ON CONSISTENCY IN TIME SERIES ANALYSIS

By P. M. Robinson¹

Harvard University

A number of statistics that arise in time series analysis can be represented as the sum of a partial realization of a possibly serially dependent and nonstationary discrete-parameter stochastic process. The almost sure and L_p , p>1, convergence of such statistics is investigated, under various moment conditions. The results are applied to the least squares estimates of multiple regressions.

1. Introduction. In time series analysis one is sometimes concerned with the partial sums $S_N = \sum_{n=1}^N \xi_n$, $N \ge 1$, of a realization of the discrete-parameter zero-mean stochastic process $\{\xi_n, n \ge 1\}$, and wishes to determine for which increasing sequences $\{s_N, N \ge 1\}$,

(1.1)
$$S_N/s_N \to 0$$
, almost surely (a.s.)

or

(1.2)
$$E|S_N/S_N|^{\nu} \to 0$$
, some ν ,

as $N \to \infty$. Both (1.1) and (1.2) imply that $S_N/s_N \to 0$ in probability, but neither (1.1) nor (1.2) imply the other.

Often $\{\xi_n, n \ge 1\}$ is the output generated by passing a set of observable or unobservable discrete-parameter time series through a filter which, possibly, is nonlinear and time-varying. Some important situations of this type might be described by the following model, which we refer to as

CONDITIONS A.

$$\xi_n = \sum_{j_1 = -\infty}^{\infty} \cdots \sum_{j_n = -\infty}^{\infty} \zeta_{j_n},$$

where ζ_{jn} is written in place of $\zeta_{j_1,\dots,j_{n}n}$,

(1.4)
$$E(\zeta_{jn} | \zeta_{jm}, m < n) = 0, \text{ a.s.},$$

all $-\infty < j_1, \dots, j_q < \infty$ and all $n \ge 2$, and for some $\nu > 1$ there exist nonnegative constants $\{\alpha_{kj}, -\infty < j < \infty, 1 \le k \le q\}, \{a_n, n \ge 1\}$ such that

$$(1.5) E|\zeta_{jn}|^{\nu} < \alpha_{j}^{\nu} a_{n}^{\nu}, \alpha_{j} = \prod_{k=1}^{q} \alpha_{kj_{k}},$$

$$\sum_{j=-\infty}^{\infty} \alpha_{kj} < \infty.$$

We take ν to be the largest value for which such relationships hold.

Received July 1975; revised March 1977.

¹ Part of this research was carried out while the author was at the University of California, Berkeley. I am grateful to the referee for comments on earlier versions. Support of NSF grant SOC75-13436 is acknowledged.

AMS 1970 subject classifications. Primary 60F15, 62M10; Secondary 60G45.

Key words and phrases. Almost sure convergence, L_p convergence, martingale inequalities, time series regression.

Conditions A generalize a familiar representation for a time series,

(1.7)
$$\xi_n = \sum_{j=-\infty}^{\infty} \gamma_j \chi_{n-j}, \qquad \sum_{j=-\infty}^{\infty} |\gamma_j| < \infty,$$

where $\{\chi_n, -\infty < n < \infty\}$ is a sequence of independent and identically distributed random variables with zero means and finite variances. Then ξ_n is strictly and second order stationary and ergodic. In Conditions A, stationarity is not implied by the homogeneity assumption (1.5); indeed, when $a_m \to \infty$ or when one can choose a_n that $\rightarrow 0$, evolutionary behavior is implied. As in Hannan and Heyde [5], independence is replaced by a martingale assumption, (1.4). In Conditions A, the summable α_{kj} are not introduced until one bounds $E|\zeta_{jn}|^{\nu}$. Cases $\nu < 2$ and $\nu > 2$, as well as $\nu = 2$, seem important, and the range of possible s_N depends on ν . The multiple summation in (1.3) is motivated primarily to handle cases in which ξ_n arises as the product of random variables, each of which may, for example, have a representation of the form (1.7). This is so when S_N is a regression sum of squares, or when S_N/N is the deviation from its expectation of a sample rth-order autocovariance of an rth-order stationary time series (see Robinson [8]). Also, a time-varying moving average of uncorrelated but heteroscedastic random variables may arise when certain continuous time stochastic processes are sampled at discrete but unequally spaced intervals (see Robinson

Define $b_{\nu 0} = 0$, $b_{\nu N} = \sum_{1}^{N} a_{n}^{\nu}$, N > 1. Throughout, K represents a positive constant, not necessarily the same one.

