

CONVERGENCE OF THE REDUCED EMPIRICAL PROCESS FOR NON-I.I.D. RANDOM VECTORS

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Any triangular array of row independent random vectors with continuous df's has a standard reduction to random vectors with values in the unit cube. The reduced empirical process belonging to the transformed random vectors is always relatively compact. Weak convergence to a (necessarily Gaussian) process holds iff the corresponding covariance kernel converges pointwise.

1. Notation and results. For every $n \geq 1$ let $X_n^j = (X_{n1}^j, \dots, X_{nk}^j)$, $j = 1, \dots, n$, be independent random k -vectors with continuous df's F_n^j . The df $F_n = (1/n) \sum_{j=1}^n F_n^j$ has marginal df's F_{ni} , $i = 1, \dots, k$ (say) which define random vectors

$$(1.1) \quad U_n^j = (U_{n1}^j, \dots, U_{nk}^j) = (F_{n1}^{-1}(X_{n1}^j), \dots, F_{nk}^{-1}(X_{nk}^j)), \quad j = 1, \dots, n,$$

with df's

$$(1.2) \quad G_n^j = F_n^j \circ (F_{n1}^{-1}, \dots, F_{nk}^{-1}), \quad j = 1, \dots, n,$$

where F_{ni}^{-1} is the left continuous inverse of F_{ni} , $i = 1, \dots, k$.

The df $G_n = (1/n) \sum_{j=1}^n G_n^j$ has uniform marginals G_{ni} and consequently

$$(1.3) \quad |G_n(t) - G_n(s)| \leq k|t - s| \quad \forall t, s \in E_k = [0, 1]^k \quad \forall n \geq 1,$$

where $|t| = \max\{t_i : i = 1, \dots, k\}$, $t = (t_1, \dots, t_k) \in R^k$.

Let G_n^e be the resulting empirical df of the U_n^j , $j = 1, \dots, n$. Then we get the reduced empirical process of X_n^j , $j = 1, \dots, n$,

$$(1.4) \quad X_n(t) = n^{1/2}(G_n^e(t) - G_n(t)), \quad t \in E_k,$$

with corresponding covariance kernel

$$(1.5) \quad EX_n(t) \cdot X_n(s) = K_n(t, s) = G_n(s \wedge t) - \frac{1}{n} \sum_{j=1}^n G_n^j(t) \cdot G_n^j(s),$$

$t, s \in E_k,$

where $s \wedge t = (s_1 \wedge t_1, \dots, s_k \wedge t_k)$.

Since $X_n(\cdot)$ takes values in D_k , see e.g. Neuhaus (1971), we can formulate a

THEOREM. (i) *For any triangular array of row independent random vectors X_n^j , $j = 1, \dots, n$, $n \geq 1$, having continuous df's*

$$(1.6) \quad \text{the sequence } (X_n(\cdot))_{n \geq 1} \text{ is relatively compact on } D_k \text{ (furnished with the Skorohod-topology).}$$

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(ii) *If for some K*

$$(1.7) \quad K_n(s, t) \rightarrow K(s, t) \quad \forall s, t \in E_k,$$

then there exists a Gaussian process X with sample paths in C_k and covariance kernel K for which

$$(1.8) \quad X_n \Rightarrow X \text{ on } D_k \text{ (weak convergence).}$$

On the other hand, if the finite dimensional distributions of X_n converge to those of some process X , then X is Gaussian and (1.7) holds with K the covariance kernel of X .

2. Proof. We shall point out that the proof of Neuhaus (1971), pages 1293–1294, where the i.i.d. case is treated extends almost literally to the non-i.i.d. case considered here. The proof *loc. cit.* is a k -dimensional extension of the “classical” 1-dimensional proof as presented in Parthasarathy (1967). Since this proof is quick and almost self-contained it seems an appropriate one in handling the reduced empirical process.

