

Stationary gap distributions for infinite systems of competing Brownian particles*

Andrey Sarantsev[†] Li-Cheng Tsai[‡]

Abstract

Consider the infinite Atlas model: a semi-infinite collection of particles driven by independent standard Brownian motions with zero drifts, except for the bottom-ranked particle which receives unit drift. We derive a continuum one-parameter family of product-of-exponentials stationary gap distributions, with exponentially *growing* density at infinity. This result shows that there are infinitely many stationary gap distributions for the Atlas model, and hence resolves a conjecture of Pal and Pitman (2008) [PP08] in the negative. This result is further generalized for infinite systems of competing Brownian particles with generic rank-based drifts.

Keywords: competing Brownian particles; infinite Atlas model; stationary distribution; gap process.

AMS MSC 2010: 60H10; 60J60; 60K35.

Submitted to EJP on February 1, 2017, final version accepted on June 18, 2017.

Supersedes arXiv:1608.00628.

1 Introduction and main results

Consider a system of infinitely many Brownian particles on the real line: $X_i(t)$, $i = 1, 2, \dots$, $t \geq 0$. Assume we can rank them from bottom upward at any time $t \geq 0$: $X_{(1)}(t) \leq X_{(2)}(t) \leq \dots$, and they satisfy the following system of SDEs:

$$dX_i(t) = 1(X_i(t) = X_{(1)}(t)) dt + dW_i(t), \quad i = 1, 2, \dots, \quad (1.1)$$

where W_1, W_2, \dots denote independent Brownian motions. In plain English, the bottom particle moves as a Brownian motion with drift one, and all other particles move as

*We would like to thank Michael Aizenman, E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas, Soumik Pal, Mykhaylo Shkolnikov, and Ramon van Handel for help and useful discussion. Andrey Sarantsev was partially supported by NSF through grants DMS 1007563, DMS 1308340, DMS 1409434, and DMS 1405210. Li-Cheng Tsai was partially supported by the NSF through DMS 1106627 and the KITP graduate fellowship through NSF grant PHY11-25915.

[†]Department of Statistics and Applied Probability, University of California, Santa Barbara
 E-mail: sarantsev@pstat.ucsb.edu

[‡]Department of Mathematics, Columbia University
 E-mail: lctsai.math@gmail.com

driftless Brownian motions. This system of Brownian particles is called the *infinite Atlas model*, for the bottom particle supporting all other particles “on its shoulders”, as the ancient Atlas hero.

1.1 Infinite systems of competing Brownian particles

Although the main interest of our work is the infinite Atlas model (1.1), our result can be naturally generalized to more general systems of competing Brownian particles. In this subsection, we rigorously define these infinite systems. Finite systems of competing Brownian particles are defined very similarly in Section 2.

Letting $\mathbb{Z}_{>0} := \{1, 2, \dots\}$, $\mathbb{R}_+ := [0, \infty)$, we adopt the notations $\mathbb{R}^\infty := \{(x_1, x_2, \dots) | x_i \in \mathbb{R}\}$ and $\mathbb{R}_+^\infty := \{(z_1, z_2, \dots) | z_i \in \mathbb{R}_+\}$ for infinite dimensional spaces. We say an infinite sequence $x = (x_i)_{i \geq 1} \in \mathbb{R}^\infty$ is *rankable* if there exists a ranking permutation $p : \mathbb{Z}_{>0} \rightarrow \mathbb{Z}_{>0}$ such that $x_{p(i)} \leq x_{p(j)}$, for all $i < j \in \mathbb{Z}_{>0}$. Not every $x \in \mathbb{R}^\infty$ is rankable; for example, the sequence $x := (x_n = \frac{1}{n})_{n=1}^\infty$ is not rankable. To ensure that such a ranking permutation is unique, we resolve ties in lexicographic order: if $x_{p(i)} = x_{p(j)}$ for $i < j$, then $\mathbf{p}_x(i) < \mathbf{p}_x(j)$. We let $\mathbf{p}_x(\cdot) : \mathbb{Z}_{>0} \rightarrow \mathbb{Z}_{>0}$ denote the unique ranking permutation for a rankable x .

Hereafter, *standard Brownian motion* refers to a one-dimensional Brownian motion with zero drift and unit diffusion coefficient. Throughout this paper, we operate on a filtered probability space $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \mathbf{P})$ with the filtration satisfying the usual conditions, and fix independent standard Brownian motions W_1, W_2, \dots with respect to the filtration $(\mathcal{F}_t)_{t \geq 0}$.

Definition 1.1. Assume $X = (X(t), t \geq 0)$ is an \mathbb{R}^∞ -valued adapted process such that $X(t) = (X_i(t))_{i \geq 1}$ is rankable for every $t \geq 0$, each coordinate $X_i = (X_i(t), t \geq 0)$ is a.s. continuous, and

$$dX_i(t) = \left[\sum_{k=1}^\infty 1(\mathbf{p}_{X(t)}(k) = i) g_k \right] dt + dW_i(t), \quad i = 1, 2, \dots \tag{1.2}$$

Then X is called an infinite system of competing Brownian particles with drift coefficients g_1, g_2, \dots . We adopt the notation $Y_k(t) := X_{\mathbf{p}_{X(t)}(k)}(t)$ for the k th ranked particle, and $Z_k(t) := Y_{k+1}(t) - Y_k(t)$ for the k th gap. The \mathbb{R}_+^∞ -valued process $Z = (Z(t), t \geq 0)$, $Z(t) = (Z_k(t))_{k \geq 1}$, is called the gap process. Each $X_i = (X_i(t), t \geq 0)$ is called the i th named particle. Throughout this paper we consider rankable initial conditions, and assume without loss of generality that the initial condition $X(0)$ is standardized. That is,

$$0 = X_1(0) \leq X_2(0) \leq X_3(0) \leq \dots$$

A sufficient condition for the existence and uniqueness of (1.2) is given by [Sar16a]. To state this result, we define the configuration space of named particles:

$$\mathcal{U} = \left\{ x = (x_i)_{i \geq 1} \in \mathbb{R}^\infty \mid \lim_{i \rightarrow \infty} x_i = \infty, \text{ and } \sum_{i=1}^\infty e^{-\alpha x_i^2} < \infty, \text{ for all } \alpha > 0 \right\}, \tag{1.3}$$

as well as the corresponding space of gaps:

$$\mathcal{V} := \{(z_k)_{k=1}^\infty \in \mathbb{R}_+^\infty \mid (0, z_1, z_1 + z_2, z_1 + z_2 + z_3, \dots) \in \mathcal{U}\}. \tag{1.4}$$

Proposition 1.2 ([Sar16a, Theorem 3.2]). Assume that $x \in \mathcal{U}$ and the drift coefficients $(g_n)_{n \geq 1}$ satisfy

$$\sum_{k=1}^\infty g_k^2 < \infty. \tag{1.5}$$

Then there exists in the weak sense a unique in law version of the infinite system (1.2) with $X(0) = x$. In this case, $X(t) \in \mathcal{U}$ for every $t \geq 0$ a.s.

Remark 1.3. If, instead of (1.5), we impose a stronger condition on the drift coefficients: the sequence of drifts eventually vanishes, that is,

$$\text{for some } n_0, \quad g_{n_0} = g_{n_0+1} = \dots = 0, \tag{1.6}$$

then the system (1.2) exists in the strong sense and is pathwise unique, see [IKS13].

Remark 1.4. The gap process $Z = Z(t)$ is invariant under adding a drift $g_\infty dt$ to each named particle. Therefore, the conditions (1.5) is readily generalized to

$$\lim_{k \rightarrow \infty} g_k = g_\infty, \quad \text{and} \quad \sum_{k=1}^{\infty} (g_k - g_\infty)^2 < \infty.$$

Similarly, the condition (1.6) is generalized to the condition $g_{n_0} = g_{n_0+1} = \dots = g_\infty$.

1.2 Main result

The question of current interest is to find stationary distributions for the gap process $Z(t)$. Let us first rigorously define this concept. Take an infinite system X of competing Brownian particles with drift coefficients g_1, g_2, \dots ; let Z be its gap process.

Definition 1.5. A probability measure π on \mathbb{R}_+^∞ is called a stationary gap distribution or a quasi-stationary distribution for the system X if there exists in the weak sense a unique in law version of (1.2) with $Z(0) \sim \pi$, and for this version we have: $Z(t) \sim \pi$ for every $t \geq 0$.

Let $\text{Exp}(\lambda)$ denote the exponential distribution with mean λ^{-1} , i.e. having density $\lambda e^{-\lambda x} dx, x > 0$. The following stationary distribution of the gap process of the Atlas model (1.1) was derived by Pal and Pitman [PP08]:

$$\bar{\pi} := \bigotimes_{k=1}^{\infty} \text{Exp}(2). \tag{1.7}$$

Samples from this distribution are configurations of particles on \mathbb{R}_+ of roughly uniform density 2, where the value 2 arises from the balancing between the unit drift $g_1 = 1$ and the push-back from the crowd of particles, as heuristically explained in [Ald03]. It was further shown in [DT15] that, under (1.7), each ranked particle $Y_k(t)$ typically deviates $O(t^{1/4})$ from its starting location $Y_k(0)$ for large t .

Here, we provide a one-parameter family of stationary gap distributions π_a , with drastically distinct behaviors: the density grows exponentially as $x \rightarrow \infty$ and each rank particle Y_k travels linearly in time (in expectation). Denote the average of the first n drift coefficients by \bar{g}_n :

$$\bar{g}_n := \frac{1}{n} (g_1 + \dots + g_n). \tag{1.8}$$

The following is the main result of this paper.

Theorem 1.6. Consider an infinite system of competing Brownian particles from (1.2) with drift coefficients satisfying (1.5). Take any real number a such that

$$a > -2 \inf_{n \geq 1} \bar{g}_n. \tag{1.9}$$

(a) The following measure π_a is supported on \mathcal{V} , and is a stationary distribution for the gap process:

$$\pi_a := \bigotimes_{n=1}^{\infty} \text{Exp}(2(g_1 + \dots + g_n) + na). \tag{1.10}$$

(b) If $Z(0) \sim \pi_a$: the system is in this stationary distribution, then

$$\mathbf{E}(Y_k(t) - Y_k(0)) = -\frac{a}{2}t, \quad t \geq 0, \quad k = 1, 2, \dots$$

We now provide some important special cases of the general systems considered in Theorem 1.6.

