

## THE EXPECTED RATIO OF THE SUM OF SQUARES TO THE SQUARE OF THE SUM<sup>1</sup>

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Let  $\{X_i, i = 1, 2, \dots\}$  be a sequence of positive i.i.d. random variables. Define  $S_n = \sum_{i=1}^n X_i$  and  $T_n = \sum_{i=1}^n X_i^2$ . We study the rate, if any, at which  $E[S_n^{-2}T_n] \rightarrow 0$ .

**1. Introduction.** Let  $X_1, X_2, \dots$  be non-negative independent random variables with common distribution function  $F$ . Assume  $F(0) < 1$  and let  $H(x) = 1 - F(x) = P(X_1 > x)$ . Define  $S_n = \sum_{i=1}^n X_i$  and  $T_n = \sum_{i=1}^n X_i^2$ . Let  $R_n = T_n S_n^{-2}$  if  $S_n > 0$ ,  $R_n = 1$  otherwise. It is shown in Theorem 4 that  $ER_n \rightarrow 0$  iff  $\int_0^\infty H(y) dy$  is slowly varying. The purpose of the rest of this paper is to study the rate at which  $ER_n \rightarrow 0$ .

The question which led us to this study was posed by Professor A. Joffe (private communication) in connection with his work (1978) on branching processes. He specifically sought moment conditions on  $X_1$  under which  $ER_n = O((\ln n)^{-1})$  as  $n \rightarrow \infty$ . The answer to this question is provided by Corollary 2; namely it is sufficient that  $E\{X_1 \ln X_1\} < \infty$  (where  $0 \ln 0 = 0$ ). Theorem 6 gives a further sufficient condition for  $ER_n = O\{(\ln n)^{-1}\}$ .

Note that the quantity  $R_n$  is related to one of the "self-normalized sums" studied by Logan, Mallows, Rice and Shepp (1973). Specifically,  $R_n = (S_n(2))^{-2}$ . Observe, however, that they assume  $EX_1 = 0$  when  $E|X_1| < \infty$ , which is incompatible with our non-negativity assumption. Cohn and Hall (1982) have also recently obtained some results on  $ER_n$  in the course of their study on weighted sums of random variables. For some readers, the main interest in our theorems will be in the methodology. The various techniques used below may also be useful in studying expectations of other random variables.

We use the following conventions throughout this paper. Let  $a_n$  and  $b_n$  be sequences of positive numbers. We write  $a_n = o(b_n)$  if  $a_n b_n^{-1} \rightarrow 0$  as  $n \rightarrow \infty$ ;  $a_n = O(b_n)$  if  $a_n b_n^{-1} \leq c$  for all  $n$  where  $c$  is a positive constant;  $a_n \sim b_n$  if  $a_n b_n^{-1} \rightarrow 1$  as  $n \rightarrow \infty$ ; and  $a_n \approx b_n$  if  $a_n = O(b_n)$  and  $b_n = O(a_n)$ . Also, the symbols  $C, C_1, C_2$ , etc. always denote positive constants. If we write  $a_n \leq C_i b_n$  for all  $n$ , we mean that there exists  $C_i > 0$  such that  $a_n \leq C_i b_n$  for all  $n$ , unless  $C_i$  has already been determined by earlier considerations. Finally, when we consider limits, we always mean as the variable goes to plus infinity.

By Jensen's inequality,  $n^{-1} \leq R_n \leq 1$  a.s. for all  $n$  so that  $n^{-1} \leq ER_n \leq 1$ . The exact behaviour of  $ER_n$  depends very much on the nature of  $H$  as the following heuristic argument suggests. Suppose  $H(x) = x^{-b}$  for  $x > 1$ , where  $b > 0$ . The random variables  $H(X_i)$  are uniformly distributed in  $(0, 1)$ . Suppose  $H(X_1), H(X_2), \dots, H(X_n)$  have the "typical" values  $n^{-1}, 2n^{-1}, \dots, 1$  in some order. Then a direct calculation gives

$$R_n = \sum_{k=1}^n k^{-2b-1} \left\{ \sum_{k=1}^n k^{-b-1} \right\}^{-2}.$$

The asymptotic behaviour (in the sense of  $\approx$ ) of these typical values of  $R_n$  is given in Table 1 for various values of  $b$ . Table 1 also gives the asymptotic behaviour of  $ER_n$  as determined later in the paper. Note that the latter entries are somewhat larger than the former when  $1 \leq b < 2$  but they are the same when  $b < 1$  or  $b \geq 2$ .

