## SOME EXPANSION THEOREMS FOR STOCHASTIC PROCESSES, II

By Hirohisa Hatori

**1.** Let  $F(\lambda)$  be the spectral function of a continuous (weakly) stationary process  $\mathcal{E}(t)$  with mean zero, and consider  $X(t)=f(t)+\mathcal{E}(t), -\infty < t < \infty$ , where f(t) is a numerical valued function. Assume that

(i) 
$$\int_{-\infty}^{\infty} |\lambda|^{2r+2\alpha} dF(\lambda) < \infty,$$

(ii) H(u) is of bounded variation in  $(-\infty, \infty)$ ,

(iii) 
$$\int_{-\infty}^{\infty} |u|^{r+\alpha} |dH(u)| < \infty,$$

(iv)  $|f(u)| \leq C (1+|u|^{r+\alpha})$  for all u and a positive constant C, and

(v) 
$$f(t+u) = \sum_{k=0}^{r} \frac{f^{(k)}(t)}{k!} u^{k} + o(|u|^{r+\alpha}),$$

where r is a non-negative integer and  $\alpha$  is a constant with  $0 \leq \alpha < 1$ . From the conditions (i)-(v) it follows that

(1.1) 
$$E\left\{\left\|\int_{-\infty}^{\infty} X\left(t-\frac{s}{n}\right) dH(s) - \sum_{k=0}^{r} \frac{(-1)^{k} X^{(k)}(t)}{k! n^{k}} \int_{-\infty}^{\infty} s^{k} dH(s)\right\|^{2}\right\} = o\left(\frac{1}{n^{2r+2\alpha}}\right).$$

This is an expansion theorem for the integral

$$\int_{-\infty}^{\infty} X\left(t-\frac{s}{n}\right) dH(s),$$

which has been treated by Kawata [3] for r=0 and  $\alpha=1/2$  with somewhat different conditions, and extended by the author [1] for  $r=0, 1, 2, \cdots$  and  $0 \le \alpha < 1$  with the above conditions (i)-(v). In this paper, we shall show (1.1) for  $X(t)=f(t)+\phi(t)\mathcal{E}(t)$ , where  $\phi(u)$  is a numerical valued function. If  $\phi(s)>0$ .  $-\infty < s < \infty$ , then, for this process X(t), the correlation coefficient of X(t) and X(s) is a function of t-s only. In section 2, Taylor expansion of  $\mathcal{E}(t)$  is discussed and, in section 3, the expansion theorem for

$$\int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) dH(s)$$

is given, where  $X(t) = f(t) + \phi(t) \mathcal{E}(t), -\infty < t < \infty$ .

Received March 15, 1963.

## HIROHISA HATORI

2. Let X(t) be a stochastic process with  $E\{|X(t)|^2\} < \infty, -\infty < t < \infty$ . If  $\lim_{\tau \to t} X(\tau) = X(t)$  i.e.  $E\{|X(\tau) - X(t)|^2\} \to 0$  as  $\tau \to t$ ,

then X(t) is called to be continuous at t. Let K(t) be of bounded variation in any finite interval and consider a division of an interval (A, B):

$$\Delta: A = t_0 < t_1 < \cdots < t_n = B.$$

If

1.i.m. 
$$\sum_{k=1}^{n} X(\tau_k) [K(t_k) - K(t_{k-1})] = S \qquad \left( \max_{k=1,2,\dots,n} (t_k - t_{k-1}) \to 0 \right),$$

where  $t_{k-1} \leq \tau_k \leq t_k$ ,  $(k=1, 2, \dots, n)$ , then S is denoted as

$$\int_{A}^{B} X(t) \ dK(t).$$

We can easily prove the following

LEMMA 1. Being continuous in an interval [A, B], X(t) is uniformly continuous in [A, B]. Moreover,

$$\int_{A}^{B} X(t) \, dK(t)$$

exists in the above sense.

