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How big are the l^{∞} -valued random fields?*

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Abstract

In this paper we establish path properties and a generalized uniform law of the iterated logarithm (LIL) for strictly stationary and linearly positive quadrant dependent (LPQD) or linearly negative quadrant dependent (LNQD) random fields taking values in l^{∞} -space.

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1 Introduction and Results

In the last years there has been growing interest in concepts of positive/negative dependence for families of random variables. Such concepts are of considerable use in deriving inequalities in probability and statistics. Recently, Csörgő et al. [1] and Choi and Csörgő [2] studied path properties and asymptotic properties for l^{∞} -valued Gaussian random fields, respectively. In this paper we are interested in path properties for *any* positive or negative dependent random fields with multidimensional indices taking values in l^{∞} -space.

Newman [3] introduced and discussed the following concepts of positive or negative dependence. The random field $\{X_i(\mathbf{t}); \mathbf{t} := (t_1, \dots, t_N) \in [0, \infty)^N\}_{i=1}^\infty$ is said to be *linearly positive quadrant dependent* (LPQD) if, for any positive numbers λ_i and any disjoint finite subsets A, B of \mathbb{Z}_+ (set of positive integers), the inequality

$$P\left\{\sum_{i\in A}\lambda_i X_i(\mathbf{t}_i) \ge x, \sum_{j\in B}\lambda_j X_j(\mathbf{t}_j) \ge y\right\} \ge P\left\{\sum_{i\in A}\lambda_i X_i(\mathbf{t}_i) \ge x\right\} P\left\{\sum_{j\in B}\lambda_j X_j(\mathbf{t}_j) \ge y\right\}$$
(1.1)

holds for all $x, y \in \mathbb{R}$ (set of real numbers), where $\{\mathbf{t}_j\}_{j=1}^{\infty} \subset \{\mathbf{t}\}$, which is equivalent to the inequality (Lehmann [6], pp. 1137-1138)

$$P\left\{\sum_{i\in A}\lambda_i X_i(\mathbf{t}_i) \le x, \sum_{j\in B}\lambda_j X_j(\mathbf{t}_j) \le y\right\} \ge P\left\{\sum_{i\in A}\lambda_i X_i(\mathbf{t}_i) \le x\right\} P\left\{\sum_{j\in B}\lambda_j X_j(\mathbf{t}_j) \le y\right\},$$
(1.2)

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while the random field $\{X_i(t)\}_{i=1}^{\infty}$ is said to be linearly negative quadrant dependent (LNQD) if the inequalities in (1.1) and (1.2) are reversed. For the results related to such dependence, one can refer to [3, 4, 5]. In general, two random variables X and Y have been called *positively* (resp. negatively) quadrant dependent (PQD) (resp. NQD) by Lehmann [6], if $P(X \ge x, Y \ge y) \ge (\text{resp.} \le) P(X \ge x) P(Y \ge y)$ for all $x, y \in \mathbb{R}$.

The objective of this paper is to establish a generalized uniform law of the iterated logarithm and investigate path properties for LPQD or LNQD random fields taking values in l^{∞} -space, whose description now follows. For $\mathbf{s} := (s_1, \dots, s_N)$, $\mathbf{t} := (t_1, \dots, t_N) \in [0, \infty)^N$, denote

$$\begin{aligned} \mathbf{0} &= (0, \cdots, 0), \quad \mathbf{1} = (1, \cdots, 1), \quad \mathbf{s} \pm \mathbf{t} = (s_1 \pm t_1, \cdots, s_N \pm t_N), \\ \mathbf{s} &\leq \mathbf{t} \quad \text{if } s_m \leq t_m \text{ for each } m = 1, 2, \cdots, N, \\ a\mathbf{t} &= (at_1, \cdots, at_N) \text{ for } a \in (-\infty, \infty), \quad (\mathbf{s}, \mathbf{t}) = (s_1, \cdots, s_N, t_1, \cdots, t_N) \in [0, \infty)^{2N} \end{aligned}$$

Assume that $\{X_i(\mathbf{t}); \mathbf{t} \in [0, \infty)^N\}_{i=1}^\infty$ is a sequence of centered strictly stationary and LPQD (or LNQD) random fields with $X_i(\mathbf{0}) = 0$ and stationary increments

$$\sigma_i(\|\mathbf{t}\|) := \sqrt{E\{X_i(\mathbf{s} + \mathbf{t}) - X_i(\mathbf{s})\}^2}, \quad i \ge 1,$$

where $\sigma_i(t)$ are nondecreasing continuous functions of t > 0, and $\|\cdot\|$ denotes the Euclidean norm such that $\|\mathbf{t}\| = \left(\sum_{m=1}^N t_m^2\right)^{1/2}$. Put

$$\sigma_*(t) = \sup_{i \ge 1} \sigma_i(t)$$

and assume that $\sigma_*(\cdot)$ is a regularly varying function with exponent $\alpha > 0$ at ∞ . Recall that a positive function $\sigma(t)$ of t > 0 is said to be *regularly varying* with exponent $\alpha > 0$ at $b \ge 0$ if $\lim_{t\to b} \{\sigma(xt)/\sigma(t)\} = x^{\alpha}$ for x > 0.

