

Limit theorems for affine Markov walks conditioned to stay positive

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Abstract. Consider the real Markov walk $S_n = X_1 + \dots + X_n$ with increments $(X_n)_{n \geq 1}$ defined by a stochastic recursion starting at $X_0 = x$. For a starting point $y > 0$, denote by τ_y the exit time of the process $(y + S_n)_{n \geq 1}$ from the positive part of the real line. We investigate the asymptotic behaviour of the probability of the event $\tau_y \geq n$ and of the conditional law of $y + S_n$ given $\tau_y \geq n$ as $n \rightarrow +\infty$.

Résumé. On considère une marche Markovienne réelle $S_n = X_1 + \dots + X_n$ dont les accroissements $(X_n)_{n \geq 1}$ sont définis par une récursion stochastique partant de $X_0 = x$. Pour un point de départ $y > 0$, on note par τ_y le temps de sortie du processus $(y + S_n)_{n \geq 1}$ de la partie positive de la droite des réels. On s'intéresse au comportement asymptotique de la probabilité de l'évènement $\tau_y \geq n$ ainsi qu'à la loi conditionnelle de $y + S_n$ sachant $\tau_y \geq n$ quand $n \rightarrow +\infty$.

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1. Introduction

Assume that the Markov chain $(X_n)_{n \geq 0}$ is defined by the stochastic recursion

$$X_0 = x \in \mathbb{R}, \quad X_{n+1} = a_{n+1}X_n + b_{n+1}, \quad n \geq 0, \quad (1.1)$$

where $(a_i, b_i)_{i \geq 1}$ is a sequence of i.i.d. real random pairs satisfying $\mathbb{E}(|a_1|^\alpha) < 1$ and $\mathbb{E}(|b_1|^\alpha) < +\infty$, for some $\alpha > 2$. Consider the Markov walk $S_n = \sum_{k=1}^n X_k$, $n \geq 1$. Under a set of conditions ensuring the existence of the spectral gap of the transition operator of the Markov chain $(X_n)_{n \geq 0}$, it was established in Guivarc'h and Le Page [17] that there exist constants μ and $\sigma > 0$ such that, for any $t \in \mathbb{R}$,

$$\mathbb{P}_x \left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq t \right) \rightarrow \Phi(t) \quad \text{as } n \rightarrow +\infty, \quad (1.2)$$

where Φ is the standard normal distribution function and \mathbb{P}_x is the probability measure generated by $(X_n)_{n \geq 0}$ starting at $X_0 = x$. There are simple expressions of μ and σ in terms of law of the pair (a, b) : in particular $\mu = \frac{\mathbb{E}b}{1 - \mathbb{E}a}$.

For a starting point $y > 0$, define the first time when the affine Markov walk $(y + S_n)_{n \geq 1}$ becomes non-positive by setting

$$\tau_y = \min\{k \geq 1, y + S_k \leq 0\}.$$

In this paper we complete upon the results in [17] by determining the asymptotic of the probability $\mathbb{P}_x(\tau_y > n)$ and proving a conditional version of the limit theorem (1.2) for the sum $y + S_n$, given the event $\{\tau_y > n\}$ in the case when $\mu = 0$. The main challenge in obtaining these asymptotics is to prove the existence of a positive harmonic function pertaining to the associated Markov chain $(X_n, y + S_n)_{n \geq 0}$. A positive harmonic function, say V , is defined as a positive solution of the equation $\mathbf{Q}_+ V = V$, where \mathbf{Q}_+ is the restriction on $\mathbb{R} \times \mathbb{R}_+^*$ of the Markov transition kernel \mathbf{Q} of the chain $(X_n, y + S_n)_{n \geq 0}$.

From the more general results of the paper it follows that, under the same hypotheses that ensure the CLT (see Condition 1 in Section 2), if the pair (a, b) is such that $\mathbb{P}((a, b) \in (0, 1) \times (0, C]) > 0$ and $\mathbb{P}((a, b) \in (-1, 0) \times (0, C]) > 0$, for some $C > 0$, then

$$\mathbb{P}_x(\tau_y > n) \underset{n \rightarrow +\infty}{\sim} \frac{2V(x, y)}{\sqrt{2\pi n\sigma}} \quad (1.3)$$

and

$$\mathbb{P}_x\left(\frac{y + S_n}{\sigma\sqrt{n}} \leq t \mid \tau_y > n\right) \underset{n \rightarrow +\infty}{\longrightarrow} \Phi^+(t), \quad (1.4)$$

where $\Phi^+(t) = 1 - e^{-t^2/2}$ is the Rayleigh distribution function. In particular, the above mentioned results hold true if a and b are independent and a is such that $\mathbb{P}(a \in (0, 1)) > 0$ and $\mathbb{P}(a \in (-1, 0)) > 0$. Less restrictive assumptions on the pair (a, b) are formulated in our Section 2. For example, (1.3) and (1.4) hold if $a = 0$ and b satisfies Condition 1 which covers the case of independent increments.

The above mentioned results are in line with those already known in the literature for random walks with independent increments conditioned to stay in limited areas: the rate $1/\sqrt{n}$ in (1.3) and the asymptotic distribution $\Phi^+(t)$ in (1.4) are the same. We refer the reader to Iglehart [18], Bolthausen [2], Doney [11], Bertoin and Doney [1], Borovkov [3,4], Caravenna [5], Eichelsbacher and Köning [12], Garbit [13], Denisov, Vatutin and Wachtel [7], Denisov and Wachtel [8,10]. More general walks with increments forming a Markov chain have been considered by Presman [20, 21], Varopoulos [22,23], Dembo [6], Denisov and Wachtel [9] or Grama, Le Page and Peigné [16]. In [20,21] the case of sums of lattice random variables defined on finite regular Markov chains has been considered. Varopoulos [22,23] studied Markov chains with bounded increments and obtained lower and upper bounds for the probabilities of the exit time from cones. Some studies take advantage of additional properties: for instance in [9] the Markov walk has a special integrated structure; in [16] the moments of X_n are bounded by some constants not depending on the initial condition. However, to the best of our knowledge, the asymptotic behaviour of the probability $\mathbb{P}_x(\tau_y > n)$ in the case of the stochastic recursion (1.1) has not yet been considered in the literature.

Note that the Wiener–Hopf factorization, which usually is employed in the case of independent random variables, cannot be applied in a straightforward manner for Markov chains. Instead, to study the case of the stochastic recursion, we rely upon the developments in [9,10] and [16]. The main idea of the paper is given below. The existence of the positive harmonic function V is linked to the construction of a martingale approximation for the Markov walk $(S_n)_{n \geq 1}$. While the harmonicity is inherently related to the martingale properties, the difficulty is to show that the approximating martingale is integrable at the exit time of the Markov walk $(y + S_n)_{n \geq 1}$. In contrast to [10] and [16], our proof of the existence of V employs different techniques according to positivity or not of the values of $\mathbb{E}(a_1)$. The constructed harmonic function allows to deduce the properties of the exit time and the conditional distribution of the Markov walk from those of the Brownian motion using a strong approximation result for Markov chains from Grama, Le Page and Peigné [15].

The technical steps of the proofs are as follows. We first deal with the case when the starting point of the Markov walk $(y + S_n)_{n \geq 0}$ is large: $y > n^{1/2-\varepsilon}$, for some $\varepsilon > 0$. When $y > 0$ is arbitrary, the law of iterated logarithm ensures that the sequence $(|y + S_k|)_{1 \leq k \leq n^{1-\varepsilon}}$ will cross the level $n^{1/2-\varepsilon}$ with high probability. Then, by the Markov property, we are able to reduce the problem to a Markov walk with a large starting point $y' = y + S_{\nu_n}$, where ν_n is the first time when the sequence $|y + S_k|$ exceeds the level $n^{1/2-\varepsilon}$. The major difficulty, compared to [10] and [16], is that, for the affine model under consideration, the sequence $(X_{\nu_n})_{n \geq 1}$ is not bounded in \mathbb{L}^1 . To overcome this we need a control of the moments of X_n in function of the initial state $X_0 = x$ and the lag n .

We end this section by agreeing upon some basic notations. As from now and for the rest of this paper the symbols $c, c_\alpha, c_{\alpha,\beta}, \dots$ denote positive constants depending only on their indices. All these constants are likely to change their

values every occurrence. The indicator of an event A is denoted by $\mathbb{1}_A$. For any bounded measurable function f on $\mathbb{X} = \mathbb{R}^d$, $d = 1, 2$, random variable X in \mathbb{X} and event A , the integral $\int_{\mathbb{X}} f(x)\mathbb{P}(X \in dx, A)$ means the expectation $\mathbb{E}(f(X); A) = \mathbb{E}(f(X)\mathbb{1}_A)$.

2. Notations and results

Assume that on the probability space $(\Omega, \mathcal{F}, \mathbb{P})$ we are given a sequence of independent real random pairs (a_i, b_i) , $i \geq 1$, with the same law as the generic random pair (a, b) . Denote by \mathbb{E} the expectation pertaining to \mathbb{P} . Consider the Markov chain $(X_n)_{n \geq 0}$ defined by the affine transformations

$$X_{n+1} = a_{n+1}X_n + b_{n+1}, \quad n \geq 0,$$

where $X_0 = x \in \mathbb{R}$ is a starting point. The partial sum process $(S_n)_{n \geq 0}$ defined by $S_n = \sum_{i=1}^n X_i$ for all $n \geq 1$ and $S_0 = 0$ will be called affine Markov walk. Note that $(S_n)_{n \geq 0}$ itself is not a Markov chain, but the pair $(X_n, S_n)_{n \geq 0}$ forms a Markov chain.

For any $x \in \mathbb{R}$, denote by \mathbb{P}_x and \mathbb{E}_x the probability and the corresponding expectation generated by the finite dimensional distributions of $(X_n)_{n \geq 0}$ starting at $X_0 = x$.

We make use of the following condition which ensures that the affine Markov walk satisfies the central limit theorem (1.2) (cf. [17]):

Condition 1. The pair (a, b) is such that:

- (1) There exists a constant $\alpha > 2$ such that $\mathbb{E}(|a|^\alpha) < 1$ and $\mathbb{E}(|b|^\alpha) < +\infty$.
- (2) The random variable b is non-zero with positive probability, $\mathbb{P}(b \neq 0) > 0$, and centred, $\mathbb{E}(b) = 0$.

Note that Condition 1 is weaker than the conditions required in [17] in the special case $\alpha > 2$. Nevertheless, using the same techniques as in [17] it can be shown that, under Condition 1, the Markov chain $(X_n)_{n \geq 0}$ has a unique invariant measure \mathbf{m} and its partial sum S_n satisfies the central limit theorem (1.2) with

$$\mu = \int_{\mathbb{R}} x \mathbf{m}(dx) = \frac{\mathbb{E}(b)}{1 - \mathbb{E}(a)} = 0 \tag{2.1}$$

and

$$\sigma^2 = \int_{\mathbb{R}} x^2 \mathbf{m}(dx) + 2 \sum_{k=1}^{\infty} \int_{\mathbb{R}} x \mathbb{E}_x(X_k) \mathbf{m}(dx) = \frac{\mathbb{E}(b^2)}{1 - \mathbb{E}(a^2)} \frac{1 + \mathbb{E}(a)}{1 - \mathbb{E}(a)} > 0. \tag{2.2}$$

Moreover, it is easy to see that under Condition 1 the Markov chain $(X_n)_{n \geq 0}$ has no fixed point: $\mathbb{P}(ax + b = x) < 1$, for any $x \in \mathbb{R}$. Below we make use of a slightly refined result which gives the rate of convergence in the central limit theorem for S_n with an explicit dependence of the constants on the initial value $X_0 = x$ stated in Section A.3.

For any $y \in \mathbb{R}$ consider the affine Markov walk $(y + S_n)_{n \geq 0}$ starting at y and define its exit time

$$\tau_y = \min\{k \geq 1, y + S_k \leq 0\}.$$

Corollary A.7 implies the finiteness of the stopping time τ_y : under Condition 1, it holds $\mathbb{P}_x(\tau_y < +\infty) = 1$, for any $x \in \mathbb{R}$ and $y \in \mathbb{R}$.

The asymptotic behaviour of the probability $\mathbb{P}(\tau_y > n)$ is determined by the harmonic function which we proceed to introduce. For any $(x, y) \in \mathbb{R} \times \mathbb{R}$, denote by $\mathbf{Q}(x, y, \cdot)$ the transition probability of the Markov chain $(X_n, y + S_n)_{n \geq 0}$. The restriction of the measure $\mathbf{Q}(x, y, \cdot)$ on $\mathbb{R} \times \mathbb{R}_+^*$ is defined by

$$\mathbf{Q}_+(x, y, B) = \mathbf{Q}(x, y, B)$$

for any measurable set B on $\mathbb{R} \times \mathbb{R}_+^*$ and for any $(x, y) \in \mathbb{R} \times \mathbb{R}$. Let \mathcal{D} be a measurable set in $\mathbb{R} \times \mathbb{R}$ containing $\mathbb{R} \times \mathbb{R}_+^*$. For any measurable $\varphi : \mathcal{D} \rightarrow \mathbb{R}$ set $\mathbf{Q}_+\varphi(x, y) = \int_{\mathbb{R} \times \mathbb{R}_+^*} \varphi(x', y') \mathbf{Q}_+(x, y, dx' \times dy')$. A \mathbf{Q}_+ -harmonic function on \mathcal{D} is any function $V : \mathcal{D} \rightarrow \mathbb{R}$ which satisfies

$$\mathbf{Q}_+V(x, y) = V(x, y), \quad \text{for any } (x, y) \in \mathcal{D}.$$

The existence of a non-negative harmonic function is obvious: $V = 0$ is an example. To ensure the existence of a harmonic function which is positive on a set containing $\mathbb{R} \times \mathbb{R}_+^*$, we need additional assumptions.

Condition 2. For all $x \in \mathbb{R}$ and $y > 0$,

$$\mathbb{P}_x(\tau_y > 1) = \mathbb{P}(ax + b > -y) > 0.$$

Condition 3. For any $x \in \mathbb{R}$ and $y > 0$, there exists $p_0 \in (2, \alpha)$ such that for any constant $c > 0$, there exists $n_0 \geq 1$ such that,

$$\mathbb{P}_x((X_{n_0}, y + S_{n_0}) \in K_{p_0, c}, \tau_y > n_0) > 0,$$

where

$$K_{p_0, c} = \{(x, y) \in \mathbb{R} \times \mathbb{R}_+^*, y \geq c(1 + |x|^{p_0})\}.$$

It is clear that Condition 3 implies Condition 2. Moreover under either Condition 2 or Condition 3, the event $\{\tau_y > n\}$ has positive probability, for any $n \geq 1$, $x \in \mathbb{R}$ and $y > 0$.

The existence of a harmonic function is guaranteed by the following theorem. For any $x \in \mathbb{R}$ consider the process $(M_n)_{n \geq 0}$ defined by

$$M_0 = 0, \quad M_n = S_n + \frac{\mathbb{E}(a)}{1 - \mathbb{E}(a)}(X_n - x), \quad n \geq 1, \quad (2.3)$$

and the natural filtration $(\mathcal{F}_n)_{n \geq 0}$ with \mathcal{F}_0 the trivial σ -algebra and \mathcal{F}_n the σ -algebra generated by X_1, X_2, \dots, X_n . It is easy to verify that $(M_n, \mathcal{F}_n)_{n \geq 0}$ is a \mathbb{P}_x -martingale, for any $x \in \mathbb{R}$ (see Gordin [14]).

Theorem 2.1. Assume Condition 1.

(1) For any $x \in \mathbb{R}$ and $y > 0$, the random variable M_{τ_y} is integrable,

$$\mathbb{E}_x(|M_{\tau_y}|) < +\infty$$

and the function

$$V(x, y) = -\mathbb{E}_x(M_{\tau_y}), \quad x \in \mathbb{R}, y > 0,$$

is well defined on $\mathbb{R} \times \mathbb{R}_+^*$.

(2) The function V has the following properties:

- (a) For any $x \in \mathbb{R}$, the function $V(x, \cdot)$ is non-decreasing.
- (b) For any $\delta > 0$, $p \in (2, \alpha)$, $x \in \mathbb{R}$ and $y > 0$,

$$\begin{aligned} V(x, y) &\geq \max(0, (1 - \delta)y - c_{p, \delta}(1 + |x|^p)), \\ V(x, y) &\leq (1 + \delta(1 + |x|^{p-1}))y + c_{p, \delta}(1 + |x|^p). \end{aligned}$$

(c) For any $x \in \mathbb{R}$, it holds $\lim_{y \rightarrow +\infty} \frac{V(x, y)}{y} = 1$.

(3) The function V is \mathbf{Q}_+ -harmonic on $\mathbb{R} \times \mathbb{R}_+^*$: for any $x \in \mathbb{R}$ and $y > 0$,

$$\mathbf{Q}_+V(x, y) = V(x, y).$$

(4) If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then the function V is positive on $\mathbb{R} \times \mathbb{R}_+^*$.

Using the harmonic function from the previous theorem, we obtain the asymptotic of the tail probability of the exit time τ_y .

Theorem 2.2. Assume Condition 1.

(1) For any $p \in (2, \alpha)$, $x \in \mathbb{R}$ and $y > 0$,

$$\sqrt{n} \mathbb{P}_x(\tau_y > n) \leq c_p(1 + y + |x|)^p.$$

(2) If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then for any $x \in \mathbb{R}$ and $y > 0$,

$$\mathbb{P}_x(\tau_y > n) \underset{n \rightarrow +\infty}{\sim} \frac{2V(x, y)}{\sqrt{2\pi n\sigma}}.$$

Corollary 2.3. Assume Condition 1. For any $p \in (2, \alpha)$, $x \in \mathbb{R}$, $y > 0$ and $\gamma \in (0, 1/2)$,

$$\mathbb{E}_x(\tau_y^\gamma) \leq c_{p,\gamma}(1 + y + |x|)^p.$$

If in addition we assume Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then

$$\mathbb{E}_x(\tau_y^{1/2}) = +\infty.$$

Moreover, we prove that the Markov walk $(y + S_n)_{n \geq 0}$ conditioned to stay positive satisfies the following limit theorem.

Theorem 2.4. Assume either Conditions 1, 2 and $\mathbb{E}(a) \geq 0$, or Conditions 1 and 3. For any $x \in \mathbb{R}$, $y > 0$ and $t > 0$,

$$\mathbb{P}_x\left(\frac{y + S_n}{\sigma\sqrt{n}} \leq t \mid \tau_y > n\right) \underset{n \rightarrow +\infty}{\longrightarrow} \Phi^+(t),$$

where $\Phi^+(t) = 1 - e^{-\frac{t^2}{2}}$ is the Rayleigh distribution function.

Theorems 2.1, 2.2, 2.4 can be extended to some non-positive initial points y . Set

$$\mathcal{D}^- := \{(x, y) \in \mathbb{R} \times \mathbb{R}_-, \mathbb{P}_x(\tau_y > 1) = \mathbb{P}(ax + b > -y) > 0\}.$$

Theorem 2.5. Assume Condition 1.

(1) For any $(x, y) \in \mathcal{D}^-$, the random variable M_{τ_y} is integrable and the function $V(x, y) = -\mathbb{E}_x(M_{\tau_y})$, is well defined on \mathcal{D}^- .

(2) The function V is \mathcal{Q}_+ -harmonic on $\mathcal{D} = \mathcal{D}^- \cup \mathbb{R} \times \mathbb{R}_+^*$.

(3) If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then V is positive $\mathcal{D} = \mathcal{D}^- \cup \mathbb{R} \times \mathbb{R}_+^*$.

(4) For any $(x, y) \in \mathcal{D}^-$,

$$\sqrt{n} \mathbb{P}_x(\tau_y > n) \leq c_p(1 + |x|)^p.$$

(5) If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then

(a) for any $(x, y) \in \mathcal{D}^-$,

$$\mathbb{P}_x(\tau_y > n) \underset{n \rightarrow +\infty}{\sim} \frac{2V(x, y)}{\sqrt{2\pi n\sigma}}.$$

(b) for any $(x, y) \in \mathcal{D}^-$ and $t > 0$,

$$\mathbb{P}_x \left(\frac{y + S_n}{\sigma \sqrt{n}} \leq t \mid \tau_y > n \right) \xrightarrow{n \rightarrow +\infty} \Phi^+(t).$$

The study of the asymptotic behaviour of τ_y and $y + S_n$ for $y \leq 0$ can be motivated by the problem of determining the time when the population $y_0 + S_n$, starting at $y_0 > 0$, stays over a fixed level H . When $y = y_0 - H$ is in $(-H, 0]$, the time $\tau_y = \inf\{k \geq 1, y_0 + S_k \leq H\}$ is the return time of the population $y_0 + S_n$ under the level H .

