formly in  $\lambda$ . However, this approach may not succeed in controlling the asymptotic *level* of the confidence set. The problem is that the relevant asymptotic expansions may not converge uniformly over all  $\lambda$ , especially when  $\lambda$  is infinite dimensional.

An interesting strategy, proposed by Loh (1985) in a testing context, is to pick critical values so as to control the apparent level of the confidence set for  $\theta$  over a confidence set for  $\lambda$  of level  $1 - \varepsilon_n$ , where  $\varepsilon_n$  is small. When feasible, this construction ensures that the level of the confidence set for  $\theta$  is at least  $1 - \alpha - \varepsilon_n$ . One difficulty is finding a good confidence set for  $\lambda$ . If the latter is too large, then the induced confidence set for  $\theta$  is likely to be inefficient. Perhaps the notion of controlling level of a confidence set for  $\theta$  is too strong. On the other hand, controlling asymptotic coverage probability only pointwise in  $\lambda$  is clearly too weak.

## REFERENCES

Beran, R. J. (1987). Prepivoting to reduce level error of confidence sets. *Biometrika* 74 457-468. Beran, R. J. (1988). Prepivoting test statistics: A bootstrap view of asymptotic refinements. *J. Amer. Statist. Assoc.* To appear.

BICKEL, P. J. (1987). Discussion of Efron. J. Amer. Statist. Assoc. 82 191.

Ducharme, G. R., Jhun, M., Romano, J. P. and Truong, K. N. (1985). Bootstrap confidence cones for directional data. *Biometrika* 72 637-645.

LOH, W.-Y. (1985). A new method for testing separate families of hypotheses. J. Amer. Statist. Assoc. 80 362-368.

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I found Professor Hall's unified treatment of bootstrap bounds and confidence intervals very valuable. I was particularly interested by his exploration of the relation between the accelerated bias correction bootstrap bounds and the Studentized bounds, a relationship which I also studied, but only in the parametric framework, in my discussion of Efron (1987). In my discussion I want to:

- 1. Argue at least heuristically that, in the nonparametric context, the second order equivalence of  $\hat{\theta}_{ABC}$  and  $\hat{\theta}_{STUD}$  holds quite generally for  $\theta(F)$  a sufficiently smooth von Mises functional, provided that we Studentize properly. For example, it holds if the estimate  $\hat{\theta} = \theta(\hat{F})$ , where  $\hat{F}$  is the empirical d.f., is an M estimate corresponding to a nice  $\psi$  function; see Huber [(1981), Chapter 2] for examples.
- 2. Suggest that quite generally in a parametric, nonparametric or semiparametric context,  $\hat{\theta}_{ABC}$  and  $\hat{\theta}_{STUD}$  are second order equivalent provided again that  $\hat{\theta}$  is efficient and we Studentize properly, that is, by an efficient estimate of the asymptotic standard deviation of  $\hat{\theta}$ .

Since, modulo regularity conditions, checking equivalence of the theoretical points  $\hat{\theta}_{abc}$  and  $\hat{\theta}_{Stud}$  and of the corresponding bootstrap points  $\hat{\theta}_{AB}$  and  $\hat{\theta}_{STUD}$  is the same thing, we restrict ourselves to the former. Suppose H and K are as in Hall the d.f.'s of  $T \equiv \sqrt{n} (\hat{\theta} - \theta(F))/(\sigma(F))$  and  $\tilde{T} \equiv \sqrt{n} (\hat{\theta} - \theta(F))/\hat{\sigma}$  and admit Edgeworth expansions,

$$H(x) = \Phi(x) - \phi(x) \left( K_{1n} + \frac{K_{3n}}{6} (x^2 - 1) \right) + O(n^{-1}),$$
  
 $K(x) = \Phi(x) - \phi(x) \left( \tilde{K}_{1n} + \frac{\tilde{K}_{3n}}{6} (x^2 - 1) \right) + O(n^{-1}),$ 

where  $K_{in}$ ,  $\tilde{K}_{in} = O(n^{-1/2})$ , i = 1, 3, are asymptotic cumulants of T,  $\tilde{T}$ ; see Bickel (1974) for a precise formulation. In this notation, Hall shows that  $\hat{\theta}_{abc} = \hat{\theta}_{Stud} + O(n^{-3/2})$  provided that

