Research Article

Estimation of Finite Population Mean in Multivariate Stratified Sampling under Cost Function Using Goal Programming

Atta Ullah,¹ Javid Shabbir,² Zawar Hussain,² and Bander Al-Zahrani³

¹ Department of Mathematics, Comsats Institute of Information Technology, Attock 43600, Pakistan

² Department of Statistics, Quaid-i-Azam University, Islamabad 45320, Pakistan

³ Department of Statistics, King Abdulaziz University, Jeddah 21589, Saudi Arabia

Correspondence should be addressed to Atta Ullah; attaqau13@gmail.com

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In practical utilization of stratified random sampling scheme, the investigator meets a problem to select a sample that maximizes the precision of a finite population mean under cost constraint. An allocation of sample size becomes complicated when more than one characteristic is observed from each selected unit in a sample. In many real life situations, a linear cost function of a sample size n_h is not a good approximation to actual cost of sample survey when traveling cost between selected units in a stratum is significant. In this paper, sample allocation problem in multivariate stratified random sampling with proposed cost function is formulated in integer nonlinear multiobjective mathematical programming. A solution procedure is proposed using extended lexicographic goal programming approach. A numerical example is presented to illustrate the computational details and to compare the efficiency of proposed compromise allocation.

1. Introduction

It is common practice in sample survey related to agriculture, market, industries, and social research, and so forth that usually more than one characteristic is observed from each sampled unit of population. Stratified random sampling is more suitable than other survey designs used for obtaining information from heterogeneous population for reasons of economy and efficiency. The theory of stratified random sampling deals with the properties of estimator constructed from stratified random sample and with the best (optimum) choice of sample size to be selected from various strata either to maximize the precision of constructed estimator for a fixed cost or to minimize the cost of survey for fixed precision of estimator. The sample sizes selected according to above criteria are known as "optimum allocation." In general, variance of study variate varies from stratum to stratum that provides basis for selecting optimum sample size.

Tschuprow [1] and Neyman [2] independently proposed an allocation procedure that minimizes variance of sample

mean under a linear cost function of sample size n_h in stratified random sampling scheme. Neyman [2] used Lagrange multiplier optimization technique to get optimum sample size for single variable under study. In stratified sampling, sample allocation problem becomes complicated when more than one characteristic is observed from each selected unit of a finite population. An allocation which is optimum for single characteristic may not be optimum for others unless the characteristics are highly correlated. There is need to use some compromise allocation criteria which produce an optimum allocation for all characteristics in some sense, for example, an allocation that minimizes the trace of variancecovariance matrix of the estimator of population mean or an allocation that minimizes the weighted average of variances or an allocation that maximizes the total relative efficiency of the estimators as compared to corresponding individual optimum allocation (Varshney et al. [3]). Many authors such as Dalenius [4, 5], Ghosh [6], Folks and Antle [7], Chromy [8], Bethel [9], Jahan et al. [10, 11], Khan et al. [12], Khan et al. [13, 14], Ansari et al. [15], Khan et al. [16], and Varshney et al. [17] used different compromise criterion to solve allocation problem in stratified random sampling scheme.

The cost of survey is an important factor of sample allocation to various strata. The linear cost function used in stratified sampling is given as

$$C = c_0 + \sum_{h=1}^{L} c_h n_h,$$
 (1)

where C denotes total budget available for survey, c_h for h = $1, 2, \ldots, L$ represents measurement per unit cost in the *h*th stratum, c_0 represents fixed cost of survey, and n_h is number of sample units selected in the *h*th stratum. In many practical situations, measurement unit cost and travel cost within strata are important factors of survey cost. The nonlinear cost function including measurement unit cost and traveling cost within strata is good approximation to actual cost of survey. Beardwood et al. [18] suggested that the shortest rout among k randomly disperse destination within a region is asymptotically proportional to \sqrt{k} for large k. Varshney et al. [17] used nonlinear cost function for large sample size given in (2). Consider

$$C = c_0 + \sum_{h=1}^{L} c_h n_h + \sum_{h=1}^{L} t_h \sqrt{n_h},$$
 (2)

where t_h is travel cost within *h*th stratum. The problem of finding the shortest rout among n_h selected units in *h*th stratum is often called the "shortest rout problem" in the operation research literature. If rout map and its length is given for each strata, we find shortest rout among n_h units within strata that n_h is either small or large. This shortest rout is used for practical purpose with confidence (Beardwood et al. [18]).

