# Stochastic differential equations with a fractionally filtered delay: A semimartingale model for long-range dependent processes

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In this paper, we introduce a model, the stochastic fractional delay differential equation (SFDDE), which is based on the linear stochastic delay differential equation and produces stationary processes with hyperbolically decaying autocovariance functions. The model departs from the usual way of incorporating this type of long-range dependence into a short-memory model as it is obtained by applying a fractional filter to the drift term rather than to the noise term. The advantages of this approach are that the corresponding long-range dependent solutions are semimartingales and the local behavior of the sample paths is unaffected by the degree of long memory. We prove existence and uniqueness of solutions to the SFDDEs and study their spectral densities and autocovariance functions. Moreover, we define a subclass of SFDDEs which we study in detail and relate to the well-known fractionally integrated CARMA processes. Finally, we consider the task of simulating from the defining SFDDEs.

Keywords: long-range dependence; moving average processes; semimartingales; stochastic differential equations

### 1. Introduction

Models for time series producing slowly decaying autocorrelation functions (ACFs) have been of interest for more than 50 years. Such models were motivated by the empirical findings of Hurst in the 1950s that were related to the levels of the Nile River. Later, in the 1960s, Benoit Mandelbrot referred to a slowly decaying ACF as the Joseph effect or long-range dependence. Since then, a vast amount of literature on theoretical results and applications have been developed. We refer to [6,13,25,28,29] and references therein for further background.

A very popular discrete-time model for long-range dependence is the *autoregressive fractionally integrated moving average* (ARFIMA) process, introduced by Granger and Joyeux [15] and Hosking [19], which extends the ARMA process to allow for a hyperbolically decaying ACF. Let B be the backward shift operator and for  $\gamma > -1$ , define  $(1 - B)^{\gamma}$  by means of the binomial expansion,

$$(1-B)^{\gamma} = \sum_{j=0}^{\infty} \pi_j B^j$$

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where  $\pi_j = \prod_{0 < k \le j} \frac{k-1-\gamma}{k}$ . An ARFIMA process  $(X_t)_{t \in \mathbb{Z}}$  is characterized as the unique purely non-deterministic process (as defined in [9], page 189) satisfying

$$P(B)(1-B)^{\beta}X_t = Q(B)\varepsilon_t, \quad t \in \mathbb{Z}, \tag{1}$$

where P and Q are real polynomials with no zeroes on  $\{z \in \mathbb{C} : |z| \leq 1\}$ ,  $(\varepsilon_t)_{t \in \mathbb{Z}}$  is an i.i.d. sequence with  $\mathbb{E}[\varepsilon_0] = 0$ ,  $\mathbb{E}[\varepsilon_0^2] \in (0, \infty)$  and  $\beta \in (0, 1/2)$ . The ARFIMA equation (1) is sometimes represented as an ARMA equation with a fractionally integrated noise, that is,

$$P(B)X_t = Q(B)(1-B)^{-\beta}\varepsilon_t, \quad t \in \mathbb{Z}.$$
 (2)

In (1) one applies a fractional filter to  $(X_t)_{t\in\mathbb{Z}}$ , while in (2) one applies a fractional filter to  $(\varepsilon_t)_{t\in\mathbb{Z}}$ . One main feature of the solution to (1), equivalently (2), is that the autocovariance function  $\gamma_X(t) := \mathbb{E}[X_0X_t]$  satisfies

$$\gamma_X(t) \sim ct^{2\beta - 1}, \quad t \to \infty,$$
 (3)

for some constant c > 0.

A simple example of a continuous-time stationary process which exhibits long-memory in the sense of (3) is an Ornstein–Uhlenbeck process  $(X_t)_{t\in\mathbb{R}}$  driven by a fractional Lévy process, that is,  $(X_t)_{t\in\mathbb{R}}$  is the unique stationary solution to

$$dX_t = -\kappa X_t dt + dI^{\beta} L_t, \quad t \in \mathbb{R}, \tag{4}$$

where  $\kappa > 0$  and

$$I^{\beta}L_{t} := \frac{1}{\Gamma(1+\beta)} \int_{-\infty}^{t} \left[ (t-u)^{\beta} - (-u)_{+}^{\beta} \right] dL_{u}, \quad t \in \mathbb{R},$$
 (5)

with  $(L_t)_{t\in\mathbb{R}}$  being a Lévy process which satisfies  $\mathbb{E}[L_1] = 0$  and  $\mathbb{E}[L_1^2] \in (0, \infty)$ . In (5),  $\Gamma$  denotes the gamma function and we have used the notation  $x_+ = \max\{x, 0\}$  for  $x \in \mathbb{R}$ . The way to obtain long memory in (4) is by applying a fractional filter to the noise, which is in line with (2). To demonstrate the idea of this paper, consider the equation obtained from (4) but by applying a fractional filter to the drift term instead, that is,

$$X_t - X_s = -\frac{\kappa}{\Gamma(1-\beta)} \int_{-\infty}^t \left[ (t-u)^{-\beta} - (s-u)_+^{-\beta} \right] X_u \, \mathrm{d}u + L_t - L_s, \quad s < t.$$
 (6)

One can write (6) compactly as

$$dX_t = -\kappa D^{\beta} X_t dt + dL_t, \quad t \in \mathbb{R}, \tag{7}$$

with  $(D^{\beta}X_t)_{t\in\mathbb{R}}$  being a suitable fractional derivative process of  $(X_t)_{t\in\mathbb{R}}$  defined in Proposition 3.6. The equations (6)–(7) are akin to (1). It turns out that a unique purely non-deterministic process (as defined in (23)) satisfying (7) exists and has the following properties:

(i) The memory is long and controlled by  $\beta$  in the sense that  $\gamma_X(t) \sim ct^{2\beta-1}$  as  $t \to \infty$  for some c > 0.

- (ii) The  $L^2(\mathbb{P})$  Hölder continuity of the sample paths is not affected by  $\beta$  in the sense that  $\gamma_X(0) \gamma_X(t) \sim ct$  as  $t \downarrow 0$  for some c > 0 (the notion of Hölder continuity in  $L^2(\mathbb{P})$  is indeed closely related to the behavior of the ACF at zero; see Remark 3.9 for a precise relation).
- (iii)  $(X_t)_{t \in \mathbb{R}}$  is a semimartingale.

While both processes in (4) and (7) exhibit long memory in the sense of (i), one should keep in mind that models for long-memory processes obtained by applying a fractional filter to the noise will generally not meet (ii)–(iii), since they inherit various properties from the fractional Lévy process  $(I^{\beta}L_t)_{t\in\mathbb{R}}$  rather than from the underlying Lévy process  $(L_t)_{t\in\mathbb{R}}$ . In particular, this observation applies to the fractional Ornstein–Uhlenbeck process (4) which is known not to possess the semimartingale property for many choices of  $(L_t)_{t\in\mathbb{R}}$ , and for which it holds that  $\gamma_X(0) - \gamma_X(t) \sim ct^{2\beta+1}$  as  $t \downarrow 0$  for some c > 0 (see [21], Theorem 4.7, and [1], Proposition 2.5). The latter property, the behavior of  $\gamma_X$  near 0, implies an increased  $L^2(\mathbb{P})$  Hölder continuity relative to (7). See Example 4.4 for details about the models (4) and (7).

The properties (ii)–(iii) may be desirable to retain in many modeling scenarios. For instance, if a stochastic process  $(X_t)_{t\in\mathbb{R}}$  is used to model a financial asset, the semimartingale property is necessary to accommodate the No Free Lunch with Vanishing Risk condition according to the (First) Fundamental Theorem of Asset Pricing, see [11], Theorem 7.2. Moreover, if  $(X_t)_{t\in\mathbb{R}}$  is supposed to serve as a "good" integrator, it follows by the Bichteler–Dellacherie theorem ([7], Theorem 7.6) that  $(X_t)_{t\in\mathbb{R}}$  must be a semimartingale. Also, the papers [4,5] find evidence that the sample paths of electricity spot prices and intraday volatility of the E-mini S&P500 futures contract are rough, and Jusselin and Rosenbaum [20] show that the no-arbitrage assumption implies that the volatility of the macroscopic price process is rough. These findings suggest less smooth sample paths than what is induced by models such as the fractional Ornstein–Uhlenbeck process (4). In particular, the local smoothness of the sample paths should not be connected to the strength of long memory.

Several extensions to the fractional Ornstein–Uhlenbeck process (4) exist. For example, it is worth mentioning that the class of *fractionally integrated continuous-time autoregressive moving average* (FICARMA) processes were introduced in Brockwell and Marquardt [8], where it is assumed that P and Q are real polynomials with  $\deg(P) > \deg(Q)$  which have no zeroes on  $\{z \in \mathbb{C} : \operatorname{Re}(z) \geq 0\}$ . The FICARMA process associated to P and Q is then defined as the moving average process

$$X_t = \int_{-\infty}^t g(t - u) \, \mathrm{d}I^{\beta} L_u, \quad t \in \mathbb{R},$$
 (8)

with  $g:\mathbb{R}\to\mathbb{R}$  being the  $L^1$  function characterized by

$$\mathcal{F}[g](y) := \int_{\mathbb{R}} e^{-iyu} g(u) \, \mathrm{d}u = \frac{Q(iy)}{P(iy)}, \quad y \in \mathbb{R}.$$

In line with (2) for the ARFIMA process, a common way of viewing a FICARMA process is that it is obtained by applying a CARMA filter to fractional noise, that is,  $(X_t)_{t \in \mathbb{R}}$  given by (8) is the solution to the formal equation

$$P(D)X_t = O(D)DI^{\beta}L_t, \quad t \in \mathbb{R}.$$

(See, e.g., [21].) Another class, related to the FICARMA process, consists of solutions  $(X_t)_{t \in \mathbb{R}}$  to fractional *stochastic delay differential equations* (SDDEs), that is,  $(X_t)_{t \in \mathbb{R}}$  is the unique stationary solution to

$$dX_t = \int_{[0,\infty)} X_{t-u} \eta(du) dt + dI^{\beta} L_t, \quad t \in \mathbb{R},$$
(9)

for a suitable finite signed measure  $\eta$ . For details about these processes, see [2,22]. Note that the fractional Ornstein–Uhlenbeck process (4) is a FICARMA process with polynomials  $P(z) = z + \kappa$  and Q(z) = 1 and a fractional SDDE with  $\eta = -\kappa \delta_0$ ,  $\delta_0$  being the Dirac measure at zero.

The model we present includes (6) and extends this process in the same way as the fractional SDDE (9) extends the fractional Ornstein–Uhlenbeck (4). Specifically, we will be interested in a stationary process  $(X_t)_{t\in\mathbb{R}}$  satisfying

$$X_t - X_s = \int_{-\infty}^t \left( D_-^{\beta} \mathbb{1}_{(s,t]} \right) (u) \int_{[0,\infty)} X_{u-v} \eta(\mathrm{d}v) \, \mathrm{d}u + L_t - L_s$$
 (10)

almost surely for each s < t, where  $\eta$  is a given finite signed measure and

$$(D_{-}^{\beta} \mathbb{1}_{(s,t]})(u) = \frac{1}{\Gamma(1-\beta)} [(t-u)_{+}^{-\beta} - (s-u)_{+}^{-\beta}], \quad u \in \mathbb{R}.$$

We will refer to (10) as a *stochastic fractional delay differential equation* (SFDDE). Equation (10) can be compactly written as

$$dX_t = \int_{[0,\infty)} D^{\beta} X_{t-u} \eta(du) dt + dL_t, \quad t \in \mathbb{R},$$
(11)

with  $(D^{\beta}X_t)_{t\in\mathbb{R}}$  defined in Proposition 3.6. The representation (11) is, for instance, convenient in order to argue that solutions are semimartingales.

