

ON CONTINUOUS-TIME MARKOV PROCESSES WITH REWARDS, II

BY HIROHISA HATORI AND TOSHIO MORI

1. Let $X_t, t \geq 0$ be a continuous-time Markov process with the state space $S = \{1, 2, \dots, N\}$. The quantity a_{jk} is defined as follows: in a short time interval dt , the process that is now in state $j \in S$ will make a transition to state $k \in S$ with probability $a_{jk}dt + o(dt)$, ($j \neq k$). The probability of two or more transitions is $o(dt)$. Then this Markov process is described by the transition-rate matrix $A = (a_{jk})$, where the diagonal elements of A are defined by $a_{jj} = -\sum_{k \neq j} a_{jk}$, ($j = 1, 2, \dots, N$). Now, let us suppose that the system earns a reward at the rate of r_{jj} dollars per unit time during all the time it occupies state j . Suppose further that when the system makes a transition from state j to state k ($j \neq k$), it receives a reward of r_{jk} dollars. In the previous paper, we have given a limiting property of the total reward $R(t)$ that the system will earn in a time t , by assuming that the multiplicity of every root of $\det(sI - A) = 0$ is 1. In this paper, we shall prove this property in the case where the roots of the equation $\det(sI - A) = 0$ are not necessarily simple.

2. Let $\varphi_{jt}(\theta) \stackrel{\text{def}}{=} E\{e^{i\theta R(t)} | X_0 = j\}$ be the characteristic function of $R(t)$ given that $X_0 = j$ and put

$$(1) \quad \Phi_j(\theta, s) = \int_0^\infty \varphi_{jt}(\theta) e^{-st} dt \quad (j = 1, 2, \dots, N),$$

where s is a positive-valued variable. Introducing the $N \times N$ matrix

$$A(\theta) \stackrel{\text{def}}{=} \begin{pmatrix} a_{11} + i\theta r_{11} & a_{12} e^{i\theta r_{12}} & \dots & a_{1N} e^{i\theta r_{1N}} \\ a_{21} e^{i\theta r_{21}} & a_{22} + i\theta r_{22} & \dots & a_{2N} e^{i\theta r_{2N}} \\ \dots & \dots & \dots & \dots \\ \dots & \dots & \dots & \dots \\ a_{N1} e^{i\theta r_{N1}} & a_{N2} e^{i\theta r_{N2}} & \dots & a_{NN} + i\theta r_{NN} \end{pmatrix}$$

and the vectors

$$\Phi(\theta, s) \stackrel{\text{def}}{=} \begin{bmatrix} \Phi_1(\theta, s) \\ \vdots \\ \Phi_N(\theta, s) \end{bmatrix} \quad \text{and} \quad \mathbf{e} \stackrel{\text{def}}{=} \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix},$$

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we have the following lemma the proof of which has been shown in section 3 of [1].

LEMMA. *We have that*

$$(2) \quad \Phi(\theta, s) = (sI - A(\theta))^{-1}e,$$

where I is the $N \times N$ identity matrix.

We shall now prove the following

THEOREM. *If A is indecomposable, then the distribution function of the random variable $[R(t) - gt] / \sqrt{t}$ converges as $t \rightarrow \infty$ to a normal law, where g is a constant.*

Proof. Let $\alpha_0 = 0, \alpha_1, \dots, \alpha_{N-1}$ be the roots of the equation $\det(sI - A) = 0$ which need not all be distinct. It has been proved in Remark 1 of [1] that $\alpha_0 = 0$ is a simple root and $\Re(\alpha_j) < 0$ ($j = 1, 2, \dots, N-1$). Let $\zeta_0(\theta), \zeta_1(\theta), \dots, \zeta_{N-1}(\theta)$ be the roots of $\det(sI - A(\theta)) = 0$ such that

$$\zeta_0(\theta) \rightarrow 0, \quad \zeta_1(\theta) \rightarrow \alpha_1, \quad \dots, \quad \zeta_{N-1}(\theta) \rightarrow \alpha_{N-1} \quad \text{as } \theta \rightarrow 0.$$

Then there exist positive constants ε and θ_0 such that $\zeta_0(\theta)$ is analytic in θ for $|\theta| < \theta_0$, and

$$(3) \quad -\varepsilon \leq \Re(\zeta_0(\theta)), \quad \Re(\zeta_l(\theta)) < -2\varepsilon \quad (l = 1, 2, \dots, N-1) \text{ for } |\theta| < \theta_0.$$

