# A NOTE ON TRUNCATION AND SUFFICIENT STATISTICS1

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1. Introduction and summary. Generalizing earlier observations by Fisher and Hotelling, Tukey [1] showed that if a family of distributions admits a set of sufficient statistics, then the family obtained by truncation to a fixed set, or by a fixed selection, also admits the same set of sufficient statistics (this wording is Tukey's; we give a precise mathematical statement later). Tukey's proof assumed the relevant family of probability measures to be dominated by a fixed measure function and made use of the factorization theorem concerning sufficient statistics in this case. In the present short note we shall first re-prove Tukey's result without assuming domination (and, hence, without appealing to the factorization theorem). Then we shall show that, under general conditions, if a sufficient statistic has one or more of the properties of completeness, bounded completeness, or minimality, before truncation, then it preserves such after truncation.

The treatment is on the lines of the abstract discussion of sufficient statistics given by Halmos and Savage [2]. We shall assume familiarity with the results given in this latter paper. For definitions of completeness, bounded completeness, and minimality, and for a discussion of the significance of these concepts we refer to Lehmann and Scheffé [3].

2. On  $\phi$ -truncation. Let X be an abstract space of elements x, and let  $\mathfrak{F}_x$  be a (Borel) field of subsets of X. We write  $\{\mu_{\theta} : \theta \in \Omega\}$  for a family of probability measures on  $(X, \mathfrak{F}_x)$ , where  $\Omega$  is an abstract parameter space. The statistic t is a mapping of X onto another abstract space T, that is to say, we suppose for simplicity, with no loss of generality, that T is precisely the range  $\{t(x) : x \in X\}$  of the mapping t. If  $B \subseteq T$ , we write  $t^{-1}B = \{x : t(x) \in B\}$  for the origin of B. The class of all  $B \subseteq T$  such that  $t^{-1}B \in \mathfrak{F}_x$  is written  $\mathfrak{F}_t$ ; it is easy to show that  $\mathfrak{F}_t$  is a (Borel) field.

We shall write  $\mathcal{E}_{\theta}$  for expectation based on  $\mu_{\theta}$ . If f(x) is any  $(\mathfrak{F}_{x})$ -measurable function such that  $\mathcal{E}_{\theta}|f(x)|<\infty$ , for all  $\theta$  in some set  $\Lambda\subseteq\Omega$ , we shall say f(x) is  $\Lambda$ -integrable. If f(x) is  $\Omega$ -integrable, the conditional expectation  $\mathcal{E}_{\theta}(f(x)\mid t)$  is given by the Radon-Nikodym derivative

(1) 
$$\mathcal{E}_{\theta}(f(x) \mid t) = \frac{d\nu_{\theta} t^{-1}}{d\mu_{\theta} t^{-1}},$$

where the measure  $\nu_{\theta}$  is defined by

$$(2) d\nu_{\theta} = f(x) d\mu_{\theta}.$$

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For a proof of this assertion, see [2]. Note that the derivative in (1) is arbitrary on a set of  $(\mu_{\theta}t^{-1})$ -measure zero. If it transpires that for each  $\Omega$ -integrable f(x) the conditional expectation  $\mathcal{E}_{\theta}(f(x) \mid t)$  may be taken as independent of  $\theta \in \Omega$ , then t is sufficient for  $\{\mu_{\theta} : \theta \in \Omega\}$ .

Suppose that  $\phi(x)$  is a non-negative  $\Omega$ -integrable function; define  $\Omega_{\phi} = \{\theta : \mathcal{E}_{\theta}\phi(x) > 0\}$ ; and define a new family  $\{\mu_{\theta}^{\phi} ; \theta \in \Omega_{\phi}\}$  of probability measures on  $(X, \mathfrak{F}_{x})$  by the equation

(3) 
$$d\mu_{\theta}^{\phi} = \frac{\phi(x)}{\mathcal{E}_{\theta} \phi(x)} d\mu_{\theta} .$$

We call  $\{\mu_{\theta}^{\phi}; \theta \in \Omega_{\phi}\}$  the  $\phi$ -truncation of  $\{\mu_{\theta}; \theta \in \Omega\}$ . When  $\phi(x)$  is the characteristic function of some set  $A \in \mathfrak{F}_x$ , then  $\phi$ -truncation corresponds to truncation to a fixed set, in the usual sense. When  $\phi(x)$  is bounded above it is easy to see that it may be assumed to be bounded above by unity, and then  $\phi$ -truncation corresponds to what Tukey [1] has called fixed selection. However,  $\phi(x)$  may not be bounded above.

