

# ASYMPTOTIC DISTRIBUTION OF DISTANCES BETWEEN ORDER STATISTICS FROM BIVARIATE POPULATIONS

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**1. Summary.** The exact and limiting distribution of quantiles in the univariate case is well known. Mood [3] investigated the joint distribution of medians in samples from a multivariate population, showing that their distribution is asymptotically multivariate normal. Recently Siddiqui [4] considered the joint distribution of two quantiles and an auxiliary statistic and showed that asymptotically their joint distribution is trivariate normal. Further, he showed the "distances"  $X'_{i+l} - X'_i$ ,  $X'_i - X'_{i-h}$ , ( $l$  and  $h$  fixed positive integers) between quantiles in the univariate case, when appropriately normalized are asymptotically independently distributed as Chi square r.v.'s with  $2l$  and  $2h$  d.f. respectively. In this paper the joint distribution of several quantiles from a bivariate population is obtained and it is shown that the distances between quantiles in the separate component populations are independent asymptotically.

**2. Assumptions and notation.** Let  $F(x, y)$  be the absolutely continuous d.f. of the pair of random variables (r.v.'s)  $(X, Y)$ , having joint p.d.f.  $f(x, y)$  and marginal d.f.'s and p.d.f.'s  $F_1(x)$ ,  $F_2(y)$ ,  $f_1(x)$ , and  $f_2(y)$  respectively. Let  $\zeta_\alpha$ ,  $\eta_\beta$  be the unique real numbers satisfying  $F_1(\zeta_\alpha) = \alpha$ ,  $F_2(\eta_\beta) = \beta$ ,  $0 < \alpha, \beta < 1$ , with  $f_1(\zeta_\alpha) \neq 0$ ,  $f_2(\eta_\beta) \neq 0$ , and put  $q_1 = F(\zeta_\alpha, \eta_\beta)$ ,  $q_2 = \beta - q_1$ ,  $q_3 = \alpha - q_1$ ,  $q_4 = 1 - \alpha - \beta + q_1$ . Then  $\zeta_\alpha$ ,  $\eta_\beta$  are quantiles of orders  $\alpha$  and  $\beta$ , respectively, of  $F_1$  and  $F_2$ . We assume that  $F(x, y)$  has first and second partial derivatives continuous in a neighborhood of  $(\zeta_\alpha, \eta_\beta)$ . Let  $(X_i, Y_i)$ ,  $i = 1, \dots, n$ , be a random sample drawn from  $F(x, y)$ , and let  $Z_1^{(n)} \leq Z_2^{(n)} \leq \dots \leq Z_n^{(n)}$  be the ordered sample values of  $X_1, X_2, \dots, X_n$ . Similarly, let  $W_1^{(n)} \leq W_2^{(n)} \leq \dots \leq W_n^{(n)}$  be the ordered values of  $Y_1, \dots, Y_n$ . Finally,  $\{r^{(n)}\}$ ,  $\{s^{(n)}\}$  will denote sequences of positive integers depending on the sample size  $n$  in such a way that  $r^{(n)}/n \rightarrow \alpha$ ,  $s^{(n)}/n \rightarrow \beta$ . We shall first find an asymptotic approximation to the joint p.d.f. of the  $k + l + 2$  r.v.'s  $Z_{r^{(n)}+i}^{(n)}$ ,  $W_{s^{(n)}+j}^{(n)}$ ,  $i = 0, 1, \dots, k$ ;  $j = 0, 1, \dots, l$ , where  $k$  and  $l$  are fixed positive integers, and then show that the "distances"

$$(1) \quad d_{r^{(n)}+i,1}^{(n)} = Z_{r^{(n)}+i+1}^{(n)} - Z_{r^{(n)}+i}^{(n)}, \quad i = 0, 1, \dots, k$$

$$(2) \quad d'_{s^{(n)}+j,1} = W_{s^{(n)}+j+1}^{(n)} - W_{s^{(n)}+j}^{(n)}, \quad j = 0, 1, \dots, l$$

are asymptotically independent r.v.'s. Unless otherwise specified, the range on

Received 31 May 1963; revised 8 November 1963.

<sup>1</sup> This paper is part of a Ph.D. dissertation at The Pennsylvania State University.