LEMMA 2. If  $X^{(k)}(t)$   $(k=1, 2, \dots, r)$  exist in the sense that

$$X^{(k)}(t) = \lim_{h \to 0} \frac{X^{(k-1)}(t+h) - X^{(k-1)}(t)}{h} \qquad (k=1, 2, \dots, r),$$

where  $X^{(0)}(t) \equiv X(t)$ , and  $X^{(k)}(t)$   $(k=1, 2, \dots, r)$  are cantinuous in [a, b], then

$$X(b) = X(a) + \frac{X'(a)}{1!}(b-a) + \dots + \frac{X^{(r-1)}(a)}{(r-1)!}(b-a)^{r-1}$$

(2.1)

$$+\frac{1}{(r-1)!}\int_{a}^{b}(b-t)^{r-1}X^{(r)}(t)\,dt \qquad (a.s.).$$

*Proof.* The existence of the integral in (2.1) is ensured by Lemma 1. Y being any random variable with  $E\{|Y|^2\} < \infty$ , the numerical valued function  $\varphi(t) \equiv E\{X(t) \cdot \overline{Y}\}$  is differentiable r times and  $\varphi^{(r)}(t) = E\{X^{(r)}(t) \cdot \overline{Y}\}$  is continuous in [a, b]. So we have

$$\varphi(b) = \sum_{k=0}^{r-1} \frac{\varphi^{(k)}(a)}{k!} (b-a)^k + \frac{1}{(r-1)!} \int_a^b (b-t)^{r-1} \varphi^{(r)}(t) dt$$

or

$$E\left\{\left[X(b) - \sum_{k=0}^{r-1} \frac{X^{(k)}(a)}{k!} (b-a)^k - \frac{1}{(r-1)!} \int_a^b (b-t)^{r-1} X^{(r)}(t) dt\right] \cdot \bar{Y}\right\} = 0.$$

Choosing

$$X(b) - \sum_{k=0}^{r-1} \frac{X^{(k)}(a)}{k!} (b-a)^k - \frac{1}{(r-1)!} \int_a^b (b-t)^{r-1} X^{(r)}(t) dt$$

as Y, we have

$$E\left\{\left|X(b)-\sum_{k=0}^{r-1}\frac{X^{(k)}(a)}{k!}(b-a)^{k}-\frac{1}{(r-1)!}\int_{a}^{b}(b-t)^{r-1}X^{(r)}(t)\,dt\,\right|^{2}\right\}=0,$$

which implies (2.1).

In the following, let  $\mathcal{E}(t)$ ,  $-\infty < t < \infty$ , be a continuous (weakly) stationary stochastic process with  $E\{\mathcal{E}(t)\}=0$  for all t. We note that  $\mathcal{E}(t)$  is continuous in the sense stated at the beginning of this section, if and only if the covariance function  $\rho(u)$  of  $\mathcal{E}(t)$  is continuous.

THEOREM 1. Let  $F(\lambda)$  be the spectral function of  $\mathcal{E}(t)$ . If

(2.2) 
$$\int_{-\infty}^{\infty} |\lambda|^{2\tau+2\alpha} dF(\lambda) < \infty,$$

where r is a non-negative integer and  $\alpha$  is a constant with  $0 \leq \alpha < 1$ , then

(2.3) 
$$E\left\{\left| \mathcal{E}(t+u) - \sum_{k=0}^{r} \frac{\mathcal{E}^{(k)}(t)}{k!} u^{k} \right|^{2} \right\} = o(|u|^{2r+2\alpha}) \text{ as } u \to 0.$$

*Proof.* It is well known that, under the assumption (2.2),  $\mathcal{E}^{(k)}(t)$  ( $k=1, 2, \dots, r$ ) exist and are continuous (weakly) stationary processes whose covariance functions are

$$\rho_k(t) = \int_{-\infty}^{\infty} \lambda^{2k} e^{it\lambda} dF(\lambda)$$

respectively. Hence, by Lemma 2, we have with probability 1 that

(2.4)  

$$R \equiv \mathcal{E}(t+u) - \sum_{k=0}^{r} \frac{\mathcal{E}^{(k)}(t)}{k!} u^{k}$$

$$= \frac{1}{(r-1)!} \int_{t}^{t+u} (t+u-\tau)^{r-1} \mathcal{E}^{(r)}(\tau) d\tau - \frac{\mathcal{E}^{(r)}(t)}{r!} u^{r}$$

