

## SUPERCritical AGE DEPENDENT BRANCHING PROCESSES WITH GENERATION DEPENDENCE<sup>1</sup>

BY DEAN H. FEARN

*California State University, Hayward*

This paper examines the size,  $Z(t)$ , of a population as a function of time.  $Z(t)$  is just like the ordinary Bellman-Harris age dependent branching process except that the number of daughters born to an individual in the  $n$ th generation is allowed to depend on  $n$ . The renewal theory of William Feller and Laplace transform theory are used to obtain the behavior of  $EZ(t)$  as  $t$  approaches infinity, and the convergence of  $Z(t)/E(Z(t))$  in quadratic mean.

**1. Introduction.** In this paper the following branching process is treated: The process starts with one cell in the  $n$ th generation. This cell lives for a random length of time  $T_{n,1}$  with distribution function  $G$ , and then splits into a random number,  $\zeta_{n,1}$ , of baby cells in the  $n + 1$ st generation. It is assumed throughout that  $G$  is nonlattice, and  $G(0) = 0$ . Each of these cells respectively live random lengths of time  $T_{n+1,1}, \dots, T_{n+1,\zeta_{n,1}}$  and split respectively into  $\zeta_{n+1,1}, \dots, \zeta_{n+1,\zeta_{n,1}}$  cells in the  $n + 2$ nd generation. This process continues;  $Z_n(t)$  is the number of cells alive at time  $t$  having started with one cell in the  $n$ th generation. It is assumed throughout that the random variables  $T_{m,j}, \zeta_{n,k}$ , where  $m$  and  $n$  run through the nonnegative integers and  $j$  and  $k$  run through the positive integers, are mutually independent; that the random variables  $T_{n,k}$  for  $n \geq 0$  and  $k \geq 1$  have the same distribution; and that for each fixed  $n$ , the random variables  $\zeta_{n,k}, k = 1, 2, \dots$  have a probability distribution on the nonnegative integers depending only on  $n$ . Let  $G_0$  be the distribution function which puts mass one on  $t = 0$ . Also the notation

$$(1.1) \quad (A * B)(t) = \int_0^t A(t-u) dB(u) = \lim_{\tau \rightarrow 0+} \int_{-\tau}^{t+\tau} A(t-u) dB(u)$$

will be adopted whenever the right side of (1.1) exists. All integrals in this paper are Lebesgue-Stieltjes integrals. Also for any sequence  $a_n$

$$(1.2) \quad \sum_{n=k}^{k-1} a_n = 0 \quad \text{and} \quad \prod_{n=k}^{k-1} a_n = 1, \quad \text{for all integers } k.$$

The Laplace transform of a function  $A$  is

$$\mathcal{L}A(\lambda) = \int_0^\infty e^{-\lambda t} dA(t)$$

whenever this exists. If  $T_1, \dots, T_k$  are  $k$  independent random variables with distribution function  $G$ , let

$$(1.3) \quad G_k(t) = P[\sum_{j=1}^k T_j \leq t] = (G_{k-1} * G)(t).$$

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Now suppose throughout for some finite number  $B$  that

$$(1.4) \quad m_n = E(\zeta_{n,k}) \leq B, \quad n = 0, 1, \dots; k = 1, 2, \dots$$

Let  $\bar{Z}_n(t)$  be the total number of cells ever alive by time  $t$ , having started with one cell in the  $n$ th generation.

It is shown in Theorem 3.1 that, when  $G$  has a finite first moment  $\mu$ ,  $m_n \geq 1$ , and  $(\prod_{j=0}^{n-1} m_j)/(n^\rho L(n)) \rightarrow 1$  as  $n \rightarrow \infty$  with  $L$  slowly varying and  $\rho \geq 0$ , one has  $E(Z_0(t))/[(t/\mu)^\rho L(t/\mu)] \rightarrow 1$  as  $t \rightarrow \infty$ . Roughly speaking, the conclusion of this theorem means that  $E(Z_0(t))$  behaves like  $n^\rho L(n)$  where  $n$  is the number of average lifespans in the time interval  $(0, t)$ , provided  $\mu =$  one unit of time. But  $n^\rho L(n)$  behaves like  $\prod_{j=0}^{n-1} m_j$ , so  $E(Z_0(t))$  recapitulates  $\prod_{j=0}^{n-1} m_j$ , the mean of an imbedded Galton–Watson process,  $Z_n$ , which is the number of cells in the  $n$ th generation. Using the renewal theory contained in Lemmas 2.3, 2.4, 2.5 and 3.2 of this paper, this result is extended (see Theorem 3.2) to the case where  $m_n \geq m > 1$  for all  $n$ , and  $(\prod_{j=0}^{n-1} (m_j/m))/n^\rho L(n) \rightarrow 1$ , as  $n \rightarrow \infty$ , with  $\rho \geq 0$  and with  $L$  slowly varying and nondecreasing. Finally in Section 4, Theorems 4.2 and 4.3 give sufficient conditions for  $Z_0(t)/E(Z_0(t))$  to converge in quadratic mean to a nondegenerate random variable  $W$  as  $t \rightarrow \infty$ , thus giving information about the sample path behavior of  $Z_0(t)$  for large  $t$ .