Below we discuss two more restrictive assumptions which, however, are easier to verify than Conditions 2 and 3, respectively.

Condition 2bis. The law of the pair (a, b) is such that for all $C > 0$,

$$\mathbb{P}(b \geq C|a|) > 0.$$

Condition 3bis. There exists $C > 0$ such that,

$$\mathbb{P}((a, b) \in (-1, 0) \times (0, C]) > 0 \quad \text{and} \quad \mathbb{P}((a, b) \in (0, 1) \times (0, C]) > 0.$$

It is straightforward that Condition 2bis implies Condition 2. This follows from the inequality

$$\mathbb{P}(ax + b > -y) \geq \mathbb{P}(b \geq C|a|),$$

with $C = |x|$. The fact that Condition 3bis implies Condition 3 is proved in the Appendix A.1.

Under Condition 1, it is easy to see that Condition 3bis is satisfied, for example, when random variables a and b are independent and $P(a \in (-1, 0)) > 0$ and $P(a \in (0, 1)) > 0$.

Note that, while Condition 3 implies Condition 2, there is no link between Conditions 2bis and 3bis. Indeed, if a and b are independent, a is non-negative and the support of b contains \mathbb{R}_+ , then Condition 2bis holds true whereas Condition 3bis does not. At the opposite, if a and b are independent, b bounded and the support of a equals to $\{-1/2\} \cup \{1/2\}$ then Condition 3bis holds true whereas Condition 2bis does not.

The outline of the paper is as follows. The martingale approximation $(M_n)_{n \geq 0}$ of the Markov walk $(S_n)_{n \geq 0}$ and some of its properties are given in Section 3. In Section 4 we prove that the expectation of the killed Markov walk $((y + S_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is bounded uniformly in n . This allows us to prove the existence of the harmonic function and establish some of its properties in Section 5. With the help of the harmonic function and of a strong approximation result for Markov chains we prove Theorems 2.2, 2.4 and 2.5, in Sections 6, 7 and 8 respectively. Appendix is an Appendix where we collect some results used in the proofs.

3. Martingale approximation

In this section we approximate the Markov walk $(S_n)_{n \geq 0}$ by the martingale defined in (2.3) and state some related bounds.

We start by a lemma which shows that there is an exponential decay of the dependence of X_n on the initial state $x = X_0$ as n grows to infinity. This simple fact will be used repeatedly in the sequel.

Lemma 3.1. Assume Condition 1. For all $p \in [1, \alpha]$, $x \in \mathbb{R}$, and $n \geq 0$,

$$\mathbb{E}_x^{1/p}(|X_n|^p) \leq c_p + (\mathbb{E}^{1/p}(|a|^p))^n |x| \leq c_p(1 + |x|).$$

Proof. Since $X_n = \sum_{k=1}^n (b_k \prod_{i=k+1}^n a_i) + \prod_{i=1}^n a_i x$, for $n \geq 1$, with the convention $\prod_{i=n+1}^n a_i = 1$, we have by the Minkowski inequality and the independence of $(a_i, b_i)_{i \geq 1}$,

$$\mathbb{E}_x^{1/p}(|X_n|^p) \leq \sum_{k=1}^n (\mathbb{E}^{1/p}(|b|^p) \mathbb{E}^{1/p}(|a|^p)^{n-k}) + \mathbb{E}^{1/p}(|a|^p)^n |x|.$$

The conclusion of the lemma is thus a direct consequence of Condition 1. □

All over the paper we use the abbreviation

$$\rho = \frac{\mathbb{E}(a)}{1 - \mathbb{E}(a)}. \tag{3.1}$$

Using this notation and the martingale $(M_n)_{n \geq 0}$ defined in (2.3), for any $x \in \mathbb{R}$ and $y \in \mathbb{R}$, the Markov walk $(y + S_n)_{n \geq 0}$ has the following martingale representation:

$$y + S_n = y + \rho x + M_n - \rho X_n, \quad n \geq 0. \tag{3.2}$$

Define the sequence $(X_n^0)_{n \geq 0}$, by

$$X_0^0 = 0 \quad \text{and} \quad X_n^0 = \sum_{k=1}^n b_k \prod_{i=k+1}^n a_i, \quad n \geq 1, \tag{3.3}$$

with the convention $\prod_{i=k+1}^n a_i = 1$ for $k = n$. The sequence $(X_n^0)_{n \geq 0}$ corresponds to the stochastic recursion starting at 0. In the same line, we define $M_0^0 = 0$ and $M_n^0 = \sum_{k=1}^n \frac{X_k^0 - \mathbb{E}(a)X_{k-1}^0}{1 - \mathbb{E}(a)}$, for all $n \geq 1$. It is easy to see that the process $(M_n^0, \mathcal{F}_n)_{n \geq 0}$ is a zero mean \mathbb{P}_x -martingale which is related to the martingale $(M_n)_{n \geq 0}$ by the identity

$$M_n = M_n^0 + \Delta_n x, \tag{3.4}$$

where

$$\Delta_0 = 0 \quad \text{and} \quad \Delta_n = \sum_{k=1}^n \frac{\prod_{i=1}^{k-1} a_i}{1 - \mathbb{E}(a)} (a_k - \mathbb{E}(a)), \quad n \geq 1.$$

The following two lemmas will be used to control $\mathbb{E}_x(|M_n|^p)$.

Lemma 3.2. *Assume Condition 1.*

- (1) *The sequence $(\Delta_n)_{n \geq 0}$ is a centred martingale.*
- (2) *For all $p \in [1, \alpha]$ and $n \geq 0$,*

$$\mathbb{E}^{1/p}(|\Delta_n|^p) \leq c_p.$$

Proof. The first claim follows from the fact that Δ_n is a difference of two zero mean martingales. Using the Minkowski inequality for $1 \leq p \leq \alpha$, the independence of $(a_i)_{i \geq 1}$ and Condition 1 we obtain the second claim. \square

Let us introduce the martingale differences:

$$\xi_k^0 = M_k^0 - M_{k-1}^0 = \frac{X_k^0 - \mathbb{E}(a)X_{k-1}^0}{1 - \mathbb{E}(a)}, \quad k \geq 1.$$

Lemma 3.3. *Assume Condition 1. For all $p \in [1, \alpha]$ and $n \geq 0$,*

$$\mathbb{E}^{1/p}(|\xi_n^0|^p) \leq c_p \quad \text{and} \quad \mathbb{E}^{1/p}(|M_n^0|^p) \leq c_p \sqrt{n}.$$

Proof. For the increments ξ_n^0 we simply use Lemma 3.1 with $x = 0$. For the martingale $(M_n^0)_{n \geq 0}$, the upper bound is obtained by Burkholder inequality: for all $2 < p \leq \alpha$ and all $n \geq 1$,

$$\mathbb{E}^{1/p}(|M_n^0|^p) \leq c_p \mathbb{E}^{1/p} \left(\left(\sum_{k=1}^n (\xi_k^0)^2 \right)^{p/2} \right).$$

By the Hölder inequality with the exponents $u = p/2 > 1$ and $v = \frac{p}{p-2}$, we obtain

$$\mathbb{E}^{1/p}(|M_n^0|^p) \leq c_p \mathbb{E}^{1/p} \left[\left(\sum_{k=1}^n |\xi_k^0|^{2u} \right)^{\frac{p}{2u}} n^{\frac{p}{2v}} \right] \leq c_p n^{\frac{p-2}{2p}} \left(\sum_{k=1}^n c_p \right)^{1/p} = c_p \sqrt{n}.$$

This proves the claim when $2 < p \leq \alpha$. When $1 \leq p \leq 2$ the assertion follows from the case above since the L^p norm is less than the L^q norm for $q \in (2, \alpha]$. \square

Lemma 3.4. *Assume Condition 1. For all $p \in [1, \alpha]$ and $n \geq 0$,*

$$\mathbb{E}_x^{1/p}(|M_n|^p) \leq c_p(|x| + \sqrt{n}).$$

Proof. By the Minkowski inequality and equation (3.4), for all $1 \leq p \leq \alpha$, $x \in \mathbb{R}$ and $n \geq 1$,

$$\mathbb{E}_x^{1/p}(|M_n|^p) \leq \mathbb{E}^{1/p}(|\Delta_n|^p)|x| + \mathbb{E}^{1/p}(|M_n^0|^p).$$

Then, by the claim 2 of Lemma 3.2 and Lemma 3.3, the result follows. \square

4. Bound on the expectation of the killed martingale

The goal of this section is to prepare the background to prove the integrability of the random variable M_{τ_y} , which is crucial for showing the existence of the harmonic function in Section 5. We use different approaches depending on the sign on $\mathbb{E}(a)$: when $\mathbb{E}(a) \geq 0$, in Section 4.2 we prove that the expectation of the martingale $(y + \rho x + M_n)_{n \geq 0}$ killed at τ_y is uniformly bounded in n , while, when $\mathbb{E}(a) < 0$, in Section 4.3 we prove that the expectation of the same martingale killed at T_y is uniformly bounded in n , where T_y is the exit time of the martingale $(y + \rho x + M_n)_{n \geq 0}$.

4.1. Preliminary results

We first state a result concerning the first time when the process $(|y + S_n|)_{n \geq 1}$ (respectively $(|y + \rho x + M_n|)_{n \geq 1}$) crosses the level $n^{1/2-\varepsilon}$. Introduce the following stopping times: for any $n \geq 1$, $\varepsilon \in (0, 1/2)$, $x \in \mathbb{R}$ and $y \in \mathbb{R}$,

$$v_n = v_{n,\varepsilon,y} = \min\{k \geq 1, |y + S_k| > n^{1/2-\varepsilon}\} \tag{4.1}$$

and

$$v_n = v_{n,\varepsilon,x,y} = \min\{k \geq 1, |y + \rho x + M_k| > n^{1/2-\varepsilon}\}.$$

Lemma 4.1. *Assume Condition 1. Let $p \in (2, \alpha)$. There exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $\delta > 0$, $x \in \mathbb{R}$, $y > 0$ and $n \geq 1$,*

$$\mathbb{P}_x(v_n > \delta n^{1-\varepsilon}) \leq \frac{C_{p,\varepsilon,\delta}}{n^{p/2-p\varepsilon}} + c_{p,\varepsilon,\delta} e^{-c_{p,\varepsilon,\delta} n^{1-2\varepsilon}} |x|^p$$

and

$$\mathbb{P}_x(v_n > \delta n^{1-\varepsilon}) \leq \frac{C_{p,\varepsilon,\delta}}{n^{p/2-p\varepsilon}} + c_{p,\varepsilon,\delta} e^{-c_{p,\varepsilon,\delta} n^{1-2\varepsilon}} |x|^p.$$

Proof. With $\varepsilon < \min(1/2, \varepsilon_0)$, where ε_0 is defined in Corollary A.6 and $b > 0$ a constant to be chosen below, let $l = \lceil b^2 \delta n^{1-2\varepsilon} \rceil$, $K = \lceil n^\varepsilon / b^2 \rceil$ and for any $m \geq 1$, $x \in \mathbb{R}$ and $y \in \mathbb{R}$, with $z = y + \rho x$,

$$A_m(x, y) = \left\{ \max_{1 \leq k \leq m} |z + M_{kl}| \leq (1 + 2|\rho|) n^{1/2-\varepsilon} \right\}.$$

Note that by the martingale representation (3.2), we have for any $k \geq 2$, $|z + M_k| = |y + S_k + \rho(y + S_k) - \rho(y + S_{k-1})| \leq (1 + |\rho|)|y + S_k| + |\rho||y + S_{k-1}|$. Then, choosing n large enough to have $l \geq 2$,

$$\begin{aligned} \mathbb{P}_x(v_n > \delta n^{1-\varepsilon}) &= \mathbb{P}_x\left(\max_{1 \leq k \leq \lfloor \delta n^{1-\varepsilon} \rfloor} |y + S_k| \leq n^{1/2-\varepsilon}\right) \\ &\leq \mathbb{P}_x\left(\max_{2 \leq k \leq \lfloor \delta n^{1-\varepsilon} \rfloor} |z + M_k| \leq (1 + 2|\rho|)n^{1/2-\varepsilon}\right) \\ &\leq \mathbb{P}_x(A_K(x, y)). \end{aligned}$$

Moreover, we have also,

$$\mathbb{P}_x(v_n > \delta n^{1-\varepsilon}) \leq \mathbb{P}_x(A_K(x, y)).$$

Since $(X_n, y + S_n)_{n \geq 0}$ is a Markov chain,

$$\mathbb{P}_x(A_K(x, y)) = \int_{\mathbb{R}^2} \mathbb{P}_{x'}(A_1(x', y')) \mathbb{P}_x(X_{(K-1)l} \in dx', y + S_{(K-1)l} \in dy', A_{K-1}(x, y)). \tag{4.2}$$

We use the decomposition (3.4) to write that, with $c = 1 + 2|\rho|$,

$$\begin{aligned} \mathbb{P}_{x'}(A_1(x', y')) &\leq \mathbb{P}_{x'}(|z' + M_l^0| \leq 2cn^{1/2-\varepsilon}, |\Delta_l x'| \leq cn^{1/2-\varepsilon}) \\ &\quad + \mathbb{P}_{x'}(|\Delta_l x'| > cn^{1/2-\varepsilon}). \end{aligned}$$

Using (3.2) with $x = 0$, we have $M_n^0 = S_n^0 + \rho X_n^0$. By the Markov inequality,

$$\begin{aligned} \mathbb{P}_{x'}(A_1(x', y')) &\leq \mathbb{P}_{x'}(|z' + S_l^0| \leq 3cn^{1/2-\varepsilon}, |\rho| |X_l^0| \leq cn^{1/2-\varepsilon}) \\ &\quad + \mathbb{P}_{x'}(|\rho| |X_l^0| > cn^{1/2-\varepsilon}) + c_p \frac{\mathbb{E}(|\Delta_l|^p)}{n^{p/2-p\varepsilon}} |x'|^p. \end{aligned}$$

Since S_l^0 does not depend on x' , using Lemma 3.1 and the claim 2 of Lemma 3.2, we obtain

$$\mathbb{P}_{x'}(A_1(x', y')) \leq \sup_{y' \in \mathbb{R}} \mathbb{P}(|y' + S_l^0| \leq 3cn^{1/2-\varepsilon}) + \frac{c_p(1 + |x'|^p)}{n^{p/2-p\varepsilon}}.$$

Inserting this bound in (4.2), it follows that

$$\begin{aligned} \mathbb{P}_x(A_K(x, y)) &\leq \mathbb{P}_x(A_{K-1}(x, y)) \sup_{y' \in \mathbb{R}} \mathbb{P}(|y' + S_l^0| \leq 3cn^{1/2-\varepsilon}) \\ &\quad + \frac{c_p}{n^{p/2-p\varepsilon}} (1 + \mathbb{E}_x(|X_{(K-1)l}|^p)). \end{aligned}$$

Set $r_n = \frac{3cn^{1/2-\varepsilon}}{\sqrt{l}}$. Denote by $\mathbb{B}_{\frac{-y'}{\sqrt{l}}}(r_n)$ the closed ball centred in $\frac{-y'}{\sqrt{l}}$ of radius r_n . The rate of convergence in the central limit theorem from Corollary A.6 (applied with $x = 0$) implies that,

$$\sup_{y' \in \mathbb{R}} \mathbb{P}\left(\frac{S_l^0}{\sqrt{l}} \in \mathbb{B}_{\frac{-y'}{\sqrt{l}}}(r_n)\right) \leq \sup_{y' \in \mathbb{R}} \int_{\mathbb{B}_{\frac{-y'}{\sqrt{l}}}(r_n)} e^{-\frac{u^2}{2\sigma^2}} \frac{du}{\sqrt{2\pi}\sigma} + 2 \frac{c_{p,\varepsilon}}{l^\varepsilon}.$$

Moreover,

$$\sup_{y' \in \mathbb{R}} \int_{\mathbb{B}_{\frac{-y'}{\sqrt{l}}}(r_n)} e^{-\frac{u^2}{2\sigma^2}} \frac{du}{\sqrt{2\pi}\sigma} \leq \frac{2r_n}{\sqrt{2\pi}\sigma} \leq \frac{c_\delta}{b}.$$

Let $q < 1$. With b large enough in the definition of l , we have $2\frac{c_{p,\varepsilon}}{l^\varepsilon} \leq \frac{q}{2}$, $\frac{c_\delta}{b} \leq \frac{q}{2}$ and thus

$$\sup_{y' \in \mathbb{R}} \mathbb{P} \left(\frac{S_l^0}{\sqrt{l}} \in \mathbb{B}_{\frac{-y'}{\sqrt{l}}}(r_n) \right) \leq q < 1.$$

Iterating, we get

$$\mathbb{P}_x(A_K(x, y)) \leq q^{K-1} \mathbb{P}_x(A_1(x, y)) + \frac{c_p}{n^{p/2-p\varepsilon}} \sum_{k=0}^{K-2} q^k (1 + \mathbb{E}_x(|X_{(K-1-k)l}|^p)).$$

Using the fact that $q^{K-1} \mathbb{P}_x(A_1(x, y)) \leq q^{K-1} = \frac{1}{q} e^{-[n^\varepsilon/b^2] \ln(1/q)} \leq \frac{c_{p,\varepsilon,\delta}}{n^{p/2-p\varepsilon}}$, Lemma 3.1 and the fact that $(K-1-k)l \geq c_{\varepsilon,\delta} n^{1-2\varepsilon}$ for all $0 \leq k \leq K-2$, we finally obtain

$$\mathbb{P}_x(A_K(x, y)) \leq \frac{c_{p,\varepsilon,\delta}}{n^{p/2-p\varepsilon}} + c_{p,\varepsilon,\delta} e^{-c_{p,\varepsilon,\delta} n^{1-2\varepsilon}} |x|^p. \quad \square$$

4.2. Bound on the expectation of the killed martingale: The case $\mathbb{E}(a) \geq 0$

The difficulty in proving that the expectation $\mathbb{E}_x(y + \rho x + M_n; \tau_y > n)$ is bounded uniformly in n lies in the fact that whereas the killed Markov walk $(y + S_n) \mathbb{1}_{\{\tau_y > n\}}$ is non-negative, the random variable $(y + \rho x + M_n) \mathbb{1}_{\{\tau_y > n\}}$ may be not. In the case when $\mathbb{E}(a) \geq 0$ we take advantage of the properties presented in the next lemma.

Lemma 4.2. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$.*

(1) *For all $x \in \mathbb{R}$ and $y > 0$,*

$$y + \rho x + M_{\tau_y} \leq 0, \quad \mathbb{P}_x\text{-a.s.}$$

(2) *For all $x \in \mathbb{R}$ and $y > 0$,*

$$\frac{X_{\tau_y}}{1 - \mathbb{E}(a)} < y + \rho x + M_{\tau_y}, \quad \mathbb{P}_x\text{-a.s.}$$

(3) *For all $x \in \mathbb{R}$ and $y > 0$, the sequence $((y + \rho x + M_n) \mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is a submartingale with respect to \mathbb{P}_x .*

Proof. *Claim 1.* Let, for brevity, $z = y + \rho x$. Since, by the definition of τ_y ,

$$X_{\tau_y} = y + S_{\tau_y} - (y + S_{\tau_y-1}) < 0,$$

it follows from (3.2) and the bound $\mathbb{E}(a) \geq 0$ that $z + M_{\tau_y} \leq y + S_{\tau_y} \leq 0$.

Claim 2. Rewrite the martingale representation (3.2) in the form

$$z + M_n = y + S_{n-1} + \frac{X_n}{1 - \mathbb{E}(a)}. \quad (4.3)$$

So, at the exit time τ_y ,

$$\frac{X_{\tau_y}}{1 - \mathbb{E}(a)} = z + M_{\tau_y} - (y + S_{\tau_y-1}) < z + M_{\tau_y}.$$

Claim 3. Using the first claim and the fact that $(M_n)_{n \geq 0}$ is a martingale,

$$\begin{aligned} \mathbb{E}_x(z + M_{n+1}; \tau_y > n + 1 \mid \mathcal{F}_n) &= z + M_n - \mathbb{E}_x(z + M_{\tau_y}; \tau_y = n + 1 \mid \mathcal{F}_n) - \mathbb{E}_x(z + M_{n+1} \mid \mathcal{F}_n) \mathbb{1}_{\{\tau_y \leq n\}} \\ &\geq (z + M_n) \mathbb{1}_{\{\tau_y > n\}}. \end{aligned} \quad \square$$

In the next lemma we obtain a first bound for the expectation of the killed martingale $((y + \rho x + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ which is of order $n^{1/2-2\varepsilon}$, for some $\varepsilon > 0$. Using a recursive procedure we improve it subsequently to a bound not depending on n .