$$= \frac{1}{(r-1)!} \int_{t}^{t+u} (t+u-\tau)^{r-1} \left[ \mathcal{E}^{(r)}(\tau) - \mathcal{E}^{(r)}(t) \right] d\tau$$

and so

$$E\{|R|^{2}\} = \frac{1}{[(r-1)!]^{2}} E\left\{\int_{t}^{t+u} (t+u-\tau)^{r-1} \left[\mathcal{E}^{(r)}(\tau) - \mathcal{E}^{(r)}(t)\right] d\tau \\ \cdot \int_{t}^{t+u} (t+u-\sigma)^{r-1} \overline{\left[\mathcal{E}^{(r)}(\sigma) - \mathcal{E}^{(r)}(t)\right]} d\sigma\right]$$
$$= \frac{1}{[(r-1)!]^{2}} \int_{t}^{t+u} \int_{t}^{t+u} (t+u-\tau)^{r-1} (t+u-\sigma)^{r-1}$$

$$\begin{split} \cdot E\{[\mathcal{E}^{(r)}(\tau) - \mathcal{E}^{(r)}(t)] \overline{[\mathcal{E}^{(r)}(\sigma) - \mathcal{E}^{(r)}(t)]}\} d\tau d\sigma \\ &= \frac{1}{[(r-1)!]^2} \int_{t}^{t+u} \int_{t}^{t+u} \{\rho_r(\tau-\sigma) - \rho_r(\tau-t) - \rho_r(t-\sigma) + \rho_r(0)\} \\ \cdot (t+u-\tau)^{r-1}(t+u-\sigma)^{r-1} d\tau d\sigma \\ &= \frac{1}{[(r-1)!]^2} \int_{t}^{t+u} \int_{t}^{t+u} \left[ \int_{-\infty}^{\infty} \{e^{i(\tau-\sigma)\lambda} - e^{i(\tau-t)} - e^{i((t-\sigma)\lambda} + 1\} \right] \\ \cdot \lambda^{2r} dF(\lambda) \right] (t+u-\tau)^{r-1}(t+u-\sigma)^{r-1} d\tau d\sigma \\ &= \frac{1}{[(r-1)!]^2} \int_{t}^{t+u} \int_{t}^{t+u} \left[ \int_{-\infty}^{\infty} (e^{i\tau\lambda} - e^{it\lambda}) \overline{(e^{i\sigma\lambda} - e^{it\lambda})} \right] \\ \cdot \lambda^{2r} dF(\lambda) \left] (t+u-\tau)^{r-1} d\tau d\sigma \\ &= \frac{1}{[(r-1)!]^2} \int_{-\infty}^{\infty} \left| \int_{t}^{t+u} (e^{i\tau\lambda} - e^{it\lambda}) (t+u-\tau)^{r-1} d\tau \right|^2 \lambda^{2r} dF(\lambda) \\ &= \frac{1}{[(r-1)!]^2} \int_{-\infty}^{\infty} \left| \int_{0}^{u} (e^{i\xi\lambda} - 1) (u-\xi)^{r-1} d\xi \right|^2 \lambda^{2r} dF(\lambda) \\ &\leq \frac{1}{[(r-1)!]^2} \int_{-\infty}^{\infty} \left\{ \int_{0}^{|u|} |e^{i\xi\lambda} - 1|^{1-\alpha}\xi^{\alpha}(|u| - \xi)^{r-1} d\xi \right\}^2 |\lambda|^{2r+2\alpha} dF(\lambda). \end{split}$$

This estimation implies with Minkowski's inequality that

(2.6) 
$$\sqrt{E\{|R|^2\}} \leq \frac{1}{(r-1)!} \int_0^{|u|} \left\{ \int_{-\infty}^{\infty} |e^{i\xi\lambda} - 1|^{2(1-\alpha)} |\lambda|^{2r+2\alpha} dF(\lambda) \right\}^{\frac{1}{2}} \xi^{\alpha}(|u| - \xi)^{r-1} d\xi.$$

