

THE NON-SINGULARITY OF GENERALIZED SAMPLE COVARIANCE MATRICES¹

BY MORRIS L. EATON AND MICHAEL D. PERLMAN

University of Chicago²

Let $X = (X_1, \dots, X_n)$ where the $X_i: p \times 1$ are independent random vectors, and let $A: n \times n$ be positive semi-definite symmetric. This paper establishes necessary and sufficient conditions that the random matrix XAX' be positive definite w.p. 1. The results are applied to cases where A has a particular form or X_1, \dots, X_n are i.i.d. In particular, it is shown that in the i.i.d. case, the sample covariance matrix $\Sigma(X_i - \bar{X})(X_i - \bar{X})'$ is positive definite w.p. 1 iff $P[X_i \in F] = 0$ for every proper flat $F \subset R^p$.

1. Introduction. In a recent paper, Dykstra (1970) demonstrated the non-singularity w.p. 1 of the sample covariance matrix $S = \sum_1^n (Y_i - \bar{Y})(Y_i - \bar{Y})'$ for Y_1, \dots, Y_n i.i.d. $N_p(\mu, \Sigma)$ where Σ is positive definite (henceforth written $\Sigma > 0$). A similar result also appears in lecture notes of C. Stein (1969). These demonstrations depend heavily on normality. Let $Y: p \times n$ have columns Y_1, \dots, Y_n of dimension p and write $S = YA_0Y'$ where $A_0 = I - (1/n)ee'$, $e' = (1, 1, \dots, 1) \in R^n$. In this paper, we obtain conditions under which a random matrix of the form XAX' is positive definite w.p. 1 where $A: n \times n$ is positive semi-definite and $X = (X_1, \dots, X_n): p \times n$ is a random matrix whose columns are independent but not necessarily normal or identically distributed.

2. Main results. Let $n - r$ denote the rank of $A = \{a_{ij}\}$ (assumed to be positive semi-definite) and assume that $n - r \geq p$ and $a_{ii} > 0$, $i = 1, \dots, n$. We now reduce the problem of studying the non-singularity of XAX' to the case $r = 0$. Let $\mathcal{V} \subseteq R^n$ be the range of A and let $\Gamma: r \times n$ be any matrix of rank r with row space $\mathcal{V}^\perp \equiv$ the orthogonal complement of \mathcal{V} . Now, XAX' is singular iff $\exists a \neq 0$ in R^p such that $a'XAX'a = 0$ iff $\exists a \neq 0$ in R^p such that $a'X \in \mathcal{V}^\perp$ iff $\exists a \neq 0$ in R^p and $b \in R^r$ such that $a'X = b'\Gamma$ iff the matrix $\tilde{X} = \begin{pmatrix} X \\ \Gamma \end{pmatrix}: (p+r) \times n$ is singular. Thus, studying the non-singularity of XAX' is equivalent to studying the non-singularity of \tilde{X} (or equivalently of $\tilde{X}\tilde{X}'$) where $\tilde{X}: \tilde{p} \times n$ again has independent columns and $n \geq \tilde{p} = p + r$.

REMARK. The assumption that $a_{ii} > 0$ is without essential loss of generality. For if $a_{ii} = 0$, then X_i does not occur in XAX' , so conditions on X_i are irrelevant.

Received June 1971; revised August 1972.

¹ This research was supported by the National Science Foundation Research Grant No. GP-25911.

² Both authors are now at the University of Minnesota, Minneapolis.

AMS 1970 subject classifications. Primary 62H10, 60D05; Secondary 15A03.

Key words and phrases. Independent random vectors, nonsingularity of random matrices, sample covariance matrix, linear manifolds, flats.

We now study the case of $r = 0$. Let X_1, \dots, X_n in R^p be independent random vectors, $n \geq p$, and $X = (X_1, \dots, X_n)$.

