on a suitable probability space, a random variable h^* with the same distribution as h and a monotone sequence \mathfrak{B}_n^* of Borel fields, which can be chosen either increasing or decreasing, for which*

$$(3) \quad g^* = \sup_n E(h^* | \mathfrak{B}_n^*) \notin L_1.$$

Theorem 2 will be an immediate consequence of the following result, which gives sharp upper bounds on the distribution of g^* , rather than only information about the expectation of g^* as in Theorem 2.

THEOREM 3. *Let h^* be any nonnegative random variable in L_1 , and let h be the (essentially unique) nonincreasing function on the unit interval $(0, 1]$ whose*

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distribution, with respect to Lebesgue measure m on $(0, 1]$, is the same as that of h^* . Define g on $(0, 1]$ by

$$(4) \quad g(x) = \frac{1}{x} \int_0^x h(t) dt.$$

Then

(a) for any monotone sequence \mathfrak{B}_n of Borel fields contained in \mathfrak{B} , and any $\lambda > 0$,

$$(5) \quad P\{g^* > \lambda\} \leq m\{g > \lambda\},$$

where $g^* = \sup_n E(h^* | \mathfrak{B}_n)$,

(b) for every $\varepsilon > 0$ there is an increasing sequence \mathfrak{C}_n of Borel fields in the unit interval X for which

$$(6) \quad P\{g^{**} \geq k\varepsilon\} = m\{g \geq k\varepsilon\} \quad \text{for } k = 0, 1, 2, \dots, \text{ and } g^{**} \geq g - \varepsilon,$$

where g^{**} denotes $\sup_n E(h | \mathfrak{C}_n)$, and

(c) for every $\varepsilon > 0$ and every decreasing sequence of real numbers $\{Q_n, n = 1, 2, \dots\}$ with $0 \leq Q_n \leq 1$ and $Q_n \rightarrow 0$ as $n \rightarrow \infty$, there are, on a suitable probability space, a random variable f with the same distribution as h and a decreasing sequence \mathfrak{D}_n of Borel fields such that for every positive integer k ,

$$(7) \quad P\{g_1 \geq k\varepsilon\} \geq Q_k m\{g \geq k\varepsilon\},$$

where $g_1 = \sup_n E(f | \mathfrak{D}_n)$.

The proof that Theorem 3 implies Theorem 2 will use the following result of Hardy and Littlewood [2, p. 99]: For any nonnegative monotone decreasing function h on $(0, 1]$, either $h \log h, h(t) \log g(t)$, and g are all in L_1 , or none is.

Say that a distribution μ on the real line dominates a distribution ν if $\mu(x, \infty) \geq \nu(x, \infty)$ for all x . Theorem 3(a) asserts that for any nonnegative h^* in L_1 , the distribution of g , denote it by μ , dominates that of $\sup_n E[h^* | \mathfrak{B}_n]$ for any monotone increasing or decreasing sequence of Borel fields \mathfrak{B}_n . Part (b) asserts that μ is, in a very strong sense, best possible for increasing \mathfrak{B}_n . Part (c) asserts that the same distribution μ is best possible for decreasing \mathfrak{B}_n , though in a somewhat weaker sense.

Inequality (5) has the following consequence. Consider a fair gambling system, which terminates after N plays, and in which the bettor is not allowed credit, i.e., a sequence $X_0, X_1, X_2, \dots, X_N$ of nonnegative random variables which form a martingale; X_k is the bettor's fortune after k plays, and for simplicity let X_0 be constant. Suppose the bettor is allowed to choose, in advance of play, either of the following options:

Option 1. He uses the system and, at the end is paid, not his final fortune X_N , but the largest fortune $Y = \max(X_0, \dots, X_N)$ he ever had in the course of play.

Option 2. He uses the system, achieving a terminal fortune X_N . If

X_N is as high as possible, he is given X_N . If not, he is given his original fortune X_0 and tries the system repeatedly until a final fortune Z is obtained which (strictly) exceeds the final fortune X_N on his first attempt. He is then given Z .

Though the distribution of Z need not dominate that of Y , it turns out that Option 2 is always better, in the sense that $E(Z) \geq E(Y)$.

