

Weak approximation of the Brownian sheet from a Poisson process in the plane

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We show an approximation in law of the Brownian sheet by processes constructed from the Poisson process in the plane. This result was inspired by a similar result of Stroock in the one-parameter case.

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1. Introduction and main result

The purpose of this paper is to prove a weak convergence to the Brownian sheet result for processes constructed from a two-parameter Poisson process. We seek an analogous result, in the two-parameter case, of the following theorem proved by Stroock (1982).

Theorem. *Consider a standard Poisson process, $\{N(t), t \geq 0\}$, and define, for any $\varepsilon > 0$, the continuous processes*

$$y_\varepsilon = \left\{ y_\varepsilon(t) := \varepsilon \int_0^{t/\varepsilon^2} (-1)^{N(s)} ds, t \in [0, T] \right\}.$$

If (P^ε) are the laws of the y_ε in the Banach space $\mathcal{C}([0, T])$ of continuous functions on $[0, T]$, then (P^ε) converges weakly, as ε tends to zero, towards the Wiener measure.

A motivation for proving results of this type is that they provide examples of processes of finite variation that can be approximated in law by the Wiener process. These processes have very different properties from the classical examples constructed from sums of independent random variables and from stationary processes, which also converge in law to the Wiener process.

Another point of interest is that they give a nice relationship between the two more important processes.

Our result is the following:

Theorem 1.1. *Define*

$$\left\{ x_\varepsilon(s, t) := \varepsilon \int_0^{t/\varepsilon} \int_0^{s/\varepsilon} \sqrt{xy}(-1)^{N(x,y)} dx dy; (s, t) \in [0, S] \times [0, T] \right\},$$

where $\{N(x, y); x, y \geq 0\}$ is a Poisson process in the plane.

Consider P_ε the image law of x_ε in the Banach space $\mathcal{C}([0, S] \times [0, T])$ of continuous functions on $[0, S] \times [0, T]$. Then, (P_ε) converges weakly, as ε tends to zero, towards the law on $\mathcal{C}([0, S] \times [0, T])$ of a Brownian sheet.

The integrand $(-1)^{N(x,y)}$ changes the sign very quickly if there are a lot of points around it. So, when ε tends to zero, $(-1)^{N(x/\varepsilon,y/\varepsilon)}$ tends to something which has independent values at each point and, properly normalized, is approximately white noise.

One might expect that the result in the two-parameter case was that the processes defined by

$$Y_\varepsilon(s, t) := \varepsilon \int_0^{t/\varepsilon} \int_0^{s/\varepsilon} (-1)^{N(x,y)} dx dy$$

converge weakly to the Brownian sheet. But it can be proved that the $Y_\varepsilon(s, t)$ converge to zero, as ε tends to zero, in $L^2(\Omega)$, for all $(s, t) \in [0, S] \times [0, T]$.

An intuitive reason for this apparent pathology is that the speed of convergence is not the same for all points (x, y) . It is slower near the origin than further away, and the square root factor expresses that. Another reason is the following.

We can write the Stroock processes as

$$y_\varepsilon(t) = \int_0^t \frac{1}{\varepsilon} (-1)^{N(s/\varepsilon^2)} ds,$$

the process defined in Theorem 1.1 as

$$x_\varepsilon(s, t) = \int_0^t \int_0^s \frac{1}{\varepsilon^2} \sqrt{xy} (-1)^{N(x/\varepsilon,y/\varepsilon)} dx dy$$

and

$$Y_\varepsilon(s, t) = \int_0^t \int_0^s \frac{1}{\varepsilon} (-1)^{N(x/\varepsilon,y/\varepsilon)} dx dy.$$

If we consider the covariance function of the integrand process in the expression of y_ε ,

$$\begin{aligned} K_\varepsilon(t, t') &= E \left[\frac{1}{\varepsilon^2} (-1)^{N(t/\varepsilon^2)} (-1)^{N(t'/\varepsilon^2)} \right] - \frac{1}{\varepsilon^2} E[(-1)^{N(t/\varepsilon^2)}] E[(-1)^{N(t'/\varepsilon^2)}] \\ &= \frac{1}{\varepsilon^2} \exp \left[-\frac{2|t - t'|}{\varepsilon^2} \right] - \frac{1}{\varepsilon^2} \exp \left[-\frac{2(t + t')}{\varepsilon^2} \right], \end{aligned}$$

it is clear that, for any $t > 0$, as a function of t' , this covariance converges weakly, as ε tends

to zero, to δ_t , the unit mass measure at the point t (which is the ‘covariance function’ of the white noise).

On the other hand, if we compute the covariance for the integrand processes in the expression of Y_ε and x_ε , we can see that $K_\varepsilon^Y((s, t), (s', t'))$ as a function of (s', t') converges weakly to zero, while $K_\varepsilon^x((s, t), (s', t'))$ tends weakly to $\delta_{(s,t)}$. These two last facts can be proved by similar arguments to those used in the proof of Lemma 4.4 below.

