

PASSAGE OF LÉVY PROCESSES ACROSS POWER LAW BOUNDARIES AT SMALL TIMES¹

BY J. BERTOIN, R. A. DONEY AND R. A. MALLER

*Université Pierre et Marie Curie, University of Manchester
 and Australian National University*

We wish to characterize when a Lévy process X_t crosses boundaries like t^κ , $\kappa > 0$, in a one- or two-sided sense, for small times t ; thus, we inquire when $\limsup_{t \downarrow 0} |X_t|/t^\kappa$, $\limsup_{t \downarrow 0} X_t/t^\kappa$ and/or $\liminf_{t \downarrow 0} X_t/t^\kappa$ are almost surely (a.s.) finite or infinite. Necessary and sufficient conditions are given for these possibilities for all values of $\kappa > 0$. This completes and extends a line of research going back to Blumenthal and Gettoor in the 1960s. Often (for many values of κ), when the lim sups are finite a.s., they are in fact zero, but the lim sups may in some circumstances take finite, nonzero, values, a.s. In general, the process crosses one- or two-sided boundaries in quite different ways, but surprisingly this is not so for the case $\kappa = 1/2$, where a new kind of analogue of an iterated logarithm law with a square root boundary is derived. An integral test is given to distinguish the possibilities in that case.

1. Introduction. Let $X = (X_t, t \geq 0)$ be a Lévy process with characteristic triplet (γ, σ, Π) , where $\gamma \in \mathbb{R}$, $\sigma^2 \geq 0$, and the Lévy measure Π has $\int (x^2 \wedge 1)\Pi(dx)$ finite. See [1] and [15] for basic definitions and properties. Since we will only be concerned with local behavior of X_t , as $t \downarrow 0$, we can ignore the “big jumps” in X (those with modulus exceeding 1, say), and write its characteristic exponent, $\Psi(\theta) = \frac{1}{t} \log E e^{i\theta X_t}$, $\theta \in \mathbb{R}$, as

$$(1.1) \quad \Psi(\theta) = i\gamma\theta - \sigma^2\theta^2/2 + \int_{[-1,1]} (e^{i\theta x} - 1 - i\theta x)\Pi(dx).$$

The Lévy process is of bounded variation, for which we use the notation $X \in bv$, if and only if $\sigma^2 = 0$ and $\int_{|x| \leq 1} |x|\Pi(dx) < \infty$, and in that case, we denote by

$$\delta := \gamma - \int_{[-1,1]} x\Pi(dx)$$

its drift coefficient.

We will continue the work of Blumenthal and Gettoor [3] and Pruitt [13], in a sense, by investigating the possible limiting values taken by $t^{-\kappa} X_t$ as $t \downarrow 0$, where

Received March 2006; revised October 2006.

¹Supported in part by ARC Grants DP0210572 and DP0664603.

AMS 2000 subject classifications. Primary 60J30, 60F15; secondary 60G10, 60G17, 60J65.

Key words and phrases. Lévy processes, crossing power law boundaries, limiting and lim sup behavior.

$\kappa > 0$ is some parameter. Recall that Blumenthal and Gettoor [3] introduced the upper-index

$$\beta := \inf \left\{ \alpha > 0 : \int_{|x| \leq 1} |x|^\alpha \Pi(dx) < \infty \right\} \in [0, 2],$$

which plays a critical role in this framework. Indeed, assuming for simplicity that the Brownian coefficient σ^2 is zero, and further that the drift coefficient δ is also 0 when $X \in bv$, then with probability 1,

$$(1.2) \quad \limsup_{t \downarrow 0} \frac{|X_t|}{t^\kappa} = \begin{cases} 0, \\ \infty, \end{cases} \quad \text{according as } \begin{cases} \kappa < 1/\beta, \\ \kappa > 1/\beta. \end{cases}$$

See Sato ([15], Proposition 47.24, page 362) and also Pruitt [13]. Note that the critical case when $\kappa = 1/\beta$ is not covered by (1.2).

One application of this kind of study is to get information on the rate of growth of the process relative to power law functions, in both a one- and a two-sided sense, at small times. More precisely, we are concerned with the values of $\limsup_{t \downarrow 0} |X_t|/t^\kappa$ and of $\limsup_{t \downarrow 0} X_t/t^\kappa$ (and the behavior of $\liminf_{t \downarrow 0} X_t/t^\kappa$ can be deduced from the \limsup behavior by performing a sign reversal). For example, when

$$(1.3) \quad \limsup_{t \downarrow 0} \frac{|X_t|}{t^\kappa} = \infty \quad \text{a.s.},$$

then the regions $\{(t, y) \in [0, \infty) \times \mathbb{R} : |y| > at^\kappa\}$ are entered infinitely often for arbitrarily small t , a.s. (almost surely), for all $a > 0$. This can be thought of as a kind of “regularity” of X for these regions, at 0. We will refer to this kind of behavior as crossing a “two-sided” boundary. On the other hand, when

$$(1.4) \quad \limsup_{t \downarrow 0} \frac{X_t}{t^\kappa} = \infty \quad \text{a.s.},$$

we have “one-sided” (up)crossings; and similarly for downcrossings, phrased in terms of the \liminf . In general, the process crosses one- or two-sided boundaries in quite different ways, and, often (for many values of κ), when the \limsup s are finite a.s., they are in fact zero, a.s., as we will show. But the \limsup s may in some circumstances take finite nonzero values, a.s. Our aim here is to give necessary and sufficient conditions (NASC’s) which distinguish all these possibilities, for all values of $\kappa > 0$.

Let us eliminate at the outset certain cases which are trivial or easily deduced from known results. A result of Khintchine [10] (see Sato [15], Proposition 47.11, page 358) is that, for any Lévy process X with Brownian coefficient $\sigma^2 \geq 0$, we have

$$(1.5) \quad \limsup_{t \downarrow 0} \frac{|X_t|}{\sqrt{2t \log |\log t|}} = \sigma \quad \text{a.s.}$$

Thus we immediately see that (1.3) and (1.4) cannot hold for $0 < \kappa < 1/2$; we always have $\lim_{t \downarrow 0} X_t/t^\kappa = 0$ a.s. in these cases. Of course, this also agrees with (1.2), since, always, $\beta \leq 2$. More precisely, recall the decomposition $X_t = \sigma B_t + X_t^{(0)}$, where $X^{(0)}$ is a Lévy process with characteristics $(\gamma, 0, \Pi)$ and B is an independent Brownian motion. Khintchine's law of the iterated logarithm for B and (1.5) applied for $X^{(0)}$ give

$$(1.6) \quad -\liminf_{t \downarrow 0} \frac{X_t}{\sqrt{2t \log |\log t|}} = \limsup_{t \downarrow 0} \frac{X_t}{\sqrt{2t \log |\log t|}} = \sigma \quad \text{a.s.}$$

So the one- and two-sided lim sup behaviors of X are precisely determined when $\sigma^2 > 0$ [regardless of the behavior of $\Pi(\cdot)$]. With these considerations, it is clear that *throughout, we can assume*

$$(1.7) \quad \sigma^2 = 0.$$

Furthermore, we can restrict attention to the cases $\kappa \geq 1/2$.

A result of Shtatland [16] and Rogozin [14] is that $X \notin bv$ if and only if

$$-\liminf_{t \downarrow 0} \frac{X_t}{t} = \limsup_{t \downarrow 0} \frac{X_t}{t} = \infty \quad \text{a.s.,}$$

so (1.3) and (1.4) hold for all $\kappa \geq 1$, in this case (and similarly for the lim inf). On the other hand, when $X \in bv$, we have

$$\lim_{t \downarrow 0} \frac{X_t}{t} = \delta \quad \text{a.s.,}$$

where δ is the drift of X (cf. [1], page 84, or [15], page 65). Thus if $\delta > 0$, (1.4) holds for all $\kappa > 1$, but for no $\kappa \leq 1$, while if $\delta < 0$, (1.4) can hold for no $\kappa > 0$; while (1.3) holds in either case, with $\kappa > 1$, but for no $\kappa \leq 1$. Thus, when $X \in bv$, we need only consider the case $\delta = 0$.

Finally, we can also rule out the compound Poisson case, because then $X_t = 0$ for all $t \in (0, t_0)$ for some (random) $t_0 > 0$, and neither (1.3) nor (1.4) can hold.

The main statements for two-sided (resp., one-sided) boundary crossing will be given in Section 2 (resp., Section 3) and proved in Section 4 (resp., Section 5). We use throughout similar notation to [5, 6] and [7]. In particular, we write $\Pi^\#$ for the Lévy measure of $-X$, then $\Pi^{(+)}$ for the restriction of Π to $[0, \infty)$, $\Pi^{(-)}$ for the restriction of $\Pi^\#$ to $[0, \infty)$, and

$$(1.8) \quad \begin{aligned} \bar{\Pi}^{(+)}(x) &= \Pi((x, \infty)), \\ \bar{\Pi}^{(-)}(x) &= \Pi((-\infty, -x)), \\ \bar{\Pi}(x) &= \bar{\Pi}^{(+)}(x) + \bar{\Pi}^{(-)}(x), \quad x > 0, \end{aligned}$$

for the tails of $\Pi(\cdot)$. Recall that we assume (1.7) and that the Lévy measure Π is restricted to $[-1, 1]$. We will often make use of the Lévy–Itô decomposition,

which can be written as

$$(1.9) \quad X_t = \gamma t + \int_{[0,t] \times [-1,1]} x N(ds, dx), \quad t \geq 0,$$

where $N(dt, dx)$ is a Poisson random measure on $\mathbb{R}_+ \times [-1, 1]$ with intensity $dt \Pi(dx)$ and the Poissonian stochastic integral above is taken in the compensated sense. See Theorem 41 on page 31 in [12] for details.

2. Two-sided case. In this section we study two-sided crossings of power law boundaries at small times. We wish to find a necessary and sufficient condition for (1.3) for each value of $\kappa \geq 1/2$. This question is completely answered in the next two theorems, where the first can be viewed as a reinforcement of (1.2).

THEOREM 2.1. *Assume (1.7), and take $\kappa > 1/2$. When $X \in bv$, assume its drift is zero.*

(i) *If*

$$(2.1) \quad \int_0^1 \overline{\Pi}(x^\kappa) dx < \infty,$$

then we have

$$(2.2) \quad \lim_{t \downarrow 0} \frac{X_t}{t^\kappa} = 0 \quad a.s.$$

(ii) *Conversely, if (2.1) fails, then*

$$\limsup_{t \downarrow 0} \frac{|X_t - a(t)|}{t^\kappa} = \infty \quad a.s.,$$

for any nonstochastic function $a(t) : [0, \infty) \mapsto \mathbb{R}$.

REMARK 1. It is easy to check that (2.1) is equivalent to

$$\int_{[-1,1]} |x|^{1/\kappa} \Pi(dx) < \infty.$$

The latter holds for $0 < \kappa \leq 1/2$ for any Lévy process, as a fundamental property of the Lévy canonical measure ([1], page 13). Equation (2.2) always holds when $0 < \kappa < 1/2$, as mentioned in Section 1, but not necessarily when $\kappa = 1/2$.

The case $\kappa = 1/2$ which is excluded in Theorem 2.1 turns out to have interesting and unexpected features. To put these in context, let us first review some background. Khintchine [10] (see also [15], Proposition 47.12, page 358) showed that, given any function $h(t)$, positive, continuous and nondecreasing in a neighborhood of 0, and satisfying $h(t) = o(\sqrt{t \log |\log t|})$ as $t \downarrow 0$, there is a Lévy process with $\sigma^2 = 0$ such that $\limsup_{t \downarrow 0} |X_t|/h(t) = \infty$ a.s. For example,

we can take $h(t) = \sqrt{t}(\log|\log t|)^{1/4}$. The corresponding Lévy process satisfies $\limsup_{t \downarrow 0} |X_t|/\sqrt{t} = \infty$ a.s. Since (2.1) always holds when $\kappa = 1/2$, we see that the implication (2.1) \Rightarrow (2.2) is not in general true when $\kappa = 1/2$.

On the other hand, when $\kappa = 1/2$, Theorem 2.1 remains true, for example, when $X \in bv$, in the sense that both (2.1) and (2.2) then hold, as follows from the fact that $X_t = O(t)$ a.s. as $t \downarrow 0$.

Thus we can have $\limsup_{t \downarrow 0} |X_t|/\sqrt{t}$ equal to 0 or ∞ a.s., and we are led to ask for a NASC to decide between the alternatives. We give such a condition in Theorem 2.2 and furthermore show that $\limsup_{t \downarrow 0} |X_t|/\sqrt{t}$ may take a positive finite value, a.s. Remarkably, Theorem 2.2 simultaneously solves the one-sided problem. These one-sided cases are further investigated in Section 3, where it will be seen that, by contrast, the one- and two-sided situations are completely different when $\kappa \neq 1/2$.

To state the theorem, we need the notation

$$(2.3) \quad V(x) = \int_{|y| \leq x} y^2 \Pi(dy), \quad x > 0.$$

Since we assume X is not compound Poisson, we have $V(x) > 0$ for all $x > 0$.

THEOREM 2.2 (The case $\kappa = 1/2$). *Assume (1.7), and put*

$$I(\lambda) = \int_0^1 x^{-1} \exp\left(-\frac{\lambda^2}{2V(x)}\right) dx$$

and

$$\lambda_I^* := \inf\{\lambda > 0 : I(\lambda) < \infty\} \in [0, \infty]$$

(with the convention, throughout, that the inf of the empty set is $+\infty$). Then, a.s.,

$$(2.4) \quad -\liminf_{t \downarrow 0} \frac{X_t}{\sqrt{t}} = \limsup_{t \downarrow 0} \frac{X_t}{\sqrt{t}} = \limsup_{t \downarrow 0} \frac{|X_t|}{\sqrt{t}} = \lambda_I^*.$$

REMARK 2. (i) Equation (2.4) forms a nice counterpart to the iterated log version in (1.5) and (1.6).

(ii) If (2.1) holds for some $\kappa > 1/2$, then $V(x) = o(x^{2-1/\kappa})$ as $x \downarrow 0$, so $I(\lambda)$ converges for all $\lambda > 0$. Thus $\lambda_I^* = 0$ and $\lim_{t \downarrow 0} X_t/\sqrt{t} = 0$ a.s. in this case, according to Theorem 2.2. Of course, this agrees with Theorem 2.1(i).