THEOREM 1. Let Conditions A hold with $1 < \nu \le 2$ and let

$$(1.8) b_{\nu[h^{k+1}]} < K b_{\nu[h^k]}, k \ge 1$$

for some h > 1. Then in (1.1) we may choose

(1.9)
$$s_N = (b_{\nu N} (\ln N)^{1+\phi+\nu} (\ln \ln N)^{1+\phi+\delta})^{1/\nu}$$

when $\phi + \nu > 0$, $\delta > 0$ or when $\phi + \nu = 0$, $\delta > \max(0, -\psi)$.

In most cases one would choose $\phi = \psi = 0$ in (1.9). However, when $a_n = n^{-\frac{1}{2}}$, $\nu = 2$, h = 2, one has $\phi = -1$, $\psi = 0$; when $a_n = (n \ln n)^{-\frac{1}{2}}$, $n \ge 2$, $\nu = 2$, h = 2, one has $\phi = 0$, $\psi = -1$.

Martingale convergence theorems can yield stronger results, under generally stronger conditions. From Loève [6], we may be able to choose $s_N = b_{\nu N}^{1/\nu}$ for $1 < \nu < 2$, and from Neveu [7] we may be able to choose $s_N = b_{2N}^{\frac{1}{2}}(\ln b_{2N})^{\frac{1}{2}+\delta}$, $\delta > 0$, for $\nu = 2$.

This approach is not easily extended to handle a representation of the generality of Conditions A, however, and it cannot be used to improve the results of Theorems 2 and 3 below.

For $\nu > 2$, define by τ the largest power of 2 such that $\tau = 2^t < \nu$ and

$$(1.10) E(\zeta_{jn}^{\tau} | \zeta_{jm}, m < n) < K\alpha_{j}^{\tau} a_{n}^{\tau}.$$

By virtue of (1.5), this is similar to $\{\zeta_{jn}, n \ge 1\}$ being an independent sequence, up to τ th moments.

THEOREM 2. Let Conditions A hold with $\nu > 2$ and let

$$(1.11) b_{2[h^{k+1}]} < Kb_{2[h^k]}, k > 1.$$

When $\nu > 2\tau$, let

$$(1.12) (N-M)^{\nu/2\tau-1}(b_{\nu N}-b_{\nu M}) < K(b_{2N}-b_{2M})^{\nu/2}, 1 \leq M \leq N.$$

Then in (1.1), we may choose

$$(1.13) s_N = b_{2N}^{\frac{1}{2}} (\ln N)^{1/\nu} (\ln \ln N)^{(1+\delta)/\nu} , any \delta > 0.$$

We next consider L_p -convergence. Lemmas 1 and 2 below (established to prove Theorems 1 and 2) imply that in (1.2), s_N may be chosen as

$$b_{\nu N}^{1/\nu} c_N$$
, $1 \le \nu \le 2$; $b_{2N}^{\frac{1}{2}} c_N$, $\nu > 2$,

where $\{c_n\}$ is any sequence such that $c_n \to \infty$ as $n \to \infty$. An improvement is possible when $1 < \nu < 2$.

CONDITIONS B. Define

$$\zeta'_{jn} = \zeta_{jn} I(|\zeta_{jn}|^{\nu} \leq CE|\zeta_{jn}|^{\nu}), \qquad \qquad \zeta''_{jn} = \zeta_{jn} - \zeta'_{jn}.$$

For any $\varepsilon > 0$ and all **j**, n, C can be chosen such that

$$E|\zeta_{in}^{\prime\prime}|^{\nu} < \varepsilon E|\zeta_{in}|^{\nu}$$
.

This is a type of uniform integrability condition.

THEOREM 3. Let Conditions A and B hold with $1 < \nu < 2$ and let

$$(1.14) b_{2N}^{\nu}/b_{\nu N}^2 \to 0 as N \to \infty.$$

Then in (1.2) we may choose

$$s_N = b_{\nu N}^{1/\nu}$$
.

Each of the conditions (1.8), (1.11), (1.12), (1.14) requires that a_n should not increase too fast, excluding $a_n = \theta^n$, $|\theta| > 1$, but not $a_n = n^{\theta}$.

The theorems are proved in Section 2. In Section 3, the results are applied to the convergence of least squares estimates of multiple time series regression.

2. Proofs of theorems. Our results rest heavily on calculations of the order of magnitude of the ν th absolute moment of $S_{MN} = S_N - S_M$, $0 \le M < N(S_0 = 0)$.

LEMMA 1. Let Conditions A hold, for $1 < \nu \le 2$. Then

$$E|S_{MN}|^{\nu} < K(b_{\nu N} - b_{\nu M})$$
.