Example 1.7. Infinite Atlas model: $g_1 = 1$, and $g_k = 0$ for $k \geq 2$. Then $\inf_{n \geq 1} \bar{g}_n = 0$, so for $a > 0$, we have the following family of stationary distributions:

$$\pi_a := \bigotimes_{n=1}^{\infty} \text{Exp}(2 + na). \tag{1.11}$$

Example 1.8. Independent Brownian motions: $g_1 = g_2 = \dots = 0$, so $\inf_{n \geq 1} \bar{g}_n = 0$, and for $a > 0$ we have the following family of stationary distributions:

$$\pi_a := \bigotimes_{n=1}^{\infty} \text{Exp}(na).$$

Example 1.9. The “inverted Atlas” model, where the bottom particle has negative drift: $g_1 = -1, g_2 = g_3 = \dots = 0$. Then $\inf_{n \geq 1} \bar{g}_n = -1$, and for $a > 2$ we get:

$$\pi_a := \bigotimes_{n=1}^{\infty} \text{Exp}(-2 + na).$$

Remark 1.10. Actually, the condition (1.5) does not play a crucial role in the proof of Theorem 1.6. More precisely, under the weaker condition $\sup |g_n| < \infty$, our proof of Theorem 1.6 still applies for constructing a copy of the infinite system with $Z(t) \sim \pi_a$ for all $t \geq 0$. The stronger condition (1.5) is assumed merely to ensure that the solution to (1.2) is unique in law, so that the notion of stationary gap distribution is well-defined.

Theorem 1.6 shows that the stationary gap distributions for systems of competing Brownian particles (and in particular for the infinite Atlas model) are *not* unique. In fact, as we further show in Appendix A, the distributions π_a are mutually *singular* for different values of a . This result in particular resolves the conjecture [PP08, Conjecture 2] of Pal and Pitman in the negative. As mentioned previously, for any a satisfying (1.9), the distribution π_a exhibits exponentially growing density as $x \rightarrow \infty$. To see why this is true, assuming the condition (1.6) for simplicity, for $(\zeta_k)_{k=1}^{\infty} \sim \pi_a$, we note that

$$L_n := \sum_{k=1}^n \mathbf{E}(\zeta_k) = \sum_{k=1}^n \frac{1}{g_1 + \dots + g_k + ka} = a^{-1} \log n + c_n,$$

where $\{c_n\}$ is a bounded sequence. Inverting this relation yields $n = c'_n e^{aL_n}$, where $c'_n := e^{-ac_n}$. This suggests that there are typically (up to a proportion) e^{aL} particles within an interval $[0, L]$. A precise statement of this is given and proven in Appendix A.

For the discrete-time analogue of independent Brownian particles from Example 1.8, quasi-stationary distributions of the type π_a already appeared in the study of the Sherrington–Kirkpatrick model of spin glasses [RA05]. Such a distribution arises naturally for independent Brownian particles. However, it is far from obvious that similar quasi-stationary distributions should appear in the context of competing Brownian particles, since rank-based drifts introduce complicated dependence among particles.

Rather, the product-of-exponential distribution π_a arises from the study of Reflected Brownian Motion (RBM). We give a heuristic derivation of the distribution π_a using RBM in the infinite-dimensional positive orthant $\mathbb{R}_{\mp}^{\infty}$ in Section 1.5. To justify this heuristic derivation (i.e. to prove Theorem 1.6) requires taking a sequence of finite systems of competing Brownian particles with suitable drift coefficients $(g_{k,N})_{k=1}^N$ and showing that the sequence converges to the infinite system. Even for the Atlas model, where $g_1 = 1$ and $g_2 = g_3 = \dots = 0$, we need to construct $g_{k,N}$ that varies in a suitable way

over $k = 2, \dots, N$, in order to simulate the pressure caused by the exponentially dense particles at $x \rightarrow \infty$; see (2.7). This is in sharp contrast with the derivation of the measure $\bar{\pi}$ (1.7), where $(g_{k,N})_{k=1}^N$ can be taken to be $(1, 0, \dots, 0)$.

Theorem 1.6 further demonstrates a sharp contrast between finite and infinite systems of competing Brownian particles, regarding the criteria for having stationary gap distributions. For a finite system to have a stationary gap distribution, the *stability condition*

$$\bar{g}_k > \bar{g}_N, \quad k = 1, \dots, N - 1 \tag{1.12}$$

must hold (see Proposition 2.2), as (1.12) imposes a “crowding” mechanism on the rank particles. On the other hand, for an infinite systems, the stationary gap distribution π_a may exist even without any form of crowding mechanisms from the drifts. As we see in Example 1.8, the drifts are not in effect. In Example 1.9, the drifts introduce a “repelling” mechanism—the *opposite* of a crowding mechanism. The sharp contrast between finite and infinite systems is due to the additional crowding effect, in infinite systems, caused by pressure from exponentially growing density under π_a .

1.3 Conjectures

Here we state some conjectures related to Theorem 1.6. First we recall that, for more general systems of competing Brownian particles than the Atlas model, [Sar16a] derived the following stationary gap distribution

$$\pi_0 := \bigotimes_{k=1}^{\infty} \text{Exp}(2(g_1 + \dots + g_k)). \tag{1.13}$$

This is done in [Sar16a, Section 4.2] under the condition (1.5) and an additional condition that there exists $N_1 < N_2 < \dots \rightarrow \infty$ such that

$$\bar{g}_k > \bar{g}_{N_j}, \quad \text{for } k = 1, \dots, N_j - 1, \quad j \geq 1. \tag{1.14}$$

Remark 1.11. It follows from Theorem 1.6 that π_0 is supported on \mathcal{V} . To see this, fix a positive $a > 0$ satisfying the condition (1.9). Let $\zeta = (\zeta_n)_{n=1}^{\infty} \sim \pi_0$ and $\zeta' = (\zeta'_n)_{n=1}^{\infty} \sim \pi_a$ be gap processes sampled from the designated distributions. With $a > 0$, comparing (1.13) and (1.10), we find that ζ stochastically dominates ζ' . That is, there exists a coupling of ζ, ζ' under which

$$\zeta_n \geq \zeta'_n, \quad \text{for all } n = 1, 2, \dots, \quad \text{a.s.} \tag{1.15}$$

By Theorem 1.6(a), we have $\zeta' \in \mathcal{V}$ a.s. Combining this with (1.15) yields $\zeta \in \mathcal{V}$ a.s.

This stationary gap distribution (1.13) generalizes the distribution (1.7) for the Altas model. Here we use the notation π_0 to unify notation with (1.10). Note that under the conditions (1.5) and (1.14), we necessarily have $\inf_n \bar{g}_n = 0$. With this, under the preceding notations, π_a is a stationary gap distribution for all $a \in [0, \infty) = \mathbb{R}_+$, including $a = 0$. We now conjecture that, the mixtures of these measures, over different values of $a \in \mathbb{R}_+$, exhaust all stationary gap distributions:

Conjecture 1.12. *Under the conditions (1.5) and (1.14), any stationary gap distribution of an infinite system of competing Brownian particles is of the following form, for some probability measure ρ on \mathbb{R}_+ :*

$$\pi_{\rho}(\cdot) := \int_{\mathbb{R}_+} \pi_a(\cdot) \rho(da).$$

Remark 1.13. For the discrete time analog of the driftless system (i.e. $g_1 = g_2 = \dots = 0$), [RA05] has already proven the analogous statement as in Conjecture 1.12. Driftless

systems differ from the systems considered in Conjecture 1.12 in that the former does not satisfy the condition (1.14). Consequently, driftless systems lack stationary gap distribution of the type π_0 , and the statement in [RA05] involves only the parameter $a > 0$.

A natural open problem following Theorem 1.6 is the large time behavior of each rank particle $Y_k(t)$. In view of Theorem 1.6(b), here we conjecture:

Conjecture 1.14. Fix $(g_n)_{n \geq 1}$ and the parameter a as in Theorem 1.6. Initiating the system of competing Brownian particles at the configuration $X_1(0) = 0$ and $(Z_k(0))_{k=1}^\infty \sim \pi_a$, we have that, for any fixed $k \in \mathbb{Z}_{>0}$,

$$\frac{Y_k(t)}{t} \rightarrow -\frac{a}{2} \quad \text{as } t \rightarrow \infty \text{ a.s.}$$

See also [Tsa17, Corollary 1.3] for a related result on the infinite Atlas model in the stationary gap distribution π_a for $a > 0$.

1.4 Motivation and literature review

The Atlas model and the more general systems of competing Brownian particles are models of interest in mathematical finance. In particular, finite systems of competing Brownian particles (with rank-based drifts and rank-based diffusion coefficients) were introduced in [BFK05] for the purposes of stock market modeling. Weak existence and uniqueness in law for these systems follows from the earlier work of [BP87]. Specific applications to mathematical finance include the study of: stability of the capital distribution [CP10], market models with splits and mergers [KS16], and portfolio optimization in [JR15]. Furthermore, finite systems of competing Brownian are of interest in their own due to their intriguing mathematical features. There has been extensive study on various aspects of their properties, including: deriving the unique stationary gap distribution [PP08, BFIKP11]; weak convergence to this stationary distribution [IPS13, Sar15a]; the stochastic monotonicity [Sar15]; small noise limits [JR14]; propagation of chaos [JM08]; refined properties of two dimensional systems [FIKP13]; and the question of triple collision (when three or more particles occupy the same position at the same time) [IK10, IKS13, BS15, Sar15b]. The last question is important because the strong solution of a finite system of competing Brownian particles is only proved to exist until the first triple collision, [IKS13].

In addition to their role in mathematical finance, systems of competing Brownian particles arise as the continuum limit of exclusion processes [KPS12], and also serve as a discrete analogue of a nonlinear diffusion governed by a McKean-Vlasov stochastic differential equation. In fact, a nonlinear diffusion can be approximated by finite systems of competing Brownian particles, see [Shk12, JR13, Rey15, DSVZ16].