Consider following proposed nonlinear cost function:

$$\acute{C} = \sum_{h=1}^{L} c_h n_h + \sum_{h=1}^{L} t_h n_h^{\delta}, \tag{3}$$

where $\dot{C} = C - c_0$ and δ represents the effect of travel within strata to cost function. The value of δ is determined by solving shortest rout problem using methods discussed by Hiller and Lieberman [19]. The cost function in (2) becomes particular case of our proposed cost function given in (3) if $\delta = 0.5$.

Generally, Lagrange multiplier technique (LMT) is used to determine sample size. However, the constraint $2 \le n_h \le$ N_h , where n_h (h = 1, 2, 3, ..., L) is an integer neglected in using LMT. For integer value of sample size n_h , rounding rule is used which may lead to violating the optimality or feasibility conditions (or both). We need integer value of sample size n_h for practical implementation. Therefore, the authors did not try to use LMT and used integer programming for integer value of strata sample sizes n_h .

In this paper, we discuss compromise allocation based on minimization of coefficients of variation of regression estimators of population mean in multivariate stratified random sampling design under proposed nonlinear cost function (3). The problem is formulated in multiobjective integer nonlinear programming. The extended lexicographic goal programming technique is applied to solve formulated allocation problem. The GAMS-AlphaECP Rosenthal [20] optimization software is used to solve numerical example which illustrates the computational detail of allocation procedure.

2. Formulation of the Problem

Consider a population of N units divided in to L mutually exclusive strata of size N_h (h = 1, 2, ..., L) such that $\sum_{h=1}^{L}$ $N_h = N$. The simple random sample of size n_h is drawn from each stratum independently. Suppose we observe Y_{ii} (*i* = $1, 2, \dots, N_h, j = 1, 2, \dots, p$, $p \ge 2$, characteristics from each unit in hth stratum and estimate population mean of $p \ge 2$ characteristics. Let \overline{y}_{ih} and \overline{x}_{ih} be the sample means and \overline{Y}_{jh} and \overline{X}_{jh} the population means of study variable Y_{jh} and auxiliary variable X_{jh} , respectively, of *j*th characteristics in the *h*th stratum. S_{vih}^2 and S_{xih}^2 are population variance and S_{yxjh} is population covariance between the *j*th study and auxiliary variable in the *h*th stratum. $b_{jh} = s_{yxjh}/s_{xih}^2$ and $\beta_{jh} = S_{yxjh}/S_{xjh}^2$ are sample and population regression coefficients and $W_h = N_h/N$ is stratum weight.

Consider an estimator,

$$\overline{y}_{j,lrs} = \sum_{h=1}^{L} W_h \overline{y}_{j,lrh}, \tag{4}$$

where $\overline{y}_{j,lrh} = \overline{y}_{jh} + b_{jh}(\overline{X}_{jh} - \overline{x}_{jh})$. The mean square error (MSE) of $\overline{y}_{j,lrs}$ is given as

$$MSE\left(\overline{y}_{j,lrs}\right) = \sum_{h=1}^{L} W_h^2\left(\frac{1}{n_h} - \frac{1}{N_h}\right) \left[S_{yjh}^2 - 2\beta_{jh}S_{yxjh} + \beta_{jh}^2S_{xjh}^2\right],$$
(5)

$$MSE(\overline{y}_{j,lrs}) = \sum_{h=1}^{L} \frac{W_h^2 U_{jh}}{n_h} - \sum_{h=1}^{L} \frac{W_h^2 U_{jh}}{N_h},$$
 (6)

where

$$\hat{U}_{jh} = S_{yjh}^2 - 2\beta_{jh}S_{yxjh} + \beta_{jh}^2 S_{xjh}^2.$$
(7)