In order to simplify the notation we denote by $N_\mu(x, y)$ the random variable $N(x\sqrt{\mu}, y\sqrt{\mu})$. Then $\{N_\mu(x, y); (x, y) \in \mathbb{R}_+^2\}$ is a Poisson process with intensity μ . Note that

$$x_\varepsilon(s, t) = \frac{1}{\varepsilon^2} \int_0^t \int_0^s \sqrt{xy}(-1)^{N_{1/\varepsilon^2}(x,y)} dx dy.$$

Setting $n = 1/\varepsilon^2$, we are looking for the weak limit as $n \rightarrow \infty$ of

$$x_n(s, t) := n \int_0^t \int_0^s \sqrt{xy}(-1)^{N_n(x,y)} dx dy, \tag{1}$$

and we denote by P_n the image law of x_n in the space $\mathcal{C}([0, S] \times [0, T])$.

The paper is organized as follows. Section 2 is devoted to some preliminaries on two-parameter processes. The proof of tightness of the family of laws (P_n) is given in Section 3. Finally, in Section 4, we identify all the possible weak limits of subsequences of (P_n) as the Wiener measure.

A lot of the estimates in the paper contain constants (not depending on n). We use the same letter, K , for these constants, although their actual value can vary from one expression to the next.

2. Preliminaries

We will use the notation and definitions introduced in the basic work of Cairoli and Walsh (1975) on stochastic calculus in the plane. We recall some of them here.

Let (Ω, \mathcal{F}, P) be a complete probability space and let $\{\mathcal{F}_{s,t}; (s, t) \in [0, S] \times [0, T]\}$ be a family of sub- σ -fields of \mathcal{F} such that:

- (i) $\mathcal{F}_{s,t} \subseteq \mathcal{F}_{s',t'}$ for any $s \leq s', t \leq t'$;
- (ii) $\mathcal{F}_{0,0}$ contains all null sets of \mathcal{F} ;
- (iii) for each $z \in [0, S] \times [0, T]$, $\mathcal{F}_z = \bigcap_{z < z'} \mathcal{F}_{z'}$, where $z = (s, t) < z' = (s', t')$ denotes the partial order on $[0, S] \times [0, T]$, meaning that $s < s'$ and $t < t'$.

Given $(s, t) < (s', t')$, we denote by $\Delta_{s,t}X_{s',t'}$ the increment of the process X over the rectangle $((s, t), (s', t'))$, that is,

$$\Delta_{s,t}X_{s',t'} = X_{s',t'} - X_{s,t'} - X_{s',t} + X_{s,t}.$$

It is said that an \mathcal{F}_z -adapted process $X = \{X_z; z \in [0, S] \times [0, T]\}$ is a *martingale* if $E(|X_z|) < \infty$ for all $z \in [0, S] \times [0, T]$ and

$$E(X_{s',t'} - X_{s,t} | \mathcal{F}_{s,t}) = 0, \quad \text{for any } (s, t) < (s', t').$$

It is said that an \mathcal{F}_z -adapted process $X = \{X_z; z \in [0, S] \times [0, T]\}$ is a *strong martingale* if $E(|X_z|) < \infty$ for all $z \in [0, S] \times [0, T]$, $X_{s,0} = X_{0,s} = 0$ for all $s \geq 0$ and

$$E(\Delta_{s,t} X_{s',t'} | \mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}) = 0, \quad \text{for any } (s, t) < (s', t').$$

Definition 2.1. An $\mathcal{F}_{s,t}$ -Brownian sheet is a continuous, adapted process $W = \{W_{s,t}; (s, t) \in [0, S] \times [0, T]\}$ such that $W_{s,0} = W_{0,t} = 0$ almost surely (a.s.), for all $(s, t) \leq (s', t')$, the increment $\Delta_{s,t} W_{s',t'}$ is independent of $\mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}$ and is normally distributed with zero mean and variance $(s' - s)(t' - t)$.

If we do not specify the filtration, $(\mathcal{F}_{s,t})$ will be the filtration generated by the process itself, completed with the null sets of $\mathcal{F}^W = \sigma\{W_{s,t}, (s, t) \in [0, S] \times [0, T]\}$.

Definition 2.2. Let $\{\mathcal{F}_{s,t}\}$ be a family of sub- σ -fields of \mathcal{F} satisfying the previous conditions for all $(s, t) \in \mathbb{R}_+^2$. An $\mathcal{F}_{s,t}$ -Poisson process is an adapted, cadlag process $N = \{N_{s,t}; (s, t) \in \mathbb{R}_+^2\}$, such that, $N_{s,0} = N_{0,t} = 0$ a.s., for all $(s, t) \leq (s', t')$ the increment $\Delta_{s,t} N_{s',t'}$ is independent of $\mathcal{F}_{\infty,t} \vee \mathcal{F}_{s,\infty}$ and has a Poisson law of parameter $(s' - s)(t' - t)$. Here, we are denoting $\mathcal{F}_{\infty,t} := \bigvee_{s>0} \mathcal{F}_{s,t}$ and $\mathcal{F}_{s,\infty} := \bigvee_{t>0} \mathcal{F}_{s,t}$.