(iii) The convergence of $\int_{|x| \leq e^{-e}} x^2 \log|\log|x|| \Pi(dx)$ implies the convergence of $\int_0^1 \exp\{-\lambda^2/2V(x)\} dx/x$ for all $\lambda > 0$, as is easily checked; hence we have $\lim_{t \downarrow 0} |X_t|/\sqrt{t} = 0$ a.s. for all such Lévy processes. A finite positive value, a.s., for $\limsup_{t \downarrow 0} |X_t|/\sqrt{t}$ can occur only in a small class of Lévy processes whose canonical measures have $\Pi(dx)$ close to $|x|^{-3} dx$ near 0. For example, we can find a Π such that, for small x , $V(x) = 1/\log|\log x|$. Then

TABLE 1

Value of κ	NASC for $\limsup_{t \downarrow 0} X_t /t^\kappa = \infty$ a.s. (when $\sigma^2 = 0$)
$0 \leq \kappa < \frac{1}{2}$	never true
$\kappa = \frac{1}{2}$	$\lambda_I^* = \infty$ (see Theorem 2.2)
$\frac{1}{2} < \kappa \leq 1$	$\int_0^1 \bar{\Pi}(x^\kappa) dx = \infty$
$\kappa > 1, X \in bv, \delta = 0$	$\int_0^1 \bar{\Pi}(x^\kappa) dx = \infty$
$\kappa > 1, X \in bv, \delta \neq 0$	always true
$\kappa > 1, X \notin bv$	always true

$\int_0^{1/2} \exp\{-\lambda^2/2V(x)\} dx/x = \int_0^{1/2} |\log x|^{-\lambda^2/2} dx/x$ is finite for $\lambda > \sqrt{2}$ but infinite for $\lambda \leq \sqrt{2}$. Thus $\limsup_{t \downarrow 0} |X_t|/\sqrt{t} = \sqrt{2}$ a.s. for this process; in fact, $\limsup_{t \downarrow 0} X_t/\sqrt{t} = \sqrt{2}$ a.s., and $\liminf_{t \downarrow 0} X_t/\sqrt{t} = -\sqrt{2}$ a.s.

(iv) Theorem 2.2 tells us that the only possible a.s. limit, as $t \downarrow 0$, of X_t/\sqrt{t} is 0, and that this occurs iff $\lambda_I^* = 0$, that is, iff $I(\lambda) < \infty$ for all $\lambda > 0$. Similarly, the iterated log version in (1.6) gives that the only possible a.s. limit, as $t \downarrow 0$, of $X_t/\sqrt{t \log |\log t|}$ is 0, and that this occurs iff $\sigma^2 = 0$. When $\kappa > 1/2$, Theorem 2.1 gives that $\lim_{t \downarrow 0} X_t/t^\kappa = 0$ a.s. iff $\int_0^1 \bar{\Pi}(x^\kappa) dx < \infty$, provided, when $\kappa \geq 1$, the drift $\delta = 0$.

Another result that follows from our analysis is that it is possible to have $\lim_{t \downarrow 0} X_t/t^\kappa = \delta$ a.s. for a constant δ with $0 < |\delta| < \infty$, and $\kappa > 0$, iff $\kappa = 1$, $X \in bv$, δ is the drift, and $\delta \neq 0$.

(v) It is immediate from Theorem 2.1 that centering has no effect in the two-sided case, in the following sense. Assume (1.7), and, if $X \in bv$, that it has drift zero. If $\limsup_{t \downarrow 0} |X_t|/t^\kappa = \infty$ a.s., for some $\kappa > 1/2$, then we have

$$(2.5) \quad \limsup_{t \downarrow 0} \frac{|X_t - a(t)|}{t^\kappa} = \infty \quad \text{a.s., for any nonstochastic function } a(t).$$

We can show that this also holds in the case $\kappa = 1/2$; see Remark 5 following the proof of Theorem 2.2.

Finally, in this section, Table 1 summarizes the conditions for (1.3).

3. One-sided case. We wish to test for the finiteness or otherwise of $\limsup_{t \downarrow 0} X_t/t^\kappa$, so we proceed by finding conditions for

$$(3.1) \quad \limsup_{t \downarrow 0} \frac{X_t}{t^\kappa} = +\infty \quad \text{a.s.}$$

In view of the discussion in Section 1, and the fact that the case $\kappa = 1/2$ is covered in Theorem 2.2, we have only two cases to consider:

- (a) $X \notin bv, 1/2 < \kappa < 1;$
- (b) $X \in bv,$ with drift $\delta = 0, \kappa > 1.$

We deal first with case (a). For this, we need to define, for $0 < y \leq 1,$

$$W(y) := \int_0^y \int_x^1 z \Pi^{(-)}(dz) dx,$$

and then, for $\lambda > 0,$

$$(3.2) \quad J(\lambda) := \int_0^1 \exp\left\{-\lambda \left(\frac{y^{(2\kappa-1)/\kappa}}{W(y)}\right)^{\kappa/(1-\kappa)}\right\} \frac{dy}{y}.$$

Also let $\lambda_j^* := \inf\{\lambda > 0 : J(\lambda) < \infty\} \in [0, \infty].$

The following theorem has two features which are surprising at first sight. The first is that (3.1) can hold by virtue of $\bar{\Pi}^{(-)}(\cdot)$ being large compared to $\bar{\Pi}^{(+)}(\cdot),$ as in part (ii); this is of course the effect of compensation. The second is that we can have the lim sup taking a value in $(0, \infty),$ as in part (v). Both these features mimic corresponding results as $t \rightarrow \infty;$ see [7].

THEOREM 3.1. *Assume (1.7) and keep $1/2 < \kappa < 1.$ Then (3.1) holds if and only if:*

- (i) $\int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx = \infty,$ or
 - (ii) $\int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx < \infty = \int_0^1 \bar{\Pi}^{(-)}(x^\kappa) dx,$ and $\lambda_j^* = \infty.$
- When (i) and (ii) fail, we have in greater detail: suppose*
- (iii) $\int_0^1 \bar{\Pi}(x^\kappa) dx < \infty,$ or
 - (iv) $\int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx < \infty = \int_0^1 \bar{\Pi}^{(-)}(x^\kappa) dx$ and $\lambda_j^* = 0.$
- Then*

$$(3.3) \quad \limsup_{t \downarrow 0} \frac{X_t}{t^\kappa} = 0 \quad a.s.$$

Alternatively, suppose

- (v) $\int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx < \infty = \int_0^1 \bar{\Pi}^{(-)}(x^\kappa) dx$ and $\lambda_j^* \in (0, \infty).$ Then

$$(3.4) \quad \limsup_{t \downarrow 0} \frac{X_t}{t^\kappa} = c \quad a.s., \text{ for some } c \in (0, \infty).$$

REMARK 3. (i) In the special case that $\bar{\Pi}^{(-)}(x) \sim 1/(x^{1/\kappa} L(x))$ as $x \downarrow 0,$ with $L(x)$ slowly varying at 0, we have $W(x) \sim cx^{2-1/\kappa}/L(x),$ and it follows that the above results hold with $J(\lambda)$ replaced by

$$\int_0^1 \exp\{-\lambda(L(y))^{\kappa/(1-\kappa)}\} \frac{dy}{y}.$$

So, if $\ell_k(y)$ denotes the k th iterated logarithm of $1/y,$ then taking $L(y)$ equal to (i) $(\ell_3(y))^{(1-\kappa)/\kappa},$ (ii) $(\ell_1(y))^{(1-\kappa)/\kappa}$ and (iii) $(\ell_2(y))^{(1-\kappa)/\kappa},$ gives examples having $\lambda_j^* = \infty, \lambda_j^* = 0$ and $0 < \lambda_j^* < \infty,$ respectively, in the situation of Theo-

rem 3.1.

(ii) Note that $X \notin bv$ when $\int_0^1 \overline{\Pi}^{(+)}(x^\kappa) dx = \infty$ or $\int_0^1 \overline{\Pi}^{(-)}(x^\kappa) dx = \infty$ in Theorem 3.1, because $\overline{\Pi}^{(\pm)}(x^\kappa) \leq \overline{\Pi}^{(\pm)}(x)$ when $0 < x < 1$ and $0 < \kappa < 1$, so $\int_0^1 \overline{\Pi}(x) dx = \infty$.

(iii) It may seem puzzling at first that a second moment-like function, $V(\cdot)$, appears in Theorem 2.2, whereas $W(\cdot)$, a kind of integrated first moment function, appears in Theorem 3.1. Though closely related, in general, $V(x)$ is not asymptotically equivalent to $W(x)$, as $x \rightarrow 0$, and neither function is asymptotically equivalent to yet another second moment-like function on $[0, \infty)$, $U(x) := V(x) + x^2 \overline{\Pi}(x)$. $V(x)$ arises naturally in the proof of Theorem 2.2, which uses a normal approximation to certain probabilities, whereas $W(x)$ arises naturally in the proof of Theorem 3.1, which uses Laplace transforms and works with spectrally one-sided Lévy processes. It is possible to reconcile the different expressions; in fact, Theorem 2.2 remains true if V is replaced in the integral $I(\lambda)$ by U or by W . Thus these three functions are equivalent in the context of Theorem 2.2 (but not in general). We explain this in a little more detail following the proof of Theorem 3.1.

Next we turn to case (b). When $X \in bv$ we can define, for $0 < x < 1$,

$$(3.5) \quad A_+(x) = \int_0^x \overline{\Pi}^{(+)}(y) dy \quad \text{and} \quad A_-(x) = \int_0^x \overline{\Pi}^{(-)}(y) dy.$$

THEOREM 3.2. *Assume (1.7), suppose $\kappa > 1$, $X \in bv$, and its drift $\delta = 0$. If*

$$(3.6) \quad \int_{(0,1]} \frac{\Pi^{(+)}(dx)}{x^{-1/\kappa} + A_-(x)/x} = \infty,$$

then (3.1) holds. Conversely, if (3.6) fails, then $\limsup_{t \downarrow 0} X_t/t^\kappa \leq 0$ a.s.

REMARK 4. (i) It is natural to inquire whether (3.6) can be simplified by considering separately integrals containing the components of the integrand in (3.6); note that (3.6) is equivalent to

$$\int_{(0,1]} \min\left(x^{1/\kappa}, \frac{x}{A_-(x)}\right) \Pi^{(+)}(dx) = \infty.$$

This is not the case. For each $\kappa > 1$, it is possible to find a Lévy process $X \in bv$ with drift 0 for which (3.6) fails but

$$\int_{(0,1]} x^{1/\kappa} \Pi^{(+)}(dx) = \infty = \int_{(0,1]} (x/A_-(x)) \Pi^{(+)}(dx).$$

The idea is to construct a continuous increasing concave function which is linear on a sequence of intervals tending to 0, which can serve as an $A_-(x)$, and which oscillates around the function $x \mapsto x^{1-1/\kappa}$. We will omit the details of the construction.

(ii) It is possible to have $\limsup_{t \downarrow 0} X_t/t^\kappa < 0$ a.s., in the situation of Theorem 3.2, when (3.6) fails; for example, when X is the negative of a subordinator with zero drift. The value of the lim sup can then be determined by applying Lemma 5.3 in Section 4.

(iii) For another equivalence, we note that (3.6) holds if and only if

$$(3.7) \quad \int_0^1 \overline{\Pi}^{(+)}(t^\kappa + X_t^{(-)}) dt = \infty \quad \text{a.s.},$$

where $X^{(-)}$ is a subordinator with drift 0 and Lévy measure $\Pi^{(-)}$. This can be deduced from Erickson and Maller ([8], Theorem 1(i) and Example 1) observing with the notation therein that

$$A(x) = \gamma + \overline{\Pi}^{(-)}(1) - \int_x^1 \overline{\Pi}^{(-)}(y) dy = \int_0^x \overline{\Pi}^{(-)}(y) dy,$$

since there is no drift. Thus the function $A(x)$ in [8] is $A_-(x)$, in our notation. This provides a connection between the almost sure divergence of the Lévy integral in (3.7) and the upcrossing condition (3.1). Note that both (3.6) and (3.7) express in some way the dominance of the small positive jumps over the small negative jumps of X . Hence the phenomenon mentioned prior to the statement of Theorem 3.1—that (3.1) can occur in some situations even when the negative jumps dominate the positive, in a certain sense—does not occur in the situation of Theorem 3.2.

Table 2 summarizes the conditions for (3.1).

Our final theorem applies the foregoing results to give a criterion for

$$(3.8) \quad \lim_{t \downarrow 0} \frac{X_t}{t^\kappa} = +\infty \quad \text{a.s.}$$

This is a stronger kind of divergence of the normed process to ∞ , for small times. A straightforward analysis of cases, using our one- and two-sided results, shows that (3.8) never occurs if $0 < \kappa \leq 1$, if $\kappa > 1$ and $X \notin bv$, or if $\kappa > 1$ and $X \in bv$

TABLE 2

Value of κ	NASC for $\limsup_{t \downarrow 0} X_t/t^\kappa = \infty$ a.s. (when $\sigma^2 = 0$)
$0 \leq \kappa < \frac{1}{2}$	never true
$\kappa = \frac{1}{2}, X \notin bv$	see Theorem 2.2
$\frac{1}{2} < \kappa < 1, X \notin bv$	see Theorem 3.1
$\frac{1}{2} \leq \kappa \leq 1, X \in bv$	never true
$\kappa > 1, X \in bv, \delta < 0$	never true
$\kappa > 1, X \in bv, \delta = 0$	see Theorem 3.2
$\kappa > 1, X \in bv, \delta > 0$	always true
$\kappa \geq 1, X \notin bv$	always true

with negative drift. If $\kappa > 1$ and $X \in bv$ with positive drift, (3.8) always occurs. That leaves just one case to consider, in:

THEOREM 3.3. *Assume (1.7), suppose $\kappa > 1$, $X \in bv$, and its drift $\delta = 0$. Then (3.8) holds iff*

$$(3.9) \quad K_X(d) := \int_0^1 \frac{dy}{y} \exp\left\{-d \frac{(A_+(y))^{\kappa/(\kappa-1)}}{y}\right\} < \infty \quad \text{for all } d > 0$$

and

$$(3.10) \quad \int_{(0,1]} \left(\frac{x}{A_+(x)}\right) \Pi^{(-)}(dx) < \infty.$$

Concluding remarks. It is natural to ask: when is

$$(3.11) \quad \limsup_{t \downarrow 0} \frac{X_t - a(t)}{t^\kappa} < \infty \quad \text{a.s., for some nonstochastic } a(t)?$$

Phrased in such a general way, the question is not interesting since we can always make $X_t = o(a(t))$ a.s. as $t \downarrow 0$ by choosing $a(t)$ large enough by comparison with X_t [e.g., $a(t)$ such that $a(t)/\sqrt{t \log |\log t|} \rightarrow \infty$, as $t \downarrow 0$, will do, by (1.5)], so the lim sup in (3.11) becomes negative. So we would need to restrict $a(t)$ in some way. Section 3 deals with the case $a(t) = 0$. Another choice is to take $a(t)$ as a natural centering function such as EX_t or as a median of X_t . However, in our small time situation, EX_t is essentially 0 or the drift of X , so we are led back to the case $a(t) = 0$ again (and similarly for the median). Of course there may be other interesting choices of $a(t)$ in some applications, and there is the wider issue of replacing t^κ by a more general norming function. Some of our results in Sections 4 and 5 address the latter, but we will not pursue these points further here.

