ON THE MOMENTS OF SOME ONE-SIDED STOPPING RULES1

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1. Introduction. The moments of stopping rules (or stopping times) have been discussed in [1], [3], and [4], and the following results have been proved. Let x_n be independent random variables with $Ex_n = 0$, $Ex_n^2 = 1$, and $S_n = x_1 + \cdots + x_n$. For c > 0 and $m = 1, 2, \cdots$, define t_m to be the first $n \ge m$ such that $|S_n| > cn^{\frac{1}{2}}$. If $c \ge 1$, then $Et_1 = \infty$. If $P[|x_n| \le K] = 1$ for some $K < \infty$ and $n = 1, 2, \cdots$, then $Et_m < \infty$ for every m if c < 1, $Et_m^2 < \infty$ for every m if $c < 3 - 6^{\frac{1}{2}}$, and $Et_m^2 = \infty$ for all large m if $c \ge 3 - 6^{\frac{1}{2}}$.

In this note, we are interested in the following one-sided stopping rules, instead of the above stated two-sided stopping rules. For c > 0 and $1 > p \ge 0$, define

$$s = \text{first} \quad n \ge 1 \quad \text{such that} \quad S_n \ge cn^p$$
.

One of the results states that, if x_n are independent, $Ex_n = \mu > 0$, and $Ex_n^2 - \mu^2 = \sigma^2 < \infty$, then $Es^2 < \infty$ and

(1)
$$\lim_{c\to\infty} \mu^2 E s^2 / (c^2 E s^{2p}) = \lim \mu E s^2 / (c E s^{1+p}) = 1.$$

When p=0, $Es^2<\infty$ implies that $P[S_1< c, \cdots, S_n< c]=P[s>n]=o(n^{-2})$ as $n\to\infty$, which completes a result of Morimura [9]. Also (1) extends the elementary renewal theorem from first moments to second moments and generalizes some results due to Chow and Robbins [2], Hatori [6], and Heyde [7].

2. The first moment. Let $(\Omega, \mathfrak{F}, P)$ be a probability space and x_n be a sequence of integrable random variables. Let $\mathfrak{F}_1 \subset \mathfrak{F}_2 \subset \cdots \subset \mathfrak{F}$ be Borel fields such that x_n is \mathfrak{F}_n -measurable and $\mathfrak{F}_0 = \{\emptyset, \Omega\}$. Put $S_n = x_1 + \cdots + x_n$, $S_0 = 0$, $m_n = E(x_n \mid \mathfrak{F}_{n-1})$ and $T_n = \sum_{1}^{n} m_j$. Assume that for some constant $\infty > \mu > 0$ and for some null set N,

(2)
$$\lim_{n\to\infty} T_n/n = \mu, \text{ uniformly on } \Omega - N.$$

For c > 0 and $1 > p \ge 0$, define

$$s =$$
first $n \ge 1$ such that $S_n \ge cn^p$.

THEOREM 1. (i) If for some $0 < \delta < \mu/3$, $P[x_n \le m_n + n\delta] = 1$ for all large n, then $Es < \infty$.

(ii) If
$$E([(x_n - m_n)^+]^{\alpha} | \mathfrak{F}_{n-1}) \leq K < \infty$$
 for some $\alpha > 1$ and

$$E(|x_n-m_n|\mid \mathfrak{F}_{n-1})\leq K<\infty,$$

then $ES < \infty$ and

(3)
$$\lim_{c\to\infty} \mu E s/(cEs^p) = 1 = \lim_{c\to\infty} E S_s/(cEs^p).$$

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PROOF. (i) Set $t = \min(s, k)$ for $k = 1, 2, \dots$. Then by the Wald identity for martingales (see [5], p. 302; or [3]),

$$ET_t = ES_t = E(S_{t-1} + x_t) \le cEt^p + E(m_t + \delta t) + O(1).$$

Let $0 < \epsilon < \delta$. As $k \to \infty$, by (2)

$$ET_t \ge (\mu - \epsilon)Et + O(1), \qquad Em_t = O(1) + o(Et).$$

Hence

$$(\mu - \epsilon)Et \le cEt^p + \delta Et + O(1) + o(Et),$$
$$\int_{\{s \le k\}} s \, dP + kP[s > k] = Et = O(1),$$

as $k \to \infty$. Therefore $P[s < \infty] = 1$ and $Es < \infty$.

