## SOME SAMPLE FUNCTION PROPERTIES OF THE TWO-PARAMETER GAUSSIAN PROCESS

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Let  $\{X(s,t;\omega)\colon (s,t)\in [0,\infty)\times [0,\infty)\}$  be a two parameter Gaussian process with mean function zero and covariance function  $R(s_1,t_1;s_2,t_2)=\min{(s_1,s_2)}\min{(t_1,t_2)}$ . This paper derives a multiparameter law of the iterated logarithm and modulus of continuity for the process  $X(s,t;\omega)$ . Estimates are also given which enable the author to define an Itô type integral for a suitable class of functions and to solve a diffusion equation involving the process.

1. Introduction. A number of papers have appeared in the literature defining multiparameter analogs of the Brownian motion process. Čencov [3] and Yeh [12] have shown that a multiparameter process with parameter space the p-dimensional unit cube A and covariance function  $R[(u_1, \dots, u_p), (v_1, \dots, v_p)] = \min(u_1, v_1) \cdots \min(u_p, v_p)$  can be realized in the space of continuous functions on A which vanish on  $A_0 = \{(u_1, \dots, u_p) \in A : u_j = 0 \text{ for some } j, 1 \leq j \leq p\}$ . Delporte [5] and W. Park [11] construct such a process on the unit cube using a Haar function expansion and W. Park generalizes some results of C. Park, Shepp and Yeh.

In the present paper a Haar function construction was used with an arctangent transformation (Ciesielski [4]) to define a Gaussian process  $\{X(s, t; \omega) : (s, t) \in [0, \infty) \times [0, \infty)\}$  with mean function  $m(s, t) \equiv 0$  and covariance function  $R(s_1, t_1; s_2, t_2) = \min(s_1, s_2) \min(t_1, t_2)$ . The sample functions of this process are continuous and the process has independent increments (i.e. if  $0 = s_0 < \cdots < s_m = S$ ,  $0 = t_0 < \cdots < t_n = T$  partitions  $[0, S] \times [0, T]$ , the random variables  $\{\Delta X[(s_i, t_j), (s_{i-1}, t_{j-1}); \omega] : i = 1, \cdots, m, j = 1, \cdots, n\}$  where  $\Delta X[(s_i, t_j), (s_{i-1}, t_{j-1}); \omega] = X(s_i, t_j; \omega) - X(s_i, t_{j-1}; \omega) - X(s_{i-1}, t_j; \omega) + X(s_{i-1}, t_{j-1}; \omega)$ , are mutually independent).

Sample function properties of the process  $X(s, t; \omega)$  are examined, an Itô type integral is defined for a suitable class of functions and a diffusion equation is solved. Some properties of the integral and the solution of the diffusion equation are also investigated.

A different generalization of Brownian motion to a p-dimensional parameter space has been discussed by Lévy [9].

2. Sample function properties of  $X(s, t; \omega)$ . For fixed  $t = t_0$ ,  $X(s, t; \omega)$  is a one-dimensional Brownian motion process with mean function zero and covariance function  $R(s_1, s_2) = t_0 \min(s_1, s_2)$ .

We define a partial ordering on  $[0, \infty) \times [0, \infty)$  by (s', t') < (s, t) if  $s' \leq s$ ,

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 $t' \leq t$ . Let  $\mathscr{F}_{st}$  be the  $\sigma$ -field generated by the random variables  $\{X(u, v; \omega):$ (u, v) < (s, t).

Lemma 2.1. Let  $0 = s_0 < \cdots < s_m = S$ ,  $0 = t_0 < \cdots < t_n = T$  be a partition of  $[0, S] \times [0, T]$ . Let  $(s, t) \in [0, S] \times [0, T]$ . If  $s_{i-1} \ge s$  or  $t_{j-1} \ge t$ , then  $\Delta X[(s_i, t_j), (s_{i-1}, t_{j-1}); \omega]$  is independent of  $\mathcal{F}_{st}$ .

**PROOF.** Since  $X(s, t; \omega)$  is a Gaussian process and since  $\Delta X[(s_i, t_i), (s_{i-1}, t_{i-1})] =$  $\Delta X_{ij}$  independent of every finite linear combination of  $X(u_1, v_1), \dots, X(u_n, v_n)$ ,  $(u_i, v_i) < (s, t), i = 1, \dots, n$ , implies  $\Delta X_{ij}$  is independent of  $\mathcal{F}_{st}$ , it is sufficient to show that  $E[\Delta X_{ij}X(u, v)] = 0$  for (u, v) < (s, t).

LEMMA 2.2. Let  $(s_1, t_1), (s_2, t_2) \in [0, \infty) \times [0, \infty)$ . The random variable  $\Delta X[(s_2, t_2),$  $(s_1, t_1)$ ;  $\omega$ ] is Gaussian with mean zero and covariance  $|s_2 - s_1||t_2 - t_1|$ .

PROOF. Clearly  $\Delta X[(s_2, t_2), (s_1, t_1); \omega]$  is a Gaussian random variable with mean zero. The covariance can be verified by a simple computation.

Theorem 1. For  $\lambda \geq 0$ ,

$$(2.1) P[\omega: \sup_{(s,t)\in[0,S]\times[0,T]} X(s,t;\omega) \ge \lambda] \le 4P[X(S,T;\omega) \ge \lambda].$$

PROOF. Consider  $X(s, t; \omega)$  at the points  $\{(iS2^{-m}, jT2^{-n}): i = 0, 1, \dots, 2^m; \}$  $j=0,\,\cdots,\,2^n$ . Let  $Z_i(\omega)=\max_{0\leq j\leq 2^n}X(iS2^{-m},jT2^{-n};\,\omega)$  and let

$$I(\omega) = \inf \{i \colon Z_i(\omega) \ge \lambda\} \qquad (I(\omega) = +\infty \text{ if } \{i \colon Z_i(\omega) \ge \lambda\} = \emptyset),$$
  
 $J(\omega) = \inf \{j \colon X(IS2^{-m}, jT2^{-n}; \omega) \ge \lambda\} \qquad (J(\omega) = +\infty \text{ if } I(\omega) = +\infty).$ 