## 2. Basic renewal theory.

LEMMA 2.1. *For any finite interval  $I$ , there is a finite number  $B_I$  such that*

$$(2.1) \quad E(\bar{Z}_n(t)) = \sum_{k=n}^{\infty} \prod_{j=n}^{k-1} m_j G_{k-n}(t) \leq B_I$$

for  $n = 0, 1, 2, \dots$ , and for all  $t \in I$ .

PROOF. In Harris (1963, page 139), “ $\sum_{i_j} P[\nu_{i_1, \dots, i_{j-1}} \geq i_j]$ ” becomes  $m_{j-1}$ . The equation in (2.1) follows. The boundedness in (2.1) follows upon evaluating the Laplace transform of  $E(\bar{Z}_n(t))$  for a suitably large value of  $\lambda$ .

This yields the following basic lemma concerning  $Z_n(t)$ :

LEMMA 2.2. *For any finite interval  $I$ , there is a finite number  $B_I$  such that*

$$(2.2) \quad M_n(t) \equiv E(Z_n(t)) \leq B_I$$

for  $n = 0, 1, \dots$ , and all  $t \in I$ . Hence,  $Z_n(t)$  is almost surely finite.

PROOF. This is an immediate consequence of the fact that  $Z_n(t) \leq \bar{Z}_n(t)$ .

Some important lemmas will now be stated whose proofs are obtained by suitably modifying arguments used in Feller (1966), pages 181–183 and pages 346–353.

LEMMA 2.3. *Let  $Y_n$  be a sequence of functions uniformly bounded on finite intervals, such that  $(Y_n * G)(t)$  exists for all  $t$  and all  $n$ . Suppose  $X_n$  is a sequence of functions, uniformly bounded on finite intervals, and for all  $t$  satisfying the renewal equations*

$$(2.3) \quad X_n(t) = Y_n(t) + m_n(X_{n+1} * G)(t), \quad n = 0, 1, 2, \dots$$

Then, for all  $t$

$$(2.4) \quad X_n(t) = \sum_{k=n}^{\infty} (\prod_{j=n}^{k-1} m_j)(Y_k * G_{k-n})(t), \quad n = 0, 1, 2, \dots$$

PROOF. (This proof uses, with some notational changes, the argument in Feller (1966, pages 181-183).) Call the right hand side of (2.4)  $\bar{X}_n(t)$ . Using (1.4)

$$\bar{X}_n(t) \leq \sum_{k=n}^{\infty} B^{k-n} [\max_{0 \leq u \leq t} |Y_k(u)|] G_{k-n}(t).$$

The argument showing the uniform boundedness of  $E(\bar{Z}_n(t))$  on finite intervals shows that  $\bar{X}_n$  has this property since  $Y_k$  is uniformly bounded on finite intervals. A direct calculation shows  $\bar{X}_n$  satisfies (2.3).

Now let  $V_n(t) = \bar{X}_n(t) - X_n(t)$ . Then  $V_n$  is a sequence of functions, uniformly bounded on finite intervals. Using (2.3) recursively one finds that

$$V_n(t) = (\prod_{j=n}^{n+k} m_j)(V_{n+k+1} * G_{k+1})(t).$$

The uniform boundedness of  $V_n(u)$  for  $n = 0, 1, 2, \dots$  and  $u$  in  $[0, t]$ , and the fact that, as has been seen,  $\sum_{k=0}^{\infty} (\prod_{j=n}^{n+k} m_j) G_{k+1}(t) < \infty$  show that  $V_n(t) = 0$  for all  $t$  and all  $n$ . So  $X_n(t) = \bar{X}_n(t)$  and (2.4) is true for all  $t$  and all  $n$ .

LEMMA 2.4. *Let  $h$  be any positive continuous, nondecreasing function such that for every number  $a$ ,  $h(t+a)/h(t)$  approaches 1 as  $t$  increases to  $\infty$ . Let  $H$  be a right continuous function, nondecreasing on  $[0, \infty)$ . Assume that  $h(t) = h(0)$  if  $t < 0$  and  $H(t) = 0$  if  $t \leq 0$ . Then for every  $\varepsilon > 0$ ,*

$$(2.5) \quad \frac{H(t) - H(t - \varepsilon)}{h(t)} \rightarrow c\varepsilon \quad \text{as } t \rightarrow \infty,$$

*if and only if  $X$  is directly Riemann integrable implies*

$$(2.6) \quad \frac{(X * H)(t)}{h(t)} \rightarrow c \int_0^{\infty} X(u) du \quad \text{as } t \rightarrow \infty.$$

PROOF. (This proof uses, with notational changes, the methods found in Feller (1966, pages 348-350).) To see that (2.6) implies (2.5), let  $X(t) = 1$  if  $0 \leq t \leq \varepsilon$  and  $X(t) = 0$  otherwise. (2.5) follows from (2.6) by a direct calculation. It is now shown that (2.5) implies (2.6).