If we ignore the second term in RHS of (6) because it is independent of sample size n_h , then

$$MSE\left(\overline{y}_{j,lrs}\right) = \sum_{h=1}^{L} \frac{W_h^2 \dot{U}_{jh}}{n_h}.$$
(8)

Since different characteristics are measured with different units, we need to use an estimate which should be independent of measurement unit. Therefore, coefficient of variation is used instead of mean square error; that is,

$$C.V\left(\overline{y}_{j,lrs}\right) = \sqrt{\frac{MSE\left(\overline{y}_{j,lrs}\right)}{\overline{Y}_{j}^{2}}}$$
(9)

or

$$C.V\left(\overline{y}_{j,lrs}\right) = Z_j = \sqrt{\sum_{h=1}^{L} \frac{\hat{u}_{jh}}{n_h}},$$
(10)

where

$$\dot{u}_{jh} = \frac{W_h^2 U_{jh}}{\overline{Y}_j^2}.$$
(11)

A sample size n_h is determined under proposed nonlinear cost function in (3) that minimizes coefficients of variation of the estimator of population mean for each characteristics Y_j (j = 1, 2, ..., p). This problem may be formulated in multiobjective integer nonlinear programming as in (12). Consider

Minimize
$$(Z_1, Z_2, ..., Z_p)$$

Subject to $\sum_{h=1}^{L} c_h n_h + \sum_{h=1}^{L} t_h n_h^{\delta} \le \acute{C}$
 $2 \le n_h \le N_h$
(12)

 n_h are integers

$$n_h \in F; \quad h = 1, 2, \dots, L,$$

where F represents the feasible region that fulfills all constraints and sign restrictions. Any solution that exists within feasible region is implementable in practice.

3. Extended Lexicographic Goal Programming

Romero [21] proposed extended lexicographic goal programming method that provides a general framework which covers and allows the mixture of most common method of solving multiobjective decision making problems. It is also encompasses distance based multicriteria decision making technique. Romero [22] extended this work to make more general form of objective function. It is a technique used by decision makers for optimizing more than one objective under some constraints. In goal programming, all specified objectives are included in the model. The decision maker tries to minimize the potential deviations from specified objectives.

Consider the following individual optimum problem:

Minimize
$$Z_j$$

Subject to $\sum_{h=1}^{L} c_h n_h + \sum_{h=1}^{L} t_h n_{jh}^{\delta} \leq \acute{C}$
 $2 \leq n_{jh} \leq N_h$
(13)
 n_{jh} are integers, $n_{jh} \in F$.

$$h = 1, 2, \dots, L, \quad j = 1, 2, \dots, p.$$

Let Z_j^* be the individual optimum values of Z_j obtained by solving above problem. These optimum values Z_j^* specify objectives and try to achieve these objectives using multiobjective mathematical programming. Let \widehat{Z}_j be values of objectives obtained by applying multiobjective optimization method. It is obvious that $\widehat{Z}_j \ge Z_j^*$ or $\widehat{Z}_j - Z_j^* \ge 0$ is the increase in Z_j due to compromise among objectives using compromise criterion. Suppose this increase is $d_j^* \ge 0$. To achieve these specified objectives, we must have

$$\widehat{Z}_j - Z_j^* \le d_j^+ \tag{14}$$

or

$$\widehat{Z}_j - d_j^+ \le Z_j^*. \tag{15}$$

In goal programming method, we minimize the deviations d_j^+ using additional constraint equation (15). To solve multiobjective allocation problem (12), the extended lexicographic goal programming has following mathematical model:

Minimize
$$\alpha(\lambda) + (1 - \alpha) \left(\sum_{j=1}^{p} d_{j}^{+} \right)$$

Subject to $d_{j}^{+} \leq \lambda$

$$Z_{j} - d_{j}^{+} \leq Z_{j}^{*}$$

$$\sum_{h=1}^{L} c_{h} n_{hc} + \sum_{h=1}^{L} t_{h} n_{hc}^{\delta} \leq \acute{C}$$

$$2 \leq n_{hc} \leq N_{h}$$

$$n_{hc} \text{ are integers}$$

$$n_{hc} \in F, \quad n_{hc} \geq 0, \quad d_{j}^{+} \geq 0$$
(16)

where α is a constant that can assume minimum value zero and maximum value one. d_i^+ is positive deviational variable.

 $h = 1, 2, \dots, L, \quad j = 1, 2, \dots, p,$

4. Some Other Compromise Allocations

In this section, some other compromise allocations are discussed for the sake of comparison with the proposed allocation.

4.1. Cochran Compromise Allocation. Cochran [23] proposed a compromise allocation criteria by averaging the individual optimum allocation n_{jh}^* (j = 1, 2, 3) that is solution to integer nonlinear programming problem (INLPP) (13) over the characteristics.

Cochran's compromise allocation is given by

$$n_h = \frac{1}{p} \sum_{j=1}^p n_{jh}^*.$$
 (17)

4.2. Khan et al. Compromise Allocation. Khan et al. [13] compromise allocation is obtained by minimizing the weighted sum of variances. The mathematical model of Khan et al. [13] compromise allocation is given by

Minimize
$$\sum_{j=1}^{p} \dot{\alpha}_{j} Z_{j}$$

Subject to
$$\sum_{h=1}^{L} c_{h} n_{h} + \sum_{h=1}^{L} t_{h} n_{jh}^{\delta} \leq \dot{C}$$
$$2 \leq n_{jh} \leq N_{h}$$
$$n_{jh} \text{ are integers}, n_{jh} \in F.$$
$$h = 1, 2, \dots, L, \quad j = 1, 2, \dots, p,$$
(18)

where $\dot{\alpha}_j = \sum_{h=1}^L S_{jh}^2 / \sum_{j=1}^P \sum_{h=1}^L S_{jh}^2$ is the relative weights proposed by Khan et al.

5. Numerical Example

The data are taken from agricultural census in Iowa state conducted by National Agricultural Statistics Service, USDA, Washington DC as reported by Khan et al. [16]. We assume that $c_1 = 12$, $c_2 = 8$, $c_3 = 6$, $c_4 = 10$, $t_1 = 6$, $t_2 = 4$, $t_3 = 3$, and $t_4 = 5$.

Let

 Y_1 denote the quantity of corn harvested in 2002;

 Y_2 denote the quantity of oats harvested in 2002;

 X_1 denote the quantity of corn harvested in 1997;

 X_2 denote the quantity of oats harvested in 1997.

The data summary is given as $\overline{Y}_1 = 474973.90$, $\overline{X}_1 = 405654.19$, $\overline{Y}_2 = 1576.25$, and $\overline{X}_2 = 2116.70$. The detailed summary of data is given in Tables 1 and 2.

The allocation problem formulated in multiobjective integer nonlinear programming is

Minimize
$$\begin{pmatrix} Z_1 = \sqrt{\frac{0.000066}{n_1} + \frac{0.000809}{n_2} + \frac{0.001212}{n_3} + \frac{0.000332}{n_4}}{n_4} \\ Z_2 = \sqrt{\frac{0.000181}{n_1} + \frac{0.009411}{n_2} + \frac{0.023390}{n_3} + \frac{0.000610}{n_4}}{n_4}} \end{pmatrix}$$
(19)

Subject to

$$12n_{1} + 8n_{2} + 6n_{3} + 10n_{4} + 6n_{1}^{\delta}$$

$$+ 4n_{2}^{\delta} + 3n_{3}^{\delta} + 5n_{4}^{\delta} \le \acute{C}$$

$$2 \le n_{1} \le 8$$

$$2 \le n_{2} \le 34$$

$$2 \le n_{3} \le 45$$

$$2 \le n_{4} \le 12$$
(20)

$$n_1, n_2, n_3$$
, and n_4 are integers.