If we do not specify the filtration, $(\mathcal{F}_{s,t})$ will be the filtration generated by the process itself, completed with the nulls sets of $\mathcal{F}^N = \sigma\{N_{s,t}, (s, t) \in \mathbb{R}_+^2\}$.

3. Proof of tightness

To prove Theorem 1.1, we have to check that the family P_n is tight and that any weakly convergent subsequence converges to the law of a Brownian sheet. In this section we prove that P_n is tight. Using the criterion given by Bickel and Wichura (1971), and that our processes x_n are null on the axes, it suffices to prove the following lemma.

Lemma 3.1. Let $\{x_n\}$ be the family of processes defined by (1). There exists a constant K such that, for any $(s, t) < (s', t')$,

$$\sup_n E[(\Delta_{s,t} x_n(s', t'))^4] \leq K(s' - s)^2(t' - t)^2.$$

In order to prove Lemma 3.1 it will be useful to have the following result which we will also utilize in Section 4.

Lemma 3.2. Let $\{x_n\}$ be the family of processes defined by (1). Then if $(s, t) < (s', t')$,

$$E[(\Delta_{s,t} x_n(s', t'))^2] \leq 4(s' - s)(t' - t).$$

Proof.

$$\begin{aligned} E[(\Delta_{s,t}x_n(s', t'))^2] &= n^2 E \left[\left(\int_t^{t'} \int_s^{s'} \sqrt{xy} (-1)^{N_n(x,y)} dx dy \right)^2 \right] \\ &= n^2 E \left[\prod_{i=1}^2 \left(\int_t^{t'} \int_s^{s'} \sqrt{x_i y_i} (-1)^{N_n(x_i, y_i)} dx_i dy_i \right) \right] \\ &= n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} E[(-1)^{N_n(x_1, y_1) + N_n(x_2, y_2)}] dx_1 dx_2 dy_1 dy_2. \end{aligned}$$

Observe that $(-1)^{\sum_{i=1}^2 N_n(x_i, y_i)} = (-1)^{\sum_{i=1}^2 \Delta_{0,0} N_n(x_i, y_i)}$, and this last sum is equal to the sum of the increments of the Poisson process over some disjoint rectangles. Each one of these last increments appears once or twice. Obviously the rectangles which contribute to the value of $(-1)^{\sum_{i=1}^2 \Delta_{0,0} N_n(x_i, y_i)}$ are those that appear only once.

If we suppose that $x_1 \leq x_2$, there are two possible orders in the plane for the points (x_1, y_1) , (x_2, y_2) . (See Figure 1, where the black zones correspond to the rectangles that appear only once in the sum $\sum_{i=1}^2 \Delta_{0,0} N_n(x_i, y_i)$.)

Now, using the fact that the Poisson process has independent increments, and that if $Z \sim \text{Pois}(\lambda)$ then $E[(-1)^Z] = \exp(-2\lambda)$, we obtain that

$$E[(\Delta_{s,t}x_n(s', t'))^2] = 2(I_1 + I_2),$$

where

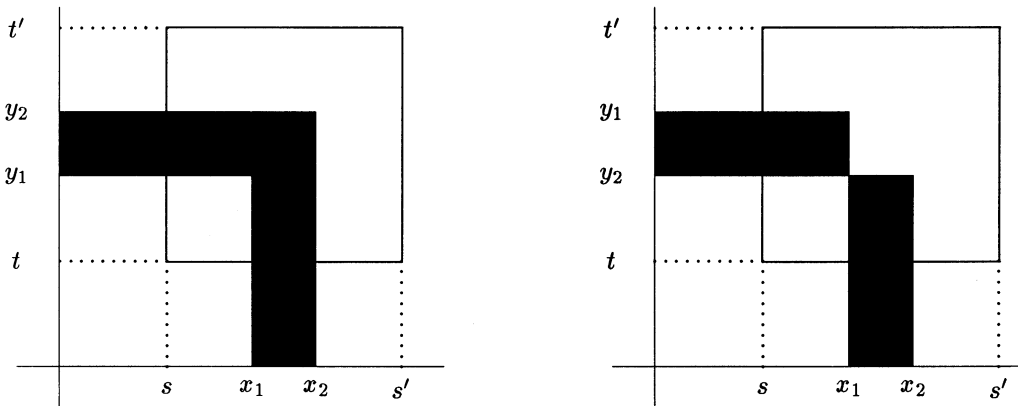


Figure 1. The two possible orders of two points in the plane.

$$I_1 = n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 y_2 - x_1 y_1)] I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2,$$

$$I_2 = n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 - x_1)y_2 - 2n(y_1 - y_2)x_1] \\ \times I_{\{x_1 \leq x_2\}} I_{\{y_2 \leq y_1\}} dx_1 dx_2 dy_1 dy_2.$$