4. Proofs for Section 2.

4.1. *Proof of Theorem 2.1.* The proof relies on a pair of technical results which we will establish first. Recall the notation $V(x)$ in (2.3).

PROPOSITION 4.1. *Let $b: \mathbb{R}_+ \rightarrow [0, \infty)$ be any nondecreasing function such that*

$$(4.1) \quad \int_0^1 \overline{\Pi}(b(x)) dx < \infty \quad \text{and} \quad \int_0^1 V(b(x)) b^{-2}(x) dx < \infty.$$

Then

$$(4.2) \quad \limsup_{t \downarrow 0} \frac{|X_t - \tilde{a}(t)|}{b(t)} \leq 1 \quad \text{a.s.,}$$

where

$$(4.3) \quad \tilde{a}(t) := \gamma t - \int_0^t ds \int_{b(s) < |x| \leq 1} x \Pi(dx), \quad t \geq 0.$$

PROOF. Recall the Lévy–Itô decomposition (1.9). In this setting, it is convenient to introduce

$$X_t^{(1)} := \int_{[0,t] \times [0,1]} \mathbf{1}_{\{|x| \leq b(s)\}} x N(ds, dx)$$

and

$$X_t^{(2)} := \gamma t + \int_{[0,t] \times [0,1]} \mathbf{1}_{\{b(s) < |x| \leq 1\}} x N(ds, dx),$$

where again the stochastic integrals are taken in the compensated sense. Plainly, $X = X^{(1)} + X^{(2)}$.

The assumption $\int_0^1 \bar{\Pi}(b(x)) dx < \infty$ implies that

$$N(\{(s, x) : 0 \leq s \leq t \text{ and } b(s) < |x| \leq 1\}) = 0$$

whenever $t > 0$ is sufficiently small a.s., and in this situation $X^{(2)}$ is just γt minus the compensator, a.s., that is,

$$X_t^{(2)} = \gamma t - \int_0^t ds \int_{b(s) < |x| \leq 1} x \Pi(dx) = \tilde{a}(t).$$

On the one hand, $X^{(1)}$ is a square-integrable martingale with oblique bracket

$$\langle X^{(1)} \rangle_t = \int_0^t ds \int_{|x| \leq b(s)} x^2 \Pi(dx) = \int_0^t V(b(s)) ds \leq tV(b(t)).$$

By Doob’s maximal inequality, we have for every $t \geq 0$

$$P\left(\sup_{0 \leq s \leq t} |X_s^{(1)}| > b(2t)\right) \leq 4tV(b(t))b^{-2}(2t).$$

On the other hand, the assumptions that $b(t)$ is nondecreasing and that $\int_0^1 dx \times V(b(x))b^{-2}(x) < \infty$ entail

$$\sum_{n=1}^{\infty} 2^{-n} V(b(2^{-n}))b^{-2}(2^{-n+1}) < \infty.$$

By the Borel–Cantelli lemma, we thus see that

$$\lim_{n \rightarrow \infty} \frac{\sup_{0 \leq s \leq 2^{-n}} |X_s^{(1)}|}{b(2^{-n+1})} \leq 1 \quad \text{a.s.}$$

Then (4.2) follows with $b(4t)$ rather than $b(t)$ in the denominator by a standard argument of monotonicity. As we can change variable to replace $b(x)$ by $b(x/4)$ in (4.1), the factor of 4 is irrelevant and we get (4.2) as stated. \square

PROPOSITION 4.2. *Suppose there are nonstochastic functions $a: \mathbb{R}_+ \rightarrow \mathbb{R}$ and $b: (0, \infty) \rightarrow (0, \infty)$, with b measurable, such that*

$$(4.4) \quad P\left(\limsup_{t \downarrow 0} \frac{|X_t - a(t)|}{b(t)} < \infty\right) > 0.$$

Then there is some finite constant C such that

$$(4.5) \quad \int_0^1 \bar{\Pi}(Cb(x)) dx < \infty.$$

PROOF. Symmetrize X by subtracting an independent equally distributed X' to get $X_t^{(s)} = X_t - X'_t, t \geq 0$. Then (4.4) and Blumenthal’s 0–1 law imply there is some finite constant C such that

$$(4.6) \quad \limsup_{t \downarrow 0} \frac{|X_t^{(s)}|}{b(t)} < \frac{C}{2} \quad \text{a.s.}$$

Suppose now that (4.5) fails. Note that $\bar{\Pi}^{(s)}(\cdot) = 2\bar{\Pi}(\cdot)$, where $\Pi^{(s)}$ is the Lévy measure of $X^{(s)}$, so that $\int_0^1 \bar{\Pi}^{(s)}(Cb(x)) dx = \infty$. Then from the Lévy–Itô decomposition, we have that, for every $\varepsilon > 0$,

$$\#\{t \in (0, \varepsilon]: |\Delta_t^{(s)}| > Cb(t)\} = \infty \quad \text{a.s.,}$$

where $\Delta_t^{(s)} = X_t^{(s)} - X_{t-}^{(s)}$. But when $|\Delta_t^{(s)}| > Cb(t)$, we have $|X_{t-}^{(s)}| > Cb(t)/2$ or $|X_t^{(s)}| > Cb(t)/2$, which contradicts (4.6). Thus (4.5) holds. \square

Finally, we will need an easy deterministic bound.

LEMMA 4.1. *Fix some $\kappa \geq 1/2$, put $b(t) = t^\kappa$, and assume*

$$(4.7) \quad \int_{|x| < 1} |x|^{1/\kappa} \Pi(dx) < \infty.$$

When $\kappa \geq 1$, $X \in bv$ and we suppose further that the drift coefficient $\delta = \gamma - \int_{|x| \leq 1} x \Pi(dx)$ is 0. Then, as $t \rightarrow 0$,

$$\tilde{a}(t) = \gamma t - \int_0^t ds \int_{s^\kappa < |x| \leq 1} x \Pi(dx) = o(t^\kappa).$$

PROOF. Suppose first $\kappa < 1$. For every $0 < \varepsilon < \eta < 1$, we have

$$\begin{aligned} \int_{\varepsilon < |x| \leq 1} |x| \Pi(dx) &\leq \varepsilon^{1-1/\kappa} \int_{|x| \leq \eta} |x|^{1/\kappa} \Pi(dx) + \int_{\eta < |x| \leq 1} |x| \Pi(dx) \\ &= \varepsilon^{1-1/\kappa} o_\eta + c(\eta) \quad \text{say,} \end{aligned}$$

where, by (4.7), $\lim_{\eta \downarrow 0} o_\eta = 0$. Since $\kappa < 1$, it follows that

$$\limsup_{t \downarrow 0} |\tilde{a}(t)| t^{-\kappa} \leq \kappa^{-1} o_\eta,$$

and as we can take η arbitrarily small, we conclude that $\tilde{a}(t) = o(t^\kappa)$.

In the case $\kappa \geq 1$, X has bounded variation with zero drift coefficient. We may rewrite $\tilde{a}(t)$ in the form

$$\tilde{a}(t) = \int_0^t ds \int_{|x| \leq s^\kappa} x \Pi(dx).$$

The assumption (4.7) entails $\int_{|x| \leq \varepsilon} |x| \Pi(dx) = o(\varepsilon^{1-1/\kappa})$ and we again conclude that $\tilde{a}(t) = o(t^\kappa)$. \square

We now have all the ingredients to establish Theorem 2.1.

PROOF OF THEOREM 2.1. Keep $\kappa > 1/2$ throughout.

(i) Suppose (2.1) holds, which is equivalent to (4.7). Writing $|x|^{1/\kappa} = |x|^{1/\kappa-2} x^2$, we see from an integration by parts (Fubini's theorem) that $\int_0^1 V(x) \times x^{1/\kappa-3} dx < \infty$. Note that the assumption that $\kappa \neq 1/2$ is crucial in this step. The change of variables $x = y^\kappa$ now gives that $\int_0^1 V(y^\kappa) y^{-2\kappa} dy < \infty$. We may thus apply Proposition 4.1 and get that

$$\limsup_{t \downarrow 0} \frac{|X_t - \tilde{a}(t)|}{t^\kappa} \leq 1 \quad \text{a.s.}$$

By Lemma 4.1, $\tilde{a}(t) = o(t^\kappa)$ as $t \downarrow 0$. We thus have shown that when (2.1) holds,

$$\limsup_{t \downarrow 0} \frac{|X_t|}{t^\kappa} \leq 1 \quad \text{a.s.}$$

For every $c > 0$, the time-changed process X_{ct} is a Lévy process with Lévy measure $c\Pi$, so we also have $\limsup_{t \downarrow 0} |X_{ct}| t^{-\kappa} \leq 1$ a.s. As we may take c as large as we wish, we conclude that (2.2) holds.

(ii) By Proposition 4.2, if

$$P\left(\limsup_{t \downarrow 0} \frac{|X_t - a(t)|}{t^\kappa} < \infty\right) > 0,$$

then $\int_0^1 \overline{\Pi}(Cx^\kappa) dx < \infty$ for some finite constant C . By an obvious change of variables, this shows that (2.1) must hold. This completes the proof of Theorem 2.1. \square

REMARK 5. It can be shown that if, for nonstochastic functions $a(t), b(t) > 0$, with $b(t)$ nondecreasing,

$$(4.8) \quad \limsup_{t \downarrow 0} \frac{|X_t - a(t)|}{b(t)} \leq C < \infty \quad \text{a.s.},$$

then $(X_t - tv(b(t)))/b(t) \xrightarrow{P} 0$, where $v(x) := \gamma - \int_{x < |y| \leq 1} y \Pi(dy)$ for $x > 0$. Since $(X_t - a(t))/b(t)$ is also stochastically bounded, under (4.8), we deduce that $|a(t) - tv(b(t))| = O(b(t))$ as $t \downarrow 0$, when (4.8) holds. But from $\int_{|x| \leq 1} x^2 \Pi(dx) < \infty$, we see that $v(x) = o(x^{-1})$ as $x \downarrow 0$. So if (4.8) holds with $b(t) = \sqrt{t}$, then $a(t) = O(\sqrt{t})$, and so $\limsup |X_t|/\sqrt{t} < \infty$ a.s. We conclude that if, in Theorem 2.2, $\lambda_I^* = \infty$, then (2.5) holds with $\kappa = 1/2$.

4.2. *Proof of Theorem 2.2.* We now turn our attention to Theorem 2.2 and develop some notation and material in this direction. Write, for $b > 0$,

$$(4.9) \quad X_t = Y_t^{(b)} + Z_t^{(b)},$$

with

$$(4.10) \quad \begin{aligned} Y_t^{(b)} &:= \int_{[0,t] \times [-1,1]} \mathbf{1}_{\{|x| \leq b\}} x N(ds, dx), \\ Z_t^{(b)} &:= \gamma t + \int_{[0,t] \times [-1,1]} \mathbf{1}_{\{b < |x|\}} x N(ds, dx), \end{aligned}$$

where $N(ds, dx)$ is a Poisson random measure on $[0, \infty) \times [-1, 1]$ with intensity $ds \Pi(dx)$, and the stochastic integrals are taken in the compensated sense.

LEMMA 4.2 (No assumptions on X). *For every $0 < r < 1$ and $\varepsilon > 0$, we have*

$$\sum_{n=1}^{\infty} P\left(\sup_{0 \leq t \leq r^n} |Z_t^{(r^{n/2})}| > \varepsilon r^{n/2}\right) < \infty,$$

and as a consequence,

$$\lim_{n \rightarrow \infty} r^{-n/2} \sup_{0 \leq t \leq r^n} |Z_t^{(r^{n/2})}| = 0 \quad a.s.$$

PROOF. Introduce, for every integer n , the set

$$A_n := [0, r^n] \times ([-1, -r^{n/2}) \cup (r^{n/2}, 1]),$$

and note that

$$\begin{aligned} &P\left(\sup_{0 \leq t \leq r^n} |Z_t^{(r^{n/2})}| > \varepsilon r^{n/2}\right) \\ &\leq P\left(\sup_{0 \leq t \leq r^n} |Z_t^{(r^{n/2})}| > \varepsilon r^{n/2}, N(A_n) = 0\right) \\ &\quad + P(N(A_n) > 0). \end{aligned}$$

We have

$$\begin{aligned} \sum_{n=1}^{\infty} P(N(A_n) > 0) &\leq \sum_{n=1}^{\infty} r^n \bar{\Pi}(r^{n/2}) \\ &\leq (1-r)^{-1} \sum_{n=1}^{\infty} \int_{r^{n+1}}^{r^n} \bar{\Pi}(\sqrt{x}) dx \\ &\leq (1-r)^{-1} \int_0^1 \bar{\Pi}(\sqrt{x}) dx, \end{aligned}$$

and this last integral is finite (always).

On the other hand, on the event $N(A_n) = 0$, we have

$$Z_t^{(r^{n/2})} = t \left(\gamma - \int_{r^{n/2} < |x| \leq 1} x \Pi(dx) \right), \quad 0 \leq t \leq r^n.$$

Again as a result of the convergence of $\int_{|x| \leq 1} x^2 \Pi(dx)$, the argument in Lemma 4.1 shows that the supremum over $0 \leq t \leq r^n$ of the absolute value of the right-hand side is $o(r^{n/2})$, and the convergence of the series follows. The second statement then follows from the Borel–Cantelli lemma. \square

In view of Lemma 4.2 we can concentrate on $Y_t^{(\sqrt{t})}$ in (4.9). We next prove:

LEMMA 4.3. *Let Y be a Lévy process of the form $Y_t = \int_0^t \int_{\mathbb{R}} z N_Y(ds, dz)$, where $N_Y(ds, dz)$ is a Poisson measure with intensity $ds \Pi_Y(dz)$, the integral being taken in the compensated sense. Assume Y satisfies $m_4 < \infty$, where $m_k := \int_{x \in \mathbb{R}} |x|^k \Pi_Y(dx)$, $k = 2, 3, \dots$*

(i) *Then*

$$\lim_{t \downarrow 0} \frac{1}{t} E|Y_t|^3 = m_3.$$

(ii) *For any $x > 0, t > 0$, we have the nonuniform bound*

$$(4.11) \quad |P(Y_t > x\sqrt{tm_2}) - \bar{F}(x)| \leq \frac{Am_3}{\sqrt{tm_2}^{3/2} (1+x)^3},$$

where A is an absolute constant and

$$\bar{F}(x) = \int_x^{\infty} e^{-y^2/2} dy / \sqrt{2\pi} = \frac{1}{2} \operatorname{erfc}(x/\sqrt{2})$$

is the tail of the standard normal distribution function.