In Section 3, we show that, for a wide range of measures  $\eta$ , there exists a unique purely nondeterministic process  $(X_t)_{t \in \mathbb{R}}$  satisfying the SFDDE (10). In addition, we study the behavior of the autocovariance function and the spectral density of  $(X_t)_{t \in \mathbb{R}}$  and verify that (i)–(ii) hold. We end Section 3 by providing an explicit (prediction) formula for computing  $\mathbb{E}[X_t \mid X_u, u \leq s]$ . In Section 4, we focus on delay measures  $\eta$  of exponential type, that is,

$$\eta(\mathrm{d}t) = -\kappa \,\delta_0(\mathrm{d}t) + f(t)\,\mathrm{d}t,\tag{12}$$

where  $f(t) = \mathbb{1}_{[0,\infty)}(t)b^{\top}e^{At}e_1$  with  $e_1 = [1,0,\ldots,0]^{\top} \in \mathbb{R}^n$ ,  $b \in \mathbb{R}^n$  and A an  $n \times n$  matrix with a spectrum contained in  $\{z \in \mathbb{C} : \text{Re}(z) < 0\}$ . Besides relating this subclass to the FICARMA processes we study two special cases of (12) in detail, namely the Ornstein–Uhlenbeck type presented in (7) and

$$dX_t = \int_0^\infty D^{\beta} X_{t-u} f(u) du dt + dL_t, \quad t \in \mathbb{R}.$$
 (13)

Equation (13) is interesting to study as it collapses to an ordinary SDDE (cf. Proposition 4.2), and hence constitutes an example of a long-range dependent solution to equation (9) with  $I^{\beta}L_t$  –

 $I^{\beta}L_s$  replaced by  $L_t - L_s$ . While (13) falls into the overall setup of [3], the results obtained in that paper do, however, not apply. Finally, based on the two examples (6) and (13), we investigate some numerical aspects in Section 5, including the task of simulating  $(X_t)_{t \in \mathbb{R}}$  from the defining equation. Section 6 contains the proofs of all the results presented in Sections 3 and 4. We start with a preliminary section which recalls a few definitions and results that will be used repeatedly.

### 2. Preliminaries

For a measure  $\mu$  on the Borel  $\sigma$ -field  $\mathcal{B}(\mathbb{R})$  on  $\mathbb{R}$ , let  $L^p(\mu)$  denote the  $L^p$  space relative to  $\mu$ . If  $\mu$  is the Lebesgue measure, we suppress the dependence on  $\mu$  and write  $L^p$  instead of  $L^p(\mu)$ . By a finite signed measure, we refer to a set function  $\mu \colon \mathcal{B}(\mathbb{R}) \to \mathbb{R}$  of the form  $\mu = \mu^+ - \mu^-$ , where  $\mu^+$  and  $\mu^-$  are two finite and mutually singular measures. Integration of a function f with respect to  $\mu$  is defined (in an obvious way) whenever  $f \in L^1(|\mu|)$  where  $|\mu| := \mu^+ + \mu^-$ . The convolution of two measurable functions  $f, g \colon \mathbb{R} \to \mathbb{C}$  is defined as

$$f * g(t) = \int_{\mathbb{R}} f(t - u)g(u) \, \mathrm{d}u$$

whenever  $f(t - \cdot)g \in L^1$ . Similarly, if  $\mu$  is a finite signed measure, we set

$$f * \mu(t) = \int_{\mathbb{R}} f(t - u)\mu(\mathrm{d}u)$$

if  $f(t - \cdot) \in L^1(|\mu|)$ . For such  $\mu$ , set

$$D(\mu) = \left\{ z \in \mathbb{C} : \int_{\mathbb{R}} e^{-\operatorname{Re}(z)u} |\mu| (\mathrm{d}u) < \infty \right\}.$$

Then we define the bilateral Laplace transform  $\mathcal{L}[\mu]: D(\mu) \to \mathbb{C}$  of  $\mu$  by

$$\mathcal{L}[\mu](z) = \int_{\mathbb{R}} e^{-zu} \mu(\mathrm{d}u), \quad z \in D(\mu),$$

and the Fourier transform by  $\mathcal{F}[\mu](y) = \mathcal{L}[f](iy)$  for  $y \in \mathbb{R}$ . If  $f \in L^1$  we will write  $\mathcal{L}[f] = \mathcal{L}[f(u) \, \mathrm{d}u]$  and  $\mathcal{F}[f] = \mathcal{F}[f(u) \, \mathrm{d}u]$ . We also note that  $\mathcal{F}[f] \in L^2$  when  $f \in L^1 \cap L^2$  and that  $\mathcal{F}$  can be extended to an isometric isomorphism from  $L^2$  onto  $L^2$  by Plancherel's theorem.

Recall that a Lévy process is the continuous-time analogue to the (discrete-time) random walk. More precisely, a one-sided Lévy process  $(L_t)_{t\geq 0}, L_0=0$ , is a stochastic process having stationary independent increments and càdlàg sample paths. From these properties, it follows that the distribution of  $L_1$  is infinitely divisible, and the distribution of  $(L_t)_{t\geq 0}$  is determined from  $L_1$  via the relation  $\mathbb{E}[e^{iyL_t}]=\exp\{t\log\mathbb{E}[e^{iyL_1}]\}$  for  $y\in\mathbb{R}$  and  $t\geq 0$ . The definition is extended to a two-sided Lévy process  $(L_t)_{t\in\mathbb{R}}$  by taking a one-sided Lévy process  $(L_t)_{t\geq 0}$  together with an independent copy  $(L_t^2)_{t\geq 0}$  and setting  $L_t=L_t^1$  if  $t\geq 0$  and  $L_t=-L_{(-t)}^2$  if t<0. If  $\mathbb{E}[L_1^2]<\infty$ ,

 $\mathbb{E}[L_1] = 0$  and  $f \in L^2$ , the integral  $\int_{\mathbb{R}} f(u) dL_u$  is well-defined as an  $L^2$  limit of integrals of step functions, and the following isometry property holds:

$$\mathbb{E}\left[\left(\int_{\mathbb{R}} f(u) dL_u\right)^2\right] = \mathbb{E}\left[L_1^2\right] \int_{\mathbb{R}} f(u)^2 du.$$

For more on Lévy processes and integrals with respect to these, see [26,31]. Finally, for two functions  $f, g: \mathbb{R} \to \mathbb{C}$  and  $a \in [-\infty, \infty]$  we write f(t) = o(g(t)), f(t) = O(g(t)) and  $f(t) \sim g(t)$  as  $t \to a$  if

$$\lim_{t \to a} \frac{f(t)}{g(t)} = 0, \qquad \limsup_{t \to a} \left| \frac{f(t)}{g(t)} \right| < \infty \quad \text{and} \quad \lim_{t \to a} \frac{f(t)}{g(t)} = 1,$$

respectively.

## 3. The stochastic fractional delay differential equation

Let  $(L_t)_{t\in\mathbb{R}}$  be a Lévy process with  $\mathbb{E}[L_1^2]<\infty$  and  $\mathbb{E}[L_1]=0$ , and let  $\beta\in(0,1/2)$ . Without loss of generality we will assume that  $\mathbb{E}[L_1^2]=1$ . Moreover, denote by  $\eta$  a finite (possibly signed) measure on  $[0,\infty)$  with

$$\int_{[0,\infty)} t |\eta| (\mathrm{d}t) < \infty \tag{14}$$

and set

$$(D_{-}^{\beta} \mathbb{1}_{(s,t]})(u) = \frac{1}{\Gamma(1-\beta)} [(t-u)_{+}^{-\beta} - (s-u)_{+}^{-\beta}], \quad u \in \mathbb{R}.$$
 (15)

(In line with [13] we write  $D_{-}^{\beta}\mathbb{1}_{(s,t]}$  rather than  $D^{\beta}\mathbb{1}_{(s,t]}$  in (15) to emphasize that it is the right-sided version of the Riemann–Liouville fractional derivative of  $\mathbb{1}_{(s,t]}$ .) Then we will say that a process  $(X_t)_{t\in\mathbb{R}}$  with  $\mathbb{E}[|X_0|]<\infty$  is a solution to the corresponding SFDDE if it is stationary and satisfies

$$X_t - X_s = \int_{-\infty}^t \left( D_-^{\beta} \mathbb{1}_{(s,t]} \right) (u) \int_{[0,\infty)} X_{u-v} \eta(\mathrm{d}v) \, \mathrm{d}u + L_t - L_s$$
 (16)

almost surely for each s < t. Note that (16) is indeed well-defined, since  $\eta$  is finite,  $(X_t)_{t \in \mathbb{R}}$  is bounded in  $L^1(\mathbb{P})$  and  $D^{\beta}_{-}\mathbb{1}_{(s,t]} \in L^1$ . As noted in the introduction, we will often write (16) shortly as

$$dX_t = \int_{[0,\infty)} D^{\beta} X_{t-u} \eta(du) dt + dL_t, \quad t \in \mathbb{R},$$
(17)

where  $(D^{\beta}X_t)_{t\in\mathbb{R}}$  is a suitable fractional derivative of  $(X_t)_{t\in\mathbb{R}}$  (defined in Proposition 3.6).

In order to study which choices of  $\eta$  lead to a stationary solution to (16) we introduce the function  $h = h_{\beta,\eta}$ :  $\{z \in \mathbb{C} : \text{Re}(z) \ge 0\} \to \mathbb{C}$  given by

$$h(z) = z^{1-\beta} - \int_{[0,\infty)} e^{-zu} \eta(du), \quad \text{Re}(z) \ge 0.$$
 (18)

Here, and in the following, we define  $z^{\gamma} = r^{\gamma} e^{i\gamma\theta}$  using the polar representation  $z = re^{i\theta}$  for r > 0 and  $\theta \in (-\pi, \pi]$ . This definition corresponds to  $z^{\gamma} = e^{\gamma \log z}$ , using the principal branch of the complex logarithm, and hence  $z \mapsto z^{\gamma}$  is analytic on  $\mathbb{C} \setminus \{z \in \mathbb{R} : z \leq 0\}$ . In particular, this means that h is analytic on  $\{z \in \mathbb{C} : \text{Re}(z) > 0\}$ .

**Proposition 3.1.** Suppose that h(z) defined in (18) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$ . Then there exists a unique  $g: \mathbb{R} \to \mathbb{R}$ , which belongs to  $L^{\gamma}$  for  $(1 - \beta)^{-1} < \gamma \leq 2$  and is vanishing on  $(-\infty, 0)$ , such that

$$\mathcal{F}[g](y) = \frac{(iy)^{-\beta}}{h(iy)}, \quad y \in \mathbb{R}.$$
 (19)

Moreover, the following statements hold:

(i) For t > 0 the Marchaud fractional derivative  $D^{\beta}g(t)$  at t of g given by

$$D^{\beta}g(t) = \frac{\beta}{\Gamma(1-\beta)} \lim_{\delta \downarrow 0} \int_{\delta}^{\infty} \frac{g(t) - g(t-u)}{u^{1+\beta}} du$$
 (20)

exists,  $D^{\beta}g \in L^1 \cap L^2$  and  $\mathcal{F}[D^{\beta}g](y) = 1/h(iy)$  for  $y \in \mathbb{R}$ .

(ii) The function g is the Riemann–Liouville fractional integral of  $D^{\beta}g$ , that is,

$$g(t) = \frac{1}{\Gamma(\beta)} \int_0^t D^{\beta} g(u) (t - u)^{\beta - 1} du, \quad t > 0.$$

(iii) The function g satisfies

$$g(t) = 1 + \int_0^t (D^{\beta} g) * \eta(u) du, \quad t \ge 0,$$
 (21)

and for  $v \in \mathbb{R}$  and with  $D_{-}^{\beta} \mathbb{1}_{(s,t]}$  given in (15),

$$g(t-v) - g(s-v) = \int_{-\infty}^{t} \left( D_{-}^{\beta} \mathbb{1}_{(s,t]} \right) (u) g * \eta(u-v) \, \mathrm{d}u + \mathbb{1}_{(s,t]}(v). \tag{22}$$

Before formulating our main result, Theorem 3.2, recall that a stationary process  $(X_t)_{t \in \mathbb{R}}$  with  $\mathbb{E}[X_0^2] < \infty$  and  $\mathbb{E}[X_0] = 0$  is said to be purely non-deterministic if

$$\bigcap_{t \in \mathbb{R}} \overline{\operatorname{sp}}\{X_s : s \le t\} = \{0\},\tag{23}$$

see [1], Section 4. Here  $\overline{\mathrm{sp}}$  denotes the  $L^2(\mathbb{P})$ -closure of the linear span.

**Theorem 3.2.** Suppose that h(z) defined in (18) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$  and let g be the function introduced in Proposition 3.1. Then the process

$$X_t = \int_{-\infty}^t g(t - u) \, \mathrm{d}L_u, \quad t \in \mathbb{R}, \tag{24}$$

is well-defined, centered and square integrable, and it is the unique purely non-deterministic solution to the SFDDE (16).