We sometimes write, for brevity,  $\{\mu_{\theta}\}$  for  $\{\mu_{\theta}; \theta \in \Omega\}$ ;  $\{\mu_{\theta}^{\phi}\}$  for  $\{\mu_{\theta}^{\phi}; \theta \in \Omega_{\phi}\}$ ; and  $\{\mu_{\theta}\}_{\phi}$  for  $\{\mu_{\theta}; \theta \in \Omega_{\phi}\}$ .

We shall also write  $\mathcal{E}_{\theta}^{\phi}$  for expectations based on  $\mu_{\theta}^{\phi}$ ; and if f(x) is any  $(\mathfrak{F}_{x})$ -measurable function such that  $\mathcal{E}_{\theta}^{\phi}|f(x)| < \infty$ , for all  $\theta$  in some set  $\Lambda \subseteq \Omega$ , we shall say f(x) is  $\Lambda^{\phi}$ -integrable. If a statement is followed by an expression like  $[\mathfrak{M}]$ , where  $\mathfrak{M}$  represents a family of probability measures, this will mean that the statement is false, at most, on a set of probability zero for each measure of  $\mathfrak{M}$ . In this connection let us notice that, since  $\{\mu_{\theta}\}_{\phi}$  dominates  $\{\mu_{\theta}^{\phi}\}_{\phi}$ , we can always replace  $[\{\mu_{\theta}\}_{\phi}]$  by  $[\{\mu_{\theta}^{\phi}t^{-1}\}_{\phi}]$  by  $[\{\mu_{\theta}^{\phi}t^{-1}\}_{\phi}]$ .

Lemma. If f(x) is  $\Omega_{\theta}^{\phi}$ -integrable, then  $\phi(x)f(x)$  is  $\Omega_{\phi}$ -integrable.

PROOF. If  $\theta \in \Omega_{\phi}$ , and we define  $X_{\phi} = \{x : \phi(x) > 0\}$ ,

$$\mathcal{E}_{\theta} \mid \phi(x)f(x) \mid = \int_{\mathcal{I}_{\phi}} \phi(x) \mid f(x) \mid d\mu_{\theta}$$

$$= \{\mathcal{E}_{\theta} \phi(x)\} \int_{\mathcal{I}_{\phi}} |f(x)| \mid d\mu_{\theta}^{\phi}$$

$$= \{\mathcal{E}_{\theta} \phi(x)\} \{\mathcal{E}_{\theta}^{\phi} \mid f(x)|\}$$

$$< \infty.$$

THEOREM. (i) If t is sufficient for  $\{\mu_{\theta} ; \theta \in \Omega\}$ , then t is sufficient for  $\{\mu_{\theta}^{\phi} ; \theta \in \Omega_{\phi}\}$ ; (ii) if, in addition, t is complete, minimal, for  $\{\mu_{\theta} ; \theta \in \Omega_{\phi}\}$ , then t is complete, minimal, for  $\{\mu_{\theta}^{\phi} ; \theta \in \Omega_{\phi}\}$ ; and a similar remark applies if t is boundedly complete—provided, in this case, that  $\phi(x)$  is bounded above.