One final easy observation. For any nonnegative martingale X_1, \dots, X_N , $E(\max(X_1, \dots, X_N)) \leq NE(X_1)$. This bound is best possible in that for every nonnegative X_1 with finite expectation and every $\varepsilon > 0$ and $N \geq 1$, there is a nonnegative martingale, X_1^*, \dots, X_N^* , where X_1^* has the same distribution as X_1 , and for which $E(\max(X_1^*, \dots, X_N^*)) > NE(X_1) - \varepsilon$.

2. Proof of Theorem 1

The Borel field \mathcal{C} will be the smallest field with respect to which some random variable Z is measurable. We first reduce the theorem to the special case in which each f_n has only two values, 0 and $v_n > 0$, and at every sample point exactly one f_n is positive. Thus, if $p_n = P\{f_n = v_n\}$, we have $0 < p_n < 1$, $\sum p_n = 1, f \equiv 0, E(g) = \sum p_n v_n = \infty$.

To achieve this reduction, write

$$F_n = \max(f_n - f, 0), \quad G_n = \min(f_n - f, 0).$$

Then $F_n \geq 0, F_n \in L_1$,

$$\begin{aligned} \sup_n F_n &\geq g - f \notin L_1, & F_n &\rightarrow 0 \text{ a.e.}, \\ \sup_n |G_n| &\leq f \in L_1, & G_n &\rightarrow 0 \text{ a.e.} \end{aligned}$$

For any Borel field \mathcal{C} , it follows from (1) that $E(G_n | \mathcal{C}) \rightarrow 0$ a.e., so that

$$\begin{aligned} P\{E(f_n | \mathcal{C}) \rightarrow E(f | \mathcal{C})\} &= P\{E(f_n - f | \mathcal{C}) \rightarrow 0\} \\ &= P\{E(F_n + G_n | \mathcal{C}) \rightarrow 0\} = P\{E(F_n | \mathcal{C}) \rightarrow 0\}. \end{aligned}$$

Thus if we find, enlarging the probability space if necessary, a Borel field \mathcal{C} for which $P\{E(F_n | \mathcal{C}) \rightarrow 0\} = 0$, it will follow that

$$P\{E(f_n | \mathcal{C}) \rightarrow E(f | \mathcal{C})\} = 0.$$

Thus we have reduced the theorem to the special case of the F_n , i.e., to the case $f = 0$.

Suppose now that $f = 0$. Denote by A_k the event

$$\{f_k \geq g - 1, f_k < g - 1 \text{ for } i < k\}.$$

The A_k are disjoint, and $\sum P(A_k) = 1$. Choose a simple function (i.e., one with only finitely many values) s_k such that s_k vanishes off $A_k, 0 \leq s_k \leq f_k$ on A_k , and

$$E(s_k) \geq \int_{A_k} f_k dP - \frac{1}{2^k}.$$

Then $\sup_k s_k = \sum_k s_k$, so that

$$\begin{aligned} E \sup_k s_k &= \sum E(s_k) \geq \sum_k \int_{A_k} f_k dP - 1 \\ &\geq \sum_k \int_{A_k} g dP - 2 = E(g) - 2 = \infty. \end{aligned}$$

Since $s_k \leq f_k$, for any \mathfrak{C} ,

$$P\{E(s_k | \mathfrak{C}) \rightarrow 0\} = 0 \quad \text{implies} \quad P\{E(f_k | \mathfrak{C}) \rightarrow 0\} = 0,$$

so that we have reduced the theorem to the case of the s_k , i.e., the case in which $f = 0$, each f_n is simple, and at each sample point at most one f_n is positive. Starting from this case we represent each f_n as the sum of a finite number of nonnegative functions, each having only two values, one of which is 0, and no two of which are simultaneously positive. Rearranging these functions into a single sequence, omitting those which are 0 with probability 1 and, if the set B on which all these functions vanish has positive probability, taking the indicator I_B as an additional function, yield a sequence f_1, f_2, \dots with the properties stated at the beginning of the section, and the reduction is complete. We now prove the theorem in the special case.

