

## ON THE MAXIMUM OF A MEASURE OF DEVIATION FROM INDEPENDENCE BETWEEN DISCRETE RANDOM VARIABLES

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The squared  $n^k$ -dimensional Euclidean distance  $f_k$  between a given joint distribution of  $k$  random variables with values in  $1, \dots, n$  and the joint distribution of independent variables with the same respective marginals has been suggested as a measure of dependence. The following facts are established for  $M_k$ , the maximum of  $f_k$  over all joint distributions for fixed  $k$ : (1)  $M_k$  is attained among the distributions with all  $k$  variables equal to a variable  $X$  that takes on just two values. (2) For  $k \leq 6$ ,  $M_k = \frac{1}{2} - (\frac{1}{2})^k$  is attained when the distribution of  $X$  is  $\{\frac{1}{2}, \frac{1}{2}\}$ . (3) For  $k \geq 7$ ,  $M_k$  is not attained at  $\{\frac{1}{2}, \frac{1}{2}\}$  and strictly exceeds  $\frac{1}{2} - (\frac{1}{2})^k$ . (4) For  $k \rightarrow \infty$ , the distributions of  $X$  where  $M_k$  is attained approach  $\{0, 1\}$ , and  $M_k \nearrow 1$ .

**1. Introduction.** Let  $X_1, \dots, X_k$  ( $k \geq 2$ ) be discrete random variables ranging over the finite sets  $C_1, \dots, C_k$  respectively, where, without loss of generality, we assume

$$C_j = \{1, 2, \dots, n_j\}, \quad j = 1, \dots, k, \quad n_j \geq 2.$$

Let

$$P(i, j) = P(X_j = i), \quad j = 1, \dots, k, \quad i = 1, \dots, n_j,$$

$$P(i_1, \dots, i_k) = P(X_1 = i_1, X_2 = i_2, \dots, X_k = i_k),$$

and let  $\mathbf{P} = (p(i_1, \dots, i_k))$ .

Consider the function

$$f_k(\mathbf{P}) = \sum_{i_1, \dots, i_k} [P(i_1, \dots, i_k) - \prod_{j=1}^k P(i_j, j)]^2.$$

If  $\mathbf{P}$  is regarded for each joint distribution as a vector in Euclidean  $(n_1 n_2 \dots n_k)$ -dimensional space, then  $f_k(\mathbf{P})$  is the squared distance between  $\mathbf{P}$  and the corresponding independent distribution, and as such,  $f_k(\mathbf{P})$  is a measure of deviation from independence between  $X_1, \dots, X_k$ . This function arises in the literature regarding the establishment of a measure of complete dependence between discrete random variables (Cramér 1924, Goodman and Kruskal 1954, Lancaster 1963, and Gilula 1981). The function  $f_k(\mathbf{P})$  vanishes if and only if  $X_1, \dots, X_k$  are independent.

In order to interpret the value that  $f_k$  attains for a given  $k$ -dimensional joint distribution, one naturally would like to compare it with the maximal value that  $f_k$  can attain for that  $k$ . Ideally, a measure of complete dependence should attain its maximum if and only if the variables are completely dependent. Cramér

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(1924), who considered the bivariate case ( $k = 2$ ), showed that  $\text{Max } f_2 = 1/4$ , and the maximum is attained when the two variables are two-valued, completely dependent, and take on their values with probability  $1/2$  each. Otherwise,  $f_2$  is strictly less than  $1/4$  (even for completely dependent variables).

We study here the maximum of  $f_k$  for general  $k$ . First, we reduce the search for a maximum to the case where all  $k$  variables are equal to one variable, say  $x$ , that takes on two values only, with probabilities, say  $p$  and  $1 - p$ . Denote  $f_k(\mathbf{P})$  for such a distribution by  $f_k(p)$ . Next, we show that when  $k \leq 6$ ,  $f_k$  is unimodal, attaining the maximum  $1/2 - (1/2)^k$  at  $p = 1/2$ ; when  $k > 6$ , the maximum is attained elsewhere, and strictly exceeds  $1/2 - (1/2)^k$ . For  $k$  approaching infinity, the pairs  $\{p_k, 1 - p_k\}$  where the maximum of  $f_k$  is attained approach  $\{0, 1\}$ , and the maximum itself approaches 1.

Summarizing, we obtain the following picture. Cramér's result, that for  $k = 2$  the maximum of  $f_k$  is attained for completely dependent variables that take on two values with probability  $1/2$  each, is valid also for  $k = 3, 4, 5$  and  $6$ . Beyond  $k = 6$ , a surprising change occurs: the maximum is still attained for completely dependent two-valued variables, but the probabilities are no longer  $(1/2, 1/2)$ . Another result is that for large  $k$  the maximum is attained close to 0 and to 1, while at  $(0, 1)$   $f_k$  vanishes. This reflects the fact that degenerate variables, being both independent and completely dependent, form a singularity of the complete dependence concept (compare with Kimeldorf and Sampson, 1978). The difficulty is therefore inherent in the concept of complete dependence rather than in the particular measure chosen to quantify it.