Then

$$P[\max_{i,j} X(iS2^{-m}, jT2^{-n}) \ge \lambda] = P[X(IS2^{-m}, JT2^{-n}) \ge \lambda]$$

$$= P[X(IS2^{-m}, JT2^{-n}) \ge \lambda, X(S, JT2^{-n}) \ge \lambda]$$

$$+ P[X(IS2^{-m}, JT2^{-n}) \ge \lambda, X(S, JT2^{-n}) < \lambda].$$

Now, using Lemma 2.1, the symmetry of the increments of  $X(s, t; \omega)$  and the fact that  $X(s, 0; \omega) = 0$  a.s. for  $s \in [0, \infty)$ , we have

$$P[X(IS2^{-m}, JT2^{-n}) \ge \lambda, X(S, JT2^{-n}) < \lambda]$$

$$= \sum_{i,j} P[I(\omega) = i, J(\omega) = j, X(S, JT2^{-n}) < \lambda]$$

$$\le \sum_{i,j} P[I(\omega) = i, J(\omega) = j, X(S, jT2^{-n}) - X(iS2^{-m}, jT2^{-n}) < 0]$$

$$= \sum_{i,j} P[I(\omega) = i, J(\omega) = j]P[\Delta X[(S, jT2^{-n}), (iS2^{-m}, 0)] < 0]$$

$$= \sum_{i,j} P[I(\omega) = i, J(\omega) = j]P[\Delta X[(S, jT2^{-n}), (iS2^{-m}, 0)] > 0]$$

$$\le \sum_{i,j} P[I(\omega) = i, J(\omega) = j, X(S, jT2^{-n}) \ge \lambda]$$

$$\le P[X(S, JT2^{-n}) \ge \lambda].$$

Combining this result with (2.2) gives

(2.4) 
$$P[\max_{i,j} X(iS2^{-m}, jT2^{-n}) \ge \lambda] \le 2P[X(S, JT2^{-n}) \ge \lambda]$$
  
  $\le 2P[\sup_i X(S, jT2^{-n}) \ge \lambda] \le 4P[X(S, T) \ge \lambda]$ 

by the corresponding theorem for one-dimensional Brownian motion. Using the continuity of the sample paths of the process  $X(s, t; \omega)$  and letting  $n \to \infty$  gives (2.1).

Let f(x, y) be a function defined on  $[0, \infty) \times [0, \infty)$ . By  $\limsup_{s,t\to\infty} f(s,t)$  we shall mean  $\lim_{s,t\to\infty} \sup_{(u,v)>(s,t)} f(u,v)$ . Theorem 2 is a multiparameter version of the law of the iterated logarithm. If the  $\limsup$  is taken as  $s,t\to\infty$ , the constant in the multiparameter version is equal to 4, however if the  $\limsup$  is taken as  $s\to\infty$  while t remains in some bounded interval  $0 < a \le t \le b < \infty$ , the result has constant equal to 2 as in the one parameter version of the theorem. This is shown in Theorem 3.

THEOREM 2. 
$$P\left[\omega: \limsup_{s,t\to\infty}\frac{X(s,t;\omega)}{[4st\log_2 st]^{\frac{1}{2}}}=1\right]=1.$$

PROOF. The proof is an analog of the standard proof of the one-dimensional theorem and will be omitted.

THEOREM 3. Let  $0 < a \le b < \infty$ , then

(i) 
$$P\left[\omega: \limsup_{s\to\infty} \sup_{a\leq t\leq b} \frac{X(s, t; \omega)}{[2st\log_s st]^{\frac{1}{2}}} = 1\right] = 1,$$

(ii) 
$$P\left[\omega: \limsup_{s\to\infty} \frac{X(s,\,t;\,\omega)}{[2st\log_2 st]^{\frac{1}{2}}} \ge 1 \quad for \ all \quad t\in[a,\,b]\right] = 1.$$

PROOF. Let  $0 < \varepsilon < 1$  and suppose that

$$(2.5) P\Big[\omega: \limsup_{s\to\infty} \sup_{a\leq t\leq b} \frac{X(s,\,t;\,\omega)}{[2st\log_2 st]^{\frac{1}{2}}} > 1 + \varepsilon\Big] > 0.$$

Divide the interval [a, b] into m equal parts each of length  $\delta = (b - a)/m$ . If (2.5) is true then for each m there exists some subinterval  $[a_m, b_m]$  contained in [a, b] such that

$$(2.6) P\Big[\omega: \limsup_{s\to\infty}\sup_{a_m\leq t\leq b_m}\frac{X(s,\,t;\,\omega)}{[2st\log_2 st]^{\frac{1}{2}}}>1+\varepsilon\Big]>0.$$

Let m be chosen so that  $\delta=(b-a)/m<(a\varepsilon)/2$  and let  $1< q<1/(1-\varepsilon/2)$ . Let  $G(x,y)=[2xy\log_2 xy]^{\frac{1}{2}}$  and  $A_{k\delta}=[\omega:\sup_{0< s\leq q^k}\sup_{a_m\leq t\leq b_m}X(s,t;\omega)>(1+\varepsilon)G(q^{k-1},a_m)]$ . Then

$$\begin{split} P[A_{k\delta}] & \leq P[\omega : \sup_{0 < s \leq q^k; \ 0 < t \leq b_m} X(s, t; \omega) > (1 + \varepsilon) G(q^{k-1}, a_m)] \\ & \leq 4P[\omega : X(q^k, a_m + \delta) > (1 + \varepsilon) G(q^{k-1}, a_m)] \\ & \leq \frac{2q^{\frac{1}{2}}[(a_m + \delta)/a_m]^{\frac{1}{2}}}{(1 + \varepsilon)[\pi \log_2 (q^{k-1}a_m)]^{\frac{1}{2}}} [(k - 1) \log q + \log a_m]^{-\gamma} \end{split}$$

where

$$\gamma = \frac{(1+\varepsilon)^2 a_m}{q(a_m+\delta)} > \frac{(1+\varepsilon)^2 (1-\varepsilon/2)}{(1+\varepsilon/2)} > 1.$$

Hence  $\sum_{k=1}^{\infty} P[A_{k\delta}] < \infty$  and by the Borel-Cantelli lemma  $P[A_{k\delta} \text{ i.o.}] = 0$  which contradicts (2.6). Thus

$$(2.7) P\Big[\omega: \limsup_{s\to\infty} \sup_{a\leq t\leq b} \frac{X(s,t;\omega)}{[2st\log_s st]^{\frac{1}{2}}} < 1 + \varepsilon\Big] = 1.$$

For any fixed t,  $a \le t \le b$ , using the one-dimensional theorem

$$P\bigg[\omega: \limsup_{s\to\infty}\frac{X(s,\,t;\,\omega)}{[2st\log_2 st]^{\frac{1}{2}}}>1-\varepsilon\bigg]=1\;.$$

Hence

$$(2.8) P\Big[\omega: \limsup_{s\to\infty} \sup_{a\leq t\leq b} \frac{X(s,t;\omega)}{[2st\log_s st]^{\frac{1}{2}}} > 1-\varepsilon\Big] = 1.$$

Combining (2.7) and (2.8) we have (i).