THEOREM 2.1. *The following are equivalent:*

- (i) $P\{X \text{ is singular}\} > 0$,
- (ii) *for some $s, 1 \leq s \leq p, \exists n - p + s$ columns of $X, \{X_{i_\alpha} \mid \alpha = 1, \dots, n - p + s\}$ and \exists an $(s - 1)$ -dimensional linear manifold $M^{(s-1)} \subset R^p$ such that $P\{X_{i_\alpha} \in M^{(s-1)}; \alpha = 1, \dots, n - p + s\} > 0$.*

PROOF. Clearly (ii) implies (i). The proof that (i) implies (ii) is deferred to Section 4.

Replacing X by \tilde{X} in Theorem 2.1 gives the necessary and sufficient condition that $XAX' > 0$ w.p. 1. We illustrate the application of Theorem 2.1 for the case $r = 1, n \geq p + 1$ and $\Gamma = e' = (1, 1, \dots, 1) \in R^n$. Recall that a q -dimensional flat, $F^{(q)}$, in R^p is the translate of a q -dimensional manifold.

THEOREM 2.2. *Let $A: n \times n$ have rank $n - 1$ and null space spanned by e . The following are equivalent:*

- (i) $P\{XAX' \text{ is singular}\} > 0$,
- (ii) *for some $t, 1 \leq t \leq p, \exists n - p + t$ columns of $X, \{X_{i_\alpha} \mid \alpha = 1, \dots, n - p + t\}$ and \exists a $(t - 1)$ -dimensional flat, $F^{(t-1)}$, such that $P\{X_{i_\alpha} \in F^{(t-1)}; \alpha = 1, \dots, n - p + t\} > 0$.*

PROOF. Apply Theorem 2.1 to $\tilde{X} = \begin{pmatrix} X \\ e \end{pmatrix}$ with p replaced by $p + 1$ and $t = s - 1$.

In the case $A_0 = I - (1/n)ee'$, Theorem 2.2 gives necessary and sufficient conditions that the sample covariance matrix be positive definite w.p. 1.

A very simple sufficient condition which guarantees (for any $r \leq n - p$) the positive definiteness of XAX' w.p. 1 is given below.

THEOREM 2.3. *Assume that under the distribution of each $X_i \in R^p$, every flat of dimension $p - 1$ has probability 0. Then $XAX' > 0$ w.p. 1.*

PROOF. Consider $\tilde{X} = \begin{pmatrix} X \\ \Gamma \end{pmatrix}: \tilde{p} \times n, \tilde{p} = p + r$. Let $1 \leq s \leq \tilde{p}$ and choose $n - \tilde{p} + s$ columns of \tilde{X} , say $\tilde{X}_1, \dots, \tilde{X}_{n-\tilde{p}+s}$, relabeling if necessary. If $M^{(s-1)} \subset R^{\tilde{p}}$ is an $(s - 1)$ -dimensional manifold, then $M^{(s-1)} = \{z \mid \Delta z = 0\}$ where $\Delta: (\tilde{p} - s + 1) \times \tilde{p}$ has rank $\tilde{p} - s + 1$. To apply Theorem 2.1, we must show that $P\{\Delta \tilde{X}_i = 0, i = 1, \dots, n - \tilde{p} + s\} = 0$. Partition $\Delta = (\Delta_1, \Delta_2), \Delta_1: (\tilde{p} - s + 1) \times p, \Delta_2: (\tilde{p} - s + 1) \times r$ and write

$$(\tilde{X}_1, \dots, \tilde{X}_{n-\tilde{p}+s}) = \begin{pmatrix} \tilde{X} \\ \Gamma \end{pmatrix}.$$

Thus, we must show that

$$(2.1) \quad P \left\{ (\Delta_1 \Delta_2) \begin{pmatrix} \tilde{X} \\ \Gamma \end{pmatrix} = \Delta_1 \tilde{X} + \Delta_2 \Gamma = 0 \right\} = 0.$$

If $\Delta_1 \neq 0$, then (2.1) = 0 from the assumption that all proper flats have probability

0. If $\Delta_1 = 0$, then $\text{rank}(\Delta_2) = \bar{p} - s + 1$ so $\Delta_2 \dot{\Gamma}$ cannot be zero: otherwise, $\dim(\text{range } \dot{\Gamma}) \leq r - (\bar{p} - s + 1)$ so that $r = \dim(\text{range } \Gamma) \leq \dim(\text{range } \dot{\Gamma}) + \dim(\text{range } \ddot{\Gamma}) \leq r - (\bar{p} - s + 1) + (\bar{p} - s) = r - 1$ where $\Gamma = (\dot{\Gamma} \ \ddot{\Gamma})$, a contradiction. Hence, in either case, (2.1) is 0.

