

# Renewal theory-A Markov process approach

By

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## § 1. Introduction.

Renewal Theory is concerned with, among other things, the point process generated by a random walk and the behaviour of solutions of the so-called renewal equation. To be precise, let  $\{X_i\}_{i=1}^{\infty}$  be i. i. d. nonnegative random variables with a common distribution  $F(\cdot)$ . Let  $X_0$  be a nonnegative random variable independent of  $\{X_i\}_1^{\infty}$ . Set  $S_n = \sum_{i=0}^n X_i$  for  $n \geq 0$ . Let  $\xi(\cdot)$  be a measurable function from  $R^+$  to  $R^+$  and be bounded a. e. on finite intervals. The equation

$$(1) \quad m(t) = \xi(t) + \int_{(0,t]} m(t-u) dF(u) \quad \text{for } t \geq 0$$

is called the *renewal equation*.

The objects of interest are: a) the asymptotic behaviour of the point process  $\{S_n\}_0^{\infty}$  and b) the asymptotic behaviour of the solution  $m(\cdot)$  to (1). The following results are well-known. Let

$$(2) \quad \begin{aligned} U(t) &= E \{ \# n : S_n \leq t \} \\ &= \sum_0^{\infty} P(S_n \leq t) \end{aligned}$$

be the so-called *renewal function*. Assume from now on that  $F(\cdot)$  is non-lattice,  $0 < \lambda^{-1} = \int_0^{\infty} u dF(u) < \infty$ .

**Theorem 1.** (Blackwell) For all  $0 < h < \infty$

$$(3) \quad U(t+h) - U(t) \longrightarrow \lambda h \quad \text{as } t \longrightarrow \infty$$

**Theorem 2.** (Feller) If  $\xi(\cdot)$  is directly Riemann integrable then the solution  $m(\cdot)$  of (1) satisfies

$$(4) \quad m(t) \longrightarrow \lambda \int_0^{\infty} \xi(u) du.$$

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Feller [4] has shown that these two theorems are equivalent.

The object of this note is to present a Markov process formulation that is equivalent to the above. This is done in terms of the Markov Process of the so-called backward *recurrence time* or what we shall call as the *age process*. We shall show that Blackwell's and Feller's theorems are equivalent to the weak convergence of this age process. If one assumes a mild smoothness on  $F(\cdot)$ , then one can strengthen this weak convergence to that in variation norm. This in turn leads to the following strengthening of (3) and (4):

$$(3)' \quad U(t+B) \longrightarrow \lambda m(B)$$

for all bounded Borel sets, where we mean, by abuse of notation,  $U(B) = \sum_0^\infty P(S_n \in B)$ ,  $m(\cdot)$  is Lebesgue measure and

$$(4)' \quad m(t) \longrightarrow \lambda \int_0^\infty \xi(u) du$$

for all bounded measurable  $\xi(\cdot)$  that are dominated by a multiple of the tail of  $F(\cdot)$ . It turns out that the smoothness of  $F(\cdot)$  is necessary as well for these stronger conclusions. Thus, in the renewal equation (1), to get the convergence of  $m(t)$  as  $t \rightarrow \infty$ , one needs either a smoothness condition on  $\xi(\cdot)$  like d.r.i. or on  $F(\cdot)$  like non-strongly singularity. The Markov process approach besides bringing out this balance between  $\xi(\cdot)$  and  $F(\cdot)$  into sharper focus, also suggests that when studying the limit behaviour of Markov Chains on general state spaces that do have a topological structure on them it is perhaps worthwhile to prove the weak convergence first as this may hold under fairly mild recurrence conditions rather than try to use the Doeblin-Harris theory as this needs stronger recurrence conditions (although, these yield stronger convergence, such as in variation norm).

The Markov Process approach has been mentioned in Doob [3]. We have learnt after this work was completed that Arjas et al [1] have also obtained results similar to ours.

## § 2. Statement of results.

Let  $F(\cdot)$  be a probability distribution on  $(0, \infty)$ . Assume throughout that  $F(\cdot)$  is non-lattice and  $0 < \lambda^{-1} = \int_0^\infty u dF(u) < \infty$ . For each  $x \in [0, T)$ , let  $X_0^{(x)}$  be a random variable with

$$(5) \quad P(X_0^{(x)} \leq t) = \frac{F(x+t) - F(x)}{1 - F(x)} \quad \text{for } t \geq 0, \text{ and } x < T$$

where  $T = \sup\{x : F(x) < 1\}$ .