Let  $\varepsilon > 0$  be given. Let  $X_n(t) = 1$  if  $(n-1)\varepsilon \leq t \leq n\varepsilon$  and  $X_n(t) = 0$  otherwise. Let  $X_n^*(t) = \int_0^{\infty} X_n(t-u) dH(u)$ . Then  $X_n^*(t) = H(t - (n-1)\varepsilon) - H(t - n\varepsilon)$ . By (2.5) one has

$$\frac{H(t - (n-1)\varepsilon) - H(t - (n-1)\varepsilon - \varepsilon)}{h(t - (n-1)\varepsilon)} \rightarrow c\varepsilon \quad \text{as } t \rightarrow \infty,$$

and also  $h(t - (n-1)\varepsilon)/h(t)$  approaches one from below as  $t \rightarrow \infty$ ; hence  $X_n^*(t)/h(t) \rightarrow c\varepsilon$  as  $t \rightarrow \infty$  for fixed  $n$ . Thus  $X_1^*(t)/h(t)$  is bounded in  $t$  due to the positive, continuous, and nondecreasing nature of  $h$ . Also because of these properties of  $h$ ,  $X_1^*(t)/h(t) \leq X_n^*(t - (n-1)\varepsilon)/h(t - (n-1)\varepsilon)$ , hence  $X_n^*(t)/h(t)$  is bounded for all  $n$  and  $t$  by (say)  $M_\varepsilon < \infty$ .

Suppose  $a_k \geq 0$  and  $\sum_{k=1}^{\infty} a_k < \infty$ . Define  $X(t) = \sum_{k=1}^{\infty} a_k X_k(t)$  and  $X^*(t) = (X * H)(t)$ . Then plainly,

$$(2.7) \quad \sum_{k=1}^n \frac{a_k X_k^*(t)}{h(t)} \leq \frac{X^*(t)}{h(t)} \leq \sum_{k=1}^n \frac{a_k X_k^*(t)}{h(t)} + M_\epsilon \sum_{k=n+1}^{\infty} a_k.$$

Let  $t \rightarrow \infty$  in (2.7), and then in the resulting series of inequalities let  $n \rightarrow \infty$ . Using the Lebesgue dominated convergence theorem to pass limits under summation and integral signs and the fact that  $\sum_{k=1}^{\infty} a_k < \infty$  one obtains

$$(2.8) \quad X^*(t)/h(t) = (X * H)(t)/h(t) \rightarrow c \int_0^\infty X(u) du \quad \text{as } t \rightarrow \infty.$$

The arbitrary nature of  $\epsilon$  and  $a_k$  allow (2.8) to hold whenever  $X$  is directly Riemann integrable, so (2.5) implies (2.6) and Lemma 2.4 is true.

The next lemma is a slight generalization of the key renewal theorem. The proof is obtained by letting  $H_n^*$  and  $f$  in the lemma below correspond to  $U$  and  $Z$  respectively in the proof on page 350 of Feller (1966) and considering  $(f * H_n^*)(x)/h(x)$  in the same way as  $(Z * U)(x)$  is considered in Feller's proof.

LEMMA 2.5. *If  $m_j$  is a sequence of positive numbers approaching 1 as  $j$  approaches  $\infty$ ; if*

$$H_n^*(t) = \sum_{k=n+1}^{\infty} (\prod_{j=n}^{k-1} m_j) G_{k-n}(t), \quad n = 0, 1, 2, \dots;$$

*if  $h(t)$  is a continuous positive nondecreasing function such that for every number  $q$ ,*

$$\frac{h(t+q)}{h(t)} \rightarrow 1 \quad \text{as } t \rightarrow \infty;$$

*then, whenever  $f$  is any right continuous, nonnegative, directly Riemann integrable function, with  $f(t) = f(0) > 0$  for  $t < 0$  and with  $f$  satisfying, for some number  $c > 0$ ,*

$$(2.9) \quad \frac{X_f(t)}{h(t)} \equiv \frac{(f * H_0^*)(t)}{h(t)} \rightarrow c \quad \text{as } t \rightarrow \infty,$$

*one has*

$$(2.10) \quad \frac{H_0^*(t) - H_0^*(t - \epsilon)}{h(t)} \rightarrow \frac{c\epsilon}{\int_0^\infty f(u) du} \quad \text{as } t \rightarrow \infty.$$

PROOF. The limit statement in (2.9) of the hypothesis of this lemma, the fact that  $f$  is directly Riemann integrable, and the positive nondecreasing nature of  $h$  together imply that  $X_f(t)/h(t)$  is bounded in  $t$  by (say)  $M_f < \infty$ . Also, the right continuity of  $f$  allows the choice of a number  $A > 0$  sufficiently small so that  $f(u) > f(0)/2 = \delta > 0$ , whenever  $0 \leq u \leq A$ . Hence for  $0 \leq v \leq A$ , and all  $t \geq 0$ ,

$$(2.11) \quad \frac{M_f}{\delta} \geq \frac{X_f(t)}{\delta h(t)} \geq \frac{\int_{t-v}^t f(t-u) dH_0^*(u)}{\delta h(t)} \geq \frac{H_0^*(t) - H_0^*(t-v)}{h(t)}.$$

Now define for any function  $E$ , and any finite interval  $I = [a, b]$ ,  $E(t+I) = E(t+b) - E(t+a)$ . The inequalities in (2.11), the nondecreasing nature of

$h$ , and the fact that any finite interval  $I$  can be broken up into subintervals of length less than or equal to  $A$ , allow  $H_0^*(t + I)/h(t)$  to be bounded in  $t$  for all finite intervals  $I$ .