We now prove (ii). Without loss of generality we may assume that  $b \le 1$ , since  $X(bs, b^{-1}t; \omega)$  has the same distribution as  $X(s, t; \omega)$  and

$$\begin{split} \left[\omega: \limsup_{s \to \infty} \frac{X(s, t; \omega)}{[2st \log_2 st]^{\frac{1}{2}}} > 1 - \varepsilon \quad \text{for all} \quad t \in [a, b] \right] \\ &= \left[\omega: \limsup_{u \to \infty} \frac{X(u, v; \omega)}{[2uv \log_2 uv]^{\frac{1}{2}}} > 1 - \varepsilon \quad \text{for all} \quad v \in [ab^{-1}, 1] \right]. \end{split}$$

Suppose

$$(2.9) P\Big[\omega: \limsup_{s\to\infty}\frac{X(s,\,t;\,\omega)}{[2st\log_s st]^{\frac{1}{2}}} \leq 1-\varepsilon \text{ for some } t\in[a,\,b]\Big] > 0.$$

As above, for each m divide the interval [a, b] into m subintervals each of length  $\delta = (b - a)/m$ . Then (2.9) implies that for each m there exists some subinterval  $[a_m, b_m]$  of length  $\delta$  such that

$$(2.10) P\Big[\omega: \limsup_{s\to\infty}\frac{X(s,\,t;\,\omega)}{[2st\log_s st]^{\frac{1}{2}}} \le 1-\varepsilon \text{ for some } t\in[a_m,\,b_m]\Big]>0.$$

Let *m* be chosen so that,  $\delta < (a\varepsilon^4)/4$  and let q > 1 be such that  $2/[(q-1)^{\frac{1}{2}}] < \varepsilon/4$ ,  $[q/(q-1)]^{\frac{1}{2}} < 1 + \varepsilon/2$ . From the proof of the law of the iterated logarithm for one-dimensional Brownian motion (Loève [10] page 560), we have

$$(2.11) X(q^n, b_m) - X(q^{n-1}, b_m) > \{(1 - \varepsilon/4)b_m^{\frac{1}{2}}[(q-1)/q]^{\frac{1}{2}}[2q^n \log_2 q^n]^{\frac{1}{2}}\} i.o.$$

The first half of the theorem implies that for  $a \le t \le b$ ,  $n \ge n_1(q, \omega)$ ,

$$|X(q^{n-1}, t)| \leq 2(t/q)^{\frac{1}{2}} [2q^n \log_2 q^n]^{\frac{1}{2}}.$$

For  $t \in [a_m, b_m]$ , let  $A_n(t) = [\omega : \Delta X[(q^n, b_m), (q^{n-1}, t)] \ge [2q^n \log_2 q^n]^{\frac{1}{2}} \gamma(t)]$ , where  $\gamma(t) = \{(1 - \varepsilon/4)b_m^{\frac{1}{2}}[(q-1)/q]^{\frac{1}{2}} - 2(t/q)^{\frac{1}{2}} - (1 - \varepsilon)t^{\frac{1}{2}}\}$  and let  $A_n = (1 - \varepsilon)t^{\frac{1}{2}}$ 

$$\begin{split} [\omega: \sup_{a_m \leq t \leq b_m} \Delta X[(q^n, b_m), (q^{n-1}, t)] & \geq [2q^n \log_2 q^n]^{\frac{1}{2}} \gamma(b_m)]. \text{ Then } A_n(t) \text{ is a subset of } A_n, \ a_m \leq t \leq b_m, \ n = 1, 2, \cdots, \text{ and } P[A_n] = P[\omega: \sup_{a_m \leq t \leq b_m} \Delta X[(q^n, b_m), (q^{n-1}, t)] \geq [2q^n \log_2 q^n]^{\frac{1}{2}} \gamma(b_m)] \leq P[\omega: X(1, 1; \omega) \geq \{[2q^n \log_2 q^n]^{\frac{1}{2}} \gamma(b_m)\}/[(q^n - q^{n-1})\delta]^{\frac{1}{2}}] \leq \text{Constant } [n \log q]^{-\alpha} \text{ where } \alpha = [q/[(q-1)\delta]]\gamma^2(b_m) = (b_m/\delta)\{(1-\varepsilon/4) - 2/(q-1)^{\frac{1}{2}} - (1-\varepsilon)[q/(q-1)]^{\frac{1}{2}}\} > (a/\delta)(\varepsilon^4/4) > 1. \text{ Hence } \sum_{n=1}^\infty P[A_n] < \infty \text{ and by the Borel-Cantelli lemma } P[A_n \text{ i.o.}] = 0. \text{ Thus if } n \geq n_2(\omega) \end{split}$$

$$(2.13) \Delta X[(q^n, b_m), (q^{n-1}, t)] < [2q^n \log_2 q^n]^{\frac{1}{2}} \gamma(t) \text{ for all } t \in [a_m, b_m].$$

If  $n \ge \max\{n_1(\omega), n_2(\omega)\}$ , then (2.12) and (2.13) hold. If (2.11) also holds it follows that

$$(2.14) X(q^n, t) > (1 - \varepsilon)t^{\frac{1}{2}} [2q^n \log_2 q^n]^{\frac{1}{2}} > (1 - \varepsilon)[2q^n t \log_2 (q^n t)]^{\frac{1}{2}}$$

for all  $t \in [a_m, b_m]$ . Hence (2.14) is true infinitely often contradicting (2.10). Thus (ii) follows.

COROLLARY. There exists a set  $\Omega_0$  of probability zero such that if  $\omega \notin \Omega_0$ ,  $X(s, t; \omega)$  satisfies the one-parameter law of the iterated logarithm (as a function of s) simultaneously on the lines t = T,  $0 < T < \infty$ .

PROOF. (i) and (ii) imply that for any finite interval [a, b],

$$(2.15) P\bigg[\omega: \limsup_{s\to\infty}\frac{X(s,\,t;\,\omega)}{[2st\log_s st]^{\frac{1}{2}}}=1 \text{ for all } t\in[a,\,b]\bigg]=1.$$

Let  $\{a_n\}$ ,  $\{b_n\}$  be sequences such that  $a_n \downarrow 0$ ,  $b_n \uparrow \infty$ . Let

$$A_n = \left[\omega : \limsup_{s \to \infty} \frac{X(s, t; \omega)}{[2st \log_2 st]^{\frac{1}{2}}} = 1 \text{ for all } t \in [a_n, b_n]\right].$$

Then  $A_n \downarrow$ , and by (2.15)  $P[A_n] = 1$  for each n, hence  $P[\lim_{n\to\infty} A_n] = P[\omega]$ :  $\limsup_{s\to\infty} X(s,t;\omega)/[2st\log_2 st]^{\frac{1}{2}} = 1$  for all  $t \in (0,\infty)$ ] =  $\lim_{n\to\infty} P[A_n] = 1$ .

Theorem 4. Let S,  $T \ge 1$ ,  $(s_1, t_1)$ ,  $(s_2, t_2) \in [0, S] \times [0, T]$ , then

$$(2.16) P\Big[\omega: \limsup_{|t_2-t_1|=\eta\downarrow 0; |s_2-s_1|=\varepsilon\downarrow 0} \frac{|\Delta X[(s_2,t_2),(s_1,t_1);\omega]|}{[2\varepsilon\eta\log(1/\varepsilon\eta)]^{\frac{1}{2}}} = 1\Big] = 1.$$

PROOF. The proof uses generalizations of the ideas and methods of the onedimensional case as given by Itô and McKean [8].