Let  $\{X_i\}_1^\infty$  be i.i.d. r.v. with distribution  $F(\cdot)$  and independent of  $X_0^{(x)}$ . Set

$$\begin{aligned}
 (6) \quad A(t) &= x + t \quad \text{for } 0 \leq t \leq X_0^{(x)} \\
 t - X_0^{(x)} & \quad X_0^{(x)} \leq t \leq X_0^{(x)} + X_1 \\
 t - S_n^{(x)} & \quad S_n^{(x)} \leq t < S_{n+1}^{(x)}, \\
 n &= 0, 1, 2, \dots
 \end{aligned}$$

where  $S_n^{(x)} = X_0^{(x)} + X_1 + \dots + X_n$  for  $n = 0, 1, 2, \dots$ . From the very definition, we get the following:

**Proposition 1.** *The stochastic process  $\{A(t); t \geq 0\}$  is a Markov Process on  $[0, \infty)$  with stationary transition probabilities.*

The transition function  $P(x, t, E) \equiv P(A(t) \in E | A(0) = x)$  satisfies the equation

$$\begin{aligned}
 (7) \quad P(x, t, E) &= \chi_E(x+t) \left( \frac{1-F(x+t)}{1-F(x)} \right) \\
 &+ \int_{(0, t_1)} P(0, t-u, E) \frac{dF(x+u)}{(1-F(x))}.
 \end{aligned}$$

One solves (7) for  $x=0$  and uses (7) to obtain it for all  $x$ . Clearly,

$$\begin{aligned}
 (8) \quad P(0, t, E) &= P \left\{ \bigcup_{n=0}^{\infty} (S_n^{(0)} \leq t < S_{n+1}^{(0)}, t - S_n^{(0)} \in E) \right\} \\
 &= \int_{(0, t_1)} \chi_E(t-u) (1-F(t-u)) U(du)
 \end{aligned}$$

where  $U(t) = \sum_0^\infty P(S_n^{(0)} \leq t)$  is the renewal function. From now on we assume  $F(\cdot)$  is non lattice.  $F(0)=0$  and  $0 < \lambda^{-1} = \int_0^\infty t dF(t) < \infty$ . Let  $\pi(E) = \lambda \int_E (1-F(u)) du$  for all Borel sets  $E$  in  $R^+$ . We now state an equivalent form of Theorems 1 and 2.

**Theorem 3.** *For all initial condition  $x$ , the Markov Process  $A(t)$  converges weakly to  $\pi(\cdot)$ .*

The proof of the equivalence is done by showing Theorem 2  $\Rightarrow$  Theorem 3  $\Rightarrow$  Theorem 1. In fact, let  $f(\cdot)$  be a bounded continuous function on  $[0, \infty)$ . Then  $a_0(t) \equiv E_0(f(A(t)))$  satisfies

$$a_0(t) = f(t)(1-F(t)) + \int_{(0, t_1)} a_0(t-u) dF(u).$$

The function  $\xi(t) \equiv f(t)(1-F(t))$  is directly Riemann integrable and so by Theorem 2  $a_0(t) \rightarrow \int_0^\infty f(t)(1-F(t)) dt$ . Since  $a_x(t) \equiv E_x f(A(t))$  satisfies

$$a_x(t) = f(x+t) \left( \frac{1-F(x+t)}{1-F(x)} \right) + \int_0^t a_0(t-u) \left( \frac{dF(x+u)}{1-F(x)} \right)$$

We get by bounded convergence theorem  $\lim_t a_x(t) = \lim_t a_0(t)$ . Thus, Theorem 2  $\Rightarrow$  Theorem 3.

For  $h$  small and positive, the function  $f(t) \equiv (1 - F(t))^{-1} \chi_{[0, h)}(t)$  is bounded and continuous a.e. So by Theorem 3,  $E_0 f(A(t)) \rightarrow \lambda h$ . But  $E_0 f(A(t)) \equiv U(t) - U(t-h)$ . Thus, Theorem 3  $\Rightarrow$  Theorem 1. We shall give in §3 an independent proof of Theorem 3, using a coupling argument.