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TABLE 1  
*Asymptotic Behaviour of Typical  $R_n$  and  
of  $ER_n$  when  $H(x) = x^{-b}$ .*

<b>b</b>	<b>Typical <math>R_n</math></b>	<b><math>ER_n</math></b>
$b > 2$	$n^{-1}$	$n^{-1}$
$b = 2$	$n^{-1} \ln n$	$n^{-1} \ln n$
$1 < b < 2$	$n^{2b-1-2}$	$n^{1-b}$
$b = 1$	$(\ln n)^{-2}$	$(\ln n)^{-1}$
$0 < b < 1$	1	1

**2. The case of finite mean.** We begin with estimates which make explicit the relationship between  $ER_n$  and the rate of growth of  $S_n$ . Define  $M_n = \max(X_1, X_2, \dots, X_n)$ .

LEMMA 1. (a) *Suppose  $a_n$  is a sequence of positive reals such that*

$$(1) \quad P[S_n - M_n \leq a_n] = o\left(na_n^{-2} \int_0^{a_n} vH(v) \, dv\right)$$

Then

$$(2) \quad \begin{aligned} ER_n &\leq 2na_n \int_0^\infty v(v + a_n)^{-3}H(v) \, dv(1 + o(1)) \\ &\leq 6na_n^{-2} \int_0^{a_n} vH(v) \, dv(1 + o(1)). \end{aligned}$$

(b) *Suppose  $a_n$  is a sequence such that  $P(S_{n-1} > a_n) \rightarrow 0$ . Then*

$$(3) \quad \begin{aligned} ER_n &\geq 2na_n \int_0^\infty v(v + a_n)^{-3}H(v) \, dv(1 + o(1)) \\ &\geq \frac{1}{4} na_n^{-2} \int_0^{a_n} vH(v) \, dv(1 + o(1)). \end{aligned}$$

PROOF. First note that for  $a > 0$

$$\begin{aligned} a \int_a^\infty v^{-2}H(v) \, dv &\leq a \int_a^\infty v^{-2}H(a) \, dv = H(a) \\ &= 2a^{-2} \int_0^a vH(a) \, dv \leq 2a^{-2} \int_0^a vH(v) \, dv. \end{aligned}$$

Therefore

$$(4) \quad \begin{aligned} a \int_0^\infty v(v + a)^{-3}H(v) \, dv &\leq a^{-2} \int_0^a vH(v) \, dv + a \int_a^\infty v^{-2}H(v) \, dv \\ &\leq 3a^{-2} \int_0^a vH(v) \, dv. \end{aligned}$$

On the other hand,

$$\begin{aligned}
 (5) \quad a^{-2} \int_0^a vH(v) \, dv &\leq 8a \int_0^a v(v+a)^{-3}H(v) \, dv \\
 &\leq 8a \int_0^\infty v(v+a)^{-3}H(v) \, dv.
 \end{aligned}$$

The second inequalities in (2) and (3) follow from (4) and (5) respectively. Observe that

$$(6) \quad ER_n \leq E(R_n, S_n - M_n \geq a_n) + P(S_n - M_n < a_n).$$

Now,

$$\begin{aligned}
 (7) \quad E(R_n, S_n - M_n \geq a_n) &\leq \sum_{j=1}^n E(X_j^2(X_j + a_n)^{-2}) = nEX_1^2(X_1 + a_n)^{-2} \\
 &= 2n \int_0^1 tP(X_1(X_1 + a_n)^{-1} > t) \, dt \\
 &= 2n \int_0^1 tP(X_1 > (1-t)^{-1}a_nt) \, dt \\
 &= 2na_n \int_0^\infty v(v + a_n)^{-3}H(v) \, dv
 \end{aligned}$$

where we have substituted  $v = (1 - t)^{-1}a_nt$ . The first inequality in (2) now follows from (6), (7), (1) and (5). The first inequality in (3) is obtained by combining (7) with the following:

$$\begin{aligned}
 (8) \quad ER_n &= nE(X_n^2S_n^{-2}) \\
 &\geq nE(X_n^2S_n^{-2} | S_{n-1} \leq a_n)P(S_{n-1} \leq a_n) \\
 &\geq nE(X_1^2(X_1 + a_n)^{-2})(1 + o(1)). \quad \square
 \end{aligned}$$

The next result indicates the relationship between  $ER_n$  and  $H$  in cases when  $EX_1 < \infty$ . We note that the upper bound applies even when  $EX_1 = \infty$  but it is not in general tight in that case, as can be seen by supposing  $xH(x) \rightarrow \infty$ .