Infinite systems arise as natural models of large systems. Specifically, infinite systems of competing Brownian particles were first introduced in [PP08] for a special case of the infinite Atlas model, and later in [Shk11, IKS13] for the general case, as well as in [Sar15c] for two-sided systems $X = (X_n)_{n \in \mathbb{Z}}$. Existence and uniqueness were established in [Shk11, IKS13, Sar16a]. As mentioned previously, these infinite models exhibit stationary gap distributions π_0 from (1.13) (in particular, $\bar{\pi}$ from (1.7) for the infinite Atlas model) of the desired product-of-exponential form. This is shown in [PP08] for the infinite Atlas model and in [Sar16a] for general systems. In the latter paper [Sar16a], the question of weak convergence of $Z(t)$ as $t \rightarrow \infty$ was also studied. As models of large systems, the infinite Atlas model is naturally related to a certain stochastic partial differential equation [DT15]. Furthermore, as mentioned previously, the driftless system already appeared in the description of the infinite volume limit of the Sherrington-Kirkpatrick model. See [RA05, AA09, Shk11] and the references therein.

There are several generalizations of these models: systems of competing Lévy particles, [Shk11, Sar16]; competing Brownian particles on the positive half-line, [Shk11, IKP13] (in the former paper, these are called *regulated systems*); competing Brownian particles with elastic collisions, [FIK13, FIKP13]; the case of asymmetric collisions, when particles behave after collision as if they had different mass, [KPS12]; second-order models, where drift and diffusion coefficients depend on both name and rank of the particle, [BFIKP11, FIK13].

1.5 A heuristic derivation of π_a

Here we give a heuristic derivation of the measure π_a , explaining how it arises from the theory of Reflected Brownian Motion (RBM). We shall not give detailed definition of an RBM here, and instead refer the readers to the classical survey [Wil95]. Recall from [Sar16a] that, under conditions of Proposition 1.2, the system $Y = (Y_k)_{k \geq 1}$ of ranked particles solves the following infinite system of SDEs:

$$dY_k(t) = g_k dt + dB_k(t) + \frac{1}{2}dL_{(k-1,k)}(t) - \frac{1}{2}dL_{(k,k+1)}(t), \quad k = 1, 2, \dots \tag{1.16}$$

Here, $L_{(k,k+1)}$ denotes the local time at zero of $Z_k = Y_{k+1} - Y_k$, we let $L_{(0,1)} := 0$ for consistency of notations, and

$$B_k(t) := \sum_{i=1}^{\infty} \int_0^t 1(\mathbf{p}_{X(t)}(i) = k) dW_i(t), \quad k = 1, 2, \dots$$

are independent standard Brownian motions. With (1.16), the process Z evolves as an RBM in the infinite-dimensional positive orthant \mathbb{R}_+^∞ :

$$dZ(t) = g dt + d\tilde{B}(t) + R dL(t), \tag{1.17}$$

where $g := (g_k)_{k=1}^\infty$, $\tilde{B}(t) := (B_{k+1}(t) - B_k(t))_{k=1}^\infty$, $L(t) := (L_{(k,k+1)}(t))_{k=1}^\infty$, and R is the reflection matrix a tridiagonal matrix given by

$$R = \begin{pmatrix} 1 & -\frac{1}{2} & 0 & 0 & \dots \\ -\frac{1}{2} & 1 & -\frac{1}{2} & 0 & \dots \\ 0 & -\frac{1}{2} & 1 & -\frac{1}{2} & \dots \\ 0 & 0 & -\frac{1}{2} & 1 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{pmatrix}.$$

For *finite*-dimensional RBM in the orthant, a sufficient condition for having product-of-exponential stationary distributions is the skew-symmetry condition (see, e.g. [Sar16a, Proposition 2.1] or [Wil95]). It is straightforward to verify that finite dimensional truncations of (1.17) (i.e. (2.2) in the following) satisfy the skew-symmetry condition, and have the stationary distribution given by

$$\bigotimes_{k=1}^{N-1} \text{Exp}(\lambda_k), \quad \lambda := R^{-1}\mu, \tag{1.18}$$

where $\lambda := (\lambda_k)_{k=1}^{N-1}$ and $\mu := (g_1 - g_2, \dots, g_{N-1} - g_N)$.

Now, even though (1.18) holds only in the finite-dimensional setting, let us *informally* adopt it for deriving stationary distributions in the infinite-dimensional setting. Rewrite (1.18) as $R\lambda = \mu$ (as it is not clear that R^{-1} is well-defined in infinite dimensions). A solution of this equation is

$$\lambda^* = (\lambda_k^*)_{k=1}^\infty, \quad \lambda_k^* := 2(g_1 + g_2 + \dots + g_k), \tag{1.19}$$

which gives rise to the measure π_0 in (1.13). This solution, however, is not unique: solving for the null vector $R\eta = 0$, we have

$$\begin{aligned} \eta_1 - \frac{1}{2}\eta_2 &= 0, \\ \frac{1}{2}\eta_{k-1} - \eta_k + \frac{1}{2}\eta_{k+1} &= 0, \quad k = 2, 3, \dots, \end{aligned}$$

which yields $\eta = (1, 2, 3, \dots)$. With this, we have the following general solution to (1.18):

$$\lambda := \lambda^* + a\eta, \quad \text{i.e. } \lambda_k := 2(g_1 + g_2 + \dots + g_k) + ka, \quad (1.20)$$

with the extra condition (1.9) on a to ensure that each component of λ is positive. The solution (1.20) then suggests that π_a should be also be a stationary distribution of Z .

1.6 Organization

In Section 2, we introduce finite systems of competing Brownian particles together with the necessary tools, and define the finite systems that will be used to prove Theorem 1.6. In Section 3, we prove Theorem 1.6 by establishing the convergence of the finite systems to the corresponding infinite system. Appendix A is devoted to establishing properties of the measure π_a mentioned in Section 1.2.

1.7 Acknowledgements

We would like to thank Michael Aizenman, E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas, Soumik Pal, Mykhaylo Shkolnikov, and Ramon van Handel for help and useful discussion.

Andrey Sarantsev was partially supported by NSF through grants DMS 1007563, DMS 1308340, DMS 1409434, and DMS 1405210. Li-Cheng Tsai was partially supported by the NSF through DMS 1106627 and the KITP graduate fellowship through NSF grant PHY11-25915.

2 Finite systems of competing Brownian particles

To define a finite system of competing Brownian particles, we fix $N \geq 2$ to be the number of particles, and let g_1, \dots, g_N denote the drift coefficients. Here $\mathbf{p}_x(\cdot) : \{1, \dots, N\} \rightarrow \{1, \dots, N\}$ denote the analogous ranking permutation for $x \in \mathbb{R}^N$, which is unique by resolving ties in the lexicographic order. Note that unlike in infinite dimensions, any $x \in \mathbb{R}^N$ is rankable.

Definition 2.1. Take an \mathbb{R}^N -valued continuous adapted process

$$X = (X(t), t \geq 0), \quad X(t) = (X_1(t), \dots, X_N(t)), \quad t \geq 0,$$

which satisfies the following SDEs: for $i = 1, \dots, N$,

$$dX_i(t) = \left[\sum_{k=1}^N 1(\mathbf{p}_{X(t)}(k) = i) g_k \right] dt + dW_i(t), \quad i = 1, \dots, N. \quad (2.1)$$

Then X is called a finite system of competing Brownian particles. Each $X_i = (X_i(t), t \geq 0)$ is called the i th named particle. As in Definition 1.1, we assume without loss of generality that the initial condition $X(0)$ is standardized: $0 = X_1(0) \leq X_2(0) \leq \dots \leq X_N(0)$. We similarly define ranked particles $Y_k = (Y_k(t), t \geq 0)$, and the gap process $Z = (Z(t), t \geq 0)$, $Z(t) = (Z_1(t), \dots, Z_{N-1}(t)) \in \mathbb{R}_+^{N-1}$ as

$$\begin{aligned} Y_k(t) &= X_{\mathbf{p}_{X(t)}(k)}(t), \quad k = 1, \dots, N, \\ Z_k(t) &= Y_{k+1}(t) - Y_k(t), \quad k = 1, \dots, N - 1. \end{aligned}$$

Stationary gap distributions

Systems with rank-based diffusion coefficients may also be considered, but for our purposes it is sufficient to consider unit diffusion coefficients. It is known from [IKS13] that, for any deterministic initial condition $x \in \mathbb{R}^N$, the equation (2.1) always has a strong solution, which is pathwise unique.

In the sequel we will also need to consider the dynamics for the ranked particles Y_k . To this end, we let $L_{(k,k+1)} = (L_{(k,k+1)}(t), t \geq 0)$ be the local time process at zero of Z_k , for $k = 1, \dots, N-1$, and call $L_{(k,k+1)}$ the *local time of collision between the ranked particles Y_k and Y_{k+1}* . For consistency of notation, we let $L_{(0,1)}(t) \equiv 0$ and $L_{(N,N+1)}(t) \equiv 0$. It was shown in [BG08, BFIKP11] that the dynamics of ranked particles is given by

$$dY_k(t) = g_k dt + dB_k(t) + \frac{1}{2}dL_{(k-1,k)}(t) - \frac{1}{2}dL_{(k,k+1)}(t), \quad k = 1, \dots, N, \quad (2.2)$$

where the following processes are i.i.d. standard Brownian motions:

$$B_k(t) := \sum_{i=1}^N \int_0^t 1(\mathbf{p}_{X(s)}(k) = i) dW_i(s), \quad k = 1, \dots, N. \quad (2.3)$$

Our strategy of proving Theorem 1.6 is to approximate the infinite system (1.2) by certain finite systems. To this end, let us recall the following result (proved in [PP08, BFIKP11, Sar16a]) on the necessary and sufficient condition for the existence of stationary gap distributions for finite systems.