Lemma 4.3. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$. Let $p \in (2, \alpha)$. For any $\varepsilon \in (0, \frac{p-2}{4p})$, $x \in \mathbb{R}$, $y > 0$ and $n \in \mathbb{N}$, we have*

$$\mathbb{E}_x(y + \rho x + M_n; \tau_y > n) \leq y + \rho x + c|x| + c_p n^{1/2-2\varepsilon}.$$

Proof. By the Doob optional stopping theorem and the claim 2 of Lemma 4.2, with $z = y + \rho x$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq z - \mathbb{E}_x\left(\frac{X_{\tau_y}}{1 - \mathbb{E}(a)}; \tau_y \leq n\right).$$

Note that $X_n = \prod_{i=1}^n a_i x + X_n^0$, with X_n^0 given by (3.3). Then, with $\varepsilon \in (0, 1/4)$,

$$\begin{aligned} \mathbb{E}_x(z + M_n; \tau_y > n) &\leq z + c \sum_{k=1}^n \prod_{i=1}^k \mathbb{E}(|a_i|)|x| + c \mathbb{E}_x\left(|X_{\tau_y}^0|; \tau_y \leq n, \max_{1 \leq k \leq n} |X_k^0| \leq n^{1/2-2\varepsilon}\right) \\ &\quad + c \mathbb{E}_x\left(|X_{\tau_y}^0|; \tau_y \leq n, \max_{1 \leq k \leq n} |X_k^0| > n^{1/2-2\varepsilon}\right). \end{aligned}$$

By the Markov inequality, for $2 < p < \alpha$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq z + c \sum_{k=1}^n \mathbb{E}^k(|a|)|x| + cn^{1/2-2\varepsilon} + c \mathbb{E}_x\left(\frac{\max_{1 \leq k \leq n} |X_k^0|^p}{n^{\frac{p-1}{2}(1-4\varepsilon)}}\right).$$

By Lemma 3.1 (with $x = 0$),

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq z + c|x| + cn^{1/2-2\varepsilon} + c_p \frac{n}{n^{\frac{p-1}{2}(1-4\varepsilon)}}.$$

Since $\varepsilon \in (0, \frac{p-2}{4p})$, we have $\frac{p-1}{2}(1-4\varepsilon) > 1/2 + 2\varepsilon$ which concludes the proof. □

Now we give an improvement of Lemma 4.3 which establishes a bound of the expectation of the killed martingale $((y + \rho x + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ depending only on the starting values x, y .

Lemma 4.4. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$. For any $\delta > 0$, $p \in (2, \alpha)$, $x \in \mathbb{R}$, $y > 0$ and $n \geq 0$,*

$$\mathbb{E}_x(y + \rho x + M_n; \tau_y > n) \leq (1 + c_p \delta (1 + |x|)^{p-1})y + c_{p,\delta} (1 + |x|)^p.$$

Moreover, with $\delta = 1$, for any $p \in (2, \alpha)$, $x \in \mathbb{R}$, $y > 0$ and $n \geq 0$,

$$\mathbb{E}_x(y + \rho x + M_n; \tau_y > n) \leq c_p (1 + y + |x|)(1 + |x|)^{p-1}.$$

Proof. Let $\delta > 0$ and $\varepsilon \in (0, \varepsilon_1]$, where $\varepsilon_1 = \min(\varepsilon_0, \frac{p-2}{4p})$ and ε_0 is defined in Lemma 4.1. Set $z = y + \rho x$. We split the proof following the values of n .

Assume first that $n \leq \delta^{-1/\varepsilon}$. A bound of $\mathbb{E}_x(z + M_n; \tau_y > n)$ is obtained immediately from Lemma 4.3: since $z = y + \rho x$, for any $y > 0$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq y + c|x| + c\sqrt{n} \leq y + c_\delta (1 + |x|)$$

and the lemma is proved when $n \leq \delta^{-1/\varepsilon}$.

Assume now that $n \geq \delta^{-1/\varepsilon}$ and $y > n^{1/2-\varepsilon}$. From Lemma 4.3, we deduce that,

$$\mathbb{E}_x(y + \rho x + M_n; \tau_y > n) \leq y + \rho x + c|x| + c_p n^{1/2-2\varepsilon} \leq (1 + c_p n^{-\varepsilon})y + c|x|,$$

which proves the lemma when $y > n^{1/2-\varepsilon}$ and n is larger than $\delta^{-1/\varepsilon}$.

Now, we turn to the last case, when $n \geq \delta^{-1/\varepsilon}$ and $0 < y \leq n^{1/2-\varepsilon}$. Introduce the following stopping time:

$$v_n^\varepsilon = v_n + \lceil n^\varepsilon \rceil.$$

We have the following obvious decomposition:

$$\mathbb{E}_x(z + M_n; \tau_y > n) = \underbrace{\mathbb{E}_x(z + M_n; \tau_y > n, v_n^\varepsilon > \lceil n^{1-\varepsilon} \rceil)}_{=: J_1} + \underbrace{\mathbb{E}_x(z + M_n; \tau_y > n, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil)}_{=: J_2}. \quad (4.4)$$

Bound of J_1 . Using the Hölder inequality for $1 < p < \alpha$, Lemma 3.4 and Lemma 4.1, we have

$$J_1 \leq c_{p,\varepsilon} \sqrt{n} (1 + y + |x|) \frac{(1 + |x|)^{p-1}}{n^{(p-1)(\frac{1}{2}-\varepsilon)}}.$$

As $\varepsilon < \frac{p-2}{4p}$, denoting $C_{p,\varepsilon}(x, y) = c_{p,\varepsilon}(1 + y + |x|)(1 + |x|)^{p-1}$, for all $n \geq 1$,

$$J_1 \leq \frac{C_{p,\varepsilon}(x, y)}{n^\varepsilon}. \quad (4.5)$$

Bound of J_2 . Using the martingale representation (3.2) for the Markov walk $(y + S_n)_{n \geq 1}$, by the Markov property,

$$J_2 = \sum_{k=1}^{\lceil n^{1-\varepsilon} \rceil} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{E}_{x'}(y' + \rho x' + M_{n-k}; \tau_{y'} > n - k) \mathbb{P}_x(X_{v_n^\varepsilon} \in dx', y + S_{v_n^\varepsilon} \in dy', \tau_y > v_n^\varepsilon, v_n^\varepsilon = k).$$

By Lemma 4.3,

$$J_2 \leq \mathbb{E}_x(z + M_{v_n^\varepsilon} + c|X_{v_n^\varepsilon}| + c_p n^{1/2-2\varepsilon}; \tau_y > v_n^\varepsilon, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil).$$

For the term $z + M_{v_n^\varepsilon}$, we use the fact that $((z + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is a submartingale (claim 3 of Lemma 4.2), while for the term $c|X_{v_n^\varepsilon}| = c|X_{v_n + \lceil n^\varepsilon \rceil}|$ we use the Markov property at v_n and Lemma 3.1. This gives

$$J_2 \leq \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; \tau_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil) + c_p \mathbb{E}_x(n^{1/2-2\varepsilon} + \mathbb{E}^{\lceil n^\varepsilon \rceil}(|a|)|X_{v_n}|; \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil).$$

Since $0 < y \leq n^{1/2-\varepsilon}$ and v_n is the first time when $(|y + S_n|)_{n \geq 1}$ exceeds $n^{1/2-\varepsilon}$, the jump X_{v_n} is necessarily positive on the event $\{\tau_y > v_n\}$. Therefore, under the condition $\mathbb{E}(a) \geq 0$, by the representation (3.2) we have $z + M_{v_n} > n^{1/2-\varepsilon}$. Using the last bound, we obtain

$$J_2 \leq \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; \tau_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil) + c_p \mathbb{E}_x\left(\frac{z + M_{v_n}}{n^\varepsilon}; \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil\right) + c_p e^{-c_p n^\varepsilon} \mathbb{E}_x(|X_{v_n}|; v_n \leq \lceil n^{1-\varepsilon} \rceil).$$

Again, using the fact that $((z + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is a submartingale and Lemma 3.1, we bound J_2 as follows,

$$J_2 \leq \left(1 + \frac{c_p}{n^\varepsilon}\right) \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; \tau_y > \lceil n^{1-\varepsilon} \rceil) + c_p e^{-c_p n^\varepsilon} n^{1-\varepsilon} (1 + |x|) - \underbrace{\mathbb{E}_x\left((z + M_{\lceil n^{1-\varepsilon} \rceil}) \left(\mathbb{1}_{\{v_n^\varepsilon > \lceil n^{1-\varepsilon} \rceil\}} + \frac{c_p}{n^\varepsilon} \mathbb{1}_{\{v_n > \lceil n^{1-\varepsilon} \rceil\}}\right); \tau_y > \lceil n^{1-\varepsilon} \rceil\right)}_{=: J_3}. \quad (4.6)$$

We bound J_3 in a same manner as J_1 ,

$$|J_3| \leq c_{p,\varepsilon} \sqrt{[n^{1-\varepsilon}]} (1+y+|x|) c_{p,\varepsilon} \frac{(1+|x|)^{p-1}}{n^{\frac{p-1}{2}-(p-1)\varepsilon}} \leq \frac{C_{p,\varepsilon}(x,y)}{n^\varepsilon}.$$

Inserting this bound in (4.6) and using (4.5) and (4.4) we find that, for any $n \geq n_0 = \lceil y^{1/(1/2-\varepsilon)} \rceil + 1$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq \left(1 + \frac{c_p}{n^\varepsilon}\right) \mathbb{E}_x(z + M_{\lfloor n^{1-\varepsilon} \rfloor}; \tau_y > \lfloor n^{1-\varepsilon} \rfloor) + \frac{C_{p,\varepsilon}(x,y)}{n^\varepsilon}.$$

Since $((z + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is a submartingale, the sequence $u_n = \mathbb{E}_x(z + M_n; \tau_y > n)$ is non-decreasing. By Lemma A.1 used with $\alpha = c_p$, $\beta = C_{p,\varepsilon}(x,y)$ and $\gamma = 0$ it follows that, for any $n \geq n_0$ and $k_0 \in \{n_0, \dots, n\}$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) \mathbb{E}_x(z + M_{k_0}; \tau_y > k_0) + \frac{C_{p,\varepsilon}(x,y)}{k_0^\varepsilon}.$$

By Lemma 4.3 and the fact that $z = y + \rho x$, we have

$$\begin{aligned} \mathbb{E}_x(z + M_n; \tau_y > n) &\leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) y + c_{p,\varepsilon} k_0^{1/2-2\varepsilon} \\ &\quad + c_{p,\varepsilon} |x| + \frac{c_{p,\varepsilon}}{k_0^\varepsilon} (1+y+|x|)(1+|x|)^{p-1} \\ &\leq \left(1 + \frac{c_{p,\varepsilon}(1+|x|)^{p-1}}{k_0^\varepsilon}\right) y + c_{p,\varepsilon,k_0} (1+|x|)^p. \end{aligned}$$

Choosing $k_0 \geq \delta^{-1/\varepsilon}$, for any $0 < y \leq n^{1/2-\varepsilon}$ and $n \geq \delta^{-1/\varepsilon}$,

$$\mathbb{E}_x(z + M_n; \tau_y > n) \leq (1 + c_{p,\varepsilon} \delta (1+|x|)^{p-1}) y + c_{p,\varepsilon, \delta} (1+|x|)^p.$$

Finally we conclude that the lemma holds true for any $n \in \mathbb{N}$. □

We can now transfer the bound provided by Lemma 4.4 to the Markov walk $(y + S_n)_{n \geq 0}$.

Corollary 4.5. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$. For any $p \in (2, \alpha)$, $x \in \mathbb{R}$, $y > 0$ and $n \in \mathbb{N}$,*

$$\mathbb{E}_x(y + S_n; \tau_y > n) \leq c_p (1+y+|x|)(1+|x|)^{p-1}.$$

Proof. Using equation (3.2), the result follows from Lemma 4.4 and Lemma 3.1. □

4.3. Bound on the expectation of the killed martingale: The case $\mathbb{E}(a) < 0$

We adapt the mainstream of the proof for the case $\mathbb{E}(a) \geq 0$ given in the previous section, highlighting the details that have to be modified.

In the discussion preceding Lemma 4.2, we noted that $(y + \rho x + M_n)\mathbb{1}_{\{\tau_y > n\}}$ may not be positive. In the case $\mathbb{E}(a) < 0$, we overcome this by introducing the exit time of the martingale $(y + \rho x + M_n)_{n \geq 0}$: for any $y \in \mathbb{R}$,

$$T_y = \min\{k \geq 1, y + \rho x + M_k \leq 0\}.$$

The importance of this new exit time is stressed by the fact that one can check that when $\mathbb{E}(a) < 0$, the sequence $((y + \rho x + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is not a submartingale (as in Lemma 4.2 when $\mathbb{E}(a) \geq 0$) but a supermartingale. Instead we prove that $((y + \rho x + M_n)\mathbb{1}_{\{T_y > n\}})_{n \geq 0}$ is a submartingale (see Lemma 4.6 below). This will play an important role in view of obtaining upper bounds. By Corollary A.7 we have $\mathbb{P}_x(T_y < +\infty) = 1$ for any $x \in \mathbb{R}$. The main point is to

show the integrability of $y + \rho x + M_{T_y}$. Under the assumption $\mathbb{E}(a) < 0$ we have $\tau_y \leq T_y$ (see Lemma 4.6 below), which together with the integrability of $y + \rho x + M_{T_y}$ and the fact $(|y + \rho x + M_n|)_{n \geq 0}$ is a submartingale, will allow us to prove in Section 5.2 that $y + \rho x + M_{\tau_y}$ is integrable.

Lemma 4.6. *Assume Condition 1.*

(1) *If $\mathbb{E}(a) < 0$, then for all $x \in \mathbb{R}$ and $y > 0$,*

$$\tau_y \leq T_y \quad \mathbb{P}_x\text{-a.s.}$$

(2) *For all $x \in \mathbb{R}$ and $y \in \mathbb{R}$, the sequence $((y + \rho x + M_n)\mathbb{1}_{\{T_y > n\}})_{n \geq 0}$ is a submartingale with respect to \mathbb{P}_x .*

Proof. *Claim 1.* We note that when $T_y > 1$, by (3.2) and (4.3), with $z = y + \rho x$,

$$\begin{aligned} y + S_{T_y} &= z + M_{T_y} - \rho X_{T_y} \leq -\rho X_{T_y}, \\ y + S_{T_y-1} &= z + M_{T_y} - \frac{X_{T_y}}{1 - \mathbb{E}(a)} \leq -\frac{X_{T_y}}{1 - \mathbb{E}(a)}. \end{aligned}$$

Since $\rho < 0$, according to the positivity or non-positivity of X_{T_y} , we have respectively $y + S_{T_y} \leq 0$ or $y + S_{T_y-1} \leq 0$. When $T_y = 1$ and $y > 0$ we have $X_1 < 0$ and so $\tau_y = 1 = T_y$.

Claim 2. In a same manner as in the proof of the claim 3 of Lemma 4.2, the claim 2 is a consequence of the fact that $z + M_{T_y} \leq 0$ and that $(M_n)_{n \geq 0}$ is a martingale. \square

The following lemma is similar to Lemma 4.3 but with T_y replacing τ_y .

Lemma 4.7. *Assume Condition 1. Let $p \in (2, \alpha)$. For any $\varepsilon \in (0, \frac{p-2}{4p})$, $x \in \mathbb{R}$, $y > -\rho x$ and $n \geq 0$, we have*

$$\mathbb{E}_x(y + \rho x + M_n; T_y > n) \leq y + \rho x + c|x| + c_p n^{1/2-2\varepsilon}.$$

Proof. Note that $z = y + \rho x > 0$. Since at the exit time T_y we have $0 \geq z + M_{T_y} \geq \xi_{T_y} = \frac{X_{T_y} - \mathbb{E}(a)X_{T_y-1}}{1 - \mathbb{E}(a)}$, by the Doob optional stopping theorem,

$$\mathbb{E}_x(z + M_n; T_y > n) \leq z + c\mathbb{E}_x(|X_{T_y}| + |X_{T_y-1}|; T_y \leq n).$$

Since $|X_{T_y}| + |X_{T_y-1}| \leq 2 \max_{1 \leq k \leq n} |X_k| + |x|$ on $\{T_y \leq n\}$, following the proof of Lemma 4.3,

$$\begin{aligned} \mathbb{E}_x(z + M_n; T_y > n) &\leq z + c \left(1 + \sum_{k=1}^n \prod_{i=1}^k \mathbb{E}(|a_i|) \right) |x| \\ &\quad + cn^{1/2-2\varepsilon} \mathbb{P} \left(\max_{1 \leq k \leq n} |X_k^0| \leq n^{1/2-2\varepsilon} \right) \\ &\quad + c\mathbb{E} \left(\max_{1 \leq k \leq n} |X_k^0|; \max_{1 \leq k \leq n} |X_k^0| > n^{1/2-2\varepsilon} \right) \\ &\leq z + c|x| + c_p n^{1/2-2\varepsilon}. \end{aligned} \quad \square$$

Lemma 4.8. *Assume Condition 1. Let $p \in (2, \alpha)$. There exists $\varepsilon_1 > 0$ such that for any $\varepsilon \in (0, \varepsilon_1)$, $x \in \mathbb{R}$, $y \in \mathbb{R}$, $n \geq 0$ and $2 \leq k_0 \leq n$,*

$$\begin{aligned} \mathbb{E}_x(y + \rho x + M_n; T_y > n) &\leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon} \right) \max(y, 0) + c_{p,\varepsilon}|x| + c_{p,\varepsilon}\sqrt{k_0} + c_{p,\varepsilon} e^{-c_{p,\varepsilon}k_0^\varepsilon} |x|^p \\ &\leq c_p (1 + \max(y, 0) + |x|^p). \end{aligned}$$

Proof. We proceed as in the proof of Lemma 4.4. Set $\varepsilon_1 = \min(\varepsilon_0, \frac{p-2}{4p})$, where ε_0 is defined in Lemma 4.1. Let $\varepsilon \in (0, \varepsilon_1]$. With $z = y + \rho x$ and $v_n^\varepsilon = v_n + \lceil n^\varepsilon \rceil$, we have

$$\begin{aligned} \mathbb{E}_x(z + M_n; T_y > n) &= \underbrace{\mathbb{E}_x(z + M_n; T_y > n, v_n^\varepsilon > \lceil n^{1-\varepsilon} \rceil)}_{=: J_1} \\ &\quad + \underbrace{\mathbb{E}_x(z + M_n; T_y > n, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil)}_{=: J_2}. \end{aligned} \quad (4.7)$$

Bound of J_1 . Let $m_\varepsilon = \lceil n^{1-\varepsilon} \rceil - \lceil n^\varepsilon \rceil$. Since on $\{v_n > m_\varepsilon\}$ it holds $z' = z + M_{m_\varepsilon} \leq n^{1/2-\varepsilon}$, by the Markov property we write that

$$J_1 \leq n^{1/2-\varepsilon} \mathbb{P}_x(v_n > m_\varepsilon) + \int_{\mathbb{R}} \mathbb{E}_{x'}(|M_{n-m_\varepsilon}|) \mathbb{P}_x(X_{m_\varepsilon} \in dx', v_n > m_\varepsilon).$$

By Lemma 3.4 and the Hölder inequality,

$$\begin{aligned} J_1 &\leq n^{1/2-\varepsilon} \mathbb{P}_x(v_n > m_\varepsilon) + \mathbb{E}_x(c(\sqrt{n-m_\varepsilon} + |X_{m_\varepsilon}|); v_n > m_\varepsilon) \\ &\leq cn^{1/2} \mathbb{P}_x(v_n > m_\varepsilon) + \mathbb{E}_x^{1/p}(|X_{m_\varepsilon}|^p) \mathbb{P}_x^{1/q}(v_n > m_\varepsilon). \end{aligned}$$

By Lemma 3.1 and Lemma 4.1 (since $m_\varepsilon \geq n^{1-\varepsilon}/c_\varepsilon$),

$$J_1 \leq \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-p\varepsilon}} + c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^{1-2\varepsilon}} |x|^p. \quad (4.8)$$

Bound of J_2 . Repeating the arguments used for bounding the term J_2 in Lemma 4.4, by the Markov property and Lemma 4.7, we get

$$J_2 \leq \mathbb{E}_x(z + M_{v_n^\varepsilon} + c|X_{v_n^\varepsilon}| + c_p n^{1/2-2\varepsilon}; T_y > v_n^\varepsilon, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil).$$

Using the claim 2 of Lemma 4.6 and Lemma 3.1,

$$\begin{aligned} J_2 &\leq \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; T_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil) \\ &\quad + \mathbb{E}_x(c_p n^{1/2-2\varepsilon}; T_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil) + c_{p,\varepsilon} e^{-c_\varepsilon n^\varepsilon} \mathbb{E}_x(|X_{v_n}|; v_n \leq \lceil n^{1-\varepsilon} \rceil). \end{aligned}$$

On the event $\{T_y > v_n\}$, we have $n^{1/2-\varepsilon} < z + M_{v_n}$. Hence

$$\begin{aligned} J_2 &\leq \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; T_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil) \\ &\quad + c_p \mathbb{E}_x\left(\frac{z + M_{v_n}}{n^\varepsilon}; T_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil\right) + c_{p,\varepsilon} e^{-c_\varepsilon n^\varepsilon} \mathbb{E}_x(|X_{v_n}|; v_n \leq \lceil n^{1-\varepsilon} \rceil). \end{aligned}$$

Coupling this with (4.8) and (4.7) and using again the claim 2 of Lemma 4.6, we obtain that

$$\begin{aligned} \mathbb{E}_x(z + M_n; T_y > n) &\leq \left(1 + \frac{c_p}{n^\varepsilon}\right) \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil}; T_y > \lceil n^{1-\varepsilon} \rceil) \\ &\quad + \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-p\varepsilon}} + c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^\varepsilon} |x|^p. \end{aligned}$$

Since $((z + M_n)\mathbb{1}_{\{T_y > n\}})_{n \geq 0}$ is a submartingale (claim 2 of Lemma 4.6), the sequence $u_n = \mathbb{E}(z + M_n; T_y > n)$ is non-decreasing. By Lemma A.1 with $\alpha = c_p$, $\beta = c_{p,\varepsilon}$, $\gamma = |x|^p$ and $\delta = c_{p,\varepsilon}$, we write that

$$\mathbb{E}_x(z + M_n; T_y > n) \leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) \mathbb{E}_x(z + M_{k_0}; T_y > k_0) + \frac{c_{p,\varepsilon}}{k_0^\varepsilon} + c_{p,\varepsilon} e^{-c_{p,\varepsilon} k_0^\varepsilon} |x|^p.$$

Using Lemma 3.4 and the fact that $z = y + \rho x$, we obtain that

$$\mathbb{E}_x(z + M_n; T_y > n) \leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) \max(y, 0) + c_{p,\varepsilon}|x| + c_{p,\varepsilon}\sqrt{k_0} + c_{p,\varepsilon} e^{-c_{p,\varepsilon}k_0^\varepsilon}|x|^p. \quad \square$$

To transfer the assertion of Lemma 4.8 to the random walk $(y + S_n)_{n \geq 0}$, we need to assume that $\mathbb{E}(a) < 0$.