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$i$  and  $j$  will henceforth be assumed as given above. For simplicity we shall omit the superscript in  $r^{(n)}$  and  $s^{(n)}$  in what follows.

We shall use the following notation:

$$\begin{aligned}
 \rho_1 &= F(x_0, y_0), & \rho_2 &= F_2(y_0) - F(x_k, y_0), & \rho_3 &= F_1(x_0) - F(x_0, y_l), \\
 (3) \quad \rho_4 &= 1 - F_1(x_k) - F_2(y_l) + F(x_k, y_l), & \rho_i^{(1)} &= \int_{-\infty}^{y_0} f(x_i, v) dv, \\
 \sigma_j^{(1)} &= \int_{-\infty}^{x_0} f(u, y_j) du, & \rho_i^{(2)} &= \int_{y_l}^{\infty} f(x_i, v) dv, & \sigma_j^{(2)} &= \int_{x_k}^{\infty} f(u, y_j) du.
 \end{aligned}$$

**3. Asymptotic joint distribution of  $(d_{r^{(n)}+i,1}^{(n)}, d_{s^{(n)}+j,1}^{(n)})$ .** First we shall derive the joint density function of  $Z_{r+i}^{(n)}, W_{s+j}^{(n)}, i = 0, 1, \dots, k; j = 0, 1, \dots, l$ . We find the probability  $P(A)$  of the event

$$\begin{aligned}
 A = \{x_i - \frac{1}{2}\Delta x_i \leq Z_{r+i}^{(n)} \leq x_i + \frac{1}{2}\Delta x_i; y_j - \frac{1}{2}\Delta y_j \leq W_{s+j}^{(n)} \leq y_j + \frac{1}{2}\Delta y_j, \\
 i = 0, 1, \dots, k; j = 0, 1, \dots, l\}.
 \end{aligned}$$

Divide the whole plane into mutually disjoint rectangles  $R_\mu (\mu = 1, 2, 3, 4), R_i^{(1)} (i = 0, 1, \dots, k), S_j^{(1)} (j = 0, 1, \dots, l),$  etc., as shown in Figure 1.

We consider the disjoint events  $A_1, A_2$  where

$$A_1 = \{\text{at least one of } (Z_{r+i}^{(n)}, W_{s+j}^{(n)}) = (X_\alpha, Y_\alpha) \text{ for } \alpha = 1, 2, \dots, n\}$$

$$A_2 = \{\text{no } Z_{r+i}^{(n)} \text{ and no } W_{s+j}^{(n)} \text{ are components of the same random vector}\}.$$