**THEOREM 1.** *For any  $H$ ,*

$$(9) \quad ER_n = O\left\{n^{-1} \int_0^n xH(x) \, dx\right\}.$$

If  $EX_1 < \infty$ ,

$$(10) \quad ER_n \approx n^{-1} \int_0^n xH(x) \, dx.$$

**PROOF.** Consider any  $a$  such that  $0 < a < \min(1, EX_1)$ . An estimate due to Chernoff (1952) shows that

$$(11) \quad P(S_{n-1} \leq an) = o(\rho^n)$$

where  $0 < \rho < 1$ . Therefore

$$\begin{aligned}
 (12) \quad P[S_n - M_n \leq an] &= P[\cup_i \{S_n - X_i \leq an\}] \\
 &\leq nP[S_{n-1} \leq an] = o(n\rho^n)
 \end{aligned}$$

so that (1) holds with  $a_n = an$ . By (2),

$$ER_n \leq 6a^{-2}n^{-1} \int_0^n vH(v) dv \{1 + o(1)\}$$

which proves (9). Similarly, if  $EX_1 < a < \infty$ , then  $P[S_n > an] \rightarrow 0$  by the law of large numbers. Applying (3) with  $a_n = an$ ,

$$\begin{aligned} ER_n &\geq 2an^2 \int_0^n v(v + an)^{-3}H(v) dv \{1 + o(1)\} \\ &\geq 2an^2(n + an)^{-3} \int_0^n vH(v) dv \{1 + o(1)\}, \end{aligned}$$

which proves (10).  $\square$

We note that the upper bound (9) can also be obtained by a truncation argument like that used later in Theorem 4. For the rest of this section we give some simplifications and improvements in Theorem 1. We begin by showing that if  $H$  satisfies a moderate smoothness condition then we may replace the integrals in (9) and (10) by quantities which do not involve integration.

**COROLLARY 1.** *Suppose there exist numbers  $a > 1$  and  $m > a^{-2}$  such that  $H(ax) \geq mH(x) > 0$  for all large  $x$ . Then  $ER_n = O\{nH(n)\}$  and if  $EX_1 < \infty$  then  $ER_n \approx nH(n)$ .*

**PROOF.** Since  $H$  is non-increasing, the hypotheses ensure that  $H$  is  $R - O$  varying as in Seneta (1976, pages 92-94). By his Theorem A.2,

$$(13) \quad nH(n) \approx n^{-1} \int_0^n yH(y) dy.$$

The results now follow from (9) and (10).  $\square$

The next corollary provides an answer to Joffe's original question as cited in the introduction. It shows in particular that  $ER_n = o\{(\ln n)^{-1}\}$  if  $E\{X_1 \ln X_1\} < \infty$ .

**COROLLARY 2.** *Let  $K$  be a function which is positive and non-decreasing for all large  $x$  and for which there exists constants  $a > 1$  and  $m > a^{-2}$  such that  $K(ax) \leq m^{-1}K(x)$  for all large  $x$ . If  $E\{K(X_1)\} < \infty$ , then  $ER_n = o\{n/K(n)\}$ .*

**PROOF.** Let  $C$  be sufficiently large that the conditions on  $K$  apply for all  $x \geq C$ . These conditions imply in particular that  $K(x) = o(x^2)$  so that

$$\int_C^\infty x\{K(x)\}^{-1} dx = \infty.$$

Also,

$$K(x)H(x) \leq \int_x^\infty K(y)F(dy) \rightarrow 0,$$

so that  $H(x) = o[\{K(x)\}^{-1}]$ . The last two statements together imply that

$$(14) \quad \int_0^n xH(x) dx = o\left\{\int_C^n x\{K(x)\}^{-1} dx\right\}.$$

The result now follows from (9) and (14) by applying (13) to  $\{K(x)\}^{-1}$ .  $\square$

If we now put further smoothness restrictions on  $H$  than we did in Corollary 1, we can obtain a more precise result. Specifically we assume  $X_1$  falls in the domain of attraction of a stable law with exponent  $b$  (notation:  $X_1 \in \mathcal{D}(b)$ ). For  $b < 2$ , this amounts to assuming that  $H(v)$  is regularly varying with exponent  $-b$ . Note that this implies  $H$  satisfies the hypotheses of Corollary 1. For  $b = 2$ ,  $X_1$  is in the domain of attraction of a normal law. A sufficient condition for  $X_1 \in \mathcal{D}(2)$  is of course that  $EX_1^2 < \infty$ . Background material for these notions may be found in Feller (1971, VIII.8).