**Remark 3.3.** Note that we cannot hope to get a uniqueness result without imposing a condition such as (23). For instance, the fact that

$$\int_{-\infty}^{t} \left[ (t-u)^{-\beta} - (s-u)_{+}^{-\beta} \right] du = 0,$$

shows together with (16) that  $(X_t + U)_{t \in \mathbb{R}}$  is a solution for any  $U \in L^1(\mathbb{P})$  as long as  $(X_t)_{t \in \mathbb{R}}$  is a solution. Moreover, uniqueness relative to condition (23) is similar to that of discrete-time ARFIMA processes, see [9], Theorem 13.2.1.

**Remark 3.4.** It is possible to generalize (16) and Theorem 3.2 to allow for a heavy-tailed distribution of the noise. Specifically, suppose that  $(L_t)_{t \in \mathbb{R}}$  is a symmetric  $\alpha$ -stable Lévy process for some  $\alpha \in (1, 2)$ , that is,  $(L_t)_{t \in \mathbb{R}}$  is a Lévy process and

$$\mathbb{E}[e^{iyL_1}] = e^{-\sigma^{\alpha}|y|^{\alpha}}, \quad y \in \mathbb{R},$$

for some  $\sigma > 0$ . To define the process  $(X_t)_{t \in \mathbb{R}}$  in (24) it is necessary and sufficient that  $g \in L^{\alpha}$ , which is indeed the case if  $\beta \in (1, 1 - 1/\alpha)$  by Proposition 3.1. From this point, using (22), we only need a stochastic Fubini result (which can be found in [1], Theorem 3.1) to verify that (16) is satisfied. One will need another notion (and proof) of uniqueness, however, as our approach relies on  $L^2$  theory. For more on stable distributions and corresponding definitions and results, we refer to [30].

**Remark 3.5.** The process (24) and other well-known long-memory processes do naturally share parts of their construction. For instance, they are typically viewed as "borderline" stationary solutions to certain equations. To be more concrete, the ARFIMA process can be viewed as an ARMA process, but where the autoregressive polynomial P is replaced by  $\tilde{P}: z \mapsto P(z)(1-z)^{\beta}$ . Although an ordinary ARMA process exists if and only if P is non-zero on the unit circle (and, in the positive case, will be a short memory process), the autoregressive function  $\tilde{P}$  of the ARFIMA model will always have a root at z = 1. The analogue to the autoregressive polynomial in the non-fractional SDDE model (that is, (16) with  $D^{\beta}_{-1}\mathbb{I}_{(s,t]}$  replaced by  $\mathbb{I}_{(s,t]}$ ) is

$$z \longmapsto z - \mathcal{L}[\eta](z),$$
 (25)

where the critical region is on the imaginary axis  $\{iy : y \in \mathbb{R}\}$  rather than on the unit circle  $\{z \in \mathbb{C} : |z| = 1\}$  (see [2]). The SFDDE corresponds to replacing (25) by  $z \mapsto z - z^{\beta} \mathcal{L}[\eta](z)$ ,

which will always have a root at z = 0. However, to ensure existence both in the ARFIMA model and in the SFDDE model, assumptions are made such that these roots will be the only ones in the critical region and their order will be  $\beta$ . For a treatment of ARFIMA processes, we refer to [9], Section 13.2.

The solution  $(X_t)_{t\in\mathbb{R}}$  of Theorem 3.2 is causal in the sense that  $X_t$  only depends on past increments of the noise  $L_t - L_s$ ,  $s \le t$ . An inspection of the proof of Theorem 3.2 reveals that one only needs to require that  $h(iy) \ne 0$  for all  $y \in \mathbb{R}$  for a (possibly non-causal) stationary solution to exist. The difference between the condition that h(z) is non-zero when Re(z) = 0 rather than when  $\text{Re}(z) \ge 0$  in terms of causality is similar to that of non-fractional SDDEs (see, e.g., [2]).

The next result shows why one may view (16) as (17). In particular, it reveals that the corresponding solution  $(X_t)_{t \in \mathbb{R}}$  is a semimartingale with respect to (the completion of) its own filtration or equivalently, in light of (16) and (24), the one generated by the increments of  $(L_t)_{t \in \mathbb{R}}$ .

**Proposition 3.6.** Suppose that h(z) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$  and let  $(X_t)_{t \in \mathbb{R}}$  be the solution to (16) given in Theorem 3.2. Then, for  $t \in \mathbb{R}$ , the limit

$$D^{\beta}X_{t} := \frac{\beta}{\Gamma(1-\beta)} \lim_{\delta \downarrow 0} \int_{\delta}^{\infty} \frac{X_{t} - X_{t-u}}{u^{1+\beta}} du$$
 (26)

exists in  $L^2(\mathbb{P})$ ,  $D^{\beta}X_t = \int_{-\infty}^t D^{\beta}g(t-u) dL_u$ , and it holds that

$$\frac{1}{\Gamma(1-\beta)} \int_{-\infty}^{t} \left[ (t-u)^{-\beta} - (s-u)_{+}^{-\beta} \right] \int_{[0,\infty)} X_{u-v} \eta(\mathrm{d}v) \,\mathrm{d}u$$

$$= \int_{s}^{t} \int_{[0,\infty)} D^{\beta} X_{u-v} \eta(\mathrm{d}v) \,\mathrm{d}u \tag{27}$$

almost surely for each s < t.

We will now provide some properties of the solution  $(X_t)_{t \in \mathbb{R}}$  to (16) given in (24). Since the autocovariance function  $\gamma_X$  takes the form

$$\gamma_X(t) = \int_{\mathbb{R}} g(t+u)g(u) \, \mathrm{d}u, \quad t \in \mathbb{R},$$
 (28)

it follows by Plancherel's theorem that  $(X_t)_{t\in\mathbb{R}}$  admits a spectral density  $f_X$  which is given by

$$f_X(y) = \left| \mathcal{F}[g](y) \right|^2 = \frac{1}{|h(iy)|^2} |y|^{-2\beta}, \quad y \in \mathbb{R}.$$
 (29)

(See the appendix for a brief recap of the spectral theory.) The following result concerning  $\gamma_X$  and  $f_X$  shows that solutions to (16) exhibit a long-memory behavior and that the degree of memory can be controlled by  $\beta$ .

**Proposition 3.7.** Suppose that h(z) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$  and let  $\gamma_X$  and  $f_X$  be the functions introduced in (28)–(29). Then it holds that

$$\gamma_X(t) \sim \frac{\Gamma(1-2\beta)}{\Gamma(\beta)\Gamma(1-\beta)\eta([0,\infty))^2} t^{2\beta-1} \quad \text{as } t \to \infty$$
and 
$$f_X(y) \sim \frac{1}{\eta([0,\infty))^2} |y|^{-2\beta} \quad \text{as } y \to 0.$$

In particular,  $\int_{\mathbb{R}} |\gamma_X(t)| dt = \infty$ .

While the behavior of  $\gamma_X(t)$  as  $t \to \infty$  is controlled by  $\beta$ , the content of Proposition 3.8 is that the behavior of  $\gamma_X(t)$  as  $t \to 0$ , and thus the  $L^2(\mathbb{P})$  Hölder continuity of the sample paths of  $(X_t)_{t \in \mathbb{R}}$  (cf. Remark 3.9), is unaffected by  $\beta$ .

**Proposition 3.8.** Suppose that h(z) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$ , let  $(X_t)_{t \in \mathbb{R}}$  be the solution to (16) and denote by  $\rho_X$  its ACF. Then it holds that  $1 - \rho_X(h) \sim h$  as  $h \downarrow 0$ .

**Remark 3.9.** Recall that for a given  $\gamma > 0$ , a centered and square integrable process  $(X_t)_{t \in \mathbb{R}}$  with stationary increments is said to be locally  $\gamma$ -Hölder continuous in  $L^2(\mathbb{P})$  if there exists a constant C > 0 such that

$$\frac{\mathbb{E}[(X_t - X_0)^2]}{t^{2\gamma}} \le C$$

for all sufficiently small t > 0. By defining the semi-variogram

$$\gamma_V(t) := \frac{1}{2} \mathbb{E}[(X_t - X_0)^2], \quad t \in \mathbb{R},$$

we see that  $(X_t)_{t\in\mathbb{R}}$  is locally  $\gamma$ -Hölder continuous if and only if  $\gamma_V(t) = O(t^{2\gamma})$  as  $t \to 0$ . When  $(X_t)_{t\in\mathbb{R}}$  is stationary we have the relation  $\gamma_V = \gamma_X(0)(1-\rho_X)$ , from which it follows that the  $L^2(\mathbb{P})$  notion of Hölder continuity can be characterized in terms of the behavior of the ACF at zero. In particular, Proposition 3.8 shows that the solution  $(X_t)_{t\in\mathbb{R}}$  to (16) is locally  $\gamma$ -Hölder continuous if and only if  $\gamma \leq 1/2$ . The behavior of the ACF at zero has been used as a measure of roughness of the sample paths in for example, [4,5].

**Remark 3.10.** As a final comment on the path properties of the solution  $(X_t)_{t\in\mathbb{R}}$  to (16), observe that

$$X_t - X_s = \int_s^t \int_{[0,\infty)} D^{\beta} X_{u-v} \eta(\mathrm{d}v) \,\mathrm{d}u + L_t - L_s$$

for each s < t almost surely by Proposition 3.6. This shows that  $(X_t)_{t \in \mathbb{R}}$  can be chosen so that it has jumps at the same time (and of the same size) as  $(L_t)_{t \in \mathbb{R}}$ . This is in contrast to models driven by a fractional Lévy process, such as (9), since  $(I^{\beta}L_t)_{t \in \mathbb{R}}$  is continuous in t (see [21], Theorem 3.4).

We end this section by providing a formula for computing  $\mathbb{E}[X_t \mid X_u, u \leq s]$  for any s < t. One should compare its form to those obtained for other fractional models (such as the one in [3], Theorem 3.2, where, as opposed to Proposition 3.11, the prediction is expressed not only in terms of its own past, but also the past noise).

**Proposition 3.11.** Suppose that h(z) is non-zero for every  $z \in \mathbb{C}$  with  $\text{Re}(z) \geq 0$  and let  $(X_t)_{t \in \mathbb{R}}$  denote the solution to (16). Then for any s < t, it holds that

$$\mathbb{E}[X_t \mid X_u, u \le s] = g(t - s)X_s + \int_{[0, t - s)} \int_{-\infty}^s X_w \int_{[0, \infty)} (D_-^{\beta} \mathbb{1}_{(s, t - u)})(v + w) \eta(\mathrm{d}v) \,\mathrm{d}w g(\mathrm{d}u),$$

where  $g(du) = \delta_0(du) + (D^{\beta}g) * \eta(u) du$  is the Lebesgue–Stieltjes measure induced by g.

# 4. Delays of exponential type

Let A be an  $n \times n$  matrix where all its eigenvalues belong to  $\{z \in \mathbb{C} : \text{Re}(z) < 0\}$ , and let  $b \in \mathbb{R}^n$  and  $\kappa \in \mathbb{R}$ . In this section, we restrict our attention to measures  $\eta$  of the form

$$\eta(\mathrm{d}t) = -\kappa \,\delta_0(\mathrm{d}t) + f(t)\,\mathrm{d}t \quad \text{with } f(t) = \mathbb{1}_{[0,\infty)}(t)b^{\mathsf{T}}e^{At}e_1,\tag{30}$$

where  $e_1 := [1, 0, \dots, 0]^{\top} \in \mathbb{R}^n$ . Note that  $e_1$  is used as a normalization; the effect of replacing  $e_1$  by any  $c \in \mathbb{R}^n$  can be incorporated in the choice of A and b. It is well known that the assumption on the eigenvalues of A imply that all the entries of  $e^{Au}$  decay exponentially fast as  $u \to \infty$ , so that  $\eta$  is a finite measure on  $[0, \infty)$  with moments of any order. Since the Fourier transform  $\mathcal{F}[f]$  of f is given by

$$\mathcal{F}[f](y) = b^{\top} (I_n i y - A)^{-1} e_1, \quad y \in \mathbb{R},$$

it admits a fraction decomposition; that is, there exist real polynomials  $Q, R : \mathbb{C} \to \mathbb{C}$ , Q being monic with the eigenvalues of A as its roots and being of larger degree than R, such that

$$\mathcal{F}[f](y) = -\frac{R(iy)}{O(iy)} \tag{31}$$

for  $y \in \mathbb{R}$ . (This is a direct consequence of the inversion formula  $B^{-1} = \operatorname{adj}(B)/\operatorname{det}(B)$ .) By assuming that Q and R have no common roots, the pair (Q, R) is unique. The following existence and uniqueness result is simply an application of Theorem 3.2 to the particular setup in question.