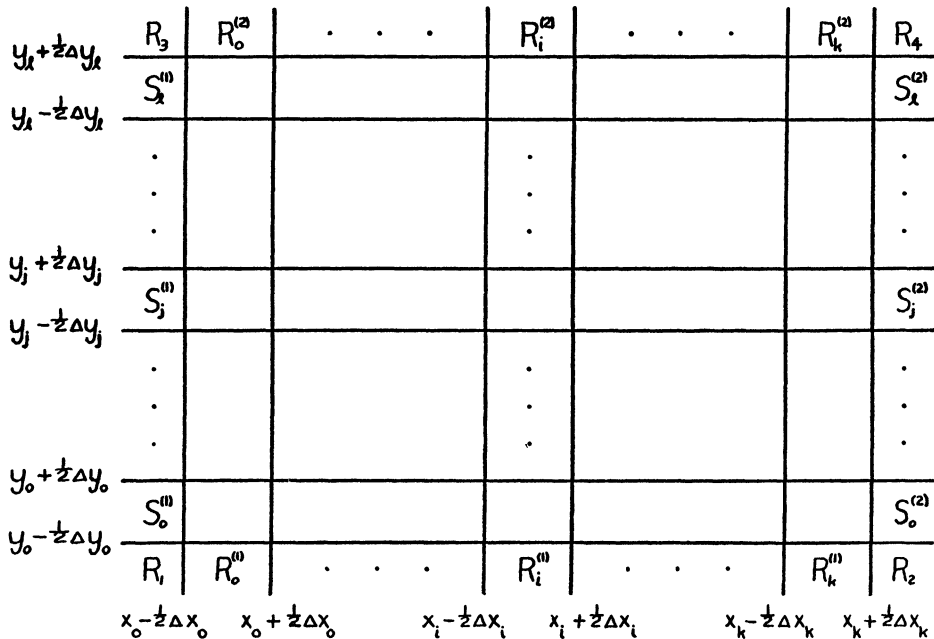


FIG. 1

Clearly,  $P(A) = P(A \cap A_1) + P(A \cap A_2)$ . It can be easily shown that  $P(A \cap A_1) = P(A \cap A_2)k_n(x, y)/n$ , where  $k_n(x, y)$  tends to a finite limit as  $n$  tends to infinity. As such

$$(4) \quad P(A) = P(A \cap A_2)[1 + O(n^{-1})].$$

In order that the event  $A \cap A_2$  be satisfied, an examination of the situation reveals that points  $V_h = (X_h, Y_h)$  can only fall in  $R_\mu (\mu = 1, 2, 3, 4)$ , along with one point in each of  $R_i^{(\nu)}, S_j^{(\nu)}$ ;  $\nu = 1$  or  $2, i = 0, 1, \dots, k$  and  $j = 0, 1, \dots, l$ . The distribution of  $V_h$  in the plane will be as follows:

Let  $n_\mu$  be the number of  $V_h$ 's in  $R_\mu (\mu = 1, 2, 3, 4)$ ; then  $\sum_{\mu=1}^4 n_\mu = n - (k + l + 2)$ . The remaining  $(k + l + 2)V_h$ 's fall in  $R_i^{(\nu)}$  and  $S_j^{(\nu)}$  in all possible  $2^{k+l+2}$  different ways arising from  $i = 0, 1, \dots, k; j = 0, 1, \dots, l$  and  $\nu = 1$  or  $2$ . Thus

$$(5) \quad P(A \cap A_2) = \sum'_{\nu_i, \nu_j} \sum_{n_1=0}^{r-\delta} \left( \prod_{\mu=1}^4 \frac{n!}{n_\mu!} P[R_\mu^{(n_\mu)}] \right) \prod_{j=0}^l \prod_{i=0}^k P[R_i^{(\nu_i)}] P[S_j^{(\nu_j)}]$$

where

$$\sum_{\mu=1}^4 n_\mu = n - (k + l + 2),$$

$$(6) \quad n_1 + n_2 = s - 1, s - 2, \dots, \text{ or } s - k - 1,$$

$$n_1 + n_3 = r - 1, r - 2, \dots, \text{ or } r - l - 1, \quad \delta = 1, 2, \dots, l + 1$$

and the prime on summation means that the summation is to be carried out through all  $2^{k+l+2}$  different combinations of  $\nu_i = 1, 2$  and  $\nu_j = 1, 2$ .

Henceforth, whenever the expression (5) appears, it will be subject to the restrictions given in (6). Thus from (4) and (5), we get

$$(7) \quad P(A) = [1 + O(n^{-1})] \sum'_{\nu_i, \nu_j} \sum_{n_1=0}^{r-\delta} \left( \prod_{\mu=1}^4 \frac{n!}{n_\mu!} P[R_\mu^{(n_\mu)}] \right) \prod_{j=0}^l \prod_{i=0}^k P[R_i^{(\nu_i)}] P[S_j^{(\nu_j)}].$$

Dividing by  $\prod_{i=0}^k \prod_{j=0}^l \Delta x_i \Delta y_j$ , and taking the limit as  $\Delta x_i \rightarrow 0, \Delta y_j \rightarrow 0$ , we obtain the joint probability density function  $g^{(n)}(\mathbf{x}, \mathbf{y})$  of  $Z_{r+i}^{(n)}, W_{s+j}^{(n)}$ . Now for  $\mu = 1, 2, 3, 4; i = 0, 1, \dots, k; j = 0, 1, \dots, l$ ; and  $\nu = 1, 2, \lim P[R_\mu] = \rho_\mu, \lim P[R_i^{(\nu)}] = \rho_i^{(\nu)}$ , and  $\lim P[S_j^{(\nu)}] = \sigma_j^{(\nu)}$  where all the limits are taken as  $\Delta x_i \rightarrow 0$  and  $\Delta y_j \rightarrow 0$ . Hence from (7), we get

$$(8) \quad g^{(n)}(\mathbf{x}, \mathbf{y}) = \sum'_{\nu_i, \nu_j} \sum_{n_1=0}^{r-\delta} \left( \prod_{\mu=1}^4 \frac{n!}{n_\mu!} \rho_\mu^{n_\mu} \right) \prod_{j=0}^l \prod_{i=0}^k \rho_i^{(\nu_i)} \sigma_j^{(\nu_j)} [1 + O(n^{-1})],$$