Let $\{\mathbb{X}(\mathbf{t}) := (X_1(\mathbf{t}), X_2(\mathbf{t}), \cdots); \mathbf{t} \in [0, \infty)^N\}$ be a centered strictly stationary and LPQD (LNQD) random field taking values in l^{∞} -space (i.e. l^{∞} -valued random field) with l^{∞} -norm $\|\cdot\|_{\infty}$ defined by $\|\mathbb{X}(\mathbf{t})\|_{\infty} = \sup_{i>1} |X_i(\mathbf{t})|$.

For each $m = 1, 2, \dots, N$, let $a_m(x)$ and $b_m(x)$ be positive nondecreasing functions of x > 0 such that $a_m(x) \le b_m(x)$ and $\lim_{x\to\infty} b_m(x) = \infty$. Denote

$$a_{x} = (a_{1}(x), \cdots, a_{N}(x)), \quad \mathbf{b}_{x} = (b_{1}(x), \cdots, b_{N}(x)),$$
$$\gamma(x) = \sqrt{2\{\log(\|\mathbf{b}_{x}\| / \|\boldsymbol{a}_{x}\|) + \log\log\|\mathbf{b}_{x}\|\}},$$

where $\log z = \log(\max\{z, e\})$.

The main results are as follows. Let $\{x_k; x_k > 0\}_{k=1}^{\infty}$ be an increasing sequence with $x_0 > 0$ and $\lim_{k\to\infty} x_k = \infty$, and let $u_k = O(v_k)$ denote $\limsup_{k\to\infty} u_k/v_k < \infty$.

Theorem 1.1. Let $\{X(t); t \in [0,\infty)^N\}$ be a centered strictly stationary and LPQD (LNQD) l^{∞} -valued random field with l^{∞} -norm $\|\cdot\|_{\infty}$ and $E|X_1(t)|^{2+\delta} < \infty$ for some $\delta \in (0,1]$, which satisfies conditions

 $\begin{array}{ll} \text{(i)} & \sum_{\substack{j \geq k+1 \\ x \geq 1}} |\operatorname{Cov}(X_i(\mathbf{1}), X_i(\mathbf{b}_j))| = O(\|\mathbf{b}_k\|^{-\lambda}) & \text{for each } i, k \geq 1 \text{ and some } \lambda > 2, \\ \text{(ii)} & \inf_{\substack{x \geq 1 \\ x \geq 1}} \sigma_*^2(x)/x > 0. \end{array}$

For each $m = 1, 2, \dots, N$, let the functions $a_m(x)$ and $b_m(x)$ satisfy conditions

- (iii) $b_m(x)/a_m(x) (> 1)$ is increasing,
- (iv) there exists $c_0 > 1$ such that $b_m(x_k) \le c_0 b_m(x_{k-1})$ for $k \ge 1$.

Then we have

$$\limsup_{x \to \infty} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbf{X}(\mathbf{s} + \mathbf{t}) - \mathbf{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\gamma(x)} = \limsup_{x \to \infty} \frac{\|\mathbf{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\gamma(x)} = 1 \quad \text{a.s.} \quad (1.3)$$

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Theorem 1.1 presents a path property for l^{∞} -valued random field, while we can obtain the following law of the iterated logarithm (LIL) without conditions (iii)-(iv) of Theorem 1.1.

Theorem 1.2. Let $\{X(t); t \in [0,\infty)^N\}$ be as in Theorem 1.1 with conditions (i)-(ii). Then

$$\lim_{x \to \infty} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\|\mathbf{b}_x\|}} = \lim_{x \to \infty} \sup_{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\|\mathbf{b}_x\|}} = 1 \quad \text{a.s.}$$
(1.4)

Note that the first result in (1.4) implies a generalized uniform law of the iterated logarithm for LPQD or LNQD l^{∞} -valued random fields, but the second one in (1.4) is a standard form of the ordinary LIL for any dependent (or independent) l^{∞} -valued random fields, which is an extension of some theorems in [1, 2, 4, 8].

Returning to our present exposition of Theorem 1.2, we present the following examples.