5.1. (a) Individual Optimum Allocation Method

5.1.1. Individual Optimum Allocation for Characteristic Y_1 . Consider

Minimize $Z_1 =$

$$\sqrt{\frac{0.000066}{n_{11}} + \frac{0.000809}{n_{12}} + \frac{0.001212}{n_{13}} + \frac{0.000332}{n_{14}}}$$
(21)

subject to

$$\begin{aligned} 12n_{11} + 8n_{12} + 6n_{13} + 10n_{14} + 6n_{11}^{\delta} \\ &+ 4n_{12}^{\delta} + 3n_{13}^{\delta} + 5n_{14}^{\delta} \leq \acute{C} \end{aligned}$$

$$\begin{split} 2 &\leq n_{11} \leq 8\\ 2 &\leq n_{12} \leq 34\\ 2 &\leq n_{13} \leq 45\\ 2 &\leq n_{14} \leq 12\\ n_{11}, n_{12}, n_{13}, \text{ and } n_{14} \text{ are integers.} \end{split}$$

(22)

5.1.2. Individual Optimum Allocation for Characteristic $\mathrm{Y}_2.$ Consider

Minimize $Z_2 =$

$$\sqrt{\frac{0.000181}{n_{21}} + \frac{0.009411}{n_{22}} + \frac{0.023390}{n_{23}} + \frac{0.000610}{n_{24}}}$$
(23)

subject to

$$12n_{21} + 8n_{22} + 6n_{23} + 10n_{24} + 6n_{21}^{\delta}$$

$$+ 4n_{22}^{\delta} + 3n_{23}^{\delta} + 5n_{24}^{\delta} \le \acute{C}$$

$$2 \le n_{21} \le 8$$

$$2 \le n_{22} \le 34$$

$$2 \le n_{23} \le 45$$

$$2 \le n_{24} \le 12$$

$$n_{21}, n_{22}, n_{23}, \text{ and } n_{24} \text{ are integers.}$$

$$(24)$$

h	N_h	W_h	S_{y1h}^2	S_{x1h}^2	S_{y2h}^2	S_{x2h}^2
1	8	0.0808	29267524195.5	21601503189.8	777174.1	1154134.2
2	34	0.3434	26079256582.8	19734615816.7	4987812.9	7056074.8
3	45	0.4545	42362842460.8	27129658750.0	1074510.6	2082871.3
4	12	0.1212	30728265336.9	17258237358.5	388378.5	732004.9

TABLE 1: Data summary.

TABLE 2: Data summary.

h	S_{x1y1h}	$S_{x^2y^{2h}}$	β_{1h}	β_{2h}	$\hat{u_{1h}}$	$\hat{u_{2h}}$
1	24360422802.3	902170.6	1.1249	0.7834	0.000066	0.000181
2	22003466630.3	5813439.5	1.1150	0.8239	0.000809	0.009411
3	33367597192.0	1285355.6	1.2300	0.6171	0.001212	0.023390
4	21033769867.3	456991.5	1.2188	0.4243	0.000332	0.000610

 Z_1^* and Z_2^* are coefficients of variation under individual allocation at different values of δ and \acute{C} given in Table 3.