(ii) For any $0 < \delta < \mu/6$, define $x_n' = \min(x_n, m_n + n\delta)$, $m_n' = E(x_n' \mid \mathfrak{F}_{n-1})$, and $T_n' = m_1' + \cdots + m_n'$. Let I(A) be the indicator function of the set A. Then

$$0 \leq m_n - m_n' = E((x_n - m_n - n\delta)I[x_n > m_n + n\delta] | \mathfrak{F}_{n-1})$$

$$(4) \qquad \leqq E((x_{n}-m_{n})I[x_{n}>m_{n}+n\delta]\mid \mathfrak{F}_{n-1})$$

$$\leqq E^{1/\alpha}([(x_{n}-m_{n})^{+}]^{\alpha}\mid \mathfrak{F}_{n-1})P^{1/\alpha'}(x_{n}-m_{n}>n\delta\mid \mathfrak{F}_{n-1}) \quad (\alpha+\alpha'=\alpha\alpha')$$

$$\leqq K(n\delta)^{-\alpha/\alpha'}.$$

Therefore $\lim_{n\to\infty} T_n'/n = \mu$ uniformly on $\Omega-N$ and $P[x_n' < m_n' + 2n\delta] = 1$ for all large n. Define

$$t = \text{first } n \ge 1 \text{ such that } x_1' + \cdots + x_n' \ge cn^p.$$

Then $s \le t$. By (i), $Et < \infty$. Therefore $Es < \infty$ and it follows by the Wald identity again ([5], p. 302; or [3]) that

$$(5) E(cs^p + x_s) \ge ES_s = ET_s \ge cEs^p.$$

Let $Z_n = \sum_{1}^{n} [(x_j - m_j)^+]^{\alpha}$. Then by Lemma 6 of [3],

(6)
$$E^{\alpha}(x_s - m_s)^+ \leq EZ_s = E\sum_{i=1}^{s} E([(x_i - m_j)^+]^{\alpha} | \mathfrak{F}_{j-1}) \leq KEs.$$

Since (2) implies that as $c \to \infty$, $Em_s = O(1) + o(Es)$ and $ET_s = O(1) + (\mu + o(1))Es$, we have

(7)
$$Ex_s = O(E^{1/\alpha}s) + o(Es) + O(1)$$

from (6); and

$$\lim_{c\to\infty} \mu E s/(cEs^p) = \lim E T_s/(cEs^p) = \lim E S_s/(cEs^p) = 1$$

from (5) and (7), since $\lim_{s\to\infty} Es = \infty$. The proof is completed.

When p = 0, part (ii) of Theorem 1 reduces to an elementary renewal theorem, which was proved in [2], in a slightly restricted form by requiring that $m_n = E(x_n)$ for each n.

3. The second moment. Assume that $Ex_n^2 < \infty$ for each n, let

$$V_n = \sum_{1}^{n} E((x_j - m_j)^2 | \mathfrak{F}_{j-1})$$

for $n=1, 2, \cdots$, and define s as before. For a random variable y, put $||y||=(Ey^2)^{\frac{1}{2}}$.

THEOREM 2. If (2) holds and $E((x_n - m_n)^2 \mid \mathfrak{F}_{n-1}) \leq K < \infty$, then $Es^2 < \infty$, $ES_s^2 < \infty$, and as $c \to \infty$,

(8)
$$ES_s^2 + ET_s^2 = EV_s + 2ES_sT_s,$$

(9)
$$\lim ES_s^2/ET_s^2 = 1,$$

(10)
$$\lim \mu^2 E s^2 / (c^2 E s^{2p}) = 1,$$

(11)
$$\lim ES_s^2/(c^2Es^{2p}) = 1,$$

(12)
$$\lim \mu E s^2 / (cE s^{1+p}) = 1.$$

PROOF. (i) First, assume that for some $0 < \delta < \mu/8$ and $0 < M < \infty$, $P[x_n \le m_n + n\delta + M] = 1$ for all large n. Set $t = \min(s, k)$ for $k = 1, 2, \cdots$. Then by Theorem 1 and Lemma 6 of [3], $E(S_t - T_t)^2 = EV_t \le KEt$. Hence by Schwarz inequality

(13)
$$E_k S_t^2 + E_k T_t^2 \le K E_k t + 2 \|T_t\|_k \cdot \|S_t\|_k,$$

where $E_k y = \int_{[s \le k]} y \, dP$ and $||y||_k = (E_k y^2)^{\frac{1}{2}}$ for a random variable y. Assume, on the contrary, that $Es^2 = \infty$. Then $\lim_{k\to\infty} E_k t^2 = \infty$ and (2) implies that

$$E_k m_t^2 = O(1) + o(E_k t^2) = o(E_k t^2),$$

as $k \to \infty$. Hence

(14)
$$||S_t||_k \le ||ct^p + m_t + \delta t + M||_k + O(1) \le c||t^p||_k + \delta ||t||_k + o(||t||_k)$$