Example 1.3. Let $\{X_i(\mathbf{t}); \mathbf{t} \in [0, \infty)^N\}_{i=1}^{\infty}$ be a sequence of centered stationary and independent l^{∞} -valued Gaussian random fields with exponent $\alpha = 1/2$ (e.g. Wiener random field). For each $i = 1, 2, \ldots, N$, let $b_i(x) = \sqrt{i}x$. Then

 $\mathbf{b}_x := (b_1(x), \cdots, b_N(x)) = (1, \sqrt{2}, \cdots, \sqrt{N})x, \quad \|\mathbf{b}_x\| = \sqrt{N(N+1)/2} x.$

Hence, by Theorem 1.2, we have the uniform law of the iterated logarithm

$$\limsup_{x \to \infty} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sqrt{\|\mathbf{b}_x\|}\sqrt{2\log\log\|\mathbf{b}_x\|}} = 1 \qquad \text{a.s.}$$

From the ordinary LIL in (1.4), one can obtain Theorem 1 in [4] for LPQD random sequence $\{\xi_n; n \ge 1\}$, as in Example 1.4 below. In (1.1) and Theorem 1.2, denote $X(n) = X_n(\mathbf{t}_n), X(n) = S_n := \xi_1 + \cdots + \xi_n$ and $\sigma(n) := \sqrt{E(S_n)^2}$, when indexed by a single time-parameter n in Theorems 1.1-1.2.

Example 1.4. Let $\{\xi_n ; n \ge 1\}$ be a centered strictly stationary and LPQD (or LNQD) random sequence with $E\xi_1^2 > 0$, which satisfies conditions

- (i) $E|\xi_1|^p < \infty$ for each p > 2,
- (ii) $\sum_{j \ge k+1} |\operatorname{Cov}(\xi_1, \xi_j)| = O(k^{-\lambda})$ for each $k \ge 1$ and some $\lambda > 2$, (iii) $\sigma^2 := E\xi_1^2 + 2\sum_{j=2}^{\infty} |\operatorname{Cov}(\xi_1, \xi_j)| < \infty$.

Then we have

$$\limsup_{n \to \infty} \frac{S_n}{\sigma(n)\sqrt{2\log\log n}} = 1 \qquad \text{a.s.,}$$

where it is easy to prove that $\sigma(n) \approx \sigma \sqrt{n}$ for *n* large enough.

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2 Proofs

In this section, let c denote a positive constant which may take different values whenever they appear in different lines. We need the following properties.

 (P_1) Two random variables X and Y are PQD (resp. NQD) if and only if $Cov(f(X), g(Y)) \ge (resp. \le) 0$ for all real-valued nondecreasing functions f and g (such that f(X) and g(Y) have finite variances) (see Lehmann [6]);

 (P_2) (Hoeffding equality): For any absolutely continuous functions f and g on the real line and for any random variables X and Y satisfying $Ef^2(X) + Eg^2(Y) < \infty$, we have

$$\operatorname{Cov}\left(f(X), g(Y)\right) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f'(x)g'(y) \left\{P(X \ge x, \ Y \ge y) - P(X \ge x)P(Y \ge y)\right\} dxdy.$$

The main ingredients of the proofs of Theorems 1.1-1.2 are Propositions 2.1-2.3 below. Note that conditions (i)-(ii) in Theorem 1.1 imply conditions (C2) and (I)-(II) in [4] and [5], respectively. Moreover, $||X(t)||_{\infty}/\sigma_*(||t||)$ is a standardized random variable. Thus Lemma 2 in [4] and Corollary 2.1 in [5] are easily changed to the following Berry-Esseen type theorem.

Proposition 2.1 (Berry-Esseen type theorem). Let $\{X(t); t \in [0, \infty)^N\}$ be as in Theorem 1.1 with conditions (i)-(ii). Then

$$\sup_{z \in \mathbb{R}} \left| P\left\{ \frac{\|\mathbb{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)} \le z \right\} - \Phi(z) \right| = O\left(\|\mathbf{b}_x\|^{-1/5}\right), \quad x \to \infty,$$

where $\Phi(\cdot)$ is a standard normal distribution function and $\|\mathbf{b}_x\| \to \infty$ as $x \to \infty$.

Denote $\mathbf{b}_k = \mathbf{b}_{x_k}$ for a positive increasing sequence $\{x_k\}_{k=1}^{\infty}$. Using Proposition 2.1 above, the following proposition is immediate from the proof of Lemma 9 in Petrov [7, p. 311].

Proposition 2.2. Let $\{X(\mathbf{t})\}$ be as in Proposition 2.1. Assume that g(x) is a positive nondecreasing function of x > 0 and that $\{\|\mathbf{b}_k\|; k \ge 1\}$ is a positive nondecreasing sequence such that $\sum_{k=1}^{\infty} \|\mathbf{b}_k\|^{-1/5} < \infty$. Then the following statements are equivalent.