The same argument was used in Feller (1966, page 350), to show that his  $U(t + I)$  was bounded in  $t$  for finite intervals  $I$ , except that there was no function  $h$  in the denominator. Following the proof in Feller (1966, page 350), the selection theorem ([3], page 263) gives a measure  $V_0$ , and a sequence  $t_k \rightarrow \infty$ , such that for each closed interval  $I$

$$(2.12) \quad H_0^*(t_k + I)/h(t_k) \rightarrow V_0(I) \quad \text{as } k \rightarrow \infty .$$

The limit statement (2.12) and the fact that by hypothesis  $m_n \rightarrow 1$  as  $n \rightarrow \infty$  allow the conclusion that

$$(2.13) \quad m_0 H_1^*(t_k + I)/h(t_k) \rightarrow V_0(I) \quad \text{as } k \rightarrow \infty .$$

A detailed argument to this effect is in [1].

Now suppose  $a$  is a positive number and that  $A$  is a continuous function vanishing outside of  $[0, a]$ . Then one has for all  $t$  and all  $t_k$ ,

$$(2.14) \quad \begin{aligned} \int_{-\infty}^{\infty} A(t - u) dH_0^*(u + t_k) \\ = m_0 \int_{-\infty}^{\infty} A(t - u) dG(u + t_k) \\ + \int_0^{\infty} (m_0 \int_{-\infty}^{\infty} A(t - u - v) dH_1^*(v + t_k)) dG(u) . \end{aligned}$$

Then dividing both sides of this equation by  $h(t_k)$  and letting  $k \rightarrow \infty$  yields the all important functional equation

$$(2.15) \quad \begin{aligned} \zeta(t) &= \int_0^{\infty} \zeta(t - u) dG(u) \quad \text{where} \\ \zeta(t) &\equiv \int_{-\infty}^{\infty} A(t - u) dV_0(u) . \end{aligned}$$

Equation (2.14), when divided by  $h(t_k)$ , corresponds to (1.14) in Feller (1966, page 350). Equation (2.15) is equation (1.15) in Feller (1966, page 350), with “ $G$ ” instead of “ $F$ ”. The rest of the proof of this lemma can now be obtained by plugging  $(H_0^*(t_k) - H_0^*(t - h))/h(t_k)$  and  $V_0$  (respectively) in for Feller’s  $U(t_k) - U(t_k - h)$  and  $V$  (respectively) on page 351 of Feller (1966), and using the argument found there. The details of this can be found in [1].

**3. Limiting behavior of  $M_n(t)$  as  $t \rightarrow \infty$ .** Lemmas 2.3, 2.4, and 2.5 constitute a body of renewal theory which will be sufficient to determine the asymptotic behavior of  $M_0(t)$  as  $t \rightarrow \infty$ , under certain assumptions on the  $m_n$ ’s. To this end the following lemma is proven.

LEMMA 3.1. For  $n = 0, 1, 2, \dots$

$$(3.1) \quad M_n(t) = 1 - G(t) + m_n(M_{n+1} * G)(t) .$$

PROOF.  $M_n(t)$  is the expected number of cells alive at time  $t$ , starting with one  $n$ th generation cell. But this is the expected number of  $n$ th generation cells alive at time  $t$  (i.e.,  $1 - G(t)$ ), plus the expected number,  $m_n(M_{n+1} * G)(t)$ , of

cells descending from an average of  $m_n$  cells in the  $n + 1$ st generation produced by the  $n$ th generation cell dying at or before time  $t$ .

LEMMA 3.2. For  $n = 0, 1, 2, \dots$

$$(3.2) \quad M_n(t) = \int_0^t 1 - G(t - u) dH_n(u) \quad \text{where}$$

$$(3.3) \quad H_n(t) = \sum_{k=n}^{\infty} (\prod_{j=n}^{k-1} m_j) G_{k-n}(t).$$

PROOF.  $1 - G(t)$  is bounded,  $\int_0^t 1 - G(t - u) dG(u)$  exists, so by Lemma 2.3 and Fubini's theorem, (3.2) is true for  $n = 0, 1, 2, \dots$ . (Here  $1 - G(t)$  plays the same role as  $Y_n(t)$  does in Lemma 2.3.)

From now on (3.3) will remain in force, as a definition of  $H_n$ .

THEOREM 3.1. If  $m_n \geq 1$  for all  $n$ ;

$$(3.4) \quad \prod_{j=0}^{n-1} m_j / (n^\rho L(n)) \rightarrow 1 \quad \text{as } n \rightarrow \infty,$$

where  $L(n)$  is slowly varying (i.e.,  $L(an)/L(n) \rightarrow 1$  as  $n \rightarrow \infty$  for any  $a > 0$ ) and  $\rho \geq 0$ ; and  $\mu = \int_0^\infty t dG(t) < \infty$ ; then

$$(3.5) \quad \frac{M_0(t)}{t^\rho L(t)} \rightarrow \frac{1}{\mu^\rho} \quad \text{as } t \rightarrow \infty.$$

PROOF. First take the Laplace transform of each side of (3.2). Next divide the resulting equation by

$$(1 - \mathcal{L}G(\lambda))^{-\rho} L\left(\frac{1}{1 - \mathcal{L}G(\lambda)}\right) \Gamma(\rho + 1).$$

Letting  $\lambda$  decrease to zero one obtains

$$\frac{\mathcal{L}M_0(\lambda)}{\Gamma(\rho + 1)\lambda^{-\rho}\mu^{-\rho}L(1/\lambda)} \rightarrow 1,$$

using Theorem 5, page 423 of Feller (1966). This is true, since  $\mathcal{L}G(\lambda)$  increases to one as  $\lambda$  decreases to zero,  $(1 - \mathcal{L}G(\lambda))/\lambda \rightarrow \mu$  as  $\lambda$  decreases to zero, and  $L$  is slowly varying. Now if it were true that  $M_0(t)$  is a monotone function of  $t$ , Theorem 2, page 421 of Feller (1966) could be applied to obtain the desired result. This monotonicity can be established by forcing each  $n$ th generation cell to split into at least one  $n + 1$ st generation cell, while splitting into  $m_n$  cells on the average. The resulting branching process has the same expected value, and is nondecreasing in  $t$ . So (3.5) is true, since the nondecreasing and right continuous nature of  $M_0(t)$  allows  $M_0(t)$  to define a measure.