= $(\delta + o(1))||t||_k$;

and from (2),

(15)
$$E_k T_t^2 = O(1) + (\mu^2 + o(1)) E_k t^2 = (\mu^2 + o(1)) E_k t^2.$$

By (13), (14) and (15), we have

$$1 + E_k S_t^2 / E_k T_t^2 \le O(\|t\|_k^{-1}) + 2\|S_t\|_k / \|T_t\|_k$$

$$\le O(\|t\|_k^{-1}) + (2\delta + o(1)) / \mu = 2\delta / \mu + o(1).$$

Since $\delta < \mu/8$, we have a contradiction when k is large. Therefore $Es^2 < \infty$ and $E_k t^2 = O(1)$. Hence

$$||S_t||_k \le ||ct^p + m_t + \delta t + M||_k + O(1)$$

$$\leq O(1) + ||m_t||_k \leq O(1) + o(E_k t^2) = O(1).$$

By Fatou's lemma, $ES_s^2 < \infty$ and from (13), $ET_s^2 < \infty$.

(ii) For the general case, let $x_n' = \min (x_n, m_n + n\delta + M)$ for arbitrary constants $\infty > M > 0$ and $0 < \delta < \mu/16$. Define m_n', T_n' and t as in the proof of part (ii) of Theorem 1. Then by (4) (for $\alpha = 2$), $0 \le m_n - m_n' \le K(n\delta)^{-1}$. Hence $P[x_n' \le m_n' + 2n\delta + M] = 1$ for all large n, and $\lim T_n'/n = \mu$ uniformly on $\Omega - N$. It is not too difficult to see that

$$E((x_n - m_n)^2 \mid \mathfrak{F}_{n-1}) - E((x_n' - m_n')^2 \mid \mathfrak{F}_{n-1}) \ge 0.$$

Therefore $E((x_n'-m_n')^2 \mid \mathfrak{F}_{n-1}) \leq K$. Since $t \geq s$ and from part (i) $Et^2 < \infty$, we have $Es^2 < \infty$. By Theorem 1 and Lemma 6 of [3] again,

$$(16) E(S_s - T_s)^2 = EV_s \le KEs.$$

For $\epsilon > 0$, (2) implies that there exists a constant $\infty > L > 0$ such that $ET_s^2 \le L + (\mu^2 + \epsilon)Es^2$. Hence $ET_s^2 < \infty$ and from (16), $ES_s^2 < \infty$. Thus (8) follows. Now by (16),

$$|ES_s^2 - ET_s^2| \le E|S_s^2 - T_s^2| \le ||S_s - T_s|| \cdot ||S_s + T_s|| \le (KEs)^{\frac{1}{2}} ||S_s + T_s||.$$

Since $ES_s^2 \ge c^2 E s^{2p}$, from (3)

$$|1 - ET_s^2/ES_s^2| \le (KEs/ES_s^2)^{\frac{1}{2}}(1 + ||T_s||/||S_s||) = o(1) + o(||T_s||/||S_s||)$$
 as $c \to \infty$. Hence (9) follows.

Since (2) implies that $ET_s^2 = O(1) + (\mu^2 + o(1))Es^2$ as $c \to \infty$, from (9)

(17)
$$\lim_{c\to\infty} \mu^2 E s^2 / E T_s^2 = 1 = \lim \mu^2 E s^2 / E S_s^2.$$
 Let $Z_n = \sum_{i=1}^n (x_i - m_i)^2$. Applying Lemma 6 of [3], we have

$$E(x_s-m_s)^2 \leq EZ_s = E\sum_{j=1}^s E((x_j-m_j)^2 \mid \mathfrak{F}_{j-1}) \leq KEs.$$

From (2), $Em_s^2 = O(1) + o(Es^2) = o(Es^2)$ as $c \to \infty$. Hence

(18)
$$Ex_s^2 = E(x_s - m_s + m_s)^2 = o(Es^2), \quad ||x_s|| = o(||s||).$$

Now from (18), as $c \to \infty$,

$$(19) c||s^p|| \leq ||S_s|| \leq ||cs^p + x_s|| \leq c||s^p|| + ||x_s|| = c||s^p|| + o(||s||).$$

Therefore (10) follows from (17) and (19), and (11) follows from (17) and (10). Now $ET_sS_s = O(1) + (\mu + o(1))EsS_s$ as $c \to \infty$. By the definition of s and (18), as $c \to \infty$,

$$(20) \quad cEs^{1+p} \le EsS_s \le cEs^{1+p} + Esx_s \le cEs^{1+p} + ||s|| \cdot ||x_s|| \le cEs^{1+p} + o(Es^2).$$

Since $EV_s \leq KEs$, from (8), (9), (10), and (11), $\lim ES_sT_s/(\mu^2Es^2) = 1$. Hence $\lim EsS_s/(\mu Es^2) = 1$ and then (20) implies (12).