**Corollary 4.1.** Let Q and R be given as in (31). Suppose that  $\kappa + b^{\top} A^{-1} e_1 \neq 0$  and

$$Q(z)[z + \kappa z^{\beta}] + R(z)z^{\beta} \neq 0$$
(32)

for all  $z \in \mathbb{C} \setminus \{0\}$  with  $\text{Re}(z) \geq 0$ . Then there exists a unique purely non-deterministic solution  $(X_t)_{t \in \mathbb{R}}$  to (16) with  $\eta$  given by (30) and it is given by (24) with  $g : \mathbb{R} \to \mathbb{R}$  characterized through

the relation

$$\mathcal{F}[g](y) = \frac{Q(iy)}{Q(iy)[iy + \kappa(iy)^{\beta}] + R(iy)(iy)^{\beta}}, \quad y \in \mathbb{R}.$$
 (33)

Before giving examples we state Proposition 4.2, which shows that the general SFDDE (16) can be written as

$$dX_t = -\kappa D^{\beta} X_t dt + \int_0^\infty X_{t-u} D^{\beta} f(u) du dt + dL_t, \quad t \in \mathbb{R},$$
 (34)

when  $\eta$  is of the form (30). In case  $\kappa = 0$ , (34) is a (non-fractional) SDDE. However, the usual existence results obtained in this setting (for instance, those in [2] and [18]) are not applicable, since the delay measure  $D^{\beta} f(u) du$  has unbounded support and zero total mass  $\int_0^{\infty} D^{\beta} f(u) du = 0$ .

**Proposition 4.2.** Let f be of the form (30). Then  $D^{\beta} f: \mathbb{R} \to \mathbb{R}$  defined by  $D^{\beta} f(t) = 0$  for  $t \leq 0$  and

$$D^{\beta} f(t) = \frac{1}{\Gamma(1-\beta)} b^{\top} \left( A e^{At} \int_0^t e^{-Au} u^{-\beta} du + t^{-\beta} I_n \right) e_1$$

for t > 0 belongs to  $L^1 \cap L^2$ . If in addition (32) holds,  $\kappa + b^{\top} A^{-1} e_1 \neq 0$  and  $(X_t)_{t \in \mathbb{R}}$  is the solution given in Corollary 4.1, then

$$\int_0^\infty D^{\beta} X_{t-u} f(u) du = \int_0^\infty X_{t-u} D^{\beta} f(u) du$$

almost surely for any  $t \in \mathbb{R}$ .

**Remark 4.3.** Due to the structure of the function g in (33) one may, in line with the interpretation of CARMA processes, think of the corresponding solution  $(X_t)_{t\in\mathbb{R}}$  as a stationary process that satisfies the formal equation

$$(Q(D)[D + \kappa D^{\beta}] + R(D)D^{\beta})X_t = Q(D)DL_t, \quad t \in \mathbb{R},$$
(35)

where D denotes differentiation with respect to t and  $D^{\beta}$  is a suitable fractional derivative. Indeed, by heuristically applying the Fourier transform  $\mathcal{F}$  to (35) and using computation rules such as  $\mathcal{F}[DX](y)=iy\mathcal{F}[X](y)$  and  $\mathcal{F}[D^{\beta}X](y)=(iy)^{\beta}\mathcal{F}[X](y)$ , one ends up concluding that  $(X_t)_{t\in\mathbb{R}}$  is of the form (24) with g characterized by (33). For two monic polynomials P and Q with  $q:=\deg(Q)=\deg(P)-1$  and all their roots contained in  $\{z\in\mathbb{C}: \operatorname{Re}(z)<0\}$ , consider the FICARMA $(q+1,\beta,q)$  process  $(X_t)_{t\in\mathbb{R}}$ . Heuristically, by applying  $\mathcal{F}$  as above,  $(X_t)_{t\in\mathbb{R}}$  may be thought of as the solution to  $P(D)D^{\beta}X_t=Q(D)DL_t$ ,  $t\in\mathbb{R}$ . By choosing the polynomial R and the constant  $\kappa$  such that  $P(z)=Q(z)[z+\kappa]+R(z)$  we can think of  $(X_t)_{t\in\mathbb{R}}$  as the solution to the formal equation

$$(Q(D)[D^{1+\beta} + \kappa D^{\beta}] + R(D)D^{\beta})X_t = Q(D)DL_t, \quad t \in \mathbb{R}.$$
 (36)

It follows that (35) and (36) are closely related, the only difference being that  $D + \kappa D^{\beta}$  is replaced by  $D^{1+\beta} + \kappa D^{\beta}$ . In particular, one may view solutions to SFDDEs corresponding to measures of the form (30) as being of the same type as FICARMA processes. While the considerations above apply only to the case where  $\deg(P) = q + 1$ , it should be possible to extend the SFDDE framework so that solutions are comparable to the FICARMA processes in the general case  $\deg(P) > q$  by following the lines of [3], where similar theory is developed for the SDDE setting.

We will now give two examples of (34).

**Example 4.4.** Consider choosing  $\eta = -\kappa \delta_0$  for some  $\kappa > 0$  so that (16) becomes

$$X_t - X_s = -\frac{\kappa}{\Gamma(1-\beta)} \int_{-\infty}^t \left[ (t-u)^{-\beta} - (s-u)_+^{-\beta} \right] X_u \, \mathrm{d}u + L_t - L_s, \quad s < t, \tag{37}$$

or, in short,

$$dX_t = -\kappa D^{\beta} X_t dt + dL_t, \quad t \in \mathbb{R}.$$
(38)

To argue that a unique purely non-deterministic solution exists, we observe that Q(z) = 1 and R(z) = 0 for all  $z \in \mathbb{C}$ . Thus, in light of Corollary 4.1 and (32), it suffices to argue that  $z + \kappa z^{\beta} \neq 0$  for all  $z \in \mathbb{C} \setminus \{0\}$  with  $Re(z) \geq 0$ . By writing such z as  $z = re^{i\theta}$  for a suitable r > 0 and  $\theta \in [-\pi/2, \pi/2]$ , the condition may be written as

$$(r\cos(\theta) + \kappa r^{\beta}\cos(\beta\theta)) + i(r\sin(\theta) + \kappa r^{\beta}\sin(\beta\theta)) \neq 0.$$
 (39)

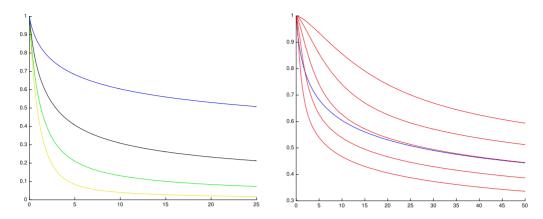
If the imaginary part of the left-hand side of (39) is zero it must be the case that  $\theta=0$ , since  $\kappa>0$  while  $\sin(\theta)$  and  $\sin(\beta\theta)$  are of the same sign. However, if  $\theta=0$ , the real part of the left-hand side of (39) is  $r+\kappa r^{\beta}>0$ . Consequently, Corollary 4.1 implies that a solution to (38) is characterized by (24) and  $\mathcal{F}[g](y)=((iy)^{\beta}\kappa+iy)^{-1}$  for  $y\in\mathbb{R}$ . In particular,  $\gamma_X$  takes the form

$$\gamma_X(t) = \int_{\mathbb{R}} \frac{e^{ity}}{y^2 + 2\kappa \sin(\frac{\beta\pi}{2})|y|^{1+\beta} + \kappa^2|y|^{2\beta}} \, \mathrm{d}y, \quad t \in \mathbb{R}.$$
 (40)

In Figure 1, we have plotted the ACF of  $(X_t)_{t\in\mathbb{R}}$  using (40) with  $\kappa=1$  and  $\beta\in\{0.1,0.2,0.3,0.4\}$ . We compare it to the ACF of the corresponding fractional Ornstein–Uhlenbeck process (equivalently, the FICARMA(1,  $\beta$ , 0) process) which was presented in (4). To do so, we use that its autocovariance function  $\gamma_{\beta}$  is given by

$$\gamma_{\beta}(t) = \int_{\mathbb{R}} \frac{e^{ity}}{|y|^{2(1+\beta)} + \kappa^2 |y|^{2\beta}} \, \mathrm{d}y, \quad t \in \mathbb{R}.$$
 (41)

From these plots it becomes evident that, although the ACFs share the same behavior at infinity, they behave differently near zero. In particular, we see that the ACF of  $(X_t)_{t \in \mathbb{R}}$  decays more rapidly around zero, which is in line with Proposition 3.8 and the fact that the  $L^2(\mathbb{P})$  Hölder continuity of the fractional Ornstein–Uhlenbeck process increases as  $\beta$  increases (cf. the introduction).



**Figure 1.** The left plot is the ACF based on (40) with  $\beta = 0.1$  (yellow),  $\beta = 0.2$  (green),  $\beta = 0.3$  (black) and  $\beta = 0.4$  (blue). With  $\beta = 0.4$  fixed, the plot on the right compares the ACF based on (40) with  $\kappa = 1$  (blue) to the ACF based on (41) for  $\kappa = 0.125, 0.25, 0.5, 1, 2$  (red) where the ACF decreases in  $\kappa$ .

**Example 4.5.** Suppose that  $\eta$  is given by (30) with  $\kappa = 0$ ,  $A = -\kappa_1$  and  $b = -\kappa_2$  for some  $\kappa_1, \kappa_2 > 0$ . In this case,  $f(t) = -\kappa_2 e^{-\kappa_1 t}$  and (34) becomes

$$dX_t = \frac{\kappa_2}{\Gamma(1-\beta)} \int_0^\infty X_{t-u} \left( \kappa_1 e^{-\kappa_1 u} \int_0^u e^{\kappa_1 v} v^{-\beta} dv - u^{-\beta} \right) du dt + dL_t, \quad t \in \mathbb{R},$$
 (42)

and since  $Q(z) = z + \kappa_1$  and  $R(z) = \kappa_2$  we have that

$$zQ(z) + R(z)z^{\beta} = z^2 + \kappa_1 z + \kappa_2 z^{\beta}.$$

To verify (32), set z = x + iy for x > 0 and  $y \in \mathbb{R}$  and note that

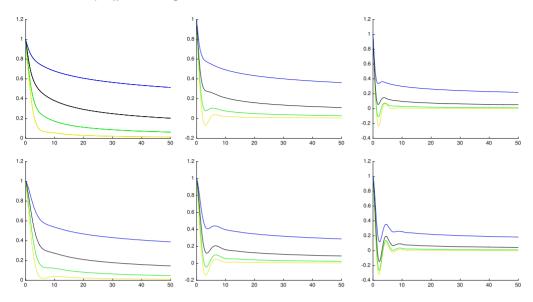
$$z^{2} + \kappa_{1}z + \kappa_{2}z^{\beta} = \left(x^{2} - y^{2} + \kappa_{1}x + \kappa_{2}\cos(\beta\theta_{z})|z|^{\beta}\right)$$
$$+ i\left(\kappa_{1}y + 2xy + \kappa_{2}\sin(\beta\theta_{z})|z|^{\beta}\right) \tag{43}$$

for a suitable  $\theta_z \in (-\pi/2, \pi/2)$ . For the imaginary part of (43) to be zero it must be the case that

$$(\kappa_1 + 2x)y = -\kappa_2 \sin(\beta \theta_z)|z|^{\beta},$$

and this can only happen if y = 0, since  $x, \kappa_1, \kappa_2 > 0$  and the sign of y is the same as that of  $\sin(\beta\theta_z)$ . However, if y = 0 it is easy to see that the real part of (43) cannot be zero for any x > 0, so we conclude that (32) holds and that there exists a stationary solution  $(X_t)_{t \in \mathbb{R}}$  given through the kernel (33). With  $\gamma_1 = \cos(\beta\pi/2)$  and  $\gamma_2 = \sin(\beta\pi/2)$  the autocovariance function  $\gamma_X$  is given by

$$\gamma_X(t) = \int_{\mathbb{R}} e^{ity} \frac{y^2 + \kappa_1^2}{y^4 + 2\kappa_2(\kappa_1 \gamma_2 |y|^{1+\beta} - \gamma_1 |y|^{2+\beta}) + \kappa_1^2 y^2 + \kappa_2^2 |y|^{2\beta}} \, \mathrm{d}y, \quad t \in \mathbb{R}.$$
 (44)



**Figure 2.** First row is ACF based on (44), second row is ACF based on (45), and the columns correspond to  $\kappa_2 = 0.5$ ,  $\kappa_2 = 1$  and  $\kappa_2 = 2$ , respectively. Within each plot, the lines correspond to  $\beta = 0.1$  (yellow),  $\beta = 0.2$  (green),  $\beta = 0.3$  (black) and  $\beta = 0.4$  (blue). In all plots,  $\kappa_1 = 1$ .

