

LAW OF LARGE NUMBERS IN THE SUPREMUM NORM FOR A CHEMICAL REACTION WITH DIFFUSION¹

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A space-time jump Markov process, modeling a chemical reaction with diffusion, is compared in the supremum norm to the usual model, the solution to a partial differential equation. Conditions are given which imply the deviation converges in probability to 0 uniformly on bounded time intervals. Estimates reflecting underlying large deviation behavior are obtained.

1. Introduction. In Arnold and Theodosopulu (1980) a stochastic model of a chemical reaction with diffusion was constructed and compared with the usual deterministic model, the solution of a partial differential equation. This model has been studied further by Kotelenez [(1982a), (1982b), (1986a), (1986b), (1987), (1988)]. The model, a space-time jump Markov process, is constructed by dividing $[0, 1]^q$ into N^q congruent cells of volume N^{-q} and modeling the concentration within a cell by a density-dependent birth and death process rescaled by l^{-1} , where l is a parameter proportional to the initial number of particles in a cell [see Kurtz (1971) or Ethier and Kurtz (1986)]. Also, particles diffuse symmetrically to neighboring cells with a jump rate proportional to N^2 . This couples the cell reactors and extends the model of Kurtz (1971) to the spatially inhomogeneous case.

In Kotelenez (1986a), for a linear reaction (branching random walks) with $q = 1$, it was shown that $N^2/l \rightarrow 0$ is sufficient to prove a law of large numbers in $L_2([0, 1])$. In Blount [(1987), (1991)] this was improved by only requiring $l \rightarrow \infty$ as $N \rightarrow \infty$. In Kotelenez (1988), for a nonlinear reaction, a law of large numbers was shown to hold in a space of distributions rather than $L_2([0, 1]^q)$. In this paper we prove the law of large numbers holds in the supremum norm for any dimension q if $\log N/l \rightarrow 0$ as $N \rightarrow \infty$. The reaction may be linear or nonlinear. We believe our result is necessary as well as sufficient but have not proved this. However, we do give a nontrivial example, Example 4.19, which suggests the conditions are necessary. We obtain an estimate, (4.18), which reflects underlying large deviation behavior.

For simplicity we prove the result for $q = 1$ with periodic boundary conditions but remark at the end of the proof on the minor notational changes needed for extending the result to $q > 1$.

References to related work may be found in the papers of Kotelenez. We have only stated previous results that compare directly with this paper.

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Note that within a proof we may use the same symbol for different constants that depend on the same parameters so long as these are not N or l . Theorems, lemmas and numbered equations are considered on the same numbering system.

2. The deterministic model. Let $R(x) = b(x) - d(x) = \sum_{i=0}^m c_i x^i$ be a polynomial for $x \in \mathbf{R}$, with $c_m < 0$ and $b(x), d(x)$ being polynomials of degree less than or equal to m with nonnegative coefficients. Assume $d(0) = 0$. For one reactant in the unit interval with periodic boundary conditions the concentration is given by $\psi(t, r)$, where, for $r \in [0, 1]$ and $t \geq 0$,

$$(2.1) \quad \begin{aligned} \frac{\partial}{\partial t} \psi(t, r) &= \Delta \psi(t, r) + R(\psi(t, r)), \\ \psi(t, 0) &= \psi(t, 1), \\ 0 &\leq \psi(0, r) < \rho < \infty. \end{aligned}$$

Δ denotes the Laplacian. Let $\psi(t) = \psi(t, \cdot)$ and note that $R(x) < 0$ for all large x . We assume ρ large enough so that $R(x) < 0$ for $x > \rho$ and that $\psi(0) \in C^3([0, 1])$, functions on $[0, 1]$ with three continuous derivatives and norm given by $\sum_{i=0}^3 \|\psi^{(i)}\|_\infty$ where $\|\cdot\|_\infty$ denotes the supremum norm. In Kotelenetz (1986b) it is shown that (2.1) has a unique mild solution in $C([0, \infty); C^3([0, 1]))$ and that $0 \leq \psi(t) < \rho$ for all $t > 0$.