$-\infty < x_0 < \infty, x_i < x_{i+1} < \infty, i = 0, 1, \dots, k - 1, -\infty < y_0 < \infty, y_j < y_{j+1} < \infty, j = 0, 1, \dots, l - 1$  where  $\mathbf{x} = (x_1, \dots, x_k)$  and  $\mathbf{y} = (y_1, \dots, y_l)$ . Let

$$(9) \quad c_n = \rho_1 + \rho_2 + \rho_3 + \rho_4 = 1 - [F_1(x_k) - F_1(x_0)] - [F_2(y_l) - F_2(y_0)] \\ + [F(x_k, y_l) - F(x_k, y_0) - F(x_0, y_l) + F(x_0, y_0)]$$

and  $\theta_\mu = \rho_\mu/c_n$ , so that  $\sum_{\mu=1}^4 \theta_\mu = 1$ .

We now consider the normalized r.v.'s  $U_i = nd_{r+i,1}^{(n)}$ ,  $V_j = nd'_{s+j,1}^{(n)}$ , ( $i = 1, \dots, k; j = 1, \dots, l$ ),  $T_1 = n^{\frac{1}{2}}(Z_r^{(n)} - \zeta)$ , and  $T_2 = n^{\frac{1}{2}}(W_s^{(n)} - \eta)$ . On applying the transformation  $u_i = n(x_i - x_{i-1})$ ,  $v_j = n(y_j - y_{j-1})$ ,  $t_1 = n^{\frac{1}{2}}(x_0 - \zeta)$ , and  $t_2 = n^{\frac{1}{2}}(y_0 - \eta)$  to (10) [with Jacobian  $n^{-(k+l+1)}$ ], we find that the asymptotic joint density function of  $U_1, \dots, U_k, V_1, \dots, V_l, T_1, T_2$  is given by

$$(10) \quad \overline{h}^{(n)}(\mathbf{u}, \mathbf{v}, t_1, t_2) = A_1^{(n)} A_2^{(n)} A_3^{(n)} [1 + O(n^{-1})]$$

where

$$(11) \quad A_1^{(n)} = \prod_{\mu=1}^4 C_n^{n_\mu} = C_n^{n-(k+l+2)}$$

$$(12) \quad A_2^{(n)} = \sum'_{v_i, v_j} \prod_{j=0}^l \prod_{i=0}^k \rho_i^{(v_i)} \sigma_j^{(v_j)}$$

$$(13) \quad A_3^{(n)} = n \sum_{n_1=0}^{r-\delta} \prod_{\mu=1}^4 \{[n - (k + l + 2)]! / n_\mu!\} \theta_\mu^{n_\mu},$$

with  $0 \leq u_i < \infty$  ( $i = 1, \dots, k$ ),  $0 \leq v_j < \infty$  ( $j = 1, \dots, l$ ),  $-\infty < t_1, t_2 < \infty$ ,  $\mathbf{u} = (u_1, \dots, u_{k-1})$ ,  $\mathbf{v} = (v_1, \dots, v_{l-1})$ ; and  $x_0 = \zeta + n^{-\frac{1}{2}}t_1$ ,  $y_0 = \eta + n^{-\frac{1}{2}}t_2$ . (Note that  $n^{-(k+l+2)}n! \sim [n - (k + l + 2)]!$ , for fixed integers  $k$  and  $l$ , as  $n \rightarrow \infty$ ).