(A)
$$\sum_{k=1}^{\infty} P\left\{\frac{\|\mathbf{X}(\mathbf{b}_{k})\|_{\infty}}{\sigma_{*}(\|\mathbf{b}_{k}\|)} > g(\|\mathbf{b}_{k}\|)\right\} < \infty,$$

(B)
$$\sum_{k=1}^{\infty} \frac{1}{g(\|\mathbf{b}_{k}\|)} \exp\left(-\frac{1}{2}g^{2}(\|\mathbf{b}_{k}\|)\right) < \infty$$

The next proposition on the large deviation probability is essential to prove our theorems for *any* strictly stationary l^{∞} -valued random field, which is proved in a way similar to those of Lemmas 2.2 and 2.3 in [8].

Proposition 2.3. Let $\{X(t); t \in [0,\infty)^N\}$ be a centered strictly stationary l^∞ -valued random field. Then, for any $\varepsilon > 0$ there exists a constant $c_{\varepsilon} > 0$ such that, for v > 1,

$$P\left\{\sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)} \ge v\right\}$$
$$\leq c_{\varepsilon} \left(P\left\{\frac{\|\mathbb{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)} \ge \frac{v}{1+\varepsilon}\right\} + \sum_{n=1}^{\infty} 2^{2N2^n} P\left\{\frac{\|\mathbb{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)} \ge \frac{v}{1+\varepsilon}\sqrt{1+2N\log 3} \cdot 2^{n/2}\right\}\right).$$

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Proof of Theorem 1.1. Since $\gamma(x) \geq \sqrt{2\log \log \|\mathbf{b}_x\|}$, we will first prove the sharper result

$$\limsup_{x \to \infty} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbf{X}(\mathbf{s} + \mathbf{t}) - \mathbf{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\|\mathbf{b}_x\|}} \le 1 \qquad \text{a.s.}$$
(2.1)

without conditions (iii)-(iv), whose result is used to prove (1.4). For $\theta > 1$ and $k \ge 1$, set $A_k = \{x; \theta^k \le \|\mathbf{b}_x\| \le \theta^{k+1}\}$. Note that $\sqrt{2\log\log\theta^k} \ge \theta^{-1}\sqrt{2\log\log\theta^{k+1}}$ since $(\log u)/u$ is decreasing for $u > e^e$. By the regularity of $\sigma_*(\cdot)$, we get $\sigma_*(\|\mathbf{b}_x\|)/\sigma_*(\theta^{k+1}) \ge \theta^{-2\alpha}$ and hence

$$\limsup_{x \to \infty} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\|\mathbf{b}_x\|}}
\le \limsup_{k \to \infty} \sup_{x \in A_k} \sup_{\|\mathbf{s}\| \le \|\mathbf{b}_x\|} \sup_{\|\mathbf{t}\| \le \|\mathbf{b}_x\|} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\theta^k}}
\le \theta^{1+2\alpha} \limsup_{k \to \infty} \sup_{\|\mathbf{s}\| \le \theta^{k+1}} \sup_{\|\mathbf{t}\| \le \theta^{k+1}} \frac{\|\mathbb{X}(\mathbf{s} + \mathbf{t}) - \mathbb{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\theta^{k+1})\sqrt{2\log\log\theta^{k+1}}}.$$
(2.2)

For convenience, let $\|\mathbf{b}_k\| = \theta^k$, where $\mathbf{b}_k := \mathbf{b}_{x_k}$ for a positive increasing sequence $\{x_k\}_{k=1}^{\infty}$. Using Proposition 2.3, it follows that for any $\varepsilon > 0$ there exists a positive constant c_{ε} such that

$$P\left\{\sup_{\|\mathbf{s}\| \le \theta^{k+1}} \sup_{\|\mathbf{t}\| \le \theta^{k+1}} \frac{\|\mathbf{X}(\mathbf{s} + \mathbf{t}) - \mathbf{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\theta^{k+1})\sqrt{2\log\log\theta^{k+1}}} > 1 + 2\varepsilon\right\}$$

$$\leq c_{\varepsilon} \left(P\left\{\frac{\|\mathbf{X}(\mathbf{b}_{k+1})\|_{\infty}}{\sigma_*(\theta^{k+1})} \ge \frac{(1 + 2\varepsilon)\sqrt{2\log\log\theta^{k+1}}}{1 + \varepsilon}\right\}$$

$$+ \sum_{n=1}^{\infty} 2^{2N2^n} P\left\{\frac{\|\mathbf{X}(\mathbf{b}_{k+1})\|_{\infty}}{\sigma_*(\theta^{k+1})} \ge \frac{(1 + 2\varepsilon)\sqrt{2\log\log\theta^{k+1}}}{1 + \varepsilon}\sqrt{1 + 2N\log3} \cdot 2^{n/2}\right\}\right).$$
(2.3)