PROOF. (i) By expanding the Lévy–Khintchine formula and using $m_4 < \infty$ we can calculate $EY_t^4 = tm_4 + 3t^2m_2^2$. So by Chebyshev’s inequality for second

and fourth moments, for $x > 0, t > 0$,

$$\frac{1}{t}P(|Y_t| > x) \leq \frac{m_2}{x^2}\mathbf{1}_{\{0 < x \leq 1\}} + \frac{m_4 + 3tm_2^2}{x^4}\mathbf{1}_{\{x > 1\}}.$$

We can also calculate

$$\frac{1}{t}E|Y_t|^3 = \frac{3}{t} \int_0^\infty x^2 P(|Y_t| > x) dx.$$

By Bertoin ([1], Example 1, page 39), $P(|Y_t| > x)/t \rightarrow \bar{\Pi}_Y(x)$, as $t \downarrow 0$, for each $x > 0$. The result (i) follows by dominated convergence.

(ii) Write $Y_t = \sum_{i=1}^n Y(i, t)$, for $n = 1, 2, \dots$, where $Y(i, t) := Y(it/n) - Y((i - 1)t/n)$ are i.i.d., each with the distribution of $Y(t/n)$. According to a nonuniform Berry–Esseen bound (Theorem 14, page 125 of [11]), for each $n = 1, 2, \dots$, (4.11) holds with the right-hand side replaced by

$$\frac{AE|Y(t/n)|^3}{\sqrt{n}(tm_2/n)^{3/2}(1+x)^3} = \frac{AE|Y(t/n)|^3/(t/n)}{\sqrt{t}m_2^{3/2}(1+x)^3}.$$

By part (i) this tends as $n \rightarrow \infty$ to the right-hand side of (4.11). \square

We now establish a result which plays a key role in the proof of Theorem 2.2.

PROPOSITION 4.3. *In the notation (4.10), we have, for $\lambda > 0, 0 < r < 1$,*

$$(4.12) \quad \begin{aligned} &\sum_{n \geq 0} P(Y_r^{(r^{n/2})} > \lambda r^{n/2}) < \infty \\ &\iff \int_0^1 \sqrt{V(x)} \exp\left(\frac{-\lambda^2}{2V(x)}\right) \frac{dx}{x} < \infty. \end{aligned}$$

PROOF. For every fixed $t > 0$, $Y_s^{(\sqrt{t})}$ is the compensated sum of jumps of X smaller in magnitude than \sqrt{t} , up to time s . It is a centered Lévy process with canonical measure $\mathbf{1}_{\{|x| \leq \sqrt{t}\}} \Pi(dx)$, $x \in \mathbb{R}$. Applying Lemma 4.3, we get $m_2 = V(\sqrt{t})$ and $m_3 = \int_{|y| \leq \sqrt{t}} |y|^3 \Pi(dy) = \rho(\sqrt{t})$, say. Then we get, for $x > 0$,

$$|P(Y_t^{(\sqrt{t})} > x\sqrt{tV(\sqrt{t})}) - \bar{F}(x)| \leq \frac{A\rho(\sqrt{t})}{\sqrt{tV^3(\sqrt{t})(1+x)^3}}.$$

Replacing x by $\lambda/\sqrt{V(\sqrt{t})}$, $a > 0$, we have

$$|P(Y_t^{(\sqrt{t})} > \lambda\sqrt{t}) - \bar{F}(\lambda/\sqrt{V(\sqrt{t})})| \leq \varepsilon(t) := \frac{A\rho(\sqrt{t})}{\sqrt{t}\lambda^3},$$

and we claim that $\sum \varepsilon(r^n) < \infty$. In fact, for some $c > 0$,

$$\begin{aligned} \sum_{n=0}^{\infty} \frac{\rho(r^{n/2})}{r^{n/2}} &= \sum_{n=0}^{\infty} \frac{1}{r^{n/2}} \sum_{j \geq n} \int_{r^{(j+1)/2} < |y| \leq r^{j/2}} |y|^3 \Pi(dy) \\ &= \sum_{j=0}^{\infty} \left(\sum_{n=0}^j r^{-n/2} \right) \int_{r^{(j+1)/2} < |y| \leq r^{j/2}} |y|^3 \Pi(dy) \\ &\leq c \sum_{j=0}^{\infty} r^{-j/2} \int_{r^{(j+1)/2} < |y| \leq r^{j/2}} |y|^3 \Pi(dy) \\ &\leq c \sum_{j=0}^{\infty} \int_{r^{(j+1)/2} < |y| \leq r^{j/2}} y^2 \Pi(dy) \\ &= c \int_{|y| \leq 1} y^2 \Pi(dy) < \infty. \end{aligned}$$

The result (4.12) follows, since the monotonicity of \bar{F} shows that the convergence of $\sum_{n \geq 1} \bar{F}(\lambda/\sqrt{V(r^{n/2})})$ is equivalent to that of

$$\int_0^1 \bar{F}(\lambda/\sqrt{V(\sqrt{x})}) \frac{dx}{x} = \sum \int_{r^{(n+1)/2}}^{r^{n/2}} \bar{F}(\lambda/\sqrt{V(\sqrt{x})}) \frac{dx}{x},$$

and it is well known that $\bar{F}(x) \sim (2\pi)^{-1/2} x^{-1} e^{-x^2/2}$ as $x \rightarrow \infty$. \square

We can now establish Theorem 2.2.

PROOF OF THEOREM 2.2. Recall the definition of $I(\cdot)$ in the statement of Theorem 2.2. We will first show that for every given $\lambda > 0$

$$(4.13) \quad I(\lambda) < \infty \implies \limsup_{t \downarrow 0} \frac{X_t}{\sqrt{t}} \leq \lambda \quad \text{a.s.}$$

To see this, observe that when $I(\lambda) < \infty$, the integral in (4.12) converges, hence so does the series. Use the maximal inequality in Theorem 12, page 50 of Petrov [11], to get, for $t > 0, b > 0, x > 0$,

$$\begin{aligned} P\left(\sup_{0 < s \leq t} Y_s^{(b)} > x\right) &= \lim_{k \rightarrow \infty} P\left(\max_{1 \leq j \leq [kt]} Y_{j/k}^{(b)} > x\right) \\ (4.14) \quad &= \lim_{k \rightarrow \infty} P\left(\max_{1 \leq j \leq [kt]} \sum_{i=1}^j \Delta(i, k, b) > x\right) \\ &\leq \limsup_{k \rightarrow \infty} 2P\left(\sum_{i=1}^{[kt]} \Delta(i, k, b) > x - \sqrt{2ktV(b)/k}\right) \\ &= 2P(Y_t^{(b)} > x - \sqrt{2tV(b)}), \end{aligned}$$

where we note that $(\Delta(i, k, b) := Y_{i/k}^{(b)} - Y_{(i-1)/k}^{(b)})_{i \geq 1}$ are i.i.d., each with expectation 0 and variance equal to $V(b)/k$. Given $\varepsilon > 0$, replace t by r^n , b by $r^{n/2}$, and x by $\lambda r^{n/2} + \sqrt{2r^n V(r^{n/2})}$, which is not larger than $(\lambda + \varepsilon)r^{n/2}$, once n is large enough, in (4.14). The convergence of the series in (4.12) then gives

$$\sum_{n \geq 0} P\left(\sup_{0 < s \leq r^n} Y_s^{(r^{n/2})} > (\lambda + \varepsilon)r^{n/2}\right) < \infty \quad \text{for all } \varepsilon > 0.$$

Hence by the Borel–Cantelli lemma

$$(4.15) \quad \limsup_{n \rightarrow \infty} \frac{\sup_{0 < t \leq r^n} Y_t^{(r^{n/2})}}{r^{n/2}} \leq \lambda \quad \text{a.s.}$$

Using (4.9), together with Lemma 4.2 and (4.15), gives

$$\limsup_{n \rightarrow \infty} \frac{\sup_{0 < t \leq r^n} X_t}{r^{n/2}} \leq \lambda \quad \text{a.s.}$$

By an argument of monotonicity, this yields

$$\limsup_{t \downarrow 0} \frac{X_t}{\sqrt{t}} \leq \frac{\lambda}{\sqrt{r}} \quad \text{a.s.}$$

Then let $r \uparrow 1$ to get $\limsup_{t \downarrow 0} X_t/\sqrt{t} \leq \lambda$ a.s.

For a reverse inequality, we show that for every $\lambda > 0$,

$$(4.16) \quad I(\lambda) = \infty \implies \limsup_{t \downarrow 0} \frac{X_t}{\sqrt{t}} \geq \lambda \quad \text{a.s.}$$

To see this, suppose that $I(\lambda) = \infty$ for a given $\lambda > 0$. Then the integral in (4.12) diverges when λ is replaced by $\lambda - \varepsilon$ for an arbitrarily small $\varepsilon > 0$, because $V(x) \geq \varepsilon \exp(-\varepsilon/2V(x))/2$, for $\varepsilon > 0$, $x > 0$. Hence, keeping in mind (4.9), Lemma 4.2 and Proposition 4.3, we deduce

$$(4.17) \quad \sum_{n \geq 0} P(X_{r^n} > \lambda' r^{n/2}) = \infty$$

for all $\lambda' < \lambda$. For a given $\varepsilon > 0$, define for every integer $n \geq 0$ the events

$$A_n = \{X_{r^n/(1-r)} - X_{r^{n+1}/(1-r)} > \lambda' r^{n/2}\},$$

$$B_n = \{|X_{r^{n+1}/(1-r)}| \leq \varepsilon r^{n/2}\}.$$

Then the $\{A_n\}_{n \geq 0}$ are independent, and each B_n is independent of the collection $\{A_n, A_{n-1}, \dots, A_0\}$. Further, $\sum_{n \geq 0} P(A_n) = \infty$ by (4.17), so $P(A_n \text{ i.o.}) = 1$. It can be deduced easily from [1], Proposition 2(i), page 16, that $X_t/\sqrt{t} \xrightarrow{P} 0$, as $t \downarrow 0$, since (1.7) is enforced. Thus $P(B_n) \rightarrow 1$ as $n \rightarrow \infty$, and then, by the

Feller–Chung lemma ([4], page 69) and Kolmogorov’s 0–1 law, we can deduce that $P(A_n \cap B_n \text{ i.o.}) = 1$. This implies $P(X_{r^n/(1-r)} > (\lambda' - \varepsilon)r^{n/2} \text{ i.o.}) = 1$, thus

$$\limsup_{t \downarrow 0} \frac{X_t}{\sqrt{t}} \geq (\lambda' - \varepsilon)\sqrt{1-r} \quad \text{a.s.,}$$

in which we can let $\lambda' \uparrow \lambda$, $\varepsilon \downarrow 0$ and $r \downarrow 0$ to get (4.16). Together with (4.13) and (4.16), this gives the statements in Theorem 2.2 [replace X by $-X$ to deduce the lim inf statements from the lim sup, noting that this leaves $V(\cdot)$ unchanged]. □

5. Proofs for Section 3.

5.1. *Proof of Theorem 3.1.* We start with some notation and technical results. Recall we assume (1.7).

Take $0 < \kappa < 1$ and suppose first

$$(5.1) \quad \int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx = \infty.$$

Define, for $0 < x < 1$,

$$\rho_\kappa(x) = \frac{1}{\kappa} \int_x^1 y^{1/\kappa-1} \bar{\Pi}^{(+)}(y) dy = \int_{x^{1/\kappa}}^1 \bar{\Pi}^{(+)}(y^\kappa) dy.$$

Since $\bar{\Pi}^{(+)}(x) > 0$ for all small x , $\rho_\kappa(x)$ is strictly decreasing in a neighborhood of 0, thus $x^{-1/\kappa} \rho_\kappa(x)$ is also strictly decreasing in a neighborhood of 0, and tends to ∞ as $x \downarrow 0$. Also define

$$U_+(x) = 2 \int_0^x y \bar{\Pi}^{(+)}(y) dy = x^2 \bar{\Pi}^{(+)}(x) + \int_0^x y^2 \Pi(dy).$$

Differentiation shows that $x^{-2}U_+(x)$ is strictly decreasing in a neighborhood of 0.

Next, given $\alpha \in (0, \kappa)$, define, for $t > 0$,

$$c(t) = \inf\{x > 0: \rho_\kappa(x)x^{-1/\kappa} + x^{-2}U_+(x) \leq \alpha/t\}.$$

Then $0 < c(t) < \infty$ for $t > 0$, $c(t)$ is strictly increasing, $\lim_{t \downarrow 0} c(t) = 0$, and

$$(5.2) \quad \frac{t\rho_\kappa(c(t))}{c^{1/\kappa}(t)} + \frac{tU_+(c(t))}{c^2(t)} = \alpha.$$

Since $\lim_{t \downarrow 0} \rho_\kappa(t) = \infty$, we have $\lim_{t \downarrow 0} c(t)/t^\kappa = \infty$.

We now point out that, when $1/2 < \kappa < 1$, (5.1) can be reinforced as follows.

LEMMA 5.1. *Keep $1/2 < \kappa < 1$. Then (5.1) implies that*

$$\int_0^1 \bar{\Pi}^{(+)}(c(t)) dt = \infty.$$

PROOF. We will first establish

$$(5.3) \quad \int_0^1 \frac{\Pi(dx)}{x^{-1/\kappa} \rho_\kappa(x) + x^{-2} U_+(x)} = \infty.$$

Suppose (5.3) fails. Since

$$f(x) := \frac{1}{x^{-1/\kappa} \rho_\kappa(x) + x^{-2} U_+(x)}$$

is nondecreasing (in fact, strictly increasing in a neighborhood of 0) with $f(0) = 0$, for every $\varepsilon > 0$ there is an $\eta > 0$ such that, for all $0 < x < \eta$,

$$\varepsilon \geq \int_x^\eta f(z) \Pi(dz) \geq f(x) (\overline{\Pi}^{(+)}(x) - \overline{\Pi}^{(+)}(\eta)),$$

giving

$$f(x) \overline{\Pi}^{(+)}(x) \leq \varepsilon + f(x) \overline{\Pi}^{(+)}(\eta) = \varepsilon + o(1) \quad \text{as } x \downarrow 0.$$

Letting $\varepsilon \downarrow 0$ shows that

$$(5.4) \quad \lim_{x \downarrow 0} f(x) \overline{\Pi}^{(+)}(x) = \lim_{x \downarrow 0} \left(\frac{\overline{\Pi}^{(+)}(x)}{x^{-1/\kappa} \rho_\kappa(x) + x^{-2} U_+(x)} \right) = 0.$$

It can be proved as in Lemma 4 of [6] that this implies

$$(5.5) \quad \lim_{x \downarrow 0} \left(\frac{x^{-1/\kappa} \rho_\kappa(x)}{x^{-2} U_+(x)} \right) = 0 \quad \text{or} \quad \liminf_{x \downarrow 0} \left(\frac{x^{-1/\kappa} \rho_\kappa(x)}{x^{-2} U_+(x)} \right) > 0.$$

(A detailed proof can be obtained from the third author.) Then, since (5.3) has been assumed not to hold,

$$(5.6) \quad \int_0^{1/2} \frac{x^{1/\kappa} \Pi(dx)}{\rho_\kappa(x)} < \infty \quad \text{or} \quad \int_0^{1/2} \frac{x^2 \Pi(dx)}{U_+(x)} < \infty.$$

Recalling that $\overline{\Pi}^{(+)}(1) = 0$, we have

$$\rho_\kappa(x) = -x^{1/\kappa} \overline{\Pi}^{(+)}(x) + \int_x^1 y^{1/\kappa} \Pi(dy) \leq \int_x^1 y^{1/\kappa} \Pi(dy),$$

so the first relation in (5.6) would imply the finiteness of

$$\int_0^{1/2} x^{1/\kappa} \left(\int_x^1 y^{1/\kappa} \Pi(dy) \right)^{-1} \Pi(dx);$$

but this is infinite by (5.1) and the Abel–Dini theorem: see, for example, [17], page 404. Thus the first relation in (5.6) is impossible. In a similar way, the second relation in (5.6) can be shown to be impossible. Thus (5.3) is proved.

