

ESTIMATION OF THE k th DERIVATIVE OF A DISTRIBUTION FUNCTION¹

BY CARL MALTZ

California State University, Long Beach

Estimation of the k th derivative of a df by means of the k th-order difference quotients of the empiric df is investigated. In particular, consistency conditions are given, the asymptotic bias, variance, and mean-squared error of the estimator are computed, and means of minimizing the latter are discussed.

1. Introduction. Let X_1, X_2, \dots, X_n be a random sample distributed according to a df F . Suppose that F possesses a k th derivative $F^{(k)}$ at a point x . In this paper we discuss the estimation of $F^{(k)}(x)$ through use of the difference quotient

$$(1) \quad F_n^{(k)}(x; h_n) = (2h_n)^{-k} \sum_{j=0}^k (-1)^j F_n(\bar{x}_j) \binom{k}{j},$$

where $\bar{x}_j = x + (k - 2j)h_n$. F_n denotes the empiric df based on X_1, \dots, X_n and $\{h_n\}$ is a suitably chosen sequence of positive numbers converging to zero. When there is no danger of confusion we will omit the subscript on h . We assume throughout that $k \geq 1$.

We investigate the consistency, asymptotic bias, variance, and mean square error of this estimator, and discuss the minimization of the latter through judicious choice of the sequence $\{h_n\}$. This generalizes results of Rosenblatt (1956) who treated the case $k = 1$. Gaffey (1959) made use of the estimator (1) and essentially proved Theorem 2 of this paper. Schuster (1969) considered a different estimator for $F^{(k)}(x)$ for which he proved a.s. uniform convergence subject to certain regularity conditions.

2. Asymptotic bias, variance and mean square error.

THEOREM 1. *Assume that F' exists at x . Then*

$$(2) \quad \text{Var} (F_n^{(k)}(x; h)) = n^{-1}(2h)^{1-2k} F'(x) \binom{2k-2}{k-1} + o(n^{-1}h^{1-2k}).$$

PROOF. We have

$$(3) \quad \begin{aligned} (2h)^{2k} \text{Var} (F_n^{(k)}(x; h)) &= \text{Var} (\sum_{j=0}^k (-1)^j F_n(\bar{x}_j) \binom{k}{j}) \\ &= \sum_{i,j=0}^k (-1)^{i+j} \binom{k}{i} \binom{k}{j} \text{Cov} (F_n(\bar{x}_i), F_n(\bar{x}_j)) \\ &= \frac{1}{n} \sum_{i,j=0}^k (-1)^{i+j} \binom{k}{i} \binom{k}{j} [F(\bar{x}_j \wedge \bar{x}_i) - F(\bar{x}_i)F(\bar{x}_j)], \end{aligned}$$

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where $\bar{x}_i \wedge \bar{x}_j$ denotes the minimum of \bar{x}_i and \bar{x}_j . By the definition of the derivative,

$$F(\bar{x}_j) = F(x + (k - 2j)h) = F(x) + (k - 2j)hF'(x) + o(h)$$

as $h \rightarrow 0$. Inserting this, together with the analogous formula for $F(\bar{x}_i)$ in (3) we obtain:

$$(4) \quad \begin{aligned} &(2h)^{2k} \text{Var} (F_n^{(k)}(x; h)) \\ &= \frac{1}{n} \sum_{i,j=0}^k (-1)^{i+j} \binom{k}{i} \binom{k}{j} \cdot [F(x) + \{(k - 2j) \wedge (k - 2i)\}F'(x)h \\ &\quad - F^2(x) - h(2k - 2j - 2i)F(x)F'(x) + o(h)]. \end{aligned}$$

Using the equations $\sum_{i=0}^k (-1)^i \binom{k}{i} = \sum_{i=0}^k (-1)^i \binom{k}{i} i = 0$ it is easily seen that (4) reduces to

$$(5) \quad \begin{aligned} &(2h)^{2k} \text{Var} (F_n^{(k)}(x; h)) \\ &= -2F'(x)hn^{-1} \sum_{i,j=0}^k (-1)^{i+j} \binom{k}{i} \binom{k}{j} (i \vee j) + o(n^{-1}h) \end{aligned}$$

where $i \vee j$ denotes the larger of i and j . The sum in (5) is equal to the constant term in the expansion in powers of z of the function

$$z(1 - z^{-1})^k \frac{d}{dz} [(1 - z)^k] (2 \sum_{m=0}^{\infty} z^{-m} - 1) = (-1)^k k(z + 1)(1 - z)^{2k-2} z^{1-k}.$$

The binomial theorem yields the value $-(\frac{2k-2}{k-1})$, and (2) follows immediately.