Let k be the positive integer such that $1 < 2^k P_1 \leq 2$, and let S_n , $n = 1, 2, \dots$, denote the set of integers $i \geq 2$ for which $2^{n+k} \leq v_i < 2^{n+k+1}$. Define:

$$\begin{aligned} r_n &= \sum_{i \in S_n} p_i, & t_n &= r_n + 2^{-(n+k)}; \\ r &= \sum_n r_n = \sum_{i \in S} p_i, \end{aligned}$$

where $S = \cup S_n = \{i : i \geq 2 \text{ and } v_i \geq 2^{k+1}\}$;

$$t = \sum t_n = r + 2^{-k}.$$

Let W, Z_0, Z_1, \dots be independent integer-valued random variables with distributions as follows:

$$P(W = 0) = 1 - t, \quad P(W = n) = t_n \quad \text{for } n > 0.$$

$$P(Z_0 = 1) = (p_1 - 2^{-k}) / (1 - t),$$

$$P(Z_0 = i) = p_i / (1 - t) \quad \text{for } i \geq 2, i \notin S,$$

$$P(Z_0 = i) = 0 \quad \text{otherwise.}$$

For $n \geq 1$,

$$P(Z_n = 1) = 2^{-(k+n)} / t_n,$$

$$P(Z_n = i) = p_i / t_n \quad \text{for } i \in S_n,$$

$$P(Z_n = i) = 0 \quad \text{otherwise.}$$

Define $X = Z_W$, and verify that $P(X = n) = p_n$ thus:

$$P(X = 1) = \sum_{n=0}^{\infty} P(W = n, Z_n = 1) = p_1 - 2^{-k} + \sum_{n=1}^{\infty} 2^{-(k+n)} = p_1;$$

for $i > 1, i \notin S,$

$$P(X = i) = P(W = 0, Z_0 = i) = p_i;$$

for $i \in S_n,$

$$P(X = i) = P(W = n, Z_n = i) = p_i.$$

Thus, if we define $\phi_n = v_n$ on $\{X = n\}, \phi_n = 0$ otherwise, $\{\phi_n\}$ has the same joint distribution as $\{f_n\},$ i.e., ϕ_n has only the two values $0, v_n,$

$$P\{\phi_n = v_n\} = p_n,$$

and at each sample point exactly one ϕ_n is positive.

For any Borel field $\mathcal{C}, E(\phi_n | \mathcal{C}) = v_n P\{X = n | \mathcal{C}\}.$ It suffices to find a \mathcal{C} for which the event $\{v_n P\{X = n | \mathcal{C}\} \geq 1$ infinitely often $\}$ has probability one. We show that the Borel field \mathcal{C} determined by Z_0, Z_1, \dots has the property. For $i \in S_n,$

$$P\{X = i | \mathcal{C}\} = 0 \quad \text{if } Z_n \neq i,$$

$$P\{X = i | \mathcal{C}\} = t_n \quad \text{if } Z_n = i.$$

Thus, if $Z_n = i, v_i P\{X = i | \mathcal{C}\} = v_i t_n.$ Since $t_n \geq 2^{-(n+k)}$ and, for $i \in S_n, v_i \geq 2^{n+k},$ we have $v_i t_n \geq 1.$ Thus, for $n \geq 1$ whenever $A_n = \{Z_n \neq 1\}$ occurs, so does $B_n = \{v_i P\{X = i | \mathcal{C}\} \geq 1$ for some $i \in S_n\}.$ The A_n are independent, with $P(A_n) = r_n/t_n.$ We show that

$$\sum_n (r_n/t_n) = \infty.$$

If $r_n < 2^{-(n+k)}, t_n < 2^{-(n+k-1)},$ so that

$$(r_n/t_n) \geq 2^{n+k-1} r_n = \sum_{i \in S_n} 2^{n+k+1} p_i/4 \geq \sum_{i \in S_n} p_i v_i/4.$$