Corollary 4.9. *Assume Condition 1 and $\mathbb{E}(a) < 0$. Let $p \in (2, \alpha)$. For any $x \in \mathbb{R}$, $y > 0$ and $n \in \mathbb{N}$,*

$$\mathbb{E}_x(y + S_n; \tau_y > n) \leq c_p(1 + y + |x|^p).$$

Proof. By (3.2) and the claim 1 of Lemma 4.6, we have

$$\mathbb{E}_x(y + S_n; \tau_y > n) = \mathbb{E}_x(y + \rho x + M_n; T_y \geq \tau_y > n) - \mathbb{E}_x(\rho X_n; \tau_y > n).$$

The result follows from Lemma 4.8. □

5. Existence of the harmonic function

In this section we prove Theorem 2.1. We split the proof into two parts according to the values of $\mathbb{E}(a)$.

5.1. Existence of the harmonic function: The case $\mathbb{E}(a) \geq 0$

We start with the following assertion.

Lemma 5.1. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$. For any $x \in \mathbb{R}$ and $y > 0$, the random variable M_{τ_y} is integrable. Moreover, for any $p \in (2, \alpha)$,*

$$\mathbb{E}_x(|M_{\tau_y}|) \leq c_p(1 + y + |x|)(1 + |x|)^{p-1}.$$

Proof. Let $z = y + \rho x$. Using the claim 1 of Lemma 4.2 and the Doob optional stopping theorem, we have

$$\begin{aligned} \mathbb{E}_x(|M_{\tau_y}|; \tau_y \leq n) &\leq -\mathbb{E}_x(z + M_n; \tau_y \leq n) + y + \rho|x| \\ &= \mathbb{E}_x(z + M_n; \tau_y > n) - z + y + \rho|x|. \end{aligned}$$

By second bound in Lemma 4.4, for all $n \geq 0$,

$$\mathbb{E}_x(|M_{\tau_y}|; \tau_y \leq n) \leq c_p(1 + y + |x|)(1 + |x|)^{p-1} =: C_p(x, y).$$

Since $(\{\tau_y \leq n\})_{n \geq 1}$ is a non-decreasing sequence of events and $\mathbb{P}_x(\tau_y < +\infty) = 1$ for any $x \in \mathbb{R}$ (by Corollary A.7), the result follows by the Lebesgue monotone convergence theorem. □

It follows from Lemma 5.1 that the function

$$V(x, y) = -\mathbb{E}_x(M_{\tau_y})$$

is well defined for any $x \in \mathbb{R}$ and $y > 0$, which also proves the claim 1 of Theorem 2.1 when $\mathbb{E}(a) \geq 0$.

The following two propositions prove the claims 2 and 3 of Theorem 2.1 when $\mathbb{E}(a) \geq 0$.

Proposition 5.2. *Assume Condition 1 and $\mathbb{E}(a) \geq 0$.*

(1) *For any $x \in \mathbb{R}$ and $y > 0$,*

$$V(x, y) = \lim_{n \rightarrow +\infty} \mathbb{E}_x(y + \rho x + M_n; \tau_y > n) = \lim_{n \rightarrow +\infty} \mathbb{E}_x(y + S_n; \tau_y > n).$$

- (2) For any $x \in \mathbb{R}$, the function $V(x, \cdot)$ is non-decreasing.
- (3) For any $\delta > 0$, $p \in (2, \alpha)$, $x \in \mathbb{R}$ and $y > 0$,

$$\max(0, y + \rho x) \leq V(x, y) \leq (1 + c_p \delta (1 + |x|)^{p-1})y + c_{p,\delta} (1 + |x|)^p.$$

1. For any $x \in \mathbb{R}$,

$$\lim_{y \rightarrow +\infty} \frac{V(x, y)}{y} = 1.$$

Proof. We use the notation $z = y + \rho x$.

Claim 1. Since, by Lemma 5.1, M_{τ_y} is integrable, we have by the Lebesgue dominated convergence theorem,

$$\mathbb{E}_x(z + M_n; \tau_y > n) = z - \mathbb{E}_x(z + M_{\tau_y}; \tau_y \leq n) \xrightarrow{n \rightarrow +\infty} -\mathbb{E}_x(M_{\tau_y}) = V(x, y).$$

To prove the second equality of the claim 1 we use Lemma 3.1 and the fact that $\tau_y < +\infty$:

$$|\mathbb{E}_x(X_n; \tau_y > n)| \leq \mathbb{E}_x^{1/2}(|X_n|^2) \sqrt{\mathbb{P}_x(\tau_y > n)} \leq c_2(1 + |x|) \sqrt{\mathbb{P}_x(\tau_y > n)} \xrightarrow{n \rightarrow +\infty} 0.$$

Using (3.2), we obtain the claim 1.

Claim 2. If $y_1 \leq y_2$, then $\tau_{y_1} \leq \tau_{y_2}$ and

$$\mathbb{E}_x(y_1 + S_n; \tau_{y_1} > n) \leq \mathbb{E}_x(y_2 + S_n; \tau_{y_1} > n) \leq \mathbb{E}_x(y_2 + S_n; \tau_{y_2} > n).$$

Taking the limit as $n \rightarrow +\infty$ we get the claim 2.

Claim 3. The upper bound follows from the claim 1 and Lemma 4.4. On the event $\{\tau_y > n\}$, we obviously have $y + S_n > 0$ and so by claim 1, $V(x, y) \geq 0$. Moreover, since $z + M_{\tau_y} \leq 0$ (by claim 1 of Lemma 4.2), we have, by claim 1,

$$V(x, y) = z - \lim_{n \rightarrow +\infty} \mathbb{E}_x(z + M_{\tau_y}; \tau_y \leq n) \geq z,$$

which proves the lower bound.

Claim 4. By the claim 3, for all $\delta > 0$, $x \in \mathbb{R}$,

$$1 \leq \liminf_{y \rightarrow +\infty} \frac{V(x, y)}{y} \leq \limsup_{y \rightarrow +\infty} \frac{V(x, y)}{y} \leq (1 + c_p \delta (1 + |x|)^{p-1}).$$

Letting $\delta \rightarrow 0$, we obtain the claim 4. □

We now prove that V is harmonic on $\mathbb{R} \times \mathbb{R}_+^*$.

Proposition 5.3. Assume Conditions 1 and $\mathbb{E}(a) \geq 0$.

(1) The function V is \mathbf{Q}_+ -harmonic on $\mathbb{R} \times \mathbb{R}_+^*$: for any $x \in \mathbb{R}$ and $y > 0$,

$$\mathbf{Q}_+ V(x, y) = V(x, y).$$

(2) If in addition we assume Condition 2, then the function V is positive on $\mathbb{R} \times \mathbb{R}_+^*$.

Proof. *Claim 1.* Denote for brevity $V_n(x, y) = \mathbb{E}_x(y + S_n; \tau_y > n)$. For all $x \in \mathbb{R}$, $y > 0$ and $n \geq 1$, by the Markov property,

$$V_{n+1}(x, y) = \mathbb{E}_x(V_n(X_1, y + S_1); \tau_y > 1).$$

By Corollary 4.5, we see that the quantity $V_n(X_1, y + S_1)$ is dominated by the random variable $c_p(1 + y + S_1 + |X_1|)(1 + |X_1|)^{p-1}$ which is integrable with respect to \mathbb{E}_x . Consequently, by the Lebesgue dominated convergence theorem and the claim 1 of Proposition 5.2,

$$V(x, y) = \mathbb{E}_x(V(X_1, y + S_1); \tau_y > 1) = \mathbf{Q}_+ V(x, y),$$

where by convention, $V(x, y)\mathbb{1}_{\{y>0\}} = 0$ if $y \leq 0$ and $x \in \mathbb{R}$.

Claim 2. Fix $x \in \mathbb{R}$ and $y > 0$. Using the claim 1 and the fact that V is non-negative on $\mathbb{R} \times \mathbb{R}_+^*$ (claim 3 of Proposition 5.2) we write

$$V(x, y) \geq \mathbb{E}_x\left(V(X_1, y + S_1); \tau_y > 1, X_1 > \frac{-y}{2(1+\rho)}\right).$$

By the lower bound of the claim 3 of Proposition 5.2 and (3.2),

$$V(x, y) \geq \mathbb{E}_x\left(y + (1+\rho)X_1; \tau_y > 1, X_1 > \frac{-y}{2(1+\rho)}\right) \geq \frac{y}{2}\mathbb{P}_x\left(X_1 > \frac{-y}{2(1+\rho)}\right).$$

By Condition 2, we conclude that, $V(x, y) > 0$ for any $x \in \mathbb{R}$ and $y > 0$. □

5.2. Existence of the harmonic function: The case $\mathbb{E}(a) < 0$

In this section we prove the harmonicity and the positivity of the function V in the case $\mathbb{E}(a) < 0$. The following analogue of Lemma 5.1 shows that the random variables M_{T_y} and M_{τ_y} are integrable.

Lemma 5.4. *Assume Condition 1.*

(1) *For any $x \in \mathbb{R}$ and $y \in \mathbb{R}$,*

$$\mathbb{E}_x(|M_{T_y}|) \leq c_p(1 + |y| + |x|^p).$$

(2) *If in addition $\mathbb{E}(a) < 0$, then for any $x \in \mathbb{R}$ and $y \in \mathbb{R}$,*

$$\mathbb{E}_x(|M_{\tau_y}|) \leq c_p(1 + |y| + |x|^p).$$

Proof. *Claim 1.* The proof of the bound of $\mathbb{E}_x(|M_{T_y}|)$ is similar to that of Lemma 5.1 using Lemma 4.8 instead of Lemma 4.4 and the fact that by Corollary A.7 we have $\mathbb{P}_x(T_y < +\infty) = 1$, $x \in \mathbb{R}$.

Claim 2. By the claim 1 of Lemma 4.6, we have $\tau_y \wedge n \leq T_y \wedge n$. Since $(|M_n|)_{n \geq 0}$ is a submartingale, with $z = y + \rho x$,

$$\mathbb{E}_x(|M_{\tau_y}; \tau_y \leq n) \leq \mathbb{E}_x(|M_{\tau_y \wedge n}|) \leq \mathbb{E}_x(|M_{T_y \wedge n}|) \leq 2|z| + 2\mathbb{E}_x(|M_{T_y}; T_y \leq n).$$

The Lebesgue monotone convergence theorem implies the claim 2. □

It follows from the claim 2 of Lemma 5.4 that, under Condition 1 and $\mathbb{E}(a) < 0$, the function

$$V(x, y) = -\mathbb{E}_x(M_{\tau_y})$$

is well defined for any $x \in \mathbb{R}$ and $y > 0$. This also implies the claim 1 of Theorem 2.1 when $\mathbb{E}(a) < 0$. To prove the positivity of the function V on $\mathbb{R} \times \mathbb{R}_+^*$, we also consider the function

$$W(x, y) = -\mathbb{E}_x(M_{T_y}),$$

which is well defined on $\mathbb{R} \times \mathbb{R}$ by the claim 1 of Lemma 5.4. Note that W exists under solely Condition 1.

Proposition 5.5. *Assume Condition 1.*

(1) For any $x \in \mathbb{R}$ and $y \in \mathbb{R}$,

$$W(x, y) = \lim_{n \rightarrow +\infty} \mathbb{E}_x(y + \rho x + M_n; T_y > n).$$

(2) For any $x \in \mathbb{R}$, the function $W(x, \cdot)$ is non-decreasing.

(3) For any $p \in (2, \alpha)$, there exists $\varepsilon_1 > 0$ such that for any $\varepsilon \in (0, \varepsilon_1]$, $k_0 \geq 2$, $x \in \mathbb{R}$ and $y \in \mathbb{R}$,

$$\max(0, y + \rho x) \leq W(x, y) \leq \left(1 + \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) \max(y, 0) + c_{p,\varepsilon}|x| + c_{p,\varepsilon}\sqrt{k_0} + c_{p,\varepsilon}e^{-c_{p,\varepsilon}k_0^\varepsilon}|x|^p.$$

(4) For any $x \in \mathbb{R}$,

$$\lim_{y \rightarrow +\infty} \frac{W(x, y)}{y} = 1.$$

(5) For any $x \in \mathbb{R}$ and $y \in \mathbb{R}$,

$$W(x, y) = \mathbb{E}_x(W(X_1, y + S_1); T_y > 1),$$

and $(W(X_n, y + S_n)\mathbb{1}_{\{T_y > n\}})_{n \geq 0}$ is a martingale.

Proof. The proof is very close to that of Proposition 5.2. The upper bound of the claim 3 is obtained taking the limit as $n \rightarrow +\infty$ in Lemma 4.8. We prove the claim 4 taking the limit as $y \rightarrow +\infty$ and then as $k_0 \rightarrow +\infty$ in the inequality of the claim 3. The proof of the claim 5 is the same as that of the claim 1 of Proposition 5.3. \square

Turning now to V , we have the following proposition.

Proposition 5.6. Assume Condition 1 and $\mathbb{E}(a) < 0$.

(1) For any $x \in \mathbb{R}$ and $y > 0$,

$$V(x, y) = \lim_{n \rightarrow +\infty} \mathbb{E}_x(y + \rho x + M_n; \tau_y > n) = \lim_{n \rightarrow +\infty} \mathbb{E}_x(y + S_n; \tau_y > n).$$

(2) For any $x \in \mathbb{R}$, the function $V(x, \cdot)$ is non-decreasing.

(3) For any $p \in (2, \alpha)$, $\delta > 0$, $x \in \mathbb{R}$ and $y > 0$,

$$0 \leq V(x, y) \leq W(x, y) \leq (1 + c_{p,\delta})y + c_{p,\delta}(1 + |x|^p).$$

(4) The function V is \mathbf{Q}_+ -harmonic on $\mathbb{R} \times \mathbb{R}_+^*$: for any $x \in \mathbb{R}$ and $y > 0$,

$$\mathbf{Q}_+V(x, y) = V(x, y)$$

and $(V(X_n, y + S_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ is a martingale.

Proof. The proofs of the claims 1, 2, 4 and of the lower bound of the claim 3, being similar to that of the previous proposition and of the Proposition 5.2, is left to the reader. The upper bound of the claim 3 is a consequence of the fact that $\tau_y \leq T_y$ (claim 1 of Lemma 4.6): with $z = y + \rho x$,

$$\begin{aligned} V(x, y) &= \lim_{n \rightarrow +\infty} \mathbb{E}_x(z + M_n; \tau_y > n) \\ &\leq \lim_{n \rightarrow +\infty} \mathbb{E}_x(z + M_n; T_y > n) = W(x, y). \end{aligned}$$

\square

Our next goal is to prove that $V(x, y) \geq \max(0, (1 - \delta)y - c_{p,\delta}(1 + |x|^p))$ from which we will deduce the positivity of V . For this we make appropriate adjustments to the proof of Lemmas 4.3 and 4.4 where the submartingale $((y + \rho x + M_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$ will be replaced by the supermartingale $(W(X_n, y + S_n)\mathbb{1}_{\{\tau_y > n\}})_{n \geq 0}$. Instead of upper bounds in Lemmas 4.3 and 4.4 the following two lemmas establish lower bounds.

