

## DISTRIBUTION ESTIMATES OF BARRIER-CROSSING PROBABILITIES OF THE YEH-WIENER PROCESS

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**Let  $Q = [0, S] \times [0, T]$  be a rectangle and  $\{X(s, t): s, t \geq 0\}$  be the two-parameter Yeh-Wiener process. This paper finds probabilities of  $X(s, t)$  crossing barriers of the type  $ast + bs + ct + d$  on the boundary  $\partial Q$ . These probabilities give lower bounds for the yet unknown probabilities of  $X(s, t)$  crossing  $ast + bs + ct + d$  on  $Q$ . The paper also discusses sharper bounds for the latter probabilities.**

1. Introduction. Let  $\{X(s, t): s, t \geq 0\}$  be the standard Yeh-Wiener process of two parameters such that it is a separable real Gaussian stochastic process satisfying:

$$(1.1) \quad X(s, t) = 0 \text{ a.s. if } s \text{ or } t \text{ is } 0,$$

$$(1.2) \quad \text{the expected value } E\{X(s, t)\} = 0 \text{ at every } s, t \geq 0,$$

$$(1.3) \quad E\{X(s, t)X(s', t')\} = \min(s, s') \cdot \min(t, t').$$

Further properties of the process are found in Yeh's [8] and [9].

For the square  $D = [0, 1] \times [0, 1]$  and its boundary  $\partial D$ , Paranjape and Park [6] showed that the probability

$$(1.4) \quad P\left\{\sup_{\partial D} X(s, t) \geq \lambda\right\} = 3N(-\lambda) - e^{4\lambda^2}N(-3\lambda), \quad \lambda \geq 0,$$

where  $N(\cdot)$  stands for the standard normal distribution function. This probability is a lower bound of the yet unknown probability,  $P\{\sup_D X(s, t) \geq \lambda\}$ . It is known (see [4] or [7]) that

$$(1.5) \quad P\left\{\sup_D X(s, t) \geq \lambda\right\} \leq 4P\{X(1, 1) \geq \lambda\} = 4N(-\lambda).$$

Recently Chan [1] showed that, for every  $\varepsilon > 0$ ,

$$(1.6) \quad P\left\{\sup_D X(s, t) \geq \lambda\right\} \leq N(\varepsilon)^{-1}P\left\{\sup_D X(s, t) \geq \lambda - \varepsilon\right\}.$$

By the same technique as he used in his paper, the upper bound can easily be improved to  $N(\varepsilon)^{-1}P\{\sup X(1, t) \geq \lambda - \varepsilon: 0 \leq t \leq 1\} = 2N(-\lambda + \varepsilon)/N(\varepsilon)$ . However it turns out to be that even this improved upper bound is not as good as  $4N(-\lambda)$  for any  $\varepsilon > 0$ . In fact

$$4N(-\lambda) < N(\varepsilon)^{-1}P\left\{\sup_{0 \leq t \leq 1} X(1, t) \geq \lambda - \varepsilon\right\}, \quad \varepsilon > 0,$$

and

$$\lim_{\epsilon \rightarrow 0^+} N(\epsilon)^{-1} P \left\{ \sup_{0 \leq t \leq 1} X(1, t) \geq \lambda - \epsilon \right\} = 4N(-\lambda) .$$

More recently Goodman [3] showed that for  $\lambda \geq 0$ ,

$$(1.7) \quad 2 \left\{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \right\} \leq P \left\{ \sup_D X(s, t) \geq \lambda \right\} .$$

Obviously the left-hand side of (1.7) is a much better lower bound of  $P\{\sup_D X(s, t) \geq \lambda\}$  than (1.4). He subsequently proves that

$$(1.8) \quad \lim_{\lambda \rightarrow \infty} \frac{2 \{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \}}{4N(-\lambda)} = 1 ,$$

thus showing that both  $2 \left\{ N(-\lambda) + \lambda \int_{\lambda}^{\infty} N(-s) ds \right\}$  and  $4N(-\lambda)$  are very good approximations of  $P\{\sup_D X(s, t) \geq \lambda\}$  for all sufficiently large  $\lambda$ .

