

## STRONG LAWS FOR QUANTILES CORRESPONDING TO MOVING BLOCKS OF RANDOM VARIABLES

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Let  $U_1, U_2, \dots$  be a sequence of independent uniform  $(0, 1)$  random variables, and for  $1 \leq k \leq n$  let  $\xi_p(n, k)$  denote the  $p$ th quantile,  $0 < p < 1$ , corresponding to the block  $U_{n-k+1}, \dots, U_n$ . In this paper we investigate the a.s. limiting behavior of  $\xi_p(n, a_n)$  when  $a_n$  is an integer sequence,  $1 \leq a_n \leq n$ , and  $\lim_{n \rightarrow \infty} a_n / \log n = \beta \in [0, \infty]$ . In addition, we investigate the a.s. limiting behavior of  $\max_{a_n \leq k \leq n} \xi_p(n, k)$  and other maxima involving the  $\xi_p(n, k)$ 's.

**1. Introduction.** Let  $U_1, U_2, \dots$  be a sequence of independent random variables, uniformly distributed on the unit interval, and for  $1 \leq i \leq k \leq n$  let  $U_{nk}(i)$  denote the  $i$ th order statistic among  $U_{n-k+1}, U_{n-k+2}, \dots, U_n$  (the block with right endpoint  $n$  and length  $k$ ). For fixed  $0 < p < 1$  and  $0 \leq \pi \leq 1$  define the corresponding  $p$ th quantile

$$(1.1) \quad \xi_p(n, k) = \begin{cases} \pi U_{nk}(pk) + (1 - \pi) U_{nk}(pk + 1), & \text{if } pk = [pk], \\ U_{nk}([pk] + 1), & \text{if } pk \neq [pk], \end{cases}$$

where  $[y]$  denotes the integer part of  $y$ . If  $a_n$  is an integer sequence satisfying  $1 \leq a_n \leq n$  for  $n \geq 1$ , then

$$(1.2) \quad \lim_{n \rightarrow \infty} \xi_p(n, a_n) = p \quad \text{a.s.}$$

when  $a_n \equiv n$ , and  $\xi_p(n, a_n) = U_n$  diverges (“oscillates” over  $[0, 1]$ ) almost surely when  $a_n \equiv 1$ . The condition “ $a_n \rightarrow \infty$ ” is necessary for (1.2) to hold, but it is not sufficient. How fast must  $a_n$  grow for (1.2) to hold? When (1.2) holds, what is the rate of convergence? When (1.2) fails, how does  $\xi_p(n, a_n)$  behave? In this paper we answer these questions for the case where  $\lim_{n \rightarrow \infty} a_n / \log n = \beta$ ,  $0 \leq \beta \leq \infty$ . In addition, we investigate the almost sure limiting behavior of  $\max_{a_n \leq k \leq n} \xi_p(n, k)$  and other maxima involving the  $\xi_p(n, k)$ 's.

**2. Results.** In Theorem 2.1 we assume  $\beta = \infty$ . We associate with each difference,  $\xi_p(n, k) - p$ , the normalizing constant

$$\phi(n, k) = k^{1/2} \left\{ 2p(1 - p) \left( \log \frac{n}{k} + \log \log k \right) \right\}^{-1/2}, \quad 3 \leq k \leq n,$$

and consider  $\alpha_{n,0} = (\xi_p(n, a_n) - p)\phi(n, a_n)$ , the maxima

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$$\begin{aligned}\alpha_{n,1} &= \max_{a_n \leq k \leq n} (\xi_p(n, k) - p) \phi(n, k), \\ \alpha_{n,2} &= \max_{0 \leq k \leq n - a_n} (\xi_p(k + a_n, a_n) - p) \phi(k + a_n, a_n), \\ \alpha_{n,3} &= \max_{\substack{0 \leq j < k \leq n \\ a_n \leq k-j}} (\xi_p(k, k-j) - p) \phi(k, k-j),\end{aligned}$$

and  $\alpha_{n,ia}$ ,  $i = 1, 2, 3$ , defined as  $\alpha_{n,i}$  with  $|\xi_p - p|$  replacing  $(\xi_p - p)$ . In Theorem 2.2 we assume  $0 \leq \beta < \infty$ . We consider  $\gamma_{n,0} = \xi_p(n, a_n)$  and the maxima

$$\begin{aligned}\gamma_{n,1} &= \max_{a_n \leq k \leq n} \xi_p(n, k), \\ \gamma_{n,2} &= \max_{0 \leq k \leq n - a_n} \xi_p(k + a_n, a_n)\end{aligned}$$

and

$$\gamma_{n,3} = \max_{\substack{0 \leq j < k \leq n \\ a_n \leq k-j}} \xi_p(k, k-j).$$

Figure 1 indicates the relationship among the various maxima we have in mind.