(Notice that we require t to be complete, etc., for the subfamily  $\{\mu_{\theta} ; \theta \in \Omega_{\phi}\}$ . It is possible that t be complete, etc., for  $\{\mu_{\theta} ; \theta \in \Omega\}$  and not so for  $\{\mu_{\theta} ; \theta \in \Omega_{\phi}\}$ .) Proof. We observe first that by (1), (2), and (3), for  $\theta \in \Omega_{\phi}$ ,

(4) 
$$\frac{d\mu_{\theta}^{\phi}t^{-1}}{du_{\theta}t^{-1}} = \frac{\varepsilon(\phi(x) \mid t)}{\varepsilon_{\theta}(x)},$$

where we may omit the suffix  $\theta$  in  $\mathcal{E}_{\theta}(\phi(x) \mid t)$  because t is sufficient for  $\{\mu_{\theta} ; \theta \in \Omega\}$ .

Let f(x) be any  $\Omega_{\phi}^{\phi}$ -integrable function. Then by the lemma,  $\phi(x)f(x)$  is  $\Omega_{\phi}$ -integrable, and by the sufficiency of t for  $\{\mu_{\theta} ; \theta \in \Omega\}$ , and hence for  $\{\mu_{\theta} ; \theta \in \Omega_{\phi} \subseteq \Omega\}$ , we may write  $\mathcal{E}\{f(x)\phi(x) \mid t\}$  independent of  $\theta \in \Omega_{\phi}$ . Thus for any  $A \in \mathcal{F}_t$  we have for all  $\theta \in \Omega_{\phi}$ ,

$$\int_{A} \mathcal{E}(f(x)\phi(x) \mid t) \ d\mu_{\theta} t^{-1} = \int_{t^{-1}A} f(x)\phi(x) \ d\mu_{\theta}$$

$$= \{\mathcal{E}_{\theta}\phi(x)\} \int_{t^{-1}A} f(x) \ d\mu_{\theta}^{\phi}, \text{ by } (2.3),$$

$$= \{\mathcal{E}_{\theta}\phi(x)\} \int_{A} \mathcal{E}_{\theta}^{\phi}(f(x) \mid t) \ d\mu_{\theta}^{\phi} t^{-1},$$

$$= \int_{A} \mathcal{E}_{\theta}^{\phi}(f(x) \mid t) \mathcal{E}(\phi(x) \mid t) \ d\mu_{\theta} t^{-1},$$

by (4). Since the last equation holds for every A  $\varepsilon$   $\mathfrak{F}_t$ , we deduce from the Radon-Nikodym theorem that

The function  $\phi(x)$  is non-negative on X, from which it follows that  $\mathcal{E}(\phi(x) \mid t) \geq 0$ ,  $[\{\mu_{\theta}t^{-1}\}_{\phi}]$ . Write  $T_{\phi} = \{t : \mathcal{E}(\phi(x) \mid t) > 0\}$ ; then for  $t \in T - T_{\phi}$ , we have  $\mathcal{E}(\phi(x) \mid t) = 0$ ,  $[\{\mu_{\theta}t^{-1}\}_{\phi}]$ . Thus, if  $\theta \in \Omega_{\phi}$ ,

$$\int_{T-T_{\phi}} d\mu_{\theta}^{\phi} t^{-1} = \int_{T-T_{\phi}} \frac{d\mu_{\theta}^{\phi} t^{-1}}{d\mu_{\theta} t^{-1}} d\mu_{\theta} t^{-1},$$

$$= \int_{T-T_{\phi}} \frac{\mathcal{E}(\phi(x) \mid t)}{\mathcal{E}_{\theta} \theta(x)} d\mu_{\theta} t^{-1}, \quad \text{by (4)},$$

$$= 0.$$

We have therefore shown that  $\mathcal{E}(\phi(x) \mid t) > 0$ ,  $[\{\mu_{\theta}^{\phi}t^{-1}\}]$ , and may deduce from (5) that

$$\mathcal{E}^{\phi}_{\theta}(f(x) \mid t) = \frac{\mathcal{E}(f(x)\phi(x) \mid t)}{\mathcal{E}(\phi(x) \mid t)}, \qquad [\{\mu^{\phi}_{\theta}t^{-1}\}].$$

Hence for any  $\Omega_{\phi}^{\phi}$ -integrable function f(x) there exists a version of  $\mathcal{E}_{\theta}^{\phi}(f(x) \mid t)$  which is independent of  $\theta \in \Omega_{\phi}$ . This is enough to prove that t is sufficient for  $\{\mu_{\theta}^{\phi} : \theta \in \Omega_{\phi}\}$ .