If $r_n \geq 2^{-(n+k)}$ for infinitely many $n,$ then $(r_n/t_n) \geq \frac{1}{2}$ for infinitely many $n,$ and the series $\sum (r_n/t_n)$ diverges. If $r_n < 2^{-(n+k)}$ for sufficiently large $n,$ say for $n \geq n_0,$

$$\sum_n (r_n/t_n) \geq \sum_{n \geq n_0} \sum_{i \in S_n} p_i v_i/4 = \sum_{i \in T} p_i v_i/4$$

where $T = \{i \geq 2, v_i \geq 2^{n_0+k}\}.$ Since

$$\sum_i p_i v_i = \infty \quad \text{and} \quad \sum_{i \in T} p_i v_i \leq 2^{n_0+k} \sum p_i \leq 2^{n_0+k},$$

we conclude that $\sum_{i \in T} p_i v_i$ diverges. Thus $\sum (r_n/t_n)$ diverges, so that, with probability 1, infinitely many $A_n,$ and hence infinitely many $B_n,$ occur. This completes the proof.

3. Proofs of other results

For part (a) of Theorem 3 use an inequality of Doob [1, p. 314] which asserts that, for every $\lambda > 0,$

$$(8) \quad \lambda P\{g_n^* \geq \lambda\} \leq \int_{\{g_n^* \geq \lambda\}} h^* dP,$$

where $g_n^* = \max_{1 \leq i \leq n} E(h^* | \mathfrak{B}_i)$. Letting $\lambda \downarrow \lambda_0 > 0$ yields

$$\lambda_0 P\{g_n^* > \lambda_0\} \leq \int_{\{g_n^* > \lambda_0\}} h^* dP.$$

Letting $n \rightarrow \infty$, and dropping the subscript in λ_0 , you obtain, for every $\lambda > 0$,

$$(9) \quad \lambda P\{g^* > \lambda\} \leq \int_{\{g^* > \lambda\}} h^* dP,$$

and letting $\lambda \uparrow \lambda_0$ yields an inequality like (9) with the event $\{g^* > \lambda\}$ replaced by $\{g^* \geq \lambda_0\}$. For any λ for which $P\{h^* > \lambda\} = 0$, we have also $P\{g^* > \lambda\} = 0$, and (5) is trivial. If $P\{h^* \geq \lambda\} > 0$, note that g is monotone, and let u be the largest number for which $g(u) \geq \lambda$. Then for any event A for which

$$\frac{1}{P(A)} \int_A h^* dP \geq \lambda,$$

we must have $P(A) \leq u$. The event $A = \{g^* \geq \lambda\}$ has the property, from the remark following (9), so that

$$(10) \quad P\{g^* \geq \lambda\} \leq u = m\{g \geq \lambda\}.$$

Letting $\lambda \downarrow \lambda_0$ yields (5), and (a) is established.

The remark on gambling systems is a consequence of $E(g^*) \leq E(g)$, which follows from (5). For, with $h^* = X_n$, and \mathfrak{B}_i the Borel field determined by $X_0, \dots, X_i, g^* = Y$, and

$$E(Z) = \int_0^1 \alpha(u) du,$$

where $\alpha(u) = g$ (smallest v with $h(v) = h(u)$). Since $\alpha(u) \geq g(u)$, $E(Z) \geq E(g)$, and the proof is complete.

For part (b), let $C_n = \{(n - 1)\varepsilon \leq g < n\varepsilon\}$, and let \mathfrak{C}_n be the Borel field determined by C_1, \dots, C_{n-1} . If C_n is nonempty, it is an interval $a < u \leq b$. When C_n occurs, $E(h | \mathfrak{C}_n) = E(h | u \leq b) = g(b) \geq (n - 1)\varepsilon$. Thus, on C_n , $E(h | \mathfrak{C}_n) \geq g - \varepsilon$, and $g^{**} = \sup_n E(h | \mathfrak{C}_n) \geq g - \varepsilon$ everywhere.