**THEOREM 2.1.** *Suppose  $1 \leq a_n \leq n$ ,  $a_n/\log n \rightarrow \infty$  as  $n \rightarrow \infty$ ,  $r = \liminf_{n \rightarrow \infty} (\log(n/a_n)/\log \log n)$  and that  $r/(r+1) = 1$  for  $r = \infty$ . Then, with probability one,  $\liminf_{n \rightarrow \infty} \alpha_n$  and  $\limsup_{n \rightarrow \infty} \alpha_n$  are as indicated in the following table:*

	$\liminf_{n \rightarrow \infty}$	$\limsup_{n \rightarrow \infty}$	
$\alpha_{n,0}$	-1	1	(a)
$\alpha_{n,1}$	$= F(a),$ $\in [-1, F(a)],$	1	(b)
$\alpha_{n,1a}$	0	1	(c)
$\alpha_{n,2}$ and $\alpha_{n,3}$	$= \left(\frac{r}{r+1}\right)^{1/2},$ $\in [-1, 0],$ $= G(a),$	1	(d)
$\alpha_{n,2a}$ and $\alpha_{n,3a}$	$\left(\frac{r}{r+1}\right)^{1/2}$	1	(e)
	(1)	(2)	

where

$$F(a) = \begin{cases} 0, & a = 0, \\ -\left(1 - \frac{\log a}{4}\right)^{-1/2}, & 0 < a \leq 1, \end{cases}$$

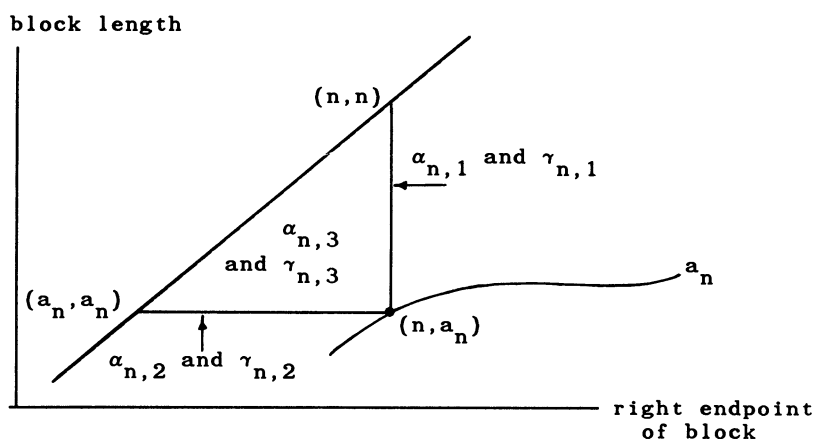


FIG. 1.

and

$$G(a) = \begin{cases} 0, & a = 0, \\ -\left(\frac{(2n+1)a-1}{n(n+1)a}\right)^{1/2}, & \frac{1}{n+1} < a \leq \frac{1}{n}, n = 1, 2, \dots \end{cases}$$

**THEOREM 2.2.** Suppose  $1 \leq a_n \leq n$  and  $a_n/\log n \rightarrow \beta \in [0, \infty)$ , as  $n \rightarrow \infty$ . Then, with probability one,  $\liminf_{n \rightarrow \infty} \gamma_n$  and  $\limsup_{n \rightarrow \infty} \gamma_n$  are as indicated in the following table:

	$\beta = 0$		$0 < \beta < \infty$		
	$\liminf_{n \rightarrow \infty}$	$\limsup_{n \rightarrow \infty}$	$\liminf_{n \rightarrow \infty}$	$\limsup_{n \rightarrow \infty}$	
$\gamma_{n,0}$	0	1	$1 - \psi_{1-p}^{-1}(\beta)$	$\psi_p^{-1}(\beta)$	(f)
$\gamma_{n,1}$	$p$	1	$p$	$\psi_p^{-1}(\beta)$	(g)
$\gamma_{n,2}$ and $\gamma_{n,3}$	1	1	$\psi_p^{-1}(\beta)$	$\psi_p^{-1}(\beta)$	(h)
	(1)	(2)	(3)	(4)	

where  $\psi_p^{-1}$  denotes the inverse of the function

$$\psi_p(\lambda) = -\left(\log\left(\left(\frac{\lambda}{p}\right)^p \left(\frac{1-\lambda}{1-p}\right)^{1-p}\right)\right)^{-1}, \quad p < \lambda < 1.$$