Proof. First we prove that

$$|S_{\scriptscriptstyle MN}|^{\scriptscriptstyle \nu} < K \sum_{j_1}' \cdots \sum_{j_q}' \alpha_{\scriptscriptstyle \rm j}^{\scriptscriptstyle 1-\nu} |\sum_{\scriptscriptstyle n} \zeta_{\scriptscriptstyle jn}|^{\scriptscriptstyle \nu},$$

where the sum over n is for $M < n \le N$ and, for $1 \le k \le q$, the sum over j_k is over those j for which $\alpha_{kj} \ne 0$. To obtain (2.1), one first writes

$$S_{MN} = \sum_{j_1}' \cdots \sum_{j_q}' \alpha_{1j_1}^{1-1/\nu} \alpha_{1j_1}^{1/\nu} \sum_{n} (\zeta_{jn}/\alpha_{1j_1})$$

whence Hölder's inequality gives

$$|S_{MN}|^{\nu} < (\sum_{j} \alpha_{1j})^{\nu-1} \sum_{j_1}' \alpha_{1j_1}^{1-\nu} |\sum_{j_2}' \cdots \sum_{j_n}' \zeta_{j_n}|^{\nu}.$$

Similarly, one has

$$|\sum_{j_2}' \cdots \sum_{j_q}' \sum_n \zeta_{j_n}|^{\nu} < (\sum_j \alpha_{2j})^{\nu-1} \sum_{j_2}' \alpha_{2j_2}^{1-\nu} |\sum_{j_3}' \cdots \sum_{j_q}' \sum_n \zeta_{j_n}|^{\nu}.$$

After using Hölder's inequality a total of q times, (2.1) follows by virtue of (1.6). Then from Von Bahr and Esseen [13],

$$E|\sum_{n} \zeta_{jn}|^{\nu} \leq K \sum_{n} E|\zeta_{jn}|^{\nu}$$
,

because of (1.4). Then the lemma is established by using (1.5) and then (1.6). \square

LEMMA 2. Let Conditions A hold for $\nu > 2$ and, when $\nu > 2\tau$, let (1.12) hold. Then

$$E|S_{MN}|^{\nu} < K(b_{2N} - b_{2M})^{\nu/2}$$
.

PROOF. Because of (2.1) one has only to show that

$$(2.2) E|\sum_{n} \zeta_{jn}|^{\nu} < K\alpha_{j}(b_{2N} - b_{2M})^{\nu/2}.$$

From (1.4), Theorem 9 of Burkholder [3] and Minkowski's inequality,

$$E|\sum_{n} \zeta_{in}|^{\nu} < KE(\sum_{n} \zeta_{in}^{2})^{\nu/2} < K(N-M)^{\nu/2-1} \sum_{n} E|\zeta_{in}|^{\nu}$$
.

Thus when t = 0, (2.2) follows from use of (1.5). When t > 0, define

$$(2.3) \eta_{jn}^{(0)} = \zeta_{jn}, \eta_{jn}^{(l)} = \eta_{jn}^{(l-1)2} - \tilde{\eta}_{jn}^{(l)}, \tilde{\eta}_{jn}^{(l)} = E(\eta_{jn}^{(l-1)2} | \zeta_{j,n-1}, \dots, \zeta_{j1}),$$

for $1 \le l \le t$. It follows from Theorem 9 of [3] that, for all θ and some K depending only on θ ,

(2.4)
$$E|\sum_{n} \eta_{jn}^{(l)}|^{\theta} < KE(\sum_{n} \eta_{jn}^{(l)2})^{\theta/2}.$$

Repeated use of (2.3), (2.4) and the c_r -inequality produces

$$(2.5) E|\sum_{n} \zeta_{jn}|^{\nu} < KE(\sum_{n} \eta_{jn}^{(t)2})^{\nu/2\tau} + K \sum_{l=1}^{t} E(\sum_{n} \bar{\eta}_{jn}^{(l)})^{\nu/2l}.$$

Now by Jensen's inequality for conditional expectations and (1.10), (1.5),

$$E(\zeta_{j_n}^r | \zeta_{j_n-1}, \cdots, \zeta_{j_1}) < K\alpha_j^r a_n^r, \qquad r < \tau.$$

Then it is easily seen that

(2.6)
$$\tilde{\eta}_{jn}^{(l)} < K \alpha_j^{2l} a_n^{2l}, \qquad 1 \leq l \leq t.$$