**THEOREM 2.** *Suppose  $X_1 \in \mathcal{D}(b)$  and  $\mu = EX_1 < \infty$ . Then,*

$$(15) \quad ER_n \sim 2n^{-1}\mu^{-2} \int_0^n vH(v) dv \quad \text{if } b = 2$$

$$(16) \quad ER_n \sim \mu^{-b}\Gamma(2 - b)\Gamma(1 + b)nH(n) \quad \text{if } 1 \leq b < 2.$$

**PROOF.** We first assume  $1 \leq b < 2$ . Fix  $a < \mu$ , let  $a_n = an$ , and let  $\epsilon > 0$ . By Lemma 1 with the estimate (12) and by the substitution  $v = nu$ ,

$$(17) \quad ER_n \leq \left[ 2n^{-1}a^{-2} \int_0^{n\epsilon} vH(v) dv + 2anH(n) \int_\epsilon^\infty H(nu)\{H(n)\}^{-1}u(a + u)^{-3} du \right] (1 + o(1)).$$

By the regular variation of  $H$  and by dominated convergence, the second integral in (17) converges as  $n \rightarrow \infty$  to

$$(18) \quad \int_\epsilon^\infty u^{1-b}(a + u)^{-3} du \leq 0.5a^{-b-1}\Gamma(2 - b)\Gamma(1 + b).$$

By Theorem 1 on page 281 of Feller (1971), the first integral at (17) is

$$(19) \quad \int_0^{n\epsilon} vH(v) dv \sim \epsilon^{2-b}(2 - b)^{-1}n^2H(n).$$

Since  $\epsilon$  is arbitrary, (18) and (19) together give

$$(20) \quad ER_n \leq a^{-b}\Gamma(2 - b)\Gamma(1 + b)nH(n)\{1 + o(1)\}.$$

On the other hand, if  $EX_1 < a < \infty$  and  $\epsilon > 0$  is sufficiently small, similar calculations show that

$$(21) \quad \begin{aligned} ER_n &\geq 2\mu n^2 \int_{n\epsilon}^\infty v(v + na)^{-3}H(v) dv\{1 + o(1)\} \\ &\geq 2\mu nH(n) \int_\epsilon^\infty u^{1-b}(a + u)^{-3} du\{1 + o(1)\} \geq a^{-b-2}\mu^2 nH(n)\Gamma(2 - b)\Gamma(1 + b) \end{aligned}$$

for large  $n$ . The result follows by letting  $a$  increase to  $\mu$  in (20) and decrease to  $\mu$  in (21).

For the case  $b = 2$ , well known facts on domains of attraction (see Feller (1971), page 577) provide that  $x^2H(x)\{\int_0^x v^2F(dv)\}^{-1} \rightarrow 0$  as  $x \rightarrow \infty$ . We deduce that

$$(22) \quad nH(n) = o\left\{n^{-1} \int_0^n vH(v) dv\right\}.$$

As in (17) with  $a < \mu$  and  $\varepsilon = 1$ , we therefore obtain

$$\begin{aligned} ER_n &\leq \left[ 2a^{-2}n^{-1} \int_0^n vH(v) dv + 2an \int_1^\infty H(nu)u(a+u)^{-3} du \right] (1 + o(1)) \\ &= \left[ 2a^{-2}n^{-1} \int_0^n vH(v) dv + 2anO(H(n)) \right] (1 + o(1)) \\ &= 2a^{-2}n^{-1} \int_0^n vH(v) dv(1 + o(1)). \end{aligned}$$

By (22),  $\int_0^x vH(v) dv$  is slowly varying. For  $a > EX_1$  and  $\varepsilon > 0$ , we see from (3) that

$$\begin{aligned} ER_n &\geq 2n^2a \int_0^{n\varepsilon} \frac{v}{(\varepsilon n + an)^3} H(v) dv(1 + o(1)) \\ &\geq 2n^{-1}a(\varepsilon + a)^{-3} \int_0^n vH(v) dv(1 + o(1)). \end{aligned}$$

These bounds on  $ER_n$  lead to (15).  $\square$

**COROLLARY 3.**  $ER_n \approx n^{-1}$  if and only if  $EX_1^2 < \infty$ . In this case,  $ER_n \sim n^{-1}EX_1^2(EX_1)^{-2}$ .