THEOREM 3.2. If  $m > 1$ , as  $n \rightarrow \infty$ ,  $m_n \geq m$  for all  $n$ ;

$$(\prod_{j=0}^{n-1} m_j^*) / n^\rho L(n) \rightarrow 1$$

as  $n \rightarrow \infty$ ;  $m_n^* = m_n/m$  for all  $n$ ;  $\rho > 0$  and  $L(n)$  is as in the hypotheses of Theorem 3.1, with  $L$  nondecreasing; and  $\alpha$  is the (positive) number satisfying

$m \int_0^\infty e^{-\alpha t} dG(t) = 1$ , then

$$(3.6) \quad \frac{M_0(t)}{e^{\alpha t} t^\rho L(t)} \rightarrow \frac{\int_0^\infty (1 - G(u)) e^{-\alpha u} du}{\bar{\mu}^\rho \int_0^\infty (1 - \bar{G}(u)) du}$$

where  $\bar{\mu} = m \int_0^\infty t e^{-\alpha t} dG(t)$  and  $\bar{G}(t) = m \int_0^t e^{-\alpha u} dG(u)$ .

PROOF. By Lemma 3.2,  $M_0(t) = \int_0^t 1 - G(t - u) dH_0(u)$ . Multiply both sides of this equation by  $e^{-\alpha t}$ . Then

$$\bar{M}_0(t) = M_0(t) e^{-\alpha t} = \int_0^t (1 - G(t - u)) e^{-\alpha(t-u)} d\bar{H}_0(u),$$

where  $\bar{H}_0(t) = \sum_{n=0}^\infty (\prod_{j=0}^{n-1} m_j^*) \bar{G}_n(t)$ , and  $\bar{G}_n(t) = m^n \int_0^t e^{-\alpha u} dG_n(u)$ , for  $n = 0, 1, 2, \dots$ . To determine the behavior of  $M_0(t)$  we determine the behavior of  $B(t) = \int_0^t 1 - \bar{G}(t - u) d\bar{H}_0(u)$ . Now  $m_n^*$  satisfies the hypotheses of Theorem 3.1, hence  $B(t)/(t^\rho L(t)) \rightarrow 1/\bar{\mu}^\rho$  as  $t \rightarrow \infty$ . Since  $1/(t^\rho L(t)) \rightarrow 0$  as  $t \rightarrow \infty$ , it is true by Theorem 3.1 that  $(1 - \bar{G}(t)) * \bar{H}_0^*(t)/(t^\rho L(t))$  approaches  $1/\bar{\mu}^\rho$  as  $t \rightarrow \infty$ , where  $\bar{H}_0^*(t) = \bar{H}_0(t) - 1$ . Now let  $f(t) = 1 - \bar{G}(t)$ ,  $h(t) = t^\rho L(t)$ ,  $X_f(t) = \int_0^t 1 - \bar{G}(t - u) d\bar{H}_0^*(u)$ ,  $c = 1/\bar{\mu}^\rho$ , and  $H_0^*(t) = \bar{H}_0^*(t)$ . Then  $f, h, c, H_0^*, X_f$  satisfy the hypotheses of Lemma 2.5. So, by (2.10)

$$\frac{\bar{H}_0^*(t) - \bar{H}_0^*(t - \varepsilon)}{t^\rho L(t)} \rightarrow \frac{\varepsilon}{\bar{\mu}^\rho \int_0^\infty (1 - \bar{G}(u)) du}$$

for each  $\varepsilon > 0$ , as  $t \rightarrow \infty$ . Now, since  $(1 - G(t))e^{-\alpha t}$  is directly Riemann integrable, it follows from Lemma 2.4 that

$$\frac{((1 - G(t))e^{-\alpha t}) * \bar{H}_0^*(t)}{t^\rho L(t)} \rightarrow \frac{\int_0^\infty (1 - G(u)) e^{-\alpha u} du}{\bar{\mu}^\rho \int_0^\infty (1 - \bar{G}(u)) du}$$

as  $t \rightarrow \infty$ . Now  $(1 - G(t))e^{-\alpha t}/(t^\rho L(t)) \rightarrow 0$  as  $t \rightarrow 0$ , so (3.6) follows.

**4. The convergence of  $W(t) \equiv Z_0(t)/M_0(t)$  in quadratic mean.** In this section it is assumed that  $m_n$  and  $\sigma_n = E(\xi_{n,k}^2) - m_n$  are both bounded by a positive finite number  $B$  independent of  $n$ . Moreover it is supposed that  $m \geq 1$ ;  $m_n \geq m$  for all  $n$ ; and  $((\prod_{j=0}^{n-1} (m_j/m))/(n^\rho L(n)) \rightarrow 1$  as  $n \rightarrow \infty$ , where  $\rho > 0$  and  $L$  is a positive nondecreasing function satisfying  $L(cn)/L(n) \rightarrow 1$  as  $n \rightarrow \infty$ , whenever  $c > 0$ . Suppose  $\alpha$  is the number defined in Theorem 3.2 (with  $\alpha = 0$  when  $m = 1$ ).