Next suppose that t is a complete sufficient statistic for  $\{\mu_{\theta} ; \theta \in \Omega_{\phi}\}$ . To prove that t is complete for  $\{\mu_{\theta}^{\phi} ; \theta \in \Omega_{\phi}\}$ , we must show that if  $\psi(t)$  is an arbitrary  $\mathfrak{F}_{t}$ -measurable function such that  $\mathfrak{E}_{\theta}^{\phi}\psi(t(x)) = 0$  for all  $\theta \in \Omega_{\phi}$ , then  $\psi(t) = 0$ ,  $[\{\mu_{\theta}^{\phi}t^{-1}\}]$ . However, if

$$\int_{T} \psi(t) \ d\mu_{\theta}^{\phi} t^{-1} = 0, \qquad \text{all } \theta \in \Omega_{\phi} ,$$

it follows from (4) that

$$\int_{T} \psi(t) \mathcal{E}(\phi(x) \mid t) \ d\mu_{\theta} t^{-1} = 0, \qquad \text{all } \theta \in \Omega_{\phi} \ .$$

But t is complete for  $\{\mu_{\theta} ; \theta \in \Omega_{\phi}\}$ , and  $\psi(t)\mathcal{E}(\phi(x) \mid t)$  is an  $(\mathfrak{F}_{t})$ -measurable function of t. Hence

$$\psi(t)\mathcal{E}(\phi(x) \mid t) = 0, \quad [\{\mu_{\theta}t^{-1}\}_{\phi}].$$

Since we have already seen that  $\mathcal{E}(\phi(x) \mid t) > 0$ ,  $[\{\mu_{\theta}i^{-1}\}_{\phi}]$ , the proof that t is complete for  $\{\mu_{\theta}^{\phi}; \theta \in \Omega_{\phi}\}$  is thus evident.

When t is boundedly complete we can employ precisely the same argument, assuming both  $\psi(t)$  and  $\phi(x)$  to be bounded so as to ensure, as is easily checked, that  $\psi(t)\mathcal{E}(\phi(x)\mid t)$  is bounded  $[\{\mu_{\theta}t^{-1}\}_{\phi}]$ .

Lastly we deal with the minimality question. Suppose that s(x) is any statistic defined on X which is sufficient for  $\{\mu_{\theta}^{\phi} : \theta \in \Omega_{\phi}\}$ . Write  $S = \{s(x) : x \in X\}$  for the abstract space on which s maps X. Then to prove that t is minimal we must demonstrate the existence of a mapping h of S on T such that t(x) = h(s(x)),  $[\{\mu_{\theta}^{\phi}t^{-1}\}]$ .

Recall  $X_{\phi} = \{x : \phi(x) > 0\}$ , and notice that since  $\theta \in \Omega_{\phi}$  implies  $\mathcal{E}_{\theta}\phi(x) > 0$ , it also implies that  $\mu_{\theta}(X_{\phi}) > 0$ . Plainly,  $\mu_{\theta}^{\phi}(X - X_{\phi}) = 0$  for all  $\theta \in \Omega_{\phi}$ ; thus it will be enough to prove the relation t(x) = h(s(x)),  $[\{\mu_{\theta}^{\phi}t^{-1}\}]$ , merely on  $X_{\phi}$ . To this end, let us define a new statistic

$$s_1(x) = s(x)$$
 if  $x \in X_{\phi}$ ,  
=  $x$  if  $x \in X - X_{\phi}$ .