**PROOF.** The “if” part and the second sentence follow from Theorem 2. If  $EX_1 < \infty$ , the “only if” part follows from (10). We delay the “only if” part with  $EX_1 = \infty$  until Theorem 5.  $\square$

We have given an integral free expression for the asymptotic behaviour of  $ER_n$  when the conditions of Corollary 1 hold. When  $EX_1^2 < \infty$ , we know  $ER_n \approx n^{-1}$ . Part of the gap between these two cases is filled by the following result.

**THEOREM 3.** Suppose  $H(x) = x^{-2}\tau(x)$  where  $\tau$  satisfies the conditions specified in Theorem 6 below. Then  $ER_n \approx n(\ln n)H(n)$ .

We do not prove this result, but only remark that the proof is similar to that of Theorem 6.

**3. The case when  $EX_1 = \infty$ .** It follows from Theorem 1 that if  $EX_1 < \infty$  then  $ER_n \rightarrow 0$ . We begin this section with a more complete result.

**THEOREM 4.**  $ER_n \rightarrow 0$  iff  $\int_0^x H(y) dy$  is slowly varying (or equivalently iff  $\int_0^x y dF(y)$  is slowly varying).

**PROOF.** If  $ER_n \rightarrow 0$ , then  $S_n^{-1}\{\max(X_1, X_2, \dots, X_n)\} \rightarrow 0$  in probability. It was shown by Breiman (1965) that this implies the slow variation of the integral. (This result is compatible with those of Logan et al (1973) and Darling (1952).)

Now suppose  $\int_0^x H(y) dy$  is slowly varying. It is shown by Feller (1971, pages 236–237) that there exists constants  $a_n > 0$  such that, for  $\varepsilon > 0$ ,

$$(23) \quad P(|a_n^{-1}S_n - 1| > \varepsilon) \leq n\varepsilon^{-2}a_n^{-2} \int_0^{a_n} x^2 dF(x) + nH(a_n) \rightarrow 0.$$

It follows that

$$\begin{aligned}
 (24) \quad ER_n &\leq 4a_n^{-2}E(T_n, \text{ all } X_i \leq a_n \text{ and } S_n \geq \frac{1}{2}a_n) \\
 &\quad + P(\text{some } X_i > a_n) + P(S_n < \frac{1}{2}a_n) \\
 &\leq 4a_n^{-2}n \int_0^{a_n} x^2 dF(x) + nH(a_n) + P(S_n < \frac{1}{2}a_n) \rightarrow 0. \quad \square
 \end{aligned}$$

The next result shows that if  $EX_1 = \infty$ ,  $ER_n$  cannot converge to zero at anything but a slow rate.

**THEOREM 5.** *Suppose  $EX_1 = \infty$ . There is a slowly varying sequence  $b_n$  such that  $\limsup_{n \rightarrow \infty} b_n ER_n = \infty$ .*

**PROOF.** By Theorem 4, we may assume  $\mu(x) \equiv \int_0^x y dF(y)$  is slowly varying. By Feller (1971, page 236), we may choose constants  $a_n$  such that

$$(25) \quad P(S_n \geq 2a_n) \rightarrow 0$$

and such that, as a function of  $n$ ,  $a_n$  is an inverse of the regularly varying function  $s(\mu(s))^{-1}$ . The result of Seneta (1976, page 21) shows that  $a_n$  is itself regularly varying with exponent 1. Using the hypothesis  $EX_1 = \infty$ , an integral comparison test yields

$$\sum_{n=1}^{\infty} (\ln n)^{-2} H(n(\ln n)^{-2}) = \infty$$

so that  $H(n(\ln n)^{-2}) > n^{-1}$  for infinitely many  $n$ . Thus

$$\begin{aligned}
 P(T_n > n^2(\ln n)^{-4}) &\geq P(\max(X_1, X_2, \dots, X_n) > n(\ln n)^{-2}) \\
 &= 1 - [F(n(\ln n)^{-2})]^n \not\rightarrow 0.
 \end{aligned}$$