Since  $|e^{i\xi_{\lambda}}-1|^{2(1-\alpha)} \rightarrow 0$  as  $\xi \rightarrow 0$  and  $|e^{i\xi_{\lambda}}-1|^{2(1-\alpha)} \leq 2^{2(1-\alpha)}$ , by the assumption (2.2) and Lebesgue's theorem it holds that

(2.7) 
$$\lim_{\xi \to 0} \int_{-\infty}^{\infty} |e^{i\xi\lambda} - 1|^{2(1-\alpha)} |\lambda|^{2r+2\alpha} dF(\lambda) = 0.$$

Hence, for any positive number  $\varepsilon$ , there exists a positive number  $\eta$  such that

$$\left\{\int_{-\infty}^{\infty} |e^{i\xi\lambda} - 1|^{2(1-\alpha)} |\lambda|^{2\tau+2\alpha} dF(\lambda)\right\}^{\frac{1}{2}} < \varepsilon \quad \text{for} \quad |\xi| < \eta.$$

And so, we have

(2.8) 
$$\sqrt{E\{|R|^2\}} < \frac{\varepsilon}{(r-1)!} \int_0^{|u|} \xi^{\alpha}(|u|-\xi)^{r-1} d\xi < \frac{\varepsilon}{(r-1)!} |u|^{r+\alpha} \text{ for } 0 < |u| < \eta,$$

which proves Theorem 1. (The proof of the case r=0 will similarly be done with slight modifications.)

REMARK 1. In the case  $\alpha = 0$ , we can prove (2. 3) for X(t) in Lemma 2. Because, choosing a positive number  $\delta$  such that

EXPANSION THEOREMS FOR STOCHASTIC PROCESSES, II

$$\sqrt{E\{|X^{(r)}(\tau) - X^{(r)}(t)|^2\}} < \varepsilon$$
 for  $|\tau - t| < \delta$ ,

we have

$$\begin{split} \sqrt{E\{|R|^2\}} &\leq \frac{1}{(r-1)!} \left| \int_{t}^{t+u} \sqrt{E\{|X^{(r)}(\tau) - X^{(r)}(t)|^2\}} \cdot |t+u-\tau|^{r-1} d\tau \right| \\ &\leq \frac{\varepsilon}{(r-1)!} \left| \int_{t}^{t+u} |t+u-\tau|^{r-1} d\tau \right| \leq \frac{\varepsilon}{(r-1)!} |u|^r. \end{split}$$

In the following, let f(t) and  $\phi(t)$  be numerical valued functions.

THEOREM 2. In addition to the assumptions of Theorem 1, we assume for some fixed t that

(2.9) 
$$f(t+u) = \sum_{k=0}^{r} \frac{f^{(k)}(t)}{k!} u^{k} + o(|u|^{r+\alpha}) \quad as \quad u \to 0$$

and

(2.10) 
$$\phi(t+u) = \sum_{k=0}^{r} \frac{\phi^{(k)}(t)}{k!} u^{k} + o(|u|^{r+\alpha}) \quad as \quad u \to 0.$$

Putting  $X(s) = f(s) + \phi(s) \mathcal{E}(s), -\infty < s < \infty$ , we have

(2.11) 
$$E\left\{\left\|X(t+u)-\sum_{k=0}^{r}\frac{X^{(k)}(t)}{k!}u^{k}\right\|^{2}\right\}=o(|u|^{2r+2\alpha}) \quad as \quad u\to 0.$$

*Proof.* It can be easily seen that  $X^{(k)}$   $(k=1, 2, \dots, r)$  exist and

$$X^{(k)}(t) = f^{(k)}(t) + \sum_{\nu=0}^{k} \binom{k}{\nu} \phi^{(\nu)}(t) \mathcal{E}^{(k-\nu)}(t) \quad (k=1, 2, \dots, r).$$

Putting

$$f(t+u) - \sum_{k=0}^{r} \frac{f^{(k)}(t)}{k!} u^{k} = \mathcal{I}_{1}, \ \phi(t+u) - \sum_{k=0}^{r} \frac{\phi^{(k)}(t)}{k!} u^{k} = \mathcal{I}_{2} \text{ and } \mathcal{E}(t+u) - \sum_{k=0}^{r} \frac{\mathcal{E}^{(k)}(t)}{k!} u^{k} = \mathcal{I}_{2}$$

by (2.9), (2.10) and Theorem 1 we have

(2.12) 
$$| \Delta_1 | = o(| u |^{r+\alpha}), | \Delta_2 | = o(| u |^{r+\alpha})$$

and

(2.13) 
$$E\{|R|^2\}=o(|u|^{2r+2\alpha})$$
 as  $u\to 0$ .