Now let us apply Proposition 2.2 with $\|\mathbf{b}_k\| = \theta^k$ and $g(\theta^k) = g_1(\theta^k)$ (or $g_2(\theta^k)$), where

$$g_1(\theta^k) := \frac{(1+2\varepsilon)\sqrt{2\log\log\theta^{k+1}}}{1+\varepsilon}, \quad g_2(\theta^k) := \frac{(1+2\varepsilon)\sqrt{2\log\log\theta^{k+1}}}{1+\varepsilon}\sqrt{1+2N\log3} \cdot 2^{n/2}$$

in (2.3). Considering the right hand side of (2.3) and equivalence of Proposition 2.2, we have

$$\sum_{k=1}^{\infty} \frac{1}{g_1(\theta^k)} e^{-g_1^2(\theta^k)/2} \le c \sum_{k=1}^{\infty} \left(\log \theta^{k+1}\right)^{-1-\varepsilon'} < \infty$$
$$\Rightarrow \sum_{k=1}^{\infty} P\left\{\frac{\|\mathbf{X}(\mathbf{b}_{k+1})\|_{\infty}}{\sigma_*(\theta^{k+1})} \ge g_1(\theta^k)\right\} < \infty,$$

where $\varepsilon'=\varepsilon/(1+\varepsilon),$ by the strict stationarity of $\mathbb{X}(\mathbf{t})$, and also

$$\frac{1}{g_2(\theta^k)} \exp\left(-\frac{1}{2}g_2^2(\theta^k)\right) \le \exp\left(-\frac{1}{2}\left(\frac{1+2\varepsilon}{1+\varepsilon}\right)^2 (2\log\log\theta^{k+1})(1+2N\log3)2^n\right) \\ \le ((k+1)\log\theta)^{-(1+\varepsilon')(1+2N\log3)2^n} \le c\,(k+1)^{-(1+\varepsilon')(1+2N\log3)2^n}$$

which implies

$$P\left\{\frac{\|\mathbb{X}(\mathbf{b}_{k+1})\|_{\infty}}{\sigma_*(\theta^{k+1})} > g_2(\theta^k)\right\} \le c \, (k+1)^{-(1+\varepsilon')(1+2N\log 3)2^n}$$

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by Proposition 2.2. Thus

$$\sum_{k=1}^{\infty} \sum_{n=1}^{\infty} 2^{2N2^n} P\Big\{\frac{\|\mathbf{X}(\mathbf{b}_{k+1})\|_{\infty}}{\sigma_*(\theta^{k+1})} > g_2(\theta^k)\Big\} \le c \sum_{k=1}^{\infty} \sum_{n=1}^{\infty} 2^{2N2^n} (k+1)^{-(1+\varepsilon')(1+2N\log 3)2^n} \\ \le c \sum_{k=1}^{\infty} \sum_{n=1}^{\infty} 2^{-N2^n \log_2(k+1)} \cdot 2^{-n} \le c \sum_{k=1}^{\infty} \sum_{n=1}^{\infty} k^{-2} \cdot 2^{-n} < \infty.$$

In conclusion, it follows from (2.3) that

$$\sum_{k=1}^{\infty} P \left\{ \sup_{\|\mathbf{s}\| \le \theta^{k+1}} \sup_{\|\mathbf{t}\| \le \theta^{k+1}} \frac{\| \mathbf{X}(\mathbf{s} + \mathbf{t}) - \mathbf{X}(\mathbf{s}) \|_{\infty}}{\sigma_*(\theta^{k+1}) \sqrt{2 \log \log \theta^{k+1}}} > 1 + 2\varepsilon \right\} < \infty$$

and the Borel-Cantelli lemma yields

$$\limsup_{k \to \infty} \sup_{\|\mathbf{s}\| \le \theta^{k+1}} \sup_{\|\mathbf{t}\| \le \theta^{k+1}} \frac{\|\mathbf{X}(\mathbf{s} + \mathbf{t}) - \mathbf{X}(\mathbf{s})\|_{\infty}}{\sigma_*(\theta^{k+1})\sqrt{2\log\log\theta^{k+1}}} \le 1 + 2\varepsilon \quad \text{a.s}$$

Combining this with (2.2) implies (2.1) since ε and θ are arbitrary.