Let H^N denote the real-valued step functions on $[0, 1]$ that are constant on the intervals $[kN^{-1}, (k+1)N^{-1})$, $0 \leq k \leq N-1$, where $N \geq 1$ is an integer; we extend functions in H^N to be periodic with period 1. For $f \in H^N$, let

$$\nabla_N^\pm f(r) = N(f(r \pm N^{-1}) - f(r))$$

and

$$\begin{aligned} \Delta_N f(r) &= -\nabla_N^+ \nabla_N^- f(r) \\ &= -\nabla_N^- \nabla_N^+ f(r) = N^2[f(r + N^{-1}) - 2f(r) + f(r - N^{-1})]. \end{aligned}$$

Rather than work directly with ψ , we use a spatially discretized version given by the solution of the differential equation

$$(2.2) \quad \begin{aligned} \frac{\partial \psi^N(t, r)}{\partial t} &= \Delta_N \psi^N(t, r) + R(\psi^N(t, r)), \\ \psi^N(t, 0) &= \psi^N(t, 1), \\ \psi^N(0, r) &= N \int_{kN^{-1}}^{(k+1)N^{-1}} \psi(0, u) du \quad \text{for } r \in [kN^{-1}, (k+1)N^{-1}). \end{aligned}$$

Set $\psi^N(t) = \psi^N(t, \cdot)$. From Kotelenetz (1986b), we have

$$(2.3) \quad 0 \leq \psi^N(t) < \rho \quad \text{for all } t$$

and

$$(2.4) \quad \sup_{[0, T]} \|\psi^N(t) - \psi(t)\|_\infty \leq C(T, R, \psi(0)) N^{-1} \quad \text{for } T > 0.$$

If N is an odd integer with $0 \leq m \leq N - 1$ and m even, let $\varphi_{0,N} \equiv 1$ and define $\varphi_{m,N}(r) = \sqrt{2} \cos(\pi mkN^{-1})$ and $\psi_{m,N}(r) = \sqrt{2} \sin(\pi mkN^{-1})$ for $r \in [kN^{-1}, (k + 1)N^{-1})$. $\{\varphi_{m,N}, \psi_{m,N}\}$ are eigenvectors of Δ_N with eigenvalues defined by $-\beta_{m,N} = -2N^2(1 - \cos(\pi mN^{-1})) \leq 0$. For $f, g \in H^N$, the $L^2([0, 1])$ inner product is given by $\langle f, g \rangle = \sum_{k=0}^{N-1} f(kN^{-1})g(kN^{-1})N^{-1}$. $\{\varphi_{m,N}, \psi_{m,N}\}$ form an orthonormal basis for $(H^N, \langle \cdot, \cdot \rangle)$. If N is even, we need the additional eigenfunction $\varphi_{N,N}(r) = \cos \pi k$ for $r \in [kN^{-1}, (k + 1)N^{-1})$.

Let $T_N(t) = \exp(\Delta_N t)$ denote the semigroup on H^N generated by Δ_N . If $f \in H^N$, we have

$$(2.5) \quad T_N(t) f = \sum e^{-\beta_{m,N} t} (\langle f, \varphi_{m,N} \rangle \varphi_{m,N} + \langle f, \psi_{m,N} \rangle \psi_{m,N}).$$

Also note that for $f, g \in H^N$,

$$\langle \nabla_N^+ f, g \rangle = \langle f, \nabla_N^- g \rangle, \quad T_N(t) \Delta_N f = \Delta_N T_N(t) f,$$

and $\Delta_N, T_N(t)$ are self-adjoint on $(H^N, \langle \cdot, \cdot \rangle)$.

3. The stochastic model. Given $N \geq 1$ and a parameter $l > 0$, let

$$n(t) = (n_0(t), \dots, n_{N-1}(t))$$

denote the jump Markov process with integer-valued components and transition rates given by

$$(3.1) \quad \begin{aligned} (n_k, n_{k+1}) &\rightarrow (n_k - 1, n_{k+1} + 1) \quad \text{at rate } N^2 n_k, \\ (n_{k-1}, n_k) &\rightarrow (n_{k-1} + 1, n_k - 1) \quad \text{at rate } N^2 n_k, \\ n_k &\rightarrow n_k + 1 \quad \text{at rate } lb(n_k l^{-1}), \\ n_k &\rightarrow n_k - 1 \quad \text{at rate } ld(n_k l^{-1}). \end{aligned}$$

Here we set $n_N = n_0$ and $n_{-1} = n_{N-1}$ and b and d are the polynomials used to define (2.1). We view $n_k(t)$ as the number of particles in the k th cell at time t , where the cells are arranged on the unit circle. The last two jump rates reflect births or deaths in a cell and the first two reflect coupling of cells through diffusion. Particles diffuse on the circle according to simple random walks with jump rate $2N^2$ and particles are produced or removed within cells. Our subsequent assumptions will imply that l is proportional to the initial number of particles in each cell and that initially there are of order Nl particles distributed on the circle. We take $n(t)$ to be right-continuous with left limits and defined on some probability space.