Thus,

(2.7)
$$\sum_{l=1}^{t} E(\sum_{n} \tilde{\eta}_{jn}^{(l)})^{\nu/2^{l}} < K\alpha_{j}^{\nu} \sum_{l=1}^{t} (b_{2lN} - b_{2lM})^{\nu/2^{l}} < K\alpha_{i}^{\nu} (b_{2N} - b_{2M})^{\nu/2}.$$

Now on expanding $\eta_{in}^{(t)}$ and using (1.5), (2.6), it is seen that

$$(2.8) E\eta_{\mathbf{j}n}^{(t)\nu/\tau} < Ka_n^{\nu}\alpha_{\mathbf{j}}^{\nu}.$$

Thus

(2.9)
$$E(\sum_{n} \eta_{jn}^{(t)2})^{\nu/2\tau} < K(N-M)^{\nu/2\tau-1} \sum_{n} E \eta_{jn}^{(t)\nu/\tau} < K(N-M)^{\nu/2\tau-1} \alpha_{j}^{\nu} (b_{\nu N} - b_{\nu M}) < K \alpha_{j}^{\nu} (b_{2N} - b_{2M})^{\nu/2},$$

the last inequality following from (1.12). Therefore the desired bound for (2.5) is provided by (2.7) and (2.9). \square

PROOF OF THEOREM 1. The method of proof is similar to one used by Stout [12]. Write $\tilde{S}_k = S_{\lfloor h^k \rfloor}$, $\tilde{s}_k = s_{\lfloor h^k \rfloor}$. There are finitely many k for which $h^k < 2$. For k > 1 such that $h^k \ge 2$, and with s_N given by (1.9),

$$\begin{split} P(|\tilde{S}_k| > \varepsilon \tilde{s}_k) & \leq E|\tilde{S}_k|^{\nu} \varepsilon^{-\nu} \tilde{s}_k^{-\nu} \\ & \leq K \varepsilon^{-\nu} (\ln \left[h^k\right])^{-1-\phi-\nu} (\ln \ln \left[h^k\right])^{-1-\phi-\delta} \\ & < K \varepsilon^{-\nu} k^{-1-\phi-\nu} (\ln k)^{-1-\phi-\delta} , \end{split}$$

from Markov's inequality, Lemma 1 and $[h^k] \ge Kh^k$. Thus, when $\phi + \nu > 0$ or when $\phi + \nu = 0$, $\phi + \delta > 0$ it follows from the Borel-Cantelli lemma that $\tilde{S}_k/\tilde{s}_k \to 0$, a.s. Now define $T_k = \max_{[h^k] < n \le [h^{k+1}]} |S_{[h^k]_n}|$. We need consider only k such that $[h^{k+1}] - [h^k] \ge 1$. From Billingsley [2, page 102], Lemma 1 implies

(2.10)
$$ET_{k}^{\nu} < K[\ln([h^{k+1}] - [h^{k}])]^{\nu}[b_{\nu[h^{k+1}]} - b_{\nu[h^{k}]}] < Kk^{\nu}b_{\nu[h^{k}]}$$

using (1.8) and $[h^{k+1}] - [h^k] \leq h^{k+1} - h^k + 1 \leq Kh^k$. Thus, much as before

$$P(T_k > \tilde{s}_k) < ET_k^{\nu} \varepsilon^{-\nu} \tilde{s}_k^{-\nu} < K \varepsilon^{-\nu} k^{-1} (\ln k)^{-1-\delta}.$$

When $\delta > 0$, it follows that $T_k/\tilde{s}_k \to 0$, a.s. Finally, for $[h^k] \leq N \leq [h^{k+1}]$,

$$S_{\scriptscriptstyle N}/s_{\scriptscriptstyle N} \leqq |S_{\scriptscriptstyle N} - \tilde{S}_{\scriptscriptstyle k}|/\bar{s}_{\scriptscriptstyle k} + |\tilde{S}_{\scriptscriptstyle k}|/\bar{s}_{\scriptscriptstyle k} \leqq T_{\scriptscriptstyle k}/\bar{s}_{\scriptscriptstyle k} + |\tilde{S}_{\scriptscriptstyle k}|/\bar{s}_{\scriptscriptstyle k} \;. \label{eq:sn_N}$$

PROOF OF THEOREM 2. The proof is almost the same. For k > 1 such that $h^k \ge 2$ and with s_N given by (1.13), it is readily shown that

$$P(|\tilde{S}_k| > \varepsilon \bar{s}_k) < K \varepsilon^{-\nu} k^{-1} (\ln k)^{-1-\delta}$$
,

so $\tilde{S}_k/\tilde{s}_k \to 0$, a.s., for $\delta > 0$. Now from [2, page 94], Lemma 2 implies