The sufficient condition of Theorem 2.3 is satisfied, for example, if the distribution of X_i has a density with respect to Lebesgue measure or is orthogonally invariant.

For pedagogical purposes we give a direct proof of Theorem 2.3 which does not appeal to Theorem 2.1. Partition

$$\tilde{X} = \begin{pmatrix} \dot{X} & \ddot{X} \\ \dot{\Gamma} & \ddot{\Gamma} \end{pmatrix}$$

with $\dot{X}: p \times (n - r)$, $\ddot{X}: p \times r$, $\dot{\Gamma}: r \times (n - r)$, and $\ddot{\Gamma}: r \times r$. Permuting columns if necessary, we may assume $\ddot{\Gamma}$ is non-singular. Now \tilde{X} is singular iff $W \equiv \dot{X} - \ddot{X} \ddot{\Gamma}^{-1} \dot{\Gamma}$ is singular (to see this, premultiply \tilde{X} by the non-singular $\begin{pmatrix} I_p & \ddot{X} \ddot{\Gamma}^{-1} \\ 0 & I_r \end{pmatrix}$), so to prove $P\{\tilde{X} \text{ singular}\} = 0$ it suffices to show that $P\{W \text{ singular} \mid \ddot{X}\} = 0$. For each $i = 1, \dots, n - r$ set $S_i = \text{Span}\{W_j: j \neq i\}$, where $W = (W_1, \dots, W_{n-r})$. Then

$$\begin{aligned} & P\{W \text{ singular} \mid \ddot{X}\} \\ (2.2) \quad & \leq \sum_{i=1}^{n-r} P\{W_i \in S_i \text{ and } \dim(S_i) \leq p - 1 \mid \ddot{X}\} \\ & = \sum_{i=1}^{n-r} E[P\{W_i \in S_i \text{ and } \dim(S_i) \leq p - 1 \mid W_j, j \neq i, \ddot{X}\} \mid \ddot{X}]. \end{aligned}$$

Since \dot{X} and \ddot{X} are independent, however, for fixed \ddot{X} the random vectors (W_1, \dots, W_{n-r}) are independent, and $P\{W_i \in F \mid \ddot{X}\} = 0$ for every proper flat F in R^p . Therefore the conditional probability in the last expression in (2.2) is zero.

If we drop the assumption that X_1, \dots, X_n are independent but assume instead that the distribution of X is absolutely continuous with respect to np -dimensional Lebesgue measure, the above argument remains valid if the statements involving independence are replaced by the observation that the conditional distribution of W_i given \ddot{X} and $W_j, j \neq i$, is absolutely continuous with respect to p -dimensional Lebesgue measure. This provides an alternate proof of the theorem of Okamoto (1973).

3. Special cases. In this section we obtain necessary and sufficient conditions that $XAX' > 0$ w.p. 1 in the special cases (a) $p = 1$ or (b) X_1, \dots, X_n are i.i.d.

THEOREM 3.1. *Let $U = (U_1, \dots, U_n)$ where the U_i are independent real-valued random variables. Let $A: n \times n$ be positive semi-definite of rank $n - r \geq 1$ with range space $\mathcal{V} \subseteq R^n$ and $a_{ii} > 0, i = 1, \dots, n$. The following are equivalent:*

- (i) $P\{UAU' > 0\} = 1,$
- (ii) $P\{U \in \mathcal{V}^\perp\} = 0,$
- (iii) U has no atoms in $\mathcal{V}^\perp.$

PROOF. Clearly, (i) \Leftrightarrow (ii) \Rightarrow (iii). We show that not (i) implies not (iii). First,

we remark that the assumption $a_{ii} > 0, i = 1, \dots, n$ is equivalent to the statement that \mathscr{V}^\perp contains no coordinate axis $e_i = (0, \dots, 0, 1, 0, \dots, 0)' \in R^n$, which is equivalent to the statement that every subset of $(n - 1)$ columns of Γ span R^r .