Let $F_t^{N,l}$ denote the completion of the σ -algebra $\sigma(n(s): s \leq t)$ and let $\delta n(t) = n(t) - n(t -)$ denote the jump at time t . If τ is an $F_t^{N,l}$ stopping time such that

$$\sup_{[0, T]} \sup_k I_{\{\tau > 0\}} n_k(t \wedge \tau) \leq M(T, N, l) < \infty$$

for all $T > 0$, then as in Blount (1991) or Kotelenez (1988) we have:

3.2 LEMMA. *The following are $F_t^{N,l}$ martingales:*

$$\begin{aligned}
 \text{(a)} \quad & n_k(t \wedge \tau) - N^2 \int_0^{t \wedge \tau} (n_{k-1}(s) - 2n_k(s) + n_{k+1}(s)) ds \\
 & - \int_0^{t \wedge \tau} lR(n_k(s)l^{-1}) ds. \\
 \text{(b)} \quad & \sum_{s \leq t \wedge \tau} (\delta n_k(s))^2 - N^2 \int_0^{t \wedge \tau} (n_{k-1}(s) + 2n_k(s) + n_{k+1}(s)) ds \\
 & - \int_0^{t \wedge \tau} l(b(n_k(s)l^{-1}) + d(n_k(s)l^{-1})) ds. \\
 \text{(c)} \quad & \sum_{s \leq t \wedge \tau} (\delta n_{k+1}(s))(\delta n_k(s)) + N^2 \int_0^{t \wedge \tau} (n_{k+1}(s) + n_k(s)) ds.
 \end{aligned}$$

Now we define the stochastic analogue of (2.1). Let

$$X^N(t, r) = n_k(t)l^{-1} \quad \text{for } r \in [kN^{-1}, (k+1)N^{-1}).$$

$X^N(t, r)$ depends on N and l [as does $n(t)$], but we suppress the l in the superscript. Now let $X^N(t) = X^N(t, \cdot)$, and note that X^N is an H^N -valued Markov process. Using Lemma 3.2(a), we can write

$$(3.3) \quad X^N(t) = X^N(0) + \int_0^t \Delta_N X^N(s) ds + \int_0^t R(X^N(s)) ds + Z^N(t),$$

where $Z^N(t \wedge \tau)$ is an $F_t^{N,l}$ martingale with values in H^N for τ as in Lemma 3.2.

4. The law of large numbers. In this section we prove the following result.

4.1 THEOREM. *Assume:*

- (i) $\|X^N(0) - \psi(0)\|_\infty \rightarrow 0$ in probability.
- (ii) $l = l(N)$ satisfies $\log N/l \rightarrow 0$ as $N \rightarrow \infty$.

Then $\sup_{[0, T]} \|X^N(t) - \psi(t)\|_\infty \rightarrow 0$ in probability for any $T > 0$.

Before giving the proof we need some preliminary results and discussion.

4.2 LEMMA. $T_N(t)$ is a positive contraction semigroup on $(H^N, \|\cdot\|_\infty)$.

PROOF. If f is constant, then $T_N(t)f = f$, since $\Delta_N f = 0$. If $f \geq 0$, then $(2N^2 + \Delta_N)f \geq 0$. Thus $T_N(t)f = \exp(-2N^2t)\exp[(2N^2 + \Delta_N)t]f \geq 0$. These two facts imply the result. \square

4.3 LEMMA. *Let $f = NI_{[k/N, (k+1)/N]}$. Then*

$$\langle (\nabla_N^+ T_N(t) f)^2 + (\nabla_N^- T_N(t) f)^2 + (T_N(t) f)^2, 1 \rangle \leq h_N(t),$$

where $\int_0^t h_N(s) ds \leq CN + t$.

PROOF. Using the observations at the end of Section 2, we have

$$\begin{aligned} & \langle (\nabla_N^\pm T_N(t) f)^2, 1 \rangle \\ &= -\langle T_N(2t) f, \Delta_N f \rangle \\ &= \sum_m [\langle f, \varphi_{m,N} \rangle^2 + \langle f, \psi_{m,N} \rangle^2] e^{-2\beta_{m,N} t} \beta_{m,N} \\ &= \sum_m (\varphi_{m,N}^2 (kN^{-1}) + \psi_{m,N}^2 (kN^{-1})) e^{-2\beta_{m,N} t} \beta_{m,N} \\ &\leq 2 \sum_m e^{-2\beta_{m,N} t} \beta_{m,N}. \end{aligned}$$

Similarly $\langle (T_N(t) f)^2, 1 \rangle \leq 1 + 2 \sum_{m>0} e^{-2\beta_{m,N} t}$. The result then holds for $h_N(t) = 1 + 4 \sum_{m>0} e^{-2\beta_{m,N} t} (\beta_{m,N} + 1)$, since $\beta_{0,N} = 0$ and $\beta_{m,N} > cm^2$ for $m > 0$ and $c > 0$, where c is independent of m and N . \square