The main purpose of this paper is to generalize the above results for more general barriers, namely, to find a formula for

$$P \left\{ \sup_{\partial D} X(s, t) - (ast + bs + ct + d) \geq 0 \right\} , \quad a, b, c, d \geq 0 ,$$

and then find a lower bound for  $P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\}$  for which (1.7) is a special case. It is apparent that for all  $a, b, c, d \geq 0$

$$(1.9) \quad P \left\{ \sup_D X(s, t) - (ast + bs + ct + d) \geq 0 \right\} \leq 4N(-d) .$$

In addition we obtain a formula for

$$P\{\sup_{\partial D} |X(s, t)| - (ast + bs + ct + d) \geq 0\} , \quad a, b, c, d \geq 0 .$$

Some results on two-parameter Brownian bridge are also included.

2. Some lemmas. To avoid unnecessary repetitions in the proofs of the theorems, the following lemmas are given. Throughout this paper  $W(t)$  and  $X(s, t)$  will denote the standard Wiener process and the Yeh-Wiener process, respectively.

LEMMA 1. (Doob [2: p. 398]). *If  $a \geq 0, b > 0, \alpha \geq 0, \beta > 0$ , then*

$$\begin{aligned} & P \left\{ \sup_{0 \leq t < \infty} [W(t) - (at + b)] \geq 0 \quad \text{or} \quad \inf_{0 \leq t < \infty} [W(t) + at + \beta] \leq 0 \right\} \\ &= \sum_{m=1}^{\infty} \exp \{ -2[m^2 ab + (m - 1)^2 \alpha \beta + m(m - 1)(a\beta + \alpha b)] \} \\ & \quad + \exp \{ -2[(m - 1)^2 ab + m^2 \alpha \beta + m(m - 1)(a\beta + \alpha b)] \} \end{aligned}$$

$$\begin{aligned}
 & - \exp \{-2[m^2(ab + \alpha\beta) + m(m - 1)\alpha\beta + m(m + 1)\alpha b]\} \\
 & - \exp \{-2[m^2(ab + \alpha\beta) + m(m - 1)a\beta + m(m - 1)\alpha b]\} .
 \end{aligned}$$

LEMMA 2. *Let  $f(t)$  be a Borel measurable function. Then for each Borel set  $E$  of real numbers,*

$$\begin{aligned}
 (2.1) \quad & P\{W(t) - f(t) \in E, 0 < t \leq 1 \mid W(1) = u\} \\
 & = P\left\{W(t) + u - (t + 1)f\left(\frac{1}{t + 1}\right) \in \frac{1}{t}E, 0 < t < \infty\right\} .
 \end{aligned}$$

*Proof.* The basic technique used here is the same as the one used by Malmquist in [5]. Observe that  $W(t)$  and  $tW(1/t)$  are equivalent processes for  $t > 0$ . Thus, the left-hand side of (2.1) reduces to

$$\begin{aligned}
 & P\left\{W\left(\frac{1}{t}\right) - \frac{1}{t}f(t) \in \frac{1}{t}E, 0 < t \leq 1 \mid W(1) = u\right\} \\
 & = P\left\{W\left(\frac{1}{t}\right) - W(1) - \left[\frac{1}{t}f(t) - u\right] \right. \\
 & \quad \left. \in \frac{1}{t}E, 0 < t \leq 1 \mid W(1) = u\right\} .
 \end{aligned}$$

Upon using the fact that  $W(1/t - 1)$  and  $W(1/t) - W(1)$  are equivalent processes for  $t > 0$ , and  $W(1/t) - W(1)$  and  $W(1)$  are independent for  $1 \geq t > 0$ , we have the result by the transformation  $1/t - 1 \rightarrow t$ .