3. Stochastic integrals and diffusion equations. Let  $\{Z(s, t; \omega) : (s, t) \in [0, \infty) \times [0, \infty)\}$  be a stochastic process.  $Z(s, t; \omega)$  is said to be a martingale with respect to  $\mathscr{F}_{st} = \sigma\{Z(u, v; \omega) : (u, v) < (s, t)\}$  if  $Z(s, t; \omega)$  is integrable for all  $(s, t) \in [0, \infty) \times [0, \infty)$  and whenever  $(s, t), (s', t') \in [0, \infty) \times [0, \infty)$  are such that  $(s', t') < (s, t), E[Z(s, t; \omega) | \mathscr{F}_{s't'}] = Z(s', t'; \omega)$  a.s.

THEOREM 5. Let  $(s_0, t_0)$  be a fixed point in  $[0, S] \times [0, T]$ . The process  $\{\Delta X[(s, t), (s_0, t_0); \omega] : s_0 \le s \le S, t_0 \le t \le T\}$  is a martingale.

PROOF. Clearly  $\Delta X[(s,t),(s_0,t_0);\omega]$  is integrable for all (s,t). Let  $\mathscr{C}_{st}=\sigma\{X(u,v;\omega):s_0\leq u\leq s,\,t_0\leq v\leq t\}$ . If  $(s_0,t_0)<(s',t')<(s,t)$ , write  $\Delta X[(s,t),t]$ 

 $(s_0, t_0)] = \Delta X[(s, t), (s', t')] + \Delta X[(s', t'), (s_0, t_0)] + \Delta X[(s', t), (s_0, t')] + \Delta X[(s, t'), (s', t_0)].$  By Lemma 2.1,  $\Delta X[(s, t), (s', t')], \Delta X[(s, t'), (s', t_0)]$  and  $\Delta X[(s', t), (s_0, t')]$  are independent of  $\mathcal{G}_{s't'}$ .  $\Delta X[(s', t'), (s_0, t_0)]$  is measurable with respect to  $\mathcal{G}_{s't'}$ . Hence  $E[\Delta X[(s, t), (s_0, t_0)]] \mathcal{G}_{s't'}] = \Delta X[(s', t'), (s_0, t_0)].$ 

We want to define an Itô type integral  $I(s, t, \omega; f)$  for the class  $\mathscr{M}$  of square integrable functions  $f(s, t; \omega)$  defined on  $[0, S] \times [0, T] \times \Omega$  having the property that  $f(s, t; \omega)$  is measurable with respect to  $\mathscr{F}_{st} = \sigma\{X(u, v; \omega) : (u, v) < (s, t)\}$  for all  $(s, t) \in [0, S] \times [0, T]$ . The following lemmas give the estimates needed to construct such an integral and to solve a differential equation involving  $X(s, t; \omega)$ .

LEMMA 3.1. Let  $X(\omega)$  and  $Y(\omega)$  be random variables on a probability space  $(\Omega, \mathcal{F}, P)$ . Let  $\mathcal{F}_0$  be a sub- $\sigma$ -field of  $\mathcal{F}$ . If (i) X is independent of Y, (ii) X is independent of  $\mathcal{F}_0$ , (iii) E[X] = 0, (iv)  $E[XY] < \infty$ , then  $E[XY | \mathcal{F}_0] = 0$  a.s.

LEMMA 3.2. Let  $X_{ij}(\omega)$  and  $Y_{ij}(\omega)$   $i=1, \dots, m; j=1, \dots, n$  be random variables such that

- (i)  $E[X_{ij}(\omega)] = 0$  for all i and j,
- (ii)  $E[X_{ij}(\omega)Y_{ij}(\omega)] < \infty$  for all i and j
- (iii)  $Y_{uv}(\omega)$   $u \leq i$ ,  $v \leq j$  and  $X_{uv}(\omega)$   $u \leq i-1$ ,  $v \leq j-1$  are independent of  $X_{uv}(\omega)$   $u \geq i$  or  $v \geq j$ . Let  $S_{ij}(\omega) = \sum_{u=1}^{i} \sum_{v=1}^{j} X_{uv}(\omega) Y_{uv}(\omega)$ , then

(3.1) 
$$P[\omega : \max_{i,j} |S_{ij}| \ge b] \le (4/b^2) E[S_{mn}]^2$$

$$(3.2) E[\max_{i,j} |S_{ij}|^2] \le 16E[S_{mn}]^2.$$

PROOF. We first show that  $\{S_{mj}: 1 \leq j \leq n\}$  is a martingale with respect to  $\mathscr{F}_{mj}$ . If  $j' \leq j$ , using Lemma 3.1 and the fact that  $S_{mj'}$  is  $\mathscr{F}_{mj'}$  measurable, we have

$$\begin{split} E[S_{mj} \, | \, \mathscr{F}_{mj'}] &= \sum_{u=1}^{m} \sum_{v=1}^{j'} E[X_{uv} \, Y_{uv} \, | \, \mathscr{F}_{mj'}] + \sum_{u=1}^{m} \sum_{v=j'+1}^{n} E[X_{uv} \, Y_{uv} \, | \, \mathscr{F}_{mj'}] \\ &= \sum_{v=1}^{m} \sum_{v=1}^{j'} X_{vv} \, Y_{uv} = S_{mj'} \, . \end{split}$$

Similarly  $E[S_{mj} | \mathscr{F}_{in}] = S_{ij}$  and thus using Jensen's inequality  $E[|S_{mj}| | \mathscr{F}_{in}] \ge |S_{ij}|$ .

Let  $Z_i(\omega) = \max_{1 \le j \le n} |S_{ij}(\omega)|$  and let  $I(\omega) = \min \{i : Z_i(\omega) \ge b\}$   $(I(\omega) = +\infty)$  if  $\{i : Z_i(\omega) \ge b\} = \emptyset$ ,  $J(\omega) = \min \{j : |S_{ij}| \ge b\}$   $(J(\omega) = +\infty)$  if  $I(\omega) = +\infty$ . Now  $\Lambda = [\omega : \max_{i,j} |S_{ij}| \ge b] = \bigcup_{i=1}^m \bigcup_{j=1}^n [\omega : I(\omega) = i, J(\omega) = j] = \bigcup_{i=1}^m \bigcup_{j=1}^n B_{ij}$  and the sets  $B_{ij}$  are disjoint.