4.4 LEMMA. *Let $m(t)$ be a bounded martingale of finite variation defined on $[t_0, t_1]$ with $m(t_0) = 0$ and satisfying:*

- (i) *m is right-continuous with left limits.*
- (ii) *$|\delta m(t)| \leq 1$ for $t_0 \leq t \leq t_1$.*
- (iii) *$\sum_{t_0 \leq s \leq t} (\delta m(s))^2 - \int_{t_0}^t g(s) ds$ is a mean 0 martingale with $0 \leq g(s) \leq h(s)$, where $h(s)$ is a bounded deterministic function and $g(s)$ is $F_t^{N,l}$ adapted. Then $E \exp(m(t_1)) \leq \exp(\frac{3}{2} \int_{t_0}^{t_1} h(s) ds)$.*

PROOF. Let $f(x) = e^x$ and note $0 \leq f''(x+y) = f(x)f(y) \leq 3f(x)$ if $|y| \leq 1$. Using change of variables for functions of bounded variation we have, for $t_0 \leq t \leq t_1$,

$$\begin{aligned} f(m(t)) &= 1 + \int_{t_0}^t f'(m(s-)) dm(s) \\ &\quad + \sum_{t_0 \leq s \leq t} [f(m(s)) - f(m(s-)) - f'(m(s-)) \delta m(s)] \\ &\leq 1 + \int_{t_0}^t f'(m(s-)) dm(s) + \frac{3}{2} \sum_{t_0 \leq s \leq t} f(m(s-)) (\delta m(s))^2, \end{aligned}$$

by Taylor's theorem, (ii) and our observations at the start of the proof. Note that $\int_{t_0}^t f'(m(s-)) dm(s)$ has mean 0 and after applying (iii) and taking expectations, we have $Ef(m(t)) \leq 1 + \frac{3}{2} \int_{t_0}^t Ef(m(s))h(s) ds$. The result then follows from Gronwall's inequality. \square

Since $0 \leq \psi(t) < \rho$ for all $t \geq 0$ and we are assuming $\|X^N(0) - \psi(0)\|_\infty \rightarrow 0$ in probability, we may, by conditioning on $\|X^N(0)\|_\infty < \rho$ if necessary, assume without loss of generality that

$$(4.5) \quad 0 \leq X^N(0) < \rho \quad \text{for all } N.$$

Also, by (2.4) it suffices to consider $\sup_{[0, T]} \|X^N(t) - \psi^N(t)\|_\infty$. Let $\tau = \inf\{t: \|X^N(t) - \psi^N(t)\|_\infty > \varepsilon_0\}$ for fixed $\varepsilon_0 \in (0, 1]$ and define

$$(4.6) \quad \begin{aligned} \bar{X}^N(t) &= X^N(t \wedge \tau) \quad \text{for } 0 \leq t \leq \tau < \infty \text{ or } 0 \leq t < \tau = \infty, \\ \bar{X}^N(t) &= X^N(t \wedge \tau) + \int_{t \wedge \tau}^t \Delta_N \bar{X}^N(s) ds \\ &\quad + \int_{t \wedge \tau}^t R(\bar{X}^N(s)) ds \quad \text{for } \tau < t < \infty. \end{aligned}$$

\bar{X}^N is obtained by running X^N until time τ and then running it deterministically afterwards (if $\tau < \infty$). We have

$$\begin{aligned} P \left[\sup_{[0, T]} \|X^N(t) - \psi^N(t)\|_\infty > \varepsilon_0 \right] \\ \leq P \left[\sup_{[0, T]} \|X^N(t \wedge \tau) - \psi^N(t \wedge \tau)\|_\infty \geq \varepsilon_0 \right] \\ \leq P \left[\sup_{[0, T]} \|\bar{X}^N(t) - \psi^N(t)\|_\infty \geq \varepsilon_0 \right], \end{aligned}$$

so we may consider $\sup_{[0, T]} \|\bar{X}^N(t) - \psi^N(t)\|_\infty$. We have

$$(4.7) \quad \bar{X}^N(t) = X^N(0) + \int_0^t \Delta_N \bar{X}^N(s) ds + \int_0^t R(\bar{X}^N(s)) ds + Z^N(t \wedge \tau).$$

By our definitions, the jumps for X^N and \bar{X}^N satisfy

$$\|\delta \bar{X}^N(t)\|_\infty = \|\delta X^N(t \wedge \tau)\|_\infty \leq l^{-1}.$$

Since $0 \leq \psi^N(t) < \rho$ for $t \geq 0$ and $l \rightarrow \infty$, we may assume that

$$(4.8) \quad 0 \leq X^N(t \wedge \tau) < \rho + 1 \quad \text{for } t \geq 0.$$

Using variation of constants we have that for $\tau < \infty$ and $t > \tau$,

$$\bar{X}^N(t) = T_N(t - \tau) X^N(\tau) + \int_\tau^t T_N(t - s) R(\bar{X}^N(s)) ds.$$