LEMMA 2.a. *If  $f(t)$  is a Borel measurable function on  $[0, 1]$ , then*

$$\begin{aligned}
 & P\left\{\sup_{0 \leq t \leq 1} |X(1, t)| - f(t) \geq 0 \mid X(1, 1) = u\right\} \\
 & = P\left\{\sup_{0 \leq t < \infty} |X(1, t) + u| - (t + 1)f\left(\frac{1}{t + 1}\right) \geq 0\right\} ,
 \end{aligned}$$

and the same holds for  $X(t, 1)$ .

LEMMA 3. *Let  $f(s, t)$  be a Borel measurable function on  $D$ . Then for each Borel set  $E$  of real numbers,*

$$\begin{aligned}
 & P\{X(s, t) - f(s, t) \in E, (s, t) \in (0, 1]^2 \mid X(1, 1) = u\} \\
 & = P\left\{X(s + 1, t + 1) - X(1, 1) \right. \\
 & \quad \left. - \left[(s + 1)(t + 1)f\left(\frac{1}{s + 1}, \frac{1}{t + 1}\right) - u\right] \in \frac{E}{st}, (s, t) \in (0, \infty)^2\right\} .
 \end{aligned}$$

*Proof.* This lemma is a two-parameter analogue of Lemma 2, and it can be proved similarly by observing that  $X(s, t)$  and  $sX(1/s, 1/t)$  are equivalent processes for  $s, t > 0$ .

LEMMA 4. *Let  $f(t)$  and  $g(t)$  be any Borel measurable functions on  $[0, 1]$ . Then for any Borel sets  $E_1$  and  $E_2$  of real numbers,*

$$(2.2) \quad \begin{aligned} &P\{X(s, 1) - f(s) \in E_1, X(1, t) - g(t) \in E_2, (s, t) \in D \mid X(1, 1) = u\} \\ &= P\{X(s, 1) - f(s) \in E_1, 0 \leq s \leq 1 \mid X(1, 1) = u\} \\ &\quad \cdot P\{X(1, t) - g(t) \in E_2, 0 \leq t \leq 1 \mid X(1, 1) = u\}. \end{aligned}$$

*Proof.* Observe first that  $X(s, 1)$  and  $sX(1/s, 1)$  are equivalent standard Wiener processes for  $s > 0$ , and so are  $X(1, t)$  and  $tX(1, 1/t)$  for  $t > 0$ . Now  $s[X(1/s, 1) - X(1, 1) + u]$  and  $t[X(1, 1/t) - X(1, 1) + u]$  are independent processes for  $1 \geq s, t > 0$ , and they are also independent of  $\{X(s, t): (s, t) \in D\}$ . Hence (2.2) gives:

$$(2.3) \quad \begin{aligned} &P\{X(s, 1) - f(s) \in E_1, X(1, t) - g(t) \in E_2, (s, t) \in D \mid X(1, 1) = u\} \\ &= P\{s[X(1/s, 1) - X(1, 1) + u] - f(s) \in E_1, 0 < s \leq 1\} \\ &\quad \cdot P\{t[X(1, 1/t) - X(1, 1) + u] - g(t) \in E_2, 0 < t \leq 1\}. \end{aligned}$$

But the two probabilities on the right-hand side of (2.3) are equal to  $P\{X(s, 1) - f(s) \in E_1, 0 \leq s \leq 1 \mid X(1, 1) = u\}$  and  $P\{X(1, t) - g(t) \in E_2, 0 \leq t \leq 1 \mid X(1, 1) = u\}$  respectively, and hence the proof is complete.

3. **Main results and proofs.** In what follows  $\{X(s, t): s, t \geq 0\}$  will be used exclusively for the Yeh-Wiener process.