The behavior of  $W(t)$  in quadratic mean will be determined by considering the asymptotic behavior of  $C_n(t, \tau) = E(Z_n(t)Z_n(t + \tau))$  as  $t \rightarrow \infty$ . First we give the renewal equation for  $C_n(t, \tau)$ .

LEMMA 4.1. *The following renewal equation holds for  $n = 0, 1, 2, \dots$  and for  $t, \tau \geq 0$ :*

$$(4.1) \quad C_n(t, \tau) = F_n(t, \tau) + m_n \int_0^t C_{n+1}(t - u, \tau) dG(u) \quad \text{where}$$

$$(4.2) \quad F_n(t, \tau) = \sigma_n \int_0^t M_{n+1}(t - u) M_{n+1}(t - u + \tau) dG(u) \\ + m_n \int_0^{t+\tau} M_{n+1}(t + \tau - u) G(u) + 1 - G(t + \tau).$$

PROOF. This lemma is easily established by, as in the proof of Lemma 3.1, conditioning on what happens when the first cell in the branching process dies.

The following lemma is a consequence of (4.1) and Lemma 2.3.

LEMMA 4.2. *For all  $n, t, \tau \geq 0$ ,*

$$(4.3) \quad C_n(t, \tau) = F_n(t, \tau) + \sum_{k=n+1}^{\infty} \left( \prod_{j=n}^{k-1} m_j \right) \int_0^t F_k(t-u, \tau) dG_{k-n}(u).$$

PROOF. Equation (4.3) follows from Lemma 4.1 once it is seen that  $F_n(t, \tau)$  (for fixed  $\tau$ ) is uniformly bounded in  $n$  on finite intervals. Due to the boundedness of  $\sigma_n$  and  $m_n$  and the fact that  $G$  is a probability distribution, it follows that  $F_n(t, \tau)$  is bounded uniformly in  $n$  on finite intervals if  $M_n(t)$  has this property. But this is true by Lemma 3.2. The asymptotic behavior of  $C_0(t, \tau)$  is given by the following theorem.

THEOREM 4.1. *If  $\rho > 1$  or  $m > 1$ , then*

$$(4.4) \quad C_0(t, \tau)/h(t, \tau) \rightarrow \sum_{k=0}^{\infty} \frac{C^2 \sigma_k \prod_{j=0}^{k-1} m_j}{(m^*)^{k+1} \left( \prod_{j=0}^k (m_j/m) \right)^2} \equiv Q \quad \text{as } t \rightarrow \infty,$$

*uniformly in  $\tau$ , where*

$$(4.5) \quad h(t, \tau) = e^{2\alpha t + \alpha \tau} t^\rho (t + \tau)^\rho L(t) L(t + \tau),$$

$$(4.6) \quad C = \lim_{t \rightarrow \infty} M_0(t)/(t^\rho L(t) e^{\alpha t}), \quad \text{and}$$

$$(4.7) \quad m^* = \left( \int_0^\infty e^{-2\alpha t} dG(t) \right)^{-1}.$$

PROOF. By Lemma 4.2, we may define  $D(t, \tau)$  by

$$(4.8) \quad \frac{C_0(t, \tau)}{h(t, \tau)} = D(t, \tau) + \sum_{k=0}^{\infty} \frac{\sigma_k \prod_{j=0}^{k-1} m_j \int_0^t M_{k+1}(t-u) M_{k+1}(t-u+\tau) dG_{k+1}(u)}{h(t, \tau)}.$$

It follows from Theorem 3.2 that  $D(t, \tau) \rightarrow 0$  uniformly in  $t$  as  $t \rightarrow \infty$ . Now, consider the series in (4.8).

First suppose  $m > 1$ . Using Lemma 3.2, the series in (4.8) is dominated termwise, for a sufficiently large  $k$ , chosen independently of  $t$  and  $\tau$ , by the series

$$\sum_{k=0}^{\infty} (k+1)^\rho \sigma_k \int_0^t \frac{M_0(t-u) M_0(t-u+\tau)}{h(t-u, \tau)} (m^*)^k e^{-2\alpha u} dG_{k+1}(u) \left( \frac{m}{m^*} \right)^k.$$

By Theorem 3.2 and the definition of  $m^*$ , the convolution in the above series is bounded in all variables. Moreover recall that  $m = \left( \int_0^\infty e^{-\alpha t} dG(t) \right)^{-1}$ ,  $m^* = \left( \int_0^\infty e^{-2\alpha t} dG(t) \right)^{-1}$  so  $m/m^* < 1$  and the above series is dominated termwise by a convergent geometric series independent of  $t$  and  $\tau$ . Hence in case  $m > 1$ , (4.4) follows by applying the Lebesgue dominated convergence theorem to (4.8) and using Theorem 3.2.