By Lemma 4.2 and the definition of R , this shows we may also assume

$$(4.9) \quad 0 \leq \bar{X}^N(t) < \rho + 1 \quad \text{for } t \geq 0.$$

Using variation of constants we have

$$\begin{aligned} \bar{X}^N(t) - \psi^N(t) &= T_N(t)(X^N(0) - \psi^N(0)) \\ &\quad + \int_0^t T_N(t - s)(R(\bar{X}^N(s)) - R(\psi^N(s))) ds + Y^N(t), \end{aligned}$$

where $Y^N(t) = \int_0^t T_N(t - s) dZ^N(s \wedge \tau)$. Note that each $Z^N(s \wedge \tau, kN^{-1})$,

$0 \leq k \leq N - 1$, is of bounded variation in s and T_N may be viewed as a continuous $N \times N$ matrix-valued function. Thus $Y^N(t, kN^{-1})$, $0 \leq k \leq N - 1$, is defined as a Stieltjes integral. Using Lemma 4.2, (2.3) and (4.9), we have

$$\begin{aligned} & \|\bar{X}^N(t) - \psi^N(t)\|_\infty \\ & \leq \|X^N(0) - \psi^N(0)\|_\infty + K \int_0^t \|\bar{X}^N(s) - \psi^N(s)\|_\infty ds + \|Y^N(t)\|_\infty, \end{aligned}$$

where K depends on ρ and the coefficients of $R(x)$. By our previous discussion and Gronwall's inequality, Theorem 4.1 will follow from the following result.

4.10 LEMMA. $\sup_{[0, T]} \|Y^N(t)\|_\infty \rightarrow 0$ in probability if $\log N/l \rightarrow 0$ as $N \rightarrow \infty$.

PROOF. Fix $\bar{t} \in (0, T]$ and $k \in \{0, 1, \dots, N - 1\}$ and let $f = NI_{[k/N, (k+1)/N]}$. Let $\bar{m}(t) = \langle \int_0^t T_N(\bar{t} - s) dZ^N(s \wedge \tau), f \rangle$ for $0 \leq t \leq \bar{t}$. Note that \bar{m} is a mean 0 martingale on $0 \leq t \leq \bar{t}$ and $\bar{m}(\bar{t}) = Y^N(\bar{t}, kN^{-1})$ since $\langle g, f \rangle = g(kN^{-1})$ for $g \in H^N$. Basic computations using Lemma 3.2 show that, for $\psi \in H^N$,

$$\begin{aligned} & \sum_{s \leq t} (\delta \langle Z^N(s \wedge \tau), \varphi \rangle)^2 \\ (4.11) \quad & - (Nl)^{-1} \int_0^{t \wedge \tau} \left[\langle X^N(s), (\nabla_N^+ \varphi)^2 + (\nabla_N^- \varphi)^2 \rangle \right. \\ & \left. + \langle b(X^N(s)) + d(X^N(s)), \varphi^2 \rangle \right] ds \end{aligned}$$

is a mean 0 martingale. Thus, for $0 \leq t \leq \bar{t}$,

$$\begin{aligned} & \sum_{s \leq t} (\delta \bar{m}(s))^2 \\ (4.12) \quad & - (Nl)^{-1} \int_0^{t \wedge \tau} \langle X^N(s), (\nabla_N^+ T_N(\bar{t} - s) f)^2 + (\nabla_N^- T_N(\bar{t} - s) f)^2 \rangle ds \\ & - (Nl)^{-1} \int_0^{t \wedge \tau} \langle b(X^N(s)) + d(X^N(s)), (T_N(\bar{t} - s) f)^2 \rangle ds \end{aligned}$$

is a mean 0 martingale. Note that $|\delta \bar{m}(s)| \leq l^{-1}$. For $\theta \in [0, 1]$, let $m(t) = \theta l \bar{m}(t)$. Then $|\delta m(t)| \leq 1$ and by Lemma 4.3, (4.8) and Lemma 4.4, we have

$$E \exp(\theta l \bar{m}(\bar{t})) \leq \exp[c(\rho) \theta^2 l (1 + \bar{t} N^{-1})].$$

Since $\bar{t} \leq T$, we may assume $\bar{t}/N \leq 1$. Thus, for $\varepsilon > 0$, we have

$$\begin{aligned} P(Y^N(\bar{t}, kN^{-1}) > \varepsilon) &= P(\theta l Y^N(\bar{t}, kN^{-1}) > \theta l \varepsilon) \\ &\leq E \exp(\theta l Y^N(\bar{t}, kN^{-1})) \exp(-\theta l \varepsilon) \\ &\leq \exp[\theta l (c(\rho) \theta - \varepsilon)]. \end{aligned}$$