To strengthen the above convergence of  $A(t)$  to  $\pi$  from weak to variation norm, we need a mild smoothness condition on  $F(\cdot)$ . With this in mind we introduce a

**Definition.** A distribution function  $F(\cdot)$  on  $R$  is *strongly singular* if for all  $n$ , the  $n$  fold convolution  $F^{(n)}(\cdot)$  is singular with respect to Lebesgue measure.

Clearly,  $F(\cdot)$  is *strongly singular* iff  $U(\cdot)$  is *singular* with respect to Lebesgue measure.

**Theorem 4.** Let  $F(\cdot)$  be not strongly singular. Then, for all  $x$ ,

$$\lim_t \|P_x(A(t) \in \cdot) - \pi(\cdot)\| = 0$$

where  $\|\cdot\|$  is variation norm.

It turns out that the converse is true as well.

**Theorem 5.** Let  $F(\cdot)$  be strongly singular. Then

$$\lim \|P_x(A(n) \in \cdot) - \pi(\cdot)\| = 2.$$

The proofs of these two theorems are in §4.

### §3. A coupling proof of the weak convergence of $A(t)$ to $\pi(\cdot)$

Here is the plan of the proof. First we show that  $\pi(\cdot)$  is stationary for  $A(\cdot)$ . Next, we construct two processes  $A_1(\cdot)$ ,  $A_2(\cdot)$  such that both are age processes with  $A_1(0)$  distributed according to  $\pi(\cdot)$  and  $A_2(0) = 0$  w.p.1. This construction is done in such a way that for every  $\varepsilon > 0$ , there exists a non-anticipating random time  $T$  such that  $0 \leq (A_1(T) - A_2(T)) < \varepsilon$ . This forces the limit behaviour of the distribution of  $A_1(\cdot)$  and  $A_2(\cdot)$  to be the same. But since  $A_1(\cdot)$  is stationary with distribution  $\pi(\cdot)$ , it follows that  $A_2(\cdot)$  converges weakly to  $\pi(\cdot)$ . Now the details. We begin by establishing the stationarity of  $\pi(\cdot)$ .

**Theorem 6.** The measure  $\pi(\cdot)$  is stationary for  $A(\cdot)$ .

*Proof.* Let  $f(\cdot)$  be bounded measurable on  $[0, \infty)$ . Let

$$\begin{aligned} m(x, t) &\equiv E_x(f(A(t))) \\ (9) \quad m(t) &= \int_0^\infty m(x, t) \pi(dx). \end{aligned}$$

We need to show that  $m(t) \equiv m(0)$ . Now  $m(x, t)$  satisfies

$$(10) \quad m(x, t) = f(x+t) \left( \frac{1-F(x+t)}{1-F(x)} \right) + \int_0^t m(0, t-u) \left( \frac{dF(x+u)}{1-F(x)} \right)$$

and we have

$$(11) \quad \lambda^{-1}m(t) = \int_0^\infty f(x+t)(1-F(x+t))dx + \int_0^\infty \left( \int_0^t m(0, t-u) dF(x+u) \right) dx$$

But, since  $m(0, t)$  satisfies the renewal equation

$$m(0, t) = f(t)(1-F(t)) + \int_0^t m(0, t-u) dF(u),$$

$$m(0, t) = \int_0^t \xi(t-u) U(du) \quad \text{where} \quad \xi(t) = f(t)(1-F(t)).$$

Thus

$$\begin{aligned} \int_0^\infty \left( \int_0^t m(0, t-u) dF(x+u) \right) dx &= \int_0^\infty m(0, t-u)(1-F(u)) du \\ &= \int_0^t \xi(u) du, \end{aligned}$$

since  $\int_0^t (1-F(t-u)) U(du) \equiv 1$ .

Now (11) shows that  $\lambda^{-1}m(t) \equiv \int_0^\infty f(u)(1-F(u)) du$ . q. e. d.