While proving the 1-dimensional version of the above theorem Shorack (1973) introduced the linearly interpolated process S_n of X_n having $S_n(i/n) = X_n(i/n)$ for $i = 0, \dots, n$ and used only the supremum metric; but we feel it is much simpler to introduce a Skorohod type metric and work directly with X_n . The crucial point of his proof is the not quite immediate inequality

$$(1.9) \quad E|S_n(t) - S_n(s)|^4 \leq 144|t - s|^2 \quad \forall s, t \in [0, 1].$$

For X_n itself such an inequality does not hold in general; nevertheless, it holds almost:

For every $k \geq 1, r \geq 1$ there exists a constant $Q = Q(k, r)$ independent of $n \geq 1$ with

$$(1.10) \quad E(X_n(t) - X_n(s))^{2r} \leq Q \cdot |t - s|^r \quad t, s \in E_k, |t - s| \geq \frac{1}{n}.$$

For the 1-dimensional case (1.10) (with $r = 2$) is elementary and was used by Shorack (1973), too. But even in the general case $k \geq 1$, (1.10) (with $r = k + 1$) and (1.3) are all that one really needs, since the proof of the statement

$$(1.11) \quad \lim_{\delta \rightarrow 0} \limsup_{n \rightarrow \infty} P(\{\sup \{|X_n(t) - X_n(s)| : |t - s| \leq \delta\} \geq \epsilon\}) = 0, \quad \forall \epsilon > 0,$$

as given in Neuhaus (1971), pages 1293–1294, works without any change if (4.1) resp. Lemma 5.2 *loc. cit.* is replaced by (1.3) resp. (1.10) above. The rest of the Theorem then follows immediately; see e.g. Shorack (1973).

Therefore, all that remains to do is verify (1.10) for general $k \geq 1, r \geq 1$. By the way, let us mention that in Sen (1970) (1.10) was claimed for $k = 1$ and every $r \geq 1$ without the restriction $|t - s| \geq (1/n)$; but for example in the 1-dimensional i.i.d. case (1.10) (with $r = 2$) becomes incorrect for $|t - s| \rightarrow 0$.

PROOF OF (1.10). For every family X_1, \dots, X_n of independent integrable rv's the r th moment

$$E(\sum_{i=1}^n X_i)^r = \sum_{i_1, \dots, i_r=1}^n E(X_{i_1} \cdots X_{i_r})$$

can be rewritten as a finite sum of terms like

$$(1.12) \quad \pm \prod_{\mu=1}^r \sum_{j=1}^n \prod_{\nu=1}^r (EX_j^{a_{\nu\mu}})^{b_{\nu\mu}}, \quad a_{\nu\mu}, b_{\nu\mu} \in \{0, 1, \dots, r\}, b_{\nu\mu} \geq 1,$$

with $\sum_{\mu=1}^r \sum_{\nu=1}^r a_{\nu\mu} b_{\nu\mu} = r$. The number of terms (1.12) in the sum as well as the $a_{\nu\mu}$'s and $b_{\nu\mu}$'s are independent of n . In the case $r = 3$ we have for example

$$(1.13) \quad E(\sum_{i=1}^n X_i)^3 = (\sum_{i=1}^n EX_i)^3 + 2 \sum_{i=1}^n (EX_i)^3 + \sum_{i=1}^n EX_i^3 - 3 \sum_{i=1}^n EX_i^2 EX_i \\ - 3 \sum_{i=1}^n (EX_i)^2 \sum_{i=1}^n EX_i + 3 \sum_{i=1}^n EX_i^2 \sum_{i=1}^n EX_i.$$

If the X_i 's have centered $B(1, p_i)$ distribution we get from (1.12), utilizing $EX_i = 0$ and $|EX_i^s| \leq p_i, s \geq 2$:

$$(1.14) \quad E(\sum_{j=1}^n X_j)^{2r} \leq \sum_{i=1}^r A_i (\sum_{j=1}^n p_j)^i \\ \text{with } A_i \geq 0, A_i \text{ independent of } n.$$

For $s, t \in E_k$ (without loss of generality $s \leq t$ componentwise) we define $X_j = 1_t(U_n^j) - 1_s(U_n^j) - p_j, p_j = G_n^j(t) - G_n^j(s)$. Then (1.10) follows from (1.14) and (1.3). \square

REMARK. Koul (1970) used still another approach in proving relative compactness of the 1-dimensional "weighted" reduced empirical process. He employed the well-known fluctuation inequalities of Billingsley (1968) (Section 12). It is clear how to get k -dimensional versions of Koul's (1970) results by using the multidimensional fluctuation inequalities of Bickel and Wichura (1971). Since this approach is rather involved one would look for a straightforward proof of relative compactness of the empirical process. This was our aim in the proof above. On the other hand, adapting our approach to the weighted empirical process seems not quite immediate and would make the proof involved, too.

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