Then note that the inverse function c^{\leftarrow} of c exists and satisfies, by (5.2),

$$c^{\leftarrow}(x) = \frac{\alpha}{x^{-1/\kappa} \rho_\kappa(x) + x^{-2} U_+(x)}.$$

Thus, by (5.3),

$$(5.7) \quad \int_0^1 c^{\leftarrow}(x)\Pi(dx) = \infty = \int_0^{c^{\leftarrow}(1)} \overline{\Pi}^{(+)}(c(x)) dx.$$

This proves our claim. \square

PROPOSITION 5.1. *For every $1/2 < \kappa < 1$, (5.1) implies*

$$\limsup_{t \downarrow 0} \frac{X_t}{t^\kappa} = \infty \quad a.s.$$

PROOF. The argument relies on the analysis of the completely asymmetric case when the Lévy measure Π has support in $[0, 1]$ or in $[-1, 0]$. Since $\kappa < 1$, we can assume $\gamma = 0$ without loss of generality, because of course $\gamma t = o(t^\kappa)$. The Lévy–Itô decomposition (1.9) then yields

$$(5.8) \quad X_t = \widehat{X}_t + \widetilde{X}_t$$

with

$$(5.9) \quad \widehat{X}_t = \int_{[0,t] \times [0,1]} xN(ds, dx) \quad \text{and} \quad \widetilde{X}_t = \int_{[0,t] \times [-1,0]} xN(ds, dx),$$

where, as usual, the Poissonian stochastic integrals are taken in the compensated sense.

Choose α so small that

$$(1 + \kappa)\alpha < 1/2 \quad \text{and} \quad \alpha/(1/2 - (1 + \kappa)\alpha)^2 \leq 1/2,$$

and then ε so small that $c(\varepsilon) < 1$. Observe that, for every $0 < t < \varepsilon$,

$$(5.10) \quad \begin{aligned} t \int_{\{c(t) < x \leq 1\}} x\Pi(dx) &= tc(t)\overline{\Pi}^{(+)}(c(t)) + t\lambda(c(t)) \\ &\leq \frac{tU_+(c(t))}{c(t)} + t\lambda(c(t)) \\ &\leq \alpha(1 + \kappa)c(t), \end{aligned}$$

where the last inequality stems from (5.2) and

$$\lambda(x) := \int_x^1 \overline{\Pi}^{(+)}(y) dy = \int_x^1 y^{1-1/\kappa} y^{1/\kappa-1} \overline{\Pi}^{(+)}(y) dy \leq \kappa x^{1-1/\kappa} \rho_\kappa(x)$$

(since $\kappa < 1$), so

$$\frac{t\lambda(c(t))}{c(t)} \leq \frac{\kappa t \rho_\kappa(c(t))}{c^{1/\kappa}(t)} \leq \kappa\alpha \quad [\text{by (5.2)}].$$

We next deduce from Lemma 5.1 that for every $\varepsilon > 0$, the Poisson random measure N has infinitely many atoms in the domain $\{(t, x) : 0 \leq t < \varepsilon \text{ and } x > c(t)\}$, a.s. Introduce

$$t_\varepsilon := \sup\{t \leq \varepsilon : N(\{t\} \times (c(t), 1]) = 1\},$$

the largest instant less than ε of such an atom. Our goal is to check that

$$(5.11) \quad P(X_{t_\varepsilon-} \geq -c(t_\varepsilon)/2) \geq 1/33$$

for every $\varepsilon > 0$ sufficiently small, so that $P(X_{t_\varepsilon} > c(t_\varepsilon)/2) > 1/33$. Since $t^\kappa = o(c(t))$, it follows that for every $a > 0$

$$P(\exists t \leq \varepsilon : X_t > at^\kappa) \geq 1/33,$$

and hence $\limsup_{t \downarrow 0} X_t/t^\kappa = \infty$ with probability at least $1/33$. The proof is then completed by an appeal to Blumenthal's 0–1 law.

In order to establish (5.11), we will work henceforth conditionally on t_ε ; recall from the Markov property of Poisson random measures that the restriction of $N(dt, dx)$ to $[0, t_\varepsilon] \times [-1, 1]$ is still a Poisson random measure with intensity $dt \Pi(dx)$.

Recalling (5.10) and discarding the jumps $\widehat{\Delta}$ of \widehat{X} such that $\widehat{\Delta}_s > c(t_\varepsilon)$ for $0 \leq s < t_\varepsilon$ in the stochastic integral (5.9), we obtain the inequality

$$(5.12) \quad X_{t_\varepsilon-} \geq \widehat{Y}_{t_\varepsilon-} - \alpha(1 + \kappa)c(t_\varepsilon) + \widetilde{X}_{t_\varepsilon-},$$

where $\widehat{Y}_{t_\varepsilon-}$ is given by the (compensated) Poissonian integral

$$\widehat{Y}_{t_\varepsilon-} := \int_{[0, t_\varepsilon] \times [0, c(t_\varepsilon)]} x N(ds, dx).$$

We chose $(1 + \kappa)\alpha < 1/2$, so by Chebyshev's inequality we have

$$\begin{aligned} P(\widehat{Y}_{t_\varepsilon-} \leq \alpha(1 + \kappa)c(t_\varepsilon) - c(t_\varepsilon)/2) &\leq P(|\widehat{Y}_{t_\varepsilon-}| > (1/2 - \alpha(1 + \kappa))c(t_\varepsilon)) \\ &\leq \frac{E|\widehat{Y}_{t_\varepsilon-}|^2}{(1/2 - \alpha(1 + \kappa))^2 c^2(t_\varepsilon)} \\ &\leq \frac{t_\varepsilon \int_{\{0 < x \leq c(t_\varepsilon)\}} x^2 \Pi(dx)}{(1/2 - \alpha(1 + \kappa))^2 c^2(t_\varepsilon)} \\ &\leq \frac{t_\varepsilon U_+(c(t_\varepsilon))}{(1/2 - \alpha(1 + \kappa))^2 c^2(t_\varepsilon)} \\ &\leq \frac{\alpha}{(1/2 - \alpha(1 + \kappa))^2}, \end{aligned}$$

where the last inequality derives from (5.2). By choice of α , the final expression does not exceed $1/2$. We conclude that

$$(5.13) \quad P(\widehat{Y}_{t_\varepsilon-} - \alpha(1 + \kappa)c(t_\varepsilon) \geq -c(t_\varepsilon)/2) \geq 1/2.$$

We will also use the fact that \widetilde{X} is a mean-zero Lévy process which is spectrally negative (i.e., with no positive jumps), so

$$\liminf_{t \downarrow 0} P(\widetilde{X}_t > 0) \geq 1/16;$$

see [9], Volume 2, Lemma IV.3.1, page 470. Furthermore, as \tilde{X} is independent of $\widehat{Y}_{t_\varepsilon-}$, we conclude from (5.12) and (5.13) that (5.11) holds provided that ε has been chosen small enough. \square

Now suppose (5.1) fails. The remaining results in Theorem 3.1 require the case $\kappa > 1/2$ of:

PROPOSITION 5.2. *Assume that Y is a spectrally negative Lévy process, has zero mean, and is not of bounded variation. Define, for $y > 0, \lambda > 0$,*

$$W_Y(y) := \int_0^y \int_x^1 z \Pi_Y^{(-)}(dz) dx$$

and

$$(5.14) \quad J_Y(\lambda) := \int_0^1 \exp\left\{-\lambda \left(\frac{y^{(2\kappa-1)/\kappa}}{W_Y(y)}\right)^{\kappa/(1-\kappa)}\right\} \frac{dy}{y},$$

where $\Pi_Y^{(-)}$ is the canonical measure of $-Y$, assumed carried on $(0, 1]$, and let $\lambda_Y^* = \inf\{\lambda > 0 : J_Y(\lambda) < \infty\}$. Then with probability 1, for $1/2 \leq \kappa < 1$,

$$\limsup_{t \downarrow 0} \frac{Y_t}{t^\kappa} \begin{cases} = \infty, \\ \in (0, \infty), \\ = 0, \end{cases} \quad \text{according as } \lambda_Y^* \begin{cases} = \infty, \\ \in (0, \infty), \\ = 0. \end{cases}$$

The proof of Proposition 5.2 requires several intermediate steps. Take Y as described; then it has characteristic exponent

$$\Psi_Y(\theta) = \int_{(0,1]} (e^{-i\theta x} - 1 + i\theta x) \Pi_Y^{(-)}(dx).$$

So we can work with the Laplace exponent

$$(5.15) \quad \psi_Y(\theta) = \Psi_Y(-i\theta) = \int_{(0,1]} (e^{-\theta x} - 1 + \theta x) \Pi_Y^{(-)}(dx),$$

such that $E e^{\theta Y_t} = e^{t\psi_Y(\theta)}$, $t \geq 0, \theta \geq 0$.

Let $T = (T_t, t \geq 0)$ denote the first-passage process of Y ; this is a subordinator whose Laplace exponent Φ is the inverse function to ψ_Y ([1], page 189), and since $Y(T_t) \equiv t$ we see that the alternatives in Proposition 5.2 can be deduced immediately from

$$\limsup_{t \downarrow 0} \frac{Y_t}{t^\kappa} \begin{cases} = \infty, \\ \in (0, \infty), \\ = 0, \end{cases} \iff \liminf_{t \downarrow 0} \frac{T_t}{t^{1/\kappa}} \begin{cases} = 0, \\ \in (0, \infty), \\ = \infty. \end{cases}$$

The subordinator T must have zero drift since if $\lim_{t \downarrow 0} T_t/t := c > 0$ a.s., then $\sup_{0 < s \leq T_t} Y_s = t$ (see [1], page 191) would give $\limsup_{t \downarrow 0} Y_t/t \leq 1/c < \infty$ a.s., thus $Y \in bv$, which is not the case. We can assume T has no jumps bigger than 1,

and further exclude the trivial case when T is compound Poisson. So the main part of the proof of Proposition 5.2 is the following, which is a kind of analogue of Theorem 1 of Zhang [18].

LEMMA 5.2. *Let T be any subordinator with zero drift whose Lévy measure Π_T is carried by $(0, 1]$ and has $\bar{\Pi}_T(0+) = \infty$, where $\bar{\Pi}_T(x) = \Pi_T\{(x, \infty)\}$ for $x > 0$. Put $m_T(x) = \int_0^x \bar{\Pi}_T(y) dy$ and for $d > 0$ let*

$$(5.16) \quad K_T(d) := \int_0^1 \frac{dy}{y} \exp\left\{-d \frac{(m_T(y))^{\gamma/(\gamma-1)}}{y}\right\} \quad \text{where } \gamma > 1.$$

Let $d_K^* := \inf\{d > 0 : K_T(d) < \infty\} \in [0, \infty]$. Then, with probability 1:

- (i) $d_K^* = \infty$ iff $\liminf_{t \downarrow 0} \frac{T_t}{t^\gamma} = 0$;
- (ii) $d_K^* = 0$ iff $\lim_{t \downarrow 0} \frac{T_t}{t^\gamma} = \infty$;
- (iii) $d_K^* \in (0, \infty)$ iff $\liminf_{t \downarrow 0} \frac{T_t}{t^\gamma} = c$, for some $c \in (0, \infty)$.

Before beginning the proof of Lemma 5.2, we need some preliminary results. To start with, we need the following lemma.