The polynomials to the associated FICARMA(2,  $\beta$ , 1) process are given by  $P(z) = z^2 + \kappa_1 z + \kappa_2$  and  $Q(z) = z + \kappa_1$  (see Remark 4.3) and the autocovariance function  $\gamma_{\beta}$  takes the form

$$\gamma_{\beta}(t) = \int_{\mathbb{R}} e^{ity} \frac{y^2 + \kappa_1^2}{|y|^{4+2\beta} + (\kappa_1^2 - 2\kappa_2)|y|^{2+2\beta} + \kappa_2^2|y|^{2\beta}} \, \mathrm{d}y, \quad t \in \mathbb{R}.$$
 (45)

In Figure 2, we have plotted the ACF based on (44) for  $\kappa_1 = 1$  and various values of  $\kappa_2$  and  $\beta$ . For comparison, we have also plotted the ACF based on (45) for the same choices of  $\kappa_1$ ,  $\kappa_2$  and  $\beta$ . From these plots, we see that both the ACF corresponding to (44) and (45) are decreasing in  $\kappa_2$ , which is similar to the role of  $\kappa$  in Example 4.4. It appears as well that a larger  $\kappa_2$  causes more curvature, although this effect is less pronounced for (44) than for (45).

## 5. Simulation from the SFDDE

In the following, we will focus on simulating from (16). We begin this simulation study by considering the Ornstein-Uhlenbeck type equation discussed in Example 4.4 with  $\kappa=1$  and under the assumption that  $(L_t)_{t\in\mathbb{R}}$  is a standard Brownian motion. Let  $c_1=100/\Delta$  and  $c_2=2000/\Delta$ . We generate a simulation of the solution process  $(X_t)_{t\in\mathbb{R}}$  on a grid of size  $\Delta=0.01$  and with  $3700/\Delta$  steps of size  $\Delta$  starting from  $-c_1-c_2$  and ending at  $1600/\Delta$ . Initially, we set

 $X_t$  equal to zero for the first  $c_1$  points in the grid and then discretize (37) using the approximation

$$\int_{\mathbb{R}} \left[ (n\Delta - u)_{+}^{-\beta} - \left( (n-1)\Delta - u \right)_{+}^{-\beta} \right] X_{u} \, du$$

$$\simeq \frac{1}{1-\beta} \Delta^{1-\beta} X_{(n-1)\Delta}$$

$$+ \sum_{k=n-c_{1}}^{n-1} \frac{X_{k\Delta} + X_{(k-1)\Delta}}{2} \int_{(k-1)\Delta}^{k\Delta} \left[ (n\Delta - u)_{+}^{-\beta} - \left( (n-1)\Delta - u \right)_{+}^{-\beta} \right] du$$

$$= \frac{1}{1-\beta} \Delta^{1-\beta} X_{(n-1)\Delta} + \frac{1}{1-\beta} \sum_{k=n-c_{1}}^{n-1} \frac{X_{k\Delta} + X_{(k-1)\Delta}}{2}$$

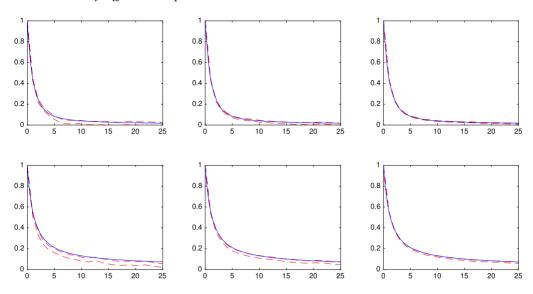
$$\cdot \left( 2 \left( (n-k-1)\Delta \right)^{1-\beta} - \left( (n-k)\Delta \right)^{1-\beta} - \left( (n-k-2)\Delta \right)^{1-\beta} \right)$$

for  $n = -c_2 + 1, ..., 3700/\Delta - c_2 - c_1$ . Next, we disregard the first  $c_1 + c_2$  values of the simulated sample path to obtain an approximate sample from the stationary distribution. We assume that the process is observed on a unit grid resulting in simulated values  $X_1, ..., X_{1600}$ . This is repeated 200 times, and in every repetition the sample ACF based on  $X_1, ..., X_L$  is computed for t = 1, ..., 25 and L = 100, 400, 1600. In long-memory models, the sample mean  $\bar{X}_L$  can be a poor approximation to the true mean  $\mathbb{E}[X_0]$  even for large L, and this may result in considerable negative (finite sample) bias in the sample ACF (see, e.g., [23]). Due to this bias, it may be difficult to see if we succeed in simulating from (16), and hence we will assume that  $\mathbb{E}[X_0]$  is known to be zero when computing the sample ACF. We calculate the 95% confidence interval

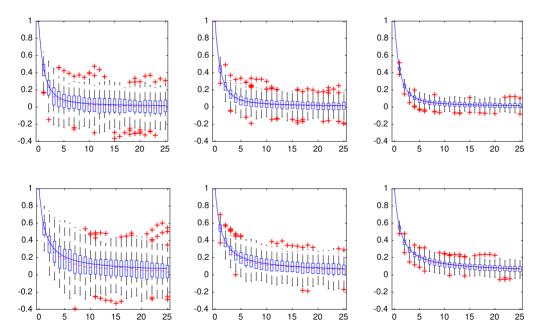
$$\left[\bar{\rho}(k) - 1.96 \frac{\hat{\sigma}(k)}{\sqrt{200}}, \bar{\rho}(k) + 1.96 \frac{\hat{\sigma}(k)}{\sqrt{200}}\right],$$

for the mean of the sample ACF based on L observations at lag k. Here  $\bar{\rho}(k)$  is the sample mean and  $\hat{\sigma}(k)$  is the sample standard deviations of the ACF at lag k based on the 200 replications. In Figure 3, the theoretical ACFs and the corresponding 95% confidence intervals for the mean of the sample ACFs are plotted for  $\beta=0.1,0.2$  and L=100,400,1600. We see that, when correcting for the bias induced by an unknown mean  $\mathbb{E}[X_0]$ , simulation from equation (37) results in a fairly unbiased estimator of the ACF for small values of  $\beta$ . When  $\beta>0.25$ , in the case where the ACF of  $(X_t)_{t\in\mathbb{R}}$  is not even in  $L^2$ , the results are more unstable as it requires large values of  $c_1$  and  $c_2$  to ensure that the simulation results in a good approximation to the stationary distribution of  $(X_t)_{t\in\mathbb{R}}$ . Moreover, even after correcting for the bias induced by an unknown mean of the observed process, the sample ACF for the ARFIMA process shows considerable finite sample bias when  $\beta>0.25$ , see [23], and hence we may expect this to apply to solutions to (16) as well.

In Figure 4, we have plotted box plots for the 200 replications of the sample ACF for  $\beta$  = 0.1, 0.2 and L = 100, 400, 1600. We see that the sample ACFs have the expected convergence when L grows and that the distribution is more concentrated in the case where less memory is present.



**Figure 3.** Theoretical ACF and 95% confidence intervals of the mean of the sample ACF based on 200 replications of  $X_1, \ldots, X_L$ . Columns correspond to L = 100, L = 400 and L = 1600, respectively, and rows correspond to  $\beta = 0.1$  and  $\beta = 0.2$ , respectively. The model is (37).



**Figure 4.** Box plots for the sample ACF based on 200 replications of  $X_1, \ldots, X_L$  together with the theoretical ACF. Columns correspond to L = 100, L = 400 and L = 1600, respectively, and rows correspond to  $\beta = 0.1$  and  $\beta = 0.2$ , respectively. The model is (37).

Following the same approach as above, we simulate the solution to the equation discussed in Example 4.5. Specifically, the simulation is based on equation (16), restricted to the case where  $\eta(\mathrm{d}t) = -e^{-t}\,\mathrm{d}t$  and  $(L_t)_{t\in\mathbb{R}}$  is a standard Brownian motion. In this case, we use the approximation

$$\int_{\mathbb{R}} \left[ (n\Delta - u)_{+}^{-\beta} - \left( (n-1)\Delta - u \right)_{+}^{-\beta} \right] \int_{0}^{\infty} X_{u-v} e^{-v} \, dv \, du$$

$$= \int_{0}^{\infty} X_{n\Delta-v} \int_{0}^{v} \left[ (u-\Delta)_{+}^{-\beta} - u_{+}^{-\beta} \right] e^{u-v} \, du \, dv$$

$$\simeq \frac{1}{2} \Delta X_{(n-1)\Delta} f(\Delta)$$

$$+ \sum_{k=2}^{c_{1}} \frac{1}{4} \Delta (X_{(n-k)\Delta} + X_{(n-k+1)\Delta}) \left( \varphi(k\Delta) + \varphi((k-1)\Delta) \right)$$

where  $\varphi \colon \mathbb{R} \to \mathbb{R}$  is given by

$$\varphi(v) = \int_0^v \left[ (u - \Delta)_+^{-\beta} - u^{-\beta} \right] e^{u - v} du.$$

We approximate  $\varphi$  recursively by noting that

$$\varphi(k\Delta) = \int_0^{k\Delta} \left[ (u - \Delta)_+^{-\beta} - u^{-\beta} \right] e^{u - k\Delta} \, dv$$

$$\simeq \frac{1 + e^{-\Delta}}{2} \int_{(k-1)\Delta}^{k\Delta} \left[ (u - \Delta)_+^{-\beta} - u_+^{-\beta} \right] \, dv + e^{-\Delta} \varphi \left( (k-1)\Delta \right)$$

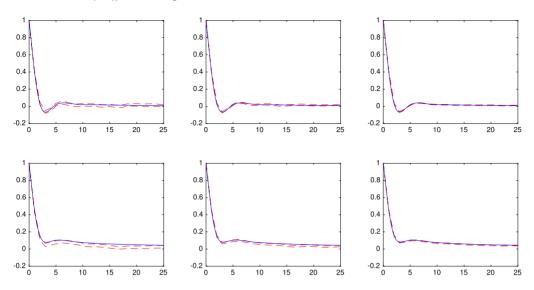
$$= \frac{1}{1 - \beta} \frac{1 + e^{-\Delta}}{2} \left[ \left( (k-1)\Delta \right)^{1-\beta} - (k\Delta)^{1-\beta} \right] + e^{-\Delta} \varphi \left( (k-1)\Delta \right)$$

for  $k \ge 1$ . The theoretical ACFs and corresponding 95% confidence intervals are plotted in Figure 5 and the box plots in Figure 6. The findings are consistent with the first example that we considered in the sense of convergence of the sample ACF and the effect of memory (the value of  $\beta$ ).

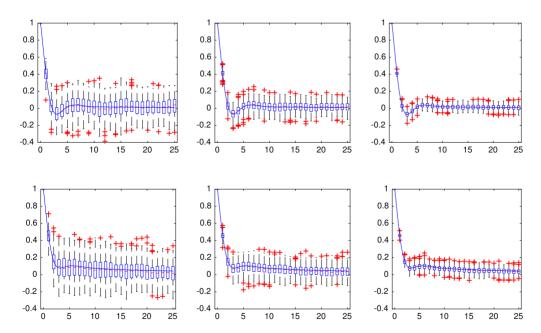
### 6. Proofs

**Proof of Proposition 3.1.** For  $\gamma > 0$  define  $h_{\gamma}(z) = z^{\gamma}/h(z)$  for each  $z \in \mathbb{C} \setminus \{0\}$  with  $\text{Re}(z) \geq 0$ . By continuity of h and the asymptotics  $|h_{\gamma}(z)| \sim |\eta([0,\infty))|^{-1}|z|^{\gamma}$ ,  $|z| \to 0$ , and  $|h_{\gamma}(z)| \sim |z|^{\gamma-1}$ ,  $|z| \to \infty$ , it follows that

$$\sup_{x>0} \int_{\mathbb{R}} \left| h_{\gamma}(x+iy) \right|^2 \mathrm{d}y < \infty \tag{46}$$



**Figure 5.** Theoretical ACF and 95% confidence intervals of the mean of the sample ACF sample based on 200 replications of  $X_1, \ldots, X_L$ . Columns correspond to L = 100, L = 400 and L = 1600, respectively, and rows correspond to  $\beta = 0.1$  and  $\beta = 0.2$ , respectively. The model is (42).