Lemma 5.7. *Assume Condition 1 and $\mathbb{E}(a) < 0$. For any $p \in (2, \alpha)$, there exists $\varepsilon_1 > 0$ such that for any $\varepsilon \in (0, \varepsilon_1]$, $x \in \mathbb{R}$, $y > 0$ and $n \in \mathbb{N}$,*

$$\mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) \geq W(x, y) - c_{p,\varepsilon} n^{1/2-2\varepsilon} - c_{p,\varepsilon} |x|^p.$$

Proof. By the claim 1 of Lemma 4.6 and the claim 5 of Lemma 5.5, as in the proof of Lemma 4.3,

$$\mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) = W(x, y) - \mathbb{E}_x(W(X_{\tau_y}, y + S_{\tau_y}); T_y > \tau_y, \tau_y \leq n).$$

Using the claim 3 of Proposition 5.5 and the fact that $y + S_{\tau_y} \leq 0$,

$$\begin{aligned} & \mathbb{E}_x(W(X_{\tau_y}, y + S_{\tau_y}); T_y > \tau_y, \tau_y \leq n) \\ & \leq \mathbb{E}_x(c_{p,\varepsilon} |X_{\tau_y}| + c_{p,\varepsilon} \sqrt{k_0} + c_{p,\varepsilon} e^{-c_{p,\varepsilon} k_0^\varepsilon} |X_{\tau_y}|^p; \tau_y \leq n). \end{aligned}$$

Taking $k_0 = \lceil n^{1-4\varepsilon} \rceil$, the end of the proof is the same as the proof of Lemma 4.3. \square

Lemma 5.8. *Assume Condition 1 and $\mathbb{E}(a) < 0$. For any $p \in (2, \alpha)$ there exists $\varepsilon_1 > 0$ such that for any $\varepsilon \in (0, \varepsilon_1]$, $k_0 \geq 2$, $x \in \mathbb{R}$ and $y > 0$,*

$$\mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) \geq y \left(1 - \frac{c_{p,\varepsilon}}{k_0^\varepsilon}\right) - c_{p,\varepsilon} k_0^2 (1 + |x|^p).$$

Proof. The proof is similar to that of Lemma 4.4. With $v_n^\varepsilon = v_n + \lceil n^\varepsilon \rceil$, we have

$$J_0 = \mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) \geq \mathbb{E}_x(W(X_n, y + S_n); \tau_y > n, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil).$$

Using the Markov property, Lemma 5.7 and the fact that $n - v_n^\varepsilon \leq n$,

$$J_0 \geq \mathbb{E}_x(W(X_{v_n^\varepsilon}, y + S_{v_n^\varepsilon}) - c_{p,\varepsilon} n^{1/2-2\varepsilon} - c_{p,\varepsilon} |X_{v_n^\varepsilon}|^p; \tau_y > v_n^\varepsilon, v_n^\varepsilon \leq \lceil n^{1-\varepsilon} \rceil).$$

By the claim 1 of Lemma 4.6, on $\{\tau_y > v_n\}$ we have $z + M_{v_n} > n^{1/2-\varepsilon}$, where $z = y + \rho x$. Moreover, using the fact that $(W(X_n, y + S_n) \mathbb{1}_{\{T_y > n\}})_{n \geq 1}$ is a non-negative martingale (claim 3 and 5 of Proposition 5.5) and the fact that $\tau_y \leq T_y$ a.s. (claim 1 of Lemma 4.6) we can see that $(W(X_n, y + S_n) \mathbb{1}_{\{\tau_y > n\}})_{n \geq 1}$ is a supermartingale. From this and as in the bound of the term J_2 of Lemma 4.4, we obtain that

$$\begin{aligned} J_0 & \geq \mathbb{E}_x(W(X_{\lceil n^{1-\varepsilon} \rceil}, y + S_{\lceil n^{1-\varepsilon} \rceil}); \tau_y > \lceil n^{1-\varepsilon} \rceil) \\ & \quad - \mathbb{E}_x(W(X_{\lceil n^{1-\varepsilon} \rceil}, y + S_{\lceil n^{1-\varepsilon} \rceil}); \tau_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon > \lceil n^{1-\varepsilon} \rceil) \\ & \quad - \frac{c_{p,\varepsilon}}{n^\varepsilon} \mathbb{E}_x(z + M_{v_n}; T_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil) - c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^\varepsilon} (1 + |x|^p). \end{aligned} \tag{5.1}$$

Using the claim 3 of Proposition 5.5 with $k_0 = n$ and the martingale representation (3.2), the absolute value of the second term in the r.h.s. of (5.1) does not exceed

$$c_{p,\varepsilon} \mathbb{E}_x(z + M_{\lceil n^{1-\varepsilon} \rceil} + \sqrt{n} + |X_{\lceil n^{1-\varepsilon} \rceil}| + e^{-c_{p,\varepsilon} n^\varepsilon} |X_{\lceil n^{1-\varepsilon} \rceil}|^p; \tau_y > \lceil n^{1-\varepsilon} \rceil, v_n^\varepsilon > \lceil n^{1-\varepsilon} \rceil).$$

Since $((z + M_n) \mathbb{1}_{\{T_y > n\}})_{n \geq 0}$ is a submartingale, by claim 2 of Lemma 4.6, the absolute value of the third term is less than

$$\frac{c_{p,\varepsilon}}{n^\varepsilon} \mathbb{E}_x(z + M_n; T_y > n).$$

These bounds imply

$$\begin{aligned}
 J_0 &\geq \mathbb{E}_x(W(X_{[n^{1-\varepsilon}]}, y + S_{[n^{1-\varepsilon}]}); \tau_y > [n^{1-\varepsilon}]) \\
 &\quad - c_{p,\varepsilon} \mathbb{E}_x(z + M_{[n^{1-\varepsilon}]} + \sqrt{n} + |X_{[n^{1-\varepsilon}]}|; \tau_y > [n^{1-\varepsilon}], v_n^\varepsilon > [n^{1-\varepsilon}]) \\
 &\quad - c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^\varepsilon} \mathbb{E}_x(|X_{[n^{1-\varepsilon}]}|^p; \tau_y > [n^{1-\varepsilon}], v_n^\varepsilon > [n^{1-\varepsilon}]) \\
 &\quad - \frac{c_{p,\varepsilon}}{n^\varepsilon} \mathbb{E}_x(z + M_n; T_y > n) - c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^\varepsilon} (1 + |x|^p).
 \end{aligned} \tag{5.2}$$

Using the Markov property with the intermediate time $m_\varepsilon = [n^{1-\varepsilon}] - [n^\varepsilon]$, Lemmas 3.4 and 3.1 and the fact that $v_n^\varepsilon = v_n + [n^\varepsilon]$, the absolute value of the second term in the r.h.s. of (5.2) is bounded by

$$c_{p,\varepsilon} \mathbb{E}_x(|z + M_{m_\varepsilon}| + cn^{\varepsilon/2} + c|X_{m_\varepsilon}| + \sqrt{n} + c(1 + |X_{m_\varepsilon}|); \tau_y > m_\varepsilon, v_n > m_\varepsilon),$$

which, in turn, using the fact that $z + M_{m_\varepsilon} \leq n^{1/2-\varepsilon}$ on $\{v_n > m_\varepsilon\}$, is less than

$$c_{p,\varepsilon} \mathbb{E}_x(\sqrt{n} + |X_{m_\varepsilon}|; \tau_y > m_\varepsilon, v_n > m_\varepsilon).$$

The absolute value of the third term in the r.h.s. of (5.2) is bounded using Lemma 3.1 by $c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^\varepsilon} (1 + |x|^p)$. The fourth term is bounded by Lemma 4.8. Collecting these bounds, we obtain

$$\begin{aligned}
 J_0 &\geq \mathbb{E}_x(W(X_{[n^{1-\varepsilon}]}, y + S_{[n^{1-\varepsilon}]}); \tau_y > [n^{1-\varepsilon}]) \\
 &\quad - c_{p,\varepsilon} \mathbb{E}_x(\sqrt{n} + |X_{m_\varepsilon}|; \tau_y > m_\varepsilon, v_n > m_\varepsilon) - \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + y + |x|^p).
 \end{aligned} \tag{5.3}$$

Coupling the Hölder inequality with Lemma 3.1 and Lemma 4.1, we find that the second term in the r.h.s. of (5.3) does not exceed

$$c_{p,\varepsilon} (\sqrt{n} + \mathbb{E}_x^{1/p}(|X_{m_\varepsilon}|^p)) \mathbb{P}_x^{1/q} \left(v_n > \frac{n^{1-\varepsilon}}{c_\varepsilon} \right) \leq c_{p,\varepsilon} (\sqrt{n} + |x|) \frac{c_{p,\varepsilon} (1 + |x|)^{p-1}}{n^{\frac{p-1}{2} - (p-1)\varepsilon}}.$$

Implementing this into (5.3),

$$J_0 \geq \mathbb{E}_x(W(X_{[n^{1-\varepsilon}]}, y + S_{[n^{1-\varepsilon}]}); \tau_y > [n^{1-\varepsilon}]) - \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + y + |x|^p).$$

Since $(W(X_n, y + S_n) \mathbb{1}_{\{\tau_y > n\}})_{n \geq 1}$ is a supermartingale, Lemma A.2 implies that

$$J_0 \geq \mathbb{E}_x(W(X_{k_0}, y + S_{k_0}); \tau_y > k_0) - \frac{c_{p,\varepsilon}}{k_0^\varepsilon} (1 + y + |x|^p).$$

Using the lower bound of the claim 3 of Proposition 5.5 and Lemma 3.4, we deduce that

$$\mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) \geq y \mathbb{P}_x(\tau_y > k_0) - y \frac{c_{p,\varepsilon}}{k_0^\varepsilon} - c_{p,\varepsilon} \sqrt{k_0} - c_{p,\varepsilon} |x|^p.$$

Now, when $y \rightarrow +\infty$, one can see that $\mathbb{P}_x(\tau_y > k_0) \rightarrow 1$: more precisely,

$$\mathbb{P}_x(\tau_y > k_0) \geq \mathbb{P}_x \left(\max_{1 \leq k \leq k_0} |X_k| < \frac{y}{k_0} \right) \geq 1 - c \frac{k_0^2 (1 + |x|)}{y}.$$

Finally,

$$\mathbb{E}_x(W(X_n, y + S_n); \tau_y > n) \geq y \left(1 - \frac{c_{p,\varepsilon}}{k_0^\varepsilon} \right) - c_{p,\varepsilon} k_0^2 (1 + |x|^p). \quad \square$$

Under Condition 3 we use Lemma 5.8 to prove that V is positive on $\mathbb{R} \times \mathbb{R}_+^*$.

Proposition 5.9. *Assume Conditions 1 and $\mathbb{E}(a) < 0$.*

(1) *For any $\delta > 0$, $p \in (2, \alpha)$, $x \in \mathbb{R}$, $y > 0$,*

$$V(x, y) \geq (1 - \delta)y - c_{p,\delta}(1 + |x|^p).$$

(2) *For any $x \in \mathbb{R}$,*

$$\lim_{y \rightarrow +\infty} \frac{V(x, y)}{y} = 1.$$

(3) *If in addition we assume Condition 3, then the function V is positive on $\mathbb{R} \times \mathbb{R}_+^*$.*

Proof. *Claim 1.* Using the claim 1 of Lemma 4.6 and the claims 3 and 5 of Proposition 5.5, with $z = y + \rho x$, we write

$$\begin{aligned} & \mathbb{E}_x(z + M_n; \tau_y > n) \\ & \geq \mathbb{E}_x(z + M_n; T_y > n) - \mathbb{E}_x(W(X_n, y + S_n); T_y > n, \tau_y \leq n) \\ & = \mathbb{E}_x(z + M_n; T_y > n) - W(x, y) + \mathbb{E}_x(W(X_n, y + S_n); \tau_y > n). \end{aligned}$$

Using Lemma 5.8, the claim 1 of Proposition 5.5 and the claim 1 of Proposition 5.6, we obtain

$$V(x, y) \geq y \left(1 - \frac{c_{p,\varepsilon}}{k_0^\varepsilon} \right) - c_{p,\varepsilon} k_0^2 (1 + |x|^p).$$

Taking k_0 large enough, the claim 1 is proved.

Claim 2. Taking the limit as $y \rightarrow +\infty$ and as $\delta \rightarrow 0$ in the claim 1, we obtain first that $\liminf_{y \rightarrow +\infty} V(x, y)/y \geq 1$. By the claim 3 of Proposition 5.6, we obtain also that $\limsup_{y \rightarrow +\infty} V(x, y)/y \leq 1$.

Claim 3. Fix $x \in \mathbb{R}$, $y > 0$ and $\delta_0 > 0$. By Condition 3, there exists $p_0 \in (2, \alpha)$ such that for any $c > 0$ there exists $n_0 \geq 1$ such that $\mathbb{P}_x((X_{n_0}, y + S_{n_0}) \in K_{p_0,c}, \tau_y > n_0) > 0$. Thus, using the claim 4 of Proposition 5.6,

$$V(x, y) \geq \mathbb{E}_x(V(X_{n_0}, y + S_{n_0}); (X_{n_0}, y + S_{n_0}) \in K_{p_0,c}, \tau_y > n_0).$$

Using the claim 1 with $p = p_0$ and $\delta = 1/2$ and choosing the constant $c = 2c_{p_0,\delta} + 2\delta_0$, there exists n_0 such that

$$V(x, y) \geq \delta_0 \mathbb{P}_x((X_{n_0}, y + S_{n_0}) \in K_{p_0,c}, \tau_y > n_0) > 0. \quad \square$$

6. Asymptotic for the exit time

The aim of this section is to prove Theorem 2.2. The asymptotic for the exit time of the Markov walk $(y + S_n)_{n \geq 0}$ will be deduced from the asymptotic of the exit time for the Brownian motion in Corollary A.4 using the functional approximation in Proposition A.5.

6.1. Auxiliary statements

We start by proving an analogue of Corollaries 4.5 and 4.9, where n is replaced by the stopping time ν_n defined by (4.1).

Lemma 6.1. *Assume Condition 1. For any $p \in (2, \alpha)$, there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$, $y > 0$ and $n \geq 1$,*

$$E_1 = \mathbb{E}_x(y + S_{\nu_n}; \tau_y > \nu_n, \nu_n \leq \lceil n^{1-\varepsilon} \rceil) \leq c_{p,\varepsilon} (1 + y + |x|)(1 + |x|)^{p-1}.$$

Proof. When $\tau_y > \nu_n > 1$, we note that

$$0 < X_{\nu_n} < y + S_{\nu_n}. \quad (6.1)$$

Using the martingale representation (3.2) and (6.1), we have

$$y + S_{\nu_n} \leq z + M_{\nu_n} + \max(0, -\rho)X_{\nu_n} \leq z + M_{\nu_n} + \max(0, -\rho)(y + S_{\nu_n}),$$

with $z = y + \rho x$, and so

$$0 < y + S_{\nu_n} \leq \frac{1}{1 - \max(0, -\rho)}(z + M_{\nu_n}) \leq 2(z + M_{\nu_n}).$$

Consequently, using Lemma 3.1 when $\nu_n = 1$,

$$\begin{aligned} E_1 &\leq c(1 + y + |x|) + c\mathbb{E}_x(z + M_{\nu_n}; \tau_y > \nu_n, 1 < \nu_n \leq [n^{1-\varepsilon}]) \\ &\leq c(1 + y + |x|) + \underbrace{c\mathbb{E}_x(z + M_{\nu_n}; \tau_y > \nu_n, \nu_n \leq [n^{1-\varepsilon}])}_{E'_1}. \end{aligned} \quad (6.2)$$

Now, denoting $\nu_n \wedge [n^{1-\varepsilon}] = \min(\nu_n, [n^{1-\varepsilon}])$, we write

$$\begin{aligned} E'_1 &= c\mathbb{E}_x(z + M_{\nu_n \wedge [n^{1-\varepsilon}]} - c\mathbb{E}_x(z + M_{\nu_n \wedge [n^{1-\varepsilon}]}; \tau_y \leq \nu_n \wedge [n^{1-\varepsilon}])) \\ &\quad - c\mathbb{E}_x(z + M_{[n^{1-\varepsilon}]}; \tau_y > [n^{1-\varepsilon}], \nu_n > [n^{1-\varepsilon}])). \end{aligned}$$

Since $(M_n)_{n \geq 0}$ is a centred martingale, using Lemma 5.1 when $\mathbb{E}(a) \geq 0$ and the claim 2 of Lemma 5.4 when $\mathbb{E}(a) < 0$, Lemmas 3.4, 4.1 and Hölder inequality, we obtain

$$E'_1 \leq c_{p,\varepsilon}(1 + y + |x|)(1 + |x|)^{p-1}.$$

Implementing this into (6.2), it concludes the proof. \square

Now, we can prove an upper bound of order $1/n^{1/2-\varepsilon}$ of the probability of survival $\mathbb{P}_x(\tau_y > n)$.

Lemma 6.2. *Assume Condition 1. For any $p \in (2, \alpha)$, there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$, $y > 0$ and $n \geq 1$,*

$$\mathbb{P}_x(\tau_y > n) \leq c_{p,\varepsilon} \frac{(1 + y + |x|)(1 + |x|)^{p-1}}{n^{1/2-\varepsilon}}.$$

Moreover, summing these bounds, we have

$$\sum_{k=1}^{[n^{1-\varepsilon}]} \mathbb{P}_x(\tau_y > k) \leq c_{p,\varepsilon}(1 + y + |x|)(1 + |x|)^{p-1}n^{1/2+\varepsilon}.$$

Proof. We write

$$\mathbb{P}_x(\tau_y > n) \leq \mathbb{E}_x\left(\frac{y + S_{\nu_n}}{n^{1/2-\varepsilon}}; \tau_y > \nu_n, \nu_n \leq [n^{1-\varepsilon}]\right) + \mathbb{P}_x(\nu_n > n^{1-\varepsilon}).$$

Using Lemma 6.1 and Lemma 4.1, the claim follows. \square

Before to proceed with the proof of Theorem 2.2, we need two additional technical lemmas. Recall the notation $\nu_n^{\varepsilon/6} = \nu_n + [n^{\varepsilon/6}]$.

Lemma 6.3. *Assume Condition 1. There exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $y > 0$,*

$$E_2 = \mathbb{E}_x \left(y + S_{v_n^{\varepsilon/6}}; \tau_y > v_n^{\varepsilon/6}, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}] \right) \xrightarrow{n \rightarrow +\infty} V(x, y).$$

Proof. Using the martingale approximation (3.2),

$$E_2 = \underbrace{-\rho \mathbb{E}_x \left(X_{v_n^{\varepsilon/6}}; \tau_y > v_n^{\varepsilon/6}, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}] \right)}_{=: E_{21}} + \underbrace{\mathbb{E}_x \left(z + M_{v_n^{\varepsilon/6}}; \tau_y > v_n^{\varepsilon/6}, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}] \right)}_{=: E_{22}}. \quad (6.3)$$

Bound of E_{21} . By the Markov property, Lemma 3.1 and the fact that $(y + S_{v_n})/n^{1/2-\varepsilon} > 1$,

$$\begin{aligned} |E_{21}| &\leq c \mathbb{E}_x \left(1 + e^{-cn^{\varepsilon/6}} |X_{v_n}|; \tau_y > v_n, v_n \leq [n^{1-\varepsilon}] \right) \\ &\leq \frac{c}{n^{1/2-\varepsilon}} E_1 + c e^{-cn^{\varepsilon/6}} \sum_{k=1}^{[n^{1-\varepsilon}]} \mathbb{E}_x(|X_k|). \end{aligned}$$

By Lemma 6.1, we obtain

$$|E_{21}| \leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{1/2-\varepsilon}}. \quad (6.4)$$

Bound of E_{22} . We proceed in the same way as for bounding E'_1 defined in (6.2):

$$\begin{aligned} E_{22} &= z - \mathbb{E}_x \left(z + M_{\tau_y}; \tau_y \leq v_n^{\varepsilon/6} \wedge [n^{1-\varepsilon}] \right) \\ &\quad - \mathbb{E}_x \left(z + M_{v_n^{\varepsilon/6} \wedge [n^{1-\varepsilon}]}; \tau_y > v_n^{\varepsilon/6} \wedge [n^{1-\varepsilon}], v_n^{\varepsilon/6} > [n^{1-\varepsilon}] \right). \end{aligned}$$

By the Hölder inequality, Lemma 3.4 and Lemma 4.1,

$$E_{22} \leq z - \mathbb{E}_x \left(z + M_{\tau_y}; \tau_y \leq v_n^{\varepsilon/6} \wedge [n^{1-\varepsilon}] \right) + c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_{p,\varepsilon}}}. \quad (6.5)$$

Since $v_n^{\varepsilon/6} \geq [n^{\varepsilon/6}] \rightarrow +\infty$ as $n \rightarrow +\infty$ and M_{τ_y} is integrable (using Lemma 5.1 when $\mathbb{E}(a) \geq 0$ and the claim 2 of Lemma 5.4 when $\mathbb{E}(a) < 0$), by the Lebesgue dominated convergence we deduce that

$$\lim_{n \rightarrow +\infty} E_{22} = -\mathbb{E}_x(M_{\tau_y}) = V(x, y).$$

Coupling this with equations (6.3) and (6.4), we conclude that $E_2 \xrightarrow{n \rightarrow +\infty} V(x, y)$. □

Lemma 6.4. *Assume Condition 1. There exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $y > 0$,*

$$E_3 = \mathbb{E}_x \left(y + S_{v_n^{\varepsilon/6}}; y + S_{v_n^{\varepsilon/6}} > n^{1/2-\varepsilon/6}, \tau_y > v_n^{\varepsilon/6}, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}] \right) \xrightarrow{n \rightarrow +\infty} 0.$$

Proof. The first step of the proof consists in proving that we can replace the time $v_n^{\varepsilon/6}$ in the definition of E_3 by the time v_n . More precisely, we shall prove that the following bound holds true:

$$\begin{aligned} E_3 &\leq cn^{\varepsilon/6} \underbrace{\mathbb{E}_x \left(y + S_{v_n}; y + S_{v_n} > n^{1/2-\varepsilon/2}, \tau_y > v_n, v_n \leq [n^{1-\varepsilon}] \right)}_{=: E_{31}} \\ &\quad + c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\varepsilon/6}}. \end{aligned} \quad (6.6)$$

To this end, we bound E_3 as follows:

$$\begin{aligned}
 E_3 &\leq E_{31} + \underbrace{\mathbb{E}_x(|S_{v_n^{\varepsilon/6}} - S_{v_n}|; y + S_{v_n} > n^{1/2-\varepsilon/2}; \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil)}_{=:E_{32}} \\
 &\quad + \underbrace{\mathbb{E}_x(y + S_{v_n}; y + S_{v_n} \leq n^{1/2-\varepsilon/2}, y + S_{v_n^{\varepsilon/6}} > n^{1/2-\varepsilon/6}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil)}_{=:E_{33}} \\
 &\quad + \underbrace{\mathbb{E}_x(|S_{v_n^{\varepsilon/6}} - S_{v_n}|; y + S_{v_n} \leq n^{1/2-\varepsilon/2}, y + S_{v_n^{\varepsilon/6}} > n^{1/2-\varepsilon/6}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil)}_{=:E_{34}}. \tag{6.7}
 \end{aligned}$$

