

# NOTES

## ON EXTENDED RATE OF CONVERGENCE RESULTS FOR THE INVARIANCE PRINCIPLE

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Let  $\{X_i, i = 1, 2, 3, \dots\}$  be a sequence of independent and identically distributed random variables for which  $EX_i = 0$ ,  $\text{var } X_i = \sigma^2$ ,  $E|X_i|^{2+a} = b < \infty$ ,  $a > 0$ .  $C[0, 1]$  is the space of continuous real-valued functions on the interval  $[0, 1]$  with the sup norm topology. We define a "random broken line" as follows:

$$\begin{aligned} \xi_n(t) &= \sum_{i=1}^k X_i / (\sigma n^{\frac{1}{2}}) && \text{for } t = n^{-1}k \\ \xi_n(t) &= \xi_n(n^{-1}k) + n[\xi_n(n^{-1}(k+1)) - \xi_n(n^{-1}k)](t - n^{-1}k) \\ &&& \text{for } n^{-1}k \leq t \leq n^{-1}(k+1), 1 \leq k \leq n. \end{aligned}$$

**THEOREM.** *Let  $g$  be any uniformly continuous functional,  $g: C[0, 1] \rightarrow \text{real line}$ , such that there exists a constant  $L > 0$  with the property that*

$$|\Pr \{g(\xi(t)) \leq x + h\} - \Pr \{g(\xi(t)) \leq x\}| \leq L|h|$$

where  $\xi(t)$  is the standard Wiener process. Then, there exists a constant  $A > 0$  such that for all  $n > 1$ ,

$$(1) \quad |\Pr \{g(\xi_n(t)) \leq x\} - \Pr \{g(\xi(t)) \leq x\}| \leq A (\log n)^\lambda n^{-\mu},$$

where  $\lambda = (1 + \frac{1}{2}a)/(a + 3) (< \frac{1}{2})$ ,  $\mu = \min(a, 1 + \frac{1}{2}a)/2(a + 3)$ . This result extends that of Theorem 5 of Rosenkrantz [3] where the bound (1) is, in essence, given in the form

$$|\Pr \{g(\xi_n(t)) \leq x\} - \Pr \{g(\xi(t)) \leq x\}| \leq A (\log n)^{\frac{1}{2}} n^{-\mu}$$

subject to the restriction that  $a \leq 2$ . Only minor modifications of the work of Rosenkrantz and necessary to furnish this extension; we shall just indicate these and refer the reader to [3] for full details of the proof.

The condition  $a \leq 2$  was imposed by Rosenkrantz as a consequence of employing a moment inequality of von Bahr and Esseen [1] which is valid for exponent  $r$ ,  $1 \leq r \leq 2$  (note that  $2r = 2 + a$ ). This is used to obtain the vital inequality (25) of [3]. For the case  $a > 2$  ( $r > 2$ ), we make use of a result of Dharmadhikari, Fabian and Jogdeo [2] which gives in place of (25),

$$(2) \quad E|z_{nn}|^r \leq C_r n^{\frac{1}{2}} v_n,$$

where  $C_r = [8(r - 1) \max(1, 2^{r-3})]^r$ .

Thus, combining (2) with (25) of [3], we can replace the estimate (27) of [3]

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by the general estimate

$$(3) \quad \Pr (\max_{1 \leq k \leq n} |z_{nk}| > \delta) \leq \max (2, C_r) n^{-a(r)} \delta^{-r}$$

where  $a(r) = \min (r - 1, \frac{1}{2}r)$ ,  $r > 1$ .

The details of proof of [3] are then followed exactly up to the stage of choosing  $\epsilon_n$  and  $\delta_n$  tending to zero at appropriate rates. Instead of the versions chosen in (37), we employ

$$\delta_n = n^{-2a(r)(2r+1)^{-1}} (\beta \log n)^{-(2r+1)^{-1}}, \quad \epsilon_n^2 = 2\beta\delta_n \log n,$$

where  $\beta > 1 + a(r)/(2r + 1)$ . The result of the theorem then follows as with [3].

REFERENCES

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