By virtue of (2.1), the proof of (1.3) is completed if we show that

$$\limsup_{x \to \infty} \frac{\|\mathbf{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\gamma(x)} \ge 1 \qquad \text{a.s.}$$
(2.4)

Let $\{x_k; x_k > 0\}_{k=1}^{\infty}$ be an increasing sequence such that $x_0 > 0$ and the (k-1)st point x_{k-1} is placed by the relation $b_m(x_k) - a_m(x_k) = b_m(x_{k-1}), 1 \le m \le N$, with x_k defined by induction, since $b_m(x) - a_m(x)$ is increasing by (iii). For convenience, put $\mathbf{a}_k = \mathbf{a}_{x_k}$ and $\mathbf{b}_k = \mathbf{b}_{x_k}$, and let $i_0 \ge 1$ be an integer such that $\sigma_{i_0}(\|\mathbf{b}_k\|) = \sigma_*(\|\mathbf{b}_k\|)$, where $\|\mathbf{b}_k\| := \theta^k$ as above. Then,

$$\limsup_{k \to \infty} \frac{\| \mathbb{X}(\mathbf{b}_k) \|_{\infty}}{\sigma_*(\|\mathbf{b}_k\|)\gamma(x_k)} \ge \limsup_{k \to \infty} \frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\|\mathbf{b}_k\|)\gamma(x_k)}$$
(2.5)

and the inequality (2.4) is immediate from (2.5) if we prove

$$\limsup_{k \to \infty} \frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\|\mathbf{b}_k\|)\gamma(x_k)} > 1 - 4\varepsilon \quad \text{a.s.}$$
(2.6)

for any small $\varepsilon > 0$. For each $k \ge 1$, we see that

$$U_k := \frac{X_{i_0}(\mathbf{b}_k) - X_{i_0}(\mathbf{b}_{k/2})}{\sigma_{i_0}(\|\mathbf{b}_k - \mathbf{b}_{k/2}\|)}$$

is a standardized random variable. Let $B_k = \{U_k > (1 - 2\varepsilon)\gamma(x_k)\}$. If \mathcal{N} is a standard normal random variable, then it follows from Proposition 2.1 and the strict stationarity of $\mathbb{X}(\mathbf{t})$ that

$$P(B_k) = \left(1 - P\left\{U_k \le (1 - 2\varepsilon)\gamma(x_k)\right\} - 1 + P\left\{\mathcal{N} \le (1 - 2\varepsilon)\gamma(x_k)\right\}\right) \\ + P\left\{\mathcal{N} > (1 - 2\varepsilon)\gamma(x_k)\right\} \\ \ge -c_1 \|\mathbf{b}_k - \mathbf{b}_{k/2}\|^{-1/5} + \frac{1}{\sqrt{2\pi}(1 - 2\varepsilon)^2 \gamma^2(x_k)} \exp\left(-\frac{1 - 2\varepsilon}{2}\gamma^2(x_k)\right) \\ \ge -c_1 \frac{1}{(\log \|\mathbf{b}_k\|)^{1-\varepsilon}} + c\left(\frac{\|\mathbf{a}_k\|}{\|\mathbf{b}_k\|\log \|\mathbf{b}_k\|}\right)^{1-\varepsilon} \ge c\frac{1}{(\log \|\mathbf{b}_k\|)^{1-\varepsilon}} \left(\varepsilon \frac{\|\mathbf{a}_k\|}{\|\mathbf{b}_k\|}\right)$$

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for all large k by (iii), where c and c_1 are positive constants, and further

$$\sum_{k=k_0}^{\ell} P(B_k) \ge \varepsilon \frac{1}{(\log \|\mathbf{b}_{\ell}\|)^{1-\varepsilon}} \sum_{k=k_0}^{\ell} \frac{\|\mathbf{a}_k\|}{\|\mathbf{b}_k\|}$$

for some $k_0 \ge 1$ with $k_0 \le k \le \ell$. Also there exist constants $c_2, c_3 > 1$ such that

$$\log \|\mathbf{b}_{\ell}\| \le c_2 \sum_{k=k_0}^{\ell} \log \frac{\|\mathbf{b}_k\|}{\|\mathbf{b}_{k-1}\|} \le c_2 \sum_{k=k_0}^{\ell} \log \left(c_0 + \frac{\|\boldsymbol{a}_k\|}{\|\mathbf{b}_{k-1}\|}\right) \le c_2 \sum_{k=k_0}^{\ell} \log \left(c_0 + \frac{c_3 \|\boldsymbol{a}_{k-1}\|}{\|\mathbf{b}_{k-1}\|}\right)$$
(2.7)

since $c_0 \| \mathbf{b}_{k-1} \| \ge \| \mathbf{b}_k \| - \| \mathbf{a}_k \|$ by (iv). The last inequality of (2.7) follows from the fact that