**2. Upper Bounds on  $f_k(\mathbf{P})$ .** By Cramér (1924), in maximizing  $f_k(\mathbf{P})$  attention can be confined to the case where all  $k$  variables considered are equal so that all sets  $C_j$  are the same, namely  $C_j = \{1, \dots, n\}$   $j = 1, \dots, k$ ,  $n \geq 2$  and

$$P(i_1, \dots, i_k) = \begin{cases} p_i & \text{if } i_1 = i_2 = \dots = i_k = i, \quad (i = 1, \dots, n) \\ 0 & \text{otherwise,} \end{cases}$$

where also  $P(X_j = i) = p_i$ ,  $j = 1, \dots, k$ ,  $i = 1, \dots, n$ ,  $\sum_{i=1}^n p_i = 1$ .

Let  $\mathbf{p} = (p_1, \dots, p_n)$ . Then

$$f_k(\mathbf{p}) = \sum_{i=1}^n p_i^2 - 2 \sum_{i=1}^n p_i^{k+1} + (\sum_{i=1}^n p_i^2)^k.$$

We will now show that in maximizing  $f_k(\mathbf{p})$ , attention can be further confined to the case where sets  $C_j$  contain only two elements, or equivalently, where the  $k$  variables considered are all Bernoulli variables.

$f_k(\mathbf{p})$  is continuous in the compact simplex  $\sum_{i=1}^n p_i = 1$ ,  $p_i \geq 0$  and therefore attains its maximum there. Let  $\mathbf{q}$  be an interior point, of the simplex, different from its centroid  $(1/n, \dots, 1/n)$ . Assuming  $n \geq 3$ , there exists a vector  $\mathbf{u}$ , orthogonal to  $(1, \dots, 1)$  and to  $\mathbf{q}$ , but not to  $(q_1^k, \dots, q_n^k)$ .

\* As is easily seen, the gradient of  $f_k(q)$  at  $q$  is a linear combination of  $\mathbf{q}$  and  $(q_1^k, \dots, q_n^k)$  with nonzero coefficients. Therefore, the directional derivative of  $f_k(\mathbf{p})$  at  $\mathbf{q}$  in direction  $\mathbf{u}$  will not be zero, and the maximum is not attained there.

The centroid is also ruled out as a maximum point, since  $f_k(0, 1/(n - 1), \dots, 1/(n - 1))$  is easily shown to exceed  $f_k(1/n, \dots, 1/n)$ .

We must conclude then that for  $\mathbf{p}$  to be a maximum point for  $f_k$  in the simplex  $\sum_{i=1}^n p_i = 1, n \geq 3$ , it must be a boundary point, namely at least one coordinate of  $\mathbf{p}$  must be zero. By induction we obtain the desired reduction to the case where  $n = 2$ .

Attention is now focused, therefore, on the function

$$f_k(p) = p^2 + (1 - p)^2 + [p^2 + (1 - p)^2]^k - 2p^{k+1} - 2(1 - p)^{k+1}.$$

This function is a polynomial in  $p$ , that is symmetric around  $p = 1/2$ . Therefore it is a polynomial in  $X = (p - 1/2)^2$ . As  $X$  increases from 0 to  $1/4$ ,  $f_k(p)$  goes through its range. Carrying out the substitution yields

$$f_k(p) = g_k((p - 1/2)^2),$$

$$g_k(X) = 1/2 + 2X + \sum_{i=0}^k 2^{2i-k} \binom{k}{i} X^i - \sum_{j=0}^{(k+1)/2} 2^{2j-k+1} \binom{k+1}{2j} X^j.$$

Let  $g'_k, g''_k, g'''_k$  denote the first, second, and third derivative of  $g_k(X)$ . While  $f_k$  is stationary at  $p = 1/2$ , this is not the case for  $g_k$  at the corresponding  $X = 0$ . Indeed,  $g'_k(0)$ , the linear coefficient, is  $2(1 - (k^2/2^{k-1}))$ . This is negative for  $k = 2, 3, 4, 5, 6$  and positive for  $k \geq 7$ . So  $g_k$  is increasing at 0 if  $k \geq 7$ , and decreasing if  $2 \leq k \leq 6$ . Clearly,  $f_k$  does not attain its maximum at  $p = 1/2$  when  $k \geq 7$ . To establish that, if  $2 \leq k \leq 6$ ,  $g_k$  and  $f_k$  do attain their maxima at  $X = 0$  and  $p = 1/2$  respectively, consider  $g'''_k$ . It is a polynomial whose coefficients will all be positive if this is the case for the coefficients of order  $\geq 3$  of  $g_k$ .