Combining this with (25), we see that there exists  $C > 0$  such that

$$\begin{aligned}
 ER_n &\geq E(R_n, S_n < 2a_n, T_n > n^2(\ln n)^{-4}) \\
 &\geq (2a_n)^{-2}n^2(\ln n)^{-4}[P(T_n > n^2(\ln n)^{-4}) - P(S_n \geq 2a_n)] \\
 &\geq Cn^2a_n^{-2}(\ln n)^{-4}
 \end{aligned}$$

for infinitely many  $n$ . The result follows with  $b_n = a_n^2n^{-2}(\ln n)^5$ .  $\square$

We have shown that, if the integral in Theorem 4 is slowly varying and if  $EX_1 = \infty$ , then  $ER_n$  converges slowly to zero. In the next theorem, we put some smoothness conditions on  $H$  and then obtain a precise rate.

**THEOREM 6.** *Suppose  $H(x) = x^{-1}\tau(x) > 0$  for all  $x > 0$  where  $\tau$  satisfies*

$$(26) \quad m \leq \frac{\tau(x^\lambda)}{\tau(x)} \leq M$$

for all  $\lambda \in [1, a]$  and  $x \geq A$  where  $m, M, a$  and  $A$  are constants satisfying

$$0 < a^{-1} < m < 1 < M < \infty.$$

Then  $ER_n \approx (\ln n)^{-1}$ .

**PROOF.** First observe that  $K(x) \equiv \tau(e^x)$  is  $R - 0$  varying and the hypotheses of Theorem A.2 of Seneta (1976, page 94) are met. For  $y \geq x \geq \ln A$  we therefore have

$$(27) \quad m \left[ \frac{\ln x}{\ln y} \right]^\beta \tau(x) \leq \tau(y) \leq M \left[ \frac{\ln y}{\ln x} \right]^\alpha \tau(x)$$

where  $\alpha = (\ln M)(\ln a)^{-1}$  and  $\beta = -(\ln m)(\ln a)^{-1} < 1$ . It is obvious from (27) that

$$(28) \quad \tau(x) = O((\ln x)^\alpha)$$

and

$$(29) \quad (\ln x)^{-\beta} = O(\tau(x)).$$

By Seneta's theorem we also have

$$(30) \quad \int_A^x H(y) dy = \int_A^x y^{-1} \tau(y) dy = \int_{\ln A}^{\ln x} \tau(e^t) dt \approx \tau(x) \ln x.$$

For any  $q > 0$  and for  $0 < p < q(q + 1)^{-1}$ ,

$$\begin{aligned} \int_0^x y^q H(y) dy &\leq \int_0^{x^p} y^q dy + \int_{x^p}^x y^{q-1} \tau(y) dy \\ &\leq (q + 1)^{-1} x^{(q+1)p} + q^{-1} x^q \sup\{\tau(y) : x^p \leq y \leq x\} \\ &\leq q^{-1} x^q \tau(x) m^{-1} p^{-\beta} (1 + o(1)), \end{aligned}$$

where we have used (29) and (27). A similar lower bound shows that in fact for any  $q > 0$

$$(31) \quad \int_0^x y^q H(y) dy \approx x^q \tau(x).$$

Define  $b_n = n\tau(n) \ln n$ . Then

$$(32) \quad n(\ln n)^{\alpha+2} \geq b_n \geq C_1 n (\ln n)^{1-\beta}$$

for large  $n$ . Thus,

$$(33) \quad (\ln(\delta b_n))(\ln n)^{-1} \rightarrow 1$$

as  $n \rightarrow \infty$  for any  $\delta > 0$ . It follows from (27) that

$$(34) \quad \frac{\tau(\delta b_n)}{\tau(n)} \geq m \left[ \frac{\ln n}{\ln(\delta b_n)} \right]^\beta \geq \frac{m}{2}$$

for sufficiently large  $n$  and that

$$(35) \quad \tau(\delta b_n) \approx \tau(n).$$

Now define  $X'_i = X_i$  if  $X_i \leq b_n$ , 0 otherwise. Let  $S'_n = \sum_{i=1}^n X'_i$  and let

$$(36) \quad a_n = ES'_n = n \int_0^{b_n} x dF(x) = -nb_n H(b_n) + n \int_0^{b_n} H(x) dx.$$

