

A NOTE ON THE CLASSICAL OCCUPANCY PROBLEM¹

BY C. J. PARK

University of Wisconsin

Assume that n balls are randomly distributed into N equiprobable cells. The ball is presumed to have probability p , $0 < p < 1$ of staying in the cell and $(1 - p)$ of falling through. Let S_0 denote the number of empty cells. In this note we establish the asymptotic normality of S_0 as n and N tend to infinity so that $np/N \rightarrow c > 0$, $np/N^{\frac{1}{2}} \rightarrow \infty$ and $n/N \rightarrow 0$, or $3np/N - \log N \rightarrow -\infty$ and $n/N \rightarrow \infty$. We accomplish this by estimating the factorial cumulants of S_0 .

1. Introduction and summary. Assume that n balls are randomly distributed into N cells with equal probabilities, i.e. each ball has probability $1/N$ of falling into i th cell, $i = 1, 2, \dots, N$. The ball is presumed to have probability p , $0 < p < 1$ of staying in the cell and $(1 - p)$ of falling through. Let S_0 denote the number of empty cells. In this note we will show that the asymptotic distribution of S_0 is normal as n and N tend to infinity with one of the following conditions being satisfied:

- (i) $np/N \rightarrow c$, $0 < c < \infty$,
- (ii) $n/N \rightarrow 0$ and $np/N^{\frac{1}{2}} \rightarrow \infty$,
- (iii) $n/N \rightarrow \infty$ and $3np/N - \log N \rightarrow -\infty$.

We establish the asymptotic normality of S_0 by estimating the factorial cumulants of S_0 and utilizing the similar method given by Harris and Park [4]. For the special case when $p = 1$, the asymptotic distribution of S_0 has been extensively studied (see for example [5], [6], [8] and [9]). Harkness [3] gives numerous examples of situations for which the distribution of S_0 can be applied (see also the references therein).

2. Asymptotic normality of S_0 . The probability distribution of S_0 is well known (see for example [3]) and given by

$$P[S_0 = x; n, N, p] = \binom{N}{x} \sum_{k=0}^{N-x} (-1)^k \binom{n-k}{k} \left(1 - \frac{(k+x)p}{N}\right)^n,$$

$x = 0, 1, \dots, N.$

The m th factorial moment of S_0 is given by

$$(1) \quad \mu_{[m]} = N^{(m)} \left(1 - \frac{mp}{N}\right)^n, \quad m = 0, 1, \dots,$$

where $N^{(m)} = N(N-1) \dots (N-m+1)$. Consequently the factorial moment

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generating function can be written as,

$$(2) \quad \varphi_{n,N}(t) = \sum_{m=0}^{\infty} \frac{\mu[m]}{m!} t^m = \sum_{m=0}^N \binom{N}{m} \left(1 - \frac{mp}{N}\right)^n t^m .$$

Let $K_{n,N}(t)$ be the corresponding factorial cumulant generating function, then

$$(3) \quad K_{n,N}(t) = \log \varphi_{n,N}(t) = \sum_{m=1}^{\infty} k_{[m]} \frac{t^m}{m!} ,$$

where $k_{[m]} = k_{[m]}(n, N)$ is the m th factorial cumulant of S_0 . The factorial cumulants are related to the cumulants in the same way as the factorial moments are related to the moments, that is,

$$(4) \quad k_m = \sum_{j=1}^m \alpha_{j,m} k_{[j]} ,$$

where $\alpha_{j,m}$ are the Stirling numbers of the second kind. To establish the asymptotic normality of S_0 , we will show that for $m > 2$

$$k_m k_2^{-m/2} \rightarrow 0$$

as n and N tend to infinity. Now we introduce the following theorem.

THEOREM 1. *The m th cumulant of S_0 ,*

$$k_m = O(N) \text{ as } N \rightarrow \infty , \quad \text{for } m = 1, 2, \dots .$$

PROOF. Let

$$P(t) = (1 + t)^N = \sum_{\nu=0}^N \binom{N}{\nu} t^\nu ,$$

a polynomial of degree N with every root -1 . Then let

$$\begin{aligned} P_1(t) &= P(t) - p \frac{t}{N} P'(t) \\ &= \sum_{\nu=0}^N \binom{N}{\nu} \left(1 - p \frac{\nu}{N}\right) t^\nu . \end{aligned}$$