To find the joint asymptotic density of  $Z_r^{(n)}$ ,  $W_s^{(n)}$ ,  $d_{r+i,1}^{(n)}$  and  $d'_{s+j,1}^{(n)}$ , we consider

$$C_n = \left\{ 1 - \left[ F_1 \left( x_0 + \sum_{i=1}^k u_i / n \right) - F_1(x_0) \right] - \left[ F_2 \left( y_0 + \sum_{j=1}^l v_j / n \right) - F_2(y_0) \right] + \left[ F \left( x_0 + \sum_{i=1}^k u_i / n, y_0 + \sum_{j=1}^l v_j / n \right) - F \left( x_0 + \sum_{i=1}^k u_i / n, y_0 \right) \right] - \left[ F \left( x_0, y_0 + \sum_{j=1}^l v_j / n \right) - F(x_0, y_0) \right] \right\}_{x_0=\zeta+n^{-\frac{1}{2}}t_1, y_0=\eta+n^{-\frac{1}{2}}t_2}$$

Since  $F(x, y)$ ,  $F_1(x)$  and  $F_2(y)$  have continuous derivatives with respect to  $x$  and  $y$ , we apply the law of mean, and obtain after some simplification,

$$(14) \quad c_n = \left[ 1 - \left( \sum_{i=1}^k u_i / n \right) f_1 \left( \zeta + \frac{t_1}{n^{\frac{1}{2}}} \right) - \left( \sum_{j=1}^l v_j / n \right) f_2 \left( \eta + \frac{t_2}{n^{\frac{1}{2}}} \right) \right] + o(n^{-1})$$

so that

$$(15) \quad A_1^{(n)} \rightarrow \exp - \left[ f_1(\zeta) \sum_{i=1}^k u_i + f_2(\eta) \sum_{j=1}^l v_j \right].$$

Also, it is easily seen that

$$A_2^{(n)} = \prod_{j=0}^l (\sigma_j^{(1)} + \sigma_j^{(2)}) \prod_{i=0}^k (\rho_i^{(1)} + \rho_i^{(2)}).$$

The limiting value of  $A_2^{(n)}$  (evaluated at  $x_0 = \zeta + n^{-\frac{1}{2}}t_1$ ,  $y_0 = \eta + n^{-\frac{1}{2}}t_2$ ) is obtained as follows. Referring to (3), we see that

$$\begin{aligned}
 \rho_i^{(1)} + \rho_i^{(2)} &= \left( \int_{-\infty}^{y_0} f(x_i, v) dv + \int_{y_i}^{\infty} f(x_i, v) dv \right)_0 \\
 &= \left( f_1(x_i) - \int_{y_0}^{y_i} f(x_i, v) dv \right)_0 \\
 (16) \quad &= \left( f_1(x_i) - n^{-1} \sum_{j=1}^l v_j f(x_i, \xi) \right)_0, \quad (y_0 < \xi < y_i) \\
 &= f_1(\zeta) + \frac{x_i - \zeta}{n^{\frac{1}{2}}} \left( \frac{\partial f_1(x)}{\partial x} \right)_{x=\zeta} + O(n^{-1}) \rightarrow f_1(\zeta),
 \end{aligned}$$

where  $(\ )_0$  means that  $x_0 = \zeta + n^{-\frac{1}{2}}t_1$ ,  $y_0 = \eta + n^{-\frac{1}{2}}t_2$  are to be substituted in the expression within the brackets. In these calculations we used the fact that

$$x_i = \sum_{\alpha=1}^i u_i/n + t_1/n^{\frac{1}{2}} + \zeta, \quad y_l = \sum_{\beta=1}^l v_\beta/n + t_2/n^{\frac{1}{2}} + \eta.$$

Similarly,

$$(17) \quad \sigma_j^{(1)} + \sigma_j^{(2)} \rightarrow f_2(\eta).$$

It follows now from (16) and (17) that

$$(18) \quad A_2^{(n)} \rightarrow [f_1(\zeta)]^{k+1} [f_2(\eta)]^{l+1}.$$

Finally, the limiting expression for  $h^{(n)}(\mathbf{u}, \mathbf{v}, t_1, t_2)$  will follow on evaluating  $\lim A_3^{(n)}$ . To do this, we proceed as follows. From (9) and (14) (and noting that  $\theta_\mu = C_n^{-1} \rho_\mu$ ), it follows that  $c_n = 1 + O(1/n)$ ,  $\theta_\mu = \rho_\mu [1 + O(1/n)]$ . Using the law of the mean and recalling the definition of  $q_\mu$ ,  $\mu = 1, 2, 3, 4$ , it is easily shown that  $\rho_\mu = q_\mu [1 + O(n^{-\frac{1}{2}})]$ , and hence that  $\theta_\mu = q_\mu [1 + O(n^{-\frac{1}{2}})]$ ,  $\mu = 1, 2, 3, 4$ . Using the theorem in the Appendix, it readily follows that