Proposition 2.2. *Recall the notation \bar{g}_k from (1.8). There exists a stationary distribution for the gap process if and only if the stability condition (1.12). In this case, this stationary distribution is unique and is given by*

$$\pi = \bigotimes_{k=1}^{N-1} \text{Exp}(2(g_1 + \dots + g_k - k\bar{g}_N)) = \bigotimes_{k=1}^{N-1} \text{Exp}(2k(\bar{g}_k - \bar{g}_N)). \quad (2.4)$$

In addition, if the system is initiated from this stationary distribution, that is, $Z(0) \sim \pi$, then

$$\mathbf{E}(Y_k(t) - Y_k(0)) = \bar{g}_N t, \quad k = 1, \dots, N, \quad t \geq 0. \quad (2.5)$$

Now, let us define the finite systems that will be used in the proof of Theorem 1.6. For every $m \geq 2$, we let $X^{(m)} = (X_k^{(m)})_{k=1}^{m^2}$ be a system of m^2 competing Brownian particles:

$$dX_i^{(m)}(t) = \left[\sum_{i=1}^{m^2} 1(\mathbf{p}_{X^{(m)}(t)}(k) = i) g_k^{(m)} \right] dt + dW_i(t), \quad i = 1, \dots, m^2, \quad (2.6)$$

with the following drift coefficients:

$$g_k^{(m)} := \begin{cases} g_k, & k = 1, \dots, m; \\ b_m, & k = m+1, \dots, m^2, \end{cases} \quad (2.7)$$

$$\text{where } b_m := -\frac{m^2}{2(m^2 - m)}a - \frac{g_1 + \dots + g_m}{m^2 - m}. \quad (2.8)$$

This specific choice of b_m ensures that $\bar{g}^{(m)} := \frac{1}{m^2}(g_1^{(m)} + \dots + g_{m^2}^{(m)}) = -\frac{a}{2}$. Letting

$$\lambda_k^{(m)} := 2(g_1^{(m)} + \dots + g_k^{(m)} - k\bar{g}^{(m)}), \quad k = 1, \dots, m^2 - 1, \quad (2.9)$$

after elementary calculations we get:

$$\lambda_k^{(m)} = \lambda_k = 2(g_1 + \dots + g_k) + ak, \quad \text{for } k = 1, \dots, m, \quad (2.10)$$

$$\lambda_k^{(m)} = \frac{m^2-k}{m-1}(2\bar{g}_m + a), \quad \text{for } k = m+1, \dots, m^2-1. \quad (2.11)$$

The assumption (1.9) ensures that $\lambda_k^{(m)} > 0$, for $m = 1, \dots, m^2-1$. This, by (2.9), is equivalent to $\bar{g}_k^{(m)} > \bar{g}_{m^2}^{(m)}$, so by Proposition 2.2, $X^{(m)}$ has the following stationary gap distribution:

$$\pi_a^{(m)} := \bigotimes_{k=1}^{m^2-1} \text{Exp}(\lambda_k^{(m)}). \quad (2.12)$$

3 Proof of Theorem 1.6

For a dimension $d \geq 1$ and a $T \in \mathbb{R}_+$, let $C([0, T], \mathbb{R}^d)$ be the space of continuous functions $[0, T] \rightarrow \mathbb{R}^d$, and for $d = 1$, we simply write $C[0, T]$. Hereafter, we endow this space with the standard uniform topology. Let $Y^{(m)} = (Y_k^{(m)})_{k=1}^{m^2}$ and $Z^{(m)} = (Z_k^{(m)})_{k=1}^{m^2-1}$ denote the corresponding ranked particles and the gap process for the system $X^{(m)}$. We initiate $X^{(m)}$ at the stationary gap distribution $\pi_a^{(m)}$, (2.12). That is, we let

$$X_1^{(m)}(0) := 0 \leq X_2^{(m)}(0) \leq X_3^{(m)}(0) \leq \dots; \text{ and } (Z_k^{(m)}(0))_{k=1}^{m^2-1} \sim \pi_a^{(m)}.$$

3.1 Proof of Part (a)

Step 1. Recall the definition of \mathcal{V} from (1.4). Let us first prove that the probability distribution π_a is supported on \mathcal{V} . Indeed, denote $g_* := \sup_{n \geq 1} |g_n| < \infty$ and let $b := 2g_* + a$. Then $b > 0$ by (1.9), and $\lambda_n := 2(g_1 + \dots + g_n) + na \leq bn$. Therefore, $\lambda_n^{-1} \geq b^{-1}n^{-1}$. For some bounded sequence $(c_n)_{n \geq 1}$ of real numbers, we get:

$$\Lambda_n := \sum_{k=1}^n \lambda_k^{-1} \geq b^{-1} \sum_{k=1}^n k^{-1} = b^{-1} \log n + c_n.$$

Applying the inequality $(a_1 + a_2)^2 \geq a_1^2/2 - a_2^2$ for all real a_1, a_2 , we have:

$$\Lambda_n^2 \geq \frac{1}{2b^2} \log^2 n - c_n^2 \geq b' \log^2 n - c' \text{ for some constants } b', c' > 0.$$

Thus, we have:

$$\sum_{n=1}^{\infty} e^{-\alpha \Lambda_n^2} \leq \sum_{n=1}^{\infty} \exp(-\alpha b' \log^2 n + \alpha c') < \infty, \text{ for all } \alpha > 0.$$

Applying [Sar16a, Lemma 4.5], we complete the proof that the distribution π_a is supported on \mathcal{V} .

Step 2. For $n' \geq n$, we let $[x]_{\downarrow n} : (x_1, \dots, x_{n'}) \mapsto (x_1, \dots, x_n)$ denote the projection onto the first n coordinates. Fixing arbitrary n and $T \in \mathbb{R}_+$, our goal is to show that $[X^{(m)}]_{\downarrow n}$ converges to a limit process $[X]_{\downarrow n}$ as $m \rightarrow \infty$, such that X solves (2.1) and has a stationary gap distribution given by π_a . Toward this end, we will need to truncate the large system $(X_k^{(m)})_{k=1}^{m^2}$ at some fixed dimension. This is done with the help of the following lemma. Hereafter, to simplify notation, we use the letter c for any generic positive constant that depends only on g_1, g_2, \dots, a and T . Slightly abusing notation, we use the same letter c even if there are multiple such constants within the same formula.

Lemma 3.1. Fix any $T \in \mathbb{R}_+$. There exists $c \in (0, \infty)$ (depending only on $a, T, g_n, n \geq 1$, as mentioned previously), such that:

$$\mathbf{P}\left(\sup_{0 \leq t \leq T} X_k^{(m)}(t) \geq u\right) \leq ce^{c(ck-u)}, \quad \text{for } k = 1, \dots, m, u \in \mathbb{R}, \quad (3.1)$$

$$\mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) \leq ce^{-c(\log k - u)_+^2} + ck^{-2}e^{cu}, \quad \text{for } k = 1, \dots, m^2, u \in \mathbb{R}. \quad (3.2)$$

Remark 3.2. The following proof actually applies even if the term k^{-2} in (3.2) is replaced by $k^{-\ell}$, for arbitrarily large ℓ , but doing so makes various constants depend also on ℓ . Here we prove (3.2) only for $\ell = 2$ as it suffices for our purpose.

Proof. Throughout this proof, for \mathbb{R} -valued random variables X, Y , the notation $X \succeq Y$ means that X stochastically dominates Y , and likewise for $X \preceq Y$. Define the standard Gaussian density and the tail distribution function:

$$\psi(y) := \frac{1}{\sqrt{2\pi}} e^{-y^2/2} \quad \text{and} \quad \Psi(x) := \int_x^\infty \psi(y) dy.$$

We begin by showing (3.1). Since $|g_k| \leq g_* < \infty$, with b_m defined in (2.8), we have that $b_m \rightarrow -a/2$ as $m \rightarrow \infty$. This implies that $\{b_m\}_{m \geq 1}$ is bounded, and hence there exists a constant g_{**} such that

$$\text{for all } m \geq 2, \quad k = 1, \dots, m^2, \quad |g_k^{(m)}| \leq g_{**} < \infty. \tag{3.3}$$

Consequently, $X_k^{(m)}$ solves the equation (2.6) with drift coefficient being at most g_{**} , thereby

$$\mathbf{P}\left(\sup_{0 \leq t \leq T} X_k^{(m)}(t) \geq u\right) \leq \mathbf{P}\left(\sup_{0 \leq t \leq T} (X_k^{(m)}(0) + W(t) + g_{**}t) \geq u\right), \tag{3.4}$$

where $W(t)$ is a standard Brownian motion. Using the reflection principle $\mathbf{P}(\sup_{0 \leq t \leq T} W(t) \geq a) = 2\Psi((\frac{a}{\sqrt{T}})_+)$ to bound the l.h.s. of (3.4), we further obtain

$$\mathbf{P}\left(\sup_{0 \leq t \leq T} X_k^{(m)}(t) \geq u\right) \leq 2\mathbf{E}\Psi\left(\frac{u - g_{**}T - X_k^{(m)}(0)}{\sqrt{T}}\right).$$

Now, fix $k \in \{1, \dots, m\}$. By (2.10) and (1.9), we have that $\lambda_k^{(m)} = \lambda_k \geq c_*k \geq c_*$, where $c_* := a + 2 \inf_{n \geq 1} \bar{g}_n > 0$. With this, letting $(\zeta_k)_{k=1}^\infty \sim \bigotimes_{k=1}^\infty \text{Exp}(c_*)$, we have $X_k^{(m)}(0) \preceq \xi_k := \sum_{j=1}^k \zeta_j$. Since $x \mapsto \Psi(\frac{u - g_{**}T - x}{\sqrt{T}})$ is increasing, by the preceding stochastic dominance we have

$$\mathbf{P}\left(\sup_{0 \leq t \leq T} X_k^{(m)}(t) \geq u\right) \leq 2\mathbf{E}\Psi\left(\frac{1}{\sqrt{T}}(u - g_{**}T - \xi_k)\right). \tag{3.5}$$

For the Gaussian tail function $\Psi(y)$ we have the following elementary inequality

$$\Psi(y) \leq ce^{-(y_+)^2/2} \leq ce^{-cy\sqrt{T}/2}, \tag{3.6}$$

where the second inequality follows from the fact that Gaussian tails decay faster than any exponential tail. Use this to further bound the r.h.s. of (3.5):

$$\mathbf{E}\left(\Psi\left(\frac{1}{\sqrt{T}}(u - g_{**}T - \xi_k)\right)\right) \leq c\mathbf{E}\left(e^{-c(u - \xi_k)/2} e^{g_{**}T/2}\right) \leq ce^{-cu/2} \prod_{j=1}^k \mathbf{E}(e^{c\zeta_j/2}),$$

and combine this result with (3.5). Recall the following elementary formula

$$\mathbf{E}(e^{v\zeta_j}) = \frac{c_*}{c_* - v}. \tag{3.7}$$

Further using this for $v = c_*/2$ (i.e. $\mathbf{E}(e^{c_*\zeta_j/2}) = 2$), we arrive at

$$\mathbf{P}\left(\sup_{0 \leq t \leq T} X_k^{(m)}(t) \geq u\right) \leq ce^{-c_*u/2} \prod_{j=1}^k \mathbf{E}(e^{c_*\zeta_j/2}) = ce^{-cu/2} 2^k.$$

This concludes the desired bound (3.1).