The last integral  $n \int_0^{b_n} H(x) dx \approx n\tau(b_n) \ln b_n \approx n\tau(n) \ln n = b_n$  by (30), (33), and (35) and similarly  $nb_n H(b_n) \approx n\tau(n)$ . Thus,

$$(37) \quad a_n \approx b_n.$$

The next step is to estimate the rate of growth of  $S_n$  by a truncation argument and Chebyshev's inequality (see also Feller (1971), pages 236-237). For any  $\epsilon > 0$ ,

$$\begin{aligned} (38) \quad P[|S_n - a_n| > \epsilon a_n] &\leq P[|S'_n - a_n| > \epsilon a_n] + nH(b_n) \\ &\leq \epsilon^{-2} a_n^{-2} n \int_0^{b_n} x^2 dF(x) + nH(b_n) \\ &\leq 2\epsilon^{-2} a_n^{-2} n \int_0^{b_n} xH(x) dx + nH(b_n) \\ &\leq C_2 (a_n^{-2} nb_n \tau(b_n)) + nb_n^{-1} \tau(b_n) \\ &\leq C_3 nb_n^{-1} \tau(n) = C_3 (\ln n)^{-1}, \end{aligned}$$

for  $n$  sufficiently large, where we have used (31), (37), and (35). A similar truncation at  $b_n$  combined with the argument at (24) yields the upper bound



(39)  $ER_n = O(\ln n)^{-1}$ .

Next, choose  $\delta \in \left(0, \frac{m}{4C_3}\right)$  where  $C_3$  is given by (38) with  $\varepsilon = 1$ . By (35),

(40)  $nH(\delta b_n) = \frac{n}{\delta n \tau(n) \ln n} \tau(\delta b_n) \rightarrow 0$

as  $n \rightarrow \infty$ . On the other hand, by (34),

(41)  $nH(\delta b_n) \geq \frac{m}{2\delta \ln n}$

for large  $n$ . By Bonferroni's inequality, (40) and (41),

(42) 
$$P[\max(X_1, X_2, \dots, X_n) > \delta b_n] \geq nH(\delta b_n) - \binom{n}{2} (H(\delta b_n))^2$$

$$\geq \frac{1}{2} nH(\delta b_n) \geq \frac{m}{4\delta \ln n}$$

for large  $n$ . By (38) with  $\varepsilon = 1$ , (42) and (37),

(43) 
$$ER_n \geq \frac{\delta^2 b_n^2}{4a_n^2} P[T_n > \delta^2 b_n^2, S_n \leq 2a_n] \geq C_4 \{P[T_n > \delta^2 b_n^2] - P[S_n > 2a_n]\}$$

$$\geq C_4 \{P[\max(X_1, \dots, X_n) > \delta b_n] - C_3(\ln n)^{-1}\}$$

$$\geq C_4 \left\{ \frac{m}{4\delta \ln n} - \frac{C_3}{\ln n} \right\} = C_4 \left( \frac{m}{4\delta} - C_3 \right) (\ln n)^{-1}$$

for sufficiently large  $n$ . The theorem now follows from (39) and (43).  $\square$

REMARK. It is obvious that if  $EX_1 < \infty$  then  $\liminf xH(x)\ln x = 0$ . We see from (29) that this is not true here; thus the hypotheses of the theorem imply  $EX_1 = \infty$ . By (27) the hypotheses also imply that  $\int_0^\infty H(y) dy$  is slowly varying.

We note that (38) gives us the following rate of convergence result for a generalized weak law of large numbers.

COROLLARY 4. Let  $H$  satisfy the hypotheses of Theorem 6 and define  $a_n$  by (36). For  $\varepsilon > 0$ ,

$$P[|a_n^{-1}S_n - 1| > \varepsilon] = O((\ln n)^{-1}).$$

**4. Further ramifications and examples.** Theorem 2 implies  $nER_n \rightarrow 1 + \alpha^{-1}$  when  $X_1$  has a Gamma  $(\alpha, 1)$  distribution. Simulations (when  $\alpha = 0.5$ ) indicate that the convergence occurs very rapidly. We also simulated the value of  $R_n$  for variables having densities of the type  $f(x) = c(1+x)^{-b-1}$  for  $x > 0$ ,  $f(x) = 0$  otherwise. For example, when  $b = 1.5$ , the asymptotic behaviour should be given by (16), but graphs indicated a highly erratic behaviour. This results from the fact that  $R_n/ER_n \rightarrow 0$  in probability in this case. In other words, the asymptotic behaviour of samples (and hence sample means) of  $R_n$  differs profoundly from that of  $ER_n$ . This can be shown whenever  $H(x)$  is regularly varying with exponent  $-b$ ,  $1 < b < 2$ , by first demonstrating that