Thus we can choose θ so that

$$P[Y^N(\bar{t}, kN^{-1}) > \varepsilon] \leq e^{-a\varepsilon^2 l} \quad \text{for } a = a(\rho) > 0,$$

independently of N , l , k and \bar{t} . The same holds for $P[-Y^N(\bar{t}, kN^{-1}) > \varepsilon]$. Thus, for $0 < t \leq T$ and $k \in \{0, 1, \dots, N-1\}$, we have

$$P[|Y^N(t, kN^{-1})| > \varepsilon] \leq 2e^{-a\varepsilon^2 l}.$$

Since $\|Y^N(t)\|_\infty = \sup_k |Y^N(t, kN^{-1})|$ and $Y^N(0) = 0$, we have

$$(4.13) \quad P[\|Y^N(t)\|_\infty > \varepsilon] \leq 2Ne^{-a\varepsilon^2 l}$$

for $0 \leq t \leq T$ and $a = a(\rho) > 0$. We now show that (4.13) holds with $\|Y^N(t)\|_\infty$ replaced by $\sup_{[0, T]} \|Y^N(t)\|_\infty$ and N (on the right) replaced by N^3 . We have $Y^N(t) = \int_0^t \Delta_N Y^N(s) ds + Z^N(t \wedge \tau)$. Thus for $nTN^{-2} \leq t \leq (n+1)TN^{-2}$ and $0 \leq n \leq N^2 - 1$, we have

$$Y^N(t) = Y^N(nTN^{-2}) + \int_{nTN^{-2}}^t \Delta_N Y^N(s) ds + \tilde{m}(t),$$

where $\tilde{m}(t) = Z^N(t \wedge \tau) - Z^N(nTN^{-2} \wedge \tau)$ for $nTN^{-2} \leq t \leq (n+1)TN^{-2}$. Taking norms and using the definition of Δ_N gives

$$\|Y^N(t)\|_\infty \leq \|Y^N(nTN^{-2})\|_\infty + 4N^2 \int_{nTN^{-2}}^t \|Y^N(s)\|_\infty ds + \|\tilde{m}(t)\|_\infty.$$

Applying Gronwall's inequality shows

$$(4.14) \quad \sup_{[nTN^{-2}, (n+1)TN^{-2}]} \|Y^N(t)\|_\infty \leq \left(\|Y^N(nTN^{-2})\|_\infty + \sup_{[nTN^{-2}, (n+1)TN^{-2}]} \|\tilde{m}(t)\|_\infty \right) e^{4T}.$$

Fix $k \in \{0, 1, \dots, N-1\}$ and $\theta \in [0, 1]$ and let $m(t) = \theta l \tilde{m}(t, kN^{-1})$. By Lemma 3.2(b) and the fact that $\delta Z^N(t, kN^{-1}) = \delta X^N(t, kN^{-1}) = l^{-1} \delta n_k(t)$,

$$\begin{aligned} \sum_{nTN^{-2} \leq s \leq t} (\delta m(s))^2 - \theta^2 l N^2 \int_{nTN^{-2} \wedge \tau}^{t \wedge \tau} [X^N(s, (k-1)N^{-1}) + 2X^N(s, kN^{-1}) \\ + X^N(s, (k+1)N^{-1})] ds \\ - \theta^2 l \int_{nTN^{-2} \wedge \tau}^{t \wedge \tau} [b(X^N(s, kN^{-1})) + d(X^N(s, kN^{-1}))] ds \end{aligned}$$

is a mean 0 martingale for $nTN^{-2} \leq t \leq (n+1)TN^{-2}$. Also $|\delta m(t)| \leq 1$. By Lemma 4.4 and (4.8), we have

$$E \exp[m((n+1)TN^{-2})] \leq \exp(c(\rho)\theta^2 l T).$$

Thus, applying Doob's inequality, we have

$$\begin{aligned} P \left[\sup_{[nTN^{-2}, (n+1)TN^{-2}]} \tilde{m}(t, kN^{-1}) \geq \varepsilon \right] &\leq E \exp[m((n+1)TN^{-2})] \exp(-\theta l \varepsilon) \\ &\leq \exp(\theta l (c(\rho) T \theta - \varepsilon)) \\ &\leq e^{-a\varepsilon^2 l} \quad \text{where } a = a(\rho, T) > 0. \end{aligned}$$