We now turn to the proof of (3.2). Similarly to the preceding, here we have

$$\mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) \leq 2\mathbf{E}\Psi\left(\frac{X_k^{(m)}(0) - g_{**}T - u}{\sqrt{T}}\right).$$

With $\lambda_k^{(m)}$ defined in (2.10)–(2.11), we clearly have that $\lambda_k^{(m)} \leq \tilde{c}_*k$, for $\tilde{c}_* := a + 2 \sup_n \bar{g}_n < \infty$. Consequently, letting $(\tilde{\zeta}_k)_{k=1}^\infty \sim \bigotimes_{k=1}^\infty \text{Exp}(\tilde{c}_*k)$, we have $X_k^{(m)}(0) \succeq \tilde{\xi}_k := \sum_{j=1}^k \tilde{\zeta}_j$, and hence

$$\mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) \leq 2\mathbf{E}\Psi\left(\frac{\tilde{\xi}_k - g_{**}T - u}{\sqrt{T}}\right). \tag{3.8}$$

Fix $k_* \geq 2/\tilde{c}_*^2$. We consider the cases $k \leq k_*$ and $k > k_*$ separately. For the former, as $x \mapsto \Psi(x)$ is decreasing and $\tilde{\xi}_k > 0$, we bound the r.h.s. of (3.8) by $2\Psi\left(\frac{-g_{**}T - u}{\sqrt{T}}\right)$. By (3.6), the last expression is bounded by ce^{cu} , so

$$\mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) \leq \frac{k_*^2}{k^2} e^{cu} = \frac{ce^{cu}}{k^2}, \quad \text{for } k = 1, \dots, k_*.$$

This concludes the desired inequality (3.2) for $k \leq k_*$.

The case $k > k_*$ requires more refined estimates. Fixing $k \in \{k_* + 1, \dots, m^2\}$, we begin by establishing a bound on the lower tail of $\tilde{\xi}_k$. To this end, we consider the “truncated” variable

$$\tilde{\xi}'_k := \tilde{\xi}_k - \tilde{\xi}_{k_*} = \sum_{j=k_*+1}^k \tilde{\zeta}_j, \tag{3.9}$$

together with the centered moment generating function

$$f_k(v) := \mathbf{E}\left(e^{v(\tilde{\xi}'_k - \mathbf{E}(\tilde{\xi}'_k))}\right). \tag{3.10}$$

Recall that $\tilde{\zeta}_j \sim \text{Exp}(\tilde{c}_*j)$, and $\mathbf{E}\tilde{\zeta}_j = (\tilde{c}_*j)^{-1}$. With $\tilde{\xi}'_k$ defined in (3.9), using (3.7), we calculate this function (3.10) explicitly as

$$f_k(v) = \mathbf{E} \exp\left(v \sum_{j=k_*+1}^k (\tilde{\zeta}_j - \mathbf{E}\tilde{\zeta}_j)\right) = \prod_{j=k_*+1}^k \left(e^{-v\mathbf{E}\tilde{\zeta}_j} \mathbf{E}\left(e^{v\tilde{\zeta}_j}\right)\right) = \prod_{j=k_*+1}^k \frac{e^{-v/(\tilde{c}_*j)}}{1 - v/(\tilde{c}_*j)},$$

defined for all $|v| < \tilde{c}_*(k_* + 1)$. We further express this as

$$f_k(v) = \exp\left(\sum_{j=k_*+1}^k \left(-\log\left(1 - \frac{v}{\tilde{c}_*j}\right) - \frac{v}{\tilde{c}_*j}\right)\right). \tag{3.11}$$

To bound the r.h.s. of (3.11), apply Taylor’s formula $f(y) = f(0) + f'(0)y + \int_0^y (y-z)f'(z)dz$ with $f(y) = \log(1+y) - y$ to obtain

$$\left|\log(1+y) - y\right| = \left|\int_0^y \frac{z}{1+z} dz\right| \leq k_* \left|\int_0^y z dz\right| \leq cy^2, \quad \text{for } |y| \leq \frac{k_*}{1+k_*}.$$

Apply this inequality for $y = v/(\tilde{c}_*j)$ in (3.11), for $j = k_* + 1, \dots, k$. With $\sum_{j=1}^\infty j^{-2} < \infty$, we obtain $f_k(v) \leq e^{cv^2}$ for $|v| \leq \tilde{c}_*k_*$. Combine the result with the Chernov bound to obtain $\mathbf{P}(|\xi'_k - \mathbf{E}(\xi'_k)| \geq x) \leq e^{-xv+cv^2}$, and substitute in $v = \tilde{c}_*k_*$. We arrive at

$$\mathbf{P}(|\tilde{\xi}'_k - \mathbf{E}(\tilde{\xi}'_k)| \geq x) \leq e^{-\tilde{c}_*k_*x} e^{c(\tilde{c}_*k_*)^2} \leq ce^{-\tilde{c}_*k_*x}. \tag{3.12}$$

This yields a tail bound on the variable $\tilde{\xi}'_k$. To relate the bound back to a lower tail bound on $\tilde{\xi}_k$, we use $\tilde{\xi}_k \geq \tilde{\xi}'_k$, followed by using (3.12), whereby obtaining

$$\mathbf{P}(\tilde{\xi}_k \leq x) \leq \mathbf{P}(\tilde{\xi}'_k \leq x) \leq ce^{-\tilde{c}_*k_*(\mathbf{E}(\tilde{\xi}'_k)-x)}.$$

Further, as $\mathbf{E}(\tilde{\xi}'_k)$ and $\mathbf{E}(\tilde{\xi}_k)$ differ by $\mathbf{E}(\tilde{\xi}_{k_*}) \leq c$, we conclude

$$\mathbf{P}(\tilde{\xi}_k \leq x) \leq ce^{-\tilde{c}_*k_*(\mathbf{E}(\tilde{\xi}_k)-x)}. \tag{3.13}$$

Going back to proving (3.2), we let $F_k(x) := \mathbf{P}(\tilde{\xi}_k \leq x)$ and $G_k(x) := 1 - F_k(x)$ denote the cumulative distribution function and the tail distribution function of $\tilde{\xi}_k$, respectively. Let $\mu_k := \mathbf{E}(\tilde{\xi}_k)$ denote the expected value. By (3.8) we have

$$\begin{aligned} \mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) &\leq \mathbf{E}\Psi\left(\frac{\tilde{\xi}_k - g_{**}T - u}{\sqrt{T}}\right) = \int_{\mathbb{R}} \Psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) dF_k(x) \\ &= - \int_{[\mu_k, \infty)} \Psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) dG_k(x) + \int_{(-\infty, \mu_k]} \Psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) dF_k(x). \end{aligned}$$

Apply integration by parts (with $f(x) = \Psi(\frac{x - g_{**}T - u}{\sqrt{T}})$). Using $G_k(\infty) = F_k(-\infty) = 0$, we get:

$$- \int_{[\mu_k, \infty)} f(x) dG_k(x) = f(\mu_k)G_k(\mu_k) + \int_{[\mu_k, \infty)} f'(x)G_k(x) dx, \tag{3.14}$$

$$\int_{(-\infty, \mu_k]} f(x) dF_k(x) = f(\mu_k)F_k(\mu_k) - \int_{(-\infty, \mu_k]} f'(x)F_k(x) dx. \tag{3.15}$$

With $f'(x) = -\frac{1}{\sqrt{T}}\psi(\frac{x - g_{**}T - u}{\sqrt{T}}) < 0$, we drop the last term in (3.14). Further using $F_k(\mu_k) + G_k(\mu_k) = 1$, summing (3.14)–(3.15), we obtain:

$$\mathbf{P}\left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u\right) \leq \Psi\left(\frac{\mu_k - g_{**}T - u}{\sqrt{T}}\right) + \int_{(-\infty, \mu_k]} \frac{1}{\sqrt{T}}\psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) F_k(x) dx. \tag{3.16}$$

Next, to further bound the r.h.s. of (3.16), we first note that

$$\mu_k = \tilde{c}_* \sum_{j=1}^k j^{-1} \geq \tilde{c}_* \log k - c. \tag{3.17}$$

Using this and (3.6), we bound the first term on the r.h.s. of (3.16) as

$$\Psi\left(\frac{\mu_k - g_{**}T - u}{\sqrt{T}}\right) \leq ce^{-c(\log k - c - u)_+^2} \leq ce^{-\frac{c}{2}(\log k - u)_+^2}. \tag{3.18}$$

(Here we put $\frac{c}{2}$ just to clarify that the second inequality follows by making the constant in the exponential smaller.) As for the integral term in (3.16), we use the tail estimate (3.13) to bound $F_k(x)$ by $ce^{-\tilde{c}_*k_*(\mu_k - x)}$, and replace the integral over $(-\infty, \mu_k]$ by an integral over the entire \mathbb{R} , followed by the change of variable $\frac{\mu_k - x}{\sqrt{T}} \mapsto x$. This yields

$$\int_{(-\infty, \mu_k]} \frac{1}{\sqrt{T}}\psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) F_k(x) dx \leq c \int_{\mathbb{R}} \psi\left(\frac{\mu_k - g_{**}T - u}{\sqrt{T}} - x\right) e^{-\tilde{c}_*k_*\sqrt{T}x} dx.$$