LEMMA 5.3. *Let $S = (S_t, t \geq 0)$ be a subordinator, and let a and γ be positive constants. Suppose the series*

$$(5.17) \quad \sum_{n \geq 1} P(S_{r^n} \leq ar^{n\gamma})$$

diverges for some $r \in (0, 1)$. Then

$$(5.18) \quad \liminf_{t \downarrow 0} \frac{S_t}{t^\gamma} \leq \frac{a}{(1-r)^\gamma} \quad \text{a.s.}$$

Conversely, if the series in (5.17) converges for some $r \in (0, 1)$, then

$$\liminf_{t \downarrow 0} \frac{S_t}{t^\gamma} > ar^\gamma \quad \text{a.s.}$$

PROOF. Suppose the series in (5.17) diverges for an $r \in (0, 1)$. Fix $m \geq 1$, and for $n > m$ define events

$$A_n = \{S_{r^n} \leq ar^{n\gamma}, S_{r^k} > ar^{k\gamma} \text{ for } m \leq k < n\}.$$

Then the $\{A_n\}_{n > m}$ are disjoint, and

$$\begin{aligned} P(A_n) &\geq P(S_{r^n} \leq ar^{n\gamma}, S_{r^k} - S_{r^n} > ar^{k\gamma} \text{ for } m \leq k < n) \\ &= P(S_{r^n} \leq ar^{n\gamma}) P\left(\frac{S_{r^k-r^n}}{(r^k-r^n)^\gamma} > a \frac{r^{k\gamma}}{(r^k-r^n)^\gamma} \text{ for } m \leq k < n\right) \\ &\geq P(S_{r^n} \leq ar^{n\gamma}) P\left(\frac{S_{r^k-r^n}}{(r^k-r^n)^\gamma} > \frac{a}{(1-r)^\gamma} \text{ for } m \leq k < n\right). \end{aligned}$$

Suppose now that $\liminf S_t/t^\gamma > a/(1-r)^\gamma$ a.s. Then for some m_0 we will have

$$P(S_t/t^\gamma > a/(1-r)^\gamma \text{ for all } t \in (0, r^{m_0})) \geq 1/2.$$

[If this were not the case, we would have $P(B_n) \geq 1/2$ for all $n \geq 1$, where B_n denotes the event $\{\exists t \in (0, r^n) \text{ with } S_t/t^\gamma \leq a/(1-r)^\gamma\}$; this would imply $P(\bigcap_n B_n) \geq 1/2$, a contradiction.] Thus for $n > m_0$ we have

$$P\left(\frac{S_{r^k-r^n}}{(r^k-r^n)^\gamma} > \frac{a}{(1-r)^\gamma} \text{ for } m_0 \leq k < n\right) \geq 1/2.$$

Then

$$\begin{aligned} 1 &\geq \sum_{n>m_0} P(A_n) \\ &\geq \sum_{n>m_0} P(S_{r^n} \leq ar^{n\gamma}) P\left(\frac{S_{r^k-r^n}}{(r^k-r^n)^\gamma} > \frac{a}{(1-r)^\gamma} \text{ for } m_0 \leq k < n\right) \\ &\geq \frac{1}{2} \sum_{n>m_0} P(S_{r^n} \leq ar^{n\gamma}), \end{aligned}$$

which is impossible since we assumed the series in (5.17) diverges. This gives (5.18), and the converse follows easily from the Borel–Cantelli lemma and an argument of monotonicity. \square

Applying Lemma 5.3 to T_t , we see that the alternatives in (i)–(iii) of Lemma 5.2 hold iff for some $r < 1$, for all, none, or some but not all, $a > 0$,

$$(5.19) \quad \sum_{n \geq 1} P(T_{r^n} \leq ar^{n\gamma}) < \infty.$$

The next step is to get bounds for the probability in (5.19). One way is easy. Since $\bar{\Pi}_T(0+) = \infty$, $\bar{\Pi}_T(x)$ is strictly positive, and thus $m_T(x)$ is strictly increasing, on a neighborhood of 0. We write $h(\cdot)$ for the inverse function to $m_T(\cdot)$.

LEMMA 5.4. *Let T be a subordinator with zero drift whose Lévy measure Π_T has support on $[0, 1]$ and satisfies $\bar{\Pi}_T(0+) = \infty$. Then there is an absolute constant K such that, for any $c > 0$ and $\gamma > 0$,*

$$(5.20) \quad P(T_t \leq ct^\gamma) \leq \exp\left\{-\frac{ct^\gamma}{h(2ct^{\gamma-1}/K)}\right\}, \quad t > 0.$$

PROOF. We can write

$$\Phi(\lambda) = -\frac{1}{t} \log Ee^{-\lambda T_t} = \int_{(0,1]} (1 - e^{-\lambda x}) \Pi_T(dx), \quad \lambda > 0.$$

Markov’s inequality gives, for any $\lambda > 0, c > 0,$

$$\begin{aligned} P(T_t \leq ct^\gamma) &\leq e^{\lambda ct^\gamma} E(e^{-\lambda T_t}) \\ &= \exp\{-\lambda t(\lambda^{-1}\Phi(\lambda) - ct^{\gamma-1})\} \\ &\leq \exp\{-\lambda t(Km_T(1/\lambda) - ct^{\gamma-1})\} \quad \text{for some } K > 0, \end{aligned}$$

where we have used [1], Proposition 1, page 74. Now choose $\lambda = 1/h(2ct^{\gamma-1}/K)$ and we have (5.20). \square

The corresponding lower bound is trickier:

LEMMA 5.5. *Suppose that T is as in Lemma 5.2, and additionally satisfies $\lim_{t \downarrow 0} P(T_t \leq Ct^\gamma) = 0$ for some $C > 0$ and $\gamma > 1$. Then, for sufficiently small $t > 0,$*

$$(5.21) \quad P(T_t \leq ct^\gamma) \geq \frac{1}{4} \exp\left\{-\frac{ct^\gamma}{h(ct^{\gamma-1}/4)}\right\},$$

where $c = bC,$ and $b \in (0, \infty)$ is an absolute constant.

PROOF. Take $\gamma > 1$ and assume $\lim_{t \downarrow 0} P(T_t \leq Ct^\gamma) = 0,$ where $C > 0.$ Write, for each fixed $t > 0,$ $T_t = T_t^{(1)} + T_t^{(2)},$ where the distributions of the independent random variables $T_t^{(1)}$ and $T_t^{(2)}$ are specified by

$$\log Ee^{-\lambda T_t^{(1)}} = -t \int_{(0, h(Ct^{\gamma-1}/4)]} (1 - e^{-\lambda x}) \Pi_T(dx)$$

and

$$\log Ee^{-\lambda T_t^{(2)}} = -t \int_{(h(Ct^{\gamma-1}/4), 1]} (1 - e^{-\lambda x}) \Pi_T(dx).$$

Observe that

$$ET_t^{(1)} = t \int_{(0, h(Ct^{\gamma-1}/4)]} x \Pi_T(dx) \leq tm_T(h(Ct^{\gamma-1}/4)) = Ct^\gamma/4,$$

so that

$$\begin{aligned} P(T_t > Ct^\gamma) &\leq P(T_t^{(1)} > Ct^\gamma) + P(T_t^{(2)} \neq 0) \\ &\leq \frac{ET_t^{(1)}}{Ct^\gamma} + 1 - P(T_t^{(2)} = 0) \\ &\leq 5/4 - P(T_t^{(2)} = 0). \end{aligned}$$

Thus for all sufficiently small $t, t \leq t_0(C),$ say,

$$P(T_t^{(2)} = 0) \leq 1/4 + P(T_t \leq Ct^\gamma) \leq 1/2,$$

because of our assumption that $\lim_{t \downarrow 0} P(T_t \leq Ct^\gamma) = 0$. Now

$$P(T_t^{(2)} = 0) = \exp(-t\bar{\Pi}_T(h(Ct^{\gamma-1}/4))),$$

so we get, for $t \leq t_0(C)$,

$$\begin{aligned} \log 2 &\leq t\bar{\Pi}_T(h(Ct^{\gamma-1}/4)) \\ &\leq \frac{t}{h(Ct^{\gamma-1}/4)}m_T(h(Ct^{\gamma-1}/4)) \\ &= \frac{Ct^\gamma/4}{h(Ct^{\gamma-1}/4)}. \end{aligned}$$

Thus, replacing t by $(C/4)^{1/(\gamma-1)}t$, we have $h(t^{\gamma-1}) \leq at^\gamma$, for all $t \leq t_1(C) := t_0(C)(4/C)^{1/(\gamma-1)}$, where $a = (4/C)^{1/(\gamma-1)}/\log 2$.

Now keep $0 < t \leq t_1$, and let $c > 0$ satisfy $3ac^{1/(\gamma-1)}A = 1/4$, where A is the constant in (4.11); thus, $c = C(\log 2)^{\gamma-1}/((3A)^{\gamma-1}4^\gamma) = bC$, say, where b is an absolute constant. Write $\eta = \eta(t) = h(ct^{\gamma-1}/4)$. Define processes $(Y_t^{(i)})_{t \geq 0}$, $i = 1, 2, 3$, such that $(T_t)_{t \geq 0}$ and $(Y_t^{(1)})_{t \geq 0}$ are independent, and $(Y_t^{(2)})_{t \geq 0}$ and $(Y_t^{(3)})_{t \geq 0}$ are independent, and are such that $\log E(e^{-\lambda Y_t^{(i)}}) = -t \int_{(0,1]} (1 - e^{-\lambda x}) \Pi_Y^{(i)}(dx)$, $i = 1, 2, 3$, where

$$\begin{aligned} \Pi_Y^{(1)}(dx) &= \bar{\Pi}_T(\eta)\delta_\eta(dx), \\ \Pi_Y^{(2)}(dx) &= \bar{\Pi}_T(\eta)\delta_\eta(dx) + \mathbf{1}_{(0,\eta]} \Pi_T(dx), \\ \Pi_Y^{(3)}(dx) &= \mathbf{1}_{(\eta,1]} \Pi_T(dx), \end{aligned}$$

and $\delta_\eta(dx)$ is the point mass at η . Then we have $T_t + Y_t^{(1)} \stackrel{d}{=} Y_t^{(2)} + Y_t^{(3)}$, and

$$\begin{aligned} P(T_t \leq ct^\gamma) &\geq P(T_t + Y_t^{(1)} \leq ct^\gamma) \\ &\geq P(Y_t^{(3)} = 0)P(Y_t^{(2)} \leq ct^\gamma) \\ &= e^{-t\bar{\Pi}_Y^{(3)}(\eta)}P(Y_t^{(2)} \leq ct^\gamma). \end{aligned}$$

Since $t\bar{\Pi}_Y^{(3)}(\eta) = \eta^{-1}t\eta\bar{\Pi}_T(\eta) \leq \eta^{-1}tm_T(\eta) \leq ct^\gamma/(4h(ct^{\gamma-1}/4))$, (5.21) will follow when we show that $\liminf_{t \downarrow 0} P(Y_t^{(2)} \leq ct^\gamma) \geq 1/4$. By construction we have

$$\begin{aligned} EY_t^{(2)} &= t\left(\int_0^\eta x\Pi_T(dx) + \eta\bar{\Pi}_T(\eta)\right) \\ &= tm_T(\eta) = \frac{ct^\gamma}{4}, \end{aligned}$$

so if we put $Z_t = Y_t^{(2)} - ct^\gamma/4$ and write $t\sigma_t^2 = EZ_t^2$ we can apply Chebyshev's inequality to get

$$P(Y_t^{(2)} \leq ct^\gamma) \geq P\left(Z_t \leq \frac{ct^\gamma}{2}\right) \geq \frac{5}{9}$$

when $t\sigma_t^2 \leq c^2t^{2\gamma}/9$. To deal with the opposite case, $t\sigma_t^2 > c^2t^{2\gamma}/9$, we use the Normal approximation in Lemma 4.3. In the notation of that lemma, $m_3 = \int_{|x| \leq \eta} |x|^3 \Pi_Y^{(2)}(dx)$ and $m_2 = \sigma_t^2$, in the present situation. Since, then, $m_3 \leq \eta\sigma_t^2$, and $\eta = h(ct^{\gamma-1}/4)$, we get, for $t \leq t_1(C)/(4/C)^{1/(\gamma-1)} := t_2(C)$,

$$\begin{aligned} \sup_{x \in \mathbb{R}} |P(Z_t \leq x\sqrt{t}\sigma_t) - 1 + \bar{F}(x)| &\leq \frac{A\eta\sigma_t^2}{\sqrt{t}\sigma_t^3} \leq \frac{3Ah(ct^{\gamma-1}/4)}{ct^\gamma} \\ &\leq 3ac^{1/(\gamma-1)}A = 1/4. \end{aligned}$$

The last equality follows by our choice of c . Finally, taking $x = ct^\gamma/(2\sqrt{t}\sigma_t)$ gives $P(Z_t \leq ct^\gamma/2) \geq 1/4$, hence (5.21). \square

We are now able to establish Lemma 5.2. Recall that $h(\cdot)$ is inverse to $m_T(\cdot)$.

PROOF OF LEMMA 5.2. (i) Suppose that $K_T(d) < \infty$ for some $d > 0$ and write $x_n = h(cr^{n(\gamma-1)})$, where $c > 0$, $\gamma > 1$ and $0 < r < 1$. Note that since $h(x)/x$ is increasing we have $x_{n+1} \leq r^{\gamma-1}x_n$. Also we have $m_T(x_n) = Rm_T(x_{n+1})$ where $R = r^{1-\gamma} > 1$. So for $y \in [x_{n+1}, x_n]$,

$$\begin{aligned} \frac{m_T(y)^{\gamma/(\gamma-1)}}{y} &\leq \frac{m_T(x_n)^{\gamma/(\gamma-1)}}{x_{n+1}} \\ &= \frac{R^{\gamma/(\gamma-1)}m_T(x_{n+1})^{\gamma/(\gamma-1)}}{x_{n+1}} \\ &= \frac{(Rc)^{\gamma/(\gamma-1)}r^{(n+1)\gamma}}{h(cr^{(n+1)(\gamma-1)})}. \end{aligned}$$

Thus

$$\begin{aligned} \int_{x_{n+1}}^{x_n} \frac{dy}{y} \exp\left\{-d \frac{m_T(y)^{\gamma/(\gamma-1)}}{y}\right\} &\geq \exp\left\{-\frac{d(Rc)^{\gamma/(\gamma-1)}r^{(n+1)\gamma}}{h(cr^{(n+1)(\gamma-1)})}\right\} \log \frac{x_n}{x_{n+1}} \\ &\geq (\log R) \exp\left\{-\frac{c'r^{(n+1)\gamma}}{h(2c'r^{(n+1)(\gamma-1)}/K)}\right\}, \end{aligned}$$

where K is the constant in Lemma 5.4 and we have chosen

$$c = \left(\frac{K}{2d}\right)^{\gamma-1} R^{-\gamma} \quad \text{and} \quad c' = Kc/2.$$

Then $K_T(d) < \infty$ gives $\sum_1^\infty P(T_{r^n} \leq c'r^{n\gamma}) < \infty$, and so $\liminf_{t \downarrow 0} T_t/t^\gamma \geq c'r^\kappa > 0$ a.s., by Lemma 5.3. Thus we see that $\liminf_{t \downarrow 0} T_t/t^\gamma = 0$ a.s. implies that $K_T(d) = \infty$ for every $d > 0$, hence $d_K^* = \infty$.