$$\frac{\|\boldsymbol{a}_{k}\|}{\|\boldsymbol{a}_{k-1}\|} \le \frac{\|\mathbf{b}_{k}\|}{\|\mathbf{b}_{k-1}\|} \le \frac{c_{0} \|\mathbf{b}_{k}\|}{\|\mathbf{b}_{k}\| - \|\boldsymbol{a}_{k}\|} = \frac{c_{0}}{1 - (\|\boldsymbol{a}_{k}\| / \|\mathbf{b}_{k}\|)} \le c_{3}$$

by (iii). Thus, by (2.7), there exists a constant K > 1 such that

$$\log \|\mathbf{b}_{\ell}\| \le c_2 \sum_{k=k_0}^{\ell} \log \left(c_0 + \frac{c_3 \|\boldsymbol{a}_k\|}{\|\mathbf{b}_{k-1}\|} \right) \le K \sum_{k=k_0}^{\ell} \frac{c_3^2 \|\boldsymbol{a}_k\|}{\|\mathbf{b}_k\|}.$$

Therefore, we have $\sum_{k=1}^{\ell} P(B_k) \ge \varepsilon (\log \|\mathbf{b}_{\ell}\|)^{\varepsilon} / (Kc_3^2) \to \infty$ as $\ell \to \infty$; that is, we get $\sum_{k=1}^{\infty} P(B_k) = \infty$.

Next, let $B'_k = \{U_k > (1 - 3\varepsilon)\gamma(x_k)\}$. We will show that $P(B'_k, i.o.) = 1$. Choose a differential function f(x) on \mathbb{R} such that $|f'(x)| \le \kappa$ for some $0 < \kappa < \infty$ and

$$0 \le I\left\{x > (1 - 2\varepsilon)\gamma(x_k)\right\} \le f(x) \le I\left\{x > (1 - 3\varepsilon)\gamma(x_k)\right\} \le 1,$$
(2.8)

where $I\{\cdot\}$ is an indicator function. In order to prove $P(B'_k, i.o.) = 1$, it is enough to show that $\sum_{k=1}^{\infty} f(U_k) = \infty$ a.s. Since $\sum_{k=1}^{\infty} P(B_k) = \infty$ in the above statement, it follows from (2.8) that $\sum_{k=1}^{\infty} Ef(U_k) \ge \sum_{k=1}^{\infty} P(B_k) = \infty$. By Markov inequality, we have

$$P\left\{\sum_{k=1}^{\infty} f(U_{k}) < \frac{1}{2} \sum_{k=1}^{n} Ef(U_{k})\right\} \le P\left\{\left|\sum_{k=1}^{n} f(U_{k}) - \sum_{k=1}^{n} Ef(U_{k})\right| > \frac{1}{2} \sum_{k=1}^{n} Ef(U_{k})\right\}\right.$$

$$\le 4 \operatorname{Var}\left(\sum_{k=1}^{n} f(U_{k})\right) \left/ \left(\sum_{k=1}^{n} Ef(U_{k})\right)^{2}$$

$$\le \frac{4}{\sum_{k=1}^{n} Ef(U_{k})} + \frac{8 \sum_{k=1}^{\infty} \sum_{j=k+1}^{\infty} |\operatorname{Cov}\left(f(U_{k}), f(U_{j})\right)|}{\left(\sum_{k=1}^{n} Ef(U_{k})\right)^{2}}.$$
(2.9)

Noting that U_k and U_j are LPQD (resp. LNQD) from the definition of LPQD (resp.

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LNQD), it follows from (P_1) , (P_2) , condition (i) and the regularity of $\sigma_*(\cdot)$ that

$$\sum_{k=1}^{\infty} \sum_{j=k+1}^{\infty} \left| \operatorname{Cov} \left(f(U_k), f(U_j) \right) \right|$$

$$\leq \kappa^2 \sum_{k=1}^{\infty} \sum_{j=k+1}^{\infty} \left| \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \left(P\{U_k \ge x, U_j \ge y\} - P\{U_k \ge x\} P\{U_j \ge y\} \right) dx dy \right|$$

$$\leq c \sum_{k=1}^{\infty} \frac{\|\mathbf{b}_k - \mathbf{b}_{k/2}\|}{\sigma_{i_0}^2(\|\mathbf{b}_k - \mathbf{b}_{k/2}\|)} \sum_{j=k+1}^{\infty} \left| \operatorname{Cov} \left(X_{i_0}(\mathbf{1}), X_{i_0}(\mathbf{b}_j) - X_{i_0}(\mathbf{b}_{j/2}) \right) \right|$$

$$\leq c \sum_{k=1}^{\infty} (\theta^k)^{1-2\alpha} \|\mathbf{b}_{(k+1)/2}\| \sum_{j\ge k+1} \left| \operatorname{Cov} \left(X_{i_0}(\mathbf{1}), X_{i_0}(\mathbf{b}_j) \right) \right|$$

$$\leq c \sum_{k=1}^{\infty} \theta^{-(\lambda-2+2\alpha)k} < \infty.$$
(2.10)