For part (c), set $Q_0 = 1$, and define $p_n = Q_{n-1} - Q_n$, so that $p_n \geq 0$, $\sum_1^\infty p_n = 1$. Let α be a random variable, independent of g, h (this may require extending the probability space) with $P\{\alpha = n\} = p_n$. The Borel field \mathfrak{D}_n will specify the value of α and, when $\alpha = k$, will specify, for every $i, 1 \leq i \leq k - n$, whether C_i , defined in the proof of part (b), occurs. More formally, if I_i is the indicator or characteristic function of C_i (I_i has 1 as its value on C_i and 0 off C_i), and J_k is the indicator of $\{\alpha = k\}$, \mathfrak{D}_n is the Borel field determined by the functions $J_k, k = 1, 2, \dots$, and those functions $J_k I_i$, for which $k > n$ and $1 \leq i \leq k - n$. Then, for any i, k, n with $i > k - n > 0$ we have, on $C_i \cap \{\alpha = k\}$,

$$E(h | \mathfrak{D}_n) = E(h | \alpha, \cup_{j>k-n} C_j) \geq (k - n)\varepsilon.$$

For $k \leq i$, we choose $n = 1$; for $k > i$, we choose $n = k - i + 1$. We then see that on $C_i \cap \{\alpha = k\}$ either $g_1 \geq (k - 1)\varepsilon$ or $g_1 \geq (i - 1)\varepsilon$ according as $k \leq i$ or $k > i$. We conclude that, on $(\cup_{i>j} C_i) \cap \{\alpha > j\}$, $g_1 \geq j\varepsilon$. Since $\cup_{i>j} C_i = \{g \geq j\varepsilon\}$, we obtain

$$P\{g_1 \geq j\varepsilon\} \geq P\{g \geq j\varepsilon\} P\{\alpha > j\} = Q_j P\{g \geq j\varepsilon\},$$

which is assertion (7).

The proof of Theorem 2 is now easy. We may suppose that h is a non-increasing function on the unit interval, and that probability is Lebesgue measure. Since $h \log h$ is not in L_1 , the result of Hardy and Littlewood referred to following (7) implies that g , defined as in Theorem 3, is not in L_1 . To choose \mathfrak{G}_n^* increasing, with $g^* \notin L_1$, choose as \mathfrak{G}_n^* the \mathfrak{C}_n of part (b) of Theorem 3. Then $g^* \geq g - \varepsilon$, so that $g^* \notin L_1$. To choose \mathfrak{G}_n^* decreasing, note first that, since $g \notin L_1$, $\sum_k m(g \geq k\varepsilon) = \infty$. We may then choose a monotone sequence Q_k converging to 0 with $1 \geq Q_k \geq 0$, $k = 1, 2, \dots$, for which

$$(11) \quad \sum_k Q_k m(g \geq k\varepsilon) = \infty.$$

For this choice of Q_k , and \mathfrak{G}_n^* chosen as the \mathfrak{D}_n of (c) of Theorem 3, the g^* of Theorem 2 is the g_1 of Theorem 3. From (7) and (11), clearly,

$$\sum P\{g^* \geq k\varepsilon\} = \infty,$$

so that $g^* \notin L_1$. This completes the proof.

As for the final remarks about nonnegative martingales, let

$$Y = \max(X_1, \dots, X_N),$$

and note that $Y \leq \sum X_i$, so $E(Y) \leq NE(X_1)$. To find a process where $E(Y) > NE(X_1) - \varepsilon$, let X_1^*, \dots, X_N^* be the successive fortunes of a gambler who at time j gambles as follows. He stakes his entire fortune X_j^* on a long shot, so that with small probability, namely t^{-1} , his fortune increases to tX_j^* , and with high probability, namely $1 - t^{-1}$, his fortune decreases to 0. It is easy to verify that $E(Y^*) = nE(X_1) - (n - 1)t^{-1}E(X_1)$. This completes the proofs.

Added in proof. Theorem 2, with a particularly interesting choice of \mathfrak{G}_n^* , has also been obtained by D. L. BURKHOLDER, in *Successive conditional expectations of an integrable function*, Ann. Math. Statistics, vol. 33 (1962), pp. 887-893.

BIBLIOGRAPHY

1. J. L. DOOB, *Stochastic processes*, New York, Wiley, 1953.
2. G. H. HARDY AND J. E. LITTLEWOOD, *A maximal theorem with function-theoretic applications*, Acta Math., vol. 54 (1930), pp. 81-116.

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