Since

by (2.12) and (2.13) we have

$$\sqrt{E\left\{\left|X(t+u)-\sum_{k=0}^{r}\frac{X^{(k)}(t)}{k!}u^{k}\right|^{2}\right\}} \\
(2.15) \leq |\mathcal{L}_{1}| + \sum_{\substack{0 \leq k, l \leq r \\ k+l > r}} \frac{|\phi^{(k)}(t)|}{k!}\frac{\sqrt{\rho_{k}(0)}}{l!}|u|^{k+l} + |\mathcal{L}_{2}|\sum_{k=0}^{r}\frac{\sqrt{\rho_{k}(0)}}{k!} \cdot |u|^{k} \\
+ \sqrt{E\{|R|^{2}\}}\left\{\sum_{k=0}^{r}\frac{|\phi^{(k)}(t)|}{k!}|u|^{k} + |\mathcal{L}_{2}|\right\}$$

 $= o(|u|^{r+\alpha}),$ 

which proves Theorem 2.

In the following, we shall give some examples as applications of Lemma 2.

EXAMPLE 1. For  $\mathcal{E}(t)$  in Theorem 1, we assume that

(2.16) 
$$\int_{-\infty}^{\infty} \lambda^{2n} dF(\lambda) < \infty \qquad (n=1, 2, \cdots).$$

Since

$$E\{|R_{n}|^{2}\} = \frac{1}{[(n-1)!]^{2}} \int_{-\infty}^{\infty} \left| \int_{0}^{u} (e^{i\xi\lambda} - 1)(u-\xi)^{n-1}d\xi \right|^{2} \lambda^{2n} dF(\lambda)$$
$$= \int_{-\infty}^{\infty} \left| e^{iu\lambda} - \sum_{k=0}^{n} \frac{(iu\lambda)^{k}}{k!} \right|^{2} dF(\lambda),$$

where

$$R_n = \mathcal{E}(t+u) - \sum_{k=0}^n \frac{\mathcal{E}^{(k)}(t)}{k!} u^k,$$

we get that

(2.17) 
$$\mathscr{E}(t+u) = \sum_{n=0}^{\infty} \frac{\mathscr{E}^{(n)}(t)}{n!} u^n \equiv \lim_{n \to \infty} \sum_{k=0}^n \frac{\mathscr{E}^{(k)}(t)}{k!} u^k$$

if and only if

(2.18) 
$$\lim_{n\to\infty}\int_{-\infty}^{\infty}\left|e^{iu\lambda}-\sum_{k=0}^{n}\frac{(iu\lambda)^{k}}{k!}\right|^{2}dF(\lambda)=0.$$

EXAMPLE 2. For  $\mathcal{E}(t)$  in Theorem 1, we assume that there exists a positive constant  $\delta$  such that

(2.19) 
$$\int_{-\infty}^{\infty} \lambda^{2n} dF(\lambda) = o\left(\frac{[n!]^2}{\delta^{2n}}\right) \quad \text{as} \quad n \to \infty.$$