Next suppose  $\rho > 1$  and  $m = 1$ . Now by our assumptions concerning  $\prod_{j=0}^{n-1} m_j$ , there are constants  $B_1 > 0$  and  $B_2 < \infty$  such that  $B_1 n^\rho L(n) \leq \prod_{j=0}^{n-1} m_j \leq B_2 n^\rho L(n)$



for  $n = 1, 2, \dots$ . From (3.2) and (3.3),

$$\begin{aligned} M_n(t) &= \sum_{k=n}^{\infty} \prod_{j=n}^{k-1} m_j (G_{k-n}(t) - G_{k-n+1}(t)) \\ &= \sum_{l=0}^n \frac{\prod_{j=0}^{l+n-1} m_j}{\prod_{j=0}^{n-1} m_j} (G_l(t) - G_{l+1}(t)) \\ &\quad + \sum_{l=n+1}^{\infty} \frac{\prod_{j=0}^{l+n-1} m_j}{\prod_{j=0}^{l-1} m_j \prod_{j=0}^{n-1} m_j} \prod_{j=0}^{l-1} m_j (G_l(t) - G_{l+1}(t)). \end{aligned}$$

Hence, for  $n = 1, 2, 3, \dots$

$$(4.9) \quad M_n(t) \leq \frac{B_2}{B_1} \left( \sum_{l=0}^n \left( \frac{l+n}{n} \right)^{\rho} \frac{L(l+n)}{L(n)} \right) (G_l(t) - G_{l+1}(t)) + \frac{1}{\prod_{j=0}^{n-1} m_j} \\ \times \left( \frac{B_2}{B_1} \sum_{l=n+1}^{\infty} \left( \frac{l+n}{l} \right)^{\rho} \frac{L(l+n)}{L(l)} \right) \prod_{j=0}^{l-1} m_j (G_l(t) - G_{l+1}(t)).$$

Since  $L$  is slowly varying and nondecreasing, there is a finite number  $B_3$  such that  $(t+s)^{\rho} L(t+s)/L(s) < B_3$  for  $0 \leq t \leq s$  and  $s \geq 1$ . This entails, after putting the appropriate additional nonnegative term into the second summation in (4.9),

$$M_n(t) \leq \frac{B_2 B_3}{B_1} (1 - G_{n+1}(t)) + \frac{1}{\prod_{j=0}^{n-1} m_j} \frac{B_2 B_3}{B_1} \sum_{l=0}^{\infty} \prod_{j=0}^{l-1} m_j (G_l(t) - G_{l+1}(t)).$$

But this means that, for all  $n \geq 1$  and all  $t \geq 0$ ,

$$(4.10) \quad M_n(t) \leq B_4 + \frac{B_5}{n^{\rho} L(n)} M_0(t),$$

where  $B_4 = B_2 B_3 / B_1$  and  $B_5 = B_4 / B_1$ . This follows using  $n = 0$  in (3.2) and (3.3). Let  $E(t, \tau)$  denote the series in (4.8). Define  $E_0(t, \tau)$ ,  $E_1(t, \tau)$ , and  $E_2(t, \tau)$  by

$$\begin{aligned} E(t, \tau) &= [E_0(t, \tau) + E_1(t, \tau) + E_2(t, \tau)]/h(t, \tau) \\ E_1(t, \tau) &= B_4 \sum_{k=0}^{\infty} \sigma_k \prod_{j=0}^{k-1} m_j (M_{k+1} * G_{k+1})(t) \\ &\quad + B_4 \sum_{k=0}^{\infty} \sigma_k \prod_{j=0}^{k-1} m_j \int_0^t M_{k+1}(t+u) dG_{k+1}(u) \\ E_2(t, \tau) &= \sum_{k=0}^{\infty} \sigma_k \prod_{j=0}^{k-1} m_j \int_0^t (M_{k+1}(t-u) - B_4)(M_{k+1}(t-u+\tau) - B_4) dG_{k+1}(u). \end{aligned}$$

Then, plainly using Theorem 5, page 423 and Theorem 2, page 421 of Feller (1966), one has  $E_0(t, \tau)/h(t, \tau) \rightarrow 0$  uniformly in  $\tau$  as  $t \rightarrow \infty$ . Upon plugging the right hand side of (4.10) into the definition of  $E_1(t, \tau)$ , realizing that the resulting quantity defines a measure, one may check that this resulting quantity when divided by  $h(t, \tau)$  approaches zero uniformly in  $\tau$  as  $t \rightarrow \infty$ , using the same theorems from Feller (1966), and Theorem 3.1. Now, by plugging the right hand side of (4.10) into the definition of  $E_2(t, \tau)/h(t, \tau)$ , we obtain a series whose terms are, by Theorem 3.1, dominated by terms of the form  $B_6/k^{\rho} L(k)$  where  $B_6$  is a positive finite number independent of  $t$  and  $\tau$ . Since  $\rho > 1$ , the Lebesgue dominated convergence theorem allows the evaluation of  $\lim_{t \rightarrow \infty} E_2(t, \tau)/h(t, \tau)$  termwise, giving (4.4) with  $m^* = m = 1$ .

**THEOREM 4.2.** *Under the hypotheses of Theorem 4.1, the quantity  $W(t) = Z_0(t)/M_0(t)$  converges in quadratic mean to a random variable  $W$  as  $t$  approaches infinity.*

**PROOF.** First we notice that

$$E(W(t + \tau) - W(t))^2 = \frac{C_0(t + \tau, 0)}{(M_0(t + \tau))^2} - \frac{2C_0(t, \tau)}{M_0(t)M_0(t + \tau)} + \frac{C_0(t, 0)}{(M_0(t))^2}$$

and the right hand side of this equality approaches zero uniformly in  $\tau$  as  $t$  approaches infinity, using Theorem 4.1. Theorem 4.2 now follows by completeness of  $L_2$ .