THEOREM 1. *If  $a, b, c, d \geq 0$ , then with  $\bar{a} = a + b + c + d$ ,*

$$\begin{aligned} &P\left\{\sup_{s, t \geq 0} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ &= N(-\bar{a}) + e^{-2(a+b)(c+d)}N(a + b - c - d) \\ &\quad + e^{-2(a+c)(b+d)}N(a - b + c - d) \\ &\quad - e^{2(d-a)(b+c+2d)}N(a - b - c - 3d). \end{aligned}$$

*Proof.* First observe that

$$(3.1) \quad \begin{aligned} P_1 &\equiv P\left\{\sup_{s, t \geq 0} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ &= P\left\{\sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0\right\} \\ &\quad + P\left\{\sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0\right\} \end{aligned}$$

$$\begin{aligned}
 & - P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) \right. \\
 & \left. - [(a + c)t + (b + d)] \geq 0 \right\}.
 \end{aligned}$$

Since  $X(s, 1)$  and  $X(1, t)$  are equivalent to the standard Wiener process  $W(t)$ , the first two probabilities on the right of (3.1) can be evaluated explicitly.

Now,

$$\begin{aligned}
 P_2 & \equiv P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) \right. \\
 & \left. - [(a + c)t + (b + d)] \geq 0 \right\} \\
 & = P\{X(1, 1) \geq \bar{a}\} \\
 & + \int_{-\infty}^{a+b+c+d} P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \right. \\
 & \left. \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\} dN(u).
 \end{aligned}$$

Due to the fact that

$$\begin{aligned}
 & P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0, \sup_{0 \leq t \leq 1} X(1, t) \right. \\
 & \left. - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\} \\
 (3.2) \quad & = P \left\{ \sup_{0 \leq s \leq 1} X(s, 1) - [(a + b)s + (c + d)] \geq 0 \mid X(1, 1) = u \right\} \\
 & \cdot P \left\{ \sup_{0 \leq t \leq 1} X(1, t) - [(a + c)t + (b + d)] \geq 0 \mid X(1, 1) = u \right\},
 \end{aligned}$$

we may use Lemma 2 to get

$$\begin{aligned}
 P_2 & = N(-\bar{a}) \\
 & + \int_{-\infty}^{a+b+c+d} P \left\{ \sup_{s \geq 0} X(s, 1) - [(c + d)s + (\bar{a} - u)] \geq 0 \right\} \\
 & \cdot P \left\{ \sup_{t \geq 0} X(1, t) - [(b + d)t + (\bar{a} - u)] \geq 0 \right\} dN(u) \\
 & = N(-\bar{a}) \\
 & + \int_{-\infty}^{a+b+c+d} e^{-2(c+d)(\bar{a}-u)} e^{-2(b+d)(\bar{a}-u)} dN(u) \\
 & = N(-\bar{a}) + e^{2(d-a)(b+c+2d)} N(a - b - c - 3d).
 \end{aligned}$$

The result now readily follows.

**COROLLARY.** *If  $d \geq 0$ , then*

$$P\left\{\sup_{jD} X(s, t) \geq d\right\} = 3N(-d) - e^{4d^2}N(-3d).$$

This corollary agrees with the result in [6: p. 877].

**THEOREM 2.** *If  $\{Y(s, t): (s, t) \in D\}$  is the two-parameter Brownian bridge, i.e.,  $\{Y(s, t): (s, t) \in D\} = \{X(s, t): (s, t) \in D | X(1, 1) = 0\}$  and  $a, b, c, d \geq 0$ , then*

$$\begin{aligned} P\{\sup_{jD} Y(s, t) - (ast + bs + ct + d) \geq 0\} \\ = e^{-2(b+d)\bar{a}} + e^{-2(b+d)\bar{a}} - e^{-2(b+c+2d)\bar{a}}. \end{aligned}$$

*Proof.* This follows from (3.2) by setting  $u = 0$ .