Bound of E_{32} . By the Markov property and Lemma 3.1,

$$\begin{aligned}
 E_{32} &\leq \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{E}_{x'}(|S_{\lceil n^{\varepsilon/6} \rceil}|) \mathbb{P}_x(X_{v_n} \in dx', y + S_{v_n} \in dy', y + S_{v_n} > n^{1/2-\varepsilon/2}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil) \\
 &\leq \mathbb{E}_x(cn^{\varepsilon/6}(1 + |X_{v_n}|); y + S_{v_n} > n^{1/2-\varepsilon/2}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil).
 \end{aligned}$$

If $\tau_y > v_n > 1$, by (6.1), we have $|X_{v_n}| = X_{v_n} < y + S_{v_n}$. Using this bound when $v_n > 1$ and the Markov inequality when $v_n = 1$,

$$\begin{aligned}
 E_{32} &\leq \mathbb{E}_x(cn^{\varepsilon/6}(1 + |X_1|); y + X_1 > n^{1/2-\varepsilon/2}, v_n = 1) + cn^{\varepsilon/6} E_{31} \\
 &\leq c \frac{(1 + y + |x|)(1 + |x|)}{n^{1/2-c\varepsilon}} + cn^{\varepsilon/6} E_{31}. \tag{6.8}
 \end{aligned}$$

Bound of E_{33} . By the Markov property,

$$\begin{aligned}
 E_{33} &\leq \int_{\mathbb{R} \times \mathbb{R}_+^*} y' \mathbb{P}_{x'}(y' + S_{\lceil n^{\varepsilon/6} \rceil} > n^{1/2-\varepsilon/6}) \\
 &\quad \times \mathbb{P}_x(X_{v_n} \in dx', y + S_{v_n} \in dy', y + S_{v_n} \leq n^{1/2-\varepsilon/2}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil).
 \end{aligned}$$

When $y' \leq n^{1/2-\varepsilon/2}$, by the Markov inequality, we have,

$$\mathbb{P}_{x'}(y' + S_{\lceil n^{\varepsilon/6} \rceil} > n^{1/2-\varepsilon/6}) \leq \mathbb{P}_{x'}\left(|S_{\lceil n^{\varepsilon/6} \rceil}| > \frac{n^{1/2-\varepsilon/6}}{c_\varepsilon}\right) \leq \frac{c_\varepsilon n^{\varepsilon/6}(1 + |x'|)}{n^{1/2-\varepsilon/6}}.$$

On the event $\{y + S_{v_n} \leq n^{1/2-\varepsilon/2}, \tau_y > v_n\}$, we obviously have $x' = X_{v_n} \leq n^{1/2-\varepsilon/2}$. From these bounds, using the positivity of X_{v_n} for $v_n > 1$, see (6.1), we obtain

$$E_{33} \leq \mathbb{E}_x\left((y + S_1) \frac{c_\varepsilon(1 + |X_1|)}{n^{1/2-\varepsilon/3}}; v_n = 1\right) + \frac{c_\varepsilon}{n^{\varepsilon/2-\varepsilon/3}} E_1.$$

By Lemma 6.1, we obtain

$$E_{33} \leq c_{p,\varepsilon} \frac{(1 + y + |x|)(1 + |x|)^{p-1}}{n^{\varepsilon/6}}. \tag{6.9}$$

Bound of E_{34} . Again, by the Markov property,

$$\begin{aligned}
 E_{34} &\leq \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{E}_{x'}(|S_{\lceil n^{\varepsilon/6} \rceil}|; y' + S_{\lceil n^{\varepsilon/6} \rceil} > n^{1/2-\varepsilon/6}) \\
 &\quad \times \mathbb{P}_x(X_{v_n} \in dx', y + S_{v_n} \in dy', y + S_{v_n} \leq n^{1/2-\varepsilon/2}, \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil).
 \end{aligned}$$

When $y' \leq n^{1/2-\varepsilon/2}$, using the Markov inequality and Lemma 3.1, we have

$$\mathbb{E}_{x'}(|S_{[n^{\varepsilon/6}]}; y' + S_{[n^{\varepsilon/6}]} > n^{1/2-\varepsilon/6}) \leq \mathbb{E}_{x'}\left(\frac{c_\varepsilon^{p-1}|S_{[n^{\varepsilon/6}]}|^p}{n^{\frac{p-1}{2}-(p-1)\varepsilon/6}}\right) \leq c_{p,\varepsilon} \frac{(1+|x'|)^p}{n^{\frac{p-1}{2}-c_p\varepsilon}}.$$

Then, using Lemma 3.1 again and the Markov property for the terms in the last sum of the first line below,

$$\begin{aligned} E_{34} &\leq \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-c_p\varepsilon}} + \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{k=1}^{[n^\varepsilon]} \mathbb{E}_x(|X_k|^p) + \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x(|X_k|^p; \tau_y > k) \\ &\leq \frac{c_{p,\varepsilon}(1+|x|^p)}{n^{\frac{p-1}{2}-c_p\varepsilon}} + \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{k=1}^{[n^{1-\varepsilon}]-[n^\varepsilon]} \mathbb{E}_x(1 + e^{-c_p n^\varepsilon} |X_k|^p; \tau_y > k) \\ &\leq \frac{c_{p,\varepsilon}(1+|x|^p)}{n^{\frac{p-1}{2}-c_p\varepsilon}} + c_{p,\varepsilon} e^{-c_p n^\varepsilon} (1+|x|^p) + \frac{c_{p,\varepsilon}}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{k=1}^{[n^{1-\varepsilon}]} \mathbb{P}_x(\tau_y > k). \end{aligned}$$

Using the second bound in Lemma 6.2, and taking $\varepsilon > 0$ small enough, we obtain

$$E_{34} \leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_p\varepsilon}} \xrightarrow{n \rightarrow +\infty} 0. \tag{6.10}$$

Inserting (6.8), (6.9) and (6.10) in (6.7), we conclude the proof of (6.6).

Bound of $cn^{\varepsilon/6}E_{31}$. Note that, when $v_n > 1$ and $y + S_{v_n} > n^{1/2-\varepsilon/2}$, we have $X_{v_n} = y + S_{v_n} - (y + S_{v_n-1}) > n^{1/2-\varepsilon/2} - n^{1/2-\varepsilon} \geq \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon}$. Consequently,

$$\begin{aligned} cn^{\varepsilon/6}E_{31} &\leq \underbrace{cn^{\varepsilon/6}\mathbb{E}_x(y + S_{v_n}; v_n \leq [n^\varepsilon])}_{=:E_{35}} \\ &\quad + \underbrace{cn^{\varepsilon/6}\mathbb{E}_x\left(y + S_{v_n}; X_{v_n} > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon}, \tau_y > v_n, [n^\varepsilon] < v_n \leq [n^{1-\varepsilon}]\right)}_{=:E_{36}}. \end{aligned} \tag{6.11}$$

Bound of E_{35} . Using the definition of v_n , the Markov inequality and Lemma 3.1,

$$\begin{aligned} E_{35} &\leq cn^{\varepsilon/6}\mathbb{E}_x\left(\max_{k \leq [n^\varepsilon]} |y + S_k|; \max_{k \leq [n^\varepsilon]} |y + S_k| > n^{1/2-\varepsilon}\right) \\ &\leq \frac{c_p(1+y+|x|)^2}{n^{1/2-c_p\varepsilon}}. \end{aligned} \tag{6.12}$$

Bound of E_{36} . The idea is based on the observation that, according to the first bound in Lemma 3.1, the random variables $y + S_{v_n-[n^\varepsilon]}$ and X_{v_n} are ‘‘almost’’ independent. In this line, summing over the values of v_n and bounding the indicators $\mathbb{1}_{\{v_n=k\}}$ by 1, we write

$$\begin{aligned} E_{36} &\leq cn^{\varepsilon/6} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x\left(y + S_{k-[n^\varepsilon]}; X_k > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon}, \tau_y > k\right) \\ &\quad + cn^{\varepsilon/6} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x\left(|S_k - S_{k-[n^\varepsilon]}|; X_k > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon}, \tau_y > k\right). \end{aligned}$$

By the Markov property,

$$\begin{aligned}
 E_{36} &\leq cn^{\varepsilon/6} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R} \times \mathbb{R}_+^*} y' \mathbb{P}_{x'} \left(X_{[n^\varepsilon]} > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon} \right) \\
 &\quad \times \mathbb{P}_x (X_{k-[n^\varepsilon]} \in dx', y + S_{k-[n^\varepsilon]} \in dy', \tau_y > k - [n^\varepsilon]) \\
 &\quad + cn^{\varepsilon/6} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x \left(n^\varepsilon \max_{k-[n^\varepsilon] \leq i \leq k} |X_i|; X_k > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon}, \tau_y > k \right). \tag{6.13}
 \end{aligned}$$

Recall that, under $\mathbb{P}_{x'}$, by (3.3), $X_{[n^\varepsilon]} = \prod_{i=1}^{[n^\varepsilon]} a_i x' + X_{[n^\varepsilon]}^0$. Then, since a_i 's are independent and identically distributed, by claim 1 of Condition 1 and Lemma 3.1,

$$\begin{aligned}
 \mathbb{P}_{x'} \left(X_{[n^\varepsilon]} > \frac{n^{1/2-\varepsilon/2}}{c_\varepsilon} \right) &\leq \mathbb{P} \left(\prod_{i=1}^{[n^\varepsilon]} a_i x' > \frac{n^{1/2-\varepsilon/2}}{2c_\varepsilon} \right) + \mathbb{P} \left(|X_{[n^\varepsilon]}^0| > \frac{n^{1/2-\varepsilon/2}}{2c_\varepsilon} \right) \\
 &\leq c_\varepsilon e^{-c_\varepsilon n^\varepsilon} |x'| + \frac{c_{p,\varepsilon}}{n^{\frac{p}{2}-c_p\varepsilon}}. \tag{6.14}
 \end{aligned}$$

Inserting (6.14) into (6.13) and using Cauchy–Schwartz inequality, by Corollaries 4.5 and 4.9,

$$\begin{aligned}
 E_{36} &\leq \sum_{j=1}^{[n^{1-\varepsilon}]} \left(c_\varepsilon e^{-c_\varepsilon n^\varepsilon} \mathbb{E}_x^{1/2} (|y + S_j|^2) \mathbb{E}_x^{1/2} (|X_j|^2) + \frac{c_{p,\varepsilon}}{n^{\frac{p}{2}-c_p\varepsilon}} (1+y+|x|)(1+|x|)^{p-1} \right) \\
 &\quad + cn^{\varepsilon/6} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x \left(\frac{\max_{k-[n^\varepsilon] \leq i \leq k} |X_i|^p}{n^{\frac{p-1}{2}-c_p\varepsilon}}; \tau_y > k - [n^\varepsilon] \right).
 \end{aligned}$$

Using the decomposition (3.2) and Lemmas 3.1 and 3.4

$$E_{36} \leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_p\varepsilon}} + \frac{c_p}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{k=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x (n^\varepsilon (1+|X_{k-[n^\varepsilon]}|^p); \tau_y > k - [n^\varepsilon]).$$

Re-indexing $j = k - [n^\varepsilon]$, after some elementary transformations, we get

$$\begin{aligned}
 E_{36} &\leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_p\varepsilon}} + \frac{c_p}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{j=1}^{[n^{1-\varepsilon}]} \mathbb{P}_x (\tau_y > j) \\
 &\quad + \frac{c_p}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{j=1}^{[n^\varepsilon]} \mathbb{E}_x (|X_j|^p) + \frac{c_p}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{j=[n^\varepsilon]+1}^{[n^{1-\varepsilon}]} \mathbb{E}_x (|X_j|^p; \tau_y > j - [n^\varepsilon]).
 \end{aligned}$$

Again using the Markov property, Lemma 3.1 and Lemma 6.2, we have

$$\begin{aligned}
 E_{36} &\leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_p\varepsilon}} + \frac{c_p}{n^{\frac{p-1}{2}-c_p\varepsilon}} \sum_{j=1}^{[n^{1-\varepsilon}]} \mathbb{P}_x (\tau_y > j) + c_p e^{-c_p n^\varepsilon} \sum_{j=1}^{[n^{1-\varepsilon}]} \mathbb{E}_x (|X_j|^p; \tau_y > j) \\
 &\leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{n^{\frac{p-2}{2}-c_p\varepsilon}}.
 \end{aligned}$$

Inserting this bound and (6.12) into (6.11), we obtain

$$cn^{\varepsilon/6} E_{31} \leq \frac{c_{p,\varepsilon}(1+y+|x|)^p}{n^{\frac{p-2}{2}-c_p\varepsilon}}.$$

Together with (6.6), this bound implies that

$$E_3 \leq \frac{c_p(1+y+|x|)^p}{n^{\varepsilon/6}} \xrightarrow{n \rightarrow +\infty} 0. \quad (6.15)$$

□

6.2. Proof of the claim 2 of Theorem 2.2

Assume either Conditions 1, 2 and $\mathbb{E}(a) \geq 0$, or Conditions 1 and 3. Introducing the stopping time $v_n^{\varepsilon/6} = v_n + \lceil n^{\varepsilon/6} \rceil$, we have

$$\mathbb{P}_x(\tau_y > n) = \mathbb{P}_x(\tau_y > n, v_n^{\varepsilon/6} \leq \lceil n^{1-\varepsilon} \rceil) + \mathbb{P}_x(\tau_y > n, v_n^{\varepsilon/6} > \lceil n^{1-\varepsilon} \rceil). \quad (6.16)$$

We bound the second term by Lemma 4.1: for $2 < p < \alpha$,

$$\mathbb{P}_x(\tau_y > n, v_n^{\varepsilon/6} > \lceil n^{1-\varepsilon} \rceil) \leq \mathbb{P}_x\left(v_n > \frac{n^{1-\varepsilon}}{c_\varepsilon}\right) \leq c_{p,\varepsilon} \frac{(1+|x|)^p}{n^{p/2-c_p\varepsilon}} = o\left(\frac{1}{\sqrt{n}}\right). \quad (6.17)$$

To bound the first term, we introduce more notations. Let $(B_t)_{t \geq 0}$ be the Brownian motion from Proposition A.5, A_k be the event $A_k = \{\max_{0 \leq t \leq 1} |S_{[tk]} - \sigma B_{tk}| \leq k^{1/2-2\varepsilon}\}$ where σ is defined by (2.2), and \bar{A}_k be its complement. Using the Markov property, we have

$$\begin{aligned} \mathbb{P}_x(\tau_y > n, v_n^{\varepsilon/6} \leq \lceil n^{1-\varepsilon} \rceil) &= \underbrace{\sum_{k=1}^{\lceil n^{1-\varepsilon} \rceil} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(\tau_{y'} > n-k, \bar{A}_{n-k}) \mathbb{P}_x(X_k \in dx', y+S_k \in dy', \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: J_1} \\ &+ \underbrace{\sum_{k=1}^{\lceil n^{1-\varepsilon} \rceil} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(\tau_{y'} > n-k, A_{n-k}) \mathbb{P}_x(X_k \in dx', y+S_k \in dy', \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: J_2}. \end{aligned} \quad (6.18)$$

Bound of J_1 . Taking into account that $n-k \geq \frac{n}{c_\varepsilon}$ for any $k \leq \lceil n^{1-\varepsilon} \rceil$, by Proposition A.5 with ε small enough, we find

$$\mathbb{P}_{x'}(\tau_{y'} > n-k, \bar{A}_{n-k}) \leq \mathbb{P}_{x'}(\bar{A}_{n-k}) \leq c_{p,\varepsilon}(1+|x'|)^p n^{-2\varepsilon}.$$

By the Markov property and the first bound in Lemma 3.1,

$$J_1 \leq \mathbb{E}_x\left(c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^{\varepsilon/6}} |X_{v_n}|^p + \frac{c_{p,\varepsilon}}{n^{2\varepsilon}}; \tau_y > v_n, v_n \leq \lceil n^{1-\varepsilon} \rceil\right).$$

Since $\frac{y+S_{v_n}}{n^{1/2-\varepsilon}} > 1$, using Lemma 6.1,

$$J_1 \leq c_{p,\varepsilon} e^{-c_{p,\varepsilon} n^{\varepsilon/6}} (1+|x|)^p + \frac{c_{p,\varepsilon}}{n^{1/2-\varepsilon+2\varepsilon}} E_1 \leq \frac{c_{p,\varepsilon}(1+y+|x|)(1+|x|)^{p-1}}{n^{1/2+\varepsilon}}. \quad (6.19)$$

Bound of J_2 . The idea is as follows. When $y' \leq \theta_n \sqrt{n}$, with $\theta_n = n^{-\varepsilon/6}$, we are going to control the probability $\mathbb{P}_{x'}(\tau_{y'} > n - k, A_{n-k})$ in J_2 by the claim 2 of Corollary A.4. When $y' > \theta_n \sqrt{n}$ we shall apply Lemma 6.4. Accordingly, we split J_2 into two terms as follows:

$$\begin{aligned}
 J_2 &= \underbrace{\sum_{k=1}^{\lfloor n^{1-\varepsilon} \rfloor} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(\tau_{y'} > n - k, A_{n-k}) \mathbb{P}_x(X_k \in dx', y + S_k \in dy', y + S_k > n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: J_3} \\
 &+ \underbrace{\sum_{k=1}^{\lfloor n^{1-\varepsilon} \rfloor} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(\tau_{y'} > n - k, A_{n-k}) \mathbb{P}_x(X_k \in dx', y + S_k \in dy', y + S_k \leq n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: J_4}.
 \end{aligned} \tag{6.20}$$

Bound of J_3 . Let $\tau_{y'}^{\text{bm}}$ be the exit time of the Brownian motion defined by (A.10) and $y'_+ = y' + (n - k)^{1/2-2\varepsilon}$. Since

$$\mathbb{P}_{x'}(\tau_{y'} > n - k, A_{n-k}) \leq \mathbb{P}_{x'}(\tau_{y'_+}^{\text{bm}} > n - k), \tag{6.21}$$

using the claim 1 of Corollary A.4 with $y'_+ > 0$, we get

$$J_3 \leq \sum_{k=1}^{\lfloor n^{1-\varepsilon} \rfloor} \mathbb{E}_x \left(c \frac{y + S_k + (n - k)^{1/2-2\varepsilon}}{\sqrt{n - k}}; y + S_k > n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k \right).$$

Since $\frac{c}{\sqrt{n-k}} \leq \frac{c_\varepsilon}{\sqrt{n}}$ and $y + S_k + (n - k)^{1/2-2\varepsilon} \leq 2(y + S_k)$ on the event $\{y + S_k > n^{1/2-\varepsilon/6}\}$, using Lemma 6.4, we have

$$J_3 \leq \frac{c_\varepsilon}{\sqrt{n}} E_3 = o\left(\frac{1}{\sqrt{n}}\right). \tag{6.22}$$

Upper bound of J_4 . Since $\frac{n}{c_\varepsilon} \leq n - k \leq n$, we have $y'_+ \leq c_\varepsilon(n - k)^{1/2-\varepsilon/6}$ when $y' \leq n^{1/2-\varepsilon/6}$. Using (6.21), from the claim 2 of Corollary A.4 with $\theta_m = c_\varepsilon m^{-\varepsilon/6}$, we deduce that

$$\begin{aligned}
 J_4 &\leq \sum_{k=1}^{\lfloor n^{1-\varepsilon} \rfloor} \mathbb{E}_x \left(\frac{2}{\sqrt{2\pi(n - k)}\sigma} (y + S_k + (n - k)^{1/2-2\varepsilon})(1 + c\theta_{n-k}^2); \right. \\
 &\left. y + S_k \leq n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k \right).
 \end{aligned} \tag{6.23}$$

Taking into account that $\frac{1}{\sqrt{n-k}} \leq \frac{1}{\sqrt{n}}(1 + \frac{c_\varepsilon}{n^\varepsilon})$, $\theta_{n-k} \leq \frac{c_\varepsilon}{n^{\varepsilon/6}}$ and $1 < \frac{y+S_k v_n}{n^{1/2-\varepsilon}}$, we obtain

$$J_4 \leq \frac{2}{\sqrt{2\pi n \sigma}} \left(1 + \frac{c_\varepsilon}{n^{\varepsilon/3}} \right) E_2 + \frac{c_\varepsilon}{n^{1/2+\varepsilon}} E_1. \tag{6.24}$$

Using Lemma 6.1 and Lemma 6.3, we get the following upper bound,

$$J_4 \leq \frac{2V(x, y)}{\sqrt{2\pi n \sigma}} (1 + o(1)). \tag{6.25}$$

Lower bound of J_4 . In the same way as for the upper bound of J_4 , with $y'_- = y + S_{v_n^{\varepsilon/6}} - (n - v_n^{\varepsilon/6})^{1/2-2\varepsilon} > 0$ on the event $\{(n - v_n^{\varepsilon/6})^{1/2-2\varepsilon} < y + S_{v_n^{\varepsilon/6}}\}$, we have

$$\begin{aligned} J_4 &\geq \frac{2}{\sqrt{2\pi n\sigma}} \left(1 - \frac{c_\varepsilon}{n^{\varepsilon/3}}\right) \mathbb{E}_x(y'_-; (n - v_n^{\varepsilon/6})^{1/2-2\varepsilon} < y + S_{v_n^{\varepsilon/6}} \leq n^{1/2-\varepsilon/6}, \tau_y > v_n^{\varepsilon/6}, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}]) \\ &\quad - \sum_{k=1}^{\lfloor n^{1-\varepsilon} \rfloor} \int_{\mathbb{R}} \mathbb{P}_{x'}(\bar{A}_{n-k}) \mathbb{P}_x(X_k \in dx', \tau_y > k, v_n^{\varepsilon/6} = k). \end{aligned} \quad (6.26)$$

Using the fact that $-y'_- \geq 0$ on $\{(n - v_n^{\varepsilon/6})^{1/2-2\varepsilon} \geq y + S_{v_n^{\varepsilon/6}}\}$, we obtain in a same way as for the upper bound of J_1 ,

$$\begin{aligned} J_4 &\geq \frac{2}{\sqrt{2\pi n\sigma}} \left(1 - \frac{c_\varepsilon}{n^{\varepsilon/3}}\right) E_2 - \frac{2}{\sqrt{2\pi n\sigma}} \mathbb{E}_x \left(n^{1/2-2\varepsilon} \frac{y + S_{v_n}}{n^{1/2-\varepsilon}}; \tau_y > v_n, v_n \leq [n^{1-\varepsilon}] \right) \\ &\quad - \frac{2}{\sqrt{2\pi n\sigma}} E_3 - \frac{c_{p,\varepsilon}(1+y+|x|)(1+|x|)^{p-1}}{n^{1/2+\varepsilon}} \\ &\geq \frac{2}{\sqrt{2\pi n\sigma}} \left(1 - \frac{c_\varepsilon}{n^{\varepsilon/3}}\right) E_2 - \frac{c}{n^{1/2+\varepsilon}} E_1 - \frac{c}{\sqrt{n}} E_3 - \frac{c_{p,\varepsilon}(1+y+|x|)(1+|x|)^{p-1}}{n^{1/2+\varepsilon}}. \end{aligned}$$

Consequently, using the results of Lemma 6.3, Lemma 6.1 and Lemma 6.4 we conclude that

$$J_4 \geq \frac{2V(x, y)}{\sqrt{2\pi n\sigma}} (1 - o(1)). \quad (6.27)$$

Coupling the obtained lower bound with the upper bound in (6.25) we obtain $J_4 \sim \frac{2V(x, y)}{\sqrt{2\pi n\sigma}}$. With the decomposition of J_2 in (6.20) and the bound of J_3 in (6.22) we get $J_2 \sim \frac{2V(x, y)}{\sqrt{2\pi n\sigma}}$. Finally, the claim 2 of Theorem 2.2 follows from (6.16), (6.17), (6.18) and (6.19).