Since the expression for  $g_k$  involves minus signs only for terms of order  $\leq [(k + 1)/2]$ , the only coefficients of  $g_k$  of order  $\geq 3$  that could be negative, when  $k \leq 6$ , are the third order coefficients of  $g_5$  and of  $g_6$ . Since they turn out to be positive,  $g'''_k$  is positive on  $0 < X$  for  $2 \leq k \leq 6$ .

By  $g'''_k > 0$ ,  $g''_k$  can at most have a simple zero on the positive axis, and  $g'_k$  at most two, when  $2 \leq k \leq 6$ .

At  $p = 0$ ,  $f_k$  and its derivative vanish. By symmetry, this holds at  $p = 1$  as well. This is a regular point of  $X = (p - 1/2)^2$ , hence  $g_k$  and  $g'_k$  vanish at  $X = 1/4$ . For  $0 \leq k \leq 6$ ,  $g'_k$  can vanish only once more on the positive axis. But since  $g'_k(0)$  is negative,  $g_k$  cannot start up again to a maximum, and then decrease to zero, without incurring two points of zero derivative. Hence there is no maximum, not even a local one, except for  $X = 0$ , where  $g_k(X) = 1/2 - (1/2)^k$  for  $k \leq 6$ . We conclude that  $f_k(p)$  is unimodal on  $[0, 1]$  and its maximum is  $1/2 - (1/2)^k, 2 \leq k \leq 6$ .

For  $k \geq 7$ , no explicit forms for the location and value of  $\text{Max } f_k$  were obtained; however, a study of the limiting behavior of  $f_k$ , and some numerical results serve to complete the picture.

Note that for  $0 < p < 1, \lim f_k(p) = p^2 + (1 - p)^2$ , and  $\max_{0 \leq p \leq 1} p^2 + (1 - p)^2 = 1$  is attained at  $p = 0$  or  $p = 1$ . Clearly, therefore, 1 is the least upper bound of  $f_k(p)$ , over all  $p$  and  $k$ . In fact, we also have  $\lim_{k \rightarrow \infty} \text{Max}_{0 \leq p \leq 1} f_k(p) = 1$ . This follows from the existence of sequences  $(\epsilon_k)$  such that  $\lim_{k \rightarrow \infty} f_k(\epsilon_k) = 1$ : choose

any sequence  $(\varepsilon_k)$  such that

$$\varepsilon_k \rightarrow 0 \text{ but } k\varepsilon_k \rightarrow \infty$$

(for example,  $\varepsilon_k = k^{-1/2}$ ). Such a choice of  $\varepsilon_k$  implies, as  $k \rightarrow \infty$ , (a)  $\varepsilon_k^2 \rightarrow 0$ , (b)  $(1 - \varepsilon_k)^2 \rightarrow 1$ , (c)  $\varepsilon_k^{k+1} \rightarrow 0$ , (d)  $(1 - \varepsilon_k)^{k+1} \rightarrow 0$ , (e)  $[(1 - \varepsilon_k)^2 + \varepsilon_k^2]^k \rightarrow 0$ . While results (a), (b), (c) are immediate, (d) and (e) follow since

$$(1 - c\varepsilon_k)^k = [(1 - c\varepsilon_k)^{1/\varepsilon_k}]^{k\varepsilon_k},$$

but  $(1 - c\varepsilon_k)^{1/\varepsilon_k} \rightarrow e^{-c}$  and  $(e^{-c})^{k\varepsilon_k} \rightarrow 0$ .

Numerical calculations of  $f_k$  seemed to indicate that  $f_k$  is bimodal, and its larger mode  $p_k$  is of the form  $k^{-1/(ak-b)}$ . When  $(\log k)/(\log p_k)$  was plotted against  $k$  for  $k = 7, \dots, 25$  the line  $k - 3$  gives a very close fit. Thus  $p_k = k^{-1/(k-3)}$  and  $q_k = 1 - k^{-1/(k-3)}$  are empirical formulas for the modes of the  $f_k$  for  $k \geq 7$ .

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### REFERENCES

- [1] CRAMÉR, H. (1924). Remarks on correlation. *Scand. Actuar. J.* **16** 220-240.
- [2] GILULA, Z. (1981). A note on measuring the degree of complete dependence between two discrete random variables measured on a nominal scale. *Comm. Statist. Theory Methods* **21** 2047-2055.
- [3] GOODMAN, L. A. and KRUSKAL, W. H. (1954). Measures of association for cross-classifications. *J. Amer. Statist. Assoc.* **49** 732-764.
- [4] KIMELDORF, G. and SAMPSON, A. R. (1978). Monotone dependence. *Ann. Statist.* **6** 895-903.
- [5] LANCASTER, H. O. (1963). Correlation and complete dependence of random variables. *Ann. Math. Statist.* **34** 1315-1321.

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