Then we have (2.17) for any t and u with  $|u| < \delta$ , because

$$E\{|R_n|^2\} = \frac{1}{[(n-1)!]^2} \int_{-\infty}^{\infty} \left| \int_{0}^{u} (e^{i\xi\lambda} - 1)(u-\xi)^{n-1} d\xi \right|^2 \lambda^{2n} dF(\lambda)$$

$$\leq \frac{4}{[(n-1)!]^2} \int_{-\infty}^{\infty} \left\{ \int_{0}^{|u|} (|u|-\xi)^{n-1} d\xi \right\}^2 \lambda^{2n} dF(\lambda)$$

$$= \frac{4|u|^{2n}}{[n!]^2} \int_{-\infty}^{\infty} \lambda^{2n} dF(\lambda)$$

$$\leq \frac{4\delta^{2n}}{[n!]^2} \int_{-\infty}^{\infty} \lambda^{2n} dF(\lambda)$$

$$= o(1) \quad \text{as} \quad n \to \infty \quad \text{for} \quad |u| < \delta.$$

EXAMPLE 3. Let  $\{t_n\}$  be a sequence of real numbers with  $t_n \neq t_0$   $(n=1, 2, \cdots)$  and  $t_n \rightarrow t$  as  $n \rightarrow \infty$ . Under the assumption (2.19), the random variable  $\mathcal{E}(t)$  for any fixed t is determined with probability 1 by the sequence  $\{\mathcal{E}(t_n); n=1, 2, \cdots\}$  of the random variables.

EXAMPLE 4. Under the assumption (2.16) for any fixed positive integer n, we have with probability 1 that

(2.21) 
$$\lim_{h\to 0} \frac{1}{h^n} \sum_{k=0}^n \binom{n}{k} (-1)^{n-k} \mathcal{E}(t+kh) = \mathcal{E}^{(n)}(t).$$

Since, by Theorem 1,

(2.22)  $E\{|R_k|^2\}=o(h^{2n})$  as  $h\to 0$ ,

where

(2.23) 
$$\mathscr{E}(t+kh) = \sum_{\nu=0}^{n} \frac{\mathscr{E}^{(\nu)}(t)}{\nu !} (kh)^{\nu} + R_{k},$$

we have

(2.24) 
$$\sum_{k=0}^{n} \binom{n}{k} (-1)^{n-k} \mathcal{E}(t+kh)$$
$$= \sum_{\nu=0}^{n} \left[ \sum_{k=0}^{n} \binom{n}{k} (-1)^{n-k} k^{\nu} \right] \frac{\mathcal{E}^{(\nu)}(t)}{\nu !} h^{\nu} + \sum_{k=0}^{n} \binom{n}{k} (-1)^{n-k} R_{k}$$
$$= \mathcal{E}^{(n)}(t) h^{n} + \sum_{k=0}^{n} \binom{n}{k} (-1)^{n-k} R_{k}$$

and so

$$E\left\{\left|\frac{1}{h^n}\sum_{k=0}^n\binom{n}{k}(-1)^{n-k}\mathcal{E}(t+kh)-\mathcal{E}^{(n)}(t)\right|^2\right\}$$

(2.25)

$$\leq \left\{\frac{1}{|\boldsymbol{h}|^n} \sum_{k=0}^n \binom{n}{k} \sqrt{\{|\boldsymbol{R}_k|^2\}}\right\}^2 = o(1) \quad \text{as} \quad \boldsymbol{h} \to 0,$$

which gives (2.21).

## HIROHISA HATORI

REMARK 2. The result of Example 4 has been proved by Kawata [2] in a different method. From Remark 1, we have (2.21) by assuming the existence and the continuity of  $\mathcal{E}^{(n)}(t)$  at the point t without the stationarity of  $\mathcal{E}(t)$ .

3. In the preceeding section, we have defined the integral

$$\int_{A}^{B} X(t) \, dK(t).$$

If it holds that

$$\lim_{\substack{A \to -\infty \\ B \to \infty}} \int_{A}^{B} X(t) \ dK(t) = I,$$

then I is denoted as

$$\int_{-\infty}^{\infty} X(t) \ dK(t).$$

LEMMA 3. If X(t) is continuous in  $(-\infty, \infty)$  and there exists a non-negative valued function g(t) such that