**THEOREM 3.** *If  $a, b, c \geq 0$  and  $d > 0$ , then with  $\bar{a} = a + b + c + d$  and  $\bar{c} = c + d$ ,*

$$P\left\{\sup_{jD} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1\right\} = 2f(a, b, c, d),$$

where

$$\begin{aligned} f(a, b, c, d) = & N(-\bar{a}) + \sum_{K=1}^{\infty} (-1)^{k+1} \left[ e^{-2(a+b)\bar{c}k^2} \int_{-\bar{a}-2\bar{c}k}^{\bar{a}-2\bar{c}k} dN(u) \right. \\ & \left. + e^{-2(a+c)(b+d)k^2} \int_{-\bar{a}-2(b+d)k}^{\bar{a}-2(b+d)k} dN(u) \right] \\ & - \sum_{j,k=1}^{\infty} (-1)^{j+k} e^{-2\bar{a}[\bar{c}j^2+(b+d)k^2]} \left\{ e^{2[\bar{c}j+(b+d)k]^2} \right. \\ & \left. \times \int_{-\bar{a}-2[\bar{c}j+(b+d)k]}^{\bar{a}-2[\bar{c}j+(b+d)k]} dN(u) + e^{2[\bar{c}j-(b+d)k]^2} \int_{-\bar{a}-2[\bar{c}j-(b+d)k]}^{\bar{a}-2[\bar{c}j-(b+d)k]} dN(u) \right\}. \end{aligned}$$

*Proof.* Observe that

$$\begin{aligned} P_s(u) & \equiv P\left\{\sup_{jD} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1 \mid X(1, 1) = u\right\} \\ & = P\left\{\sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a+b)s + (c+d)} \geq 1 \text{ or} \right. \\ & \quad \left. \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a+c)t + (b+d)} \geq 1 \mid X(1, 1) = u\right\}. \end{aligned}$$

Upon applying Lemma 4, we obtain

$$\begin{aligned} P_s(u) = & P\left\{\sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a+b)s + (c+d)} \geq 1 \mid X(1, 1) = u\right\} \\ & + P\left\{\sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a+c)t + (b+d)} \geq 1 \mid X(1, 1) = u\right\} \end{aligned}$$

$$\begin{aligned}
 & - P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & \cdot P\left\{ \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u \right\}.
 \end{aligned}$$

Due to Lemma 2.a., it follows

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = P\left\{ \sup_{0 \leq s < \infty} \frac{|X(s, 1) + u|}{(c + d)s + \bar{a}} \geq 1 \right\} \\
 & = P\left\{ \sup_{0 \leq s < \infty} X(s, 1) - [(c + d)s + (\bar{a} - u)] \geq 0 \right. \\
 & \quad \left. \text{or } \inf_{0 \leq s < \infty} X(s, 1) + [(c + d)s + (\bar{a} + u)] \leq 0 \right\}.
 \end{aligned}$$

Lemma 1 applied to the last expression gives:

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s < \infty} X(s, 1) - [(c + d)s + (\bar{a} - u)] \geq 0 \right. \\
 & \quad \left. \text{or } \inf_{0 \leq s < \infty} X(s, 1) + [(c + d)s + (\bar{a} + u)] \leq 0 \right\} \\
 & = \sum_{m=1}^{\infty} \{ e^{-2\bar{a}(c+d)(2m-1)^2} [e^{2(c+d)(2m-1)u} + e^{-2(c+d)(2m-1)u}] \\
 & \quad - e^{-2\bar{a}(c+d)(2m)^2} [e^{2(c+d)(2m)u} + e^{-2(c+d)(2m)u}] \}.
 \end{aligned}$$

Therefore

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq s \leq 1} \frac{|X(s, 1)|}{(a + b)s + (c + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = \sum_{j=1}^{\infty} (-1)^{j+1} e^{-2\bar{a}(c+d)j^2} [e^{2(c+d)ju} + e^{-2(c+d)ju}]
 \end{aligned}$$

and

$$\begin{aligned}
 & P\left\{ \sup_{0 \leq t \leq 1} \frac{|X(1, t)|}{(a + c)t + (b + d)} \geq 1 \mid X(1, 1) = u \right\} \\
 & = \sum_{k=1}^{\infty} (-1)^{k+1} e^{-2\bar{a}(b+d)k^2} [e^{2(b+d)ku} + e^{-2(b+d)ku}].
 \end{aligned}$$