Now we construct the two processes  $A_1(\cdot)$ ,  $A_2(\cdot)$ . Recall the set up in the beginning of the section 2. Since  $A_1(0)$  is distributed according to  $\pi(\cdot)$  so is  $X_0$ . Let  $\{X_i^1\}_1^\infty$  and  $\{X_i^2\}_0^\infty$  be two independent sequences of i.i.d.r.v. with distribution  $F(\cdot)$ . The sequence  $Y_i = X_i^1 - X_i^2$  for  $i=1, 2, \dots$ , being i.i.d. mean zero non-lattice random variables, given any  $\varepsilon > 0$ , there exists a random variable  $N(\varepsilon)$  such that

$$0 < \sum_{i=0}^{N(\varepsilon)} X_i^1 - \sum_{i=0}^{N(\varepsilon)} X_i^2 \equiv \Delta < \varepsilon.$$

Let  $\{X_i^3\}_1^\infty$  be a sequence of random variables defined by

$$X_i^3 = \begin{cases} X_i^1 & 0 \leq i \leq N(\varepsilon) \\ X_i^2 & i > N(\varepsilon) \end{cases}$$

Let  $\{A_3(t); t \geq 0\}$  be the age process associated with  $\{X_i^3\}_0^\infty$ . As stochastic processes  $A_3(\cdot)$  and  $A_1(\cdot)$  are clearly equivalent. Also the processes  $\{A_2(\cdot)\}$  and  $\{A_3(\cdot)\}$  are coupled in the sense that

$$A_2(t) = A_3(t - \Delta) \quad \text{for} \quad t \geq T \equiv \sum_0^{N(\varepsilon)} X_i^2 \quad \text{and if}$$

$$A^3(t) > \Delta, \quad \text{then} \quad A_2(t) = A_3(t) - \Delta.$$

*Proof of Theorem 3.* Now let  $f(\cdot)$  be bounded and uniformly continuous on  $(0, \infty)$ . Then,

$$\begin{aligned}
Ef(A_2(t)) &= E\{f(A_2(t)); T \leq t\} \\
&\quad + E\{f(A_2(t)); T > t, A_3(t) < \varepsilon\} \\
&\quad + E\{f(A_2(t)); T > t, A_3(t) \geq \varepsilon\}
\end{aligned}$$

and doing a similar decomposition of  $Ef(A_3(t))$ , we see that

$$|Ef(A_2(t)) - Ef(A_3(t))| \leq 2\|f\|(P(T > t) + P(A_3(t) < \varepsilon)) + \eta(\varepsilon)$$

where

$$\eta(\varepsilon) \equiv \sup_{|x-y| \leq \varepsilon} |f(x) - f(y)|$$

Since  $A_3(t)$  has distribution  $\pi(\cdot)$  and since  $f(\cdot)$  is uniformly continuous, given a  $\delta > 0$ , we can choose  $\varepsilon > 0$  such that

$$P(A_3(t) < \varepsilon) < \delta \quad \text{for all } t$$

and

$$\eta(\varepsilon) < \delta.$$

Now  $P(T > t) \rightarrow 0$  as  $t \uparrow \infty$ . Thus,

$$\overline{\lim}_t |Ef(A_2(t)) - Ef(A_3(t))| \leq (2\|f\| + 1)\delta.$$

Since  $\delta$  is arbitrary we are done.

#### §4. Proofs of Theorems 4 and 5.

Our proofs are based on the following ergodic theorem for Markov chains on general state spaces formulated and proved in a manner suitable to us by Athreya and Ney [2].

**Theorem 7.** Let  $\{X_n\}_0^\infty$  be a Markov Chain on a measurable space  $(S, \mathcal{S})$  with transition function  $P(x, E)$ . Assume

(1)  $\exists A \in \mathcal{S}$  such that

$$P(X_n \in A \text{ for some } n \geq 1 | x_0 = x) \equiv 1$$

(2)  $\exists n_0, \lambda$  and a probability measure  $\varphi(\cdot)$  on  $A$

$$P(X_{n_0} \in E | X_0 = x) \geq \lambda \varphi(E) \quad \text{for all } x \text{ in } A, E \subset A$$

(3) g.c.d. of such  $n_0$  as  $A$  varies is one.

Then, there exists a  $\sigma$  finite measure  $\pi(\cdot)$  that is invariant for  $P(\cdot, \cdot)$  and is unique upto a multiplicative constant. Further, when that  $\pi(\cdot)$  is finite and normalized to be a probability measure,  $\overline{\lim}_n \|P_x(X_n \in \cdot) - \pi(\cdot)\| = 0$  for all  $x$ .

Their proof is based on a simple regeneration lemma and the discrete renewal equation. We refer to [2] for details.