6.3. Proof of the claim 1 of Theorem 2.2

Assume Condition 1. All the necessary bounds are obtained in the previous Section 6.2. It is easy to see that they hold under solely Condition 1. We highlight how to gather them. By (6.16), (6.17), (6.18) and (6.20), we have,

$$\mathbb{P}_x(\tau_y > n) \leq c_{p,\varepsilon} \frac{(1+|x|^p)}{\sqrt{n}} + J_1 + J_3 + J_4.$$

Then, by (6.19), (6.22), and (6.24),

$$\mathbb{P}_x(\tau_y > n) \leq c_{p,\varepsilon} \frac{(1+y+|x|)(1+|x|)^{p-1}}{\sqrt{n}} + \frac{c_\varepsilon}{\sqrt{n}} E_3 + \frac{c_\varepsilon}{\sqrt{n}} (E_2 + E_1).$$

Now, by Lemma 6.1, (6.3) and (6.15),

$$\mathbb{P}_x(\tau_y > n) \leq c_{p,\varepsilon} \frac{(1+y+|x|)^p}{\sqrt{n}} + \frac{c_\varepsilon}{\sqrt{n}} (E_{21} + E_{22}).$$

Finally, using (6.4), (6.5) and Lemmas 5.1 and 5.4 we have,

$$\begin{aligned} \mathbb{P}_x(\tau_y > n) &\leq \frac{c_\varepsilon}{\sqrt{n}} (z - \mathbb{E}_x(z + M_{\tau_y}; \tau_y \leq v_n^{\varepsilon/6} \wedge [n^{1-\varepsilon}])) + c_{p,\varepsilon} \frac{(1+y+|x|)^p}{\sqrt{n}} \\ &\leq \frac{c_\varepsilon}{\sqrt{n}} \mathbb{E}_x(|M_{\tau_y}|) + c_{p,\varepsilon} \frac{(1+y+|x|)^p}{\sqrt{n}} \leq c_p \frac{(1+y+|x|)^p}{\sqrt{n}}. \end{aligned}$$

6.4. Proof of Corollary 2.3

For any $p > 0$,

$$\mathbb{E}_x(\tau_y^p) = \sum_{k=0}^{+\infty} \mathbb{P}_x(\tau_y > k)((k+1)^p - k^p).$$

Now the first and the second assertions of the corollary follow respectively from the claim 1 and 2 of Theorem 2.2.

7. Asymptotic for conditioned Markov walk

In this section we prove Theorem 2.4. We will deduce the asymptotic of the Markov walk $(y + S_n)_{n \geq 0}$ conditioned to stay positive from the corresponding result for the Brownian motion given by Proposition A.3. As in Section 6, we will use the functional approximation of Proposition A.5. We will refer frequently to Section 6 in order to shorten the exposition.

Proof of Theorem 2.4. Introducing $v_n^{\varepsilon/6} = v_n + [n^{\varepsilon/6}]$ and taking into account Condition 2 or 3, we have

$$\begin{aligned} \mathbb{P}_x(y + S_n \leq t\sqrt{n} \mid \tau_y > n) &= \underbrace{\frac{\mathbb{P}_x(y + S_n \leq t\sqrt{n}, \tau_y > n, v_n^{\varepsilon/6} > [n^{1-\varepsilon}])}{\mathbb{P}_x(\tau_y > n)}}_{=:L_1} \\ &+ \underbrace{\frac{\mathbb{P}_x(y + S_n \leq t\sqrt{n}, \tau_y > n, v_n^{\varepsilon/6} \leq [n^{1-\varepsilon}])}{\mathbb{P}_x(\tau_y > n)}}_{=:L_2}. \end{aligned} \quad (7.1)$$

Bound of L_1 . Using Lemma 4.1 and Theorem 2.2,

$$L_1 \leq \frac{\mathbb{P}_x(v_n > \frac{n^{1-\varepsilon}}{c_\varepsilon})}{\mathbb{P}_x(\tau_y > n)} \leq \frac{c_{p,\varepsilon}(1 + |x|)^p}{n^{\frac{p}{2} - c_{p,\varepsilon}} \mathbb{P}_x(\tau_y > n)} \xrightarrow{n \rightarrow +\infty} 0. \quad (7.2)$$

Bound of L_2 . As in Section 6, setting $A_k = \{\max_{0 \leq t \leq 1} |S_{[tk]} - \sigma B_{tk}| \leq k^{1/2-2\varepsilon}\}$, by the Markov property,

$$\begin{aligned} &\mathbb{P}_x(\tau_y > n)L_2 \\ &= \underbrace{\sum_{k=1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(y' + S_{n-k} \leq t\sqrt{n}, \tau_{y'} > n-k, \bar{A}_{n-k}) \mathbb{P}_x(X_k \in dx', y + S_k \in dy', \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: \mathbb{P}_x(\tau_y > n)L_3} \\ &+ \sum_{k=1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(y' + S_{n-k} \leq t\sqrt{n}, \tau_{y'} > n-k, A_{n-k}) \mathbb{P}_x(X_k \in dx', \\ &\quad \underbrace{y + S_k \in dy', y + S_k > n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: \mathbb{P}_x(\tau_y > n)L_4} \\ &+ \sum_{k=1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'}(y' + S_{n-k} \leq t\sqrt{n}, \tau_{y'} > n-k, A_{n-k}) \mathbb{P}_x(X_k \in dx', \\ &\quad \underbrace{y + S_k \in dy', y + S_k \leq n^{1/2-\varepsilon/6}, \tau_y > k, v_n^{\varepsilon/6} = k)}_{=: \mathbb{P}_x(\tau_y > n)L_5}). \end{aligned} \quad (7.3)$$

Bound of L_3 . Using the bound of J_1 in (6.19) and Theorem 2.2,

$$L_3 \leq \frac{J_1}{\mathbb{P}_x(\tau_y > n)} \leq \frac{c_{p,\varepsilon}(1+y+|x|)(1+|x|)^{p-1}}{n^{1/2+\varepsilon}\mathbb{P}_x(\tau_y > n)} \xrightarrow{n \rightarrow +\infty} 0. \tag{7.4}$$

Bound of L_4 . Using the bound of J_3 in (6.22) and Theorem 2.2, we have

$$L_4 \leq \frac{J_3}{\mathbb{P}_x(\tau_y > n)} = o(1). \tag{7.5}$$

Upper bound of L_5 . Define $t_+ = t + \frac{2}{(n-k)^{2\varepsilon}}$ and $y'_+ = y' + (n-k)^{1/2-2\varepsilon}$. By Proposition A.3,

$$\begin{aligned} & \mathbb{P}_{x'}(y' + S_{n-k} \leq t\sqrt{n}, \tau_{y'} > n-k, A_{n-k}) \\ & \leq \mathbb{P}(y'_+ + \sigma B_{n-k} \leq t_+\sqrt{n}, \tau_{y'_+}^{\text{bm}} > n-k) \\ & = \frac{1}{\sqrt{2\pi(n-k)}\sigma} \int_0^{t_+\sqrt{n}} e^{-\frac{(s-y'_+)^2}{2(n-k)\sigma^2}} - e^{-\frac{(s+y'_+)^2}{2(n-k)\sigma^2}} \, ds. \end{aligned}$$

Note that for any $y' \leq n^{1/2-\varepsilon/6}$ we have $y'_+/\sqrt{n} \leq \frac{2}{n^{\varepsilon/6}}$ and for any $k \leq [n^{1-\varepsilon}]$ we have $n(1 - \frac{1}{n^\varepsilon}) \leq n-k \leq n$. Using these remarks with the fact that $\text{sh}(x) \leq x(1 + \frac{x^2}{6} \text{ch}(x))$ for any $x \geq 0$, we obtain after some calculations that

$$\begin{aligned} & \mathbb{P}_{x'}(y' + S_{n-k} \leq t\sqrt{n}, \tau_{y'} > n-k, A_{n-k}) \\ & \leq \frac{2y'_+}{\sqrt{2\pi n}\sigma} \left(1 + \frac{c_\varepsilon}{n^\varepsilon}\right) \int_0^{t_+\sqrt{n}} \frac{s e^{-\frac{s^2+(y'_+)^2}{2(n-k)\sigma^2}}}{(n-k)\sigma^2} \left[1 + \frac{s^2(y'_+)^2}{6(n-k)^2\sigma^4} \text{ch}\left(\frac{sy'_+}{(n-k)\sigma^2}\right)\right] \, ds \\ & \leq \frac{2y'_+}{\sqrt{2\pi n}\sigma} \left(1 + \frac{c_{t,\varepsilon}}{n^{\varepsilon/3}}\right) (1 - e^{-\frac{t^2}{2\sigma^2}}). \end{aligned}$$

Consequently, using the same arguments as in the proof of Theorem 2.2 in Section 6 (see the developments from (6.23) to (6.25)), we obtain, with $\Phi_\sigma^+(t) = 1 - e^{-\frac{t^2}{2\sigma^2}}$,

$$L_5 \leq \left(1 + \frac{c_{t,\varepsilon}}{n^{\varepsilon/3}}\right) \Phi_\sigma^+(t) \frac{2V(x, y)}{\sqrt{2\pi n}\sigma \mathbb{P}_x(\tau_y > n)} (1 + o(1)),$$

which by the claim 2 of Theorem 2.2 implies that

$$L_5 = \Phi_\sigma^+(t)(1 + o(1)). \tag{7.6}$$

Lower bound of L_5 . In the same way as for the upper bound, with $y'_- = y' - (n-k)^{1/2-2\varepsilon}$ and $t_- = t - \frac{2}{(n-k)^{2\varepsilon}}$, we have

$$\begin{aligned} & \mathbb{P}_x(\tau_y > n)L_5 \\ & \geq \sum_{k=1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R}_+^*} \mathbb{P}(y'_- + \sigma B_{n-k} \leq t_-\sqrt{n}, \tau_{y'_-}^{\text{bm}} > n-k) \mathbb{P}_x(y + S_k \in \, \text{d}y', \\ & \quad (n-k)^{1/2-2\varepsilon} < y + S_k \leq n^{1/2-\varepsilon/6}, \tau_y > k, \nu_n^{\varepsilon/6} = k) \\ & \quad - \sum_{k=1}^{[n^{1-\varepsilon}]} \int_{\mathbb{R}} \mathbb{P}_{x'}(\bar{A}_{n-k}) \mathbb{P}_x(X_k \in \, \text{d}x', \tau_y > k, \nu_n^{\varepsilon/6} = k). \end{aligned}$$

Using Lemma A.3 with y'_- , which is positive when $(n-k)^{1/2-2\varepsilon} < y' \leq n^{1/2-\varepsilon/6}$, we obtain after calculation that

$$\mathbb{P}(y'_- + \sigma B_{n-k} \leq t_- \sqrt{n}, \tau_{y'_-}^{\text{bm}} > n-k) \geq \frac{2y'_-}{\sqrt{2\pi n\sigma}} \left(1 - \frac{c_{t,\varepsilon}}{n^{\varepsilon/3}}\right) \Phi_\sigma^+(t).$$

Copying the proof of the bound of J_1 in (6.19) and using the same arguments as in the proof of Theorem 2.2 in Section 6 (see the developments from (6.26) to (6.27)), we get

$$L_5 \geq \Phi_\sigma^+(t) \frac{2V(x, y)}{\sqrt{2\pi n\sigma} \mathbb{P}_x(\tau_y > n)} (1 - o(1)) = \Phi_\sigma^+(t) (1 - o(1)).$$

Coupling this with (7.6) we obtain that

$$L_5 = \Phi_\sigma^+(t) (1 + o(1)).$$

Inserting this and (7.4) and (7.5) into (7.3), we deduce that $L_2 \underset{n \rightarrow +\infty}{\sim} \Phi_\sigma^+(t)$. By (7.1) and (7.2), we finally have

$$\mathbb{P}_x(y + S_n \leq t \sqrt{n} \mid \tau_y > n) \underset{n \rightarrow +\infty}{\longrightarrow} \Phi_\sigma^+(t).$$

Changing t into $t\sigma$, this concludes the proof.

8. The case of non-positive initial point

In this section, we prove Theorem 2.5.

Lemma 8.1. *Assume Condition 1. For any $(x, y) \in \mathcal{D}^-$, the random variable M_{τ_y} is integrable and the function $V(x, y) = -\mathbb{E}_x(M_{\tau_y})$, is well defined on \mathcal{D}^- .*

Proof. If $\mathbb{E}(a) \geq 0$, by the Markov inequality, with $z = y + \rho x$,

$$\begin{aligned} \mathbb{E}_x(z + M_n; \tau_y > n) &= \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{E}_{x'}(y' + \rho x' + M_{n-1}; \tau_{y'} > n-1) \\ &\quad \times \mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1). \end{aligned}$$

Since $y + S_1 > 0$ on $\{\tau_y > 1\}$, by Lemma 4.4,

$$\begin{aligned} \mathbb{E}_x(z + M_n; \tau_y > n) &\leq c_p \mathbb{E}_x((1 + y + S_1 + |X_1|)(1 + |X_1|)^{p-1}; \tau_y > 1) \\ &\leq c_p \mathbb{E}_x((1 + |X_1|)^p) \\ &\leq c_p (1 + |x|)^p. \end{aligned} \tag{8.1}$$

Moreover

$$\begin{aligned} \mathbb{E}_x(|M_{\tau_y}|; \tau_y \leq n) &\leq |z| + \sum_{k=2}^n \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{E}_{x'}(|y' + \rho x' + M_{k-1}|; \tau_{y'} = k-1) \\ &\quad \times \mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1) \\ &\quad + \mathbb{E}_x(|M_1|; \tau_y = 1). \end{aligned}$$

Since $y + S_1 > 0$ on $\{\tau_y > 1\}$, by Lemma 4.2,

$$\begin{aligned} \mathbb{E}_x(|M_{\tau_y}|; \tau_y \leq n) &\leq c(1 + |y| + |x|) - \mathbb{E}_x(z + M_{\tau_y}; \tau_y \leq n) \\ &\leq c(1 + |y| + |x|) + \mathbb{E}_x(z + M_n; \tau_y > n). \end{aligned}$$

Using (8.1), we deduce that $\mathbb{E}_x(|M_{\tau_y}|; \tau_y \leq n) \leq c_p(1 + |y| + |x|^p)$. Consequently, by the Lebesgue monotone convergence theorem, the assertion is proved when $\mathbb{E}(a) \geq 0$. When $\mathbb{E}(a) < 0$, the assertion follows from Lemma 5.4. \square

Lemma 8.2. *Assume Condition 1. The function V is \mathbf{Q}_+ -harmonic on $\mathcal{D} = \mathcal{D}^- \cup \mathbb{R} \times \mathbb{R}_+^*$. If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then the function V is positive on $\mathcal{D} = \mathcal{D}^- \cup \mathbb{R} \times \mathbb{R}_+^*$.*

Proof. Note that by Corollary A.7, we have $\mathbb{P}_x(\tau_y < +\infty) = 1$, for any $x \in \mathbb{R}$ and $y \in \mathbb{R}$. Therefore, by the Lebesgue dominated convergence theorem,

$$V(x, y) = -\mathbb{E}_x(M_{\tau_y}) = z - \lim_{n \rightarrow \infty} \mathbb{E}_x(z + M_{\tau_y}; \tau_y \leq n) = \lim_{n \rightarrow \infty} \mathbb{E}_x(z + M_n; \tau_y > n),$$

for any $(x, y) \in \mathcal{D}^-$. The fact that V is \mathbf{Q}_+ -harmonic on \mathcal{D} can be proved in the same way as in the proof of Proposition 5.3. Therefore, for any $(x, y) \in \mathcal{D}^-$,

$$V(x, y) = \mathbb{E}_x(V(X_1, y + S_1); \tau_y > 1). \tag{8.2}$$

By the claim 2 of Proposition 5.3 and the claim 3 of Proposition 5.9, on $\{\tau_y > 1\}$, the random variable $V(X_1, y + S_1)$ is positive almost surely. Since by the definition of \mathcal{D}^- , we have $\mathbb{P}_x(\tau_y > 1) > 0$, we conclude that $V(x, y) > 0$ for any $(x, y) \in \mathcal{D}^-$. \square

Lemma 8.3. *Assume Condition 1.*

(1) *For any $(x, y) \in \mathcal{D}^-$,*

$$\sqrt{n}\mathbb{P}_x(\tau_y > n) \leq c_p(1 + |x|)^p.$$

(2) *If in addition we assume either Condition 2 and $\mathbb{E}(a) \geq 0$, or Condition 3, then for any $(x, y) \in \mathcal{D}^-$,*

$$\mathbb{P}_x(\tau_y > n) \underset{n \rightarrow +\infty}{\sim} \frac{2V(x, y)}{\sqrt{2\pi n\sigma}}.$$

Proof. By the Markov property,

$$\sqrt{n}\mathbb{P}_x(\tau_y > n) = \int_{\mathbb{R} \times \mathbb{R}_+^*} \sqrt{n}\mathbb{P}_{x'}(\tau_{y'} > n - 1)\mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1).$$

By Theorem 2.2, for any $y' > 0$, we have $\sqrt{n}\mathbb{P}_{x'}(\tau_{y'} > n - 1) \leq c_p(1 + y' + |x'|)^p$ and moreover, for any $y \leq 0$,

$$\mathbb{E}_x(c_p(1 + y + S_1 + |X_1|)^p; \tau_y > 1) \leq c_p(1 + |x|)^p.$$

Then, we obtain the claim 1 and by the Lebesgue dominated convergence theorem and the claim 2 of Theorem 2.2,

$$\lim_{n \rightarrow \infty} \sqrt{n}\mathbb{P}_x(\tau_y > n) = \mathbb{E}_x\left(\frac{2V(X_1, y + S_1)}{\sqrt{2\pi\sigma}}; \tau_y > 1\right).$$