Conversely, assume that $\liminf_{t \downarrow 0} T_t/t^\gamma > C > 0$ a.s. Then by Lemma 5.3, $\sum_1^\infty P(T_{r^n} \leq (1-r)^\gamma Cr^{n\gamma}) < \infty$ for some $0 < r < 1$. Then $P(T_{r^n} \leq (1-r)^\gamma Cr^{n\gamma}) \rightarrow 0$, Lemma 5.5 applies, and we have

$$\sum_1^\infty \exp\left\{-\frac{cr^{n\gamma}}{h(cr^{n(\gamma-1)}/4)}\right\} < \infty,$$

where $c = (1-r)^\gamma bC$, $b > 0$. Putting $x_n = h(cr^{n(\gamma-1)}/4)$ (similar to but not the same x_n as in the previous paragraph), and $c' = 4^{\gamma/(\gamma-1)}/c^{1/(\gamma-1)}$, we see that

$$(5.22) \quad \sum_1^\infty \exp\left\{-\frac{c'm_T(x_n)^{\gamma/(\gamma-1)}}{x_n}\right\} < \infty.$$

We have $m_T(x_{n-1}) = Rm_T(x_n)$ where $R = r^{1-\gamma} > 1$. Take $L > R$ and let $k_n = \min(k \geq 1 : x_{n-1}L^{-k} \leq x_n)$, so $x_{n-1}L^{-k_n} \leq x_n$. Then for any $d > 0$

$$\begin{aligned} & \int_{x_n}^{x_{n-1}} \exp\left\{-d\frac{m_T(y)^{\gamma/(\gamma-1)}}{y}\right\} y^{-1} dy \\ & \leq \sum_{i=1}^{k_n} \int_{x_{n-1}L^{-i}}^{x_{n-1}L^{1-i}} \exp\left\{-dm_T(y)^{1/(\gamma-1)}\frac{m_T(y)}{y}\right\} y^{-1} dy \\ & \leq \sum_{i=1}^{k_n} \int_{x_{n-1}L^{-i}}^{x_{n-1}L^{1-i}} \exp\left\{-dm_T(x_n)^{1/(\gamma-1)}\frac{m_T(x_{n-1}L^{-i})}{x_{n-1}L^{-i}}\right\} y^{-1} dy \\ & \leq \log L \sum_{i=1}^\infty \exp\left\{-dm_T(x_n)^{1/(\gamma-1)}\frac{L^i m_T(x_n L^{-1})}{x_{n-1}}\right\} \\ & \leq \log L \sum_{i=1}^\infty \exp\left\{-dm_T(x_n)^{1/(\gamma-1)}\frac{L^{i-1} m_T(x_n)}{x_{n-1}}\right\} \\ & = \log L \sum_{i=1}^\infty \exp\left\{-dL^{i-1} R^{-\gamma/(\gamma-1)}\frac{m_T(x_{n-1})^{\gamma/(\gamma-1)}}{x_{n-1}}\right\}. \end{aligned}$$

Choose $d = c'LR^{\gamma/(\gamma-1)} = (4R)^{\gamma/(\gamma-1)}L/((1-r)^\gamma bC)^{1/(\gamma-1)}$, to write this last sum as $\sum_i a_n^L$, where $a_n = \exp(-c'm_T(x_{n-1})^{\gamma/(\gamma-1)}/x_{n-1})$. It is then easy to see that for all large enough n , the n th term in the sum is bounded above by a constant multiple of a_n . It follows from (5.22) that $\sum_n a_n < \infty$, hence $K_T(d) < \infty$, and part (i) follows.

(ii) If $K_T(d) < \infty$ for all $d > 0$, then, because $c' \rightarrow \infty$ as $d \rightarrow 0$ at the end of the proof of the forward part of part (i), we have $\liminf_{t \downarrow 0} T_t/t^\gamma = \infty$ a.s., that is,

$\lim_{t \downarrow 0} T_t/t^\gamma = \infty$ a.s. Conversely, if this holds, then because $d \rightarrow 0$ as $C \rightarrow \infty$ at the end of the proof of the converse part of part (i), we have $K_T(d) < \infty$ for all $d > 0$.

(iii) Since we know from Kolmogorov’s 0–1 law that $\liminf_{t \downarrow 0} T_t/t^\gamma$ is a.s. constant, part (iii) follows from parts (i) and (ii).

This completes the proof of Lemma 5.2. \square

PROOF OF PROPOSITION 5.2. To finish the proof of Proposition 5.2, we need only show that $J_Y(\lambda) = \infty$ ($< \infty$) for all $\lambda > 0$ is equivalent to $K_T(d) = \infty$ ($< \infty$) for all $d > 0$, respectively, where $K_T(d)$ is evaluated for the first-passage process T of Y , and $\gamma = 1/\kappa$. We have from (5.15), after integrating by parts,

$$\begin{aligned} \psi_Y(\theta) &= \int_0^1 (e^{-\theta x} - 1 + \theta x) \Pi_Y^{(-)}(dx) \\ &= \theta \int_0^1 (1 - e^{-\theta x}) \overline{\Pi}_Y^{(-)}(x) dx, \end{aligned}$$

and differentiating (5.15) gives

$$\psi'_Y(\theta) = \int_0^1 x(1 - e^{-\theta x}) \Pi_Y^{(-)}(dx).$$

So we see that $\theta^{-1}\psi_Y(\theta)$ and $\psi'_Y(\theta)$ are Laplace exponents of driftless subordinators, and using the estimate in [1], page 74, twice, we get

$$\psi_Y(\theta) \asymp \theta^2 \widetilde{W}_Y(1/\theta) \quad \text{and} \quad \psi'_Y(\theta) \asymp \theta W_Y(1/\theta),$$

where $\widetilde{W}_Y(x) = \int_0^x A_Y(y) dy$ and $A_Y(x) := \int_x^1 \overline{\Pi}_Y^{(-)}(y) dy$, for $x > 0$. [Recall the definition of W_Y just prior to (5.14), and use “ \asymp ” to mean that the ratio of the quantities on each side of the symbol is bounded above and below by finite positive constants for all values of the argument.] However, putting $U_Y(x) = \int_0^x 2z \overline{\Pi}_Y^{(-)}(z) dz$, for $x > 0$, we see that

$$W_Y(x) = \int_0^x \int_y^1 z \Pi_Y^{(-)}(dz) dy = \frac{1}{2} U_Y(x) + \widetilde{W}_Y(x)$$

and

$$\widetilde{W}_Y(x) = \frac{1}{2} U_Y(x) + x A_Y(x);$$

thus

$$\widetilde{W}_Y(x) \leq W_Y(x) = U_Y(x) + x A_Y(x) \leq 2 \widetilde{W}_Y(x).$$

Hence we have

$$(5.23) \quad \theta^2 W_Y(1/\theta) \asymp \psi_Y(\theta) \asymp \theta \psi'_Y(\theta).$$

We deduce that $J_Y(\lambda) = \infty (< \infty)$ for all $\lambda > 0$ is equivalent to $\tilde{J}_Y(\lambda) = \infty (< \infty)$ for all $\lambda > 0$, respectively, where

$$\begin{aligned} \tilde{J}_Y(\lambda) &= \int_0^1 \exp\{-\lambda y^{-1/(1-\kappa)} \psi_Y(1/y)^{-\kappa/(1-\kappa)}\} \frac{dy}{y} \\ &= \int_1^\infty \exp\{-\lambda y^{1/(1-\kappa)} \psi_Y(y)^{-\kappa/(1-\kappa)}\} \frac{dy}{y}. \end{aligned}$$

But we know that Φ , the exponent of the first-passage process T , is the inverse of ψ_Y , so making the obvious change of variable gives

$$\tilde{J}_Y(\lambda) = \int_{\psi_Y(1)}^\infty \exp\{-\lambda \Phi(z)^{1/(1-\kappa)} z^{-\kappa/(1-\kappa)}\} \frac{\Phi'(z) dz}{\Phi(z)}.$$

From (5.23) we deduce that $z\Phi'(z)/\Phi(z) \asymp 1$ for all $z > 0$, so $J_Y(\lambda) = \infty (< \infty)$ for all $\lambda > 0$ is equivalent to $\hat{J}_Y(\lambda) = \infty (< \infty)$ for all $\lambda > 0$, respectively, where

$$\begin{aligned} \hat{J}_Y(\lambda) &= \int_1^\infty \exp\{-\lambda \Phi(z)^{1/(1-\kappa)} z^{-\kappa/(1-\kappa)}\} \frac{dz}{z}, \\ &= \int_0^1 \exp\{-\lambda \Phi(z^{-1})^{1/(1-\kappa)} z^{\kappa/(1-\kappa)}\} \frac{dz}{z}. \end{aligned}$$

Since $\Phi(z^{-1})$ is bounded above and below by multiples of $z^{-1}m_T(z)$ ([1], page 74), our claim is established. \square

We are now able to complete the proof of Theorem 3.1.

PROOF OF THEOREM 3.1. The implications (i) \Rightarrow (3.1) and (iii) \Rightarrow (3.3) stem from Proposition 5.1 and Theorem 2.1, respectively. So we can focus on the situation when

$$\int_0^1 \bar{\Pi}^{(+)}(x^\kappa) dx < \infty = \int_0^1 \bar{\Pi}^{(-)}(x^\kappa) dx.$$

Recall the decomposition (5.8) where \hat{X}_t has canonical measure $\Pi^{(+)}(dx)$. Thus from Theorem 2.1, \hat{X}_t is $o(t^\kappa)$ a.s. as $t \downarrow 0$. Further, \tilde{X} is spectrally negative with mean zero. When (ii), (iv) or (v) holds, $\tilde{X} \notin bv$ [see Remark 3(ii)]. The implications (ii) \Rightarrow (3.1), (iv) \Rightarrow (3.3) and (v) \Rightarrow (3.4) now follow from Proposition 5.2 applied to \tilde{X} . \square

REMARK 6. Concerning Remark 3(iii): perusal of the proof of Theorem 2.2 shows that we can add to X a compound Poisson process with masses $f_\pm(t)$, say, at $\pm\sqrt{t}$, provided $\sum_{n \geq 1} \sqrt{t_n} f_\pm(t_n)$ converges, and the proof remains valid. The effect of this is essentially only to change the kind of truncation that is being applied, without changing the value of the lim sup, and in the final result this shows up only in an alteration to $V(x)$. Choosing $f(t)$ appropriately, the new $V(\cdot)$ becomes

$U(\cdot)$ or $W(\cdot)$, which are thus equivalent in the context of Theorem 2.2. Note that we allow $\kappa = 1/2$ in Proposition 5.2. We will omit further details, but the above shows there is no contradiction with Theorem 2.2.

5.2. *Proof of Theorem 3.2.* Theorem 3.2 follows by taking $a(x) = x^\kappa$, $\kappa > 1$, in Propositions 5.3 and 5.4 below, which are a kind of generalization of Theorem 9 in Chapter III of [1]. Recall the definition of $A_-(\cdot)$ in (3.5).

PROPOSITION 5.3. *Assume $X \in bv$ and $\delta = 0$. Suppose $a(x)$ is a positive nonstochastic measurable function on $[0, \infty)$ with $a(x)/x$ nondecreasing and $a(0) = 0$. Let $a^\leftarrow(x)$ be its inverse function. Suppose*

$$(5.24) \quad \int_0^1 \frac{\Pi(dx)}{1/a^\leftarrow(x) + A_-(x)/x} = \infty.$$

Then we have

$$(5.25) \quad \limsup_{t \downarrow 0} \frac{X_t}{a(t)} = \infty \quad a.s.$$

PROOF. Assume X and a as specified. Then the function $a(x)$ is strictly increasing, so $a^\leftarrow(x)$ is well defined, positive, continuous and nondecreasing on $[0, \infty)$, with $a^\leftarrow(0) = 0$ and $a^\leftarrow(\infty) = \infty$. Note that the function

$$\frac{1}{a^\leftarrow(x)} + \frac{A_-(x)}{x} = \frac{1}{a^\leftarrow(x)} + \int_0^1 \bar{\Pi}^{(-)}(xy) dy$$

is continuous and nonincreasing, tends to ∞ as $x \rightarrow 0$, and to 0 as $x \rightarrow \infty$. Choose $\alpha \in (0, 1/2)$ arbitrarily small so that

$$2\alpha(2(1/2 - \alpha)^{-2} + 1) \leq 1,$$

and define, for $t > 0$,

$$b(t) = \inf \left\{ x > 0 : \frac{1}{a^\leftarrow(x)} + \frac{A_-(x)}{x} \leq \frac{\alpha}{t} \right\}.$$

Then $0 < b(t) < \infty$ for $t > 0$, $b(t)$ is strictly increasing, $\lim_{t \downarrow 0} b(t) = 0$, and

$$(5.26) \quad \frac{t}{a^\leftarrow(b(t))} + \frac{tA_-(b(t))}{b(t)} = \alpha.$$

Also $b(t) \geq a(t/\alpha)$, and the inverse function $b^\leftarrow(x)$ exists and satisfies

$$(5.27) \quad b^\leftarrow(x) = \frac{\alpha}{1/a^\leftarrow(x) + A_-(x)/x}.$$

Thus, by (5.24),

$$(5.28) \quad \int_0^1 b^\leftarrow(x) \Pi(dx) = \infty = \int_0^1 \bar{\Pi}^{(+)}(b(x)) dx.$$

Set

$$U_-(x) := 2 \int_0^x y \bar{\Pi}^{(-)}(y) dy.$$

Then we have the upper-bounds

$$t \bar{\Pi}^{(-)}(b(t)) \leq \frac{t A_-(b(t))}{b(t)} \leq \alpha$$

and

$$\frac{t U_-(b(t))}{b^2(t)} \leq \frac{2t A_-(b(t))}{b(t)} \leq 2\alpha.$$

Since $X \in bv$ and $\delta = 0$ we can express X in terms of its positive and negative jumps, $\Delta_s^{(+)} = \max(0, \Delta_s)$ and $\Delta_s^{(-)} = \Delta_s^{(+)} - \Delta_s$:

$$(5.29) \quad X_t = \sum_{0 < s \leq t} \Delta_s^{(+)} - \sum_{0 < s \leq t} \Delta_s^{(-)} = X_t^{(+)} - X_t^{(-)} \quad \text{say.}$$

Recall that $\Delta_s^{(\pm)} \leq 1$ a.s. We then have

$$\begin{aligned} &P(X_t^{(-)} > b(t)/2) \\ &\leq P\left(\sum_{0 < s \leq t} (\Delta_s^{(-)} \wedge b(t)) > b(t)/2\right) + P(\Delta_s^{(-)} > b(t) \text{ for some } s \leq t) \\ &\leq P\left(\sum_{0 < s \leq t} (\Delta_s^{(-)} \wedge b(t)) - t A_-(b(t)) > (1/2 - \alpha)b(t)\right) + t \bar{\Pi}^{(-)}(b(t)). \end{aligned}$$

Observe that the random variable $\sum_{0 < s \leq t} (\Delta_s^{(-)} \wedge b(t)) - t A_-(b(t))$ is centered with variance at most $t U_-(b(t))$. Hence

$$P\left(\sum_{0 < s \leq t} (\Delta_s^{(-)} \wedge b(t)) - t A_-(b(t)) > (1/2 - \alpha)b(t)\right) \leq \frac{t U_-(b(t))}{(1/2 - \alpha)^2 b^2(t)},$$

so that, by the choice of α , we finally arrive at

$$(5.30) \quad P(X_t^{(-)} > b(t)/2) \leq \left(\frac{2}{(1/2 - \alpha)^2} + 1\right)\alpha \leq 1/2.$$

By (5.28), $P(X_t^{(+)} > b(t) \text{ i.o.}) \geq P(\Delta_t^{(+)} > b(t) \text{ i.o.}) = 1$. Choose $t_n \downarrow 0$ such that $P(X_{t_n}^{(+)} > b(t_n) \text{ i.o.}) = 1$. Since the subordinators $X^{(+)}$ and $X^{(-)}$ are independent, we have

$$\begin{aligned} &P(X_{t_n} > b(t_n)/2 \text{ i.o.}) \\ &\geq \lim_{m \rightarrow \infty} P(X_{t_n}^{(+)} > b(t_n), X_{t_n}^{(-)} \leq b(t_n)/2 \text{ for some } n > m) \\ &\geq (1/2)P(X_{t_n}^{(+)} > b(t_n) \text{ i.o.}) \quad [\text{by (5.30)}] \\ &= 1/2. \end{aligned}$$

In the last inequality we used the Feller–Chung lemma ([4], page 69). Thus, by Kolmogorov’s 0–1 law, $\limsup_{t \downarrow 0} X_t/b(t) \geq 1/2$, a.s. Now since $a(x)/x$ is nondecreasing, we have for $\alpha < 1$, $b(t)/a(t) \geq a(t/\alpha)/a(t) \geq 1/\alpha$, so $\limsup_{t \downarrow 0} X_t/a(t) \geq 2/\alpha$ a.s. Letting $\alpha \downarrow 0$ gives $\limsup_{t \downarrow 0} X_t/a(t) = \infty$ a.s., as claimed in (5.25). \square

We now state a strong version of the converse of Proposition 5.3 which completes the proof of Theorem 3.2.