$$\begin{aligned}
 (19) \quad n \sum_{n_1=0}^n \prod_{\mu=1}^4 [(n - k - l - 2)! / n_\mu!] \theta_\mu^n & \\
 &= \left[ k \exp \left\{ -\frac{1}{2} D \Sigma^{-1} D' \right\} \right] [1 + O(n^{-\frac{1}{2}})]
 \end{aligned}$$

where  $k = (2\pi |\Sigma|)^{-\frac{1}{2}}$ ,  $D = [t_1 f_1(\zeta) \ t_2 f_2(\eta)]$ , and

$$\Sigma = \begin{pmatrix} \alpha(1 - \alpha) & \alpha\beta - q_1 \\ \alpha\beta - q_1 & \beta(1 - \beta) \end{pmatrix}.$$

Thus from (10), (14), (18) and (19), we finally obtain that the limiting joint density function of  $U_1, \dots, U_k, V_1, \dots, V_l, T_1, T_2$  is given by

$$h(\mathbf{u}, \mathbf{v}, t_1, t_2) = g(t_1, t_2) \left[ \prod_{i=1}^k h_1(u_i) \right] \left[ \prod_{j=1}^l h_2(v_j) \right],$$

where  $g(t_1, t_2) = (2\pi|\Sigma|^{\frac{1}{2}})^{-1}[\exp\{-\frac{1}{2}D\Sigma^{-1}D'\}]f_1(\zeta)f_2(\eta)$ , (agreeing with a result obtained by Siddiqui [4], p. 148) and

$$h_1(u_i) = f_1(\zeta)e^{-u_i f_1(\zeta)}, \quad h_2(v_j) = f_2(\eta)e^{-v_j f_2(\eta)}.$$

Hence, the distances  $\{d_{r+i,1}^{(n)}\}_{i=0}^{k-1}$ ,  $\{d_{s+j,1}^{(n)}\}_{j=0}^{l-1}$  are mutually asymptotically independently distributed, independently of  $Z_r^{(n)}$  and  $W_s^{(n)}$ . Finally we infer the following

**THEOREM.**  $d_{r,k}^{(n)}$ ,  $d_{s,l}^{(n)}$  and  $d^{(n)} = Z_r^{(n)} - W_s^{(n)}$  are asymptotically independent.

**PROOF.** Since  $d_{r,j}^{(n)} = \sum_{i=0}^{k-1} d_{r+i,1}^{(n)}$  and  $d_{s,l}^{(n)} = \sum_{j=0}^{l-1} d_{s+j,1}^{(n)}$ , the asserted asymptotic independence follows from the asymptotic mutual independence of the r.v.'s  $d_{r+i,1}^{(n)}$  and  $d_{s+j,1}^{(n)}$ .

APPENDIX

Here we obtain a theorem on which the derivation of (19) is based. We use the following trivariate normal approximation to the multinomial probability law, as given by Gnedenko [2], p. 85.

**LEMMA.** Let  $f_n(n_1, n_2, n_3, n_4) = n! \prod_{i=1}^4 \lambda_i^{n_i} / n_i!$ , with  $0 < \lambda_i < 1, 0 \leq n_i \leq n, i = 1, 2, 3, 4$ , and  $n^{-1} \sum_{i=1}^4 n_i = \sum_{i=1}^4 \lambda_i = 1$ , and set  $v_i = n^{-\frac{1}{2}}(n_i - n\lambda_i)$ , so that  $v_4 = -(v_1 + v_2 + v_3)$ . Then uniformly in all the  $n_i$  for which the corresponding  $v_i$  lie in the arbitrary finite intervals  $c_i \leq v_i \leq d_i$ ,