Using (8.2) we conclude the proof. \square

Lemma 8.4. *Assume either Conditions 1, 2 and $\mathbb{E}(a) \geq 0$, or Conditions 1 and 3. For any $(x, y) \in \mathcal{D}^-$ and $t > 0$,*

$$\mathbb{P}_x\left(\frac{y + S_n}{\sigma\sqrt{n}} \leq t \mid \tau_y > n\right) \underset{n \rightarrow +\infty}{\longrightarrow} 1 - e^{-\frac{t^2}{2}}.$$

Proof. Similarly as in the proof of Lemma 8.3, we write,

$$\begin{aligned}
 & \mathbb{P}_x \left(\frac{y + S_n}{\sigma \sqrt{n}} \leq t \mid \tau_y > n \right) \\
 &= \frac{1}{\mathbb{P}_x(\tau_y > n)} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'} \left(\frac{y' + S_{n-1}}{\sigma \sqrt{n-1}} \leq t; \tau_{y'} > n-1 \right) \\
 & \quad \times \mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1) \\
 &= \frac{1}{\sqrt{n} \mathbb{P}_x(\tau_y > n)} \int_{\mathbb{R} \times \mathbb{R}_+^*} \mathbb{P}_{x'} \left(\frac{y' + S_{n-1}}{\sigma \sqrt{n-1}} \leq t \mid \tau_{y'} > n-1 \right) \sqrt{n} \mathbb{P}_{x'}(\tau_{y'} > n-1) \\
 & \quad \times \mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1).
 \end{aligned}$$

Since, by Lemma 8.3, $\sqrt{n} \mathbb{P}_{x'}(\tau_{y'} > n-1) \leq c_p(1 + |x'|)^p$, applying the Lebesgue dominated convergence theorem, Theorem 2.2, Theorem 2.4 and Lemma 8.3, we have

$$\begin{aligned}
 & \lim_{n \rightarrow \infty} \mathbb{P}_x \left(\frac{y + S_n}{\sigma \sqrt{n}} \leq t \mid \tau_y > n \right) \\
 &= \frac{\sqrt{2\pi}\sigma}{2V(x, y)} \int_{\mathbb{R} \times \mathbb{R}_+^*} (1 - e^{-\frac{t^2}{2}}) \frac{2V(x', y')}{\sqrt{2\pi}\sigma} \mathbb{P}_x(X_1 \in dx', y + S_1 \in dy', \tau_y > 1).
 \end{aligned}$$

Using (8.2) concludes the proof. \square

Appendix

A.1. Proof of the fact Condition 3bis implies Condition 3

We suppose that Condition 3bis holds. Then, there exists $\delta > 0$ such that

$$\mathbb{P}((a, b) \in [-1 + \delta, 0] \times [\delta, C]) > 0 \tag{A.1}$$

and

$$\mathbb{P}((a, b) \in [0, 1 - \delta] \times [\delta, C]) > 0. \tag{A.2}$$

For any $x \in \mathbb{R}$, set $C_x = \max(|x|, \frac{C}{\delta})$ and

$$\mathcal{A}_n = \{\delta \leq X_1 \leq C_x, \delta \leq X_2 \leq C_{X_1}, \dots, \delta \leq X_n \leq C_{X_{n-1}}\}.$$

Using (A.1) for $x < 0$ and (A.2) for $x \geq 0$, we obtain that $\mathbb{P}_x(\mathcal{A}_1) > 0$. By the Markov property, we deduce that $\mathbb{P}_x(\mathcal{A}_n) > 0$. Moreover, it is easy to see that, on \mathcal{A}_n , we have $y + S_k \geq y + k\delta > 0$, for all $k \leq n$, and $|X_n| \leq C_x$. Taking $n = n_0$ large enough, we conclude that Condition 3 holds under Condition 3bis.

A.2. Convergence of recursively bounded monotonic sequences

The following lemmas give sufficient conditions for a monotonic sequence to be bounded.

Lemma A.1. *Let $(u_n)_{n \geq 1}$ be a non-decreasing sequence of reals such that there exist $n_0 \geq 2$, $\varepsilon \in (0, 1)$, $\alpha, \beta, \gamma \geq 0$ and $\delta > 0$ such that for any $n \geq n_0$,*

$$u_n \leq \left(1 + \frac{\alpha}{n^\varepsilon}\right) u_{\lfloor n^{1-\varepsilon} \rfloor} + \frac{\beta}{n^\varepsilon} + \gamma e^{-\delta n^\varepsilon}. \tag{A.3}$$

Then, for any $n \geq n_0$ and any integer $k_0 \in \{n_0, \dots, n\}$,

$$\begin{aligned} u_n &\leq \exp\left(\frac{\alpha}{k_0^\varepsilon} \frac{2^\varepsilon 2^{\varepsilon^2}}{2^{\varepsilon^2} - 1}\right) \left(u_{k_0} + \frac{\beta}{k_0^\varepsilon} \frac{2^\varepsilon 2^{\varepsilon^2}}{2^{\varepsilon^2} - 1} + \gamma \frac{\exp(-\delta \frac{k_0^\varepsilon}{2^\varepsilon})}{1 - e^{-\delta(2^{\varepsilon^2} - 1)}}\right) \\ &\leq \left(1 + \frac{C_{\alpha, \varepsilon}}{k_0^\varepsilon}\right) u_{k_0} + \beta \frac{C_{\alpha, \varepsilon}}{k_0^\varepsilon} + \gamma C_{\alpha, \delta, \varepsilon} e^{-c_{\alpha, \delta, \varepsilon} k_0^\varepsilon}. \end{aligned}$$

In particular, choosing k_0 constant, it follows that $(u_n)_{n \geq 1}$ is bounded.

Proof. Fix $n \geq n_0$ and $k_0 \in \{n_0, \dots, n\}$ and consider for all $j \geq 0$,

$$p_j = \lceil n^{(1-\varepsilon)^j} \rceil.$$

The sequence $(p_j)_{j \geq 0}$ starts at $n_0 = n$, is non-increasing and converge to 1. So there exists $m = m(k_0) \in \mathbb{N}$ such that $p_m \geq k_0 \geq p_{m+1}$. Since $n^{(1-\varepsilon)^j} / 2 \geq k_0 / 2 \geq 1$, for all $j \in \{0, \dots, m\}$, we have

$$n^{(1-\varepsilon)^j} \geq p_j \geq n^{(1-\varepsilon)^j} - 1 \geq \frac{n^{(1-\varepsilon)^j}}{2}. \tag{A.4}$$

Using (A.3) and the fact that $(u_n)_{n \geq 2}$ is non-decreasing, we write for all $j = 0, \dots, m$,

$$u_{p_j} \leq \left(1 + \frac{\alpha}{p_j^\varepsilon}\right) u_{p_{j+1}} + \frac{\beta}{p_j^\varepsilon} + \gamma e^{-\delta p_j^\varepsilon} \leq \left(1 + \frac{\alpha}{p_j^\varepsilon}\right) \left(u_{p_{j+1}} + \frac{\beta}{p_j^\varepsilon} + \gamma e^{-\delta p_j^\varepsilon}\right).$$

Iterating, we obtain that

$$u_n \leq A_m(u_{p_{m+1}} + \beta B_m + \gamma C_m),$$

where $A_m = \prod_{j=0}^m (1 + \frac{\alpha}{p_j^\varepsilon})$, $B_m = \sum_{j=0}^m \frac{1}{p_j^\varepsilon}$ and $C_m = \sum_{j=0}^m e^{-\delta p_j^\varepsilon}$. Since $p_{m+1} \leq k_0$ and since $(u_n)_{n \geq 2}$ is non-decreasing,

$$u_n \leq A_m(u_{k_0} + \beta B_m + \gamma C_m). \tag{A.5}$$

Now, we bound A_m as follows,

$$A_m \leq \prod_{j=0}^m e^{\frac{\alpha}{p_j^\varepsilon}} = e^{\alpha B_m}. \tag{A.6}$$

Denoting $\eta_j = n^{-(1-\varepsilon)^j \varepsilon}$, using (A.4), we have $B_m \leq 2^\varepsilon \sum_{j=0}^m \eta_j$. Moreover, for all $j \leq m$, we note that $\frac{\eta_j}{\eta_{j+1}} = \frac{1}{n^{\varepsilon^2(1-\varepsilon)^j}} \leq \frac{1}{k_0^{\varepsilon^2}} \leq \frac{1}{2^{\varepsilon^2}} < 1$ and so

$$\eta_j \leq \frac{\eta_m}{2^{\varepsilon^2(m-j)}} \leq \frac{1}{p_m^\varepsilon 2^{\varepsilon^2(m-j)}} \leq \frac{1}{k_0^\varepsilon 2^{\varepsilon^2(m-j)}}. \tag{A.7}$$

Therefore, B_m is bounded as follows:

$$B_m \leq \frac{2^\varepsilon}{k_0^\varepsilon} \sum_{k=0}^m \left(\frac{1}{2^{\varepsilon^2}}\right)^k \leq \frac{1}{k_0^\varepsilon} \frac{2^\varepsilon 2^{\varepsilon^2}}{2^{\varepsilon^2} - 1}. \tag{A.8}$$

Using (A.4) and (A.7), we have

$$C_m \leq \sum_{j=0}^m e^{-\frac{\delta}{2^{\varepsilon} \eta_j}} \leq \sum_{j=0}^m \exp\left(-\frac{\delta k_0^{\varepsilon} 2^{\varepsilon^2(m-j)}}{2^{\varepsilon}}\right).$$

Since for any $u \geq 0$ and $k \in \mathbb{N}$, we have $(1 + u)^k \geq 1 + ku$, it follows that

$$C_m \leq e^{-\frac{\delta k_0^{\varepsilon}}{2^{\varepsilon}}} \sum_{k=0}^m \exp(-\delta k(2^{\varepsilon^2} - 1)) \leq \frac{e^{-\frac{\delta k_0^{\varepsilon}}{2^{\varepsilon}}}}{1 - e^{-\delta(2^{\varepsilon^2} - 1)}}. \tag{A.9}$$

Putting together (A.6), (A.8) and (A.9) into (A.5) proves the lemma. □

Lemma A.2. *Let $(u_n)_{n \geq 1}$ be a non-increasing sequence of reals such that there exist $n_0 \geq 2$, $\varepsilon \in (0, 1)$ and $\beta \geq 0$ such that for any $n \geq n_0$,*

$$u_n \geq u_{\lfloor n^{1-\varepsilon} \rfloor} - \frac{\beta}{n^{\varepsilon}}.$$

Then, for any $n \geq n_0$ and any integer $k_0 \in \{n_0, \dots, n\}$,

$$u_n \geq u_{k_0} - \frac{\beta}{k_0^{\varepsilon}} \frac{2^{\varepsilon} 2^{\varepsilon^2}}{2^{\varepsilon^2} - 1} = u_{k_0} - c_{\varepsilon} \frac{\beta}{k_0^{\varepsilon}}.$$

In particular, choosing k_0 constant, it follows that $(u_n)_{n \geq 1}$ is bounded.

Proof. For the proof it is enough to use Lemma A.1 with u_n replaced by $-u_n$. □

A.3. Results on the Brownian case and strong approximation

Consider the standard Brownian motion $(B_t)_{t \geq 0}$ living on the probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Define the exit time

$$\tau_y^{\text{bm}} = \inf\{t \geq 0, y + \sigma B_t \leq 0\}, \tag{A.10}$$

where $\sigma > 0$. The following assertions are due to Lévy [19].

Proposition A.3. *For any $y > 0$, $0 \leq a \leq b$ and $n \geq 1$,*

$$\mathbb{P}(\tau_y^{\text{bm}} > n) = \frac{2}{\sqrt{2\pi n \sigma}} \int_0^y e^{-\frac{s^2}{2n\sigma^2}} ds$$

and

$$\mathbb{P}(\tau_y^{\text{bm}} > n, y + \sigma B_n \in [a, b]) = \frac{1}{\sqrt{2\pi n \sigma}} \int_a^b \left(e^{-\frac{(s-y)^2}{2n\sigma^2}} - e^{-\frac{(s+y)^2}{2n\sigma^2}} \right) ds.$$

From this one can deduce easily:

Corollary A.4.

(1) *For any $y > 0$,*

$$\mathbb{P}(\tau_y^{\text{bm}} > n) \leq c \frac{y}{\sqrt{n}}.$$

(2) For any sequence of real numbers $(\theta_n)_{n \geq 0}$ such that $\theta_n \xrightarrow{n \rightarrow +\infty} 0$,

$$\sup_{y \in [0; \theta_n \sqrt{n}]} \left(\frac{\mathbb{P}(\tau_y^{\text{bm}} > n)}{\frac{2y}{\sqrt{2\pi n \sigma}}} - 1 \right) = O(\theta_n^2).$$

To transfer the results from the Brownian motion to the Markov walk, we use a functional approximation given in Theorem 3.3 from Grama, Le Page and Peigné [15]. We have to construct an adapted Banach space \mathcal{B} and verify the hypotheses **M1–M5** in [15] which are necessary to apply Theorem 3.3. Fix $p \in (2, \alpha)$ and let ε, θ, c_0 and δ be positive numbers such that $c_0 + \varepsilon < \theta < 2c_0 < \alpha - \varepsilon$ and $2 < 2 + 2\delta < (2 + 2\delta)\theta \leq p$. Define the Banach space $\mathcal{B} = \mathcal{L}_{\varepsilon, c_0, \theta}$ as the set of continuous function f from \mathbb{R} to \mathbb{C} such that $\|f\| = |f|_\theta + [f]_{\varepsilon, c_0} < +\infty$, where

$$|f|_\theta = \sup_{x \in \mathbb{R}} \frac{|f(x)|}{1 + |x|^\theta}, \quad [f]_{\varepsilon, c_0} = \sup_{\substack{(x, y) \in \mathbb{R}^2 \\ x \neq y}} \frac{|f(x) - f(y)|}{|x - y|^\varepsilon (1 + |x|^{c_0})(1 + |y|^{c_0})}.$$

For example, one can take $\varepsilon < \min(\frac{p-2}{4}, \frac{1}{2})$, $c_0 = 1$, $\theta = 1 + 2\varepsilon$ and $2 + 2\delta = \frac{p}{1+2\varepsilon}$. Using the techniques from [17] one can verify that, under Condition 1, the Banach space \mathcal{B} and the perturbed operator $\mathbf{P}_t f(x) = \mathbb{E}_x(f(X_1) e^{itX_1})$ satisfy Hypotheses **M1–M5** in [15]. The hypothesis **M1** is verified straightforwardly. In particular the norm of the Dirac measure δ_x is bounded: $\|\delta_x\|_{\mathcal{B} \rightarrow \mathcal{B}} \leq 1 + |x|^\theta$, for each $x \in \mathbb{R}$. We refer to Proposition 4 and Corollary 3 of [17] for **M2–M3**. For **M4**, we have

$$\mu_\delta(x) = \sup_{k \geq 1} \mathbb{E}_x^{1/2+2\delta}(|X_n|^{2+2\delta}) \leq c_\delta(1 + |x|).$$

Hypothesis **M5** follows from Proposition 1 of [17] and Lemma 3.1.

With these considerations, the $C(x) = C_1(1 + \mu_\delta(x) + \|\delta_x\|)^{2+2\delta}$ in Theorem 3.3 established in [15] is less than $c_p(1 + |x|)^p$, where C_1 is a constant. Therefore Theorem 3.3 can be reformulated in the case of the stochastic recursion as follows.

Proposition A.5. Assume Condition 1. For any $p \in (2, \alpha)$, there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $n \geq 1$, without loss of generality (on an extension of the initial probability space) one can reconstruct the sequence $(S_n)_{n \geq 0}$ with a continuous time Brownian motion $(B_t)_{t \in \mathbb{R}_+}$, such that

$$\mathbb{P}_x \left(\sup_{0 \leq t \leq 1} |S_{[tn]} - \sigma B_{tn}| > n^{1/2-\varepsilon} \right) \leq \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p,$$

where σ is given by (2.2).

This proposition plays the crucial role in the proof of Theorem 2.2 and Theorem 2.4 (cf. Sections 6 and 7). The following straightforward consequence of Proposition A.5 is used in the proof of Lemma 4.1 in Section 4. Set $\Phi(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t e^{-\frac{u^2}{2}} du$.

Corollary A.6. Assume Condition 1. For any $p \in (2, \alpha)$, there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $n \geq 1$,

$$\sup_{u \in \mathbb{R}} \left| \mathbb{P}_x \left(\frac{S_n}{\sqrt{n}} \leq u \right) - \Phi \left(\frac{u}{\sigma} \right) \right| \leq \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p.$$

Proof. Let $\varepsilon \in (0, 1/2)$ and $A_n = \{\sup_{0 \leq t \leq 1} |S_{[tn]} - \sigma B_{tn}| > n^{1/2-\varepsilon}\}$. For any $x \in \mathbb{R}$ and any $u \in \mathbb{R}$,

$$\mathbb{P}_x \left(\frac{S_n}{\sqrt{n}} \leq u \right) \leq \mathbb{P}_x(A_n) + \mathbb{P}_x \left(\frac{\sigma B_n}{\sqrt{n}} \leq u + \frac{1}{n^\varepsilon} \right),$$

where the last probability does not exceed $\Phi\left(\frac{u}{\sigma}\right) + c_\varepsilon n^{-\varepsilon}$. Using Proposition A.5, we conclude that there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$ and $x \in \mathbb{R}$,

$$\mathbb{P}_x\left(\frac{S_n}{\sqrt{n}} \leq u\right) \leq \Phi\left(\frac{u}{\sigma}\right) + \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p.$$

In the same way we obtain a lower bound and the assertion follows. □

A.4. Finiteness of the exit times

Corollary A.7. *Assume Condition 1. For any $x \in \mathbb{R}$ and $y \in \mathbb{R}$,*

$$\mathbb{P}_x(\tau_y < +\infty) = 1 \quad \text{and} \quad \mathbb{P}_x(T_y < +\infty) = 1.$$

Proof. Let $y > 0$ and $\varepsilon \in (0, 1/2)$. Set $A_n = \{\sup_{0 \leq t \leq 1} |S_{[tn]} - \sigma B_{tn}| \leq n^{1/2-\varepsilon}\}$. Using Proposition A.5, there exists $\varepsilon_0 > 0$ such that for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $y > 0$,

$$\begin{aligned} \mathbb{P}_x(\tau_y > n) &\leq \mathbb{P}_x(\tau_y > n, A_n) + \mathbb{P}_x(\bar{A}_n) \\ &\leq \mathbb{P}(\tau_{y+n^{1/2-\varepsilon}}^{\text{bm}} > n) + \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p. \end{aligned}$$

Since, by the claim 1 of Corollary A.4, $\mathbb{P}(\tau_{y+n^{1/2-\varepsilon}}^{\text{bm}} > n) \leq c \frac{y+n^{1/2-\varepsilon}}{\sqrt{n}} \leq (1+y) \frac{c}{n^\varepsilon}$, taking the limit as $n \rightarrow +\infty$ we conclude that $\mathbb{P}_x(\tau_y < +\infty) = 1$.

Let $D_n = \{\max_{1 \leq k \leq n} |S_k - M_k| \leq n^{1/2-\varepsilon}\}$. Obviously

$$\begin{aligned} \mathbb{P}_x(T_y > n) &\leq \mathbb{P}_x(T_y > n, A_n, D_n) + \mathbb{P}_x(\bar{A}_n) + \mathbb{P}_x(\bar{D}_n) \\ &\leq \mathbb{P}(\tau_{y+2n^{1/2-\varepsilon}}^{\text{bm}} > n) + \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p + \mathbb{P}_x\left(\max_{1 \leq k \leq n} |\rho X_k| > n^{1/2-\varepsilon}\right). \end{aligned}$$

Using the claim 1 of Corollary A.4, the Markov inequality and Lemma 3.1, for any $\varepsilon \in (0, \varepsilon_0]$, $x \in \mathbb{R}$ and $y > 0$,

$$\mathbb{P}_x(T_y > n) \leq (1+y) \frac{c}{n^\varepsilon} + \frac{c_{p,\varepsilon}}{n^\varepsilon} (1 + |x|)^p + c_p \frac{1 + |x|^p}{n^{\frac{p-2}{2}-p\varepsilon}}.$$

Choosing ε small enough and taking the limit as $n \rightarrow +\infty$ we conclude the second assertion when $y > 0$.

When $y \leq 0$, the results follow since the applications $y \mapsto \tau_y$ and $y \mapsto T_y$ are non-decreasing. □

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