By doing a change of variables in order to obtain $y_1 \leq y_2$, it is easy to see that $I_1 \leq I_2$. Then,

$$E[(\Delta_{s,t} x_n(s', t'))^2] \leq 4n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 - x_1)y_1 - 2n(y_2 - y_1)x_1] \\ \times I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2.$$

By using the fact that $x_2 \leq s'$, $y_2 \leq t'$ and then integrating with respect to these two variables, we obtain that the last expression is less than or equal to

$$\sqrt{s'}\sqrt{t'} \int_s^{s'} \frac{1}{\sqrt{x_1}} dx_1 \int_t^{t'} \frac{1}{\sqrt{y_1}} dy_1 = 4\sqrt{s'}(\sqrt{s'} - \sqrt{s})\sqrt{t'}(\sqrt{t'} - \sqrt{t}) \\ \leq 4(s' - s)(t' - t).$$

□

We are now ready to prove Lemma 3.1.

Proof of Lemma 3.1. By arguments of additivity it is enough to prove the lemma for the case where s and t are strictly positive, and $t' - t < t$ and $s' - s < s$.

We have that

$$E[(\Delta_{s,t} x_n(s', t'))^4] = n^4 E \left[\left(\int_t^{t'} \int_s^{s'} \sqrt{xy} (-1)^{N_n(x,y)} dx dy \right)^4 \right] \\ = n^4 E \left[\prod_{i=1}^4 \left(\int_t^{t'} \int_s^{s'} \sqrt{x_i y_i} (-1)^{N_n(x_i, y_i)} dx_i dy_i \right) \right].$$

Observe that $(-1)^{\sum_{i=1}^4 N_n(x_i, y_i)} = (-1)^{\sum_{i=1}^4 \Delta_{0,0} N_n(x_i, y_i)}$, and that

$$\sum_{i=1}^4 \Delta_{0,0} N_n(x_i, y_i) = \sum_{i=1}^4 \Delta_{s,t} N_n(x_i, y_i) + \sum_{i=1}^4 \Delta_{s,0} N_n(x_i, t) + \sum_{i=1}^4 \Delta_{0,t} N_n(s, y_i) + 4\Delta_{0,0} N_n(s, t).$$

So,

$$(-1)^{\sum_{i=1}^4 \Delta_{0,0} N_n(x_i, y_i)} = (-1)^{\sum_{i=1}^4 \Delta_{s,t} N_n(x_i, y_i)} (-1)^{\sum_{i=1}^4 \Delta_{s,0} N_n(x_i, t)} (-1)^{\sum_{i=1}^4 \Delta_{0,t} N_n(s, y_i)},$$

and these three factors are independent. If we suppose $x_1 \leq x_2 \leq x_3 \leq x_4$ and $y_1 \leq y_2 \leq y_3 \leq y_4$, we have that

$$E[(-1)^{\sum_{i=1}^4 \Delta_{s,0} N_n(x_i,t)}]E[(-1)^{\sum_{i=1}^4 \Delta_{0,t} N_n(s,y_i)}] = \exp[-2nt[(x_4 - x_3) + (x_2 - x_1)]] \exp[-2ns[(y_4 - y_3) + (y_2 - y_1)]];$$

by using the fact that $2t > t'$ and $2s > s'$, the last expression is less than or equal to

$$\exp[-nt'[(x_4 - x_3) + (x_2 - x_1)]] \exp[-ns'[(y_4 - y_3) + (y_2 - y_1)]] \leq \exp[-n[(x_4 - x_3)y_3 + (x_2 - x_1)y_1]] \exp[-n[(y_4 - y_3)x_3 + (y_2 - y_1)x_1]].$$

Finally, we can bound $E[(-1)^{\sum_{i=1}^4 \Delta_{s,t} N_n(x_i,y_i)}]$ by 1. So,

$$\begin{aligned} E[(\Delta_{s,t} x_n(s', t'))^4] &\leq K n^4 \int_{[s,s']^4 \times [t,t']^4} \prod_{i=1}^4 \sqrt{x_i y_i} \exp[-n(x_4 - x_3)y_3 + n(x_2 - x_1)y_1] \\ &\quad \times I_{\{x_1 \leq x_2 \leq x_3 \leq x_4\}} \exp[-n(y_4 - y_3)x_3 + n(y_2 - y_1)x_1] I_{\{y_1 \leq y_2 \leq y_3 \leq y_4\}} dx_1 \dots dy_4 \\ &\leq K \left(n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-n(x_2 - x_1)y_1 - n(y_2 - y_1)x_1] \right. \\ &\quad \left. \times I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2 \right)^2 \\ &\leq K \left(n^2 \int_{[\frac{s}{2},s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 - x_1)y_1 - 2n(y_2 - y_1)x_1] \right. \\ &\quad \left. \times I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2 \right)^2. \end{aligned}$$

Using the computations of Lemma 3.2, the last expression is bounded by $K(t' - t)^2(s' - s)^2$. □

4. Identification of the limit law

We have proved that the family P_n is tight. Now, we must show that the law of all possible weak limits is the law of a Brownian sheet.