$$f_n(n_1, n_2, n_3, n_4) = [e^{-Q/2} / (2n\pi)^{\frac{3}{2}} (\lambda_1 \lambda_2 \lambda_3 \lambda_4)^{\frac{1}{2}}] [1 + O(n^{-1})],$$

where  $a_{ij} = \lambda_i^{-1}$  for  $i \neq j$ ,  $a_{ii} = \lambda_i^{-1} + \lambda_i^{-1}$  and  $Q = \sum_{i=1}^3 \sum_{j=1}^3 a_{ij} v_i v_j$ .

For  $i = 1, 2, 3, 4$ , we define the quantiles  $p_i^{(n)}$  by  $p_1^{(n)} = F(x, y)$ ,  $p_2^{(n)} = F_2(y) - F(x, y)$ ,  $p_3^{(n)} = F_1(x) - F(x, y)$  and  $p_4^{(n)} = 1 - F_1(x) - F_2(y) + F(x, y)$ , with  $x = \zeta + n^{-\frac{1}{2}}t_1, y = \eta + n^{-\frac{1}{2}}t_2$ . Clearly,  $p_i^{(n)} \rightarrow q_i$  so that  $p_1^{(n)} + p_3^{(n)} \rightarrow \alpha, p_2^{(n)} + p_4^{(n)} \rightarrow 1 - \alpha, p_1^{(n)} + p_2^{(n)} \rightarrow \beta, p_3^{(n)} + p_4^{(n)} \rightarrow 1 - \beta$ , and  $p_2^{(n)} p_3^{(n)} - p_1^{(n)} p_4^{(n)} \rightarrow \alpha\beta - q_1$ . (Note the similarity between the  $p_i^{(n)}$  here and the  $\rho_i^{(n)}$  defined in Section 2). We now prove the following theorem.

**THEOREM.** For  $i = 1, 2, 3, 4$ , let  $v_i = n^{-\frac{1}{2}}(n_i - np_i^{(n)})$  with  $p_i^{(n)}$  as defined above. Then

$$n \sum_{n_1=0}^n f_n(n_1, n_2, n_3, n_4) \rightarrow (2\pi|\Sigma|^{\frac{1}{2}})^{-1} \exp\{-\frac{1}{2}D\Sigma^{-1}D'\},$$

where  $D = [t_1 f_1(\zeta) \quad t_2 f_2(\eta)]$ , and

$$\Sigma = \begin{pmatrix} \alpha(1 - \alpha) & \alpha\beta - q_1 \\ \alpha\beta - q_1 & \beta(1 - \beta) \end{pmatrix}.$$

**PROOF.** By the lemma,

$$(20) \quad n \sum_{n_1=0}^n f_n(n_1, n_2, n_3, n_4) = \left[ H_n \sum_{n_1=0}^n n^{-\frac{1}{2}} e^{-Q_n} \right] [1 + O(n^{-1})]$$

where  $H_n^{-1} = (2\pi)^{\frac{3}{2}} (p_1^{(n)} p_2^{(n)} p_3^{(n)} p_4^{(n)})^{\frac{1}{2}}$ , and  $Q_n$  is  $Q$  with  $\lambda_i$  replaced by  $p_i^{(n)}$ . For  $v_i$  in the arbitrary finite intervals  $c_i \leq v_i \leq d_i$ , we find that

$$v_1 + v_2 = n^{\frac{1}{2}} [\beta - (p_1^{(n)} + p_2^{(n)})] + \epsilon_2/n^{\frac{1}{2}}$$

$$v_1 + v_3 = n^{\frac{1}{2}} [\alpha - (p_1^{(n)} + p_3^{(n)})] + \epsilon_1/n^{\frac{1}{2}}$$