(2.11)
$$P(T_k > \varepsilon) < K\varepsilon^{-\nu} (b_{2[h^{k+1}]} - b_{2[h^k]})^{\nu/2}.$$

Therefore, using (1.11), one has

$$P(T_k > \varepsilon \tilde{s}_k) < K \varepsilon^{-\nu} k^{-1} (\ln k)^{-1-\delta}$$
,

and so $T_k/s_k \to 0$, a.s., and the proof is completed as before. \square

Serfling [10] proves inequalities similar to (2.10) and (2.11) for $\nu \ge 2$. For $\nu > 2$, his assumed bound for $E|S_{MN}|^{\nu}$ (and derived bound for $E\max_{M < n \le N} |S_{Mn}|^{\nu}$), depend on M and N only through N-M, and he requires $2b_{2N} \le b_{2,2N}$. These assumptions exclude many cases where $a_n \to \infty$ or 0. Under Serfling's assumptions, there is no loss in generality in taking h=2 in the above proof (see Stout [12, page 210]).

Serfling [11] lists several alternative $\{\xi_n\}$ for which results similar to those of Lemma 2 and Theorem 2 are available.

PROOF OF THEOREM 3. Because of (2.1) we need only show that for any $\varepsilon > 0$ which does not depend on j,

$$(2.12) E|\sum_{n=1}^{N} \zeta_{in}|^{\nu} < \varepsilon \alpha_{i}^{\nu} b_{\nu N}$$

for sufficiently large N. Our method of proof then extends one given by Chow [4]. From [3, Theorem 9] and Conditions B,

$$E|\sum_{n} \zeta_{jn}|^{\nu} < KE(\sum_{n} \zeta_{jn}^{2})^{\nu/2} < EK(\sum_{n} \zeta_{jn}^{\prime 2})^{\nu/2} + KE(\sum_{n} \zeta_{jn}^{\prime \prime 2})^{\nu/2}$$

$$\leq KE(C^{2} \sum_{n} E|\zeta_{jn}^{\prime}|^{2})^{\nu/2} + K \sum_{n} E|\zeta_{jn}^{\prime \prime}|^{\nu}$$

$$< \alpha_{1} b_{\nu N} (Kb_{2N}^{\nu/2}/b_{\nu N} + K\varepsilon)$$

so (2.12) follows from (1.14). \square

3. Multiple regression. The preceding results may be applied to estimates of the model

$$y_n = \beta_1 z_{1n} + \cdots + \beta_p z_{pn} + x_n$$
, $n \ge 1$.

Define

(3.1)
$$\boldsymbol{\beta} = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_p \end{bmatrix}, \quad \mathbf{y}_N = \begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix}, \quad \mathbf{x}_N = \begin{bmatrix} x_1 \\ \vdots \\ x_N \end{bmatrix}, \quad \mathbf{Z}_N = \begin{bmatrix} z_{11} & \cdots & z_{p1} \\ \vdots & & \vdots \\ z_{1N} & \cdots & z_{pN} \end{bmatrix},$$
$$\hat{\boldsymbol{\beta}}_N = (\hat{\boldsymbol{\beta}}_{iN}) = (\mathbf{Z}_N^T \mathbf{Z}_N)^{-1} \mathbf{Z}_N^T \mathbf{y}_N = \boldsymbol{\beta} + (\mathbf{Z}_N^T \mathbf{Z}_N)^{-1} \mathbf{Z}_N^T \mathbf{x}_N$$

which is the least squares estimate of β . A variety of results may be generated from Theorems 1-3. Three examples will be given, under somewhat simplified conditions.

The first assumes the z_{in} are random variables with representations

$$Z_{in} = \sum_{j=-\infty}^{\infty} u_{ijn}, \qquad n \ge 1, \ 1 \le i \le p,$$

where there exist real constants α_{ij} , a_{in} , a_{n} such that $\sum_{i=-\infty}^{\infty} |\alpha_{ij}| < \infty$,

(3.2)
$$E(u_{ijn} | u_{ijm}, m < n; x_m, m \le n) = 0, \text{ a.s.},$$

$$(3.3) E(u_{hj_1n}u_{ij_2n}|u_{hj_1m},u_{hj_2m},m< n) = \alpha_{hj}\alpha_{ij}a_{hn}a_{in}, a.s., j_1=j_2=j,$$

$$(3.4) j_1 \neq j_2,$$

(3.5)
$$E(|u_{ijn}|^{\mu}|u_{hjm}, m < n, 1 \le h \le p; x_m, m \le n) < K|\alpha_{ij}a_{in}|^{\mu}$$

$$(3.6) E|x_n|^{\nu} < |a_n|^{\nu}.$$

These are Conditions C.