Not (i) implies, by Theorem 2.1, that there is an $s, 1 \leq s \leq \bar{p} = r + 1$, there are $n - \bar{p} + s$ columns $\{\check{X}_{i_\alpha} | \alpha = 1, \dots, n - \bar{p} + s\}$ of $\check{X} = \binom{x}{r}$ and there is an $(s - 1)$ -dimensional manifold $M^{(s-1)} \subset R^{\bar{p}}$ such that

$$(3.1) \quad P\{\check{X}_{i_\alpha} \in M^{(s-1)}, \alpha = 1, \dots, n - \bar{p} + s\} > 0.$$

If $n - \bar{p} + s$ columns of \check{X} lie in $M^{(s-1)}$ with positive probability, then the corresponding $n - \bar{p} + s$ columns of Γ lie in a manifold contained in R^r of dimension $\leq s - 1$. If $s \leq r$, then some $n - 1$ columns of Γ lie in a manifold of dimension $\leq r - 1$ which cannot happen. Thus, $s = r + 1$ so $M^{(s-1)} = \{z | z \in R^{r+1}, \delta'z = 0\}$ where $\delta \neq 0, \delta \in R^{r+1}$. Hence (3.1) becomes

$$(3.2) \quad P\{\delta\check{X} = \delta_1 U + \check{\delta}\Gamma = 0\} > 0$$

where $\delta = (\delta_1, \check{\delta}), \delta_1 \in R^1$. Since Γ has full rank, (3.2) implies that $\delta_1 \neq 0$, so

$$P\{U = -(1/\delta_1)\check{\delta}\Gamma \in \mathscr{V}^\perp\} > 0.$$

This completes the proof.

We now turn to the case of i.i.d. random vectors in R^p . The following is an immediate consequence of Theorem 2.1.

THEOREM 3.2. *Let $X = (X_1, \dots, X_n)$ where the X_i are i.i.d. random vectors in $R^p, n \geq p$. The following are equivalent:*

- (i) $P\{X \text{ is non-singular}\} = 1,$
- (ii) $P\{X_1 \in M\} = 0$ for all proper manifolds $M \subset R^p$.

For $w \in R^p, w \neq 0$, and $c \in R^1$, let $F(w, c) = \{x \in R^p | w'x = c\}$ be the $(p - 1)$ -dimensional flat determined by w and c .

THEOREM 3.3. *Let $X = (X_1, \dots, X_n)$ where the $X_i \in R^p$ are i.i.d. and let $A: n \times n$ be positive semi-definite with range space $\mathscr{V} \subseteq R^n$ and rank $n - r \geq p$. The following are equivalent:*

- (i) $P\{XAX' > 0\} = 1,$
- (ii) $\prod_{i=1}^n P\{X_i \in F(w, v_i)\} = 0$ for all $w \in R^p, w \neq 0$ and $v = (v_1, \dots, v_n) \in \mathscr{V}^\perp$.

PROOF. If (ii) does not hold, then $\exists w \neq 0$ in R^p and $v = b'\Gamma \in \mathscr{V}^\perp$ such that $P\{w'X = b'\Gamma\} > 0$. Hence $P\{XAX' \text{ is singular}\} = P\{\binom{x}{r} \text{ is singular}\} > 0$ and (i) does not hold.