The following result is a key step in the proof of Theorem 4.

**Theorem 8.** Let  $F(\cdot)$  be not strongly singular. Then, for all sufficiently small and positive  $\delta$  the sequence  $\{X_n \equiv A(n\delta)\}_0^\infty$  satisfies the hypotheses of Theorem 7.

Assuming the validity of Theorem 7 and 8, we now present the proof of Theorem 4.

*Proof of Theorem 4.* Since  $\pi(\cdot)$  is invariant for the process  $A(\cdot)$ , we have for  $n\delta \leq t < (n+1)\delta$ ,

$$P_x(A(t) \in E) - \pi(E) = \int_0^\infty P(y, t - n\delta, E) P_x(A(n\delta) \in dy) \\ - \int_0^\infty P(y, t - n\delta, E) \pi(dy),$$

and hence,

$$\|P_x(A(t) \in \cdot) - \pi(\cdot)\| \leq \|P_x(A(n\delta) \in \cdot) - \pi(\cdot)\|$$

which goes to zero as  $n \rightarrow \infty$  by Theorem 6 and 7.

*Proof of Theorem 5.* If  $F(\cdot)$  is strongly singular then the renewal measure  $U(\cdot)$  is singular and hence there exists a set  $B$  such that  $U(B) = 0$  and  $m(B^c) = 0$ . Let  $B_0 = \bigcap_{n=1}^\infty (n - B)$ . Then,  $m(B_0^c) = 0$  and  $U(n - B_0) \leq U(n - (n - B)) = 0$  for each  $n$ . Now, since

$$P_0(A(t) \in E) = \int \chi_E(u)(1 - F(u))U(t - du),$$

and since  $U(n - B_0) = 0$ ,

$$P_0(A(n) \in B_0) = 0.$$

On the other hand,  $\pi(\cdot)$  being absolutely continuous with respect to  $m(\cdot)$ ,  $\pi(B_0^c) = 0$ . Hence,  $\|P(A(n) \in \cdot) - \pi(\cdot)\| = 2$ , for all  $n$  thus proving Theorem 4.

All that remains now is to prove Theorem 8. We need the following lemma and its corollaries in the proof of Theorem 8.

**Lemma.** Let  $f(\cdot)$  be a non-negative measurable function on  $R$  that is not zero a. e. Then  $\exists a < b, \delta > 0$  such that

$$h(x) \equiv \int f(x - y)f(y)dy > \delta \quad \text{for all } a < x < b.$$

*Proof.* Since  $h(x) \geq h_k(x) \equiv \int_{|y| \leq k} f_k(x - y)f_k(y)dy$  for all  $k > 0$ , where  $f_k(u) = \min(f(u), k)$ , it is enough to prove the assertion for  $h_k(\cdot)$  for some  $k$ . It is easily verified that the convolution  $(f_1 * f_2)(\cdot)$  is continuous if  $f_1 \in L_1$  and  $f_2 \in L_\infty$ . Thus,  $h_k(\cdot)$  is continuous. Finally, since  $f \neq 0$  a. e.  $\exists x_0$  and  $k$  such that  $h_k(x_0) > 0$ .

**Corollary 1.** Let  $F(\cdot)$  be a probability distribution on  $R$  that is not singular with respect to Lebesgue measure. Then  $(F * F)(\cdot)$  has an absolutely continuous component whose density is bounded away from zero in a non-degenerate interval.

**Corollary 2.** Since  $F(\cdot)$  is not strongly singular there exists  $n_0$  such that  $F^{(n_0)}(\cdot)$  is not singular with respect to Lebesgue measure. By corollary 1,  $F^{(2n_0)}(\cdot)$  has a density that is bounded below in a non-degenerate interval. But for any  $a < b$ ,  $U(b) - U(a) \geq F^{(2n_0)}(b) - F^{(2n_0)}(a) \geq \text{const.}(b-a)$  if  $a, b$  lie in such an interval.