Define Ω_N as the $p \times p$ matrix with (h, i)th element $\sum_{n=1}^N \omega_{hin}$, where

$$\omega_{hin} = E z_{hn} z_{in} = \left(\sum_{j=-\infty}^{\infty} \alpha_{hj} \alpha_{ij}\right) a_{hn} a_{in}$$

and define

(3.7)
$$b_{\nu iN} = \sum_{n=1}^{N} |a_{in}|^{\nu}, \quad \mathbf{B}_{N} = \operatorname{diag}\{b_{21N}, \dots, b_{2pN}\}.$$

THEOREM 4. Let Conditions C hold with $\mu > 4$, $\nu > 2$, $\mu \ge \nu$, and let

(3.8)
$$\lim_{N\to\infty} \det \left\{ \mathbf{B}_N^{-\frac{1}{2}} \mathbf{\Omega}_N \mathbf{B}_N^{-\frac{1}{2}} \right\} \neq 0,$$

$$(3.9) b_{2,h,2N} < Kb_{2hN}, 1 \le h \le p, N \ge 1,$$

(3.10)
$$\limsup_{N\to\infty} (\ln N)^{4/\mu} (\ln \ln N)^{(4+\delta)/\mu} \max_{1\leq n\leq 2N} a_{hn}^2/b_{2hN} < \infty$$
,

$$1 \leq h \leq p, \ \delta > 0$$

(3.11) $\limsup_{N\to\infty} (\ln N)^{2/\nu} (\ln \ln N)^{(2+\delta)/\nu} \max_{1\leq n\leq 2N} a_n^2/b_{2iN} < \infty$, $\delta > 0$. Then $\hat{\beta}_{iN} \to \hat{\beta}_i$, a.s. as $N\to\infty$.

PROOF. It is readily deduced from (3.1) that

$$(3.12) |\hat{\beta}_{iN} - \beta_i| \leq ||\mathbf{B}_N^{\frac{1}{2}}(\mathbf{Z}_N^T \mathbf{Z}_N)^{-1} \mathbf{B}_N^{\frac{1}{2}}|| \sum_{h=1}^p b_{2hN}^{-\frac{1}{2}} b_{2iN}^{-\frac{1}{2}} |\sum_{n=1}^N z_{hN} x_n|,$$

where $||\cdot||$ is the Euclidean norm. From (3.8) it suffices to prove

$$\mathbf{B}_{N}^{-\frac{1}{2}}(\mathbf{Z}_{N}^{T}\mathbf{Z}_{N}-\mathbf{\Omega}_{N})\mathbf{B}_{N}^{-\frac{1}{2}}\rightarrow\mathbf{0},\quad \text{a.s.},$$

(3.14)
$$b_{2hN}^{-\frac{1}{2}} b_{2iN}^{-\frac{1}{2}} \sum_{n=1}^{N} z_{hn} x_n \to 0, \quad \text{a.s.}, \qquad 1 \le h \le p.$$

Consider (3.13) first. Denote the (k, l)th element of $\mathbf{B}_N^{-\frac{1}{2}}(\mathbf{Z}_N^T\mathbf{Z}_N)^{-1}\mathbf{B}_N^{-\frac{1}{2}}$ by $b_{2kN}^{-\frac{1}{2}}b_{2lN}^{-\frac{1}{2}}\sum_{1}^{N}\xi_n$, where $\xi_n=\sum_{j_1,j_2=-\infty}^{\infty}\zeta_{j_n}$,

$$\zeta_{jn} = \zeta_{j_1 j_2 n} = u_{k j_n} u_{l j_n} - a_{k n} a_{l n} \alpha_{k j} \alpha_{l j},$$
 $j_1 = j_2 = j,$
 $= u_{k j_1 n} u_{l j_2 n},$ $j_1 \neq j_2.$

Then (1.4) is satisfied because of (3.3), (3.4), and (1.5) is satisfied with

$$\begin{split} E|\zeta_{\mathfrak{j}n}|^{\mu/2} &< K(E|u_{kj_1n}|^{\mu}E|u_{lj_2n}|^{\mu})^{\frac{1}{2}} + K|a_{kn}a_{ln}\alpha_{kj_1}\alpha_{lj_1}|^{\mu/2} \\ &< K|a_{kn}a_{ln}\alpha_{kj_1}\alpha_{lj_2}|^{\mu/2} \;, \end{split}$$