From Lemma, the following expression for $\Phi(\theta, s)$ is obtained:

$$(4) \quad \Phi(\theta, s) = \frac{\sigma(\theta)}{s - \zeta_0(\theta)} + \frac{g(s, \theta)}{(s - \zeta_1(\theta))(s - \zeta_2(\theta)) \cdots (s - \zeta_{N-1}(\theta))},$$

where $\sigma(\theta)$ is an N -dimensional vector-valued function which is continuous at $\theta = 0$, and

$$(5) \quad g(s, \theta) = g_0(\theta)s^{N-2} + g_1(\theta)s^{N-3} + \cdots + g_{N-2}(\theta)$$

is a polynomial whose degree is at most $N-2$ and whose coefficients $g_0(\theta), g_1(\theta), \dots, g_{N-2}(\theta)$ are vector-valued function of θ which are continuous at $\theta = 0$. It is easy to see that the polynomial (5) may be written as follows:

$$(6) \quad \begin{aligned} g(s, \theta) = & \tau_1(\theta) \cdot (s - \zeta_2(\theta))(s - \zeta_3(\theta)) \cdots (s - \zeta_{N-1}(\theta)) \\ & + \tau_2(\theta) \cdot (s - \zeta_3(\theta)) \cdots (s - \zeta_{N-1}(\theta)) \\ & + \cdots + \tau_{N-2}(\theta) \cdot (s - \zeta_{N-1}(\theta)) + \tau_{N-1}(\theta), \end{aligned}$$

where the vector-valued functions $\tau_1(\theta), \tau_2(\theta), \dots, \tau_{N-1}(\theta)$ are continuous functions of g_0, \dots, g_{N-2} and $\zeta_1, \dots, \zeta_{N-1}$. It follows from the continuity of g 's and ζ 's that $\tau_j(\theta)$ is continuous at $\theta = 0$ ($j = 1, 2, \dots, N-1$), and therefore there exists a constant $K < +\infty$ such that

$$(7) \quad |\tau_{jk}(\theta)| \leq K \quad \text{for } |\theta| < \theta_0 \quad (j=1, 2, \dots, N-1; k=1, 2, \dots, N),$$

where $\tau_{jk}(\theta)$ is the k -th component of $\boldsymbol{\tau}_j(\theta)$. From (4) and (6), we have

$$(8) \quad \Phi(\theta, s) = \frac{\boldsymbol{\sigma}(\theta)}{s - \zeta_0(\theta)} + \frac{\boldsymbol{\tau}_1(\theta)}{s - \zeta_1(\theta)} + \frac{\boldsymbol{\tau}_2(\theta)}{(s - \zeta_1(\theta))(s - \zeta_2(\theta))} \\ + \dots + \frac{\boldsymbol{\tau}_{N-1}(\theta)}{(s - \zeta_1(\theta))(s - \zeta_2(\theta)) \dots (s - \zeta_{N-1}(\theta))}.$$

If we define $\gamma_k(t, \theta)$ ($k=1, \dots, N-1; t \geq 0; |\theta| < \theta_0$) to be

$$\gamma_1(t, \theta) = e^{\zeta_1(\theta)t}$$

and

$$\gamma_k(t, \theta) = \int_0^t e^{\zeta_k(\theta)u} \gamma_{k-1}(t-u, \theta) du \quad (k \geq 2),$$

then we have from (8) that

$$(9) \quad \boldsymbol{\varphi}_t(\theta) \stackrel{\text{def}}{=} \begin{bmatrix} \varphi_{1t}(\theta) \\ \vdots \\ \varphi_{Nt}(\theta) \end{bmatrix} = e^{\zeta_0(\theta)t} \boldsymbol{\sigma}(\theta) + \gamma_1(t, \theta) \boldsymbol{\tau}_1(\theta) + \gamma_2(t, \theta) \boldsymbol{\tau}_2(\theta) + \dots + \gamma_{N-1}(t, \theta) \boldsymbol{\tau}_{N-1}(\theta).$$

Since $\Re(\zeta_l(\theta)) < -2\varepsilon$ for $|\theta| < \theta_0$ and $l \neq 0$, it is easily verified by induction that

$$(10) \quad |\gamma_k(t, \theta)| \leq \frac{t^{k-1}}{(k-1)!} e^{-2\varepsilon t} \quad (k=1, 2, \dots, N-1).$$