5.2. Proposed Compromise Allocation. We used extended lexicographic goal programming model (16) for sample allocation to different strata taking into account two characteristics Y_1 and Y_2 . Consider

Minimize
$$0.2(\lambda) + (1 - 0.2)(d_1^+ + d_2^+)$$
 (25)

subject to

$$d_1^+ \le \lambda$$
$$d_2^+ \le \lambda$$
$$\sqrt{\frac{0.000066}{n_{1c}} + \frac{0.000809}{n_{2c}} + \frac{0.001212}{n_{3c}} + \frac{0.000332}{n_{4c}} - d_1^+ \le Z_1^*}$$
$$\sqrt{\frac{0.000181}{n_{1c}} + \frac{0.009411}{n_{2c}} + \frac{0.023390}{n_{3c}} + \frac{0.000610}{n_{4c}}} - d_2^+ \le Z_2^*$$

$$\begin{split} 12n_{1c} + 8n_{2c} + 6n_{3c} + 10n_{4c} + 6n_{1c}^{\delta} \\ &+ 4n_{2c}^{\delta} + 3n_{3c}^{\delta} + 5n_{4c}^{\delta} \leq \acute{C} \\ &2 \leq n_{1c} \leq 8 \\ &2 \leq n_{2c} \leq 34 \\ &2 \leq n_{3c} \leq 45 \\ &2 \leq n_{4c} \leq 12 \\ n_{hc} (h = 1, 2, 3, 4) \, \epsilon F \text{ are integer} \\ &d_1^+, d_2^+ \geq 0. \end{split}$$

(26)

Let \widehat{Z}_1 and \widehat{Z}_2 be the coefficients of variation at various values of constants δ and \acute{C} under proposed allocation given in Table 6.

5.3. Khan et al. Compromise Allocation. We have applied model (18) to find compromise allocation proposed by Khan et al. Consider

$$\operatorname{Minimize}\left(\begin{array}{c} 0.999944\sqrt{\frac{0.000066}{n_{1}} + \frac{0.000809}{n_{2}} + \frac{0.001212}{n_{3}} + \frac{0.000332}{n_{4}}}{n_{4}}} \\ + 0.000056\sqrt{\frac{0.000181}{n_{1}} + \frac{0.009411}{n_{2}} + \frac{0.023390}{n_{3}} + \frac{0.000610}{n_{4}}}{n_{4}}}\right)$$
(27)

subject to

$$\begin{split} 12n_1 + 8n_2 + 6n_3 + 10n_4 + 6n_1^\delta \\ &+ 4n_2^\delta + 3n_3^\delta + 5n_4^\delta \leq \acute{C} \\ &2 \leq n_1 \leq 8 \\ &2 \leq n_2 \leq 34 \\ &2 \leq n_3 \leq 45 \end{split}$$

 $2 \le n_4 \le 12$

 n_1, n_2, n_3 , and n_4 are integers.

(28)

The values \widehat{Z}_1 and \widehat{Z}_2 are the coefficients of variation under Khan et al. compromise allocation obtained by solving above model at different values of constants δ and \acute{C} given in Table 5.

δ	Ć	Allocation	n_1	n_2	n_3	n_4	Used Ć	Z_1^*	Z_2^*	$Trace = Z_1^* + Z_2^*$
0.5	300	Y_1	2	9	18	5	298.00	0.01602	0.05057	0.06659
0.5	500	Y_2	2	10	22	2	298.27	0.01830	0.04899	0.06729
1	500	Y_1	4	11	21	7	498.00	0.01398	0.04584	0.05982
1	500	Y_2	2	13	29	3	498.00	0.01601	0.04271	0.05872
1.5	850	Y_1	4	12	15	9	847.56	0.01420	0.04950	0.06376
	850	Y_2	2	14	21	3	833.18	0.01610	0.04560	0.06170
2	1500	Y_1	3	10	13	7	1470.00	0.01560	0.05374	0.06934
	1500	Y_2	2	10	16	3	1467.00	0.01733	0.05193	0.06926

TABLE 3: Individual optimum allocation.

TABLE 4: Cochran's compromise allocation.