(1) 
$$m = \Phi^{-1}H(0) + O(n^{-1}) = K_{1n} - \frac{K_{3n}}{6} = \tilde{K}_{1n} - \frac{\tilde{K}_{3n}}{6} + O(n^{-1})$$

and

(2) 
$$a = -\frac{1}{6}(K_{3n} + \tilde{K}_{3n}) + O(n^{-1}).$$

Note that without any further calculation, the corresponding quantities for  $\hat{\theta}_{ABC}$ , not only  $\hat{m}$  but also  $\hat{a}$  can be estimated (but not in an invariant way) directly. In particular, we can take  $\hat{a} = -\frac{1}{6}(\hat{K}_{3n} + \tilde{K}_{3n})$ , where the  $K_{3n}$ ,  $\tilde{K}_{3n}$  are the bootstrap cumulants of appropriately truncated versions of T and  $\tilde{T}$ . For the fully nonparametric model and special  $\theta(F)$  and exponential families Hall, following Efron (1987), derives a further approximation for a, if  $\hat{\theta}$  is an efficient estimate of  $\theta(F)$  and  $\hat{\sigma}$  is an efficient estimate of  $\sigma(F)$ , the asymptotic s.d. of  $\sqrt{n} \hat{\theta}$ .

To extend Hall's approach in the fully nonparametric model, suppose  $\hat{\theta}$  admits a von Mises (stochastic) expansion

$$\hat{\theta} = \theta(F) + \int \psi(x, F) d\hat{F}(x) + \frac{1}{2} \int \gamma(x, y, F) d\hat{F}(x) d\hat{F}(y) + o_P(n^{-1}),$$

where

$$\int \Phi(x,F) \, dF(x) = 0,$$

γ is symmetric and

$$\int \gamma(x, y, F) dF(y) = 0 \text{ for all } x.$$

Compare Withers (1983). The efficient estimate of  $\sigma^2(F) = \int \psi^2(x, F) dF(x)$  is  $\sigma^2(\hat{F}) = \int \psi^2(x, \hat{F}) d\hat{F}(x)$ . Under additional regularity conditions we can check  $\sigma^2(\hat{F}) = \sigma^2(F) + \int v(x, F) d\hat{F} + O_P(n^{-1})$ , where

$$v(X, F) = 2 \int \psi(y, F) \gamma(x, y, F) dF(y) + \psi^{2}(x, F) - \sigma^{2}(F).$$

Since

$$ilde{T} = T iggl(1 - rac{\sigma^{-2}}{2}(F)igl(\hat{\sigma}^2(F) - \sigma^2(F)igr)iggr) + O_P(n^{-1})$$

if we take  $K_{3n}$ ,  $\tilde{K}_{3n}$  to be the cumulants of the appropriate linear or quadratic approximations to T,  $\tilde{T}$ , we get

(3) 
$$\tilde{K}_{3n} - K_{3n} = -3n^{-1/2} \Big[ \text{SKEW}(\psi) + \sigma^{-3}(F) \\ \times \int \int \psi(x, F) \psi(y, F) \gamma(x, y, F) \, dF(x) \, dF(y) \Big] \\ + O(n^{-1})$$

and from Withers (1983),

(4) 
$$K_{3n} = n^{-1/2} \left[ \text{SKEW}(\psi) + \frac{3}{2} \sigma^{-3}(F) \times \int \psi(x) \psi(y) \gamma(x, y) \, dF(x) \, dF(y) \right] + O(n^{-1}).$$

From (3) and (4) we get

(5) 
$$K_{3n} + \tilde{K}_{3n} = -SKEW(\psi)$$

and hence

(6) 
$$a = \frac{\text{SKEW}}{6}(\psi).$$

These computations and Efron [(1987), Section 6] suggest that if **P** is any model (parametric, fully nonparametric or semiparametric), the map  $F \to \theta(F)$  is sufficiently smooth,  $\hat{\theta}$  is an efficient (on **P**) estimate of  $\theta(F)$  and  $\hat{\sigma}$  is an efficient (on **P**) estimate of the asymptotic variance of  $n^{1/2}\hat{\theta}$ , then (6) should hold. Of course, to make any such assertions precise we need to assume existence of stochastic (von Mises) and Edgeworth expansions for T and  $\tilde{T}$ . For appropriate conditions and the necessary formulae, see Pfanzagl [(1985), Chapter 10], where an extensive discussion of questions such as those treated in Section 4 may also be found.

## REFERENCES

BICKEL, P. J. (1974). Edgeworth expansions in nonparametric statistics. Ann. Statist. 2 1-20.

EFRON, B. (1987). Better bootstrap confidence intervals (with discussion). J. Amer. Statist. Assoc. 82 171-200.

HUBER, P. J. (1981). Robust Statistics. Wiley, New York.

PFANZAGL, J. (1985). Asymptotic Expansions for General Statistical Models. Lecture Notes in Statist. 31. Springer, Berlin.

WITHERS, C. S. (1983). Expansions for the distribution and quantiles of a regular functional of the empirical distribution with application to nonparametric confidence intervals. *Ann. Statist.* 11 577-587.

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