Next, assume (ii) holds. Setting $v = 0$, it follows that $P\{X_1 \in M\} = 0$ for all proper manifolds $M \subset R^p$ and therefore $P\{X_1 \in F\} = 0$ for all flats of dimension $\leq p - 2$. If (i) does not hold, then by Theorem 2.1, $\exists s, 1 \leq s \leq \bar{p} = p + r, \exists n - \bar{p} + s$ columns of $\check{X} = \binom{x}{r}$, say $\check{X}_1, \dots, \check{X}_{n-\bar{p}+s}$ (relabeling if necessary) and \exists a manifold $M^{(s-1)} \subset R^{\bar{p}}$ such that

$$(3.3) \quad P\{\check{X}_i \in M^{(s-1)}, i = 1, \dots, n - \bar{p} + s\} > 0.$$

Note that s must be ≥ 2 since $P\{X_1 = 0\} = 0$. Now, $M^{(s-1)} = \{z \mid \Delta z = 0\}$ where $\Delta: (\bar{p} + s - 1) \times \bar{p}$ has rank $\bar{p} + s - 1$, so (3.3) becomes

$$(3.4) \quad P \left\{ \Delta \begin{pmatrix} \dot{X} \\ \dot{\Gamma} \end{pmatrix} = 0 \right\} = P\{\Delta_1 X_i + \Delta_2 \gamma_i = 0, i = 1, \dots, n - \bar{p} + s\} > 0$$

where $\Delta = (\Delta_1 \Delta_2)$,

$$\begin{pmatrix} \dot{X} \\ \dot{\Gamma} \end{pmatrix} = (\dot{X}_1, \dots, \dot{X}_{n-\bar{p}+s}),$$

and $\Gamma = (\gamma_1, \dots, \gamma_{n-\bar{p}+s})$. As in the proof of Theorem 2.3, $\Delta_1 \neq 0$. However, if $\text{rank } \Delta_1 \geq 2$ then $\{x \mid x \in R^p, \Delta_1 x = -\Delta_2 \gamma_i\}$ is a flat of dimension $\leq p - 2$, contradicting (3.4). Thus $\text{rank } \Delta_1 = 1$ so $\Delta_1 = aw'$, $a: (\bar{p} - s + 1) \times 1$ and $w: p \times 1$.

Since

$$\Delta \begin{pmatrix} \dot{X} \\ \dot{\Gamma} \end{pmatrix} = 0$$

implies that

$$\Delta \begin{pmatrix} X \\ \Gamma \end{pmatrix}$$

is singular, (3.4) yields

$$(3.5) \quad \begin{aligned} 0 &< P \left\{ \Delta \begin{pmatrix} X \\ \Gamma \end{pmatrix} \text{ singular} \right\} \\ &= P\{aw'X + \Delta_2 \Gamma \text{ singular}\} \\ &= P \left\{ (a \ I) \begin{pmatrix} w'X \\ \Delta_2 \Gamma \end{pmatrix} \text{ singular} \right\}, \end{aligned}$$

where $I: (\bar{p} - s + 1) \times (\bar{p} - s + 1)$, and

$$\begin{pmatrix} w'X \\ \Delta_2 \Gamma \end{pmatrix}: (\bar{p} - s + 2) \times n.$$

Using the fact that $s \geq 2$ and $(a \ I)$ has rank $\bar{p} - s + 1$, we conclude that

$$(3.6) \quad P \left\{ \begin{pmatrix} w'X \\ \Delta_2 \Gamma \end{pmatrix} \text{ singular} \right\} > 0.$$

Since $U \equiv w'X: 1 \times n$ has independent components, Theorem 3.1 implies that there is a $y: (\bar{p} - s + 1) \times 1$ such that

$$(3.7) \quad P\{U = y'\Delta_2 \Gamma\} > 0.$$

Setting $v = y'\Delta_2 \Gamma \in \mathscr{S}^1$, (3.7) becomes

$$\prod_{i=1}^n P\{X_i \in F(w, v_i)\} > 0,$$

contradicting (ii). Hence (i) must hold.

As an immediate consequence of Theorem 3.3, we obtain

THEOREM 3.4. *Let $X = (X_1, \dots, X_n)$ where the $X_i \in R^p$ are i.i.d., and let*

$A: n \times n, n \geq p + 1$, have rank $n - 1$ with null space spanned by $e \in R^n$. The following are equivalent:

- (i) $P\{XAX' \text{ non-singular}\} = 1$,
- (ii) $P\{X_1 \in F\} = 0$ for all proper flats $F \subset R^p$.