Since  $X(1, 1)$  is the standard normal random variable, the result now follows by:

$$\begin{aligned}
 & P\left\{ \sup_{\partial D} \frac{|X(s, t)|}{ast + bs + ct + d} \geq 1 \right\} \\
 & = P\{|X(1, 1)| \geq \bar{a}\} + \int_{-\bar{a}}^{\bar{a}} P_s(u) dN(u)
 \end{aligned}$$

$$= 2N(-\bar{a}) + \int_{-\bar{a}}^{\bar{a}} P_3(u) dN(u) .$$

**THEOREM 4.** *If  $a, b, c, d \geq 0$  and  $\bar{a} = a + b + c + d, \bar{b} = b + c + d, \bar{c} = c + d$ , then for  $u < \bar{a}$ ,*

$$P_4 \equiv P \left\{ \sup_D X(s, t) - (ast + bs + ct + d) \geq 0 \mid X(1, 1) = u \right\} \\ \cong \begin{cases} \frac{\bar{b}}{b} [e^{2\bar{c}(u-\bar{a})} - e^{-2\bar{b}(u-\bar{a})}] + e^{2\bar{b}(u-\bar{a})} , & b > 0 \\ e^{2\bar{c}(u-\bar{a})} [1 + 2\bar{c}(\bar{a} - u)] , & b = 0 . \end{cases}$$

*Proof.* Upon applying Lemma 3, we obtain

$$(3.3) \quad P_4 = P \left\{ \sup_{s, t \geq 0} X(s + 1, t + 1) - X(1, t + 1) + X(1, t + 1) - X(1, 1) \right. \\ \left. - [d(s + 1)(t + 1) + c(s + 1) + b(t + 1) + a - u] \geq 0 \right\} .$$

Consider the fact that  $X(s + 1, t + 1) - X(1, t + 1)$  and  $X(1, t + 1) - X(1, 1)$  are independent processes equivalent to  $X(s, t + 1)$  and  $X(1, t)$ , respectively. The latter  $X(1, t)$  will be denoted by  $X^*(1, t)$  to signify that it is independent of  $X(s, t + 1)$ . Due to the fact that  $c(s + 1) \leq c(s + 1)(t + 1)$  for all  $c, s, t \geq 0$ , it follows from (3.3)

$$(3.4) \quad P_4 \geq P \left\{ \sup_{s, t \geq 0} X(s, t + 1) + X^*(1, t) - [\bar{c}(t + 1)s + \bar{b}t + \bar{a} - u] \geq 0 \right\} \\ \geq \int_{u-\bar{a}}^{\infty} P \left\{ \sup_{s \geq 0} X(s, t + 1) - [\bar{c}(t + 1)s - r] \geq 0 \mid \sup_{t \geq 0} X^*(1, t) \right. \\ \left. - (\bar{b}t + \bar{a} - u) = r \right\} p(r, u) dr ,$$

where  $p(r, u)$  is the probability density of

$$P \left\{ \sup_{t \geq 0} X^*(1, t) - (\bar{b}t + \bar{a} - u) \leq r \right\} \\ = \begin{cases} 1 - e^{-2\bar{b}(\bar{a}+r-u)} , & u - \bar{a} \leq r \\ 0 , & \text{otherwise} . \end{cases}$$

Thus

$$(3.5) \quad p(r, u) = \begin{cases} 2\bar{b}e^{-2\bar{b}(\bar{a}+r-u)} , & u - \bar{a} \leq r \\ 0 , & \text{otherwise} . \end{cases}$$

Observe that the probability in the integrand of (3.4) becomes

$$(3.6) \quad P\left\{\sup_{s \geq 0} X(s, t+1) - [\bar{c}(t+1)s - r] \geq 0\right\} = \begin{cases} e^{2\bar{c}r}, & r \leq 0 \\ 1, & r > 0. \end{cases}$$

Therefore, (3.5) and (3.6) together with (3.4) give

$$P_4 \geq \int_{u-\bar{a}}^0 e^{2\bar{c}r} 2\bar{b} e^{-2\bar{b}(\bar{a}+r-u)} dr + \int_0^\infty 2\bar{b} e^{-2\bar{b}(\bar{a}+r-u)} dr,$$

from which the result readily follows.