where  $\epsilon_1$  and  $\epsilon_2$  depend on  $n$  in such a way that both tend to zero with increasing  $n$ . Further, let  $u_1 = n^{\frac{1}{2}}[\beta - F_2(y)] + \epsilon_2/n^{\frac{1}{2}}$ ,  $u_2 = n^{\frac{1}{2}}[\alpha - F_1(x)] + \epsilon_1/n^{\frac{1}{2}}$ , so that  $v_2 = u_1 - v_1$ ,  $v_3 = u_2 - v_1$ . Putting  $\pi_1^{(n)} = \sum_{i=1}^4 [p_i^{(n)}]^{-1}$ ,  $\pi_j^{(n)} = [p_j^{(n)}]^{-1} + [p_4^{(n)}]^{-1}$ ,  $j = 1, 2$ , after algebraic simplifications we see that  $Q_n = Q_n^{(1)} + Q_n^{(2)}$ , where

$$Q_n^{(1)} = \pi_1^{(n)} [v_1 - (u_1\pi_2^{(n)} + u_2\pi_3^{(n)})/\pi_1^{(n)}]^2$$

and

$$Q_n^{(2)} = (\pi_1^{(n)})^{-1} \{ \pi_2^{(n)} [\pi_1^{(n)} - \pi_2^{(n)}] u_1^2 + 2 [\pi_1^{(n)}/p_4^{(n)} - \pi_2^{(n)} \pi_3^{(n)}] u_1 u_2 + \pi_3^{(n)} [\pi_1^{(n)} - \pi_3^{(n)}] u_2^2 \}.$$

Substituting in (20), we obtain

$$n \sum_{n_1=0}^n f_n(n_1, n_2, n_3, n_4) = H_n L_n M_n [1 + O(n^{-1})],$$

where  $L_n = \exp(-Q_n^{(2)}/2)$  and  $M_n = \sum_{n_1=0}^n n^{-\frac{1}{2}} \exp(-Q_n^{(1)}/2)$ . From the expansions  $F_1(\zeta + n^{-\frac{1}{2}}t_1) = F_1(\zeta) + n^{-\frac{1}{2}}t_1 f_1'(\zeta) + (2n)^{-1} t_1^2 f_1''(\theta_1)$ ,  $\zeta < \theta_1 < \zeta + n^{-\frac{1}{2}}t_1$ , and  $F_2(\eta + n^{-\frac{1}{2}}t_2) = F_2(\eta) + n^{-\frac{1}{2}}t_2 f_2'(\eta) + (2n)^{-1} t_2^2 f_2''(\theta_2)$ ,  $\eta < \theta_2 < \eta + n^{-\frac{1}{2}}t_2$ , we obtain easily that

$$u_1 = -t_2 f_2'(\eta) + O(n^{-\frac{1}{2}}) \rightarrow -t_2 f_2'(\eta)$$

$$u_2 = -t_1 f_1'(\zeta) + O(n^{-\frac{1}{2}}) \rightarrow -t_1 f_1'(\zeta).$$

Simple calculations now show that  $L_n \rightarrow \exp(-\frac{1}{2}D\Sigma^{-1}D')$  and  $H_n^{-1} \rightarrow (2\pi)^{\frac{1}{2}}(q_1 q_2 q_3 q_4)^{\frac{1}{2}}$ . Noting that  $M_n$  is a Riemann sum, and that  $c_i, d_i$  are arbitrary, it can be shown that  $M_n \rightarrow (2\pi\psi_1^{-1})^{\frac{1}{2}}$ , where  $\psi_1 = \lim \pi_1^{(n)} = \sum_{i=1}^4 q_i^{-1}$ . Since  $\psi_1 \prod_{i=1}^4 q_i = \alpha\beta(1 - \alpha)(1 - \beta) - (\alpha\beta - q_1)^2 = |\Sigma|$ , the theorem follows.

REFERENCES

[1] CRAMÉR, H. (1946). *Mathematical Methods in Statistics*. Princeton Univ. Press.  
 [2] GNEDENKO, B. V. (1962). *The Theory of Probability*. Chelsea, New York.  
 [3] MOOD, A. M. (1941). On the joint distribution of the medians in samples from a multivariate population. *Ann. Math. Statist.* **12** 268-278.  
 [4] SIDDIQUI, M. M. (1960). Distribution of quantiles in samples from a bivariate population. *J. Res. NBS* **64B (Math. and Math. Phys)** 145-150.  
 [5] WILKS, S. S. (1948). Order statistics. *Bull. Amer. Math. Soc.* **54** 6-50.