For $\nu \geq 1$, define

$$P_{\nu+1}(t) = P_\nu(t) - \left(p \frac{t}{N}\right) P'_\nu(t) ;$$

then we readily see that

$$P_n(t) = \sum_{\nu=0}^N \binom{N}{\nu} \left(1 - p \frac{\nu}{N}\right)^n t^\nu = \varphi_{n,N}(t)$$

where $\varphi_{n,N}(t)$ is defined in (2). Now define

$$Q_n(t) = \frac{N}{p} P_{n+1}(t) = \frac{N}{p} P_n(t) - t P'_n(t) .$$

Then it can be verified (cf. Lemma 1 and Lemma 2 in [4]) that for every $n \geq 1$ $Q_n(t)$ has N real roots and all of its roots ≤ -1 because $P_n(t)$ is a polynomial

of degree N and has N real roots ≤ -1 . Hence, $N^{-1} \log P_n(t) = N^{-1} \log \varphi_{n,N}(t) = N^{-1}K_{n,N}(t)$ is analytic in $|t| < 1$. Thus for $|t| < 1$,

$$\begin{aligned} \operatorname{Re} (N^{-1} \log P_n(t)) &= N^{-1} \log |P_n(t)| \\ &\leq N^{-1} \log \sum_{\nu=0}^N \binom{N}{\nu} |t|^\nu = \log (1 + |t|) \leq \log 2. \end{aligned}$$

We can now apply a well-known theorem of Carathéodory (see [1], [2] and [7]), that is, if $f(z) = \sum_{j=1}^\infty \alpha_j z^j$, $|z| < 1$ and $\operatorname{Re} [f(z)] \leq 1$ for $|z| < 1$, then $|\alpha_j| < 2$ for all j . Thus, since

$$K_{n,N}(t) = \sum_{m=1}^\infty k_{[m]} t^m / m!,$$

we have

$$|k_{[m]}| \leq Nm! \log 4;$$

thus the theorem follows from (4).

Now from (1), we have

$$E(S_0) = \mu(S_0) = N \left(1 - \frac{p}{N}\right)^n,$$

$$\begin{aligned} \operatorname{Var} (S_0) &= \sigma^2(S_0) \\ &= N^2 \left(\left(1 - \frac{2p}{N}\right)^n - \left(1 - \frac{p}{N}\right)^{2n} \right) + N \left(\left(1 - \frac{p}{N}\right)^n - \left(1 - \frac{2p}{N}\right)^n \right). \end{aligned}$$

We now establish the limiting distribution of

$$S_0^* = (S_0 - \mu(S_0)) / \sigma(S_0).$$

THEOREM 2. *If one of the conditions (i)—(iii) in Section 1 is satisfied, the limiting distribution of S_0^* , as n and N tend to infinity, is the standard normal distribution.*

PROOF. To establish the theorem it suffices to show that $k_2^{-m/2} \rightarrow 0$ for $m > 2$. From Theorem 1, this is equivalent to showing that $Nk_2^{-3/2} \rightarrow 0$. Let $n/N = \alpha(n, N)$ and since $\alpha(n, N) = o(N)$, we have

$$k_2 = \sigma^2(S_0) = Ne^{-\alpha p} (1 - e^{-\alpha p} - \alpha p e^{-\alpha p}) + O(\psi(\alpha))$$

where $\psi(\alpha) = \max(\alpha, \alpha^2)$. Thus, the conclusion holds for $\alpha \rightarrow 0$ as n and N tend to infinity with $np/N \rightarrow \infty$, and for $\alpha \rightarrow \infty$ as n and N tend to infinity with $3np/N - \log N \rightarrow -\infty$. The conclusion clearly holds if α has a positive limit as n and N tend to infinity.

REMARK. The probability distribution of S_0 can be written as

$$P[S_0 = x; n, N, p] = \sum_{t=0}^n P(S_0 = x; t, N, 1) \binom{n}{t} p^t (1 - p)^{n-t}$$

where $P(S_0 = x; t, N, 1)$ denotes the probability distribution of the number of empty cells when t balls are randomly distributed into N equi-probable cells and $p = 1$.

The limiting distribution of the number of cells occupied by i balls, $i \neq 0$, is under investigation and we hope to report the result in the future.

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MATH RESEARCH CENTER
UNIVERSITY OF WISCONSIN
MADISON, WISCONSIN 53706