Since s(x) is obviously a one-valued function of  $s_1(x)$ , it follows that  $s_1(x)$  is also a sufficient statistic for  $\{\mu_{\theta}^{\phi} : \theta \in \Omega_{\phi}\}$  (by Theorem 6.4 of Bahadur [4]). We write  $S_1 = \{s_1(x) : x \in X\}$  and  $\mathfrak{F}_{s_1}$  for the (Borel) field of all subsets  $B \subseteq S_1$  such that  $s_1^{-1}B \in \mathfrak{F}_x$ .

If we set

$$\psi(x) = \frac{1}{\phi(x)}, \qquad x \in X_{\phi},$$

$$= 0, \qquad x \in X - X_{\phi}.$$

then  $\psi(x)$  is a non-negative  $(\mathfrak{F}_x)$ -measurable function on X; and for  $\theta \in \Omega_{\phi}$ ,

(6) 
$$\begin{split} \mathcal{E}_{\theta}^{\phi}\psi(x) &= \int_{\mathbf{x}_{\phi}} \frac{1}{\phi(x)} d\mu_{\theta}^{\phi} ,\\ &= \int_{\mathbf{x}_{\phi}} \frac{1}{\phi(x)} \frac{\phi(x)}{\xi_{\theta} \phi(x)} d\mu_{\theta} , \quad \text{by (3)},\\ &= \frac{\mu_{\theta} (X_{0})}{\xi_{\theta} \phi(x)} < \infty . \end{split}$$

It follows from (6) that  $\psi(x)$  is  $\Omega_{\phi}^{\phi}$ -integrable, and so we may consider the  $\psi$ -truncation of  $\{\mu_{\theta}^{\phi} : \theta \in \Omega_{\phi}\}$ . We shall employ obvious extensions of our notation, and observe that by (6),  $\Omega_{\phi\psi}^{\phi} = \{\theta : \theta \in \Omega_{\phi} , \mathcal{E}_{\theta}^{\phi}\psi(x) > 0\} = \Omega_{\phi}$ . Hence, for all  $\theta \in \Omega_{\phi}$ ,

(7) 
$$d\mu_{\theta}^{\phi\psi} = 0, \qquad \text{if } x \in X - X_{\phi},$$

$$= \frac{\psi(x)}{\varepsilon_{\theta}^{\phi}\psi(x)} d\mu_{\theta}^{\phi},$$

$$= \frac{\varepsilon_{\theta}\phi(x)}{\phi(x)\mu_{\theta}(X_{\phi})} \frac{\phi(x)}{\varepsilon_{\theta}\phi(x)} d\mu_{\theta}, \qquad \text{by (3) and (6),}$$

$$= \frac{d\mu_{\theta}}{\mu_{\theta}(X_{\phi})}, \qquad \text{if } x \in X_{\phi}.$$

Because  $s_1$  is sufficient for  $\{\mu_{\theta}^{\phi} : \theta \in \Omega_{\phi}\}$ , and because  $\Omega_{\phi\psi}^{\phi} = \Omega_{\phi}$ , it follows from the first part of our theorem that  $s_1$  is also sufficient for the  $\psi$ -truncation  $\{\mu_{\theta}^{\phi\psi} : \theta \in \Omega_{\phi}\}$ . Let f(x) be any  $\Omega_{\phi}$ -integrable function. Then we notice that for  $\theta \in \Omega_{\phi}$ , by (7) and (8),

$$\mathcal{E}_{\theta}^{\phi\psi} |f(x)| = \frac{1}{\mu_{\theta}(X_{\phi})} \int_{X_{\phi}} |f(x)| d\mu_{\theta} < \infty,$$

i.e., f(x) is  $\Omega_{\theta}^{\phi \psi}$ -integrable; and by the sufficiency of  $s_1$  there must exist a function  $g(s_1) = \mathcal{E}^{\phi \psi}(f(x) \mid s_1)$  which is independent of  $\theta \in \Omega_{\phi}$ , is such that  $g(s_1(x))$  is  $(\mathfrak{F}_x)$ -measurable, and is such that for any  $B \in \mathfrak{F}_{s_1}$  and all  $\theta \in \Omega_{\phi}$ ,

$$\int_{s_1^{-1}B} f(x) \ d\mu_{\theta}^{\phi\psi} = \int_{s_1^{-1}B} g(s_1(x)) \ d\mu_{\theta}^{\phi\psi}.$$