4. Corollaries and comments. In this section we assume that x_n is a sequence of random variables and p=0. Define S_n , m_n , T_n , \mathfrak{F}_n and s as in Section 2. Corollary 1. If (2) holds, $E(x_n^2) < \infty$ for each n, and

(21)
$$E((x_n-m_n)^2 \mid \mathfrak{F}_{n-1}) \leq K < \infty,$$

386 Y. S. CHOW

then $Es^2 < \infty$ and

(22)
$$\lim_{c\to\infty} Es^{\alpha}/c^{\alpha} = \mu^{-\alpha} \quad \text{for} \quad 0 \le \alpha \le 2.$$

PROOF. Since (21) implies $E(x_n - m_n)^2 \le K$, from (2) and (21) it follows [8] that $\lim_{n \to \infty} S_n/n = \mu$ a.e. Hence

$$1 \leq \lim \inf_{c \to \infty} S_s/c \leq \lim \inf \mu s/c \leq \lim \sup \mu s/c = \lim \sup \mu(s-1)/c$$
$$= \lim \sup S_{s-1}/c \leq 1.$$

Therefore $\lim s/c = \mu^{-1}$ a.e. Since p = 0, from (10) we have that $E(s/c)^2 \le M < \infty$ for all large c. Hence (see [5], p. 629) for every $0 \le \alpha < 2$, $(s/c)^{\alpha}$ is uniformly integrable and

(23)
$$\lim_{c\to\infty} E|\mu^{-1} - s/c|^{\alpha} = 0, \quad \lim Es^{\alpha}/c^{\alpha} = \mu^{-\alpha}.$$

Thus (22) follows from (23) and (10).

COROLLARY 2. Let x_n be a sequence of independent, identically distributed random variables such that $Ex_1 > 0$ and $E(x_1 - Ex_1)^2 < \infty$. Then for every c > 0, as $n \to \infty$,

(24)
$$P[S_1 < c, \cdots, S_n < c] = o(n^{-2}).$$

PROOF. Since $[s > n] = [S_1 < c, \dots, S_n < c]$, $Es^2 < \infty$ implies (24) and thus Corollary 2 follows from Corollary 1.

(22) has been proved by Hatori [6] for every $\alpha > 0$, by requiring, in addition to the assumptions of Corollary 1, that x_n be independent, $P[x_n \ge 0] = 1$ and $m_n \ge L > 0$ for each n.

Under the conditions of Corollary 2, Morimura [9] proves that

$$P[S_1 < c, \cdots, S_n < c] = O(n^{-\delta})$$

for $0 \le \delta < (1+5^{\frac{1}{2}})/2$ and that there exists an example such that for some D > 0 and for each $\epsilon > 0$, $P[S_1 < c, \dots, S_n < c] \ge Dn^{-2-\epsilon}$ when n is large enough. Thus (24) is the best possible. Clearly, Corollary 2 completes Morimura's work.

The counter example in [9] satisfies the condition $Es^{2+\epsilon} = \infty$ for every $\epsilon > 0$, since $P[s > n] \neq o(n^{-2-\epsilon})$. Therefore (22) can not be extended to the cases where $\alpha > 2$, without some conditions such as $P[x_n \ge 0] = 1$ imposed in [6].

COROLLARY 3. Let x_n be a sequence of independent, identically distributed random variables such that $0 < Ex_1 = \mu \le \infty$ and $E(x_1^-)^2 < \infty$. Then (22) holds.

PROOF. Let $0 \le \alpha \le 2$. For $0 < \mu' < \mu$, choose $0 < M < \infty$ so that $Ex_1' = D > \mu'$, where $x_n' = \min(x_n, M)$. Define $S_n' = x_1' + \cdots + x_n'$ and $t = \text{first } n \ge 1$ such that $S_n' \ge c$. By Corollary 1, $\limsup Es^{\alpha}/c^{\alpha} \le \lim Et^{\alpha}/c^{\alpha} = D^{-\alpha}$. Since μ' is arbitrary, $\limsup_{c \to \infty} Es^{\alpha}/c^{\alpha} \le \mu^{-\alpha}$. Hence (22) holds for $\mu = \infty$ and $E(s/c)^2 \le M < \infty$ for all large c. Now assume $0 < \mu < \infty$. By the strong law of large numbers, $\lim_n S_n/n = \mu$ a.e. Hence for $0 \le \alpha < 2$, as in the proof of

Corollary 1, (23) holds and therefore (22) holds. For the case $\alpha=2$, by Theorem 2 of [2] $\lim_c Es/c = \mu^{-1}$. Hence $\limsup_c E(s/c)^2 \ge \lim_c E^2(s/c) = \mu^2$. Therefore $\lim_c E(s/c)^2 = \mu^{-2}$ and the proof is completed.

Corollary 3 has been recently proved by Heyde [7] under the stronger condition that $E(x_1^-)^3 < \infty$.

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