**Figure 6.** Box plots for the sample ACF based on 200 replications of  $X_1, \ldots, X_L$  together with the theoretical ACF. Columns correspond to L = 100, L = 400 and L = 1600, respectively, and rows correspond to  $\beta = 0.1$  and  $\beta = 0.2$ , respectively. The model is (42).

for  $\gamma \in (-1/2, 1/2)$ . In other words,  $h_{\gamma}$  is a certain Hardy function, and thus there exists a function  $f_{\gamma} : \mathbb{R} \to \mathbb{R}$  in  $L^2$  which is vanishing on  $(-\infty, 0)$  and has  $\mathcal{L}[f_{\gamma}](z) = h_{\gamma}(z)$  when Re(z) > 0, see [2,12,14]. Note that  $f_{\gamma}$  is indeed real-valued, since  $h_{\gamma}(x-iy) = h_{\gamma}(x+iy)$  for  $y \in \mathbb{R}$  and a fixed x > 0. We can apply [24], Proposition 2.3, to deduce that there exists a function  $g \in L^2$  satisfying (19) and that it can be represented as the (left-sided) Riemann–Liouville fractional integral of  $f_0$ , that is,

$$g(t) = \frac{1}{\Gamma(\beta)} \int_0^t f_0(u) (t - u)^{\beta - 1} du, \quad t > 0.$$

Conversely, [24], Theorem 2.1, ensures that  $D^{\beta}g$  given by (20) is a well-defined limit and that  $D^{\beta}g = f_0$ . In particular, we have shown (ii) and if we can argue that  $f_0 \in L^1$ , we have shown (i) as well. This follows from the assumption in (14), since then we have that  $y \mapsto \mathcal{L}[f_0](x+iy)$  is differentiable for any  $x \ge 0$  (except at 0 when x = 0) and

$$\mathcal{L}[u \mapsto u f_0(u)](x + iy) = i \frac{d}{dy} \mathcal{L}[f_0](x + iy)$$

$$= \frac{\mathcal{L}[u \eta (du)](x + iy) + (1 - \beta)(x + iy)^{-\beta}}{h(x + iy)^2}.$$
(47)

The function  $\mathcal{L}[u \mapsto uf_0(u)]$  is analytic on  $\{z \in \mathbb{C} : \text{Re}(z) < 0\}$  and from the identity (47) it is not too difficult to see that it also satisfies the Hardy condition (46). This means  $u \mapsto uf_0(u)$  belongs to  $L^2$ , and hence we have that  $f_0$  belongs to  $L^1$ . Since g is the Riemann–Liouville integral of  $f_0$  of order  $\beta$  and  $f_0 \in L^1 \cap L^2$ , [3], Proposition 4.3, implies that  $g \in L^{\gamma}$  for  $(1 - \beta)^{-1} < \gamma \le 2$ .

It is straightforward to verify (22) and to obtain the identity

$$\int_{s}^{t} \left(D^{\beta} g\right) * \eta(u - \cdot) du = \int_{\mathbb{R}} \left(D_{-}^{\beta} \mathbb{1}_{(s,t]}\right) (u) g * \eta(u - \cdot) du$$

almost everywhere by comparing their Fourier transforms. This establishes the relation

$$g(t-v) - g(s-v) = \int_{s}^{t} (D^{\beta}g) * \eta(u-v) du + \mathbb{1}_{(s,t]}(v).$$

By letting  $s \to -\infty$ , and using that  $D^{\beta}g$  and g are both vanishing on  $(-\infty, 0)$ , we deduce that

$$g(t) = \mathbb{1}_{[0,\infty)}(t) \left( 1 + \int_0^t \left( D^{\beta} g \right) * \eta(u) \, \mathrm{d}u \right),$$

for almost all  $t \in \mathbb{R}$  which shows (21) and, thus, finishes the proof.

**Proof of Theorem 3.2.** Since  $g \in L^2$ , according to Proposition 3.1, and  $\mathbb{E}[L_1^2] < \infty$  and  $\mathbb{E}[L_1] = 0$ ,

$$X_t = \int_{-\infty}^t g(t-u) \, \mathrm{d}L_u, \quad t \in \mathbb{R},$$

is a well-defined process (e.g., in the sense of [26]) which is stationary with mean zero and finite second moments. By integrating both sides of (22) with respect to  $(L_t)_{t \in \mathbb{R}}$ , we obtain

$$X_t - X_s = \int_{\mathbb{R}} \left( \int_{\mathbb{R}} \left( D_-^{\beta} \mathbb{1}_{(s,t]} \right) (u) g * \eta(u - r) du \right) dL_r + L_t - L_s.$$

By a stochastic Fubini result (e.g., [1], Theorem 3.1) we can change the order of integration (twice) and obtain

$$\int_{\mathbb{R}} \left( \int_{\mathbb{R}} \left( D_{-}^{\beta} \mathbb{1}_{(s,t]} \right) (u) g * \eta(u-r) du \right) dL_{r} = \int_{\mathbb{R}} \left( D_{-}^{\beta} \mathbb{1}_{(s,t]} \right) (u) X * \eta(u) du.$$

This shows that  $(X_t)_{t\in\mathbb{R}}$  is a solution to (16). To show uniqueness, note that the spectral process  $\Lambda_X$  (with spectral distribution, say,  $F_X$ ) of any purely non-deterministic solution  $(X_t)_{t\in\mathbb{R}}$  satisfies

$$\int_{\mathbb{R}} \mathcal{F}[\mathbb{1}_{(s,t]}](-y)(iy)^{\beta} h(iy) \Lambda_X(\mathrm{d}y) = L_t - L_s \tag{48}$$

almost surely for all choices of s < t. This follows from the results in the supplementary material on spectral representations (see [10]). Using the fact that  $(X_t)_{t \in \mathbb{R}}$  is purely non-deterministic,  $F_X$  is absolutely continuous with respect to the Lebesgue measure, and hence we can extend (48) from  $\mathbb{1}_{(s,t]}$  to any function  $f \in L^2$  using an approximation of f with simple functions of the form  $s = \sum_{i=1}^n \alpha_i \mathbb{1}_{(t_{i-1},t_i]}$  for  $\alpha_i \in \mathbb{C}$  and  $t_0 < t_1 < \cdots < t_n$ . Specifically, we establish that

$$\int_{\mathbb{R}} \mathcal{F}[f](-y)(iy)^{\beta} h(iy) \Lambda_X(dy) = \int_{\mathbb{R}} f(u) dL_u$$
 (49)

almost surely for any  $f \in L^2$ . In particular, we may take  $f = g(t - \cdot)$ , g being the solution kernel characterized in (19), so that  $\mathcal{F}[f](y) = e^{-ity}(-iy)^{-\beta}/h(-iy)$  and (49) thus implies that  $X_t = \int_{-\infty}^t g(t-u) dL_u$ , which ends the proof.

**Proof of Proposition 3.6.** We start by arguing that the limit in (26) exists and is equal to  $\int_{-\infty}^{t} D^{\beta} g(t-u) dL_u$ . For a given  $\delta > 0$  it follows by a stochastic Fubini result that

$$\frac{\beta}{\Gamma(1-\beta)} \int_{\delta}^{\infty} \frac{X_t - X_{t-u}}{u^{1+\beta}} du = \int_{\mathbb{R}} D_{\delta}^{\beta} g(t-r) dL_r, \tag{50}$$

where

$$D_{\delta}^{\beta}g(t) = \frac{\beta}{\Gamma(1-\beta)} \int_{s}^{\infty} \frac{g(t) - g(t-u)}{u^{1+\beta}} du, \quad t > 0,$$

and  $D_{\delta}^{\beta}g(t)=0$  for  $t\leq 0$ . Suppose for the moment that  $(L_t)_{t\in\mathbb{R}}$  is a Brownian motion, so that  $(X_t)_{t\in\mathbb{R}}$  is  $\gamma$ -Hölder continuous for all  $\gamma\in(0,1/2)$  by (16). Then, almost surely,  $u\mapsto(X_t-X_{t-u})/u^{1+\beta}$  is in  $L^1$  and the relation (50) thus shows that

$$\int_{\mathbb{R}} \left[ D_{\delta}^{\beta} g(t-r) - D_{\delta'}^{\beta} g(t-r) \right] dL_r \stackrel{\mathbb{P}}{\longrightarrow} 0 \quad \text{as } \delta, \delta' \to 0,$$

which in turn implies that  $(D^{\beta}_{\delta}g)_{\delta>0}$  has a limit in  $L^2$ . We also know that this limit must be  $D^{\beta}g$ , since  $D^{\beta}_{\delta}g \to D^{\beta}g$  pointwise as  $\delta \downarrow 0$  by (20). Having established this convergence, which does not rely on  $(L_t)_{t\in\mathbb{R}}$  being a Brownian motion, it follows immediately from (50) and the isometry property of the integral map  $\int_{\mathbb{R}} \cdot dL$  that the limit in (26) exists and that  $D^{\beta}X_t = \int_{-\infty}^t D^{\beta}g(t-u)\,dL_u$ . To show (27), we start by recalling the definition of  $D^{\beta}_{-1}\mathbb{I}_{(s,t]}$  in (15) and that  $\mathcal{F}[D^{\beta}_{-1}\mathbb{I}_{(s,t]}](y) = (-iy)^{\beta}\mathcal{F}[\mathbb{I}_{(s,t]}](y)$ . This identity can be shown by using that the improper integral  $\int_0^\infty e^{\pm iv}v^{\gamma-1}\,dv$  is equal to  $\Gamma(\gamma)e^{\pm i\pi\gamma/2}$  for any  $\gamma\in(0,1)$ . Now observe that

$$\mathcal{F}\left[\int_{\mathbb{R}} \left(D_{-}^{\beta} \mathbb{1}_{(s,t]}\right)(u)g * \eta(u - \cdot) du\right](y) = (-iy)^{\beta} \mathcal{F}[\mathbb{1}_{(s,t]}](y)\mathcal{F}[g](-y)\mathcal{F}[\eta](-y)$$

$$= \mathcal{F}[\mathbb{1}_{(s,t]}](y)\mathcal{F}\left[\left(D^{\beta}g\right) * \eta\right](-y)$$

$$= \mathcal{F}\left[\int_{s}^{t} \left(D^{\beta}g\right) * \eta(u - \cdot) du\right](y),$$

and hence  $\int_{\mathbb{R}} (D_{-}^{\beta} \mathbb{1}_{(s,t]})(u)g * \eta(u - \cdot) du = \int_{s}^{t} (D^{\beta}g) * \eta(u - \cdot) du$  almost everywhere. Consequently, using that  $D^{\beta}X_{t} = \int_{-\infty}^{t} D^{\beta}g(t - u) dL_{u}$  and applying a stochastic Fubini result twice,

$$\int_{s}^{t} \left(D^{\beta}X\right) * \eta(u) du = \int_{\mathbb{R}} \int_{s}^{t} \left(D^{\beta}g\right) * \eta(u-r) du dL_{r}$$

$$= \int_{\mathbb{R}} \int_{\mathbb{R}} \left(D_{-}^{\beta}\mathbb{1}_{(s,t]}\right) (u)g * \eta(u-r) du dL_{r}$$

$$= \frac{1}{\Gamma(1-\beta)} \int_{\mathbb{R}} \left[ (t-u)_{+}^{-\beta} - (s-u)_{+}^{-\beta} \right] X * \eta(u) du.$$

The semimartingale property of  $(X_t)_{t \in \mathbb{R}}$  is now an immediate consequence of (16).