**REMARKS.** When  $\lim_{n \rightarrow \infty} a_n/\log n = \beta$ , statements (a) and (f) say that “ $\beta = \infty$ ” is necessary and sufficient for (1.2) to hold. Convergence rates in this case are given by (a). The oscillatory behavior of  $\xi_p(n, a_n)$  (when  $0 \leq \beta < \infty$ ) is indicated by (f). Note that  $\psi_p^{-1}(\beta) \rightarrow p$  as  $\beta \rightarrow \infty$  and that  $\psi_p^{-1}(\beta) \rightarrow 1$  as  $\beta \rightarrow 0$ . When  $a_n = [\beta \log n]$ ,  $0 < \beta < \infty$ , (3h) and (4h) applied to  $\gamma_{n,2}$  give Book and Truax’s

“Erdős–Rényi law for sample quantiles” in [3]. In (1b) and (1d),  $\liminf$  is not given when  $a_n/n$  diverges. However, it can be shown by example that “ $\liminf$ ” is not determined by  $\liminf a_n/n = b_1$  and  $\limsup a_n/n = b_2$  when  $b_1 \neq b_2$ , without additional assumptions on  $a_n$ .

**3. Proofs.** Throughout this section  $\#[A]$  denotes the number of elements in the set  $A$ , and  $C$  denotes various positive constants whose exact values do not matter so that, for example,  $1 + C = C$  might appear in this notation. For notational convenience we define for  $3 \leq k \leq n$ ,  $\phi^{-1}(n, k) = (\phi(n, k))^{-1}$  and  $d(n, k) = \{2k(\log(n/k) + \log \log k)\}^{1/2}$ , and for  $0 \leq u < v \leq 1$ ,  $B_{nk}(u, v) = \#[i: n - k + 1 \leq i \leq n, u \leq U_i < v]$ .

**LEMMA 3.1.** *If  $1 \leq a_n \leq n$  and  $a_n/\log n \rightarrow \infty$  as  $n \rightarrow \infty$ , then*

$$P\left\{\limsup_{n \rightarrow \infty} \max_{a_n \leq k \leq n} |B_{nk}(0, p) - kp|k^{-1}\phi(n, k) = 1\right\} = 1.$$

**PROOF.** For  $n \geq 1$  define

$$(3.1) \quad Y_n = \frac{p - B_{n1}(0, p)}{(p(1-p))^{1/2}} \quad \text{and} \quad S_n = Y_1 + \cdots + Y_n.$$

Since  $|B_{nk}(0, p) - kp|k^{-1}\phi(n, k) = |S_n - S_{n-k}|/d(n, k)$  for  $3 \leq k \leq n$ , the above lemma follows from Theorem 5.1 of [5].  $\square$

For  $\lambda > 0$  and  $3 \leq k \leq n$  define the events

$$\Lambda_{nk\lambda} = \left\{|B_{nk}(p, p + \lambda\phi^{-1}(n, k)) - k\lambda\phi^{-1}(n, k)| > 4(k\lambda\phi^{-1}(n, k)\log n)^{1/2}\right\},$$

$$\Lambda'_{nk\lambda} = \left\{|B_{nk}(p - \lambda\phi^{-1}(n, k), p) - k\lambda\phi^{-1}(n, k)| > 4(k\lambda\phi^{-1}(n, k)\log n)^{1/2}\right\}$$

and

$$E_\lambda = \left(\limsup_{n \rightarrow \infty} \bigcup_{k=a_n}^n \Lambda_{nk\lambda}\right) \cup \left(\limsup_{n \rightarrow \infty} \bigcup_{k=a_n}^n \Lambda'_{nk\lambda}\right) = E_{\lambda 1} \cup E_{\lambda 2}.$$

**LEMMA 3.2.** *Suppose that  $\lambda > 0$ . If  $1 \leq a_n \leq n$  and  $a_n/\log n \rightarrow \infty$  as  $n \rightarrow \infty$ , then  $P(E_\lambda) = 0$ .*