The same holds for $-\tilde{m}(t, k/N)$. But this shows

$$(4.15) \quad P \left[\sup_{[nTN^{-2}, (n+1)TN^{-2}]} \|\tilde{m}(t)\|_\infty > \varepsilon \right] \leq 2Ne^{-a\varepsilon^2 t}.$$

By (4.13), (4.14) and (4.15) we have

$$(4.16) \quad P \left[e^{-4T} \sup_{[nTN^{-2}, (n+1)TN^{-2}]} \|Y^N(t)\|_\infty > \varepsilon \right] \leq 4Ne^{-a\varepsilon^2 t}.$$

We also have

$$P \left[e^{-4T} \sup_{[0, T]} \|Y^N(t)\|_\infty > \varepsilon \right] \leq \sum_{n=0}^{N^2-1} P \left[e^{-4T} \sup_{[nTN^{-2}, (n+1)TN^{-2}]} \|Y^N(t)\|_\infty > \varepsilon \right].$$

By (4.16) this implies

$$(4.17) \quad P \left[e^{-4T} \sup_{[0, T]} \|Y^N(t)\|_\infty > \varepsilon \right] \leq 4N^3 e^{-a\varepsilon^2 t},$$

where $a = a(\rho, T) > 0$. And we have $\sup_{[0, T]} \|Y^N(t)\|_\infty$ converges to 0 in probability if $\log N/l \rightarrow 0$ as $N \rightarrow \infty$. This completes the proof of the lemma and Theorem 4.1. \square

REMARK. Our result is easily extended to the q -dimensional hypercube, $[0, 1]^q$. In this case $[0, 1]^q$ is divided into N^q disjoint congruent cells of volume N^{-q} . Particles jump to each neighboring cell at rate N^2 , or total rate $2qN^2$. Δ_N is given by

$$\Delta_N f(r_1, \dots, r_q) = \sum_{i=1}^q -\nabla_{N,i}^+ \nabla_{N,i}^- f(r_1, \dots, r_q),$$

where

$$\nabla_{N,i}^\pm f(r_1, \dots, r_q) = N \left[f(r_1, \dots, r_i \pm N^{-1}, \dots, r_q) - f(r_1, \dots, r_q) \right].$$

Δ_N has N^q bounded eigenvectors formed by taking products of the one-dimensional eigenvectors.

Our final estimate, (4.17), now becomes

$$(4.18) \quad P \left[e^{-4T} \sup_{[0, T]} \|Y^N(t)\|_\infty > \varepsilon \right] \leq 4N^{q+2} e^{-a\varepsilon^2 t},$$

where the N^q arises from taking the supremum over all cells, as in (4.13) and (4.15). An important fact is that replacing N by N^q in Lemma 4.3 causes no change since this N^q is cancelled by replacing N^{-1} by N^{-q} in (4.12). Our result would also hold as well for reflecting boundary conditions. For a linear reaction one need only replace ρ by ρ_T .

Finally, we note that the idea of defining the martingale

$$\bar{m}(t) = \left\langle \int_0^t T_N(\bar{t} - s) dZ^N(s \wedge \tau), NI_{[k/N, (k+1)/N]} \right\rangle$$

to obtain information about $Y(\bar{t}, kN^{-1}) = \bar{m}(\bar{t})$ was used to bound the moments of X^N in Kotelenez (1988), but we have refined it here by bounding the moment generating function of $Y(\bar{t}, kN^{-1})$.

4.19 EXAMPLE. Assume the initial cell numbers $\{n_k(0): 0 \leq k \leq N - 1\}$ are independent Poisson random variables, each with mean l . Also assume there is no reaction, only diffusion; that is, $b(x) = d(x) \equiv 0$. Then $X^N(t)$ has a stationary distribution with $\{lX^N(t, kN^{-1}): 0 \leq k \leq N - 1\}$ being independent and mean l Poisson random variables for each $t \geq 0$. In this case, $\psi(t, r) \equiv 1$. Given $\varepsilon > 0$,

$$\begin{aligned} P(\|X^N(0) - \psi(0)\|_\infty \geq \varepsilon) &\leq \sum_{k=0}^{N-1} P(n_k(0) \geq l(1 + \varepsilon)) + P(n_k(0) \leq l(1 - \varepsilon)) \\ &= NP(n_1(0) \geq l(1 + \varepsilon)) + NP(n_1(0) \leq l(1 - \varepsilon)) \\ &\leq 2N \exp[\theta l(\theta - \varepsilon)] \end{aligned}$$

for $\theta > 0$, after applying Markov's inequality and using the moment generating function of $n_1(0)$. Setting $\theta = \varepsilon/2$ shows that $\|X^N(0) - \psi(0)\|_\infty \rightarrow 0$ in probability if $\log N/l \rightarrow 0$ as $N \rightarrow \infty$. Thus, by Theorem 4.1, $\sup_{[0, T]} \|X^N(t) - \psi(t)\|_\infty \rightarrow 0$ in probability if $\log N/l \rightarrow 0$ as $N \rightarrow \infty$.