**THEOREM 4.3.** *Under the hypotheses of Theorem 4.1, if  $W$  is the random variable in Theorem 4.2, and if  $m^*$  is as in Theorem 4.1 then  $EW = 1$ ,  $\text{Var } W = \sum_{k=0}^{\infty} \sigma_k (\prod_{j=0}^{k-1} m_j) / [m^{*k+1} (\prod_{j=0}^k (m_j/m))^2] - 1$ , and  $\text{Var } W > 0$  if  $\text{Var } \zeta_{n,1} > 0$  for some  $n$ .*

**PROOF.**  $EW(t) - EW = E(W(t) - W)$ . The right hand side of this inequality approaches zero as  $t$  approaches infinity, by Theorem 4.2, since  $L_2$  convergence implies  $L_1$  convergence. This means that  $1 = EW(t) = EW$ . Also, by Minkowski's inequality,

$$(\text{Var } W(t))^{\frac{1}{2}} - (E(W(t) - W)^2)^{\frac{1}{2}} \leq (\text{Var } W)^{\frac{1}{2}} \leq (\text{Var } W(t))^{\frac{1}{2}} + (E(W(t) - W)^2)^{\frac{1}{2}}$$

and so, letting  $t$  approach infinity,

$$\text{Var } W = \lim_{t \rightarrow \infty} \text{Var } W(t) = \lim_{t \rightarrow \infty} \frac{C_0(t, 0)}{(M_0(t))^2} - 1.$$

The latter limit is the desired result for  $\text{Var } W$ , by Theorem 4.1. Now assume  $\text{Var } \zeta_{N,1} > 0$ .

$$\begin{aligned} E(W^2) &= \sum_{k=0}^{\infty} (\sigma_k \prod_{j=0}^{k-1} m_j) / [m^{*k+1} (\prod_{j=0}^k (m_j/m))^2] \\ &= \sigma + \sum_{k=0}^{\infty} [(1/\prod_{j=0}^{k-1} m_j) - (1/\prod_{j=0}^k m_j)] / (m^*/m^2)^{k+1} \end{aligned}$$

where  $\sigma = \sum_{k=0}^{\infty} [\text{Var}(\zeta_{k,1}) \prod_{j=0}^{k-1} m_j] / [m^{*k+1} \prod_{j=0}^k (m_j/m)^2]$ . Note that  $\sigma$  is positive since  $\text{Var}(\zeta_{N,1}) > 0$ . By Jensen's inequality (or Holder's inequality),  $m^*/m^2 = (\int_0^{\infty} e^{-\alpha t} dG(t))^2 / \int_0^{\infty} e^{-2\alpha t} dG(t) \leq 1$ . Thus,

$$1 + \text{Var } W = EW^2 \geq \sigma + \sum_{k=0}^{\infty} (1/\prod_{j=0}^{k-1} m_j) - (1/\prod_{j=0}^k m_j) = \sigma + 1,$$

proving that  $\text{Var } W$  is positive.

Notice that if  $m = m^* = 1$  in the expression for  $E(W^2)$ , then  $\text{Var}(W)$  is  $\sum_{k=0}^{\infty} \text{Var}(\zeta_{k,1}) / (m_k^2 \prod_{j=0}^{k-1} m_j)$ . This recapitulates the corresponding result in Theorem 1 of [2] (provided (48) of [2] is corrected to read " $\text{Var } W = \sum_{k=0}^{\infty} \sigma_k^2 / (m_k^2 \prod_{j=0}^{k-1} m_j)$ ") as in the left hand side of (49) in [2]. This expression after notational changes was also obtained in Theorem 4 of [6]. It is anticipated that by relating the branching process  $Z_0(t)$  to  $Z_n$ , the size of the  $n$ th generation of cells in the branching process  $Z_0(t)$ , one may show that if  $0 \leq \rho \leq 1$  and

$m = 1$ ,  $P[Z_0(t) \neq 0] \text{Var } W(t) \rightarrow \gamma > 0$  as  $t \rightarrow \infty$ . (This is the so-called critical case result.) This approach was successful in Goldstein (1971) and is feasible since it has been shown in [2] that under certain conditions with generation dependence,  $P[Z_n \neq 0] \text{Var}(Z_n/EZ_n) \rightarrow 2$  as  $n$  approaches infinity. Weiner's (1972) article contains a survey of results and methods for obtaining critical case results.

Finally, the following corollary to Theorems 4.1 and 4.3 establishes that the convergence in quadratic mean of  $W(t)$  to a random variable  $W$  is not vacuous under the hypotheses of Theorem 4.1.

**COROLLARY 4.1.** *Under the hypotheses of Theorem 4.1,  $P[Z_0(t) \rightarrow \infty \text{ as } t \rightarrow \infty] > 0$ .*

**PROOF.** Clearly, since  $M_0(t) \rightarrow \infty$  as  $t \rightarrow \infty$  under the hypotheses of Theorem 4.1,  $P[Z_0(t) \rightarrow \infty \text{ as } t \rightarrow \infty] \geq P[W > 0]$ . From Theorem 4.3,  $EW = 1$ , insuring that  $P[W > 0] > 0$ .

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DEPARTMENT OF STATISTICS  
CALIFORNIA STATE UNIVERSITY  
HAYWARD, CALIFORNIA 94542