because of (3.3), (3.5). Since $\sum_{j} |\alpha_{lj}| < \infty$, $1 \le l \le p$, all of Conditions A are satisfied. Next, note that, because of (3.5), the result $E|\sum \xi_n|^{\mu/2} < K(\sum a_{kn}^2 a_{ln}^2)^{\mu/4}$ does not require a condition like (1.12). With regard to (1.11) note that

$$\textstyle \sum_{1}^{N} a_{kn}^{2} a_{ln}^{2} \leqq \max_{1 \leqq n \leqq 2N} a_{kn}^{2} b_{2lN} \,, \qquad \sum_{N+1}^{2N} a_{kn}^{2} a_{ln}^{2} \leqq \max_{1 \leqq n \leqq 2N} a_{kn}^{2} b_{2lN} \,,$$

using (3.9) in the second case. Therefore, from Theorem 2,

$$\textstyle \sum_{1}^{N} (z_{kn} z_{ln} - \omega_{kln}) / (\max_{1 \leq n \leq 2N} |a_{kn}| b_{2lN}^{\frac{1}{2}} (\ln N)^{2/\mu} (\ln \ln N)^{(2+\delta)/\mu}) \to 0 \;, \quad \text{a.s.}$$

Then (3.12) is a consequence of (3.10). To establish (3.14) note that $\sum_{1}^{N} z_{hn} x_{n}$ is of the form $\sum_{1}^{N} \hat{\xi}_{n}$, $\hat{\xi}_{n} = \sum_{-\infty}^{\infty} \zeta_{jn}$, where $\zeta_{jn} = u_{hjn} x_{n}$. Then Conditions A are satisfied because of (3.2), (3.5), (3.6), with $E|\zeta_{jn}|^{\nu} < |\alpha_{hj} a_{hn} a_{n}|^{\nu}$, $\nu \leq \mu$. Note now that $E|\sum_{n} \hat{\xi}_{n}|^{\nu} < K(\sum_{n} a_{n}^{2} a_{hn}^{2})^{\nu/2}$ does not require a condition like (1.13), because of (3.5), (3.6) and $\mu \geq \nu$. Note also that

$$\sum_{1}^{N} a_n^2 a_{hn}^2 \le \max_{1 \le n \le 2N} a_n^2 b_{2hN}$$
, $\sum_{N+1}^{2N} a_n^2 a_{hn}^2 \le \max_{1 \le n \le 2N} a_n^2 b_{2hN}$.

Therefore, from Theorem 2,

$$\sum_{1}^{N} z_{hn} x_{n} / (\max_{1 \le n \le 2N} |a_{n}| b_{2hN}^{\frac{1}{2}} (\ln N)^{1/\nu} (\ln \ln N)^{(1+\delta)/\nu}) \to 0, \quad \text{a.s.},$$

and so (3.14) follows from (3.11). \square

An example in which the regressors are "decreasing" while the residuals are "increasing" but Theorem 4 applies is when $a_{nn} = n^{-\frac{1}{2}}$, $a_n = \ln \ln n$.

Notice that some of the weak dependence assumptions on the z_{hn} , introduced to establish (3.13), are used to prove (3.14) without any weak dependence assumptions on x_n . By way of contrast, Anderson and Taylor [1, Theorem 2], for the case p = 1, assume x_n is a martingale difference sequence and assume the a.s. divergency of $\sum_{1}^{N} z_n^2$, with no restrictions on its rate of increase.

Now let the z_{hn} be fixed constants, and let $x_n = \sum_{j=-\infty}^{\infty} v_{jn}$, where α_j , a_n are such that $\sum_{j=-\infty}^{\infty} |\alpha_j| < \infty$, $E(v_{jn} | v_{jm}, m < n) = 0$, a.s., $E(|v_{jn}|^{\nu} | v_{jm}, m < n) < |\alpha_j a_n|^{\nu}$. For $\nu > 2$, let

(3.15)
$$E(|v_{jn}|^{\nu} | v_{jm}, m < n) < K |\alpha_j a_n|^{\nu}.$$

These are Conditions D. We replace a_{hn} by z_{hn} in the definitions of $b_{\nu hN}$ and \mathbf{B}_N in (3.7).

THEOREM 5. Let Conditions D, (3.7) and (3.9) hold. For $1 < \nu \le 2$, let

$$\limsup_{N \to \infty} (\ln N)^{(2+2\nu)/\nu} (\ln \ln N)^{(2+\delta)/\nu} \max_{1 \le n \le 2N} a_n^2/b_{2iN} < \infty$$

and for $\nu > 2$ let (3.11) hold. Then $\hat{\beta}_{iN} \to \beta_i$, a.s. as $N \to \infty$.