(44)  $nH(d_n^{1/2}) \rightarrow 0$

where  $d_n = n^3H(n) \approx n^2ER_n \approx S_n^2ER_n$  and then verifying by truncating each  $X_i$  at  $d_n^{1/2}$  that (44) implies  $d_n^{-1}T_n \rightarrow 0$  in probability. Thus, under these conditions, the rate of convergence apparent in (16) represents for a single sequence (or average of a fixed number of sequences) an empirically unverifiable phenomenon.

Many of the foregoing results do not depend heavily on the independence of the random variables  $X_1, X_2, \dots$ . Assume now that the random variables are identically distributed but

possibly dependent. Take  $a_n = an$  in Theorem 1. In order to draw the conclusion  $ER_n = O\{n^{-1} \int_0^n vH(v) dv\}$  we may replace  $o(\cdot)$  in the condition (1) of Lemma 1 by  $O(\cdot)$ . It is sufficient, then, that  $P\{S_n < a_n\} = O(n^{-1})$ . This may be seen to hold under a variety of dependence assumptions using only Chebyshev's inequality, for example when the sequence of random variables is pairwise independent, or when second moments are finite and the random variables are orthogonal. Similarly, the condition of Lemma 1(b) may be replaced in the case  $a_n = an$ ,  $a > EX_1$  by the condition  $\limsup P\{S_n > an\} < 1$ . This is also an easy consequence of Chebyshev's inequality in many cases.

Consider the class of examples obtained by letting  $H(x) = x^{-b}(\ln x)^r$  for large  $x$  where  $b \geq 0$  and  $-\infty < r < \infty$  (with  $r < 0$  if  $b = 0$ ). If  $b > 2$  or if  $b = 2$  and  $r < -1$ , then  $EX_1^2 < \infty$  and  $ER_n \sim n^{-1}\mu^{-2}EX_1^2$  by Corollary 3. By Theorem 1,  $ER_n \approx n^{-1}(\ln \ln n)$  if  $b = 2$  and  $r = -1$  and  $ER_n \approx n^{-1}(\ln n)^{r+1}$  if  $b = 2$  and  $r > -1$ . If  $1 < b < 2$  or if  $r < -1$  and  $b = 1$ , then  $ER_n \approx n^{1-b}(\ln n)^r$ . If  $b = 1$  and  $r = -1$ , we obtain the answer from the proof of Theorem 6. Let  $b_n = n \ln \ln n$  and calculate directly that  $a_n \approx b_n$ . Then deduce that  $ER_n \approx \{(\ln n)(\ln \ln n)\}^{-1}$ . If  $b = 1$  and  $r > -1$ , Theorem 6 implies that  $ER_n \approx (\ln n)^{-1}$ . Finally, if  $b < 1$ , it can be shown that  $ER_n \approx 1$  by the results of Logan et al (1973) and Darling (1952).

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#### REFERENCES

- [1] BREIMAN, L. (1965). On some limit theorems similar to the arc-sin law. *Theor. Probability Appl.* **10** 323-331.
- [2] CHERNOFF, H. (1952). A measure of asymptotic efficiency for tests of a hypothesis based on the sum of observations. *Ann. Math. Statist.* **23** 493-507.
- [3] COHN, H. and HALL, P. (1982). On the limit behaviour of weighted sums of random variables. *Z. Wahrsch. verw. Gebiete.* **53** 319-332.
- [4] DARLING, D. A. (1952). The influence of the maximum term in the addition of independent random variables. *Trans. of the AMS* **73** 95-107.
- [5] FELLER, W. (1971). *An Introduction to Probability Theory and its Applications.* 2. (second ed.) Wiley, New York.
- [6] JOFFE, A. (1978). Remarks on the structure of trees with applications to supercritical Galton-Watson Processes. *Adv. in Prob. and Related Topics* **5** 263-268. Edited by A. Joffe and P. Ney, Dekker, N.Y.
- [7] LOGAN, B. F., MALLOWS, C. L., RICE, S. O., and SHEPP, L. A. (1973). Limit distributions of self-normalized sums. *Ann. Probability* **1** 788-809.
- [8] SENETA, E. (1976). *Regularly Varying Functions.* Springer, Berlin.

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