The last integral is explicitly evaluated to be $e^{(\tilde{c}_*k_*\sqrt{T})^2/2} e^{-\tilde{c}_*k_*(\mu_k - g_{**}T - u)} \leq ce^{-\tilde{c}_*k_*(\mu_k - u)}$. Combining this with the estimate of μ_k (3.17), followed by using $\tilde{c}_*^2k_* \geq 2$, we conclude

$$\int_{(-\infty, \mu_k]} \frac{1}{\sqrt{T}}\psi\left(\frac{x - g_{**}T - u}{\sqrt{T}}\right) F_k(x) dx \leq ce^{-\tilde{c}_*^2k_* \log k + \tilde{c}_*k_*u} \leq ck^{-2}e^{cu}. \tag{3.19}$$

Inserting (3.18)–(3.19) into (3.16), we conclude the desired estimate (3.2) for $k > k_*$. \square

Step 3. We now return to showing the convergence of $[X^{(m)}]_{\downarrow n}$. For $f \in C([0, T], \mathbb{R}^n)$, we let

$$\text{osc}_\delta(f) := \sup \{ \|f(t) - f(s)\|_2 \mid s, t \in [0, T], |t - s| \leq \delta \}$$

denote the modulus of continuity of f , measured in the Euclidean norm $\|x\|_2 := \sqrt{x_1^2 + \dots + x_n^2}$. Since $X^{(m)} = (X_i^{(m)})_{i=1}^{m^2}$ solves the equation (2.6) with drift coefficients bounded by g_{**} , we have

$$\sup_{m \geq 1} \mathbf{P} \left(\text{osc}_\delta([X^{(m)}]_{\downarrow n}) \geq \varepsilon \right) \rightarrow 0, \text{ as } \delta \rightarrow 0, \tag{3.20}$$

for any fixed $\varepsilon > 0$. With this, by Arzelà–Ascoli theorem, it is standard to show that $\{[X^{(m)}]_{\downarrow n}\}_{m \geq 1}$ is tight in $C([0, T], \mathbb{R}^n)$. By the Skorohod representation theorem, after passing to a subsequence and a different probability space, we have

$$[X_k^{(m)}]_{\downarrow n} \rightarrow (X_k)_{k=1}^n \text{ in } C([0, T], \mathbb{R}^n), \text{ as } m \rightarrow \infty, \text{ a.s.} \tag{3.21}$$

The limit process $X = (X_i)_{i \geq 1}$ is taken to be independent of T and n by a standard diagonal argument.

Step 4. We now proceed to show that X has gap distribution π_a . Fix $T > 0$ and $n = 1, 2, \dots$. We do this by first establishing the convergence of $[Y^{(m)}]_{\downarrow n}$. By (3.1) we have

$$\mathbf{P} \left(\max_{k=1, \dots, n} \sup_{t \leq T} X_k^{(m)}(t) > u \right) \leq \sum_{k=1}^n \mathbf{P} \left(\sup_{t \leq T} X_k^{(m)}(t) > u \right) \leq cn e^{c(cn-u)} \rightarrow 0, \text{ as } u \rightarrow \infty.$$

Fix an arbitrarily small $\varepsilon > 0$. With n, T already being fixed, we now choose a sufficiently large $u \in \mathbb{R}_+$ such that

$$\mathbf{P}(X_k^{(m)}(t) \leq u, \forall k = 1, \dots, n, t \leq T) \geq 1 - \varepsilon/2, \text{ for all } m \geq n. \tag{3.22}$$

That is, with probability at least $1 - \varepsilon/2$, the first n named particles $X_1^{(m)}, \dots, X_n^{(m)}$ always stay below the level u within the time interval $[0, T]$. With this u , we further apply (3.2) to obtain

$$\begin{aligned} \mathbf{P} \left(\min_{n' \leq k \leq m^2} \inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u \right) &\leq \sum_{k=n'}^{m^2} \mathbf{P} \left(\inf_{0 \leq t \leq T} X_k^{(m)}(t) \leq u \right) \\ &\leq c \sum_{k=n'}^{\infty} (e^{-c(\log k - u)_+^2} + k^{-2} e^{cu}). \end{aligned} \tag{3.23}$$

Since $\sum_{k=1}^{\infty} (e^{-c(\log k - u)_+^2} + k^{-2} e^{cu}) < \infty$, the last expression in (3.23) clearly tends to zero as $n' \rightarrow \infty$. Fix some $\tilde{n} \geq n$ such that

$$\mathbf{P}(X_k^{(m)}(t) \geq u, \forall k > \tilde{n}, t \leq T) \geq 1 - \varepsilon/2. \tag{3.24}$$

That is, with probability at least $1 - \varepsilon/2$, none of the name particles $X_{\tilde{n}+1}, X_{\tilde{n}+2}, \dots$ ever reaches a level below u within $[0, T]$. Let $\mathcal{R} : \mathbb{R}^{\tilde{n}} \rightarrow \mathbb{R}^{\tilde{n}}$, $\mathcal{R}(x) := (x_{p_x(i)})_{i=1}^{\tilde{n}}$, denote the ranking map of an \tilde{n} -tuple. By (3.22) and (3.24), we have that

$$\mathbf{P} \left([\mathcal{R}([X^{(m)}(t)]_{\downarrow \tilde{n}})]_{\downarrow n} = [Y^{(m)}(t)]_{\downarrow n}, \forall t \leq T \right) \geq 1 - \varepsilon.$$

Namely, with probability at least $1 - \varepsilon$, to obtain the first n ranked particles $Y_1^{(m)}(t), \dots, Y_n^{(m)}(t)$ within the system of m^2 particles $X_1^{(m)}, \dots, X_{m^2}^{(m)}$, it suffices to rank *only* $X_1^{(m)}(t)$,

$\dots, X_{\tilde{n}}^{(m)}(t)$ (as opposed to ranking all m^2 named particles), and take the first n resulting particles. Because the map \mathcal{R} is continuous, and $[X^{(m)}]_{\downarrow \tilde{n}} \rightarrow [X]_{\downarrow \tilde{n}}$ in $C([0, T], \mathbb{R}^{\tilde{n}})$ (from (3.21)), it follows that

$$[Y^{(m)}]_{\downarrow n} \rightarrow [Y]_{\downarrow n} \text{ in } C([0, T], \mathbb{R}^n) \text{ a.s., where } Y_k(t) := X_{\mathbf{p}_{X(t)}(k)}(t). \tag{3.25}$$

Further, since we have $(Z_k^{(m)}(t))_{k=1}^m \sim \pi_a^{(m)}$, letting $m \rightarrow \infty$ we further obtain that

$$(Z_k(t))_{k=1}^\infty \sim \bigotimes_{k=1}^\infty \text{Exp}(\lambda_k), \quad \forall t \in \mathbb{R}_+, \quad \text{where } Z_k(t) := Y_{k+1}(t) - Y_k(t).$$

We have thus concluded that the gap process $Z(t)$ of the system X is distributed according to π_a for all $t \in \mathbb{R}_+$.

Step 5. Finally, we still need to show that X solves (1.2). Doing so requires first showing that Y solves the corresponding equation (1.16). Indeed, the ranked process $Y^{(m)}$ solves the following finite system of SDEs:

$$Y_k^{(m)}(t) = Y_k^{(m)}(0) + g_k^{(m)}t + B_k^{(m)}(t) + \frac{1}{2}L_{(k-1,k)}^{(m)}(t) - \frac{1}{2}L_{(k,k+1)}^{(m)}(t), \quad k = 1, \dots, m^2, \tag{3.26}$$

where the local time $L_{(k,k+1)}^{(m)}$ and the Brownian motion $B_k^{(m)}$ are defined in Section 2. Note that although we take $(W_k)_{k=1}^\infty$ to be fixed (i.e. independent of m), the Brownian motions $B_k^{(m)}$ (defined in (2.3)) still depend on m . However, with $[B^{(m)}]_{\downarrow n}$ being tight in $C([0, T], \mathbb{R}^n)$, applying again the Skorohod representation theorem (after passing to a finer subsequence and yet another probability space), we assume without loss of generality $[B^{(m)}]_{\downarrow n} \rightarrow [B]_{\downarrow n}$ in $C([0, T], \mathbb{R}^n)$, where $B(t) = (B_k(t))_{k=1}^\infty$, and $B_k(t)$, $k = 1, 2, \dots$ are independent standard Brownian motions.

Now, as we already have that $Y^{(m)}$ and $B^{(m)}$ converge, taking $m \rightarrow \infty$ in (3.26) for $k = 1$ yields

$$\frac{1}{2}L_{(1,2)}^{(m)} \rightarrow \frac{1}{2}L_{(1,2)} \text{ in } C[0, T] \text{ a.s., where } \frac{1}{2}L_{(1,2)}(t) := -Y_1(t) + Y_1(0) + g_1t + B_1(t).$$

Performing this $m \rightarrow \infty$ procedure inductively for $k = 2, 3, \dots$, we further obtain

$$\frac{1}{2}L_{(k,k+1)}^{(m)} \rightarrow \frac{1}{2}L_{(k,k+1)} \text{ in } C[0, T] \text{ a.s.,}$$

where $L_{(k,k+1)}(t)$ is defined inductively through the following relation

$$\frac{1}{2}L_{(k,k+1)}(t) := \frac{1}{2}L_{(k-1,k)}(t) - Y_k(t) + Y_k(0) + g_k t + B_k(t).$$

Next, each $L_{(k,k+1)}$ is continuous, nondecreasing, and starts from zero: $L_{(k,k+1)}(0) = 0$. This is so because each $L_{(k,k+1)}^{(m)}$ has all these properties, and they are preserved in limits under the uniform topology of $C[0, T]$. Furthermore, $L_{(k,k+1)}$ increases only when $Y_k = Y_{k+1}$. To see this, we consider a generic $t \in [0, T]$ such that $Y_k(t) < Y_{k+1}(t)$. By the continuity of $Y_k(t)$ and $Y_{k+1}(t)$, we must also have $Y_k(s) < Y_{k+1}(s)$ for $s \in [t', t'']$, for some $t' < t'' \in [0, T]$. With this, for all large enough m , we have $Y_k^{(m)}(s) < Y_{k+1}^{(m)}(s)$, $s \in [t', t'']$. From the properties of the local time for finite systems, we get: $L_{(k,k+1)}^{(m)}(t') = L_{(k,k+1)}^{(m)}(t'')$. Letting $m \rightarrow \infty$ yields $L_{(k,k+1)}(t') = L_{(k,k+1)}(t'')$, which proves that $L_{(k,k+1)}$ increases only when $Y_k = Y_{k+1}$. With the aforementioned properties of $L_{(k,k+1)}$, we conclude that $L_{(k,k+1)}$ is the local time of collision between Y_k and Y_{k+1} , and hence that Y solves (1.16).