**Proof of Proposition 3.7.** Using (29) and that  $h(0) = -\eta([0, \infty))$ , it follows that  $f_X(y) \sim |y|^{-2\beta}/\eta([0, \infty))^2$  as  $y \to 0$ . To show the asymptotic behavior of  $\gamma_X$  at  $\infty$  we start by recalling that, for  $u, v \in \mathbb{R}$ ,

$$\int_{u \vee v}^{\infty} (s - u)^{\beta - 1} (s - v)^{\beta - 1} ds = \frac{\Gamma(\beta)\Gamma(1 - 2\beta)}{\Gamma(1 - \beta)} |u - v|^{2\beta - 1}$$

by [17], page 404. Having this relation in mind we use Proposition 3.1(ii) and (28) to do the computations

$$\gamma_X(t) = \frac{1}{\Gamma(\beta)^2} \int_{\mathbb{R}} \int_{\mathbb{R}} \int_{\mathbb{R}} D^{\beta} g(u) D^{\beta} g(v) (s+t-u)_+^{\beta-1} (s-v)_+^{\beta-1} dv du ds$$

$$= \frac{1}{\Gamma(\beta)^2} \int_{\mathbb{R}} \int_{\mathbb{R}} D^{\beta} g(u) D^{\beta} g(v)$$

$$\cdot \int_{(u-t)\vee v}^{\infty} (s-(u-t))^{\beta-1} (s-v)^{\beta-1} ds dv du$$

$$= \frac{\Gamma(1-2\beta)}{\Gamma(\beta)\Gamma(1-\beta)} \int_{\mathbb{R}} \int_{\mathbb{R}} D^{\beta} g(u) D^{\beta} g(v) |u-v-t|^{2\beta-1} dv du$$

$$= \frac{\Gamma(1-2\beta)}{\Gamma(\beta)\Gamma(1-\beta)} \int_{\mathbb{R}} \gamma(u) |u-t|^{2\beta-1} du,$$
(51)

where  $\gamma(u) = \int_{\mathbb{R}} D^{\beta} g(u+v) D^{\beta} g(v) dv$ . Note that  $\gamma \in L^1$  since  $D^{\beta} g \in L^1$  by Proposition 3.1 and, using Plancherel's theorem,

$$\gamma(u) = \int_{\mathbb{R}} e^{-iuy} \left| \mathcal{F} \left[ D^{\beta} g \right](y) \right|^2 dy = \mathcal{F} \left[ \left| h(i \cdot) \right|^{-2} \right](u).$$

In particular  $\int_{\mathbb{R}} \gamma(u) du = |h(0)|^{-2} = \eta([0, \infty))^{-2}$ , and hence it follows from (51) that we have shown the result if we can argue that

$$\frac{\int_{\mathbb{R}} \gamma(u)|u-t|^{2\beta-1} du}{t^{2\beta-1}} = \int_{\mathbb{R}} \frac{\gamma(u)}{|\frac{u}{t}-1|^{1-2\beta}} du \to \int_{\mathbb{R}} \gamma(u) du \quad \text{as } t \to \infty.$$
 (52)

It is clear by Lebesgue's theorem on dominated convergence that

$$\int_{-\infty}^{0} \frac{\gamma(u)}{|\frac{u}{t} - 1|^{1 - 2\beta}} du \to \int_{-\infty}^{0} \gamma(u) du \quad \text{as } t \to \infty.$$

Moreover, since  $|h(i \cdot)|^{-2}$  is continuous at 0 and differentiable on  $(-\infty, 0)$  and  $(0, \infty)$  with integrable derivatives, it is absolutely continuous on  $\mathbb{R}$  with a density  $\phi$  in  $L^1$ . As a consequence,  $\gamma(u) = \mathcal{F}[\phi](u)/(iu)$  and, thus,

$$\int_{t/2}^{\infty} \frac{\gamma(u)}{\left|\frac{u}{t} - 1\right|^{1-2\beta}} du = \int_{1/2}^{\infty} \frac{t\gamma(tu)}{|u - 1|^{1-2\beta}} du = -i \int_{1/2}^{\infty} \frac{\mathcal{F}[\phi](tu)}{u|u - 1|^{1-2\beta}} du.$$
 (53)

By the Riemann-Lebesgue lemma and Lebesgue's theorem on dominated convergence it follows that the right-hand side of expression in (53) tends to zero as t tends to infinity. Finally, integration by parts and the symmetry of  $\gamma$  yields

$$\int_0^{t/2} \gamma(u) \left( 1 - \frac{1}{\left| \frac{u}{t} - 1 \right|^{1 - 2\beta}} \right) du = \int_0^{t/2} t \gamma(tu) \left( 1 - \frac{1}{(1 - u)^{1 - 2\beta}} \right) du$$

$$= \left( 2^{1 - 2\beta} - 1 \right) \int_{-\infty}^{-t/2} \gamma(u) du$$

$$- \int_0^{t/2} \frac{1 - 2\beta}{(1 - u)^{2 - 2\beta}} \int_{-\infty}^{-tu} \gamma(v) dv du,$$

where both terms on the right-hand side converge to zero as t tends to infinity. Thus, we have shown (52), and this completes the proof.

**Proof of Proposition 3.8.** Observe that it is sufficient to argue  $\mathbb{E}[(X_t - X_0)^2] \sim t$  as  $t \downarrow 0$ . By using the spectral representation  $X_t = \int_{\mathbb{R}} e^{ity} \Lambda_X(\mathrm{d}y)$  and the isometry property of the integral map  $\int_{\mathbb{R}} \cdot \mathrm{d}\Lambda_X \colon L^2(F_X) \to L^2(\mathbb{P})$ , see [16], page 389, we have that

$$\frac{\mathbb{E}[(X_t - X_0)^2]}{t} = t^{-2} \int_{\mathbb{R}} |1 - e^{iy}|^2 f_X(y/t) \, \mathrm{d}y$$

$$= \int_{\mathbb{R}} \frac{|1 - e^{iy}|^2}{|y|^{2\beta} |(iy)^{1-\beta} - t^{1-\beta} \mathcal{F}[\eta](y/t)|^2} \, \mathrm{d}y. \tag{54}$$

Consider now a  $y \in \mathbb{R}$  satisfying  $|y| \ge C_1 t$  with  $C_1 := (2|\eta|([0,\infty)))^{1/(1-\beta)}$ . In this case,  $|y|^{1-\beta}/2 - |t^{1-\beta}\mathcal{F}[\eta](y/t)| \ge 0$ , and we thus get by the reversed triangle inequality that

$$\frac{|1 - e^{iy}|^2}{|y|^{2\beta}|(iy)^{1-\beta} - t^{1-\beta}\mathcal{F}[\eta](y/t)|^2} \le 2\frac{|1 - e^{iy}|^2}{y^2}.$$

If  $|y| < C_1 t$ , we note that the assumption on the function in (18) implies that

$$C_2 := \inf_{|x| < C_1} |(ix)^{1-\beta} - \mathcal{F}[\eta](x)| > 0,$$

which shows that

$$|(iy)^{1-\beta} - t^{1-\beta}\mathcal{F}[\eta](y/t)| \ge t^{1-\beta}C_2 \ge \frac{C_2}{C_1^{1-\beta}}|y|^{1-\beta}.$$

This establishes that

$$\frac{|1 - e^{iy}|^2}{|y|^{2\beta}|(iy)^{1-\beta} - t^{1-\beta}\mathcal{F}[\eta](y/t)|^2} \le \frac{C_1^{2(1-\beta)}}{C_2^2} \frac{|1 - e^{iy}|^2}{y^2}.$$

Consequently, it follows from (54) and Lebesgue's theorem on dominated convergence that

$$\frac{\mathbb{E}[(X_t - X_0)^2]}{t} \to \int_{\mathbb{R}} \frac{|1 - e^{iy}|^2}{y^2} \, \mathrm{d}y = \int_{\mathbb{R}} \left| \mathcal{F}[\mathbb{1}_{(0,1]}](y) \right|^2 \, \mathrm{d}y = 1 \quad \text{as } t \downarrow 0,$$

which was to be shown.

**Proof of Proposition 3.11.** We start by arguing that the first term on the right-hand side of the formula is well-defined. In order to do so, it suffices to argue that

$$\mathbb{E}\left[\int_{0}^{t-s} \int_{-\infty}^{s} |X_{w}| \int_{[0,\infty)} \left| \left(D_{-}^{\beta} \mathbb{1}_{(s,t-u)}\right)(v+w) \right| |\eta|(\mathrm{d}v) \, \mathrm{d}w|g|(\mathrm{d}u) \right] \\
\leq \mathbb{E}\left[|X_{0}|\right] \int_{0}^{t-s} \int_{[0,\infty)} \int_{-\infty}^{s} \left| \left(D_{-}^{\beta} \mathbb{1}_{(s,t-u)}\right)(v+w) \right| \, \mathrm{d}w|\eta|(\mathrm{d}v)|g|(\mathrm{d}u) \tag{55}$$

is finite. This is implied by the facts that

$$\Gamma(1-\beta) \int_{-\infty}^{s} \left| \left( D_{-}^{\beta} \mathbb{1}_{(s,t-u)} \right) (v+w) \right| dw$$

$$\leq \int_{u+s-t}^{0} (t-s-u+w)^{-\beta} dw + \int_{0}^{1} \left[ w^{-\beta} - (t-s-u+w)^{-\beta} \right] dw$$

$$+ (1+\beta) \int_{1}^{\infty} w^{-1-\beta} (t-s-u) dw$$

$$= \frac{1}{1-\beta} \left( 2(t-s-u)^{1-\beta} + 1 - (t-s-u+1)^{1-\beta} \right) + \frac{(1+\beta)}{\beta} (t-s-u)$$

$$\leq \frac{2}{1-\beta} (t-s)^{1-\beta} + \frac{(1+\beta)}{\beta} (t-s)$$

for  $u \in [0, t - s]$  and g(du) is a finite measure (since  $D^{\beta}g \in L^1$  by Proposition 3.1). Now fix an arbitrary  $z \in \mathbb{C}$  with Re(z) > 0. It follows from (16) that

$$\mathcal{L}[X\mathbb{1}_{(s,\infty)}](z) = X_s \mathcal{L}[\mathbb{1}_{(s,\infty)}](z) + \mathcal{L}[\mathbb{1}_{(s,\infty)}(L_{\cdot} - L_s)](z)$$

$$+ \mathcal{L}\left[\mathbb{1}_{(s,\infty)} \int_{\mathbb{R}} X_u \int_{[0,\infty)} \left(D_{-}^{\beta} \mathbb{1}_{(s,\cdot)}\right) (u+v) \eta(\mathrm{d}v) \,\mathrm{d}u\right](z).$$
 (56)

By noting that  $(D_{-}^{\beta} \mathbb{1}_{(s,t]})(u) = 0$  when  $t \le s < u$  we obtain

$$\mathcal{L}\left[\mathbb{1}_{(s,\infty)} \int_{s}^{\infty} X_{u} \int_{[0,\infty)} \left(D_{-}^{\beta} \mathbb{1}_{(s,\cdot)}\right) (u+v) \eta(\mathrm{d}v) \, \mathrm{d}u\right](z)$$

$$= \frac{1}{\Gamma(1-\beta)} \mathcal{L}\left[\int_{s}^{\infty} X_{u} \int_{[0,\infty)} (\cdot -u - v)_{+}^{-\beta} \eta(\mathrm{d}v) \, \mathrm{d}u\right](z)$$

$$= \mathcal{L}[\mathbb{1}_{(s,\infty)} X](z) \mathcal{L}[\eta](z) z^{\beta-1}.$$

Combining this observation with (56), we get the relation

$$\begin{split} & \left(z - z^{\beta} \mathcal{L}[\eta](z)\right) \mathcal{L}[\mathbb{1}_{(s,\infty)} X](z) \\ &= z X_{s} \mathcal{L}[\mathbb{1}_{(s,\infty)}](z) + z \mathcal{L}\left[\mathbb{1}_{(s,\infty)} (L - L_{s})\right](z) \\ &+ z \mathcal{L}\left[\mathbb{1}_{(s,\infty)} \int_{-\infty}^{s} X_{u} \int_{[0,\infty)} \left(D_{-}^{\beta} \mathbb{1}_{(s,\cdot)}\right) (u + v) \eta(\mathrm{d}v) \, \mathrm{d}u\right](z), \end{split}$$

which implies

$$\mathcal{L}[\mathbb{1}_{(s,\infty)}X](z)$$

$$= \mathcal{L}[g](z)\mathcal{L}[X_s\delta_0(s-\cdot)](z) + z\mathcal{L}[g](z)\mathcal{L}[\mathbb{1}_{(s,\infty)}(L-L_s)](z)$$

$$+z\mathcal{L}[g](z)\mathcal{L}\left[\mathbb{1}_{(s,\infty)}\int_{-\infty}^{s}X_{u}\int_{[0,\infty)}\left(D_{-}^{\beta}\mathbb{1}_{(s,\cdot]}\right)(u+v)\eta(\mathrm{d}v)\,\mathrm{d}u\right](z)$$

$$=\mathcal{L}\left[g(\cdot-s)X_{s}\right](z)+\mathcal{L}\left[\int_{s}^{\cdot}g(\cdot-u)\mathrm{d}L_{u}\right](z)$$

$$+\mathcal{L}\left[\int_{0}^{\cdot-s}\int_{-\infty}^{s}X_{w}\int_{[0,\infty)}\left(D_{-}^{\beta}\mathbb{1}_{(s,\cdot-u)}\right)(v+w)\eta(\mathrm{d}v)\,\mathrm{d}wg(\mathrm{d}u)\right](z).$$

This establishes the identity

$$X_{t} = g(t - s)X_{s} + \int_{s}^{t} g(t - u) dL_{u}$$

$$+ \int_{0}^{t - s} \int_{-\infty}^{s} X_{w} \int_{[0, \infty)} \left(D_{-}^{\beta} \mathbb{1}_{(s, t - u)}\right) (v + w) \eta(dv) dw g(du)$$
(57)

almost surely for Lebesgue almost all t > s. Since both sides of (57) are continuous in  $L^1(\mathbb{P})$ , the identity holds for each fixed pair s < t almost surely as well. By applying the conditional mean  $\mathbb{E}[\cdot \mid X_u, u \le s]$  on both sides of (57), we obtain the result.