Let $\{P_{n_i}\}_i$ be a subsequence of $\{P_n\}_n$ (which we will also denote by $\{P_n\}$) weakly convergent to some probability P . We want to show that P is the Wiener measure, that is, the canonical process $\{X_{s,t}(x) =: x(s, t)\}$ is a Brownian sheet under the probability P .

There exist various possible characterizations of a Brownian sheet; see, for example, Tudor (1980) or Florit and Nualart (1996). In particular, in Theorem 2.2 of Florit and Nualart (1996) necessary and sufficient conditions for a process to be a Brownian motion, with respect to an arbitrary filtration, are proved. If we just consider the case in which the underlying filtration is the natural one, we realize that we can weaken the hypotheses of

Theorem 2.2 of Florit and Nualart (1996). Then, we obtain the following characterization of the two-parameter Wiener process.

Theorem 4.1. *Let $X = \{X_{s,t}; (s, t) \in [0, S] \times [0, T]\}$ be a continuous process such that $X_{s,0} = X_{0,t} = 0$. And let $(\mathcal{F}_{s,t})$ be the natural filtration of X .*

Then, the following statements are equivalent:

- (i) *X is a Brownian sheet.*
- (ii) *For any $0 < s \leq s', 0 < t \leq t', E(\Delta_{s,t}X_{s',t'} | \mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}) = 0$ and $E[(\Delta_{s,t}X_{s',t'})^2 | \mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}] = (s' - s)(t' - t)$.*

The difference between this theorem and Theorem 2.2 of Florit and Nualart (1996) is that in order to obtain a Brownian sheet (with respect to its natural filtration) we need only to check the properties of strong martingale and ‘quadratic variation’ for increments over rectangles without intersection with axes.

So, in order to prove that the limit law is the Wiener measure, it suffices to prove the following two propositions.

Proposition 4.2. *Suppose that $\{P_n\}$ are the laws in $\mathcal{C}([0, S] \times [0, T])$ of the processes x_n defined by (1), and assume that $\{P_{n_i}\}$ is a subsequence weakly convergent to P . Let X be the canonical process and let $\{\mathcal{F}_{s,t}\}$ be its natural filtration. Then for any $0 < s \leq s', 0 < t \leq t', E_P(\Delta_{s,t}X_{s',t'} | \mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}) = 0$.*

Proposition 4.3. *Under the hypotheses of the above proposition, we have that*

$$E_P[(\Delta_{s,t}X_{s',t'})^2 | \mathcal{F}_{s,t} \vee \mathcal{F}_{s,T}] = (s' - s)(t' - t), \quad \text{for any } 0 < s \leq s', 0 < t \leq t'.$$

Proof of Proposition 4.2. We only need to prove that for any $(s_1, t_1), \dots, (s_m, t_m)$ and $\delta > 0$, with $s_i \leq S, t_i \leq t - \delta$ or $s_i \leq s - \delta, t_i \leq T$, for $i = 1, \dots, m$, and for any bounded continuous $\varphi: \mathbb{R}^m \rightarrow \mathbb{R}$,

$$E_P[\varphi(X_{s_1,t_1}, \dots, X_{s_m,t_m})(\Delta_{s,t}X_{s',t'})] = 0.$$

Since $P_n \xrightarrow{w} P$, and taking into account Lemma 3.1, we have that

$$\lim_{n \rightarrow \infty} E_{P_n}[\varphi(x(s_1, t_1), \dots, x(s_m, t_m))(\Delta_{s,t}x(s', t'))] = E_P[\varphi(x(s_1, t_1), \dots, x(s_m, t_m))(\Delta_{s,t}x(s', t'))].$$

Thus, it suffices to prove that

$$\lim_{n \rightarrow \infty} E_{P_n}[\varphi(x(s_1, t_1), \dots, x(s_m, t_m))(\Delta_{s,t}x(s', t'))] = 0.$$

We have that

$$\begin{aligned}
 & |E_{P_n}[\varphi(x(s_1, t_1), \dots, x(s_m, t_m))(\Delta_{s,t}x(s', t')))] \\
 &= |E[\varphi(x_n(s_1, t_1), \dots, x_n(s_m, t_m))(\Delta_{s,t}x_n(s', t')))] \\
 &= |E[\varphi(x_n(s_1, t_1), \dots, x_n(s_m, t_m))E[\Delta_{s,t}x_n(s', t')|\mathcal{F}_{s,t,\delta}^n]]] \\
 &\leq (E[\varphi^2(x_n(s_1, t_1), \dots, x_n(s_m, t_m))]^{1/2}(E[Y_n^2])^{1/2} \leq K(E[Y_n^2])^{1/2},
 \end{aligned}$$

where $\mathcal{F}_{s,t,\delta}^n = \mathcal{F}_{S,t-\delta}^n \vee \mathcal{F}_{s-\delta,T}^n$ and

$$Y_n := E \left[n \int_{[s,s'] \times [t,t']} \sqrt{xy}(-1)^{N_n(x,y)} dx dy \middle| \mathcal{F}_{s,t,\delta}^n \right].$$