**PROOF.** We will prove that  $P(E_{\lambda 1}) = 0$ . The proof that  $P(E_{\lambda 2}) = 0$  is similar and is omitted. If  $3 \leq k \leq n$ , then by Bernstein's inequality (see [9])

$$(3.2) \quad \begin{aligned} P(\Lambda_{nk\lambda}) &\leq 2 \exp \left\{ \frac{-16k\lambda\phi^{-1}(n, k)\log n}{2k\lambda\phi^{-1}(n, k) + \frac{8}{3}(k\lambda\phi^{-1}(n, k)\log n)^{1/2}} \right\} \\ &\leq 2 \exp \left\{ -\frac{1}{2} \min(8 \log n, 6(k\lambda\phi^{-1}(n, k)\log n)^{1/2}) \right\}. \end{aligned}$$

Since  $a_n/\log n \rightarrow \infty$ , we have  $(k\lambda\phi^{-1}(n, k)\log n)^{1/2}(\log n)^{-1} \rightarrow \infty$  uniformly in

$k$  for  $a_n \leq k \leq n$ . Thus from (3.2),  $P(\Lambda_{nk\lambda}) \leq 2n^{-4}$  for  $a_n \leq k \leq n$  and  $n$  sufficiently large so that

$$\sum_{n=3}^{\infty} P\left(\bigcup_{k=a_n}^n \Lambda_{nk\lambda}\right) \leq C + \sum_{n=3}^{\infty} (n - a_n)n^{-4} < \infty. \quad \square$$

**LEMMA 3.3.** *If  $1 \leq a_n \leq n$  and  $a_n/\log n \rightarrow \infty$  as  $n \rightarrow \infty$ , then*

$$(3.3) \quad \xi_p(n, k) = 2p - k^{-1}B_{nk}(0, p) + e_{nk}, \quad 1 \leq k \leq n,$$

where as  $n \rightarrow \infty$ ,  $\max_{a_n \leq k \leq n} |e_{nk}| \phi(n, k) = o(1)$  a.s.

**PROOF.** Fix  $0 < \delta < 1$ , an integer  $M > 2/\delta$ , and  $\omega \in D = \bigcap_{j=1}^{M+1} (E_{j\delta}^c \cap A)$ , where  $A$  is the event in Lemma 3.1. By Lemmas 3.1 and 3.2,  $P(D) = 1$ . Since  $k\phi^{-1}(n, k)/(k\phi^{-1}(n, k)\log n)^{1/2} \rightarrow \infty$  as  $n \rightarrow \infty$ , uniformly in  $k$  for  $a_n \leq k \leq n$ , we have for all large  $n$ ,  $n > n_0$ , and  $a_n \leq k \leq n$ ,

$$(3.4) \quad |B_{nk}(0, p) - kp| \leq 2k\phi^{-1}(n, k),$$

$$(3.5) \quad \omega \notin \bigcup_{j=1}^{M+1} (\Lambda_{n, k, j\delta} \cup \Lambda'_{n, k, j\delta})$$

and

$$(3.6) \quad \delta k\phi^{-1}(n, k) > 4((M+1)\delta k\phi^{-1}(n, k)\log n)^{1/2}.$$

Fix  $n > n_0$  and suppose that  $a_n \leq k \leq n$ . By (3.4) there exists  $1 \leq j_{nk} = j \leq M$  such that either

$$(3.7) \quad kp - j\delta k\phi^{-1}(n, k) \leq B_{nk}(0, p) < kp - (j-1)\delta k\phi^{-1}(n, k)$$

or

$$(3.7a) \quad kp + (j-1)\delta k\phi^{-1}(n, k) \leq B_{nk}(0, p) < kp + j\delta k\phi^{-1}(n, k).$$

Suppose that (3.7) holds [the proof is the same if (3.7a) holds]. Then, by (3.5) and (3.6)

$$\begin{aligned} & B_{nk}(0, p + (j+1)\delta\phi^{-1}(n, k)) \\ &= B_{nk}(0, p) + B_{nk}(p, p + (j+1)\delta\phi^{-1}(n, k)) \\ &\geq kp - j\delta k\phi^{-1}(n, k) + (j+1)\delta k\phi^{-1}(n, k) \\ &\quad - 4((j+1)\delta k\phi^{-1}(n, k)\log n)^{1/2} \\ &\geq kp + \delta k\phi^{-1}(n, k) - 4((M+1)\delta k\phi^{-1}(n, k)\log n)^{1/2} > kp, \end{aligned}$$

so that  $\xi_p(n, k) \leq p + (j+1)\delta\phi^{-1}(n, k)$ . By (3.3) and (3.7)

$$\begin{aligned} e_{nk} &= \xi_p(n, k) - 2p + k^{-1}B_{nk}(0, p) \\ &\leq (j+1)\delta\phi^{-1}(n, k) - p + k^{-1}(kp - (j-1)\delta k\phi^{-1}(n, k)) \\ &= 2\delta\phi^{-1}(n, k). \end{aligned}$$