For any (i,j), (0,0) < (i,j) < (m,n)  $\int_{B_{ij}} S_{mj}^2 dP = E[S_{ij}^2 I_{B_{ij}}] + E[(S_{mj} - S_{ij})^2 I_{B_{ij}}] \ge \int_{B_{ij}} S_{ij}^2 dP$  since  $E[S_{ij}(S_{mj} - S_{ij})I_{B_{ij}}] = 0$ . Thus

(3.3) 
$$E[S_{mJ}^2] \ge \int_{\Lambda} S_{mJ}^2 dP = \sum_{i=1}^m \sum_{j=1}^n \int_{B_{ij}} S_{mj}^2 dP \ge \sum_{i=1}^m \sum_{j=1}^n \int_{B_{ij}} S_{ij}^2 dP$$
 
$$\ge b^2 P[\Lambda]$$

and since  $\{S_{mj}: 1 \leq j \leq n\}$  is a martingale with respect to  $\mathscr{F}_{mj}$ , we have

(3.4) 
$$E[S_{mJ}^2] \le E[\max_{1 \le j \le n} S_{mj}^2] \le 4E[S_{mn}^2]$$

(Doob [6] Chapter VII, Theorem 3.4).

(3.1) follows from (3.3) and (3.4).

To prove (3.2), we note that

$$P[\omega : \max_{i,j} |S_{ij}| \ge b] = \sum_{i=1}^{m} \sum_{j=1}^{n} P[B_{ij}] \le (1/b) \sum_{i=1}^{m} \sum_{j=1}^{n} \int_{B_{ij}} |S_{ij}| dP$$

$$\le (1/b) \sum_{i=1}^{m} \sum_{j=1}^{n} \int_{B_{ij}} E[|S_{mj}|| \mathscr{F}_{in}] dP = (1/b) \sum_{i=1}^{m} \sum_{j=1}^{n} \int_{B_{ij}} |S_{mj}| dP$$

$$= (1/b) \int_{\Lambda} |S_{m,l}| dP$$

since  $B_{ij} \in \mathscr{F}_{in}$ . Using Theorem 3.4' (Chapter VII [6]) and (3.4), we have  $E[\max_{i,j} |S_{ij}|^2] = E[(\max_{i,j} |S_{ij}|^2] \le 4E[|S_{mJ}|^2] \le 16E[S_{mn}^2]$ .

Let  $f(s,t;\omega)$  be an  $\mathscr{F}_{st}$  measurable function for all s and t. Let  $0=s_0 < s_1 < \cdots < s_m = S$ ,  $0=t_0 < t_1 < \cdots < t_n = T$  partition  $[0,S] \times [0,T]$ . Consider the product  $f(s_{i-1},t_{j-1};\omega)f(s_{u-1},t_{v-1};\omega)\Delta X_{ij}\Delta X_{uv}$ . Without loss of generality we may assume that  $s_{u-1} \leq s_{i-1}$ . Then  $f(s_{u-1},t_{v-1};\omega)$  and  $f(s_{i-1},t_{j-1};\omega)$  are measurable with respect to the  $\sigma$ -field  $\mathscr{F}_{s_{i-1}\tau}$  where  $\tau = \max{(t_{v-1},t_{j-1})}$  and either (i)  $\Delta X_{uv}$  is measurable with respect to  $\mathscr{F}_{s_{i-1}\tau}$  and  $\Delta X_{ij}$  is independent of  $\mathscr{F}_{s_{i-1}\tau}$  or (ii)  $\Delta X_{uv}$  and  $\Delta X_{ij}$  are independent of  $\mathscr{F}_{s_{i-1}\tau}$  and also independent of each other.

LEMMA 3.3. Let  $f(s, t; \omega)$ ,  $g(s, t, \omega) \in L^2[[0, S] \times [0, T] \times \Omega]$  be functions which are  $\mathscr{F}_{st}$  measurable for all s and t. Then

- (i)  $E[|f(s_{i-1}, t_{j-1}; \omega)g(s_{u-1}, t_{v-1}; \omega)\Delta X_{ij}\Delta X_{uv}|] < \infty$ ,
- (ii)  $E[f(s_{i-1}, t_{j-1}; \omega)g(s_{u-1}, t_{v-1}; \omega)\Delta X_{ij}\Delta X_{uv}] = 0 \ i \neq u \ or \ t \neq v$
- (iii)  $E[f(s_{i-1}, t_{j-1}; \omega)g(s_{i-1}, t_{j-1}; \omega)(\Delta X_{ij})^2] = E[f(s_{i-1}, t_{j-1}; \omega)g(s_{i-1}, t_{j-1}; \omega)] \times (s_i s_{i-1})(t_j t_{j-1}).$

PROOF. These results follow using Lemma 2.1 and the remark preceding the lemma.

Let  $\mathscr C$  be the class of measurable functions  $f(s,\,t;\,\omega)$  defined on  $[0,\,S]\times [0,\,T]\times\Omega$  having the property that  $f(s,\,t;\,\omega)$  is measurable with respect to  $\mathscr F_{st}=\sigma\{X(u,\,v;\,\omega):(u,\,v)<(s,\,t)\}$  for all  $(s,\,t)\in[0,\,S]\times[0,\,T]$ . Let  $H(s_0,\,\ldots,\,s_m;\,t_0,\,\ldots,\,t_n)$   $0=s_0<\ldots< s_m=S,\,0=t_0<\ldots< t_n=T,$  be the class of functions  $f(s,\,t;\,\omega)\in\mathscr M=\mathscr C\cap L^2[[0,\,S]\times[0,\,T]\times\Omega]$  which have  $f(s,\,t;\,\omega)=f(s_{i-1},\,t_{j-1};\,\omega)$  for  $s_{i-1}\leq s< s_i,\,t_{j-1}\leq t< t_j,\,H=\bigcup H(s_0,\,\ldots,\,s_m;\,t_0,\,\ldots,\,t_n).$  The integral is first defined for functions in H as follows: if  $s_k\leq s\leq s_{k+1},\,t_p\leq t\leq t_{p+1},$ 

$$\begin{split} I(s,\,t,\,\omega;f) &= \sum_{i=1}^k \sum_{j=1}^p f(s_{i-1},\,t_{j-1};\,\omega) \Delta X[(s_i,\,t_j),\,(s_{i-1},\,t_{j-1})] \\ &+ \sum_{i=1}^k f(s_{i-1},\,t_p;\,\omega) \Delta X[(s_i,\,t),\,(s_{i-1},\,t_p)] \\ &+ \sum_{j=1}^p f(s_k,\,t_{j-1};\,\omega) \Delta X[(s_i,\,t_j),\,(s_k,\,t_{j-1})] \\ &+ f(s_k,\,t_p;\,\omega) \Delta X[(s,\,t),\,(s_k,\,t_p)] \;. \end{split}$$

This definition can easily be shown to be independent of the choice of partition used in defining the integral.