Thus, it suffices to prove that Y_n converges to zero in L^2 when n goes to infinity. Observe that $\Delta_{s-\delta,t-\delta}N_n(s, t)$ is independent of $\mathcal{F}_{s,t,\delta}^n$, and

$$\begin{aligned}
 & E \left[n \int_{[s,s'] \times [t,t']} \sqrt{xy}(-1)^{N_n(x,y)} dx dy \middle| \mathcal{F}_{s,t,\delta}^n \right] \\
 &= E[(-1)^{\Delta_{s-\delta,t-\delta}N_n(s,t)}] E \left[n \int_{[s,s'] \times [t,t']} \sqrt{xy}(-1)^{N_n(x,y) - \Delta_{s-\delta,t-\delta}N_n(s,t)} dx dy \middle| \mathcal{F}_{s,t,\delta}^n \right],
 \end{aligned}$$

which clearly goes to zero in L^2 because the conditional expectation is L^2 -bounded by Lemma 3.2, and $E[(-1)^{\Delta_{s-\delta,t-\delta}N_n(s,t)}] = \exp[-2\delta^2 n]$ which tends to zero as $n \rightarrow \infty$. \square

Proof of Proposition 4.3. We have to prove that for all $(s_1, t_1), \dots, (s_m, t_m)$ with $s_i \leq S, t_i \leq t$ or $s_i \leq s, t_i \leq T$, for $i = 1, \dots, m$, and for all bounded continuous $\varphi: \mathbb{R}^m \rightarrow \mathbb{R}$,

$$E_P[\varphi(X_{s_1,t_1}, \dots, X_{s_m,t_m})((\Delta_{s,t}X_{s',t'})^2 - (s' - s)(t' - t))] = 0.$$

Recall that it suffices to prove it for $s, t > 0$.

Since P_n converges weakly to P and using Lemma 3.1 it is enough to check that

$$E[\varphi(x_n(s_1, t_1), \dots, x_n(s_m, t_m))((\Delta_{s,t}x_n(s', t'))^2 - (s' - s)(t' - t))]$$

converges to zero when n tends to infinity. But this last expression is equal to

$$E[\varphi(x_n(s_1, t_1), \dots, x_n(s_m, t_m))(E[(\Delta_{s,t}x_n(s', t'))^2|\mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n] - (s' - s)(t' - t))].$$

Finally, in order to prove that this expression tends to zero, it suffices to show that,

$$E[(\Delta_{s,t}x_n(s', t'))^2|\mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n] \xrightarrow{L^2} (s' - s)(t' - t) \quad \text{as } n \rightarrow \infty. \tag{2}$$

This last convergence can be done using the following two facts.

Fact 1. $E[E[(\Delta_{s,t}x_n(s', t'))^2|\mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n]] = E[(\Delta_{s,t}x_n(s', t'))^2] \rightarrow (s' - s)(t' - t)$, as $n \rightarrow \infty$.

This result is proved in Lemma 4.4.

Fact 2. There exist some constants C_n converging to $(s' - s)^2(t' - t)^2$, when n goes to infinity, such that

$$E[E[(\Delta_{s,t}x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n]]^2 \leq C_n.$$

This result is proved in Lemma 4.5.

Facts 1 and 2 imply the convergence stated in (2) because if we assume that they are true,

$$\begin{aligned} 0 &\leq E[E[(\Delta_{s,t}x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n] - (s' - s)(t' - t)]^2 \\ &\leq C_n - 2(s' - s)(t' - t)E[(\Delta_{s,t}x_n(s', t'))^2] + (s' - s)^2(t' - t)^2, \end{aligned}$$

and the right-hand side of this expression obviously converges to 0. This finishes the proof of Proposition 4.3. □

Lemma 4.4. *In the previous situation*

$$\lim_{n \rightarrow \infty} E[(\Delta_{s,t}x_n(s', t'))^2] = (s' - s)(t' - t).$$

Proof. In the proof of Lemma 3.2 we have shown that

$$E[(\Delta_{s,t}x_n(s', t'))^2] = 2(I_1 + I_2),$$

where

$$\begin{aligned} I_1 &= n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 y_2 - x_1 y_1)] I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2, \\ I_2 &= n^2 \int_{[s,s']^2 \times [t,t']^2} \sqrt{x_1 x_2 y_1 y_2} \exp[-2n(x_2 - x_1)y_1 - 2n(y_2 - y_1)x_1] \\ &\quad \times I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 dx_2 dy_1 dy_2. \end{aligned}$$