(3.1) 
$$\sqrt{E\{|X(t)|^2\}} \leq g(t)$$
 for all t

and

(3.2) 
$$\int_{-\infty}^{\infty} g(t) |dK(t)| < \infty,$$

then there exists

$$\int_{-\infty}^{\infty} X(t) \ dK(t).$$

Proof. Since

(3.3) 
$$E\left\{\left\|\int_{A}^{A'} X(t) \, dK(t)\right\|^{2}\right\} \leq \left\{\int_{A}^{A'} g(t) \, dK(t) \, \|\right\}^{2},$$

we have by (3.2) that

(3.4) 
$$\lim_{A,A'\to\infty} \int_{A}^{A'} X(t) \, dK(t) = 0$$

and

(3.5) 
$$\lim_{B,B'\to\infty}\int_{B}^{B'}X(t)\,dK(t)=0,$$

which ensure the existence of the integral

$$\int_{-\infty}^{\infty} X(t) \, dK(t).$$

THEOREM 3. Let H(u) be a numerical valued function of bounded variation in  $(-\infty, \infty)$  and we assume that

(3.6) 
$$\int_{-\infty}^{\infty} |u|^{r+\alpha} |dH(u)| < \infty,$$

where r is a non-negative integer and  $\alpha$  is a constant with  $0 \leq \alpha < 1$ . If  $X(t), -\infty < t$  $<\infty$ , is a continuous stochastic process such that

(3.7) 
$$\sqrt{E\{|X(s)|^2\}} \leq C(1+|s|^{r+\alpha})$$

for all s and some constant C and

(3.8) 
$$E\left\{\left|X(t+u)-\sum_{k=0}^{r}\frac{X^{(k)}(t)}{k!}u^{k}\right|^{2}\right\}=o(|u|^{2r+2\alpha}) \quad as \quad u\to 0,$$

then it holds that

(3

$$=o\left(\frac{1}{n^{2r+2\alpha}}\right)$$
 as  $n\to\infty$ .

Proof. (3.6) and (3.7) ensure with Lemma 3 the existence of the integral

$$\int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) dH(s).$$

Put

$$I = \int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) dH(s) - \sum_{k=0}^{r} \frac{(-1)^{k} X^{(k)}(t)}{k! n^{k}} \int_{-\infty}^{\infty} s^{k} dH(s)$$
$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) - \sum_{k=0}^{r} \frac{X^{(k)}(t)}{k! n^{k}} \left(-\frac{s}{n}\right)^{k} dH(s).$$

(3.10)

$$= \int_{-\infty}^{\infty} \left[ X\left(t - \frac{s}{n}\right) - \sum_{k=0}^{\prime} \frac{X^{(k)}(t)}{k!} \left(-\frac{s}{n}\right)^{k} \right] dH(s).$$

Then we have that

$$n^{r+\alpha}\sqrt{E\{|I|^2\}} \leq n^{r+\alpha} \int_{-\infty}^{\infty} \sqrt{E\left\{\left|X\left(t-\frac{s}{n}\right)-\sum_{k=0}^{r} \frac{X^{(k)}(t)}{k!}\left(-\frac{s}{n}\right)^k\right|^2\right\}} |dH(s)|$$
(3.11)

$$= \int_{-\infty}^{\infty} \psi\left(-\frac{s}{n}\right) \cdot |s|^{r+\alpha} |dH(s)|,$$

where

$$\psi(u) = |u|^{-r-\alpha} \sqrt{E\left\{ \left| X(t+u) - \sum_{k=0}^{r} \frac{X^{(k)}(t)}{k!} u^{k} \right|^{2} \right\}}$$

for  $u \neq 0$  and  $\psi(0) = 0$ .

Since, by (3.8),

 $(3.12) \qquad \qquad \phi(u) = o(1) \qquad \text{as} \quad u \to 0,$ 

we can choose a positive number  $\delta$  such that

 $(3.13) \qquad \qquad \psi(u) \leq 1 \qquad \qquad \text{for} \quad |u| < \delta.$ 