The integral so defined has the following properties. For  $f, g \in H$   $(L^2 = L^2[[0, S] \times [0, T] \times \Omega])$ 

- (i)  $E[I(s, t, \omega; f)] = 0$  for all  $(s, t) \in [0, S] \times [0, T]$ ,
- (ii)  $I(s, t, \omega; \alpha f + \beta g) = \alpha I(s, t, \omega; f) + \beta I(s, t, \omega; g),$
- (iii)  $I(s, t, \omega; 1) = \Delta X[(s, t), (0, 0); \omega] = X(s, t; \omega),$
- (iv)  $I(s, t, \omega; f)$  is a continuous function of (s, t) with probability one,
- (v)  $E[I(s, t, \omega; f)I(s, t, \omega; g)] = (f, g)_{L^2}$
- (vi)  $P[\omega : \sup_{0 \le s \le S; \ 0 \le t \le T} |I(s, t, \omega; f)| \ge b] \le (4/b^2) ||f(s, t; \omega)||_{L^2}^2$
- (vii) if  $f(s, t; \omega) = h(s, t; \omega)$  for all  $\omega \in \Omega_1$  where  $\Omega_1$  is any *P*-measurable subset of  $\Omega$ , then  $I(s, t, \omega; f) = I(s, t, \omega; h)$   $0 \le s \le S$ ,  $0 \le t \le T$  a.e. in  $\Omega_1$ .

These properties follow using standard techniques and the results of Lemmas 3.2 and 3.3.

As in the one-dimensional case, the integral so defined is a linear isometric operator from H to  $L^2(\Omega)$  and can be extended to a linear operator from the closure of H to  $L^2(\Omega)$ . It can be shown that the closure of the class of  $(s, t; \omega)$  step functions in  $\mathscr{G}$  includes all functions in  $\mathscr{M} = \mathscr{G} \cap L^2[[0, S] \times [0, T] \times \Omega]$ . The extension satisfies (i)—(vii) for  $f, g \in \mathscr{M}$  (Itô [7]).

LEMMA 3.4. Let  $f \in \mathcal{M}$  and let b > 0, then

- (i)  $E[\sup_{(s,t)\in[0,S]\times[0,T]}[I(s,t,\omega;f)]^2] \leq 16E[I(S,T,\omega;f)]^2$
- (ii)  $bP[\omega: \sup_{(s,t) \in [0,S] \times [0,T]} [I(s,t,\omega;f)]^2 \ge b] \le 16E[I(S,T,\omega;f)]^2$ .

PROOF. The lemma follows using Lemma 3.2.

THEOREM 6. Let  $Y(s, t; \omega) = I(s, t, \omega; f) = \int_0^t \int_0^s f(u, v; \omega) dX(u, v; \omega)$ . Then  $\{Y(s, t; \omega) : (s, t) \in [0, S] \times [0, T]\}$  is a martingale.

PROOF. Let  $\{f_n(s,t;\omega)\}$  be a sequence in H such that  $||f_n-f||_{L^2}\to 0$  as  $n\to\infty$ . Let  $Y_n(s,t;\omega)=I(s,t,\omega;f_n)$ .  $Y_n(s,t;\omega)$   $n=1,2,\cdots$ , is measurable with respect to  $\mathscr{F}_{st}$  for all  $(s,t)\in[0,S]\times[0,T]$  and  $||Y_n-Y||_{\Omega}\to 0$  as  $n\to\infty$ . Hence  $Y(s,t;\omega)$  is measurable with respect to  $\mathscr{F}_{st}$  for all  $(s,t)\in[0,S]\times[0,T]$ .

Let (s', t') < (s, t). Without loss of generality we may assume that  $f_n \in H(s_0, \dots, s_m; t_0, \dots, t_n)$  where  $s' = s_\alpha$ ,  $t' = t_\beta$ ,  $s = s_r$ ,  $t = t_p$  are partition points. We can write

$$\begin{split} E[Y_{n}(s, t; \omega) | \mathscr{F}_{s't'}] &= \sum_{i=1}^{\alpha} \sum_{j=1}^{\beta} E[f_{n}(s_{i-1}, t_{j-1}; \omega) \Delta X_{ij} | \mathscr{F}_{s't'}] \\ &+ \sum_{i=1}^{r} \sum_{j=\beta+1}^{p} E[f_{n}(s_{i-1}, t_{j-1}; \omega) \Delta X_{ij} | \mathscr{F}_{s't'}] \\ &+ \sum_{i=\alpha+1}^{r} \sum_{j=1}^{\beta} E[f_{n}(s_{i-1}, t_{j-1}; \omega) \Delta X_{ij} | \mathscr{F}_{s't'}] \\ &= I + II + III . \end{split}$$

In I both  $f_n(s_{i-1}, t_{j-1}; \omega)$  and  $\Delta X_{ij}$  are measurable with respect to  $\mathscr{F}_{s't'}$ , hence  $I = Y_n(s', t'; \omega)$  a.s. To show that II and III are zero, for each term in II and III we consider the  $\sigma$ -field  $\mathscr{F}_{\tau\tau}$  where  $\gamma = \max[s', s_{i-1}], \tau = \max[t', t_{j-1}]$ .  $\mathscr{F}_{s't'}$  is a sub- $\sigma$ -field of  $\mathscr{F}_{\tau\tau}$ ,  $f_n(s_{i-1}, t_{j-1}; \omega)$  is measurable with respect to  $\mathscr{F}_{\tau\tau}$  and  $\Delta X_{ij}$  is independent of  $\mathscr{F}_{\tau\tau}$ . Thus, for a typical term in II or III,  $E[f_n(s_{i-1}, t_{j-1}; \omega)\Delta X_{ij} | \mathscr{F}_{s't'}] = E\{E[f_n(s_{i-1}, t_{j-1}; \omega)\Delta X_{ij} | \mathscr{F}_{s't'}] = E[f_n(s_{i-1}, t_{j-1}; \omega)E[\Delta X_{ij}] | \mathscr{F}_{s't'}] = 0$  a.s. Hence  $E[Y_n(s, t; \omega) | \mathscr{F}_{s't'}] = Y_n(s', t'; \omega)$  a.s. Letting  $n \to \infty$  gives the desired result.

As an application of the stochastic integral we find a solution of the diffusion equation

(3.5) 
$$\Delta Y[(s, t), (0, 0); \omega] = \int_0^t \int_0^s m[u, v; Y(u, v; \omega)] du dv + \int_0^t \int_0^s \sigma[u, v; Y(u, v; \omega)] dX(u, v; \omega).$$

The following hypotheses are made.

- (i)  $m(\cdot, \cdot; \cdot)$  and  $\sigma(\cdot, \cdot; \cdot)$  are Baire functions of  $(s, t; \eta)$  for  $(s, t) \in [0, S] \times [0, T], -\infty < \eta < \infty$ .
- (ii) There is a constant K such that  $|m(s, t; \eta)| \le K(1 + \eta^2)^{\frac{1}{2}}$ ,  $0 \le \sigma(s, t; \eta) \le K(1 + \eta^2)^{\frac{1}{2}}$ .
- (iii)  $m(\cdot, \cdot; \cdot)$  and  $\sigma(\cdot, \cdot; \cdot)$  satisfy a uniform Lipschitz condition in  $\eta$  i.e.  $|m(s, t; \eta_2) m(s, t, \eta_1)| \le K|\eta_2 \eta_1|$ ,  $|\sigma(s, t; \eta_2) \sigma(s, t; \eta_1)| \le K|\eta_2 \eta_1|$  where K is independent of s, t and  $\eta$ .