It follows from (7), (9) and (10) that

$$(11) \quad |\varphi_{jt}(\theta) - e^{\zeta_0(\theta)t} \sigma_j(\theta)| \leq K e^{-2\varepsilon t} \left(1 + t + \dots + \frac{t^{N-2}}{(N-2)!} \right) \quad (j=1, \dots, N; |\theta| < \theta_0),$$

where $\sigma_j(\theta)$ is the j -th component of $\boldsymbol{\sigma}(\theta)$. Since $\boldsymbol{\varphi}_t(0) = \mathbf{e}$, we have $\boldsymbol{\sigma}(0) = \mathbf{e}$. From (11), we have for every θ and sufficiently large t

$$\left| \varphi_{jt} \left(\frac{\theta}{t} \right) - e^{\zeta_0(\theta/t)t} \sigma_j \left(\frac{\theta}{t} \right) \right| \leq K e^{-2\varepsilon t} \left(1 + t + \dots + \frac{t^{N-2}}{(N-2)!} \right).$$

This shows that the characteristic function $\varphi_{jt}(\theta/t)$ converges as $t \rightarrow \infty$ for all θ to the continuous function $e^{\zeta'_0(0)\theta}$, because $\zeta_0(0) = 0$ and so $\zeta_0(\theta/t) \cdot t$ converges to $\zeta'_0(0) \cdot \theta$ as $t \rightarrow \infty$. It follows that $e^{\zeta'_0(0)\theta}$ must be a characteristic function, and therefore that $\zeta'_0(0)$ is a pure imaginary. Now define the real number $g \stackrel{\text{def}}{=} -i\zeta'_0(0)$ and consider the family $\{[R(t) - gt]/\sqrt{t}\}_{t>0}$ of random variables. The characteristic function of the conditional distribution of the random variable $[R(t) - gt]/\sqrt{t}$ given that $X_0 = j$ is

$$(12) \quad \phi_{jt}(\theta) \stackrel{\text{def}}{=} E\{e^{i\theta[R(t) - gt]/\sqrt{t}} | X_0 = j\} = e^{-ig\sqrt{t}\theta} \varphi_{jt} \left(\frac{\theta}{\sqrt{t}} \right).$$

From (11) and (12), we have for every θ

$$(13) \quad \left| \phi_{jt}(\theta) - e^{-ig\sqrt{t}} e^{i\theta\zeta_0(\theta/\sqrt{t})t} \sigma_j\left(\frac{\theta}{\sqrt{t}}\right) \right| = O(t^{N-2}e^{-2t})$$

as $t \rightarrow \infty$. On the other hand, we have for every θ

$$(14) \quad \begin{aligned} & e^{-ig\sqrt{t}} e^{i\theta\zeta_0(\theta/\sqrt{t})t} \sigma_j\left(\frac{\theta}{\sqrt{t}}\right) \\ &= e^{-ig\sqrt{t}} e^{i\theta\zeta_0'(0)\sqrt{t} + \zeta_0''(0)\theta^2/2 + O(1/\sqrt{t})} \cdot \sigma_j\left(\frac{\theta}{\sqrt{t}}\right) \\ &\rightarrow e^{\zeta_0''(0)\theta^2/2} \end{aligned}$$

as $t \rightarrow \infty$, because $\zeta_0(0) = 0$ and $g = -i\zeta_0'(0)$. It follows from (13) and (14) that

$$\phi_{jt}(\theta) \rightarrow e^{\zeta_0''(0)\theta^2/2} \quad \text{as } t \rightarrow \infty,$$

and, by the argument similar to the one used to show that $\zeta_0'(0)$ is pure imaginary, we know that $\zeta_0''(0) \leq 0$. Therefore we have

$$\begin{aligned} \phi_t(\theta) &\stackrel{\text{def}}{=} E\{e^{i\theta[R(t)-gt]/\sqrt{t}}\} = \sum_{j=1}^N \phi_{jt}(\theta) \cdot P\{X_0=j\} \\ &\rightarrow e^{\zeta_0''(0)\theta^2/2} \quad \text{as } t \rightarrow \infty. \end{aligned}$$

This shows that $[R(t)-gt]/\sqrt{t}$ converges in distribution to the normal distribution $N(0, -\zeta_0''(0))$.

REFERENCE

- [1] HATORI, H., On continuous-time Markov processes with rewards, I. *Kōdai Math. Sem. Rep.* **18** (1966), 212-218.

TOKYO COLLEGE OF SCIENCE, AND
CHŪBU INSTITUTE OF TECHNOLOGY.