We now return to proving that X solves (1.2). This is done by taking the $m \rightarrow \infty$ limit of the finite system of equations (2.6) similarly to the way we did it for $Y^{(m)}$. However, unlike (3.26), the diffusion coefficients in (2.6) are generally discontinuous due the exchange of ranks. We resolve this problem following [Sar16a], by first showing:

Lemma 3.3. Define the following random set:

$$\mathcal{N} := \{t \in [0, T] \mid Y(t) \text{ has a tie}\} = \{t \in [0, T] \mid \exists k \geq 1 : Y_k(t) = Y_{k+1}(t)\}.$$

Then \mathbf{P} -a.s. the Lebesgue measure of \mathcal{N} is equal to zero.

Proof. As Y solves (1.16) (as proven above), the desired result follows once we show that the infinite system (1.16) can be reduced to a *finite system* at any given level $u \in \mathbb{R}$. More precisely, fixing arbitrary $u \in \mathbb{R}$ and $T \in \mathbb{R}_+$, we aim at showing

$$\inf_{0 \leq t \leq T} Y_k(t) < u, \text{ for only finitely many } k, \text{ a.s.} \tag{3.27}$$

Once this is established, the desired result follows by the same argument in [Sar16a, Lemma 3.9]. Turning to showing (3.27), because $Y_k(t) \leq Y_{k+1}(t)$ a.s., we need only to show that

$$\mathbf{P} \left(\inf_{0 \leq t \leq T} Y_k(t) < u \right) \rightarrow 0, \quad k \rightarrow \infty. \tag{3.28}$$

As $Y_k^{(m)} \rightarrow Y_k$ in $C[0, T]$, and the set $\{y(\cdot) \mid \inf_{0 \leq t \leq T} y(t) < u\}$ is open in $C[0, T]$, we have

$$\mathbf{P} \left(\inf_{0 \leq t \leq T} Y_k(t) < u \right) \leq \varliminf_{m \rightarrow \infty} \mathbf{P} \left(\inf_{0 \leq t \leq T} Y_k^{(m)}(t) < u \right). \tag{3.29}$$

Next, since $Y_k^{(m)}(t)$ is the k th ranked particle in $X^{(m)}(t)$, it follows that

$$Y_k^{(m)}(t) \geq \min_{j=k, \dots, m^2} X_j^{(m)}(t), \text{ so } \mathbf{P} \left(\inf_{0 \leq t \leq T} Y_k^{(m)}(t) < u \right) \leq \sum_{j=k}^{m^2} \mathbf{P} \left(\inf_{0 \leq t \leq T} X_j^{(m)}(t) < u \right).$$

Now, applying (3.2) to bound the r.h.s., we arrive at

$$\mathbf{P} \left(\inf_{0 \leq t \leq T} Y_k^{(m)}(t) < u \right) \leq c \sum_{j=k}^{m^2} (e^{-c(\log j - u)_+^2} + j^{-2} e^{cu}) \leq c \sum_{j=k}^{\infty} (e^{-c(\log j - u)_+^2} + j^{-2} e^{cu}). \tag{3.30}$$

The last term in (3.30) is independent of m and tends to zero as $k \rightarrow \infty$ (as explained previously after (3.23)). Consequently, inserting (3.30) into (3.29) yields the desired result (3.28). \square

With this Lemma 3.3, the rest of the proof follows by the same argument to the end of the proof of [Sar16a, Theorem 3.3], starting from [Sar16a, Lemma 6.5] up to the end of the proof of this theorem. That is, as $m \rightarrow \infty$, the solution $X_k^{(m)}$ of the finite system (2.1) converges to a solution of the infinite system (1.1).

3.2 Proof of Part (b)

Fix $t \in \mathbb{R}_+$ and an integer k . It follows from Proposition 2.2 that

$$\mathbf{E}(Y_k^{(m)}(t) - Y_k^{(m)}(0)) = -\frac{a}{2}t, \tag{3.31}$$

because $\bar{g}^{(m)} = -a/2$ for all m . Our goal is to pass (3.31) to the limit $m \rightarrow \infty$. To this end, since we already have $Y_k^{(m)}(s) \rightarrow Y_k(s)$ a.s. for $s = t$ and $s = 0$, it suffices to establish the L^2 -boundedness of $\{Y_k^{(m)}(t)\}_{m \geq 1}$:

$$\sup_{m \geq 1} \mathbf{E}(Y_k^{(m)}(t))^2 < \infty. \tag{3.32}$$

Indeed, (3.32) guarantees the uniform integrability of $\{Y_k^{(m)}(t)\}_{m \geq 1}$, so almost sure convergence implies convergence of expectations. Further, since

$$Y_k^{(m)}(t) = Y_1^{(m)}(t) + Z_{k-1}^{(m)}(t) + \dots + Z_1^{(m)}(t)$$

and $Z_k^{(m)}(t) \sim \text{Exp}(\lambda_k)$ for $k \leq m$, (3.31) clearly follows once we prove it for $k = 1$.

Proceeding to prove (3.31) for $k = 1$, we recall that \preceq denotes stochastic dominance. Apply comparison techniques from [Sar15, Corollary 3.7] to obtain that, for all $m \geq 1$, $Y_1^{(m)}(t) \preceq g_1 t + W(t)$, where W is some standard Brownian motion. From this it follows that

$$\sup_{m \geq 1} \mathbf{E}(Y_k^{(m)}(t)_+)^2 \leq \mathbf{E}((g_1 t + W(t))_+)^2 < \infty. \tag{3.33}$$

Next, to bound $\mathbf{E}(Y_k^{(m)}(t)_-)^2$, we fix $u \in \mathbb{R}_+$ and write

$$\mathbf{P}(Y_1^{(m)}(t) \leq -u) = \mathbf{P}(X_k^{(m)}(t) \leq -u) \leq \sum_{k=1}^{\infty} \mathbf{P}\left(\inf_{k \geq 1} X_k^{(m)}(t) \leq -u\right). \tag{3.34}$$

Combining this with (3.2), followed by using $e^{-c(u+\log k)^2} \leq e^{-cu^2 - c(\log k)^2}$, we obtain

$$\mathbf{P}(Y_1^{(m)}(t) \leq -u) \leq c \sum_{k=1}^{\infty} (e^{-cu^2 - c(\log k)^2} + k^{-2} e^{-cu}) \leq cS_1 e^{-cu^2} + cS_2 e^{-cu}, \tag{3.35}$$

$$\text{where } S_1 := \sum_{k=1}^{\infty} e^{-c(\log k)^2}, \quad S_2 := \sum_{k=1}^{\infty} k^{-2}. \tag{3.36}$$

As both these series in (3.36) converge, (3.35) shows that $Y_1^{(m)}(t)$ has an exponential lower tail which is uniformly in m , so in particular $\sup_{m \geq 1} \mathbf{E}(Y_1^{(m)}(t)_-)^2 < \infty$. Combining this with (3.33) yields the desired result (3.32).

A

Here we provide bounds on the number of particles under the measure π_a .

Proposition A.1. Fix g_1, g_2, \dots satisfying the condition (1.6) and fix $a \in \mathbb{R}$ satisfying (1.9). Let

$$0 = \xi_1 < \xi_2 < \xi_3 < \dots \in \mathbb{R}_+$$

be a random configuration of points with the gap distribution $(\zeta_k := \xi_{k+1} - \xi_k)_{k=1}^{\infty} \sim \pi_a$. Let $N(x) := \#\{i \geq 1 \mid \xi_i \leq x\}$ denote the random number of ξ_k -particles on the interval $[0, x]$. Then

$$0 < \liminf_{x \rightarrow \infty} e^{-ax} N(x) \leq \limsup_{x \rightarrow \infty} e^{-ax} N(x) < \infty \quad \text{a.s.}$$

Proof. Let $\lambda_k := 2(g_1 + \dots + g_k) + ka$. Under the conditions (1.6) and (1.9), we have

$$\sum_{k=1}^n \mathbf{E}\zeta_k = \sum_{k=1}^n \frac{1}{\lambda_k} = a^{-1} \log n + c_n, \tag{A.1}$$

where $(c_k)_{k \geq 1}$ is a bounded deterministic sequence, and $\sum_{k=1}^{\infty} \text{Var}(\zeta_k) = \sum_{k=1}^{\infty} \frac{1}{\lambda_k^2} < \infty$. With the last condition, [Str11, Theorem 1.4.2] implies that the series $\sum_{k=1}^{\infty} (\zeta_k - \mathbf{E}\zeta_k)$ converges a.s. Combining this with (A.1) yields

$$\sup_{n \geq 1} \left| \xi_n - a^{-1} \log n \right| < \infty \quad \text{a.s.},$$

which clearly implies the desired result. □

We next show that the measures π_a are all mutually singular for different values of a .