δ	Ć	n_{1c}	n_{2c}	<i>n</i> _{3c}	n_{4c}	Used Ć	\widehat{Z}_1	\widehat{Z}_2	Trace = $\widehat{Z}_1 + \widehat{Z}_2$
0.5	300	2	10	20	4	308.55	0.01605	0.04851	0.06456
1.0	500	3	12	25	5	498.00	0.01429	0.04361	0.05790
1.5	850	3	13	18	6	829.25	0.01438	0.04675	0.06113
2.0	1500	3	10	15	5	1510.00	0.01581	0.05180	0.06761

TABLE 5: Khan et al. compromise allocation.

δ	Ć	n_{1c}	n_{2c}	<i>n</i> _{3c}	n_{4c}	Used Ć	\widehat{Z}_1	\widehat{Z}_2	Trace = $\widehat{Z}_1 + \widehat{Z}_2$
0.5	300	2	9	18	5	298.00	0.01602	0.05057	0.06659
1	500	4	11	21	7	498.00	0.01398	0.04584	0.05982
1.5	850	4	13	15	8	844.91	0.01418	0.04903	0.06321
2.0	1500	4	12	12	5	1495.00	0.01585	0.05386	0.06971

TABLE 6: Proposed compromise allocation.

δ	Ć	n_{1c}	n_{2c}	<i>n</i> _{3c}	n_{4c}	Used Ć	\widehat{Z}_1	\widehat{Z}_2	Trace = $\widehat{Z}_1 + \widehat{Z}_2$
0.5	300	2	11	19	3	299.49	0.01676	0.04879	0.06557
1.0	500	2	16	23	4	495.00	0.01480	0.04299	0.05779
1.5	850	2	11	22	5	848.58	0.01473	0.04583	0.06056
2.0	1500	3	9	15	6	1491.00	0.01575	0.05260	0.06835

6. Discussion

In this section, a comparative study of proposed compromise allocation with Cochran compromise allocation, Khan et al. compromise allocation, and individual optimum allocation has been made. The comparison is based on trace of variancecovariance matrix of the estimates of finite population means under compromise allocations. We assume that characteristics are independent; therefore, covariances are zero. Table 3 gives a individual optimum allocation. Tables 4 and 5 give Cochran compromise allocation and Khan compromise allocation as discussed in Section 5. The proposed compromise allocation is given in Table 6.

Table 4 shows that Cochran compromise allocation gives high trace values for $\delta = 1, 1.5$ as compared to proposed compromise allocation given in Table 6. For $\delta = 0.5, 2.0$, Cochran compromise allocation gives slightly low value of trace but is infeasible because corresponding cost exceeds the available cost. Table 5 shows that Khan et al. compromise allocation gives higher trace values than proposed compromise

TABLE 7: PRE of proposed compromise allocation to individual optimum allocation.

δ	Ć	Y_1	Y_2	δ	Ć	Y_1	Y_2
0.5	300	101.59	102.65	1.5	850	105.28	101.88
1.0	500	103.51	101.61	2.0	1500	101.45	101.33

allocation. The performance comparison of proposed compromise allocation relative to individual optimum allocation of one characteristic is used for both characteristics given in Table 7 based on percentage relative efficiency (PRE) expression given as

$$PRE = \frac{T_I}{T_C} \times 100, \qquad (29)$$

where T_I is the value of trace using individual optimum allocation and T_C is the value of trace using proposed compromise allocation. Table 7 shows that proposed compromise allocation provides more efficient estimates of population means as compared to individual optimum allocation.

7. Conclusion

On the basis of the comparison made in Section 6, we can conclude that the extended lexicographic goal programming approach always secures a feasible solution which is not granted Cochran's compromise method and it provides better results comparative to Khan et al. compromise approach and individual optimum allocation approach from the point of view of efficiency.

Conflict of Interests

The authors declare that there is no conflict of interests regarding the publication of this paper.

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