In particular, condition (ii) is necessary and sufficient for the non-singularity w.p. 1 of the sample covariance matrix $\sum (X_i - \bar{X})(X_i - \bar{X})'$ in the i.i.d. case.

Theorem 3.3 has an interesting geometric interpretation. Let $\mathcal{S}(X) \subseteq R^n$ be the random subspace spanned by the row vectors $X_{(1)}, \dots, X_{(p)}$ of X . Theorem 3.3 states that

$$P\{\dim [\mathcal{S}(X) \cap \mathcal{V}^\perp] \geq 1\} > 0$$

iff there exists a fixed vector $v \in \mathcal{V}^\perp$ and a fixed nonzero linear combination $\sum_1^p w_i X_{(i)}$ of $X_{(1)}, \dots, X_{(p)}$ such that

$$P\{\sum_1^p w_i X_{(i)} = v\} > 0.$$

4. Proof of Theorem 2.1. Throughout this section V^p will denote a p -dimensional real vector space and $M^{(d)}$ (with or without subscripts) a manifold of dimension d in V^p .

LEMMA 4.1. *Let $Z \in V^p$ be a random vector with $P\{Z = 0\} = 0$. Then for each $d = 1, \dots, p - 1$ there are at most countably many manifolds, say $\{M_i^{(d)} \mid i = 1, \dots\}$, such that*

(a)
$$P\{Z \in M_i^{(d)}\} > 0$$

and

(b)
$$P\{Z \in M^{(d-1)}\} = 0 \quad \text{for every } M^{(d-1)} \subset M_i^{(d)}.$$

PROOF. If $M_1^{(d)}, \dots, M_m^{(d)}$ satisfy (a) and (b), then

$$P\{\bigcup_1^m M_\alpha^{(d)}\} = \sum_1^m P\{M_\alpha^{(d)}\}.$$

This shows that the collection of d -manifolds satisfying (a) and (b) is at most countable.

If v_1, \dots, v_β are vectors in V^p , $S(v_1, \dots, v_\beta)$ will denote the span of v_1, \dots, v_β . The following result yields Theorem 2.1.

THEOREM 4.2. *For every $j, 1 \leq j \leq p$, for every $k \geq j$, and for every set of independent random vectors X_1, \dots, X_k in V^p , the following are equivalent:*

- (i) $P\{\dim S(X_1, \dots, X_k) \leq j - 1\} > 0$.
- (ii) *For some $s, 1 \leq s \leq j$, the following assertion, $A(s)$, holds: there exist $k - j + s$ vectors $\{X_{i_\alpha} \mid \alpha = 1, \dots, k - j + s\}$ and there exists $M^{(s-1)}$ such that $P\{X_{i_\alpha} \in M^{(s-1)}, \alpha = 1, \dots, k - j + s\} > 0$.*

PROOF. Clearly (ii) implies (i). We prove that (i) implies (ii) by induction on p . This is easily verified for $p = 1$. For $p \geq 2$, assume the result is true for all $p', 1 \leq p' \leq p - 1$.

To prove the result for dimension p , we use a secondary induction on j . The result is clear for $j = 1$. For $2 \leq j \leq p$, assume (i) implies (ii) for all j' , $1 \leq j' \leq j - 1$. Now fix $k \geq j$ and X_1, \dots, X_k in V^p . Assume that

$$(4.1) \quad P\{\dim S(X_1, \dots, X_k) \leq j - 1\} > 0$$

but that $A(1), \dots, A(j - 1)$ do not hold; we will show that $A(j)$ obtains. Since $A(1)$ does not hold, there is some X_i , say X_1 , such that $P\{X_1 = 0\} = 0$. Using the secondary induction hypothesis for $j' \equiv j - 1$ and $k' \equiv k - 1$, not $A(1), \dots$, not $A(j - 1)$ implies

$$P\{\dim S(X_2, \dots, X_k) \leq j - 2\} = 0$$

so

$$(4.2) \quad P\{\dim S(X_2, \dots, X_k) \geq j - 1\} = 1.$$

Combining (4.1) and (4.2) gives

$$P\{\dim S(X_2, \dots, X_k) = j - 1, \dim S(X_1, \dots, X_k) \leq j - 1\} > 0$$

so

$$(4.3) \quad P\{X_1 \in S(X_2, \dots, X_k), \dim S(X_2, \dots, X_k) = j - 1\} > 0.$$