The following is a special case ( $u = 0$ ) of Theorem 4, which has broad application in Kolmogorov-Smirnov statistics.

**THEOREM 4.a.** *If  $\{Y(s, t): (s, t) \in D\}$  is the two-parameter Brownian bridge and if  $a, b, c, d \geq 0$ , then*

$$P\{\sup_D Y(s, t) - (ast + bs + ct + d) \geq 0\} \geq \begin{cases} \frac{\bar{b}}{b}(e^{-2\bar{a}\bar{c}} - e^{-2\bar{a}\bar{b}}) + e^{-2\bar{a}\bar{b}}, & b > 0 \\ (1 + 2\bar{a}\bar{c})e^{-2\bar{a}\bar{b}}, & b = 0. \end{cases}$$

**THEOREM 5.** *If  $a, b, c, d \geq 0$ , then*

$$P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\} \geq \begin{cases} N(-\bar{a}) + \frac{\bar{b}}{b}[N(\bar{a} - 2\bar{c})e^{-2\bar{c}(\bar{a}+b)} - N(\bar{a} - 2\bar{b})e^{-2\bar{a}\bar{b}}] \\ \quad + N(\bar{a} - 2\bar{b})e^{-2\bar{a}\bar{b}}, & b > 0 \\ N(-\bar{a}) + \frac{2\bar{c}}{\sqrt{2\pi}}e^{-\bar{a}^2/2} + N(\bar{a} - 2\bar{c})(1 + 2\bar{a}\bar{c} - 4\bar{c}^2)e^{-2\bar{a}\bar{c}}, & b = 0. \end{cases}$$

In particular,

$$P\{\sup_D X(s, t) - \lambda \geq 0\} \geq 2\left[(1 - \lambda^2)N(-\lambda) + \frac{\lambda}{\sqrt{2\pi}}e^{-\lambda^2/2}\right] = 2\left[N(-\lambda) + \lambda \int_\lambda^\infty N(-s)ds\right], \quad \lambda \geq 0.$$

*Proof.* The theorem now can be established by integrating lower estimates of the conditional probability  $P_4$  in Theorem 4 with respect to  $dP\{X(1, 1) \leq u\} = dN(u) = (2\pi)^{-1/2} \exp(-u^2/2)du$ . The special case when  $a = b = c = 0$  and  $d = \lambda$  agrees with Goodman's result (Theorem 3 in [3]).

In order to find sharper upper bounds for the barrier-crossing

probabilities we introduce the following: Let  $f(s, t)$  be a continuous function on  $D$ . If  $\sup_D X(s, t) - f(s, t) \geq 0$ , then define  $\tau_f = (s_0, t_0)$  where

$$s_0 = \inf \{s \in [0, 1] \mid X(s, t) = f(s, t) \text{ for some } t \in [0, 1]\},$$

$$t_0 = \inf \{t \in [0, 1] \mid X(s_0, t) = f(s_0, t)\},$$

while if  $\sup_D X(s, t) - f(s, t) < 0$ , then set  $\tau_f = (\infty, \infty)$ . Thus with the convention that  $(s_1, t_1) \leq (s_2, t_2)$  if and only if  $s_1 \leq s_2$  and  $t_1 \leq t_2$ , we have that

$$P\left\{\sup_D X(s, t) - f(s, t) \geq 0\right\} = P\left\{\tau_f \leq (1, 1)\right\}.$$