On the other hand, we have

$$\begin{split} \psi(u) &\leq |u|^{-r-\alpha} \left[ \sqrt{E\{|X(t+u)|^2\}} + \sum_{k=0}^r \frac{\sqrt{E\{|X^{(k)}(t)|^2\}}}{k!} |u|^k \right] \\ &\leq |u|^{-r-\alpha} \left[ C(1+|t+u|^{r+\alpha}) + \sum_{k=0}^r \frac{\sqrt{E\{|X^{(k)}(t)|^2\}}}{k! \ \delta^{r+\alpha-k}} |u|^{r+\alpha} \right] \\ &\leq C \left[ \delta^{-r-\alpha} + \left(1 + \frac{|t|}{\delta}\right)^{r+\alpha} \right] + \sum_{k=0}^r \frac{\sqrt{E\{|X^{(k)}(t)|^2\}}}{k! \ \delta^{r+\alpha-k}} \\ &\equiv K \qquad \text{for} \quad |u| \geq \delta, \end{split}$$

which implies with (3.13) that

(3.15) 
$$\psi\left(-\frac{s}{n}\right) < K+1 \qquad \text{for all } s,$$

Next, (3.12) shows that

(3.16) 
$$\lim_{n \to \infty} \psi\left(-\frac{s}{n}\right) = 0 \qquad \text{for any fixed } s.$$

Since K is a constant independent of s and n, it is obtained by (3.6), (3.15), (3.16) and Lebesgue's theorem that

(3.17) 
$$\lim_{n \to \infty} \int_{-\infty}^{\infty} \psi\left(-\frac{s}{n}\right) |s|^{r+\alpha} |dH(s)| = 0$$

and so

(3.18) 
$$\lim_{n \to \infty} n^{r+\alpha} \sqrt{E\{|I|^2\}} = 0,$$

which gives (3.9).

From Theorem 2 and Theorem 3, we get the following

THEOREM 4. Let f(s) and  $\phi(s)$  be continuous in  $(-\infty, \infty)$ , H(s) be of bounded variation and  $\mathcal{E}(s), -\infty < s < \infty$ , be continuous (weakly) stationary stochastic process whose spectral function is  $F(\lambda)$ . If, in addition to the assumptions (2.2), (2.9) (2.10) and (3.6), we assume that

(3.19) 
$$|f(s)| \leq C(1+|s|^{r+\alpha})$$

and

(3.20) 
$$|\phi(s)| \leq C(1+|s|^{r+\alpha})$$

136

(3.14)

for all s and some positive constant C, then we have

(3.21)  
$$E\left\{\left|\int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) dH(s) - \sum_{k=0}^{r} \frac{(-1)^{k} X^{(k)}(t)}{k! n^{k}} \int_{-\infty}^{\infty} s^{k} dH(s)\right|^{2}\right\}$$
$$= o\left(\frac{1}{n^{2r+2\alpha}}\right) \qquad as \qquad n \to \infty,$$

where  $X(s) = f(s) + \phi(s) \mathcal{E}(s), -\infty < s < \infty$ .

REMARK 3. The continuity of f(s) and  $\phi(s)$  is used to ensure the existence of the integral

$$\int_{-\infty}^{\infty} X\left(t - \frac{s}{n}\right) dH(s)$$

only.

REMARK 4. Let  $\mathcal{E}_{\nu}(t)$  ( $\nu=1, 2, \dots, N$ ) be stationary processes satisfying the conditions similar to the one on  $\mathcal{E}(t)$  in Theorem 4 and  $\phi_{\nu}(t)$  ( $\nu=1, 2, \dots, N$ ) be numerical valued functions satisfying the conditions similar to the one on  $\phi(t)$  in Theorem 4. Then we have (3.21) for

$$X(t) = f(t) + \sum_{\nu=1}^{N} \phi_{\nu}(t) \mathcal{E}_{\nu}(t), \quad -\infty < t < \infty.$$

The author expresses his sincerest thanks to Prof. I. Amemiya who has given valuable advices.

## References

- HATORI, H., Some expansion theorems for stochastic processes, I. Kōdai Math. Sem. Rep. 15 (1963), 111-120.
- [2] KAWATA, T., Remarks on prediction problem in the theory of stationary stochastic processes. Tôhoku Math. Journ. (2) 6 (1954), 13-20.
- [3] KAWATA, T., Some convergence theorems for stationary stochastic processes. Ann. of Math. Stat. 30 (1959), 1192-1214.

TOKYO COLLEGE OF SCIENCE.