Assuming hypotheses (i), (ii) and (iii) it is possible using Lemma 3.4 and an iterative process as in the one-dimensional case (Doob [6] page 277–282) to construct a process  $Y(s, t; \omega)$  having the following properties.

- (A) The  $Y(s, t; \omega)$  sample functions are almost all continuous in  $[0, S] \times [0, T]$ .
- (B)  $\int_0^T \int_0^s E[Y(s, t; \omega)]^2 ds dt < \infty$ .
- (C) For each  $(s_0, t_0) \in [0, S] \times [0, T]$ ,  $Y(s_0, t_0; \omega)$  is independent of  $\Delta X[(s_i, t_j), (s_{i-1}, t_{j-1}); \omega]$  for  $s_{i-1} \ge s_0$  or  $t_{j-1} \ge t_0$ .
- (D)  $Y(s, t; \omega)$  is measurable with respect to  $\mathscr{B}_{st} = \sigma\{X(u, v; \omega) : (u, v) < (s, t)\} \vee \sigma\{Y(0, 0; \omega), Y(u, 0; \omega), Y(0, v; \omega) : (u, v) < (s, t)\}.$
- (E) For each  $(s, t) \in [0, S] \times [0, T]$ ,  $Y(s, t; \omega)$  satisfies (3.5). The process  $Y(s, t; \omega)$  is essentially uniquely determined.
- **4.** Sample variation. Let  $\{X(s, t: \omega) : (s, t) \in [0, S] \times [0, T]\}$  be a Gaussian stochastic process of real-valued random variables with mean function m(s, t) and covariance function  $R(s_1, t_1; s_2, t_2) = E[X(s_1, t_1; \omega)X(s_2, t_2; \omega)] m(s_1, t_1)m(s_2, t_2)$  satisfying the following conditions:
  - (i) m(s, t) has bounded second order partial derivatives on  $[0, S] \times [0, T]$ .
- (ii)  $R(s_1, t_1; s_2, t_2)$  is continuous for  $0 \le s_1, s_2 \le S$ ,  $0 \le t_1, t_2 \le T$  and has uniformly bounded third order partial derivatives for  $s_1 \ne s_2, t_1 \ne t_2$ .

Define

$$D_{-}^{+}(s,t) = \lim_{u \to s^{+}; v \to t^{-}} \frac{R(s,t;s,t) - R(s,t;u,t) - R(s,t;s,v) + R(s,t;u,v)}{(s-u)(t-v)}.$$

Similarly define  $D_{+}^{-}(s, t)$ ,  $D_{+}^{+}(s, t)$  and  $D_{-}^{-}(s, t)$ . Let  $f(s, t) = D_{-}^{-}(s, t) - D_{-}^{+}(s, t) - D_{+}^{-}(s, t) + D_{+}^{+}(s, t)$ .

*Note*. Unless otherwise indicated the sums in this section will be on  $j, k = 1, \dots, 2^n$  and the indices will be omitted.

THEOREM 7. If  $X(s, t; \omega)$  satisfies the above conditions then with probability one,

(4.1) 
$$\lim_{n\to\infty} \sum \left[\Delta X[(jS2^{-n}, kT2^{-n}), ((j-1)S2^{-n}, (k-1)T2^{-n}); \omega]\right]^2$$
$$= \int_0^T \int_0^S f(s, t) \, ds \, dt.$$

PROOF. The proof uses Taylor's series expansions and follows the methods used by Baxter [1].

COROLLARY 1. If  $\{X(s, t; \omega) : (s, t) \in [0, S] \times [0, T]\}$  satisfies the assumptions of Theorem 7 and if

$$R(s_1, t_1; s_2, t_2) = u(s_1, t_1)v(s_2, t_2) \qquad s_1 \leq s_2, \quad t_1 \leq t_2$$

$$= u(s_1, t_2)v(s_2, t_1) \qquad s_1 \leq s_2, \quad t_1 \geq t_2$$

$$= u(s_2, t_1)v(s_1, t_2) \qquad s_1 \geq s_2, \quad t_1 \leq t_2$$

$$= u(s_2, t_2)v(s_1, t_1) \qquad s_1 \geq s_2, \quad t_1 \geq t_2$$

then with probability one,

$$\begin{split} \lim_{n\to\infty} \sum \left[ \Delta X [(jS2^{-n}, kT2^{-n}), ((j-1)S2^{-n}, (k-1)T2^{-n}); \omega] \right]^2 \\ &= \int_0^T \int_0^S \left[ u(s, t) \frac{\partial^2 v(s, t)}{\partial s \, \partial t} - \frac{\partial u(s, t)}{\partial s} \frac{\partial v(s, t)}{\partial t} \right. \\ &- \frac{\partial u(s, t)}{\partial t} \frac{\partial v(s, t)}{\partial s} + v(s, t) \frac{\partial^2 u(s, t)}{\partial s \, \partial t} \right] ds \, dt \, . \end{split}$$

COROLLARY 2. Let  $\{X(s,t;\omega):(s,t)\in[a,b]\times[c,d]\}$  be a Gaussian process satisfying the conditions of Theorem 7. Then with probability one,

(4.2) 
$$\lim_{n\to\infty} \sum \left[ \Delta X [(a+j(c-a)2^{-n}, b+k(d-b)2^{-n}), \times (a+(j-1)(c-a)2^{-n}, b+(k-1)(d-b)2^{-n}); \omega] \right]^{2}$$

$$= \int_{b}^{d} \int_{a}^{c} f(s,t) \, ds \, dt.$$

COROLLARY 3. Let  $\{X(s, t; \omega) : (s, t) \in [0, S] \times [0, T]\}$  be a Gaussian process with mean function zero and covariance function  $R(s_1, t_1; s_2, t_2) = \min(s_1, s_2) \min(t_1, t_2)$ . Then with probability one,

$$\lim_{n\to\infty} \sum \left[\Delta X[(jS2^{-n}, kT2^{-n}), ((j-1)S2^{-n}, (k-1)T2^{-n}); \omega]\right]^2 = ST.$$

THEOREM 8. Let  $\{Y(s, t; \omega) : (s, t) \in [0, S] \times [0, T]\}$  be a solution of the diffusion equation (3.5), then with probability one,

(4.3) 
$$\lim_{n\to\infty} \sum \left[ \Delta Y[(jS2^{-n}, kT2^{-n}), ((j-1)S2^{-n}, (k-1)T2^{-n}); \omega] \right]^2$$
$$= \int_0^T \int_0^S \sigma^2[u, v; Y(u, v; \omega)] du dv.$$

PROOF. Let  $K = kT2^{-n}$ ,  $K - 1 = (k - 1)T2^{-n}$ ,  $J = jS2^{-n}$ ,  $J - 1 = (j - 1)S2^{-n}$ , then

$$\sum \left[\Delta Y[(J,K), (J-1,K-1)]\right]^{2} = \sum \left[\int_{K-1}^{K} \int_{J-1}^{J} m[u,v; Y(u,v)] du dv\right]^{2} + 2 \sum \left[\int_{K-1}^{K} \int_{J-1}^{J} m[u,v; Y(u,v)] du dv\right] \left[\int_{K-1}^{K} \int_{J-1}^{J} \sigma[u,v; Y(u,v)] dX(u,v)\right] + \sum \left[\int_{K-1}^{K} \int_{J-1}^{J} \sigma[u,v; Y(u,v)] dX(u,v)\right]^{2} = A_{n} + 2B_{n} + C_{n}.$$

It can be shown that  $A_n \to 0$ ,  $|B_n| \to 0$  a.s. as  $n \to \infty$ . Hence the existence of the limit on the left-hand side is an event whose probability is the same for all functions  $m(s, t; \eta)$  satisfying (i), (ii) and (iii) and the value of the limit when it exists is independent of  $m(s, t; \eta)$ . We shall use  $m(s, t; \eta) \equiv 0$  in our calculations.