Now (4.3) can be written as

$$(4.4) \quad 0 < \int_C P\{X_1 \in S(X_2, \dots, X_k) | X_2, \dots, X_k\} dQ$$

where Q is the probability distribution of (X_2, \dots, X_k) and $C = D \cap E$, where D and E are the following events:

$$D = \{\dim S(X_2, \dots, X_k) = j - 1\}$$

$$E = \{P\{X_1 \in S(X_2, \dots, X_k) | X_2, \dots, X_k\} > 0\}.$$

By Lemma 4.1 with $Z = X_1$,

$$C = \bigcup_{d=1}^{j-1} \bigcup_{i=1}^{\infty} [\{M_i^{(d)} \subseteq S(X_2, \dots, X_k)\} \cap D].$$

From (4.4), there exists a d and an i such that

$$(4.5) \quad P\{X_1 \in M_i^{(d)}\} P[\{M_i^{(d)} \subseteq S(X_2, \dots, X_k)\} \cap D] > 0.$$

If $d = j - 1$, then $A(j)$ obtains with $M^{(j-1)} = M_i^{(d)}$. If $1 \leq d \leq j - 2$, then we have

$$(4.6) \quad P[\{M_i^{(d)} \subseteq S(X_2, \dots, X_k)\} \cap D] > 0.$$

Let Π denote the orthogonal projection onto $(M_i^{(d)})^\perp$. Clearly,

$$\Pi(S(X_2, \dots, X_k)) = S(\Pi X_2, \dots, \Pi X_k),$$

so (4.6) implies that

$$(4.7) \quad P\{\dim S(\Pi X_2, \dots, \Pi X_k) \leq j - 1 - d\} > 0.$$

Set $p' = p - d$, $j' = j - d$ and $k' = k - 1$. Since $\Pi X_\alpha \in (M_i^{(d)})^\perp$ and $\dim (M_i^{(d)})^\perp = p' \leq p - 1$, we can apply our original induction hypothesis.

Hence for some s' , $1 \leq s' \leq j'$, there exists $k' - j' + s'$ vectors $\{\Pi X_{i_\alpha} \mid \alpha = 1, \dots, k' - j' + s'\}$ and a manifold $M^{(s'-1)} \subset (M_i^{(d)})^\perp$ such that

$$P\{\Pi X_{i_\alpha} \in M^{(s'-1)}, \alpha = 1, \dots, k' - j' + s'\} > 0.$$

Now, let $M_0 = M^{(s'-1)} \oplus M_i^{(d)} \subset V^p$ so

$$(4.8) \quad P\{X_{i_\alpha} \in M_0, \alpha = 1, \dots, k' - j' + s'\} > 0.$$

Combining (4.8) and (4.5), we have

$$(4.9) \quad P\{X_1 \in M_0, X_{i_\alpha} \in M_0, \alpha = 1, \dots, k' - j' + s'\} > 0.$$

Since we have assumed that $A(1), \dots, A(j-1)$ do not hold, it follows that $1 + k' - j' + s' = k$ and $\dim(M_0) = j - 1$. Thus $A(j)$ obtains, completing the proof.

Theorem (2.1) follows by setting $j = p$ and $k = n$ in Theorem 4.2.

REFERENCES

- [1] DYKSTRA, R. L. (1970). Establishing the positive definiteness of the sample covariance matrix. *Ann. Math. Statist.* **41** 2153-2154.
- [2] OKAMOTO, M. (1973). Distinctness of the eigenvalues of a quadratic form in a multivariate sample. *Ann. Statist.* **1** 763-765.
- [3] STEIN, C. (1969). *Multivariate Analysis I*. (Lecture notes recorded by M. L. Eaton). Technical Report No. 42, Department of Statistics, Stanford Univ.

DEPARTMENT OF STATISTICS
UNIVERSITY OF MINNESOTA
MINNEAPOLIS, MINNESOTA 55455