The proof applies Theorems 1 and 2 in a way that is similar to that of Theorem 4, so it is omitted. Anderson and Taylor [1, Theorem 1] proved that if the x_n are i.i.d. normal variables, then a necessary and sufficient condition for $\hat{\boldsymbol{\beta}}_N \to \boldsymbol{\beta}$ a.s. is $(\mathbf{Z}_N^T \mathbf{Z}_N)^{-1} \to 0$. Our restriction (3.11) is relaxed as ν increases, however. Finally, we give an L_r -convergence theorem.

THEOREM 6. Let Conditions D hold. In case $1 < \nu < 2$, let

(3.16)
$$\lim_{N\to\infty} \max_{1\leq n\leq N} a_n^2 b_{\nu h N}^{2/\nu} / (b_{2hN} b_{2iN}) = 0, \qquad 1\leq h\leq p,$$

or let the $|v_{in}/\alpha_i a_n|^{\nu}$ be uniformly integrable and

(3.17)
$$\lim \sup_{N \to \infty} \max_{1 \le n \le N} a_n^2 b_{\nu h N}^{2/\nu} / (b_{2hN} b_{2iN}) < \infty , \qquad 1 \le h \le p .$$

In case $\nu \geq 2$, let

(3.18)
$$\lim_{N\to\infty} \max_{1\leq n\leq N} a_n^2/b_{2iN} = 0.$$

Then, for $\nu > 1$, $E|\hat{\beta}_{iN} - \beta_i|^{\nu} \to 0$ as $N \to \infty$.

PROOF. We use (3.12) and thus have to investigate

$$(3.19) E|\sum_{n=1}^{N} Z_{hn} X_n|^{\nu}.$$

From Lemma 1, this is bounded by

$$(3.20) K \sum_{n=1}^{N} |a_n a_{hn}|^{\nu} \leq K \max_{1 \leq n \leq N} |a_n|^{\nu} b_{\nu hN}, 1 < \nu < 2.$$

Then the theorem is proved under (3.16). From Theorem 3, (3.19) is of smaller order than (3.20) as $N \to \infty$ under uniform integrability of $|v_{jn}/\alpha_j a_n|^{\nu}$ and (3.17), so that theorem is proved under these conditions. From (3.15) and Lemma 2,

(3.19) is bounded by

$$K(\sum_{n=1}^{N} a_n^2 a_{hn}^2)^{\nu/2} < K \max_{1 \le n \le N} |a_n|^{\nu} b_{2hN}^{\nu/2}$$
.

Then because of (3.18), the theorem is proved for $\nu \geq 2$. \square

REFERENCES

- [1] Anderson, T. W. and Taylor, J. B. (1976). Strong consistency of least squares estimates in normal linear regression. *Ann. Statist.* 4 788-790.
- [2] BILLINGSLEY, P. (1968). Convergence of Probability Measures. Wiley, New York.
- [3] BURKHOLDER, D. L. (1966). Martingale transforms. Ann. Math. Statist. 37 1494-1504.
- [4] CHOW, Y. S. (1971). On the L_p -convergence for $n^{-1/p}S_n$, 0 . Ann. Math. Statist. 42 393-394.
- [5] HANNAN, E. J. and HEYDE, C. C. (1972). On limit theorems for quadratic functions of discrete time series. Ann. Math. Statist. 43 2058-2066.
- [6] Loève, M. (1951). On almost sure convergence. *Proc. Second Berkeley Symp. Math. Statist. Prob.* 1 279-304, Univ. of California Press.
- [7] NEVEU, J. (1965). Mathematical Foundations of the Calculus of Probability. Holden-Day, San Francisco.
- [8] Robinson, P. M. (1977). Distributed lag approximation to linear time invariant systems. Ann. Statist., forthcoming.
- [9] ROBINSON, P. M. (1977). Estimation of a time series model from unequally spaced data. Stoch. Proc. Appl., forthcoming.
- [10] SERFLING, R. J. (1970). Moment inequalities for the maximum cumulative sum. Ann. Math. Statist. 41 1227-1234.
- [11] Serfling, R. J. (1970). Convergence properties of S_N under moment restrictions. *Ann. Math. Statist.* 41 1235-1248.
- [12] STOUT, W. F. (1974). Almost Sure Convergence. Academic Press, New York.
- [13] Von Bahr, B. and Esseen, C.-G. (1965). Inequalities for the rth absolute moment of a sum of random variables $1 \le r \le 2$. Ann. Math. Statist. 36 299-303.

DEPARTMENT OF STATISTICS
HARVARD UNIVERSITY
CAMBRIDGE, MASSACHUSETTS 02138