**Proof of Corollary 4.1.** In this setup, it follows that the function h in (18) is given by

$$h(z) = z^{1-\beta} + \kappa + \frac{R(z)}{Q(z)},$$

where  $Q(z) \neq 0$  whenever  $Re(z) \geq 0$  by the assumption on A. This shows that h is non-zero (on  $\{z \in \mathbb{C} : Re(z) \geq 0\}$ ) if and only if

$$Q(z)\left[z^{1-\beta} + \kappa\right] + R(z) \neq 0 \quad \text{for all } z \in \mathbb{C} \text{ with } \operatorname{Re}(z) \geq 0.$$
 (58)

Condition (58) may equivalently be formulated as  $Q(z)[z + \kappa z^{\beta}] + R(z)z^{\beta} \neq 0$  for all  $z \in \mathbb{C} \setminus \{0\}$  with  $\text{Re}(z) \geq 0$  and  $h(0) = \kappa + b^{\top}A^{-1}e_1 \neq 0$ , which by Theorem 3.2 shows that a unique solution to (34) exists. It also provides the form of the solution, namely (24) with

$$\mathcal{F}[g](y) = \frac{(iy)^{-\beta}}{(iy)^{1-\beta} + \kappa + \frac{R(iy)}{O(iy)}} = \frac{Q(iy)}{Q(iy)[iy + \kappa(iy)^{\beta}] + R(iy)(iy)^{\beta}}, \quad y \in \mathbb{R}.$$

This finishes the proof.

**Proof of Proposition 4.2.** We will first show that  $D^{\beta} f \in L^1$ . By using that  $\int_0^{\infty} e^{Au} du = -A^{-1}$  we can rewrite  $D^{\beta} f$  as

$$D^{\beta} f(t) = \frac{1}{\Gamma(1-\beta)} b^{\top} A \left( \int_0^t e^{Au} \left[ (t-u)^{-\beta} - t^{-\beta} \right] du - \int_t^{\infty} e^{Au} t^{-\beta} du \right) e_1, \quad t > 0,$$

from which we see that it suffices to argue that (each entry of)

$$t \longmapsto \int_0^t e^{Au} \left[ (t-u)^{-\beta} - t^{-\beta} \right] du$$

belongs to  $L^1$ . Since  $u \mapsto e^{Au}$  is continuous and with all entries decaying exponentially fast as  $u \to \infty$ , this follows from the fact that, for a given  $\gamma > 0$ ,

$$\int_0^\infty \int_0^t e^{-\gamma u} \left| (t-u)^{-\beta} - t^{-\beta} \right| du dt$$

$$\leq \int_0^\infty e^{-\gamma u} \left( \int_u^{u+1} \left[ (t-u)^{-\beta} + t^{-\beta} \right] dt + \beta u \int_1^\infty t^{-\beta-1} dt \right) du < \infty.$$

Here we have used the mean value theorem to establish the inequality

$$\left| (t-u)^{-\beta} - t^{-\beta} \right| \le \beta u (t-u)^{-\beta-1}$$

for 0 < u < t. To show that  $D^{\beta} f \in L^2$ , note that it is the left-sided Riemann–Liouville fractional derivative of f, that is,

$$D^{\beta} f(t) = \frac{1}{\Gamma(1-\beta)} \frac{\mathrm{d}}{\mathrm{d}t} \int_0^t f(t-u) u^{-\beta} \, \mathrm{d}u, \quad t > 0.$$

Consequently, it follows by [27], Theorem 7.1, that the Fourier transform  $\mathcal{F}[D^{\beta}f]$  of f is given by

$$\mathcal{F}[D^{\beta}f](y) = (iy)^{\beta}\mathcal{F}[f](y) = (iy)^{\beta}b^{\top}(iy - A)^{-1}e_1, \quad y \in \mathbb{R},$$

in particular it belongs to  $L^2$  (e.g., by Cramer's rule), and thus  $D^{\beta} f \in L^2$ . By comparing Fourier transforms, we establish that  $(D^{\beta}g) * f = g * (D^{\beta}f)$ , and hence it holds that

$$\int_0^\infty D^{\beta} X_{t-u} f(u) \, du = \int_{\mathbb{R}} (D^{\beta} g) * f(t-r) \, dL_r = \int_0^\infty X_{t-u} D^{\beta} f(u) \, du$$

using Proposition 3.6 and a stochastic Fubini result. This finishes the proof.

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# **Supplementary Material**

**Spectral representations of continuous-time stationary processes** (DOI: 10.3150/18-BEJ1086SUPP; .pdf). The supplement provides an exposition of the spectral representation and related results for continuous-time stationary, centered and square integrable processes.

### References

- [1] Barndorff-Nielsen, O.E. and Basse-O'Connor, A. (2011). Quasi Ornstein-Uhlenbeck processes. *Bernoulli* 17 916–941. MR2817611 https://doi.org/10.3150/10-BEJ311
- [2] Basse-O'Connor, A., Nielsen, M.S., Pedersen, J. and Rohde, V. (2019). Stochastic delay differential equations and related autoregressive models. *Stochastics*. https://doi.org/10.1080/17442508.2019. 1635601
- [3] Basse-O'Connor, A., Nielsen, M.S., Pedersen, J. and Rohde, V. (2019). Multivariate stochastic delay differential equations and CAR representations of CARMA processes. *Stochastic Process. Appl.* 129 4119–4143. MR3997674 https://doi.org/10.1016/j.spa.2018.11.011
- [4] Bennedsen, M. (2015). Rough electricity: A new fractal multi-factor model of electricity spot prices. CREATES Research Paper 42.
- [5] Bennedsen, M., Lunde, A. and Pakkanen, M.S. (2016). Decoupling the short-and long-term behavior of stochastic volatility.
- [6] Beran, J., Feng, Y., Ghosh, S. and Kulik, R. (2013). Long-Memory Processes: Probabilistic Properties and Statistical Methods. Heidelberg: Springer. MR3075595 https://doi.org/10.1007/978-3-642-35512-7
- Bichteler, K. (1981). Stochastic integration and L<sup>p</sup>-theory of semimartingales. Ann. Probab. 9 49–89.
   MR0606798
- [8] Brockwell, P. and Marquardt, T. (2005). Lévy-driven and fractionally integrated ARMA processes with continuous time parameter. *Statist. Sinica* **15** 477–494. MR2190215
- [9] Brockwell, P.J. and Davis, R.A. (2006). *Time Series: Theory and Methods. Springer Series in Statistics*. New York: Springer. Reprint of the second (1991) edition. MR2839251
- [10] Davis, R.A., Nielsen, M.S. and Rohde, V. (2020). Supplement to "Stochastic differential equations with a fractionally filtered delay: A semimartingale model for long-range dependent processes." https://doi.org/10.3150/18-BEJ1086SUPP.
- [11] Delbaen, F. and Schachermayer, W. (1994). A general version of the fundamental theorem of asset pricing. *Math. Ann.* 300 463–520. MR1304434 https://doi.org/10.1007/BF01450498
- [12] Doetsch, G. (1937). Bedingungen für die Darstellbarkeit einer Funktion als Laplace-integral und eine Umkehrformel für die Laplace-Transformation. *Math. Z.* 42 263–286. MR1545675 https://doi.org/10. 1007/BF01160078
- [13] Doukhan, P., Oppenheim, G. and Taqqu, M.S. (eds.) (2003). Theory and Applications of Long-Range Dependence. Boston, MA: Birkhäuser Boston. MR1956041
- [14] Dym, H. and McKean, H.P. (1976). Gaussian Processes, Function Theory, and the Inverse Spectral Problem. Probability and Mathematical Statistics 31. New York: Academic Press. MR0448523
- [15] Granger, C.W.J. and Joyeux, R. (1980). An introduction to long-memory time series models and fractional differencing. *J. Time Series Anal.* 1 15–29. MR0605572 https://doi.org/10.1111/j.1467-9892. 1980.tb00297.x
- [16] Grimmett, G.R. and Stirzaker, D.R. (2001). Probability and Random Processes, 3rd ed. New York: Oxford Univ. Press. MR2059709

- [17] Gripenberg, G. and Norros, I. (1996). On the prediction of fractional Brownian motion. J. Appl. Probab. 33 400–410. MR1385349 https://doi.org/10.2307/3215063
- [18] Gushchin, A.A. and Küchler, U. (2000). On stationary solutions of delay differential equations driven by a Lévy process. *Stochastic Process. Appl.* 88 195–211. MR1767844 https://doi.org/10.1016/ S0304-4149(99)00126-X
- [19] Hosking, J.R.M. (1981). Fractional differencing. *Biometrika* 68 165–176. MR0614953 https://doi.org/10.1093/biomet/68.1.165
- [20] Jusselin, P. and Rosenbaum, M. (2018). No-arbitrage implies power-law market impact and rough volatility. Preprint. Available at arXiv:1805.07134.
- [21] Marquardt, T. (2006). Fractional Lévy processes with an application to long memory moving average processes. *Bernoulli* 12 1099–1126. MR2274856 https://doi.org/10.3150/bj/1165269152
- [22] Mohammed, S.E.A. and Scheutzow, M.K.R. (1990). Lyapunov exponents and stationary solutions for affine stochastic delay equations. *Stoch. Stoch. Rep.* 29 259–283. MR1041039 https://doi.org/10.1080/ 17442509008833617
- [23] Newbold, P. and Agiakloglou, C. (1993). Bias in the sample autocorrelations of fractional noise. Biometrika 80 698–702.
- [24] Pipiras, V. and Taqqu, M.S. (2003). Fractional calculus and its connections to fractional Brownian motion. In *Theory and Applications of Long-Range Dependence* 165–201. Boston, MA: Birkhäuser. MR1956050
- [25] Pipiras, V. and Taqqu, M.S. (2017). Long-Range Dependence and Self-Similarity. Cambridge Series in Statistical and Probabilistic Mathematics 45. Cambridge: Cambridge Univ. Press. MR3729426
- [26] Rajput, B.S. and Rosiński, J. (1989). Spectral representations of infinitely divisible processes. *Probab. Theory Related Fields* 82 451–487. MR1001524 https://doi.org/10.1007/BF00339998
- [27] Samko, S.G., Kilbas, A.A. and Marichev, O.I. (1993). Fractional Integrals and Derivatives: Theory and Applications. Yverdon: Gordon & Breach. Edited and with a foreword by S.M. Nikol'skiĭ, translated from the 1987 Russian original, revised by the authors. MR1347689
- [28] Samorodnitsky, G. (2006). Long range dependence. Found. Trends Stoch. Syst. 1 163–257. MR2379935 https://doi.org/10.1561/0900000004
- [29] Samorodnitsky, G. (2016). Stochastic Processes and Long Range Dependence. Springer Series in Operations Research and Financial Engineering. Cham: Springer. MR3561100 https://doi.org/10.1007/978-3-319-45575-4
- [30] Samorodnitsky, G. and Taqqu, M.S. (1994). Stable Non-Gaussian Random Processes: Stochastic Models with Infinite Variance. Stochastic Modeling. New York: CRC Press. MR1280932
- [31] Sato, K. (1999). Lévy Processes and Infinitely Divisible Distributions. Cambridge Studies in Advanced Mathematics 68. Cambridge: Cambridge Univ. Press. Translated from the 1990 Japanese original, revised by the author. MR1739520

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