From a similar calculation,  $e_{nk} \geq -2\delta\phi^{-1}(n, k)$ . Since  $\delta$ ,  $k$  and  $\omega$  are arbitrary, the lemma follows.  $\square$

**PROOF OF (1a) AND (2a).** With  $S_n$  defined as in (3.1), Lemma 3.3 says that the collection of limit points,  $L$ , of  $\alpha_{n,0}$  is almost surely that of  $(S_n - S_{n-a_n})/d(n, a_n)$ . By Theorem 1 of [8] we can define a standard Wiener process  $\{W(T), T \geq 0\}$  and the sequence  $Y_1, Y_2, \dots$  on a new probability space (say  $\Omega'$ ) so as to have  $S_n = W(n) + O(\log n)$  a.s. Since  $a_n/\log n \rightarrow \infty$  we get  $(\log n)/d(n, a_n) \rightarrow 0$ , so that the collection of limit points of  $\alpha_{n,0}$  is almost surely that of  $(W(n) - W(n - a_n))/d(n, a_n)$ . By Corollary 2.3 of [6],  $P\{L = [-1, 1]\} = 1$ .  $\square$

**PROOF OF (2b)–(2e).** With  $S_n$  defined as in (3.1) it follows by Lemma 3.3 and Theorem 5.3 of [5] that

$$\limsup_{n \rightarrow \infty} \alpha_{n,3a} = \limsup_{n \rightarrow \infty} \max_{\substack{0 \leq j < k \leq n \\ a_n \leq k-j}} \frac{|S_k - S_j|}{d(k, k-j)} = 1 \quad \text{a.s.}$$

Thus, statements (2b)–(2e) follow from (2a) since  $\alpha_{n,0} \leq \alpha_{n,i} \leq \alpha_{n,3a}$  for  $i = 1, 1a, 2, 2a, 3$ .  $\square$

**PROOF OF (1b)–(1e).** Again, we go to  $\Omega'$ . By Lemma 3.3 and Theorem 1 of [8] we have, as in the proof of (1a) and (2a),

$$\liminf_{n \rightarrow \infty} \alpha_{n,2} = \liminf_{n \rightarrow \infty} \max_{0 \leq k \leq n-a_n} \frac{W(k + a_n) - W(k)}{d(k + a_n, a_n)} \quad \text{a.s.}$$

By Theorem 2 of [7], to prove our result for  $\liminf_{n \rightarrow \infty} \alpha_{n,2}$  it suffices to prove that

$$\begin{aligned} (3.8) \quad & \liminf_{n \rightarrow \infty} \max_{0 \leq k \leq n-a_n} \frac{W(k + a_n) - W(k)}{d(k + a_n, a_n)} \\ &= \liminf_{T \rightarrow \infty} \max_{0 \leq t \leq T-a_T} \frac{W(t + a_T) - W(t)}{d(t + a_T, a_T)} \quad \text{a.s.,} \end{aligned}$$

where  $a_T = a_{[T]}$  for  $T \geq 1$  and  $a_T = T$  for  $0 < T < 1$ . For  $T > 1$  and  $0 \leq t \leq T - a_T$ ,

$$\begin{aligned} & \frac{W([t] + a_{[T]}) - W([t])}{d([t] + a_{[T]}, a_{[T]})} \\ &= \frac{W(t + a_T) - W(t)}{d(t + a_T, a_T)} \frac{d(t + a_T, a_T)}{d([t] + a_{[T]}, a_{[T]})} \\ & \quad - \frac{W(t + a_T) - W([t] + a_{[T]})}{d([t] + a_{[T]}, a_{[T]})} + \frac{W(t) - W([t])}{d([t] + a_{[T]}, a_{[T]})} \\ &= \left( \frac{W(t + a_T) - W(t)}{d(t + a_T, a_T)} \right) A(t, T) + B(t, T) + C(t, T). \end{aligned}$$

By Theorem of 3.3B of [5], if  $M_T \rightarrow \infty$ , then

$$(3.9) \quad \lim_{T \rightarrow \infty} \max_{\substack{0 \leq s < t \leq T \\ |t-s| \leq 1}} \frac{|W(t) - W(s)|}{(M_T \log T)^{1/2}} = 0 \quad \text{a.s.}$$

