

INVARIANCE OF MAXIMUM LIKELIHOOD ESTIMATORS

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One of the distinguishing features of the method of maximum likelihood in statistical estimation is the fact that it enjoys a certain invariance property. Briefly stated, if $\hat{\theta}$ is a maximum likelihood estimator for θ , then $u(\hat{\theta})$ is a maximum likelihood estimator for $u(\theta)$ where u is some function of θ . Some textbooks on the subject avoid any explicit mention of properties that u must possess in order for invariance to hold. When a proof of the property is given, it is at least assumed, either explicitly or implicitly that u is 1-1 thereby defining a unique inverse.

Now if the assumption that u be 1-1 is really necessary, then the invariance principle could not be invoked to find the maximum likelihood estimator for even as common a case as the variance, $p(1 - p)$, of a Bernoulli random variable. Indeed, there may be some doubt as to the meaning of maximum likelihood in such a case. The purpose of this note is to point out that the notion of a maximum likelihood estimator for $u(\theta)$ when u is not 1-1 can and should be made explicit. The method used for accomplishing this task has the desirable feature that it coincides with the usual method employed when u is 1-1.

Suppose that parameter θ is restricted to lie in some set Θ and let $L(\theta)$ denote the likelihood function, a mapping from Θ to the real line. Assume that the maximum likelihood estimator $\hat{\theta}$ exists so that, $\hat{\theta} \in \Theta$ and $L(\hat{\theta}) \geq L(\theta)$ for all $\theta \in \Theta$. Let u be an arbitrary transformation from Θ to some set Λ . For convenience, we suppose that Λ is the range of u and we adopt the notation $\lambda = u(\theta)$.

Since u is a function, $u(\hat{\theta})$ is a unique member, say $\hat{\lambda}$, of Λ . For each $\lambda \in \Lambda$, let $\Theta_\lambda = \{\theta; u(\theta) = \lambda\}$ and $M(\lambda) = \sup_{\theta \in \Theta_\lambda} L(\theta)$. Then M is a real-valued function on Λ to be called the *likelihood function induced by u* . Clearly, $M(\hat{\lambda}) = L(\hat{\theta})$ and the fact that $\hat{\lambda}$ maximizes M is a trivial consequence of the inequality $M(\lambda) = \sup_{\theta \in \Theta_\lambda} L(\theta) \leq \sup_{\theta \in \Theta} L(\theta) = L(\hat{\theta}) = M(\hat{\lambda})$ for all $\lambda \in \Lambda$. In this sense, it is reasonable to call $\hat{\lambda} = u(\hat{\theta})$ the maximum likelihood estimator for $u(\theta)$.

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