This can also be proved by an elementary argument as in Blount (1987). Conversely, we now show that if $N \rightarrow \infty$ and $\limsup(l/\log N) < \infty$, then one can choose $\varepsilon > 0$ such that

$$\limsup P[\|X^N(t) - \psi(t)\|_\infty \geq \varepsilon] \rightarrow 1 \quad \text{as } N \rightarrow \infty$$

for each fixed t . This will show that requiring $\log N/l \rightarrow 0$ as $N \rightarrow \infty$ is necessary for this special case.

By stationarity it suffices to consider $t = 0$. Let A_k be the event $\{n_k(0) \geq l(1 + \varepsilon)\}$. Then

$$\begin{aligned} (4.20) \quad 1 - P(\|X^N(0) - \psi(0)\|_\infty \geq \varepsilon) &\leq 1 - P\left(\bigcup_0^{N-1} A_k\right) = (1 - P(A_1))^N \leq \exp(-NP(A_1)). \end{aligned}$$

Let $[\cdot]$ denote the greatest integer function. Then

$$\begin{aligned} NP(A_1) &\geq NP(n_1(0) = [l(1 + \varepsilon)] + 1) \\ &\geq Ne^{-l} l^{l(1+\varepsilon)} / ([l(1 + \varepsilon)] + 1)! \\ &\geq cNe^{-l} l^{l(1+\varepsilon)} e^{l(1+\varepsilon)} / (l^{1/2}(l(1 + \varepsilon) + 1)^{(1+\varepsilon)+1}), \end{aligned}$$

where we have applied Stirling's formula and $c = c(l)$ is bounded away from 0 for all large l . Thus,

$$\begin{aligned} NP(A_1) &\geq cNe^{\varepsilon l} / (l^{3/2}(1 + \varepsilon + l^{-1})^{l(1+\varepsilon+l^{-1})}) \\ &\geq cNl^{-3/2} \exp(-l(\varepsilon^2 + 2\varepsilon l^{-1} + l^{-1} + l^{-2})), \end{aligned}$$

where we have used $(1+x)^{-1} \geq e^{-x}$ for $x \geq 0$. If $\limsup(l/\log N) < \infty$, we can choose ε so that $\limsup NP(A_1) = \infty$. By (4.20), this implies $\limsup P(\|X^N(0) - \psi(0)\|_\infty \geq \varepsilon) = 1$.

Note that in this example one can obtain convergence by holding N fixed and letting $l \rightarrow \infty$ arbitrarily. However, this is because $\psi(t)$ is spatially homogeneous. If $\Delta\psi \neq 0$, then letting $N \rightarrow \infty$ is necessary.

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REFERENCES

- ARNOLD, L. and THEODOSOPULU, M. (1980). Deterministic limit of the stochastic model of chemical reactions with diffusion. *Adv. in Appl. Probab.* **12** 367–379.
- BLOUNT, D. J. (1987). Comparison of a stochastic model of a chemical reaction with diffusion and the deterministic model. Ph.D. dissertation, Univ. Wisconsin-Madison.
- BLOUNT, D. J. (1991). Comparison of stochastic and deterministic models of a linear chemical reaction with diffusion. *Ann. Probab.* **19** 1440–1462.
- ETHIER, S. N. and KURTZ, T. G. (1986). *Markov Processes: Characterization and Convergence*. Wiley, New York.
- KOTELENEZ, P. (1982a). Ph.D. dissertation, Universität Bremen Forschungsschwerpunkt Dynamische Systemes.
- KOTELENEZ, P. (1982b). A submartingale type inequality with applications to stochastic evolution equations. *Stochastics* **8** 139–151.
- KOTELENEZ, P. (1986a). Law of large numbers and central limit theorem for linear chemical reactions with diffusion. *Ann. Probab.* **14** 173–193.
- KOTELENEZ, P. (1986b). Gaussian approximation to the nonlinear reaction–diffusion equation. Report 146, Universität Bremen Forschungsschwerpunkt Dynamische Systems.
- KOTELENEZ, P. (1987). Fluctuations near homogeneous states of chemical reactions with diffusion. *Adv. in Appl. Probab.* **19** 352–370.
- KOTELENEZ, P. (1988). High density limit theorems for nonlinear chemical reactions with diffusion. *Probab. Theory Related Fields* **78** 11–37.
- KURTZ, T. G. (1971). Limit theorems for sequences of jump Markov processes approximating ordinary differential processes. *J. Appl. Probab.* **9** 344–356.

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