Now,  $A(t, T) \rightarrow 1$  uniformly in  $t$ ,  $0 \leq t \leq T - a_T$ , as  $T \rightarrow \infty$ . By (3.9), with probability one,  $B(t, T) \rightarrow 0$  and  $C(t, T) \rightarrow 0$  uniformly in  $t$ ,  $0 \leq t \leq T - a_T$ , as  $T \rightarrow \infty$ . This proves (3.8). A similar proof (which also uses Theorems 1 and 3 of [7]) yields the remaining lim inf statements of Theorem 2.1.  $\square$

**PROOF OF (4g) AND (4h).** The following relationship is evident from Figure 1 or Lemma 2.1 of [5]:

$$(3.10) \quad \limsup_{n \rightarrow \infty} \gamma_{n,2} \leq \limsup_{n \rightarrow \infty} \gamma_{n,3} = \limsup_{n \rightarrow \infty} \gamma_{n,1}.$$

(Note: To use Lemma 2.1, one needs to replace the original sequence,  $a_n$ , by a nondecreasing sequence.) Define for  $p < \lambda < 1$ ,

$$h(\lambda) = (\lambda/p)^p ((1-\lambda)/(1-p))^{1-p}.$$

By Theorem 1 of [1] (see also Lemma 1 of [2]) there exist constants  $0 < C_{1\lambda} < C_{2\lambda} < \infty$  such that for all large  $k$  and  $n \geq k$ ,

$$(3.11) \quad C_{1\lambda} k^{-1/2} h(\lambda)^k \leq P\{\xi_p(n, k) > \lambda\} \leq C_{2\lambda} k^{-1/2} h(\lambda)^k.$$

Fix  $\psi_p^{-1}(\beta) < \lambda < 1$ . By (3.11), for  $n$  sufficiently large

$$(3.12) \quad \begin{aligned} P(G_n) &= P\left(\bigcup_{k=a_n}^n \{\xi_p(n, k) > \lambda\}\right) \leq C \sum_{k=a_n}^n h(\lambda)^k \\ &= C \sum_{k=a_n}^n \exp(-k/\psi_p(\lambda)) \leq C \int_{a_n-1}^n \exp(-x/\psi_p(\lambda)) dx \\ &\leq C \exp(-a_n/\psi_p(\lambda)) \leq C n^{-((a_n/\psi_p(\lambda))/\log n)}. \end{aligned}$$

Since  $0 < \psi_p(\lambda) < \beta$  and  $a_n/\log n \rightarrow \beta$ , it follows from (3.12) that  $\sum_{n=1}^{\infty} P(G_n) < \infty$ . Since  $\lambda$  is arbitrary we get  $\limsup_{n \rightarrow \infty} \gamma_{n,1} \leq \psi_p^{-1}(\beta)$  almost surely. Thus, (4g) and (4h) follow from (3.10), (3.16) in the proof of (4f), and the observation that  $\gamma_{n,i} \geq \gamma_{n,0}$  for  $1 \leq i \leq 3$ .  $\square$

**PROOF OF (3h).** Fix  $p < \lambda < \psi_p^{-1}(\beta)$ . By (3.11), for  $n$  sufficiently large

$$\begin{aligned} P(J_n) &= P\left(\bigcap_{j=1}^{[n/a_n]} \{\xi_p(ja_n, a_n) < \lambda\}\right) \\ &\leq \prod_{j=1}^{[n/a_n]} (1 - C a_n^{-1/2} h(\lambda)^{a_n}) \\ &\leq \exp(-C a_n^{-1/2} h(\lambda)^{a_n} [n/a_n]) \\ &\leq \exp(-C a_n^{-3/2} n^{1-((a_n/\psi_p(\lambda))/\log n)}). \end{aligned}$$

Since  $\psi_p(\lambda) > \beta$  and  $a_n/\log n \rightarrow \beta$ ,  $\sum_{n=1}^{\infty} P(J_n) < \infty$ . Since  $\lambda$  is arbitrary, (3h) follows from (4h).  $\square$

**PROOF OF (1g) AND (3g).** We will prove that

$$(3.13) \quad \text{if } 1 \leq a_n \leq n, \text{ then } \liminf_{n \rightarrow \infty} \gamma_{n,1} = p \text{ a.s.}$$

(Note: The assumption that  $\lim a_n/\log n$  exists is not needed in this result.) Fix  $0 < \delta < 1 < B$  so that  $p^{2[\delta \log n]} \geq n^{-1}$ , all  $n \geq 1$ , and  $\psi_p(p + \delta) < B$ . By (4g),  $\limsup_{n \rightarrow \infty} \max_{[B \log n] \leq k \leq n} \xi_p(n, k) \leq p + \delta$  almost surely, so that

$$(3.14) \quad \liminf_{n \rightarrow \infty} \gamma_{n,1} \leq \max \left( \liminf_{n \rightarrow \infty} \max_{1 \leq k \leq [B \log n]} \xi_p(n, k), p + \delta \right) \text{ a.s.}$$