Since $\sum_{k=1}^{\infty} Ef(U_k) = \infty$ as above, letting $n \to \infty$ in (2.9) yields $P\left\{\sum_{k=1}^{\infty} f(U_k) < \infty\right\} = 0$ by (2.10). This proves $\sum_{k=1}^{\infty} f(U_k) = \infty$ a.s. and consequently $P(B'_k, i.o.) = 1$. Let

$$C_{k} = \left\{ \frac{X_{i_{0}}(\mathbf{b}_{k/2})}{\sigma_{i_{0}}(\|\mathbf{b}_{k/2}\|)} \ge -2\gamma(x_{k/2}) \right\}.$$

Since $P(B'_k, i.o.) = 1$, it follows from (2.1) that $P(B'_k \cap C_k, i.o.) = 1$. It is easy to see that

$$P\left\{\frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\|\mathbf{b}_k\|)} > (1-4\varepsilon)\gamma(x_k), \ i.o.\right\}$$

$$\geq P\left\{\frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\|\mathbf{b}_k\|)} > (1-3\varepsilon)\gamma(x_k) - 2\gamma(x_{k/2}), \ i.o.\right\}$$

$$\geq P\left\{B'_k \cap C_k, \ i.o.\right\} = 1$$

for k large enough, by the stationarity of $\mathbb{X}(t).$ This implies (2.6) and hence (2.4) holds true.

Proof of Theorem 1.2. Since we have proved (2.1) without conditions (iii)-(iv) of Theorem 1.1, it is enough to show that

$$\limsup_{x \to \infty} \frac{\|\mathbb{X}(\mathbf{b}_x)\|_{\infty}}{\sigma_*(\|\mathbf{b}_x\|)\sqrt{2\log\log\|\mathbf{b}_x\|}} \ge 1 \qquad \text{a.s.}$$
(2.11)

Set $\mathbf{b}_k = \mathbf{b}_{x_k}$ for a positive increasing sequence $\{x_k\}_{k=1}^{\infty}$, and let $i_0 \ge 1$ be an integer such that $\sigma_{i_0}(||\mathbf{b}_k||) = \sigma_*(||\mathbf{b}_k||)$. Then

$$\limsup_{k \to \infty} \frac{\| \mathbb{X}(\mathbf{b}_k) \|_{\infty}}{\sigma_*(\| \mathbf{b}_k \|) \sqrt{2 \log \log \| \mathbf{b}_k \|}} \ge \limsup_{k \to \infty} \frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\| \mathbf{b}_k \|) \sqrt{2 \log \log \| \mathbf{b}_k \|}}$$
(2.12)

and (2.11) is immediate from (2.12) if we prove

$$\limsup_{k \to \infty} \frac{X_{i_0}(\mathbf{b}_k)}{\sigma_{i_0}(\|\mathbf{b}_k\|)\sqrt{2\log\log\|\mathbf{b}_k\|}} > 1 - 4\varepsilon \quad \text{ a.s.}$$

for any small $\varepsilon > 0$. For $\theta > 1$, set $\|\mathbf{b}_k\| = \theta^k$ and $B_k^* = \{U_k > (1 - 2\varepsilon)\sqrt{2\log\log\|\mathbf{b}_k - \mathbf{b}_{k/2}\|}\}$ as in the proof of (2.6). Then $\|\mathbf{b}_k - \mathbf{b}_{k/2}\| \approx \theta^k$ for sufficiently large k. If we apply Proposition 2.2 with $g(\|\mathbf{b}_k - \mathbf{b}_{k/2}\|) = (1 - 2\varepsilon)\sqrt{2\log\log\|\mathbf{b}_k - \mathbf{b}_{k/2}\|}$, then

$$\sum_{k=1}^{\infty} \frac{1}{g\left(\|\mathbf{b}_k - \mathbf{b}_{k/2}\|\right)} \exp\left(-\frac{1}{2}g^2\left(\|\mathbf{b}_k - \mathbf{b}_{k/2}\|\right)\right) \ge c \sum_{k=1}^{\infty} k^{-1+\varepsilon} = \infty$$

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and hence $\sum_{k=1}^{\infty} P(B_k^*) = \infty$ by the strict stationarity of $X_i(\mathbf{t})$ for $i \ge 1$. The remainder of the proof is the same as the corresponding proof in (2.8)-(2.10). The details are omitted. This completes the proof of Theorem 1.2.

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