*Proof of Theorem 8.* Let  $I_n = [(n-1)h, nh]$ ,  $n=1, 2, \dots$ , where  $h > 0$  will be chosen later. Let  $E$  be a Borel set in  $I_1 = [0, h]$  and  $n_0$  a fixed integer. Then

$$P(x_1, n_0h, E) \equiv P(A(n_0h) \in E | A(0) = x)$$

$$= \text{Prob} \bigcap_{n=0}^{\infty} \{(n_0-1)h \leq S_n < n_0h, X_{n+1} > n_0h - S_n$$

$$\text{and } n_0h - S_n \in E\}$$

where  $S_n = X_0 + X_1 + \dots + X_n$ , for  $n=0, 1, 2, \dots$ , with  $X_i$  being independent and

$$P(X_0 \leq t) \equiv \frac{F(x+t) - F(x)}{1 - F(x)}$$

$$P(X_i \leq t) = F(t) \quad \text{for } i \geq 1.$$

$$\begin{aligned} P(A(n_0h) \in E | A(0) = x) &= \sum_{n=0}^{\infty} P(S_n \in I_{n_0}, X_{n+1} > n_0h - S_n, n_0h - S_n \in E) \\ &= \sum_{n=0}^{\infty} \int_{I_{n_0}} (1 - F(n_0h - y)) \chi_E(n_0h - y) P(S_n \in dy) \\ &= \int_{E - \{0\}} (1 - F(u)) U_{n_0}^x(du) \end{aligned}$$

where  $U_{n_0}^x(B) = \sum_0^{\infty} P(S_n \in n_0h - B)$  for any Borel set  $B$  in  $I_1 = [0, h]$ . Now,

$$U_{n_0}^x(B) = \int_0^{\infty} U(n_0h - z - B) F_x(dz)$$

where  $U(\cdot)$  is the usual renewal measure (i.e., with  $x=0$ ). Since

$$F_x(z) = \frac{F(x+z) - F(x)}{1 - F(x)},$$

$$\begin{aligned} U_{n_0}^x(B) &\geq \int_x^{\infty} U(n_0h + x - v - B) F(dv) \\ &\geq \int_h^{\infty} \left( \int_{n_0h + x - v - B} k(w) dw \right) F(dv) \end{aligned}$$

where  $k(\cdot)$  is the density of the absolutely continuous component of  $U(\cdot)$  with respect to Lebesgue measure. Making a change of variables yields

$$U_{n_0}^x(B) \geq \int_h^{\infty} \left( \int_B k(n_0h + x - v - s) ds \right) F(dv).$$

By corollary 2  $\exists, 0 < a < b < \infty$  and  $\delta > 0$ , such that  $k(x) > \delta$  for  $a < x < b$ . Now choose  $h$  small so that  $(b-a) > 10h$ . Since  $F(0) = 0$ ,  $\exists, I_r \equiv [(r-1)h, rh]$  such that  $F(I_r) > 0$ . As  $x$  and  $s$  vary in  $[0, h]$ , and  $v$  varies in  $[rh, (r+1)h]$ ,  $n_0h + x - v - s$



has to lie between  $(n_0-r-2)h$  and  $(n_0+1-r)h$ . Since  $(b-a)>10h$  we can find integers  $n_1$  and  $n_2$  such that  $(n_2-n_1)>10$  and  $a<n_1h<n_2h<b$ . Thus for  $n_1+r+2\leq n_0\leq n_2+r-1$ :

$$U_{n_0}^x(B)\geq\delta F(I_r)m(B)$$

where  $m(\cdot)$  is Lebesgue measure. Hence, there is an  $\alpha>0$  such that for all  $0\leq x<h$ , Borel sets  $E$  in  $[0, h)$ :

$$P\{A(n_0h)\in E\mid A(0)=x\}\geq\alpha\int_E(1-F(u))du$$

for all  $n_1+r+2\leq n_0\leq n_2+r-1$ . Also by the definition of the age process we do get the recurrence condition

$$P(A(nh)\in[0, h)\quad\text{for some } n\geq 1\mid A(0)=x)\equiv 1$$

for all  $x$  in  $R^+$  and all  $h>0$ . Thus, for the Markov Chain  $\{Z_n\equiv A(nh)\}$  we have produced a set  $A\equiv[0, h)$ , an integer  $N$  such that

$$P(Z_n\in E\mid Z_0=x)\geq\alpha\int_E(1-F(u))du$$

for all  $x$  in  $A$ ,  $E\subset A$ ,  $n=N, N+1, N+2$ . This is precisely what makes  $\{Z_n\}_0^\infty$  satisfy the hypothesis of Theorem 7.

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