We can write the integral I_2 as

$$\begin{aligned} I_2 &= \int_s^{s'} \frac{1}{2\sqrt{x_1}} \int_t^{t'} \frac{1}{2\sqrt{y_1}} \left[\int_{y_1}^{t'} (2nx_1 \exp[-2nx_1(y_2 - y_1)]) \sqrt{y_2} dy_2 \right. \\ &\quad \left. \times \int_{x_1}^{s'} (2ny_1 \exp[-2ny_1(x_2 - x_1)]) \sqrt{x_2} dx_2 \right] dy_1 dx_1. \end{aligned}$$

The last integral tends to $\sqrt{x_1}$ because $2ny_1 \exp[-2ny_1(x_2 - x_1)]$ is a probability density that gives an approximation of the identity as $n \rightarrow \infty$, and the penultimate tends to $\sqrt{y_1}$ as $n \rightarrow \infty$. The convergence is bounded, because the two integrals are bounded by $\sqrt{s'}$ and $\sqrt{t'}$ respectively and $1/\sqrt{x_1 y_1}$ is integrable, so, by the dominated convergence theorem, we obtain

$$\lim_{n \rightarrow \infty} I_2 = \frac{1}{4} \int_s^{s'} \int_t^{t'} dy_1 dx_1 = \frac{(s' - s)(t' - t)}{4}.$$

In the same way,

$$I_1 = \int_s^{s'} \frac{1}{2\sqrt{x_2}} \int_t^{t'} \frac{1}{2\sqrt{y_1}} \left[\int_{y_1}^{t'} (2nx_2 \exp[-2nx_2(y_2 - y_1)])\sqrt{y_2} dy_2 \right. \\ \left. \times \int_s^{x_2} (2ny_1 \exp[-2ny_1(x_2 - x_1)])\sqrt{x_1} dx_1 \right] dy_1 dx_2,$$

and we have

$$\lim_{n \rightarrow \infty} I_1 = \frac{(s' - s)(t' - t)}{4}.$$

□

Lemma 4.5. *In the previous situation there exist some constants C_n converging to $(s' - s)^2(t' - t)^2$ when n tends to infinity, such that*

$$E[E((\Delta_{s,t}x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n)]^2] \leq C_n.$$

Proof. We recall that it suffices to prove the case $s, t > 0$. Thus, we will suppose from now on that $s, t > 0$. By measurability and independence, we have that

$$E((\Delta_{s,t}x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n) \\ = n^2 \int_{[s,s']^2 \times [t,t']^2} (-1)^{\sum_{i=1}^2 (\Delta_{0,t}N_n(s, y_i) + \Delta_{s,0}N_n(x_i, t))} \times \prod_{i=1}^2 \sqrt{x_i y_i} E[(-1)^{\sum_{i=1}^2 \Delta_{s,t}N_n(x_i, y_i)}] dx_1 \dots dy_2;$$

then

$$E[E((\Delta_{s,t}x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n)]^2 \\ = E \left[n^4 \int_{[s,s']^4 \times [t,t']^4} (-1)^{\sum_{i=1}^4 (\Delta_{0,t}N_n(s, y_i) + \Delta_{s,0}N_n(x_i, t))} \right. \\ \left. \times \prod_{i=1}^4 \sqrt{x_i y_i} E[(-1)^{\sum_{i=1}^2 \Delta_{s,t}N_n(x_i, y_i)}] E[(-1)^{\sum_{i=3}^4 \Delta_{s,t}N_n(x_i, y_i)}] dx_1 \dots dy_4 \right].$$

By the arguments of the proof of Lemma 3.2, we have that

$$E[(-1)^{\sum_{i=1}^2 \Delta_{s,t}N_n(x_i, y_i)}] \leq \exp[-2n(|x_2 - x_1|(\min\{y_1, y_2\} - t) + |y_2 - y_1|(\min\{x_2, x_2\} - s))], \tag{3}$$

and that

$$E[(-1)^{\sum_{i=3}^4 \Delta_{s,t} N_n(x_i, y_i)}] \leq \exp[-2n(|x_4 - x_3|(\min\{y_3, y_4\} - t) + |y_4 - y_3|(\min\{x_3, x_4\} - s))]. \tag{4}$$

Then,

$$\begin{aligned} & E[E((\Delta_{s,t} x_n(s', t'))^2 | \mathcal{F}_{S,t}^n \vee \mathcal{F}_{s,T}^n)]^2 \\ & \leq 16n^4 \int_{[s,s']^4 \times [t,t']^4} |E(-1)^{\sum_{i=1}^4 (\Delta_{0,t} N_n(s, y_i) + \Delta_{s,0} N_n(x_i, t))}| \\ & \quad \times \prod_{i=1}^4 \sqrt{x_i y_i} \exp[-2n((x_2 - x_1)(y_1 - t)(y_2 - y_1)(x_1 - s))] \tag{5} \\ & \quad \times \exp[-2n((x_4 - x_3)(y_3 - t) + (y_4 - y_3)(x_3 - s))] \\ & \quad \times I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} I_{\{x_3 \leq x_4\}} I_{\{y_3 \leq y_4\}} dx_1 \dots dy_4. \end{aligned}$$

We can divide the last integral in two parts: the integral over $A = (\{y_1 \leq y_2 \leq y_3 \leq y_4\} \cup \{y_3 \leq y_4 \leq y_1 \leq y_2\}) \cap (\{x_1 \leq x_2 \leq x_3 \leq x_4\} \cup \{x_3 \leq x_4 \leq x_1 \leq x_2\})$; and the integral over A^c .