**THEOREM 6.** *If  $c, d \geq 0$ , then*

$$P\left\{\sup_D X(s, t) - (ct + d) \geq 0\right\}$$

$$\leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\}$$

$$= 2[1 - N(c + d) + \exp(-2cd)N(a - b)].$$

*Proof.* Let  $\tau$  stand for  $\tau_f$  when  $f(s, t) = ct + d$ . Define

$$F(s, t) \equiv P\{\tau \leq (s, t)\}.$$

Then

$$(3.7) \quad \begin{aligned} F(1, 1) &= P\left\{\sup_D X(s, t) - (ct + d) \geq 0\right\} \\ &= P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\} \\ &\quad + P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) < 0, \sup_D X(s, t) - (ct + d) \geq 0\right\} \\ &= P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\} \\ &\quad + \int_0^1 P\left\{\sup_{0 \leq t' \leq 1} X(1, t') - (ct' + d) < 0 \mid \tau = (s, t)\right\} dF(s, t) \\ &\leq P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (ct + d) \geq 0\right\} \\ &\quad + \int_0^1 P\left\{X(1, t) - (ct + d) < 0 \mid \tau = (s, t)\right\} dF(s, t). \end{aligned}$$

On account of the fact that  $\tau = (s, t)$  implies  $X(s, t) = ct + d$  and  $X(1, t) - X(s, t)$  is independent of the conditioning  $\tau = (s, t)$ , it follows that

$$(3.8) \quad \int_0^1 P\{X(1, t) - (ct + d) < 0 \mid \tau = (s, t)\} dF(st) \\ = \int_0^1 P\{X(1, t) - X(s, t) < 0\} dF(s, t) = \frac{1}{2} F(1, 1).$$

The theorem now follows readily from (3.7) and (3.8).

COROLLARY 6.1. *If  $b, c, d \geq 0$ , then*

$$(3.9) \quad P\left\{\sup_D X(s, t) - (bs + ct + d) \geq 0\right\} \\ \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (b^*t + d) \geq 0\right\}, \quad b^* = \max\{b, c\}.$$

*Proof.* The result follows immediately by observing that

$$P\left\{\sup_D X(s, t) - (bs + ct + d) \geq 0\right\} \\ \leq \min\left\{P\left[\sup_D X(s, t) - (bs + d) \geq 0\right], \right. \\ \left. P\left[\sup_D X(s, t) - (ct + d) \geq 0\right]\right\}.$$

The right-hand side of (3.9) can also serve as an upper bound of  $P\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\}$ , and it is certainly a substantial improvement over (1.9). We state this fact formally as a corollary.

COROLLARY 6.2. *If  $a, b, c, d \geq 0$ , then*

$$(3.10) \quad P\left\{\sup_D X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - (b^*t + d) \geq 0\right\} \\ \leq 2P\left\{\sup_{0 \leq t \leq 1} X(1, t) - d \geq 0\right\} = 4N(-d),$$

where  $b^* = \max\{b, c\}$ .

4. **Supremum over rectangular regions.** Some adjustments are needed to apply the results for the more general rectangular region  $Q = [0, S] \times [0, T]$ . The conversion formulas are given by:

$$(4.1) \quad P\left\{\sup_{\partial Q} X(s, t) - (ast + bs + ct + d) \geq 0\right\} \\ = P\left\{\sup_{\partial D} X(s, t) - (a'st + b's + c't + d') \geq 0\right\},$$

where  $a' = a\sqrt{ST}$ ,  $b' = b\sqrt{S/T}$ ,  $c' = c\sqrt{T/S}$ , and  $d' = d/\sqrt{ST}$ .

$$\begin{aligned}
 (4.2) \quad & P \left\{ \sup_Q X(s, t) - (ast + bs + ct + d) \geq 0 \mid X(S, T) = u \right\} \\
 & = P \left\{ \sup_D X(s, t) - (a'st + b's + c't + d') \geq 0 \mid X(1, 1) = u' \right\},
 \end{aligned}$$

where  $a', b', c', d'$  are as in (4.1) and  $u' = u/\sqrt{ST}$ . In (4.1), if  $\partial Q$  is replaced by  $Q$ , then  $D$  replaces  $\partial D$ .

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