PROPOSITION 5.4. *The notation and assumptions are the same as in Proposition 5.3. If*

$$(5.31) \quad \int_0^1 \frac{\Pi(dx)}{1/a^{\leftarrow}(x) + A_-(x)/x} < \infty,$$

then we have

$$(5.32) \quad \limsup_{t \downarrow 0} \frac{X_t}{a(t)} \leq 0 \quad \text{a.s.}$$

We will establish Proposition 5.4 using a coupling technique similar to that in [2]. For this purpose, we first need a technical lemma, which is intuitively obvious once the notation has been assimilated. Let Y be a Lévy process and $((t_i, x_i), i \in I)$ a countable family in $(0, \infty) \times (0, \infty)$ such that the t_i ’s are pairwise distinct. Let $(Y^i, i \in I)$ be a family of i.i.d. copies of Y , and set for each $i \in I$

$$\rho_i := \inf\{s \geq 0 : Y_s^i \geq x_i\} \wedge a^{\leftarrow}(x_i),$$

where $a(\cdot)$ is as in the statement of Proposition 5.3. More generally, we could as well take for ρ_i any stopping time in the natural filtration of Y^i , depending possibly on the family $((t_i, x_i), i \in I)$.

Now assume that

$$(5.33) \quad T_t := \sum_{t_i \leq t} \rho_i < \infty \quad \text{for all } t \geq 0 \quad \text{and} \quad \sum_{i \in I} \rho_i = \infty \quad \text{a.s.}$$

Then $T = (T_t, t \geq 0)$ is a right-continuous nondecreasing process and (5.33) enables us to construct a process Y' by pasting together the paths $(Y_s^i, 0 \leq s \leq \rho_i)$ as follows. If $t = T_u$ for some $u \geq 0$, then we set

$$Y'_t = \sum_{t_i \leq u} Y^i(\rho_i).$$

Otherwise, there exists a unique $u > 0$ such that $T_{u-} \leq t < T_u$, and thus a unique index $j \in I$ for which $T_u - T_{u-} = \rho_j$, and we set

$$Y'_t = \sum_{t_i < u} Y^i(\rho_i) + Y^j(t - T_{u-}).$$

LEMMA 5.6. *Under the assumptions above, Y' is a version of Y ; in particular its law does not depend on the family $((t_i, x_i), i \in I)$.*

PROOF. The statement follows readily from the strong Markov property in the case when the family $(t_i, i \in I)$ is discrete in $[0, \infty)$. The general case is deduced by approximation. \square

We will apply Lemma 5.6 in the following framework. Consider a subordinator $X^{(-)}$ with no drift and Lévy measure $\Pi^{(-)}$; $X^{(-)}$ will play the role of the Lévy process Y above. Let also $X^{(+)}$ be an independent subordinator with no drift and Lévy measure $\Pi^{(+)}$. We write $((t_i, x_i), i \in I)$ for the family of the times and sizes of the jumps of $X^{(+)}$. By the Lévy–Itô decomposition, $((t_i, x_i), i \in I)$ is the family of the atoms of a Poisson random measure on $\mathbb{R}_+ \times \mathbb{R}_+$ with intensity $dt \otimes \Pi^{(+)}(dx)$.

Next, mark each jump of $X^{(+)}$, say (t_i, x_i) , using an independent copy $X^{(-,i)}$ of $X^{(-)}$. In other words, $((t_i, x_i, X^{(-,i)}), i \in I)$ is the family of atoms of a Poisson random measure on $\mathbb{R}_+ \times \mathbb{R}_+ \times \mathbb{D}$ with intensity $dt \otimes \Pi^{(+)}(dx) \otimes \mathbb{P}^{(-)}$, where \mathbb{D} stands for the space of càdlàg paths on $[0, \infty)$ and $\mathbb{P}^{(-)}$ for the law of $X^{(-)}$. Finally, define for every $i \in I$,

$$\rho_i := \inf\{s \geq 0 : X_s^{(-,i)} \geq x_i\} \wedge a^{\leftarrow}(x_i).$$

LEMMA 5.7. *In the notation above, the family $((t_i, \rho_i), i \in I)$ fulfills (5.33). Further, the process*

$$T_t := \sum_{t_i \leq t} \rho_i, \quad t \geq 0,$$

is a subordinator with no drift.

PROOF. Plainly,

$$\sum_{i \in I} \delta_{(t_i, \rho_i)}$$

is a Poisson random measure on $\mathbb{R}_+ \times \mathbb{R}_+$ with intensity $dt \otimes \mu(dy)$, where

$$\mu(dy) := \int_{(0, \infty)} \Pi^{(+)}(dx) \mathbb{P}^{(-)}(\tau_x \wedge a^{\leftarrow}(x) \in dy),$$

and τ_x denotes the first-passage time of $X^{(-)}$ in $[x, \infty)$. So it suffices to check that $\int_{(0, \infty)} (1 \wedge y) \mu(dy) < \infty$.

In this direction, recall (e.g., Proposition III.1 in [1]) that there is some absolute constant c such that

$$\mathbb{E}^{(-)}(\tau_x) \leq \frac{cx}{A_-(x)} \quad \forall x > 0.$$

As a consequence, we have

$$\begin{aligned} \int_{(0,\infty)} y\mu(dy) &= \int_{(0,\infty)} \Pi^{(+)}(dx) \mathbb{E}^{(-)}(\tau_x \wedge a^{\leftarrow}(x)) \\ &\leq \int_{(0,\infty)} \Pi^{(+)}(dx) (\mathbb{E}^{(-)}(\tau_x) \wedge a^{\leftarrow}(x)) \\ &\leq c \int_{(0,\infty)} \Pi^{(+)}(dx) \left(\frac{x}{A_-(x)} \wedge a^{\leftarrow}(x) \right). \end{aligned}$$

Recall that we assume that $\Pi^{(+)}$ has support in $[0, 1]$. It is readily checked that convergence of the integral in (5.31) is equivalent to

$$\int_{(0,\infty)} \Pi^{(+)}(dx) \left(\frac{x}{A_-(x)} \wedge a^{\leftarrow}(x) \right) < \infty.$$

Our claim is established. \square

We can thus construct a process X' , as in Lemma 5.6, by pasting together the paths $(X_s^{(-,i)}, 0 \leq s \leq \rho_i)$. This enables us to complete the proof of Proposition 5.4.

PROOF OF PROPOSITION 5.4. Since $X^{(-,i)}$ is independent of (t_i, x_i) , an application of Lemma 5.6 shows that X' is a subordinator which is independent of $X^{(+)}$ and has the same law as $X^{(-)}$. As a consequence, we may suppose that the Lévy process X is given in the form $X = X^{(+)} - X'$.

Set $Y_t := X'_t + a(t)$. For every jump (t_i, x_i) of $X^{(+)}$, we have by construction

$$\begin{aligned} Y(T_{t_i}) - Y(T_{t_i-}) &= X^{(-,i)}(\rho_i) + a(T_{t_i-} + \rho_i) - a(T_{t_i-}) \\ &\geq X^{(-,i)}(\rho_i) + a(\rho_i) \quad [\text{as } a(x)/x \text{ increases}] \\ &\geq x_i \quad (\text{by definition of } \rho_i). \end{aligned}$$

By summation (recall that $X^{(+)}$ has no drift), we get that $Y(T_t) \geq X_t^{(+)}$ for all $t \geq 0$.

As $T = (T_t, t \geq 0)$ is a subordinator with no drift, we know from the result of Shtatland [16] that $T_t = o(t)$ as $t \rightarrow 0$, a.s., thus with probability 1, we have for every $\varepsilon > 0$

$$X_t^{(+)} \leq X'_{\varepsilon t} + a(\varepsilon t) \quad \forall t \geq 0 \text{ sufficiently small.}$$

Since $a(x)/x$ increases, we deduce that for t sufficiently small

$$\frac{X_t}{a(t)} \leq \frac{X_t^{(+)} - X'_{\varepsilon t}}{a(\varepsilon t)/\varepsilon} \leq \varepsilon,$$

which completes the proof. \square

5.3. *Proof of Theorem 3.3.* Suppose (3.8) holds so that $X_t > 0$ for all $t \leq$ some (random) $t_0 > 0$. Thus X is irregular for $(-\infty, 0)$ and (3.10) follows from [2]. Also $X_t^{(+)} := \sum_{0 < s \leq t} \Delta_s^{(+)} \geq X_t$ a.s., so (3.8) holds with X_t replaced by $X_t^{(+)}$, and since $X_t^{(+)}$ is a subordinator with zero drift we can apply Lemma 5.2 with $\gamma = \kappa$ to get (3.9).

Conversely, [2] has that (3.10) implies $\sum_{0 < s \leq t} \Delta_s^{(-)} = o(\sum_{0 < s \leq t} \Delta_s^{(+)})$, a.s., as $t \downarrow 0$, so, for arbitrary $\varepsilon > 0$, $X_t \geq (1 - \varepsilon) \sum_{0 < s \leq t} \Delta_s^{(+)} := (1 - \varepsilon) X_t^{(+)}$, a.s., when $t \leq$ some (random) $t_0(\varepsilon) > 0$. But (3.9) and Lemma 5.2 imply $\lim_{t \downarrow 0} \sum_{0 < s \leq t} \Delta_s^{(+)} / t^\kappa = \infty$ a.s., hence (3.8).

Acknowledgments. We are grateful to the participants in the Kioloa Research Workshop, February 2005, where this work began, for stimulating input. The second author also thanks the Leverhulme Trust for their financial support, and the second and third authors thank the University of Western Australia for support and provision of facilities in early 2006. Thanks also to two referees for an extremely careful reading of the paper, many helpful suggestions and for pointing out some errors and obscurities in an earlier version; and likewise also to Mr. M. Savov.

REFERENCES

- [1] BERTOIN, J. (1996). *Lévy Processes*. Cambridge Univ. Press. [MR1406564](#)
- [2] BERTOIN, J. (1997). Regularity of the half-line for Lévy processes. *Bull. Sci. Math.* **121** 345–354. [MR1465812](#)
- [3] BLUMENTHAL, R. M. and GETTOOR, R. K. (1961). Sample functions of stochastic processes with stationary independent increments. *J. Math. Mech.* **10** 492–516. [MR0123362](#)
- [4] CHOW, Y. S. and TEICHER, H. (1988). *Probability Theory: Independence, Interchangeability, Martingales*, 2nd ed. Springer, Berlin. [MR0953964](#)
- [5] DONEY, R. A. and MALLER, R. A. (2002). Stability and attraction to normality for Lévy processes at zero and infinity. *J. Theoret. Probab.* **15** 751–792. [MR1922446](#)
- [6] DONEY, R. A. and MALLER, R. A. (2002). Stability of the overshoot for Lévy processes. *Ann. Probab.* **30** 188–212. [MR1894105](#)
- [7] DONEY, R. A. and MALLER, R. A. (2005). Passage times of random walks and Lévy processes across power law boundaries. *Probab. Theory Related Fields* **133** 57–70. [MR2197137](#)
- [8] ERICKSON, K. B. and MALLER, R. A. (2007). Finiteness of integrals of functions of Lévy processes. *Proc. London Math. Soc.* **94** 386–420.
- [9] GIHMANN, I. I. and SKOROHOD, A. V. (1975). *Theory of Stochastic Processes*. Springer, Berlin. [MR0375463](#)
- [10] KHINTCHINE, A. YA. (1939). Sur la croissance locale des processus stochastiques homogènes à accroissements indépendants. *Izv. Akad. Nauk SSSR* **3** 487–508. [MR0002054](#)
- [11] PETROV, V. V. (1975). *Sums of Independent Random Variables*. Springer, Berlin. [MR0388499](#)
- [12] PROTTER, P. E. (1990). *Stochastic Integration and Differential Equations. A New Approach*. Springer, Berlin. [MR1037262](#)
- [13] PRUITT, W. E. (1981). The growth of random walks and Lévy processes. *Ann. Probab.* **9** 948–956. [MR0632968](#)
- [14] ROGOZIN, B. A. (1968). Local behaviour of processes with independent increments. *Theory Probab. Appl.* **13** 482–486. [MR0242261](#)

- [15] SATO, K. (1999). *Lévy Processes and Infinitely Divisible Distributions*. Cambridge Univ. Press.
- [16] SHATLAND, E. S. (1965). On local properties of processes with independent increments. *Theory Probab. Appl.* **10** 317–322. [MR0183022](#)
- [17] STROMBERG, K. (1981). *Introduction to Classical Real Analysis*. Wadsworth International, Belmont, CA. [MR0604364](#)
- [18] ZHANG, C.-H. (1986) The lower limit of a normalized random walk. *Ann. Probab.* **14** 560–581. [MR0832024](#)

J. BERTOIN
LABORATOIRE DES PROBABILITÉS
UNIVERSITÉ PIERRE ET MARIE CURIE
PARIS
FRANCE
E-MAIL: jbe@ccr.jussieu.fr

R. A. DONEY
SCHOOL OF MATHEMATICS
UNIVERSITY OF MANCHESTER
MANCHESTER M60 1QD
UNITED KINGDOM
E-MAIL: rad@ma.man.ac.uk

R. A. MALLER
CENTRE FOR MATHEMATICAL ANALYSIS
AND
SCHOOL OF FINANCE AND APPLIED STATISTICS
AUSTRALIAN NATIONAL UNIVERSITY
CANBERRA ACT0200
AUSTRALIA
E-MAIL: ross.maller@anu.edu.au