If we integrate over A ,

$$\begin{aligned} & |E(-1)^{\sum_{i=1}^4 (\Delta_{0,t} N_n(s, y_i) + \Delta_{s,0} N_n(x_i, t))}| \\ & = \exp[-2ns[(y_4 - y_3) + (y_2 - y_1)] - 2nt[(x_4 - x_3) + (x_2 - x_1)]], \end{aligned}$$

and the integral given in (5), over A , can be bounded by

$$16 \left(n^2 \int_{[s,s']^2 \times [t,t']^2} \prod_{i=1}^2 \sqrt{x_i y_i} \exp[-2n(x_2 - x_1)y_1 - 2n(y_2 - y_1)x_1] I_{\{x_1 \leq x_2\}} I_{\{y_1 \leq y_2\}} dx_1 \dots dy_2 \right)^2$$

which, as we have shown in Lemma 4.4, converges to

$$16 \left(\frac{(s' - s)(t' - t)}{4} \right)^2 = (s' - s)^2 (t' - t)^2.$$

When we integrate over A^c , the integral converges to zero. Indeed, if we have $y_1 \leq y_3 \leq y_2 \leq y_4$ (or $y_1 \leq y_3 \leq y_4 \leq y_2$) we can bound (3) by $\exp[-2n(y_2 - y_3)(x_1 - s)]$ (or $\exp[-2n(y_4 - y_3)(x_1 - s)]$) and (4) by 1.

If we have $y_3 \leq y_1 \leq y_4 \leq y_2$ (or $y_3 \leq y_1 \leq y_2 \leq y_4$) we likewise bound (4) by $\exp[-2n(y_4 - y_1)(x_3 - s)]$ (or $\exp[-2n(y_2 - y_1)(x_3 - s)]$) and (3) by 1.

Then, by doing a change of variables in order to obtain $x_1 \leq x_2 \leq x_3 \leq x_4$ and $y_1 \leq y_2 \leq y_3 \leq y_4$, we majorize the integral given in (5), over $(\{y_1 \leq y_2 \leq y_3 \leq y_4\} \cup \{y_3 \leq y_4 \leq y_1 \leq y_2\})^c$, by

$$Kn^4 \int_{[s,s']^4 \times [t,t']^4} \prod_{i=1}^4 \sqrt{x_i y_i} \exp[-2nt[(x_4 - x_3) + (x_2 - x_1)] - 2ns[(y_4 - y_3) + (y_2 - y_1)] - 2n(y_3 - y_2)(x_1 - s)] I_{\{x_1 \leq x_2 \leq x_3 \leq x_4\}} I_{\{y_1 \leq y_2 \leq y_3 \leq y_4\}} dx_1 \dots dy_4.$$

Finally, using the fact that $s, t > 0$, this integral equals to

$$K \frac{1}{s^2 t^2} \int_s^{s'} \sqrt{x_1} \int_s^{s'} \sqrt{x_3} \int_t^{t'} \sqrt{y_2} \int_{y_2}^{t'} \sqrt{y_3} \left\{ \int_{x_3}^{s'} 2nt \exp[-2nt(x_4 - x_3)] \sqrt{x_4} dx_4 \right. \\ \times \int_{x_1}^{s'} 2nt \exp[-2nt(x_2 - x_1)] \sqrt{x_2} dx_2 \int_{y_3}^{t'} 2ns \exp[-2ns(y_4 - y_3)] \sqrt{y_4} dy_4 \\ \left. \times \int_t^{y_2} 2ns \exp[-2ns(y_2 - y_1)] \sqrt{y_1} dy_1 \right\} \exp[-2n(y_3 - y_2)(x_1 - s)] dy_3 dy_2 dx_3 dx_1,$$

which goes to zero by dominated convergence.

Interchanging the roles of the variables x_i and the y_i for $i = 1, \dots, 4$, we will obtain a similar integral.

This completes the proof. □

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References

Bickel, P.J. and Wichura, M.J. (1971) Convergence criteria for multiparameter stochastic processes and some applications. *Ann. Math. Statist.*, **42**, 1656–1670.
 Cairoli, R. and Walsh, J.B. (1975) Stochastic integrals in the plane. *Acta Math.*, **134**, 111–183.
 Florit, C. and Nualart, D. (1996) Diffusion approximation for hyperbolic stochastic differential equations. *Stochastic Process. Appl.*, **65**, 1–15.
 Stroock, D. (1982) *Lectures on Topics in Stochastic Differential Equations*. Berlin: Springer-Verlag.
 Tudor, C. (1980) Remarks on the martingale problem in the two dimensional time parameter. *Rev. Roumaine Math. Pures Appl.*, **25**, 1551–1556.

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