The sum of squares of the increments of the process can be decomposed in the following manner.

$$\sum \left[ \int_{K-1}^{K} \int_{J-1}^{J} \sigma[u, v; Y(u, v)] dX(u, v) \right]^{2}$$

$$= \frac{1}{d} \sum_{L} \left[ \int_{K-1}^{K} \int_{J-1}^{J} F(\sigma, Y, u, v, K, J) dX(u, v) \right]^{2}$$

where  $F(\sigma, Y, u, v, K, J) = \{\sigma[u, v; Y(u, v)] - \sigma[u, K-1; Y(u, K-1)]\} + \{\sigma[u, v; Y(u, v)] - \sigma[J-1, v; Y(J-1, v)]\} + \{\sigma[u, K-1; Y(u, K-1)] - \sigma[J-1, K-1; Y(J-1, K-1)]\} + \{\sigma[J-1, v; Y(J-1, v)] - \sigma[J-1, K-1; Y(J-1, K-1)]\} + 2\sigma[J-1, K-1; Y(J-1, K-1)].$  When expanded this sum of squares gives fifteen terms. It can be shown, using methods similar to those employed by Berman [2], that the absolute values of all terms in the sum converge to zero almost surely as  $n \to \infty$  with the exception of the term

$$\begin{split} S_n &= \sum \left[ \int_{K-1}^K \int_{J-1}^J \sigma[J-1,K-1;\ Y(J-1,K-1)]\ dX(u,v) \right]^2 \\ &= \sum \sigma^2[(j-1)S2^{-n},(k-1)T2^{-n};\ Y((j-1)S2^{-n},(k-1)T2^{-n})] \\ &\times \left[ \Delta X[(jS2^{-n},kT2^{-n}),((j-1)S2^{-n},(k-1)T2^{-n})] \right]^2. \end{split}$$

Let m be an arbitrary fixed positive integer. Then for n > m,  $\Gamma = \{(s,t): (\alpha-1)S2^{-m} \le s \le \alpha S2^{-m}, (\beta-1)T2^{-m} \le t \le \beta T2^{-m}\}, \sum_{\alpha,\beta=1}^{2^m} \min_{\Gamma} \sigma^2[s,t; Y(s,t)] \sum_{j,k=1}^{2^{n-m}} [\Delta X[(\{(\alpha-1)2^{-m}+j2^{-n}\}S,\{(\beta-1)2^{-m}+k2^{-n}\}T),(\{(\alpha-1)2^{-m}+(j-1)2^{-n}\}S,\{(\beta-1)2^{-m}+(k-1)2^{-n}\}T)]]^2 \le S_n \le \sum_{\alpha,\beta=1}^{2^m} \max_{\Gamma} \sigma^2[s,t; Y(s,t)] \sum_{j,k=1}^{2^{n-m}} [\Delta X[(\{(\alpha-1)2^{-m}+j2^{-n}\}S,\{(\beta-1)2^{-m}+k2^{-n}\}T),(\{(\alpha-1)2^{-m}+(j-1)2^{-n}\}S,\{(\beta-1)2^{-m}+(k-1)2^{-n}\}T)]]^2$ . Applying Corollary 2 of Theorem 7, we see that the lower bound converges a.s. as  $n \to \infty$  to

$$\sum_{\alpha,\beta=1}^{2^{m}} \min_{\Gamma} \sigma^{2}[s, t; Y(s, t)][\alpha 2^{-m} - (\alpha - 1)2^{-m}]S[\beta 2^{-m} - (\beta - 1)2^{-m}]T$$

$$= \sum_{\alpha,\beta=1}^{2^{m}} \min_{\Gamma} \sigma^{2}[s, t; Y(s, t)](S2^{-m})(T2^{-m})$$

and the upper bound converges a.s. as  $n \to \infty$  to  $\sum_{\alpha,\beta=1}^{2m} \max_{\Gamma} \sigma^2[s, t; Y(s, t)] \times (S2^{-m})(T2^{-m})$ . As  $m \to \infty$  these bounds converge a.s. to the common limit  $\int_0^T \int_0^s \sigma^2[u, v; Y(u, v)] du dv$  since  $\sigma^2[s, t; Y(s, t; \omega)]$  is a continuous function of (s, t) with probability one.

## REFERENCES

- [1] BAXTER, G. (1956). A strong limit theorem for Gaussian processes. *Proc. Amer. Math. Soc.* 7 522-27.
- [2] Berman, S. M. (1965). Sign invariant random variables and stochastic processes with sign invariant increments. Trans. Amer. Math. Soc. 119 216-43.
- [3] ČENCOV, N. N. (1956). Wiener random fields depending on several parameters. Dokl. Akad. Nauk SSSR 106 607-09.
- [4] Ciestelski, Z. (1966). Lectures on Brownian motion, heat conduction and potential theory.

  Math. Institute, Aarhus Univ., Denmark.

- [5] Delporte, J. (1966). Fonctions aléatoires de deux variables presque surement à échantillons continus sur un domaine rectangulaire borné. Z. Wahrscheinlichkeitstheorie und Verw. Gebiete 6 181-205.
- [6] DOOB, J. L. (1953). Stochastic Processes. Wiley, New York.
- [7] ITô, K. (1944). Stochastic integral. Proc. Imperial Acad. Tokyo 20 519-24.
- [8] Itô, K. and McKean, H. (1965). Diffusion Processes and Their Sample Paths. Academic Press, New York.
- [9] LÉVY, P. (1965). Processus Stochastiques et Mouvement Brownian. Gauthier-Villars, Paris.
- [10] LOÈVE, M. (1963). Probability Theory, 3rd ed. Van Nostrand, Princeton.
- [11] PARK, W. J. (1970). A multi-parameter Gaussian process. Ann. Math. Statist. 41 1582-95.
- [12] Yeh, J. (1960). Wiener measure in a space of functions of two variables. Trans. Amer. Math. Soc. 95 433-50.