Proposition A.2. *Fixing g_1, g_2, \dots satisfying the condition (1.5) and $a > a' > -2 \inf_n \bar{g}_n$, we have that the measures π_a and $\pi_{a'}$ are mutually singular.*

Proof. Under the measure π_a , we have that $\frac{1}{2n\bar{g}_n+na}Z_n$, $n = 1, 2, \dots$, are i.i.d. $\text{Exp}(1)$ variables. For $Z \sim \text{Exp}(1)$, the variable $U := \mathbf{E}(\log Z)$ is integrable (i.e. $\mathbf{E}|U| < \infty$), so, letting $\mu := \mathbf{E}(U)$, by the strong Law of Large Numbers we that

$$\frac{1}{n} \sum_{k=1}^n \log \left(\frac{Z_k}{2k\bar{g}_k + ka} \right) \rightarrow \mu, \quad \pi_a\text{-a.s.}, \tag{A.2}$$

$$\frac{1}{n} \sum_{k=1}^n \log \left(\frac{Z_k}{2k\bar{g}_k + ka'} \right) \rightarrow \mu, \quad \pi_{a'}\text{-a.s.} \tag{A.3}$$

Under the conditions (1.5) and $a, a' > -2 \inf_n g_n$, it is straightforward to show that

$$\frac{1}{n} \sum_{k=1}^n \log \left(\frac{2k\bar{g}_k + ka'}{2k\bar{g}_k + ka} \right) \rightarrow \log \left(\frac{a'}{a} \right).$$

Adding this to (A.3) yields

$$\frac{1}{n} \sum_{k=1}^n \log \left(\frac{Z_k}{2k\bar{g}_k + ka} \right) \rightarrow \mu + \log \left(\frac{a'}{a} \right), \quad \pi_{a'}\text{-a.s.}$$

This, with $a \neq a'$, concludes that π_a and $\pi_{a'}$ are mutually singular. □

References

- [Ald03] David Aldous (2003). Unpublished, available at <http://www.stat.berkeley.edu/~aldous/Research/OP/river.pdf>
- [AA09] Louis-Pierre Arguin, Michael Aizenman (2009). On the Structure of Quasi-Stationary Competing Particle Systems. *Ann. Probab.* **37** (3), 1080–1113. MR-2537550
- [BFK05] Adrian D. Banner, E. Robert Fernholz, Ioannis Karatzas (2005). Atlas Models of Equity Markets. *Ann. Appl. Probab.* **15** (4), 2996–2330. MR-2187296
- [BG08] Adrian D. Banner, Raouf Ghomrasni (2008). Local Times of Ranked Continuous Semimartingales. *Stoch. Proc. Appl.* **118** (7), 1244–1253. MR-2428716
- [BFIKP11] Adrian D. Banner, E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas, Vassilios Papathanakos (2011). Hybrid Atlas Models. *Ann. Appl. Probab.* **21** (2), 609–644. MR-2807968
- [BP87] Richard Bass, E. Pardoux (1987). Uniqueness for Diffusions with Piecewise Constant Coefficients. *Probab. Th. Rel. Fields* **76**, 557–572. MR-0917679
- [BS15] Cameron Bruggeman, Andrey Sarantsev (2015). Multiple Collisions in Systems of Competing Brownian Particles. To appear in *Bernoulli*. Available at arXiv:1309.2621.
- [CP10] Sourav Chatterjee, Soumik Pal (2010). A Phase-Transition Behavior for Brownian Motions Interacting Through Their Ranks. *Probab. Th. Rel. Fields* **147** (1–2), 123–159. MR-2594349
- [DSVZ16] Amir Dembo, Mykhaylo Shkolnikov, S.R. Srinivasa Varadhan, Ofer Zeitouni (2016). Large Deviations for Diffusions Interacting Through Their Ranks. *Comm. Pure Appl. Math.* **69** (7), 1259–1313. MR-3503022
- [DT15] Amir Dembo, Li-Cheng Tsai (2015). Equilibrium Fluctuations of the Atlas Model. To appear in *Ann. Probab.* Available at arXiv:1503.03581.
- [Fel68] William Feller (1968). *An Introduction to Probability Theory and Its Applications*. Vol. 1, Wiley. MR-0228020

Stationary gap distributions

- [FIK13] E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas (2013). A Second-Order Stock Market Model. *Ann. Finance* **9** (3), 439–454. MR-3082660
- [FIKP13] E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas, Vilmos Prokaj (2013). Planar Diffusions with Rank-Based Characteristics and Perturbed Tanaka Equations. *Probab. Th. Rel. Fields*, **156** (1–2), 343–374. MR-3055262
- [FIK13] E. Robert Fernholz, Tomoyuki Ichiba, Ioannis Karatzas (2013). Two Brownian Particles with Rank-Based Characteristics and Skew-Elastic Collisions. *Stoch. Proc. Appl.* **123** (8), 2999–3026. MR-3062434
- [FK09] E. Robert Fernholz, Ioannis Karatzas (2009) Stochastic Portfolio Theory: An Overview. *Handbook of Numerical Analysis: Mathematical Modeling and Numerical Methods in Finance*, 89–168. Elsevier. MR-1894767
- [IK10] Tomoyuki Ichiba, Ioannis Karatzas (2010). On Collisions of Brownian Particles. *Ann. Appl. Probab.* **20** (3), 951–977. MR-2680554
- [IKP13] Tomoyuki Ichiba, Ioannis Karatzas, Vilmos Prokaj (2013). Diffusions with Rank-Based Characteristics and Values in the Nonnegative Quadrant. *Bernoulli* **19** (5B), 2455–2493. MR-3160561
- [IKS13] Tomoyuki Ichiba, Ioannis Karatzas, Mykhaylo Shkolnikov (2013). Strong Solutions of Stochastic Equations with Rank-Based Coefficients. *Probab. Th. Rel. Fields* **156**, 229–248. MR-3055258
- [IPS13] Tomoyuki Ichiba, Soumik Pal, Mykhaylo Shkolnikov (2013). Convergence Rates for Rank-Based Models with Applications to Portfolio Theory. *Probab. Th. Rel. Fields* **156**, 415–448. MR-3055264
- [JM08] Benjamin Jourdain, Florent Malrieu (2008). Propagation of Chaos and Poincare Inequalities for a System of Particles Interacting Through Their cdf. *Ann. Appl. Probab.* **18** (5), 1706–1736. MR-2462546
- [JR13] Benjamin Jourdain, Julien Reygner (2013). Propagation of Chaos for Rank-Based Interacting Diffusions and Long Time Behaviour of a Scalar Quasilinear Parabolic Equation. *SPDE Anal. Comp.* **1** (3), 455–506. MR-3327514
- [JR14] Benjamin Jourdain, Julien Reygner (2014). The Small Noise Limit of Order-Based Diffusion Processes. *Electr. J. Probab.* **19** (29), 1–36. MR-3174841
- [JR15] Benjamin Jourdain, Julien Reygner (2015). Capital Distribution and Portfolio Performance in the Mean-Field Atlas Model. *Ann. Finance* **11** (2), 151–198. MR-3340241
- [KPS12] Ioannis Karatzas, Soumik Pal, Mykhaylo Shkolnikov (2016). Systems of Brownian Particles with Asymmetric Collisions. *Ann. Inst. H. Poincare* **52** (1), 323–354. MR-3449305
- [KS16] Ioannis Karatzas, Andrey Sarantsev (2016). Diverse Market Models of Competing Brownian Particles with Splits and Mergers. *Ann. Appl. Probab.* **26** (3), 1329–1361. MR-3513592
- [PP08] Soumik Pal, Jim Pitman (2008). One-Dimensional Brownian Particle Systems with Rank-Dependent Drifts. *Ann. Appl. Probab.* **18** (6), 2179–2207. MR-2473654
- [Rey15] Julien Reygner (2015). Chaoticity of the Stationary Distribution of Rank-Based Interacting Diffusions. *Electr. Comm. Probab.* **20** (60), 1–20. MR-3399811
- [RA05] Anastasia Ruzmaikina, Michael Aizenman (2005). Characterization of Invariant Measures at the Leading Edge for Competing Particle Systems. *Ann. Probab.* **33** (1), 82–113. MR-2118860
- [Sar15] Andrey Sarantsev (2015). Comparison Techniques for Competing Brownian Particles. To appear in *J. Th. Probab.* Available at arXiv:1305.1653.
- [Sar15a] Andrey Sarantsev (2015). Reflected Brownian Motion in a Convex Polyhedral Cone: Tail Estimates for the Stationary Distribution. To appear in *J. Th. Probab.* Available at arXiv:1509.01781.
- [Sar15b] Andrey Sarantsev (2015). Triple and Simultaneous Collisions of Competing Brownian Particles. *Electr. J. Probab.* **20** (29), 1–28. MR-3325099
- [Sar15c] Andrey Sarantsev (2015). Two-Sided Infinite Systems of Competing Brownian Particles. To appear in *ESAIM P&S*. Available at arXiv:1509.01859.

Stationary gap distributions

- [Sar16] Andrey Sarantsev (2016). Explicit Rates of Exponential Convergence for Reflected Jump-Diffusions on the Half-Line. *ALEA Lat. Am. J. Probab. Math. Stat.* **16** (2), 1069–1093. MR-3580089
- [Sar16a] Andrey Sarantsev (2016). Infinite Systems of Competing Brownian Particles. To appear in *Ann. Inst. H. Poincare*. Available at arXiv:1403.4229. MR-3438889
- [Shk11] Mykhaylo Shkolnikov (2011). Competing Particle Systems Evolving by Interacting Lévy Processes. *Ann. Appl. Probab.* **21** (5), 1911–1932. MR-2884054
- [Shk12] Mykhaylo Shkolnikov (2012). Large Systems of Diffusions Interacting Through Their Ranks. *Stoch. Proc. Appl.* **122** (4), 1730–1747. MR-2914770
- [Str11] Daniel W. Stroock (2011). *Probability Theory. An Analytic View*. Cambridge University Press. MR-2760872
- [Tsa17] Li-Cheng Tsai (2017). Stationary Distributions of the Atlas Model. Available at arXiv:1702.02043.
- [Wil95] Ruth J. Williams (1995). *Semimartingale Reflecting Brownian Motions in the Orthant*. *Stochastic networks*, IMA Vol. Math. Appl. **71**, 125–137. Springer-Verlag. MR-1381009