Let  $n_0 \geq 1$  be such that  $(n_0 - 1)^2 < n_0^2 - [B \log n_0^2]$ , and for  $n \geq n_0$  define  $A_n = \{\max_{1 \leq k \leq [B \log n^2]} \xi_p(n^2, k) \leq p + \delta\}$ . We will show that

$$P\left(\limsup_{n \rightarrow \infty} A_n\right) = 1.$$

Let  $k'_n = [B \log n^2]/[\delta \log n]$  and suppose  $n \geq n_0$ . The set

$$\{n^2 - [B \log n^2] + 1, \dots, n^2\}$$

can be partitioned into  $k_n = [k'_n]$  groups,

$$g_j = \{h_n + (k_n - j)[\delta \log n] + 1, \dots, h_n + (k_n - j + 1)[\delta \log n]\},$$

for  $2 \leq j \leq k_n$  and  $g_1 = \{h_n + (k_n - 1)[\delta \log n] + 1, \dots, n^2\}$ , where  $h_n = n^2 - [B \log n^2]$ . Note that  $\#[g_j] = [\delta \log n]$  for  $2 \leq j \leq k_n$  and that  $[\delta \log n] \leq \#[g_1] < 2[\delta \log n]$ . Define  $\nabla_j = \{\#[i: i \in g_j, U_i \leq p + \delta] \geq p[\delta \log n]\}$  for  $2 \leq j \leq k_n$ . Note that  $P(\nabla_j) \geq 1/2$  and  $k_n \leq 3B/\delta$  for  $n$  sufficiently large, so by choice of  $\delta$  and  $B$ ,

$$(3.15) \quad \begin{aligned} P(A_n^*) &= P\left(\{U_i \leq p + \delta \text{ all } i \in g_1\} \cap \left(\bigcap_{j=2}^{k_n} \nabla_j\right)\right) \\ &\geq p^{2[\delta \log n]} (1/2)^{k_n} \geq n^{-1} (1/2)^{3B/\delta}, \end{aligned}$$

for  $n$  sufficiently large. For  $n \geq n_0$ ,  $A_n \supset A_n^*$  so by (3.15),  $\sum_{n=n_0}^{\infty} P(A_n) = \infty$ . Since  $A_{n_0}, A_{n_0+1}, \dots$  are independent,  $P(\limsup_{n \rightarrow \infty} A_n) = 1$  as asserted. Now, since  $\delta$  was arbitrary, (3.13) follows from (3.14), and the observation that  $\gamma_{n,1} \geq \xi_p(n, n) \rightarrow p$  almost surely.  $\square$

**PROOF OF (3f) AND (4f).** Let  $K$  denote the collection of limit points of the sequence  $\gamma_{n,0}$ . We first prove that

$$(3.16) \quad P\{K \supset [p, \psi_p^{-1}(\beta)]\} = 1.$$

Suppose that  $p < \lambda_1 < \lambda_2 < \psi_p^{-1}(\beta)$ . Fix  $\varepsilon > 0$ , and for  $n \geq 1$  let  $y(n) = [n^{1+\varepsilon}]$ . For some  $n_0 \geq 1$ ,  $\xi_p(y(n_0), a_{y(n_0)}), \xi_p(y(n_0 + 1), a_{y(n_0 + 1)}), \dots$  are independent random variables. The function  $h$  defined in the proof of (4g) and (4h) is



decreasing, so by (3.11), for  $n$  sufficiently large,

$$\begin{aligned} P\{\lambda_1 < \xi_p(y(n), a_{y(n)}) \leq \lambda_2\} &\geq C(\log n)^{-1/2} (h(\lambda_1))^{a_{y(n)}} \\ &= C(\log n)^{-1/2} n^{-(a_{y(n)}/a_n)((a_n/\psi_p(\lambda_1))/\log n)} \geq Cn^{-1}, \end{aligned}$$