This implies, by (7), that

$$\int_{X_{A} \cap s_{1}^{-1}B} f(x) \ d\mu_{\theta}^{\phi \psi} = \int_{X_{A} \cap s_{1}^{-1}B} g(s_{1}(x)) \ d\mu_{\theta}^{\phi \psi},$$

and so, by (8), that

(9) 
$$\int_{X_{\theta} \cap s_1^{-1}B} f(x) \ d\mu_{\theta} = \int_{X_{\theta} \cap s_1^{-1}B} g(s_1(x)) \ d\mu_{\theta} .$$

Finally, define an  $(\mathfrak{F}_x)$ -measurable function

$$g^*(s_1(x)) = g(s_1(x))$$
 if  $s_1 \in S$ , i.e., if  $x \in X_{\phi}$   
=  $g(s_1(x))$  (=  $f(x)$ ) if  $s_1 \in X - X_{\phi}$ , i.e., if  $x \in X - X_{\phi}$ 

Thus

$$\begin{split} \int_{s_1^{-1}B} f(x) \ d\mu_{\theta} &= \int_{\mathbf{X}_{\phi} \cap s_1^{-1}B} f(x) \ d\mu_{\theta} + \int_{(\mathbf{X} - \mathbf{X}_{\phi}) \cap s_1^{-1}B} f(x) \ d\mu_{\theta} \\ &= \int_{\mathbf{X}_{\phi} \cap s_1^{-1}B} g(s_1(x)) \ d\mu_{\theta} + \int_{(\mathbf{X} - \mathbf{X}_{\phi}) \cap s_1^{-1}B} g^*(s_1(x)) \ d\mu_{\theta} \\ &= \int_{s_1^{-1}B} g^*(s_1(x)) \ d\mu_{\theta} \,, \end{split}$$

the last equality following from (9). Since f(x) is an arbitrary  $\Omega_{\phi}$ -integrable function, this last equation, being true for all  $B \in \mathfrak{F}_{s_1}$ , shows that  $s_1$  is a sufficient statistic for  $\{\mu_{\theta} : \theta \in \Omega_{\phi}\}$ . But we are given that t is minimal sufficient for  $\{\mu_{\theta} : \theta \in \Omega_{\phi}\}$ . Hence there is a mapping h of  $S_1$  onto T such that  $t(x) = h(s_1(x))$ ,  $[\{\mu_{\theta} t^{-1}\}_{\phi}]$ . If we now restrict x to  $X_{\phi}$  it is evident that t(x) = h(s(x)),  $[\{\mu_{\theta} t^{-1}\}_{\phi}]$ , as was to be proved.

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# A CENTRAL LIMIT THEOREM FOR MULTILINEAR STOCHASTIC PROCESSES

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1. Introduction. Let the random sequence X(t) be observed for  $t = 1, 2, \dots$ , and let  $S(n) = X(1) + \dots + X(n)$  be its consecutive sums. The random sequence may be said to obey the *classical* central limit theorem if, for any real number a,

(1.1) 
$$\lim_{n\to\infty} \operatorname{Prob}\left\{\frac{S(n)-ES(n)}{\sigma[S(n)]} < a\right\} = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{a} e^{-1/2x^2} dx.$$

Because of the importance of the central limit theorem in establishing the properties of statistical tests and estimates, it would appear that in order to develop a satisfactory theory of statistical inference for stochastic processes which are random sequences of dependent random variables, it is necessary to establish a central limit theorem for such processes. Diananda [2] has proved a central limit theorem for discrete parameter stochastic processes which are linear processes. We here introduce a class of stochastic processes which we call multilinear processes, for which we prove a central limit theorem. The results are capable of extension to the continuous parameter case, but we do not do so here.