THEOREM 2.

$$\begin{aligned} EF_n^{(k)}(x; h) &= F^{(k)}(x) + o(1) \\ &= F^{(k)}(x) + \frac{k}{6} h^2 F^{(k+2)}(x) + o(h^2) \end{aligned}$$

provided $F^{(k)}$ (or $F^{(k+2)}$) exists at x .

PROOF. The theorem follows immediately from (1) and the equations

$$(6) \quad EF_n(\bar{x}_j) = \sum_{r=0}^m \frac{h^r F^{(r)}(x)(k - 2j)^r}{r!} + o(h^m), \quad m = k, k + 2$$

and

$$(7) \quad \begin{aligned} &= 0 \quad r = 0, 1, \dots, k - 1, k + 1 \\ &\sum_{j=0}^k (-1)^j \binom{k}{j} (k - 2j)^r = 2^k k! \quad r = k \\ &= 2^k k! \binom{k+2}{3}! \quad r = k + 2. \end{aligned}$$

The following are immediate consequences of Theorems 1 and 2:

COROLLARY 1. *If $F^{(k)}$ exists at x , and the sequence $\{h_n\}$ satisfies the conditions $h_n \rightarrow 0$ and $n(h_n)^{2k-1} \rightarrow \infty$ as $n \rightarrow \infty$, then $F_n^{(k)}(x; h_n) \rightarrow_P F^{(k)}(x)$.*

COROLLARY 2. *If $F^{(k+2)}$ exists at x , then the mean square error (MSE) of $F_n^{(k)}(x; h)$*

is given by:

$$(8) \quad \text{MSE} (F_n^{(k)}(x; h)) = \frac{k^2 h^4}{36} (F^{(k+2)}(x))^2 + n^{-1} (2h)^{1-2k} F'(x) \binom{2k-2}{k-1} \\ + o(h^4 + n^{-1} (2h)^{1-2k}).$$

3. Minimization of asymptotic mean square error. Let us now consider a sequence h_n of the form $Kn^{-\alpha}$. It is clear that the asymptotically optimal choice of α is such that the first two terms on the right in (8) are of the same order. That is, choose α so that

$$n^{-4\alpha} = n^{-1+\alpha(2k-1)}, \quad \text{or} \quad \alpha = (2k+3)^{-1}.$$

With this choice we have

$$(9) \quad \text{MSE} (F_n^{(k)}(x; h_n)) \\ = \left[\frac{k^2}{36} K^4 (F^{(k+2)}(x))^2 + (2K)^{1-2k} F'(x) \binom{2k-2}{k-1} \right] n^{-(4/(2k+3))} + o(n^{-(4/(2k+3))}).$$

The optimal choice of K is the one minimizing the right side of (9). Assuming that $F^{(k+2)}(x)$ and $F'(x)$ are not zero, this optimal choice K_0 is easily found to be:

$$(10) \quad K_0 = \left[\frac{9 \binom{2k}{k} F'(x)}{k 2^{2k} \{F^{(k+2)}(x)\}^2} \right]^{1/(2k+3)}.$$

Unfortunately, K_0 depends on the unknown values $F'(x)$ and $F^{(k+2)}(x)$. It seems reasonable, therefore, to replace $F'(x)$ and $F^{(k+2)}(x)$ in (10) by the estimates $F_n^{(1)}(x; h_n^{(1)})$ and $F_n^{(k+2)}(x; \bar{h}_n)$ where $h_n^{(1)} = n^{-\frac{1}{2}}$ and $\bar{h}_n = n^{-(1/(2k+7))}$, if both estimates are $\neq 0$. Otherwise set $K_n = 1$.

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DEPARTMENT OF MATHEMATICS
CALIFORNIA STATE UNIVERSITY
LONG BEACH, CALIFORNIA 90815