for  $\varepsilon$  sufficiently small since  $\psi_p(\lambda_1) > \beta$ ,  $a_n/\log n \rightarrow \beta$  and  $a_{y(n)}/a_n \rightarrow 1 + \varepsilon$ . Thus, (3.16) holds. Now, consider the sequence  $U'_1 = 1 - U_1$ ,  $U'_2 = 1 - U_2, \dots$ . Let  $\xi'_{1-p}(n, a_n)$  denote a  $(1-p)$ th quantile [use (1.1) with " $\pi$ " =  $1 - \pi$ ] of the subsequence  $U'_{n-a_n+1}, U'_{n-a_n+2}, \dots, U'_n$  and let  $K'$  denote the collection of limit points of  $\xi'_{1-p}(n, a_n)$ . By applying (3.16) to the  $U'_n$  sequence we get  $P\{K' \supset [1 - p, \psi_{1-p}^{-1}(\beta)]\} = 1$ , or equivalently

$$(3.17) \quad P\{K \supset [1 - \psi_{1-p}^{-1}(\beta), p]\} = 1.$$

By (3.16), (3.17) and an application of (4g) to the  $U_n$  and  $U'_n$  sequences we get  $P\{K = [1 - \psi_{1-p}^{-1}(\beta), \psi_p^{-1}(\beta)]\} = 1$ .  $\square$

**PROOF OF (1f) AND (2f).** Define  $K$  as in the preceding proof. Suppose  $0 \leq a < b \leq 1$ . Define for  $n \geq 1$ ,  $M_n = \{a < U_i < b, \text{ all } n^2 - a_{n^2} + 1 \leq i \leq n^2\}$ . For  $n$  sufficiently large

$$P(M_n) = (b - a)^{a_{n^2}} = (b - a)^{(a_{n^2}/\log n^2) \log n^2} \geq n^{-1},$$

since  $a_{n^2}/\log n^2 \rightarrow 0$ . The events  $M_{n_0}, M_{n_0+1}, \dots$  are independent for some  $n_0 \geq 1$ , so that  $P\{K = [0, 1]\} = 1$ .  $\square$

**PROOF OF (1h), (2g) AND (2h).** Statements (2g) and (2h) follow from (2f). To prove (1h) fix  $0 < \delta < 1$  and define  $H_n = \bigcup_{j=1}^{k_n} \{1 - \delta < U_i < 1 \text{ all } (j-1)a_n < i \leq ja_n\}$  for  $n \geq 1$  where  $k_n = [n/a_n]$ . Since  $a_n/\log n \rightarrow 0$ , we have for  $n$  sufficiently large

$$P(H_n^c) = (1 - \delta^{a_n})^{k_n} \leq \exp(-\delta^{a_n} k_n) \leq \exp(-n^{1/2}),$$

so that  $\sum_{n=1}^{\infty} P(H_n^c) < \infty$ . Thus,  $P(\liminf_{n \rightarrow \infty} H_n) = 1$  (note that this fact follows also from well-known results on the longest head run in Bernoulli trials). Now note that  $\{\gamma_{n,3} \geq 1 - \delta\} \supset \{\gamma_{n,2} \geq 1 - \delta\} \supset H_n$  for  $n \geq 1$ , and that  $\delta$  is arbitrary.  $\square$

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## REFERENCES

- [1] BAHADUR, R. R. and RANGA RAO, R. (1960). On deviations of the sample mean. *Ann. Math. Statist.* **31** 1015-1027.
- [2] BOOK, S. A. (1971). Large deviation probabilities for order statistics. *Naval Res. Logist. Quart.* **18** 521-523.
- [3] BOOK, S. A. and TRUAX, D. R. (1976). An Erdős-Rényi strong law for sample quantiles. *J. Appl. Probab.* **13** 578-583.
- [4] CSÖRGŐ, S. (1979). Erdős-Rényi laws. *Ann. Statist.* **7** 772-787.

- [5] HANSON, D. L. and RUSSO, R. P. (1983). Some results on increments of the Wiener process with applications to lag sums of I.I.D. random variables. *Ann. Probab.* **11** 609–623.
- [6] HANSON, D. L. and RUSSO, R. P. (1983). Some more results on increments of the Wiener process. *Ann. Probab.* **11** 1009–1015.
- [7] HANSON, D. L. and RUSSO, R. P. (1985). Some “liminf” results for increments of a Wiener process. Technical Report 113, Dept. Statistics, Univ. Iowa.
- [8] KOMLÓS, J., MAJOR, P. and TUSNÁDY, G. (1976). An approximation of partial sums of independent r.v.’s and the sample d.f. II. *Z. Wahrsch. verw. Gebiete* **34** 33–58.
- [9] SERFLING, R. J. (1980). *Approximation Theorems of Mathematical Statistics*. Wiley, New York.

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